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Analysis of dose response for circulatory disease following radiotherapy for benign disease

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Abstract

Purpose—To assess the shape of the dose response for various circulatory disease endpoints, and modifiers by age and time since exposure.

Methods and Materials—Analysis of the US peptic ulcer data testing for heterogeneity of radiogenic risk by circulatory disease endpoint (ischemic heart, cerebrovascular, other circulatory disease).

Results—There are significant excess risks for all circulatory disease, with an excess relative risk Gy^{-1} of 0.082 (95% CI 0.031, 0.140), and ischemic heart disease, with an excess relative risk Gy^{-1} of 0.102 (95% CI 0.039, 0.174) (both p<0.01), and indications of excess risk for stroke. There are no statistically significant (p>0.2) differences between risks by endpoint, and few indications of curvature in the dose response. There are significant modifications of relative risk by time since exposure, the magnitude of which does not vary between endpoints (p>0.2). Risk modifications are similar if analysis is restricted to those receiving radiation, although relative risks are slightly larger and the risk of stroke fails to be significant. The slopes of the dose response are generally consistent with those observed in the Japanese atomic bomb survivors and in occupationally and medically exposed groups.

Conclusions—There are excess risks for a variety of circulatory diseases in this dataset, with significant modification of risk by time since exposure. The consistency of the dose-response slopes with those observed in radiotherapeutically-treated groups at much higher dose, as well as in lower-dose exposed cohorts such as the Japanese atomic bomb survivors and nuclear workers implies that there may be little sparing effects of fractionation of dose or low dose-rate exposure.

Keywords

circulatory disease; ischemic heart disease; stroke; peptic ulcer; benign disease

Conflicts of interest notification

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Introduction

Based on observations in irradiated populations, the health risks of low-level exposure to ionizing radiation have been assumed to be related primarily to cancer. At high radiation doses a variety of other well-established effects are observed, in particular damage to the heart and to the coronary and other large arteries, both in patients receiving radiotherapy (RT) and in experimental animals ¹. There are plausible, if not completely understood, mechanisms by which high doses of radiation affect the blood circulatory system ^{2;3}.

Circulatory system effects have also been observed in populations exposed at somewhat lower doses, in particular in the Life Span Study (LSS) cohort of Japanese atomic bomb survivors ⁴, in a UK group treated with spinal X-rays for ankylosing spondylitis ⁵ and in a US cohort treated with X-rays for peptic ulcer ⁶. In the peptic ulcer cohort there is excess risk of ischemic heart disease (IHD) in patients receiving mean heart doses in the range 0-7.6 Gy, with excess IHD risk above 2.6 Gy, but no excess risk for other heart disease ⁶. Carr *et al.* described relative risks of IHD in various dose groups, but did not otherwise model the dose response, or consider modifying variables.

In the atomic bomb survivors, assessments have previously been made of heterogeneity between cancer types in the adjustments of the radiogenic excess risk for gender, time since exposure or age at exposure ⁷; similar analysis has not been attempted in relation to circulatory disease^{4;8}.

In this paper we re-analyse the shape of the dose response for various circulatory disease endpoints in the US peptic ulcer dataset, and assess modifications of risk by age at exposure and time since exposure. Because of uncertainty as to the target for radiation-induced disease⁹, we consider risk in relation to dose to various target organs. Patients chosen for irradiation may have been less medically well (i.e., not fit for anesthetic/surgery) than those treated in other ways; assessing this potential bias requires that we analyze the full cohort (exposed+unexposed) as well as the exposed group only. Using the methods of Pierce and Preston ⁷ we formally evaluate heterogeneity of the shape of the dose-response and modifications by time and age in this dataset. The data used here are very similar to those used in previous analyses of this cohort ^{6;10}.

