

# Diabetes Incidence and Long-Term Exposure to Air Pollution

## A cohort study

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**OBJECTIVE**—Animal and cross-sectional epidemiological studies suggest a link between air pollution and diabetes, whereas the limited prospective data show mixed results. We studied the association between long-term exposure to traffic-related air pollution and incidence of diabetes.

**RESEARCH DESIGN AND METHODS**—We followed 57,053 participants of the Danish Diet, Cancer, and Health cohort in the Danish National Diabetes Register between baseline (1993–1997) and 27 June 2006. We estimated the mean levels of nitrogen dioxide (NO<sub>2</sub>) at the residential addresses of the cohort participants since 1971 and modeled the association between NO<sub>2</sub> and diabetes incidence with a Cox regression model, separately for two definitions of diabetes: all cases and a more strict definition where unconfirmed cases were excluded.

**RESULTS**—Over a mean follow-up of 9.7 years of 51,818 eligible subjects, there were 4,040 (7.8%) incident diabetes cases in total and 2,877 (5.5%) with confirmed diagnoses. Air pollution was not associated with all diabetes cases (hazard ratio 1.00 [95% CI 0.97–1.04] per interquartile range of 4.9 μg/m<sup>3</sup> mean NO<sub>2</sub> levels since 1971), but a borderline statistically significant association was detected with confirmed cases of diabetes (1.04 [1.00–1.08]). Among confirmed diabetes cases, effects were significantly enhanced in nonsmokers (1.12 [1.05–1.20]) and physically active people (1.10 [1.03–1.16]).

**CONCLUSIONS**—Long-term exposure to traffic-related air pollution may contribute to the development of diabetes, especially in individuals with a healthy lifestyle, nonsmokers, and physically active individuals.

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The prevalence and incidence of type 2 diabetes are rising rapidly in both the developed and developing world, presenting one of the greatest contributors to the global burden of the disease (1). The diabetes epidemic is in large part attributable to established causes related

to modern lifestyle including obesity, physical inactivity, and the growing aging populations (1). Environmental exposures linked to industrialization and urbanization, such as air pollution, have not been considered risk factors for diabetes until recently (2). In the U.S., the prevalence of

diabetes correlated with the release of toxicants into the air (3), whereas diabetic people appeared more vulnerable than nondiabetic people to cardiovascular health effects associated with exposure to air pollution (4). Diabetes and cardiovascular diseases share many risk factors, and diabetic people are at a highly increased risk for heart or circulatory disorders (5). The central biological mechanisms of air pollution damage to the heart and blood vessels include inflammation (6), which is also believed to be involved in the promotion of insulin resistance and type 2 diabetes (7). An enhanced association between air pollution and inflammation, endothelial dysfunction, prothrombotic changes, and altered heart rate variability was found in diabetic people (6). A plausible biological mechanism of air pollution promoting diabetes was provided by Sun et al. (8), showing that exposure to particulate air pollution caused increased blood glucose, inflammation in adipose tissue, and insulin resistance in high-fat diet-fed mice. Recent study confirmed that prolonged exposure to air pollution leads to insulin resistance and impaired glucose tolerance in rats and that this association is not limited to high-fat diet rats (9). Epidemiological evidence is sparse. Short-term exposure to air pollution was linked to exacerbations of diabetes leading to death (10–14) and hospitalizations (13). Prevalence of diabetes was linked to air pollution (14,15). Two prospective diabetes studies investigated the link with air pollution, with one reporting significant associations among a small number ( $n = 87$ ) of women (16) and another failing to detect association in two large cohorts, except with a single traffic proximity proxy in women (17). Limited and mixed evidence precludes conclusions about causality between air pollution and diabetes and merits more study.

We studied the association between traffic-related air pollution levels at the residence and the risk for diabetes in an elderly Danish cohort and tested for an effect modification by lifestyle, education, and comorbid conditions.

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## RESEARCH DESIGN AND METHODS

### The Danish Diet, Cancer, and Health Cohort

The Danish Diet, Cancer, and Health (DCH) cohort (18) consists of 57,053 people from Copenhagen or Aarhus aged 50–65 years, recruited from 1993 to 1997. We linked the cohort members to the 1) Danish National Diabetes Register (NDR) for diabetes diagnoses; 2) the Central Population Registry to obtain date of death or emigration and detailed residential address history (1971–2006); 3) the Danish Address Database to obtain geographical coordinates of the addresses; and 4) the National Patient Register (dating back to 1976) to obtain data on hospital admissions for chronic obstructive pulmonary disease (COPD) (ICD-10, J40–44, and ICD-8, 490–92) and asthma (ICD-10, J45–46, and ICD-8, 493) before diabetes diagnoses.

