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# **Epithelial ovarian cancer and exposure to dietary nitrate and nitrite in the NIH-AARP Diet and Health Study**

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# **Abstract**

Ovarian cancer is a leading cause of cancer death among women in the United States and it has the highest mortality rate of all gynecologic cancers. Internationally, there is a five-fold variation in incidence and mortality of ovarian cancer, which suggests a role for environmental factors, including diet. Nitrate and nitrite are found in various food items and they are precursors of Nnitroso compounds, which are known carcinogens in animal models. We evaluated dietary nitrate and nitrite intake and epithelial ovarian cancer in the National Institutes of Health (NIH)-AARP Diet and Health Study, including 151 316 women aged 50–71 years at the time of the baseline questionnaire in 1995–1996. The nitrate and nitrite intake was assessed using a 124-item validated food frequency questionnaire. Through 31 December 2006, 709 incident epithelial ovarian cancer cases with complete dietary information were identified. Using Cox proportional hazards regression to estimate hazard ratios and 95% confidence intervals (CIs), women in the highest intake quintile of dietary nitrate had a 31% increased risk (95% CI: 1.01–1.68) of epithelial ovarian cancer, compared with those in the lowest intake quintile. Although there was no association for total dietary nitrite, those in the highest intake category of animal sources of nitrite had a 34% increased risk (95% CI: 1.05–1.69) of ovarian cancer. There were no clear differences in risk by histologic subtype of ovarian cancer. Our findings suggest that a role of dietary nitrate and nitrite in ovarian cancer risk should be followed in other large cohort studies.

# **Keywords**

epithelial; nitrate and nitrite; ovarian cancer

# **Introduction**

Ovarian cancer has the highest mortality rate of all of the female cancers and more than 13 000 American women are expected to be diagnosed with the disease in 2010 (Jemal et al., 2010). Few modifiable risk factors for ovarian cancer have been identified and the etiology

**Conflicts of interest**

There are no conflicts of interest.

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of this malignancy is poorly understood (Byers et al., 1983; Mori et al., 1988; Hankinson and Danforth, 2006; Moorman et al., 2009). However, the five-fold international variation in ovarian cancer incidence and mortality (Baker and Piver, 1994) and the increase in incidence in migrants from Japan to the United States (Dunn, 1975; Herrinton *et al.*, 1994) strongly suggest a role for environmental factors, including diet. There is limited data on the relation between diet and ovarian cancer (Hankinson and Danforth, 2006). Investigation of the role of nitrate and nitrite in ovarian cancer risk is of interest as these compounds are precursors for the endogenous formation of N-nitroso compounds (NOC), which have been shown to induce tumors in animals (Bogovski and Bogovski, 1981) and potentially in humans (McKnight et al., 1999).

Nitrate is a natural component of plants and is found at high concentrations in leafy vegetables, such as lettuce and spinach, and some root vegetables, such as beets (Gangolli et al., 1994). Nitrate is also present as a contaminant in drinking water, which can be a major source of intake when levels are at/above the maximum contaminant level (MCL) of 10 mg/l nitrate-nitrogen (Chilvers and Caygill, 1984). Nitrite and nitrate salts are added to cured meats such as bacon, hot dogs, and ham to prevent the growth of spore-forming bacteria as well as to add color and flavor (Lück, 1985).

The relation between nitrate levels in municipal drinking water and ovarian cancer risk was evaluated previously in a cohort of 21 977 women in Iowa (USA) (Weyer et al., 2001). The investigators observed increasing relative risks (RRs) of 1.0, 1.52, 1.81, and 1.84 for ovarian cancer  $(n=102 \text{ cases})$  across increasing quartiles of municipal nitrate water concentrations; however, the relationship between dietary nitrate intake and ovarian cancer was null (Weyer et al., 2001). Using prospective data from the National Institutes of Health (NIH)-AARP Diet and Health Study, we evaluated the relationship between dietary nitrate and nitrite intake and epithelial ovarian cancer incidence. This large cohort allowed us to evaluate associations after accounting for factors previously shown to influence endogenous Nnitrosation reactions (Mirvish, 1996), including cigarette smoking and intake of vitamins C and E.