Data and Methods

Data

The cohort consisted of 3719 persons, comprising 1860 unexposed persons and 1859 exposed patients. Removing 8 persons in the exposed group for whom the dosimetry was incomplete, and 111 who had received megavoltage or ⁶⁰Co γ -therapy, and for whom phantom measurements were not done, led to an analysis cohort of 3600 persons. Follow-up started 5 years after radiation treatment in the exposed cohort, in contrast to the paper of Carr *et al.* ⁶ in which follow-up started 1 or 10 years after exposure. Follow-up finished with the earliest of date of death, date last known alive or December 31st 1997, yielding a total of 76,571.7 person-years of follow-up (Table 1). The distribution of persons and deaths from circulatory disease for various demographic parameters are given in Table 2. We concentrate on various circulatory disease endpoints, namely: (a) all circulatory disease; (b) IHD; (c) cerebrovascular disease; (d) all other circulatory disease.

Dosimetry

The RT for this cohort was highly standardized, and has been described in detail elsewhere ^{10;11}. Several machines were used, but, with rare exceptions, these were orthovoltage X-ray machines with 250 kVp X-ray beams and effective energies equivalent

to 1.3–1.5 mm Cu half-value layers. Treatments were anterior and posterior parallel-opposed fields (typically 13 cm \times 13 cm), centred on the stomach and under fluoroscopic control since 1949. Radiotherapy was delivered in daily fractions of 1.5 Gy at a dose rate of 0.3 Gy/ min during one or two 6–14 day treatment courses; the goal, although it varied slightly over the years, was to provide a stomach dose of 16–17 Gy at each course. More than one treatment course was received by 182 (9.8%) of the 1859 irradiated patients.

Cardiac doses were estimated from measurements in an adult male Alderson phantom. The machine used to irradiate the phantom was one of the orthovoltage machines (General Electric Maxitron 250) used to treat the patients in the study. The beam was defined by the machine collimators; a 3 mm thick lead rubber cutout provided additional shielding of adjacent organs. Thermoluminescent dosimeters (TLD) were placed on a three-dimensional grid throughout the torso of the phantom.

A marked dose gradient occurred across the heart, decreasing with increasing distance from the edge of the treatment field. The physician who treated the patients estimated that about 5% (range 0-10%) of the heart (apex) was directly in the radiation field during all treatments ⁶¹¹. The remainder of the heart received scattered radiation.

We derived a volume-weighted average heart dose by summing the in-field dose received in the apex of the heart weighted by its proportion of the organ volume (0%, 5%, 10%) and the average of the dose received in the remainder of the heart weighted by its volume proportion (respectively 100%, 95%, 90%). The doses delivered to the stomach for individual patients were obtained from RT records. The measurements were renormalized on the basis of each patient's stomach dose to obtain their volume-weighted average dose to the heart. The total average cumulative cardiac dose for each patient was obtained by summing the volume-average cardiac dose over all treatment courses. In addition to the average cardiac doses, we estimated the dose to the thyroid, kidney, pancreas and brain from TLD measurements; thyroid dose serves as a surrogate for the dose to carotid arteries.

Statistical methods

A Poisson relative risk model was fitted by Poisson maximum likelihood. Tests were based on the likelihood-ratio test ¹². Confidence intervals are based on the profile likelihood ¹². Tests of homogeneity of excess relative risk Gy^{-1} (ERR) and other parameters modifying the ERR across subtypes of circulatory disease are based on the methods of Pierce and Preston ⁷. All *p*-values are two-sided.

Results

The optimal background model is given in Table A1. As with Carr *et al.* ⁶, we used tobacco smoking and alcohol habit/quantity in the background model for circulatory disease because of their previously known effects on certain circulatory disease subtypes, although these variables were not statistically significant.

Table 3 (model 1) demonstrates that there are statistically significant increases in risk with dose for all circulatory diseases combined (p=0.001) and, more specifically, for IHD (p=0.001). For all other subtype endpoints there is no significant dose response (p>0.20). However, a 2 degrees of freedom (df) test of the dose response for stroke adjusted for time since exposure was highly statistically significant (p=0.009–0.011 for various organ doses) (Table 3 model 5, Table 4 model 6). Risk modifications are similar if analysis is restricted to those receiving radiation, although risks are slightly larger (Tables 3, 4, Figure 1) and the dose response for stroke fails to be significant using the 2 df test (p=0.101)(result not shown). When stroke is assessed in relation to heart dose there are no significant differences

in the ERR between circulatory disease subtypes (p>0.20) (Table 5 panel A, Figure 1), but if thyroid dose is used for stroke there is significant (p=0.011) heterogeneity of ERR (Table 5 panel B). There are substantial differences in the magnitude of ERRs for IHD and stroke, which are relatively small for dose to the pancreas or kidney, more substantial for heart and thyroid dose, and very large for brain dose (Table 4).