### The Danish NDR

The NDR was established in 2006 by the National Board of Health to describe and monitor the occurrence of diabetes in Denmark and provide data for epidemiological research (19). The NDR was established by linking existing nationwide administrative records in the Danish health care system: the National Patient Register, containing hospital discharge diagnoses since 1994; the National Health Insurance Registry, with information of all services provided by general and specialist practitioners since 1973; and the Register of Medicinal Product Statistics, containing all prescriptions dispensed at Danish pharmacies since 1993. Inclusion criteria for the NDR were as follows: diabetes hospital discharge diagnoses in the National Patient Register defined as ICD-10: DE10–14, DH36.0, DO24, or ICD-8: 249, 250; chiropody for diabetic patients, five blood glucose measurements within 1 year, or two blood glucose measurements per year for 5 consecutive years, registered in the National Health Insurance Registry; or second purchase of insulin or oral glucose-lowering drugs within 6 months, registered in the Register of Medicinal Product Statistics. The results of blood glucose measurements are not included in the NDR. Between 50 and 60% of the patients in NDR met more than one inclusion criterion. It was not possible to distinguish between type 1 and type 2 diabetes from NDR. Because of different dates of initiation of the underlying registers and accumulation of prevalent cases, only

incidence values after 1 January 1995 were found to be reliable (19). Thus, the incidence of diabetes was defined as the earliest record in the diabetes register occurring after 1 January 1995, between baseline (1993–1997) and 27 June 2006. In addition to the original NDR definition including all cases, we defined a more strict incidence definition, including only confirmed diabetes, by excluding individuals who were included in the NDR solely because of a blood glucose test, since a number of these people probably did not have diabetes (19).

### Exposure assessment

The Danish AirGIS human exposure modeling system (20) (described in the Supplementary Data) is a validated model (20) used to estimate outdoor levels of nitrogen dioxide (NO<sub>2</sub>) and nitrogen oxide (NO<sub>x</sub>) at the residential addresses of the cohort members, who had residential history information available for more than 80% of the time between 1971 and 27 June 2006 (97.4%). Missing values because of missing address or missing geographical coordinates were substituted by the levels calculated for the preceding address or, when the first address was missing, for the subsequent address, giving a complete series of annual mean NO<sub>2</sub> and NO<sub>x</sub> concentrations since 1971. Mean NO<sub>2</sub> levels since 1971 (calculated for each year of follow-up) was the main air pollution exposure proxy, but we also included the same measures for NO<sub>x</sub> (presented in the Supplementary Data), mean NO<sub>2</sub> levels since 1991, 1-year mean NO<sub>2</sub> at baseline, and 1-year mean NO<sub>2</sub> at follow-up, traffic proximity (presence of a major road (density ≥10,000 vehicles/day) within a 50-m radius of the baseline residence, and traffic load (total kilometers driven by vehicles) within a 100-m radius of a baseline residence).

### Follow-up

Cohort members were followed from date of recruitment or 1 January 1995 (start of the NDR) (whichever was last) until date of diabetes, as recorded in NDR, death, emigration, or 27 June 2006 (whichever came first). This was done separately for the two definitions of diabetes.

### Statistical analysis

We used a Cox proportional hazards model to model the diabetes incidence, with age as the underlying timescale. The effects of exposure to NO<sub>2</sub> on diabetes incidence was modeled as time-dependent variables and evaluated in several steps,

with adjustment for a priori defined confounders: 1) adjusted only for age; 2) additionally adjusted for year (restricted cubic spline), to model the increase in diabetes incidence over follow-up time, and full adjustment for recognized diabetes risk factors: sex, BMI, waist-to-hip ratio, smoking status, smoking duration, smoking intensity, environmental tobacco smoke, educational level, leisure-time physical activity, alcohol intake, fruit consumption, and fat consumption; and 3) additionally adjusted for self-reported hypertension, hypercholesterolemia, and previous myocardial infarction (MI), which are risk factors for diabetes but are also associated with air pollution (6) and thus suspected to be on the biological pathway from air pollution to diabetes. Confounder definition is provided in the Supplementary Data. The potential modifiers of an effect between NO<sub>2</sub> and diabetes by selected baseline confounders and comorbid conditions were evaluated by introducing interaction terms into the model and tested by the Wald test. Comorbidity with cardiovascular disease (hypertension, hypercholesterolemia, and MI) was self-reported at baseline, whereas comorbidity with COPD and asthma were defined by hospitalization for COPD or asthma, respectively, before baseline. The results are presented as hazard ratios (HRs) with 95% CIs, per interquartile range (IQR) increase (4.9 μg/m<sup>3</sup>), estimated with survival library (R 2.9.0 statistical software). The graphical presentation of a functional form of an association between NO<sub>2</sub> and confirmed diabetes was produced using restricted cubic spline in the design library (R 2.9.0 statistical software).