# **Participants and methods**

#### **Study population**

The NIH-AARP Diet and Health Study was initiated in 1995–1996 when an extensive baseline questionnaire was mailed to 3.5 million AARP members aged 50–71 years residing in one of six US states (California, Florida, Pennsylvania, New Jersey, North Carolina, and Louisiana) and two US metropolitan areas (Atlanta, Georgia, and Detroit, Michigan; Schatzkin et al., 2001). This questionnaire ascertained information on usual dietary intake over the past 12 months, use of individual and multivitamin supplements, smoking history, alcohol intake, height and weight at baseline, reproductive history, menopausal status, and other factors. A total of 617 119 persons returned the baseline questionnaire, and 567 169 questionnaires were determined to have been satisfactorily completed. The NIH-AARP Diet and Health Study was approved by the Special Studies Institutional Review Board of the National Cancer Institute.

Among the 226 733 women with satisfactory baseline questionnaires, there remained 200 571 women after excluding those with duplicate questionnaires, those who had died or moved out of the study area prior to baseline, those who withdrew from the study, and those who had questionnaires completed by proxy respondents. We further excluded those who had been previously diagnosed with cancer except for nonmelanoma skin cancer  $(n=26 546)$ and those with extreme values for total energy intake (i.e. beyond twice the interquartile range of Box-Cox log-transformed intake, corresponding to <317 and >4791 kcal/day; total:

 $n=1836$ ). We additionally excluded women who had a bilateral oophorectomy or unknown oophorectomy status at baseline ( $n=20 855$ ), no follow-up time ( $n=10$ ), or an extreme body mass index (BMI, 12 kg/m<sup>2</sup> or 45 kg/m<sup>2</sup>; n=8). After these exclusions, 151 316 women were available for analysis.

#### **Cancer ascertainment**

Incident, first primary epithelial ovarian cancer cases, topography code C56.9 of the International Classification of Diseases for Oncology, Third Edition (World Health Organization, 2001), were identified through 31 December 2006, by linkage of the cohort database to cancer registries of the eight original plus two additional states (Texas and Arizona), and the National Death Index Plus. By histologic subtype, total incident epithelial ovarian cancer cases included the International Classification of Diseases for Oncology, Third Edition histology codes 8000–8573; serous cancers included 8441, 8450, 8460, 8461, and 8462; endometrioid included 8380, 8381, 8560, and 8570; and mucinous included 8470, 8471, 8480, 8481, and 8490.

#### **Dietary intake**

The dietary component of the baseline questionnaire asked about the frequency of consumption and corresponding portion sizes of 124 food items, including 14 fruits and 23 vegetables, and fresh and processed meats, during the past 12 months. Participants were queried about their frequency of intake in 10 categories ranging from 'never' to '2+ times per day' for foods and 'never' to '6+ times per day' for beverages. Each line item was accompanied by three possible portion size categories. The food items, portion sizes, nutrient database, and Pyramid Food Servings database were utilized using methods developed by Subar *et al.* (2000) with national dietary data from the US Department of Agriculture's 1994–1996 Continuing Survey of Food Intakes by Individuals (CSFII; Friday, 2006). The Pyramid Food Servings database used a recipe file to disaggregate food mixtures into their component ingredients and assign them to food groups.

The food frequency questionnaire (FFQ) was validated using two 24 h recalls in a subset of the cohort (Thompson *et al.*, 2008). In the validation study, the investigators found that for the 26 nutrient constituents examined, when adjusted for reported energy intake, the estimated correlations with true intake ranged from 0.36 to 0.76 and the attenuation factors ranged from 0.24 to 0.68. These results compare favorably with those from other large prospective studies. For fruits and vegetables, the adjusted correlation was 0.71 in men and 0.61 in women. For vitamin C, the adjusted correlation was 0.70 for men and 0.65 for women. However, for protein, the adjusted correlation was lower, with 0.43 for men and 0.50 for women.