Table 3 (model 5) and Figure 1 demonstrate that for all circulatory disease, IHD, and stroke there are highly statistically significant (p<0.005) decreasing trends of ERR with time after exposure, which changes by 100(exp(-0.075) -1) =-7.3% (95% CI -11.8, -4.4) per year after exposure; once adjustment is made for time, there are no significant modifications in risk by age at exposure (p>0.2) (model 6). There are no statistically significant (p>0.2) differences between the three endpoints in the trends with time since exposure of ERR (Table 5 panels A, B).

Discussion

The present analysis shows that radiation of the stomach for peptic ulcer increased the risk of circulatory disease as a function of dose received. There are significant excess risks for a variety of circulatory disease endpoints, in particular IHD, stroke, and all circulatory disease. There are significant reductions of ERR with increasing time since exposure. Risk modifications are similar if analysis is restricted to those receiving radiation, although relative risks are slightly larger and the risk of stroke fails to be significant; this does not suggest any marked difference in underlying health status between those selected for radiotherapy and those given surgery.

The reduction in relative risk with increasing time since exposure is somewhat novel. The LSS is the only group in which time- or age-at-exposure-modifications of the radiogenic excess circulatory disease risk has been modeled ^{4;8}. Although no time trends (or increasing ones) in aggregate SIR have been observed in a few other cohorts 13-15, the absence of individual dosimetry combined with the substantial temporal changes in mean heart dose for treatment of breast and other cancers ¹⁶ mean that these studies are largely uninformative on this issue. There are borderline significant (p=0.04) variations in ERR with attained age for stroke in the LSS, with indications of greater ERR for those under age 60, but no variations with exposure age or gender (as here)⁴. As here, Shimizu *et al.* documented no significant modifications of ERR for IHD by age, exposure age or gender in the LSS⁴, although further analyses using underlying and contributing causes of death (rather than just underlying cause of Shimizu et al.⁴) demonstrate significant reduction in ERR with increasing exposure age (Table A2). Because of the limited exposure-age spread in the peptic ulcer cohort there is little power to detect variations of risk by this variable. Although there are no time trends in the LSS data, the trends are statistically consistent with those in the peptic ulcer cohort (Tables 3, A2, Figure 1).

The suggestions of homogeneity of ERR and speed of variation of ERR over time by different circulatory disease subtypes is also a novel finding. Possibly because there were only relatively weak indications in the LSS of modifications of ERR by gender, age and exposure age ⁴, no formal analysis of heterogeneity of modification of risk was performed there. However, not too much should be made of this, since the only endpoints in the present study with significant dose response were IHD and stroke.

One limitation of the study is that the radiation dosimetry, although of high quality in many respects, fails to account for variability in patient anatomy, e.g., the heart size/shape/position and its relation to the diaphragm and stomach. The dose received by the heart for the "same" radiation technique (as recreated on the phantom) may vary markedly between patients.

However, the treatments were set up using fluoroscopy, in order to ensure stomach exposure and reduce exposure to other organs.