**RESULTS**—Among the 57,053 cohort members, 571 were excluded because of a cancer diagnosis before enrollment, 2 with an uncertain date of cancer diagnosis, 960 with missing baseline address or missing geocode for baseline address, 1,341 with NO<sub>2</sub> annual levels available for <80% of the follow-up, 1,147 with self-reported diabetes at baseline, 173 with a diabetes record in the NDR before baseline, 6 with diabetes diagnoses from NDR between baseline and 1 January 1995, and 1,035 with missing information on covariates. Of the 51,818 eligible subjects followed over a mean of 9.7 years, 4,040 (7.8%) developed diabetes in total (incidence rate 8.0 per 1,000 person-years), among which 1,163 included solely because of blood glucose measurements were excluded for a more strict definition of diabetes, resulting

in 2,877 (5.5%) diabetes cases (incidence rate 5.7 per 1,000 person-years).

Diabetic people were older; had higher BMI, waist-to-hip ratio, and alcohol and fat intake; had lower education; consumed less fruit; and were more likely to be men, current or previous smokers, exposed to environmental tobacco smoke, physically inactive, and have hypertension, hypercholesterolemia, or MI, but only marginally more likely to have COPD and asthma than nondiabetic people (Table 1). The estimated

levels of NO<sub>2</sub> varied greatly (Supplementary Fig. 1). Air pollution levels were higher for diabetic people than for the whole cohort (Table 1, Supplementary Table 1).

A statistically significant positive association was detected for both definitions of diabetes with all air pollution exposure proxies (except with an indicator of a major road within a 50-m radius of baseline residence for all diabetes) in models adjusted only for age, with a twice-as-strong effect for the confirmed compared with

all diabetes cases (Table 2). There was no association between diabetes incidence and air pollution after adjustment for confounders when including all diabetes cases. For confirmed diabetes cases, full adjustment for lifestyle resulted in attenuated but borderline statistically significant association with mean levels of NO<sub>2</sub> since 1971 (HR 1.04 [95% CI 1.00–1.108] per IQR of 4.9 μg/m<sup>3</sup>), which was insensitive to further adjustment for hypertension, hypercholesterolemia, and MI. The

**Table 1—Characteristics of the DCH cohort (n = 51,818) by incident diabetes status at follow-up, for two definitions of diabetes incidence based on the NDR**