The nitrite and nitrate content in over 3000 foods were determined by conducting a review of the literature, focusing on US and Canadian foods, and calculating means of the published values weighted by the number of food samples analyzed (Ward et al., 2006). If values from US or Canadian foods were unavailable, we used values from other western countries. The nitrite and nitrate values for foods constituting a FFQ line item were combined by weighting the food-specific values by sex-specific intake amounts from the 1994–1996 CSFII (Subar et al., 2000). For example, the nitrate content of a line item was calculated using a weighted average of the nitrate content in the included foods where the weights were determined by intake amounts from the CSFII, specific for the age group and sex. Daily intakes of nitrate and nitrite were calculated by multiplying the frequency of consumption of each line item by its nitrate or nitrite content and summing over line items. In addition to calculating dietary nitrite and nitrate from all foods, we calculated nitrite from plant, animal, and processed meat sources separately.

Although information about the usual source of drinking water and tap water intake was not assessed for the cohort, we used the census tract location of a participant's residence at enrollment to estimate the likelihood that they may have been exposed to elevated nitrate concentrations by their drinking water supply in a Geographic Information System. First, residential addresses for all participants were geocoded to a census tract. Using ESRI's ArcInfo (Redlands, California, USA) Geographic Information System software we linked the census tract to a geospatial model developed by Nolan and Hitt (2006) that predicts nitrate contamination of ground water used for drinking water across the continental US. The predicted nitrate concentration represents average levels in an area of approximately 20 square kilometers, which is the minimum size of the US Geological Survey ground water monitoring network. Ground water provides drinking water for more than one-half of the US population (Solley et al., 1993), and is the sole source of drinking water for many rural communities and some large cities. The model includes variables that are significant predictors of nitrate in ground water including nitrogen fertilizer and manure applications, the location of orchards/vineyards, aquifer rock type, and population density, and provides a predicted nitrate concentration level (0–5,>5–10,>10 mg/l nitrate-N).

We overlaid the 1 km by 1km grid of predicted nitrate levels across the United States onto study participants' census tracts. Using the Tabulate Area Tool in the Spatial Analyst extension of ESRI's ArchInfor (Redlands), we calculated the area within each census tract that intersected areas estimated to have an average nitrate level of 10 mg/l or more, the MCL. We identified participants who resided in census tracts where 50% or more of the area was estimated to overlie groundwater with nitrate levels of 10 mg/l or greater; because their nitrate intake from water sources might have exceeded that from their diet, we excluded these participants in sensitivity analyses.

#### **Statistical analysis**

Person–years of follow-up for each participant accrued from the date of return of the baseline questionnaire to the date of ovarian cancer diagnosis, the date of moving out of the registry ascertainment area, death, or the end of the follow–up period. Cox proportional hazards regression models with person–years as the underlying time metric were used to estimate hazard ratios (HR) and 95% confidence intervals (CIs) for ovarian cancer within quintiles of dietary intake. Tests for linear trend were conducted using the median value of each exposure category as an ordinal variable in the model. Dietary variables were adjusted for energy intake using the nutrient density method (Willett, 1998), which expresses intake in units per 1000 calories. The proportional hazards assumption was tested and upheld in all analyses.

In all models, we adjusted for age (continuous), cigarette smoking status (never, current, former), race (White, Black, other, missing), family history of cancer (yes, no), BMI (<25, 25–29.9, 30–34.9, 35 kg/m<sup>2</sup>), menopausal status at baseline (premenopausal, natural <45 years, natural 45–49 years, natural 50–54 years, natural 55 years, surgical, unknown), parity (nulliparous,  $1, 2$ , or 3 or more children), age at menarche ( $12$  years,  $13-14$  years,

≥15 years). The family history of cancer variable did not specify ovarian cancer. We adjusted for total daily dietary vitamin C intake (mg/1000 kcal) in the risk model by quintiles of nitrate and nitrite but not in the joint effect model (combined effects of nitrate and nitrite and vitamin C).