The magnitude of radiation-induced circulatory disease ERR, 0.082 - 0.194 Gy⁻¹, is consistent with the value of 0.11 - 0.15 Sv⁻¹ for this endpoint in the LSS ⁴. [The contrast between the present data, in which the excess is largely of IHD, and the LSS⁴, in which this endpoint is not significantly elevated should be noted; however, doubts as to the accuracy of death certificate coding in both cohorts suggest that one should not over-emphasise this possible discrepancy.] The ERRs are also consistent with those predicted by a recent metaanalysis of occupational and environmentally exposed groups, of 0.19 Sv^{-1} (95% CI 0.14. (0.23)¹⁷. The risks of stroke are somewhat uncertain; nevertheless, the indications of much larger ERR (2.649 – 10.53 Gy⁻¹) in relation to brain dose (Table 4) than those observed in the LSS 4 (0.12 Sv⁻¹) or in a number of groups exposed to moderate and low radiation doses ¹⁷ (0.27 Sv⁻¹), suggest (albeit weakly) that dose to the heart, thyroid, pancreas or kidney, or something strongly correlated with these, may be the more appropriate dose measure than brain dose for this endpoint, at least in the present dataset. This is also supported by indications of excess stroke risk in groups treated with RT for Hodgkin's disease and breast cancer ^{18;19}; the brain doses in such patients would be expected to be much lower than those to the heart, coronary and carotid arteries. Possibly it is the dose to some combination of the heart (or other nearby target in the circulatory system) and carotid that may be relevant for this and IHD. Radiation-induced renal damage is well documented²⁰, and can result in hypertension²¹ and very likely subsequent circulatory disease. There is increased risk of diabetes mellitus in this cohort²², likely due to radiation damage to the pancreas, which is located partly in the radiation treatment field. Overt or subclinical diabetes is an added factor in increasing the risk of circulatory disease.

The relevance of this study to radiation therapy regimes is intriguing. For a substantial part of the cohort the cumulative mean heart doses exceed 2 Gy (Figure 1), larger than the vast majority of those in the LSS cohort. The mean doses are more comparable with those received therapeutically, if not quite as large as those received by some groups, e.g., childhood cancer survivors ^{23;24}. Mulrooney et al.²³ and Tukenova et al.²⁴ demonstrate that there is significant excess risk of various types of circulatory disease in two groups of childhood cancer survivors receiving 15 Gy mean heart dose. Tukenova et al.²⁴ estimated an ERR of 0.6 Gy^{-1} (95% CI 0.2, 2.5), much higher than the risk we estimate, 0.082 Gy^{-1} (Table 3). The discrepancy likely reflects the younger exposure age in the cancer-survivor group, although as above, the present dataset has little power to detect such a variation. The calculations that we performed with the latest LSS data⁴ imply a 2.8% (95% CI 0.8, 4.9) reduction in circulatory disease ERR per year of exposure age (Table A2). Applying the central estimate (2.8%) to the ERR of Tukenova et al.²⁴ adjusted for the exposure-age difference between the two cohorts (5.9 vs 41.3 years) yields an ERR of 0.22 Gy⁻¹ (95% CI (0.07, 0.91), still higher than the estimate we obtain, (0.082 Gy^{-1}) , but statistically compatible with it. The ERR in the LSS ⁴ in a subset with similar exposure-age range (30–59 years) as the present study is 0.127 Sv^{-1} (95% CI 0.068, 0.188) (Table A2), again compatible with the present study.

Excess circulatory disease risk associated with high dose RT have been established for some time ¹, although in the absence of individual radiation dosimetry, quantitative evaluation risks in most RT data is problematic – some recent high quality studies remedy this defect ^{23;24}. At high radiation doses, such as would be received by patients treated with RT, a variety of so-called deterministic or tissue-reaction effects are observed ^{1;2}. Among such effects are damage to the structures of the heart – including marked diffuse fibrotic damage, especially of the pericardium and myocardium, pericardial adhesions, microvascular damage and stenosis of the valves – and to the coronary, carotid and other large arteries; these sorts

of damage occur frequently both in patients receiving RT and in experimental animals ^{1;25}. An association between lower doses (< 0.5 Gy) and late circulatory disease has only recently been suspected and remains controversial. Recent reviews present evidence suggesting an excess radiation-induced risk for IHD and stroke at occupational and environmental dose levels ^{9;17}. A review of much biological data suggested that many inflammatory endpoints potentially relevant to circulatory disease may be differentially regulated below and above about 0.5 Gy ⁹, emphasizing the importance of assessing risks in biological and epidemiological data below this value. Although uncertainties in dosimetry in the present cohort suggest a measure of caution, there is perhaps remarkable consistency between the risks in this data, in which a highly heterogeneous dose was delivered acutely to the heart in about 10 daily fractions with those in the LSS, in which a uniform whole body dose was also delivered acutely but in a single fraction, and with the risks in occupational series in which whole body dose was delivered at low dose rate in small daily fractions; this implies that there may be little sparing effect of low dose rate exposure or dose fractionation, and that similar circulatory disease mechanisms operate over this wide range of doses and dose rates.