	Total cohort	All diabetes*	Confirmed diabetes*
<i>n</i>	51,818	4,040	2,877
Baseline cohort covariates			
Age (years) [median (IQR)]	56.1 (7.6)	57.3 (7.7)	57.1 (7.8)
Older than 56 years at baseline [ <i>n</i> (%)]	26,414 (51.0)	2,362 (58.5)	1,655 (57.5)
Males [ <i>n</i> (%)]	24,545 (47.4)	2,282 (56.5)	1,702 (59.2)
BMI [median (IQR)]	25.5 (4.9)	28.6 (5.9)	29.1 (5.9)
Underweight (BMI <18.5 kg/m <sup>2</sup> ) [ <i>n</i> (%)]	406 (0.8)	9 (0.2)	6 (0.2)
Normal (18.5 ≤ BMI < 25 kg/m <sup>2</sup> ) [ <i>n</i> (%)]	22,475 (43.4)	751 (18.6)	401 (13.9)
Overweight (25 ≤ BMI < 30 kg/m <sup>2</sup> ) [ <i>n</i> (%)]	21,638 (41.8)	1,776 (44.0)	1,262 (43.9)
Obese (BMI ≥30 kg/m <sup>2</sup> ) [ <i>n</i> (%)]	7,299 (14.1)	1,504 (37.2)	1,208 (42.0)
Waist-to-hip ratio [median (IQR)]	0.88 (0.16)	0.95 (0.14)	0.96 (0.13)
Waist-to-hip ratio >0.90 male/>0.85 female [ <i>n</i> (%)]	25,635 (49.5)	3,007 (74.4)	2,283 (79.3)
Never smoked [ <i>n</i> (%)]	18,438 (35.6)	1,202 (29.7)	823 (28.6)
Previously smoked [ <i>n</i> (%)]	14,822 (28.6)	1,273 (31.5)	891 (31.0)
Currently smoke [ <i>n</i> (%)]	18,558 (35.8)	1,565 (38.7)	1,163 (40.4)
Smoking duration (years) [median (IQR)]	32 (19)	34 (17)	35 (16)
Smoking intensity (g/day) [median (IQR)]	6 (18)	7 (20)	9 (20)
Environmental tobacco smoke [ <i>n</i> (%)]	33,160 (64.0)	2,769 (68.5)	2,016 (70.1)
Consume alcohol [ <i>n</i> (%)]	50,682 (97.8)	3,919 (97.0)	2,782 (96.7)
Alcohol use (g/day) [median (IQR)]	13.4 (25.0)	13.3 (27.3)	13.5 (28.6)
Physically active or play sports in leisure time [ <i>n</i> (%)]	28,282 (54.6)	1,784 (44.2)	1,185 (41.2)
Physical activity (h/week) [median (IQR)]	2 (2)	2 (2)	2 (2)
<8 years of education [ <i>n</i> (%)]	16,961 (32.7)	1,618 (40.0)	1,209 (42.0)
8–10 years of education [ <i>n</i> (%)]	23,995 (46.3)	1,795 (44.4)	1,250 (43.5)
>10 years of education [ <i>n</i> (%)]	10,862 (21.0)	627 (15.5)	418 (14.5)
Fruit intake (g/day) [median (IQR)]	145.2 (165.3)	142.5 (164.4)	136.8 (162.1)
Fat intake (g/day) [median (IQR)]	81.2 (37.7)	81.9 (38.0)	82.2 (38)
Hypertension [ <i>n</i> (%)]	8,200 (15.8)	1,226 (30.3)	915 (31.8)
Hypercholesterolemia [ <i>n</i> (%)]	3,737 (7.2)	482 (11.9)	346 (12.0)
MI [ <i>n</i> (%)]	1,027 (2.0)	161 (4.0)	129 (4.5)
COPD† [ <i>n</i> (%)]	2,058 (4.0)	188 (4.6)	140 (4.9)
Asthma† [ <i>n</i> (%)]	1,273 (2.5)	99 (2.4)	78 (2.7)
Air pollution exposure			
NO <sub>2</sub> (μg/m <sup>3</sup> ), 1971 to end of follow-up [median (IQR)]	14.5 (4.9)	14.7 (5.3)	15.1 (5.4)
NO <sub>2</sub> (μg/m <sup>3</sup> ), 1991 to end of follow-up [median (IQR)]	15.3 (5.6)	15.7 (6.1)	15.7 (6.1)
NO <sub>2</sub> (μg/m <sup>3</sup> ), baseline (1 year) [median (IQR)]	15.4 (5.6)	15.7 (6.1)	16.3 (6.0)
NO <sub>2</sub> (μg/m <sup>3</sup> ), end of follow-up (1 year) [median (IQR)]	15.2 (5.7)	15.7 (6.1)	16.2 (6.0)
Traffic load‡ within 100 m at baseline (10 <sup>3</sup> vehicle km/day) [median (IQR)]	0.34 (1.3)	0.40 (1.4)	0.46 (1.6)
Major road§ within 50 m at baseline [ <i>n</i> (%)]	4,184 (8.1)	373 (9.2)	287 (10.0)

IQR, difference between 75th and 25th percentile. \*All diabetes: based on original inclusion criteria in the NDR: hospital admission for diabetes, diabetes medication, reimbursement for chiropody due to diabetes, or glucose blood tests; confirmed diabetes: stricter definition of diabetes by exclusion of cases included solely because of glucose blood tests (not confirmed diabetes cases). †Defined as at least one hospital admission for COPD or asthma. ‡Total number of kilometers traveled within 100 m (sum of product of street length and traffic density for each road). §Road with annual traffic density of ≥10,000 vehicles.

**Table 2—Association between traffic-related air pollution per IQR increase in NO<sub>2</sub> (4.9 µg/m<sup>3</sup>) or traffic load within 100 m (1.2 × 10<sup>3</sup> vehicles km/day) and diabetes among 51,818 DCH cohort participants, for two definitions of diabetes incidence based on the National Diabetes Register**