To test for effect modification (statistical interaction) by total dietary vitamin C and vitamin E intake, we included these continuous variables in a multivariate model with a cross product term. We assessed multiplicative interactions by adding the relevant cross-product term to main-effects models, with the P value for interaction determined by a Wald test for the cross-product term.

To evaluate the consistency of associations, we stratified by age (above/below the median age of 62.6 years), BMI (above/below median of 27 kg/m<sup>2</sup>), and education (high school or fewer years of education; some college or more years of education). We stratified by smoking status (ever/never), red meat intake (above/below median of 26.1g/1000 kcal), and vitamin C intake (above/below median of 90.8 mg/1000 kcal) to evaluate factors potentially affecting endogenous N-nitrosation. We repeated our analyses in a subsample restricted to those women identified by geocoding as potentially having low intake levels of nitrate from drinking water. For all comparisons, P values were two-sided and  $\alpha$  <0.05 was considered statistically significant.

# **Results**

A total of 709 incident epithelial ovarian cancer cases were identified during an average of 10 years of follow-up. Of these, 374 were serous, 66 were endometrioid, and 35 were mucinous tumors. The remaining 234 epithelial ovarian cancer cases included clear and transitional cell tumors, and tumors with other rare and unknown histologies.

Women in the highest quintiles of nitrate and nitrite intake were less likely to be current smokers, had a lower BMI, and had lower intakes of total calories and processed meat as compared with women in the lowest quintiles of nitrate and nitrite intake (Table 1). Furthermore, women in the highest quintiles of nitrate and nitrite intake were more likely to be highly educated, more physically active, and consumed more fruits and vegetables than those in the lowest quintiles. Women who consumed more nitrate consumed less processed meat, whereas women who consumed higher levels of nitrite consumed higher levels of processed meat. Although women in the highest quintile of nitrite intake were less likely to be a current smoker and consumed fewer calories, they tended to be more educated, more physically active, and more likely to consume fruits, and vegetables, than those in the lowest quintile. No difference in nitrate or nitrite intake was observed for the reproductive factors evaluated (data not shown).

The mean dietary nitrate intake in the total study population was 91.9 mg/day [standard deviation (SD)=  $68.6$  mg/day] and the mean dietary nitrite intake was 1.1 mg/day (SD= $0.5$ ) mg/day). The average intake of nitrate from plant sources was 86.7 mg/day or 94.9% of total nitrate intake, whereas average intake of nitrite from plant sources was 0.73 mg/day or 66.3% of the total nitrite intake. The average intake of nitrate from animal sources was 4.7 mg/day or 5.1% of the total nitrate intake, whereas average intake of nitrite from animal sources was 0.3 mg/day or 27.3% of the total nitrite intake. The average of the combined nitrate and nitrite intake from processed meat sources was 0.84mg/day (SD=1.02 mg/day).

The major contributors to nitrate intake were lettuce (30.7%), cooked spinach (8.8%), and broccoli (5.3%). The major contributors to nitrite intake were pasta (6.7%), rice (6.2%), and hot cereal (5.6%). Processed meats contributed 10.0% to total nitrite intake and 30.5% to nitrite intake from animal sources. As the major sources of nitrate were vegetables, which also contain beneficial nutrients such as vitamin C, we evaluated the correlation of intake of vitamin C with nitrate intake in our study population. The correlation coefficient for nitrate and vitamin C was 0.39.