Acknowledgments

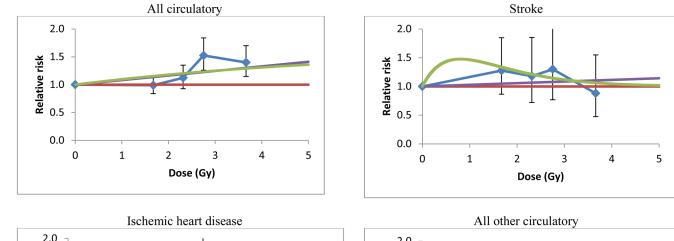
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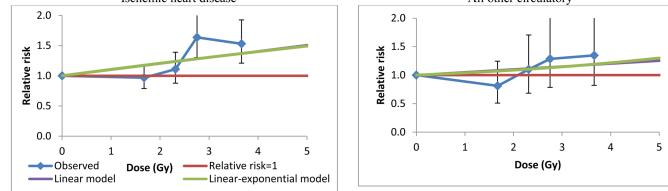
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All circulatory disease time since exposure trends

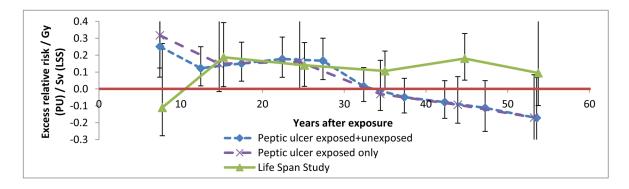


Figure 1. Relative risk vs dose (with 95% CI) for various circulatory disease endpoints in the peptic ulcer study, with fit of optimal linear and linear-exponential models, and time after exposure trends (adjusted for age at exposure) in aggregate and in the Life Span Study ⁴. All use average heart dose (5% of heart in beam) Background model is as in Tables 3+4.

Table 1

Numbers of deaths people and person years in US peptic ulcer cohort

Endpoint	Known radiation dose and >5 years of follow-up	Percentage of cardiovascular deaths
Ischemic heart disease (ICD9 410-414)	1003	68.3
Stroke (ICD9 430-438)	226	15.4
All other circulatory diseases	240	16.3
All circulatory disease (ICD9 390–459)	1469	100.0
Numbers of people	3600	
Person years	76,571.7	

Table 2

Numbers of circulatory disease deaths by radiotherapy status and distribution of other risk factors in US peptic ulcer cohort (all deaths 5 years after treatment)

Radiotherapy		y	No radiotherapy	
Category	Circulatory disease deaths	Persons at risk	Circulatory disease deaths	Persons at risk
Age at treatment				
<35	57	252	125	437
35–44	173	462	232	549
45–54	234	515	248	497
55	224	511	176	377
Year of treatment/ent	ry			
<1940	87	207	248	534
1940–44	147	373	199	459
1945–49	111	277	212	504
1950–59	296	738	122	363
1960	47	145	-	-
Gender				
Male	556	1389	585	1423
Female	132	351	196	437
Marital status				
Not stated/unknown	17	47	3	9
Never married	58	163	94	234
Married	546	1356	617	1471
Divorced separated	17	63	22	60
Widowed	50	111	45	86
Cigarette smoking sta	atus			
Unknown	120	301	140	369
Never smoked	156	418	239	499
Smoked	412	1021	402	992
Cigarette smoking qu	antity			
Unknown	291	749	395	903
1 pack/day	286	694	295	726
> 1 pack/day	111	297	91	231
Alcohol drinking stat	us			
Unknown	136	342	156	424
Never drank	218	574	308	642
Drank	334	824	317	794

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Radiotherapy		No radiotherapy		
Category	Circulatory disease deaths	Persons at risk	Circulatory disease deaths	Persons at risk
Unknown	397	1025	520	1214
5 drinks/week	169	404	167	379
6-15 drinks/week	59	143	42	123
>15 drinks/week	63	168	52	144

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Table 3

Modifiers of dose response for various categories of circulatory disease, among full cohort (exposed + unexposed) and among exposed group only. Parameter estimates (all Gy⁻¹) with 95% profile likelihood CI, unless otherwise indicated

Background model is quartic model in age, with adjustment for gender, smoking habit and quantity, alcohol habit and quantity, and calendar year period treated (year of entry).