	Adjusted for age	Fully adjusted†	Fully adjusted† + hypertension, hypercholesterolemia, and MI
All diabetes* (n = 4,040)			
NO <sub>2</sub> (µg/m <sup>3</sup> ) 1971 to follow-up	1.06 (1.03–1.09)	1.00 (0.97–1.03)	1.00 (0.97–1.04)
NO <sub>2</sub> (µg/m <sup>3</sup> ) 1991 to follow-up	1.06 (1.03–1.10)	1.00 (0.97–1.04)	1.01 (0.98–1.04)
NO <sub>2</sub> (µg/m <sup>3</sup> ) at baseline (1 year)	1.03 (1.00–1.06)	0.98 (0.95–1.01)	0.98 (0.95–1.01)
Major road‡ within 50 m	1.09 (0.98–1.22)	1.00 (0.90–1.11)	0.99 (0.89–1.10)
Traffic load§ within 100 m (10 <sup>3</sup> vehicle km/day)	1.03 (1.00–1.05)	1.00 (0.98–1.02)	1.00 (0.98–1.02)
NO <sub>2</sub> (µg/m <sup>3</sup> ) at follow-up (1 year)	1.05 (1.02–1.08)	1.01 (0.98–1.03)	1.01 (0.98–1.04)
Confirmed diabetes* (n = 2,877)			
NO <sub>2</sub> (µg/m <sup>3</sup> ) 1971 to follow-up	1.11 (1.07–1.15)	1.04 (1.00–1.08)	1.04 (1.00–1.08)
NO <sub>2</sub> (µg/m <sup>3</sup> ) 1991 to follow-up	1.10 (1.06–1.13)	1.04 (1.01–1.07)	1.04 (1.01–1.07)
NO <sub>2</sub> (µg/m <sup>3</sup> ) at baseline (1 year)	1.08 (1.05–1.11)	1.02 (0.98–1.05)	1.02 (0.98–1.05)
Major road‡ within 50 m	1.20 (1.06–1.36)	1.07 (0.95–1.21)	1.06 (0.94–1.20)
Traffic load§ within 100 m (10 <sup>3</sup> vehicle km/day)	1.05 (1.03–1.08)	1.02 (1.00–1.04)	1.02 (1.00–1.04)
NO <sub>2</sub> (µg/m <sup>3</sup> ) at follow-up (1 year)	1.10 (1.06–1.13)	1.04 (1.01–1.07)	1.04 (1.01–1.08)

Data are HR (95% CI). \*All diabetes: based on original inclusion criteria in NDR: hospital admission for diabetes, diabetes medication, reimbursement for chiropody due to diabetes, or glucose blood tests; confirmed diabetes: stricter definition of diabetes by exclusion of cases included solely because of glucose blood tests (not confirmed diabetes cases). †Adjusted for sex, BMI, waist-to-hip ratio, smoking status, smoking duration, smoking intensity, environmental tobacco smoke, educational level, physical/sports activity in leisure time (indicated yes/no and intensity), alcohol consumption (indicated yes/no and intensity), fruit consumption, fat consumption, and calendar year. ‡Road with annual traffic density of ≥10,000. §Total number of kilometers traveled within 200 m (sum of product of street length and traffic density for each road).

associations were identical with mean NO<sub>2</sub> levels since 1991 (1.04 [1.01–1.07]) and 1-year mean at end of follow-up (1.04 [1.01–1.08]), and weaker but borderline significant with traffic load within a 100-m radius of baseline residence (1.02 [1.00–1.04] per IQR of 1.2 × 10<sup>3</sup> vehicles km/day). Associations were weakly positive and statistically non-significant with 1-year mean NO<sub>2</sub> (1.02 [0.98–1.05]) and indicator of a major road within a 50-m radius (1.06 [0.94–1.20]) of baseline residence. The findings for NO<sub>x</sub> were similar to those for NO<sub>2</sub>, but the associations were weaker (Supplementary Table 2). We found indication of linear dose response between mean NO<sub>2</sub> levels since 1971 and diabetes incidence (Fig. 1).

We found enhanced associations between mean NO<sub>2</sub> levels since 1971 and confirmed diabetes cases in women; people with a high waist-to-hip ratio (>0.90 for men and 0.85 for women); people with <8 years of education and without MI, hypertension, or hypercholesterolemia; and people with asthma or COPD, but no significant effect modifications were detected (Table 3). The association with NO<sub>2</sub> was not modified by BMI. A significant effect modification was detected by physical activity (Wald test for interaction, *P* = 0.02), with effects of NO<sub>2</sub> only in physically active people

(HR 1.10 [95% CI 1.03–1.16]), and by smoking status (*P* = 0.02), with enhanced effect of NO<sub>2</sub> in nonsmokers (1.12 [1.05–1.20]) and none in previous (1.03 [0.98–1.09]) or current smokers (1.00 [0.88–1.09]).