In Table 2, we tested the hypotheses that nitrate and nitrite intake are associated with increased risk of epithelial ovarian cancer and investigated these associations by histologic subtypes. Women in the highest, compared with the lowest, quintile of nitrate intake had an elevated risk of ovarian cancer, and the trend across increasing quintiles was borderline statistically significant (Q5 vs. Q1, HR= $1.31$ ; 95% CI:  $1.01-1.68$ , Ptrend= 0.06). Although total nitrite intake was not associated with risk, we observed a positive association between

nitrite from animal sources and epithelial ovarian cancer risk (Q5 vs. Q1, HR=1.34; 95% CI: 1.05–1.69, P trend=0.02). In contrast, neither nitrite from plants nor processed meat sources were associated with ovarian cancer risk. We also evaluated food sources rich in nitrate and nitrite, including intake of red and processed meats, as well as fruits and vegetables, and no statistically significant differences in epithelial ovarian cancer risk were identified (data not shown).

In general, there were no clear differences in risk according to histologic subtypes; although for those in the highest quintile of total nitrate, the risks were elevated for serous (Q5 vs. Q1, HR=1.27; 95% CI: 0.90–1.78, P trend=0.08) and endometrioid subtypes (Q5 vs. Q1,  $HR=1.88$ ;  $95\%$  CI:  $0.80-4.43$ , Ptrend=0.18), but not for the mucinous subtype (of which there were only six cases in the highest quintile of intake). We observed a statistically significant positive trend of increasing risk of mucinous ovarian cancer across increasing quintiles of nitrite intake from processed meats ( $P$  trend=0.04); however, case numbers were small and the risk estimate in the highest quintile was not statistically significant (Q5 vs. Q1, HR=2.24; 95% CI: 0.76–6.61). Our results were unchanged when we excluded the 2.8% of the study population  $(n=4176,$  including 20 ovarian cancer cases) who resided in census tracts with predicted nitrate levels greater than 10 mg/l for more than 50% of the census tract. We also stratified our analyses by age, education, smoking status, BMI, and physical activity; results were consistent with our main findings.

As vitamin C can inhibit N-nitrosation reactions, we evaluated the joint effect of nitrate or nitrite and vitamin C intake on epithelial ovarian cancer risk (Table 3). We considered women with low nitrite or nitrate intake and high vitamin C intake as the low risk referent group based on our hypothesis. The association between high nitrate/nitrite intake and either high or low vitamin C intake did not differ, and we did not observe any statistically significant interactions. We also evaluated the joint effect of nitrate or nitrite with vitamin E intake and did not observe any statistically significant interactions (data not shown).

# **Discussion**

In this large prospective cohort of AARP members, we found that higher intake of dietary nitrate was associated with an increased risk of ovarian cancer by approximately 30% for women with the highest compared with those in the lowest quintile of intake. Although total nitrite intake was not associated with ovarian cancer risk, we identified a 34% increase in risk and a statistically positive significant trend for animal sources of nitrite. Taken together, our results suggest that the source of the NOC precursors may be important for epithelial ovarian cancer risk.

One study previously evaluated the relationship between dietary nitrate intake and ovarian cancer and the findings were null and the case numbers more modest  $(n=102;$  Weyer *et al.*, 2001). Although no prior study has evaluated sources (i.e. plant vs. animal) of dietary nitrate and nitrite intake with respect to ovarian cancer risk, other studies have reported positive associations between high meat intake and ovarian cancer risk (Forman, 1987; La Vecchia et al., 1987; Mori and Miyake, 1988; Mori et al., 1988; Bosetti et al., 2001; Zhang et al., 2002). These findings are consistent with our observations for nitrite from animal sources. However, it should be noted that in a large cohort study including more than 325 000 European women, there was no evidence of an association between consumption of animal foods (meat, fish, eggs, dairy products) and risk of ovarian cancer (Schulz et al., 2007).