				Exposed an	Exposed and unexposed		Exposed only	sd only
Model no.	Model/parameters		All circulatory disease	IHD	Stroke	Other circulatory	All circulatory disease	IHD
-		a (Gy ⁻¹)	$0.082\ (0.031,\ 0.140)$	$0.102\ (0.039,\ 0.174)$	0.028 (-0.085, 0.186)	0.050 (-0.053, 0.194)	$0.194\ (0.014,\ 0.648)$	0.376 (0.060, 1.735)
	Background[1+ $a D$]	<i>p</i> -values ^{<i>a</i>}	0.001	0.001	0.665	0.384	0.028	0.007
5		a (Gy ⁻¹)	0.104 (0.026, 0.287)	0.107 (0.025, 0.296)	1.604 (-404.3, 807.5)	$0.035 (-0143^b, 14.12)$	0.441 (0.042, 1.767 <i>b</i>)	1.043 (-0.657, 4.356b)
		$\gamma(\mathrm{Gy}^{-1)}$	-0.074 (-0.468, 0.197)	-0.016 (-0.391, 0.262)	-1.242 (-157.1, 0.852 ^b)	0.105 (-0.987b, 1.196b)	$-0.162 (-0.675, 0.140^{b})$	$-0.155 (-0.355, 0.086^{b})$
	Background[1+ $\alpha D \exp[\gamma D]$]	p -values $^{\mathcal{C}}$	0.599	0.920	0.123	0.834	0.212	0.152
ε		a (Gy ⁻¹)	0.089 (0.034, 0.151)	0.123(0.054, 0.204)	$0.004 \ (-0.120^{b}, 0.128^{b})$	0.051 (-0.056, 0.203)	0.248 (0.035, 0.935)	0.403 (0.060, 2.154)
		ĸ	-0.625 (-54.32, 0.877)	–1124 (NA, NA)	$3.532 (< -10^4, 35.56^b)$	-0.123(-5.647, 6.135b)	$-8.049 (-3469^{b}, 1.108)$	-0.535(-71.30, 3.233b)
	Background[1+ $\alpha D \exp[x_{1 \sin - female}]$]	p -values $^{\mathcal{C}}$	0.501	0.092	0.114	0.964	0.236	0.762
4		a (Gy ⁻¹)	0.062 (0.023, 0.113)	0.078 (0.028, 0.143)	0.016 (-0.050b, 0.137)	0.046 (-0.063, 0.190)	0.236 (0.067, 0.698)	0.453 (0.118, 2.047)
	Background[1+ $a D \exp[\tau(age at exposure - 41.32)]]^d$	$\boldsymbol{\tau}\left(\boldsymbol{y}^{-1}\right)$	0.051 (0.025, 0.082)	$0.050\ (0.023,\ 0.082)$	0.097 (-0.034 ^b , 0.228 ^b)	0.021 (-0.115 ^b , 0.157 ^b)	$0.040\ (0.023,\ 0.066)$	$0.035\ (0.020,\ 0.059)$
		<i>p</i> -values ^{<i>c</i>}	<0.001	<0.001	0.064	0.677	<0.001	<0.001
5		a (Gy ⁻¹)	0.115 (0.054, 0.181)	0.140 (0.067, 0.223)	0.107 (-0.085b, 0.299)	$0.004 \ (-0.048 \ b), 0.195)$	0.318 (0.115, 0.794)	0.596 (0.195, 2.250)
		$\delta(\mathbf{y}^{-1})$	-0.075 (-0.126, -0.045)	-0.066 (-0.113, -0.036)	-0.118(-0.395, -0.044)	-0.299 (-1.487b, 0.889b)	-0.048 (-0.086, -0.027)	-0.039 (-0.068, -0.021)
	Background[1+ $a D \exp[\delta(\text{time since exposure} - 20.78)]]d$	p -values $^{\mathcal{C}}$	<0.001	<0.001	0.003	0.270	<0.001	<0.001
		p -values $^{m{ heta}}$	<0.001	<0.001	0.009	0.372	<0.001	<0.001
9		a (Gy ⁻¹)	0.115 (0.053, 0.193)	$0.133\ (0.060,\ 0.229)$	0.150 (-0.125b, 0.554)	$0.001 \ (-0.020^{b}, 0.022^{b})$	0.367 (0.112, 1.270)	0.602 (0.150, 3.409)
		$r(y^{-1})$	-0.001 (-0.035, 0.033)	0.006 (-0.031, 0.044)	-0.024 (-0.124, 0.079)	-0.044 (-0.217b, 0.128b)	-0.013 (-0.076, 0.044)	-0.001 (-0.074, 0.076)
	Background[1+ $a D \exp[\tau(age at exposure -41.32) + \delta(time since exposure -20.78)]]^d$	$\delta(\mathbf{y}^{-1})$	-0.076 (-0.139, -0.036)	-0.061 (-0.119, -0.020)	-0.132 (-0.392, -0.035)	-0.434 (-1.723b, 0.856b)	-0.060 (-0.129, -0.000)	-0.040 (-0.112, 0.043)
	1	p -values f	<0.001	0.004	0.014	0.111	0.050	0.286
		p -values $^{\mathcal{S}}$	0.964	0.738	0.575	0.222	0.652	0.964