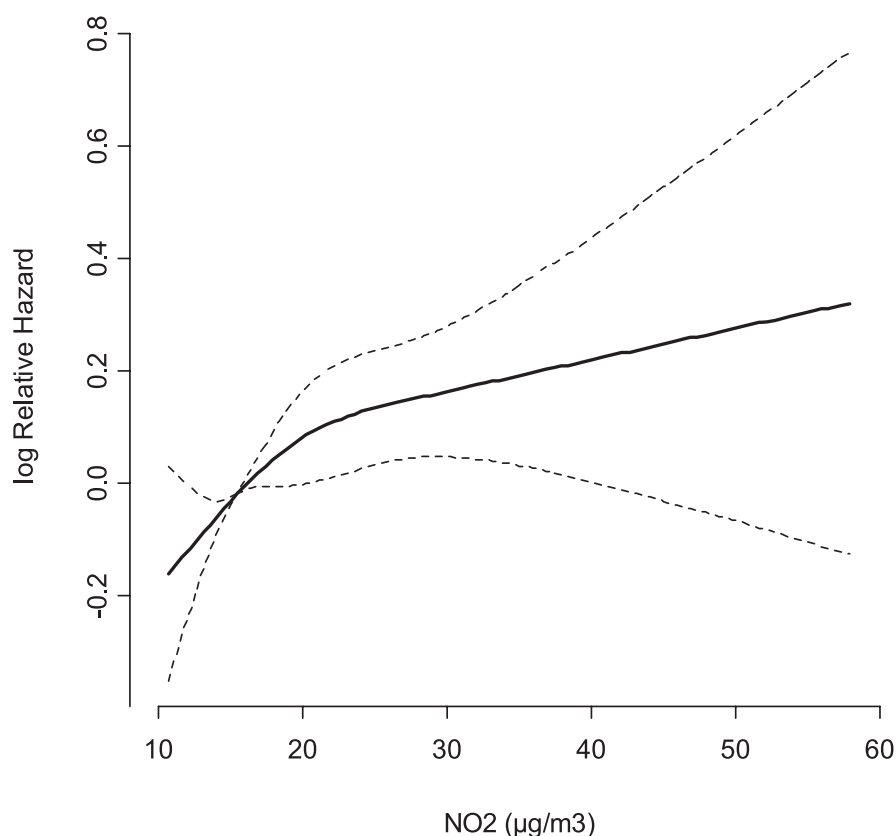
**CONCLUSIONS**—The risk for diabetes was weakly positively associated with increasing mean levels of traffic-related air pollution at the residence. The risk was highest in nonsmokers and physically active people.

This is the first study to relate air pollution to incidence of diabetes assessed objectively from a nationwide register. Previous studies used self-reports of diabetes incidence (16,17), whereas diabetes prevalence was based on administrative databases of physician billing and hospital discharges (14) or on a national survey (15). The onset of diabetes based on self-reports is loosely defined and subject to recall bias (16,17). Thus, the use of objective measures of diabetes onset based on the nationwide register, NDR, is appealing and convenient, since the entire population is covered by uniform inclusion criteria and the dropout rate is zero (19). The disadvantages of the NDR include inability to distinguish between diabetes types, although the majority of new diabetes cases in this age-group is most likely of type 2 diabetes (1). The date of inclusion in the NDR is only a proxy for

the date of formal clinical diagnosis, which was likely made some time before inclusion in the register. No information on glucose or other clinical measurements used at diagnosis are available. Finally, the NDR likely underestimates the actual diabetes burden, since people without clinical diagnoses are not included.

“All diabetes” (*n* = 4,040) definition of incidence was previously validated (19) by a study comparing register-identified patients with their general practitioners, finding sensitivity of 86% and positive predictive value of 90%. Of four inclusion criteria, hospital discharge diagnoses, diabetes medication records, and chiropody all reflect highly likely confirmed diabetes cases. In a “confirmed diabetes” definition (*n* = 2,877), we excluded the 1,163 diabetes cases included solely because of blood glucose tests, because without available results of these tests, or other records in the NDR, it is likely that many of these people did not have diabetes. Because of increasing awareness among physicians on detecting undiagnosed diabetes, it is not uncommon for elderly healthy people to have five blood glucose tests per year.

Incidence rates in this cohort, as observed in all of Denmark, increased over the period 1995–2004 and showed a tendency to decline after 2004 (19), which we adjusted for by modeling the calendar year using restricted cubic splines.



**Figure 1**—Association between exposure to  $\text{NO}_2$  levels at residence and incident diabetes defined by strict definition ( $n = 2,877$ ) (log relative hazard with 95% CI) for 51,818 Danish DCH cohort members, adjusted for sex, BMI, waist-to-hip ratio, smoking status (never, previous, current), smoking intensity, smoking duration, environmental tobacco smoke, physical activity, alcohol intake, fruit intake, fat intake, and educational level.

Our findings are generally consistent with the limited epidemiological evidence linking diabetes prevalence (14,15) and the incidence (16,17) to traffic-related air pollution. The magnitude of reported associations is smaller than that of other studies (15,16), but these were based on women only. Moreover, proximity of the residence to a road was significantly associated with risk for diabetes only in a cohort of women (Nurses' Health Study), where no significant associations were detected with air pollution in a cohort of men (Health Professionals Follow-Up Study) (17). Our results also suggest that the risk for diabetes associated with air pollution may be limited to women (Table 3). Sex-related differences in susceptibility to air pollution could be associated with physiological differences in inflammatory responses or with lifestyle and activity patterns. Brook et al. (15) documented that women spent more time in and around the home and typically worked closer to home in Canada, contributing to smaller exposure misclassification and higher air pollution estimates,

whereas no data in Denmark exist to support this.