Ingested nitrate is absorbed in the small intestine and 25% is excreted in the mouth, where oral bacteria reduce approximately 20% to nitrite (approximately 5% of ingested nitrate; Gangolli et al., 1994). There is great interindividual variability in endogenous production of

NOCs, which is mediated by several known factors. Nitrite and reactive nitrogen species react with nitrosatable compounds, mainly amines and amides, to form NOCs (Mirvish, 1995, 1996; Bartsch and Frank, 1996; Grosse et al., 2006; Schulz et al., 2007). Processed meats are not only a source of nitrate and nitrite, but also a source of amines and amides, which are also precursors of NOCs. As a result, consumption of nitrate or nitrite from processed meats should theoretically result in more exposure to NOCs than plant-based foods; however, our results for nitrite appear to be due to animal sources, no association with processed meats was observed. As vegetables contain inhibitors of in-vivo  $N$ nitrosation such as vitamin C (Mirvish, 1996), we might expect a lack of an association between plant sources of nitrate and nitrite and ovarian cancer; this observation has been pointed to in several epidemiologic studies that have reported either no association or inverse associations between dietary nitrate intake and human cancers other than ovarian cancers. However, we did not observe an interaction between nitrate/nitrite and vitamin C intake and ovarian cancer risk.

The finding that animal sources of nitrite, but not processed meat sources, are positively associated with ovarian cancer risk was somewhat unexpected. Although the top contributor to daily nitrite intake in the NIH-AARP study population is coldcuts (7.2% of the daily nitrite variable), women consume lower levels of nitrite from processed meats than men (4.6% of nitrite intake in women is from coldcuts). In fact, the major animal source of nitrite in women is milk, as it contributes to 8.8% of nitrite intake in women. Our results suggest that additional animal sources of nitrite, such as milk, eggs, and ham (not luncheon meat), in addition to coldcuts and hotdogs, may increase the risk of ovarian cancer. Milk sugar lactose has been hypothesized as a risk factor for epithelial ovarian cancer previously because of possible direct toxic effects of its metabolites on oocytes or by compensatory gonadotropin stimulation (Fairfield et al., 2004). Investigators evaluated this hypothesis in the Nurses Health Study (Fairfield *et al.*, 2004) and found that women who consumed one or more servings of skim or low-fat milk daily had a 32% higher risk of any ovarian cancer (RR=1.32; 95% CI: 0.97–1.82) and a 69% higher risk of serous ovarian cancer (RR=1.69; 95% CI: 1.12–2.56) compared with women consuming three or fewer servings monthly. Further investigation of the role of nitrite from milk sources should be considered in future studies of ovarian cancer.

Drinking water nitrate can constitute the majority of nitrate intake when levels are near or above the MCL (Chilvers and Caygill, 1984). In a previous study, a positive association was observed across increasing quartiles of nitrate in the municipal water supply and ovarian cancer risk (Weyer et al., 2001); however, these findings were based on a limited number of cases. Although information on the primary source of drinking water was not obtained in this study, our findings were not altered when we excluded individuals living in areas where nitrate contamination of ground water is highly probable. Although a limitation of our study is that we did not have surface water estimates to identify persons for exclusion, more than half of the US population obtains drinking water from groundwater sources (Solley *et al.*, 1993). Furthermore, surface waters in areas with contaminated groundwater are also likely to be affected. Lack of individual level drinking water information likely resulted in some misclassification of nitrate exposure from drinking water. Future investigation in this area, if possible, should utilize a study design that captures drinking water intake of nitrate in addition to dietary intake.

Our study had some additional limitations, including the possibility of measurement error associated with using a FFQ to assess usual diet, which, if nondifferential, could attenuate observed associations. The low response rate in this population is also a limitation as it could be a source of bias. In addition, measured nitrate values are not available for comparison for any region included in the study and the addition of measured values

specific to the NIH-AARP population would improve the perceived validity of the exposure estimate. However, our study had numerous strengths, including the use of a detailed validated questionnaire to assess dietary intake of nitrate and nitrite and the wide range of intake: among participants in our study, the median intake of nitrate in the highest quintile was over five times that in the lowest quintile, and for nitrite, the highest quintile was over two times that in the lowest quintile. Other strengths include the prospective nature of the study, completeness of follow-up, the relatively large number of ovarian cancer cases, which allowed us to explore associations by histologic subtype, and the ability to adjust for a large number of potential confounding variables. As the follow-up time for this population increases, the power to evaluate this hypothesis will also be a strength relative to other smaller cohort populations.