 ${}^{a}_{p}$ value for improvement of fit over model without linear term in dose

 $b_{Wald-based confidence limit}$

c p-value for improvement of fit over model with (unadjusted) linear term in dose

d age at exposure and time since exposure are approximately centered by subtracting off their person-year weighted means (41.32 years, 20.78 years respectively) in the full cohort (exposed+unexposed), to stabilize parameter estimates.

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 e^{-p} value for 2 degrees of freedom test of improvement of fit over model with no dose-response terms

f p-value for improvement of fit over linear model in dose with adjustment for age at exposure.

 \mathcal{E}_{P} value for improvement of fit over linear model in dose with adjustment for time since exposure

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Table 4

among full cohort (exposed + unexposed) and among exposed group only. Parameter estimates (all Gy⁻¹) with 95% profile likelihood CI, Modifiers of dose response for ischemic heart disease and stroke using dose to various organs (heart, thyroid, kidney, pancreas, brain), unless otherwise indicated Background model is quartic model in age, with adjustment for gender, smoking habit and quantity, alcohol habit and quantity, and calendar year period

Model no.	Model/parameters		Heart dose	Thyroid dose	Kidney dose	Pancreas dose	Brain dose
			Ischemic heart	Ischemic heart disease (exposed and unexposed)	osed)		
		a (Gy ⁻¹)	0.102 (0.039, 0.174)	1.696 (0.651, 2.907)	0.033 (0.012, 0.056)	$0.020\ (0.008,\ 0.035)$	10.36 (3.925, 17.82)
	Background[1+ α D]	<i>p</i> -values ^{<i>a</i>}	0.001	0.001	0.001	0.001	0.001
2		$a (\mathrm{Gy}^{-1})$	0.140 (0.067, 0.223)	2.338 (1.114, 3.719)	0.045 (0.021, 0.072)	0.028 (0.013, 0.044)	14.34 (6.800, 22.85)
		$\delta\left(\mathrm{y}^{-1} ight)$	-0.066(-0.113, -0.036)	-0.066(-0.113, -0.036)	-0.066(-0.114, -0.036)	-0.066 (-0.113, -0.036)	-0.066(-0.114, -0.036)
	Background[1+ $a D \exp[\delta(\text{time})]$	p-values b	<0.001	<0.001	<0.001	<0.001	<0.001
	since exposure $-20.78)$]] ^C	<i>p</i> -values ^{<i>d</i>}	<0.001	<0.001	<0.001	<0.001	<0.001
			Ischemic I	Ischemic heart disease (exposed only)			
3		$a (\mathrm{Gy}^{-1})$	0.376 (0.060, 1.735)	6.371 (1.041, 29.59)	0.114 (0.016, 0.514)	$0.078\ (0.013,\ 0.368)$	37.54 (5.681, 173.9)
	Background[1+ α D]	<i>p</i> -values ^{<i>a</i>}	0.007	0.006	0.010	0.006	0.008
4		a (Gy ⁻¹)	0.596 (0.195, 2.250)	9.991 (3.277, 38.42)	0.186 (0.061, 0.683)	0.122 (0.040, 0.478)	60.31 (19.60, 229.7)
		$\delta\left(\mathrm{y}^{-1} ight)$	-0.039 (-0.068, -0.021)	-0.039 (-0.068, -0.021)	-0.039 (-0.070, -0.021)	-0.038 (-0.067, -0.021)	-0.039 (-0.069, -0.021)
	Background[1+ α D exp[δ (time	p -values b	<0.001	<0.001	<0.001	<0.001	<0.001