Assigning individual exposure to air pollution with high spatial (address-specific) and temporal (residential addresses) resolution, historically back to 1971, is novel and a major strength of this study. We detected identical associations with long (mean levels since 1971 and 1991) and short (1-year mean at follow-up) exposure windows to  $\text{NO}_2$ . Puett et al. (17) on the other hand failed to detect associations with the 1-year mean particulate pollution levels before diabetes incidence. In our data, associations with baseline 1-year mean  $\text{NO}_2$  levels and more "naïve" proxies based on traffic density around the residence were weaker, likely because of exposure misclassification in individuals who moved since baseline.

The association between air pollution and diabetes incidence in his cohort was strongest in nonsmokers (Table 3). Smoking and environmental tobacco smoke are confirmed risk factors for type 2 diabetes (21), as corroborated by our (Table 1) and previous data (16,17). Inhalation of tobacco

smoke triggers similar responses in the lung as air pollution, and these two related exposures share a plausible biological mechanism leading to glucose intolerance via systemic inflammation and inflammation in adipose tissue (22). It has been hypothesized that preexisting low-level inflammation seen in smokers would further enhance the effects of air pollution (2). In contrast, our results suggest a marginal contribution of air pollution to the development of diabetes in smokers and ex-smokers.

Furthermore, in contrast to existing evidence suggesting that individuals with preexisting cardiovascular disease would be more susceptible to the adverse effects of air pollution (2,7,8), we found no statistically significant effect modification of the effects of air pollution by MI, hypertension, or hypercholesterolemia (Table 3). This finding contradicts the study of air pollution-triggered diabetes deaths, which reported effects only in individuals with preexisting cardiovascular disease (11). Whereas diabetic patients have been found to be at higher risk of developing cardiovascular disease due to air pollution than nondiabetic people (7), our results do not support the idea that cardiovascular disease enhances the susceptibility to air pollution and the associated risk of diabetes. Similarly, physical inactivity was found to be the expected risk factor for diabetes, whereas we only found association with air pollution among individuals who were physically active (Table 3). This observation may reflect smaller air pollution exposure misclassification in physically active individuals, because of more time spent outdoors and higher relative exposures and exaggerated air exchange rates during physical activity. It is possible that the association between diabetes and air pollution is only detected among individuals with a low prior risk, i.e., people who are physically active, representing a limited absolute risk, whereas this cannot be detected among individuals with a priori increased risk, suggesting additive patterns (23). Among people with a low level of education, we found the expected increased risk for diabetes as well as enhanced effects of air pollution, although without significant interaction.

We have earlier reported significant positive association between hospital admissions for COPD and asthma and long-term exposure to  $\text{NO}_2$  in this cohort (24,25). It can thus be argued that our results may be affected by diagnosis bias, due to more frequent screening for disease, including diabetes, among chronic respiratory disease patients, who are also more likely to

**Table 3—Modification of associations\* between NO<sub>2</sub> (1971 follow-up, per IQR of 4.9 μg/m<sup>3</sup>) and confirmed diabetes cases (n = 2,877) by baseline characteristics and comorbid conditions among 51,818 DCH cohort participants**

	n	Incidence rate†	HR (95% CI)	P‡
<b>BMI</b>				
Normal/underweight (BMI <25 kg/m <sup>2</sup> )	407	1.8	1.10 (1.00–1.21)	
Overweight (25 ≤ BMI < 30 kg/m <sup>2</sup> )	1,262	6.0	1.01 (0.96–1.07)	
Obese (BMI ≥30 kg/m <sup>2</sup> )	1,208	18.3	1.05 (1.00–1.12)	0.30
<b>Waist-to-hip ratio</b>				
>0.90/0.85 (male/female)	2,283	9.5	1.09 (1.01–1.18)	
≤0.90/0.85 (male/female)	594	2.3	1.04 (1.00–1.08)	0.24
<b>Sex</b>				
Males	1,702	7.3	1.01 (0.97–1.07)	
Females	1,175	4.4	1.07 (1.01–1.13)	0.15
<b>Physical activity</b>				
Yes	1,185	4.3	1.10 (1.03–1.16)	
No	1,692	7.5	1.00 (0.95–1.05)	0.02
<b>Smoking status</b>				
Never	823	4.5	1.12 (1.05–1.20)	
Previous	891	6.2	1.02 (0.95–1.10)	
Current	1,163	6.5	1.00 (0.95–1.06)	0.05
<b>Education</b>				
<8 years of education	1,209	7.4	1.06 (1.01–1.12)	
8–10 years of education	1,250	5.4	1.03 (0.98–1.09)	
≥10 years of education	418	3.9	0.98 (0.88–1.09)	0.44
<b>MI</b>				
Yes	129	14.8	0.93 (0.77–1.12)	
No	2,748	5.6	1.04 (1.01–1.08)	0.36
<b>Hypertension</b>				
Yes	915	12.1	1.03 (0.96–1.10)	
No	7,285	4.6	1.04 (1.00–1.40)	0.71
<b>Hypercholesterolemia</b>				
Yes	346	10.0	0.96 (0.85–1.09)	
No	2,531	5.4	1.05 (1.01–1.09)	0.20
<b>COPD</b>				
Yes	140	9.6	1.05 (1.01–1.09)	
No	2,737	8.0	0.99 (0.83–1.17)	0.52
<b>Asthma</b>				
Yes	78	8.1	1.05 (1.01–1.09)	
No	2,799	8.1	0.85 (0.66–1.11)	0.13