In summary, we found a modest increased risk of epithelial ovarian cancer among women with high dietary nitrate intake and among those with a high intake of nitrite from animal sources. This study suggests that intake of NOC precursors may play a role in the etiology of ovarian cancer. These findings should be evaluated in additional studies with classification by histologic type, data on recent and long-term nitrate and nitrite intake, and detailed water consumption data.

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**Table 1**

Means and proportions for baseline characteristics of women in the NIH-AARP Diet and Health Study cohort by quintiles of nitrate and nitrite intake ( = 151 316)

 $\leq$ 



Q, quintile.

**Table 2**

Multivariate hazard ratios and 95% confidence intervals according to quintile of nitrite and nitrate intake and ovarian cancer in the NIH-AARP Diet and<br>Health Study (n = 151 316), 1995–2006 Multivariate hazard ratios and 95% confidence intervals according to quintile of nitrite and nitrate intake and ovarian cancer in the NIH-AARP Diet and  $n = 151316$ , 1995–2006 Health Study (





m<sup>2</sup>), education (<high school–high school graduate, posthigh school other than college, some college, college graduate, postgraduate, and missing), smoking status (current, former, never, missing),<br>menopausal status (pre menopausal status (premenopausal, natural: < 45 years, natural: 45–49 years, natural: 50–54 years, natural: >55 years, surgical, unknown), parity (nulliparous, 1, 2, or 3 or more children), age at menarche m2), education (<high school–high school graduate, posthigh school other than college, some college, college graduate, postgraduate, and missing), smoking status (current, former, never, missing), (12 years, 13-14 years, 15 years), and total daily dietary vitamin C intake (continuous). (≤12 years, 13–14 years, ≥15 years), and total daily dietary vitamin C intake (continuous).

CI, confidence interval; HR, hazards ratio; NIH, National Institutes of Health; Q, quintile. CI, confidence interval; HR, hazards ratio; NIH, National Institutes of Health; Q, quintile.

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 $\rm{^2Units}$  are in mg/1000 kcal. Units are in mg/1000 kcal.

# **Table 3**

Multivariate hazard ratios and 95% confidence intervals according to low/high a joint intake of nitrite or nitrate and vitamin C and ovarian cancer risk ( Multivariate hazard ratios and 95% confidence intervals according to low/high<sup>a</sup> joint intake of nitrite or nitrate and vitamin C and ovarian cancer risk (n = 151 316)



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29.9, 30-34.9, 35 Adjusted for age (continuous), race (nonHispanic White, nonHispanic Black, other, and missing), total energy intake (continuous), family history of cancer, body mass index ( <25-29.9, 30–34.9, 25–29.9, 20) š

menopausal status (premenopausal, natural: <45 years, natural: 45-49 years, natural: 50-54 years, natural: >55 years, surgical, unknown), parity (nulliparous, 1, 2, or 3 or more children), age at menarche menopausal status (premenopausal, natural: <45 years, natural: 45–49 years, natural: 50–54 years, natural: >55 years, surgical, unknown), parity (nulliparous, 1, 2, or 3 or more children), age at menarche kg/m<sup>2</sup>), education (< high school-high school graduate, posthigh school other than college, some college, college graduate, postgraduate, and missing), smoking status (current, former, never, missing), kg/m<sup>2</sup>), education ( < high school–high school graduate, posthigh school other than college, some college, college graduate, postgraduate, and missing), smoking status (current, former, never, missing),  $(12 \text{ years}, 13-14 \text{ years}, 15 \text{ years}).$  $(12 \text{ years}, 13-14 \text{ years}, 15 \text{ years}).$ 

CI, confidence interval; HR, hazards ratio. CI, confidence interval; HR, hazards ratio.

 $\rm ^{4}Low-high$  distinctions based on median in<br>take values. Low–high distinctions based on median intake values.