\*Adjusted for sex, BMI, waist-to-hip ratio, smoking status, smoking duration, smoking intensity, environmental tobacco smoke, educational level, physical/sports activity in leisure time (indicated yes/no and intensity), alcohol consumption (indicated yes/no and intensity), fruit consumption, fat consumption, and calendar year. †Crude rate per 1,000 person-years. ‡For Wald or likelihood ratio test for interaction.

live in areas with higher air pollution levels than the rest of the cohort (24,25). However, this is unlikely, since we found only a marginally higher prevalence of asthma and COPD in diabetic people than in nondiabetic people (Table 1), and history of asthma or COPD hospitalizations was not significantly associated with increased risk of diabetes (results not shown). Furthermore, comorbidity with asthma or COPD did not statistically significantly modify the effect of air pollution on diabetes (Table 3).

Effects of air pollution were slightly enhanced in people with a high waist-to-hip (Table 3), corroborating the evidence

from animal models, where exposure to air pollution led to glucose intolerance only in obese rats (7). However, enhanced air pollution effects were not seen with increasing BMI. This is the first study of air pollution and diabetes to adjust for waist-to-hip ratio, which is an independent predictor of diabetes next to BMI. Excluding waist-to-hip ratio from the full model as others did (14–17) resulted in even stronger associations with air pollution (HR 1.05 [95% CI 1.02–1.10]) as did exclusion of BMI (1.08 [1.04–1.12]). Mean levels of NO<sub>2</sub> were weakly correlated with BMI (correlation coefficient  $r = 0.03$ ) and waist-to-hip ratio ( $r = 0.02$ ).

Obesity, the most important risk factor for diabetes (1), has been linked to air pollution (22) and may be on the biological pathway between air pollution and diabetes. Our results remained unchanged when adding total energy intake, including energy due to alcohol intake, into the model.

The strengths of this study include the large prospective cohort with an objective assessment of diabetes incidence and well-defined confounders, with minimal possibility of recall and information bias. The main limitation is the exposure misclassification in modeled air pollution concentrations, since these are only proxies of personal exposure. This cohort lacked information on indoor exposures, use of air purifier and air conditioning, work address, working time, as well as commuting habits and personal outdoor activity patterns. However, lack of these parameters and resulting exposure misclassification is likely to be nondifferential with respect to diabetes diagnoses. The air pollution models used to assess NO<sub>2</sub> levels were successfully validated (26,27) and applied in epidemiological studies (24,25). Another limitation is the lack of results from blood glucose tests from the National Health Insurance Registry, which precludes better ascertainment of diabetes among 1,163 people included in NDR solely due to blood glucose tests and more specific overall diabetes definition. Finally, there is little evidence coming from original research that supports biological plausibility of our findings (8,9,22), and more studies will be needed to confirm or reject the link between air pollution and diabetes.

We detect weak positive associations between diabetes incidence and traffic-related air pollution at residence, and we add to the body of evidence that air pollution may be a risk factor for diabetes. We offer several novel findings that need to be reproduced. The effects of air pollution were strongest in nonsmokers and physically active people. Facing the emerging challenges in controlling the diabetes epidemic, a possibility that air pollution contributes to the diabetes burden may have a huge public health impact.

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Z.J.A. takes full responsibility for the work as a whole, including the study design, access to data, and the decision to submit and publish the article. Z.J.A. analyzed data, designed the study, and wrote the manuscript. O.R.-N. contributed to the design of the study and preparation of the manuscript. M.K., S.S.J.,

and M.H. developed the Danish dispersions air pollution model and calculated air pollution data. S.L. secured the funding for the study and contributed to the design and interpretation of findings. A.T. and K.O. are the main investigators of the DCH cohort study. M.S. designed the study, contributed to the analysis strategy, and helped in preparing the manuscript.

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