	since exposure $-20.78)$]] ^C	<i>p</i> -values ^d	<0.001	<0.001	<0.001	<0.001	<0.001
			Stroke	Stroke (exposed and unexposed)			
5		a (Gy ⁻¹)	$0.028 \ (-0.085, \ 0.186)$	$0.422 \ (-1.455, 3.039)$	0.009 (-0.028, 0.060)	0.005 (-0.017, 0.036)	2.649 (-8.912, 18.740)
	Background[1+ $a D$]	<i>p</i> -values ^{<i>a</i>}	0.665	0.698	0.678	0.704	0.692
9		$a~(\mathrm{Gy}^{-1})$	$0.107 (-0.085^{e}, 0.299)$	1.734 (–1.437 <i>°</i> , 4.923)	$0.034 \ (-0.028^{e}, 0.097)$	0.020 (-0.017 ^e , 0.058)	$10.53 (-8.940^{e}, 30.35)$
		$\delta\left(\mathrm{y}^{-1} ight)$	-0.118(-0.395, -0.044)	-0.118 (-0.411, -0.044)	-0.118 (-0.396, -0.044)	-0.119 (-0.411, -0.044)	-0.120 (-0.413, -0.044)
	Background[1+ $a D \exp[\delta(\text{time})]$	p-values b	0.003	0.003	0.003	0.003	0.003

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 $\stackrel{b}{P}$ value for improvement of fit over model with (unadjusted) linear term in dose

 c_i time since exposure is approximately centered by subtracting off its person-year weighted mean (20.78 years) in the full cohort (exposed+unexposed), to stabilize parameter estimates.

 d_{P} value for 2 degrees of freedom test of improvement of fit over model with no dose-response terms

^eWald-based confidence limit

Table 5

Test of heterogeneity in dose response and its modification by time since exposure and age at exposure in various categories of circulatory disease and cancer

The background model for each circulatory disease subtype is a quartic model in age, with adjustment for gender, smoking habit and quantity, alcohol habit and quantity, and calendar year period treated (year of entry). Unless otherwise stated heart dose is used for all circulatory disease endpoints.

Panel	Model	<i>p</i> -value
A	Test of heterogeneity across four circulatory disease endpoints (using hear	t dose throughout), linear dose response
	Background[1+ $a D \exp[\delta$ (time since exposure - 20.78)]] ^a	<0.001 ^b
	Background[1+ $a_i D \exp[\delta$ (time since exposure - 20.78)]] ^a	$0.283^{\mathcal{C}}$
	Background[1+ $a_i D \exp[\delta_i$ (time since exposure - 20.78)]] ^a	0.372^{d}
В	Test of heterogeneity across four circulatory disease endpoints (using thyroid dos response	se for stroke, heart dose otherwise), linear dose
	Background[1+ $a D \exp[\delta$ (time since exposure - 20.78)]] ^a	<0.001 ^b
	Background[1+ $a_i D \exp[\delta$ (time since exposure - 20.78)]] ^a	$0.011^{\mathcal{C}}$
		0.370^{d}

^a time since exposure is approximately centered by subtracting off its person-year weighted mean (20.78 years) in the full cohort (exposed +unexposed), to stabilize parameter estimates.

b p-value for improvement of fit over model without radiation dose response term.

 ^{c}p -value for improvement of fit over model with the same excess relative risk coefficient (*a*) for each endpoint.

 d_{p} -value for improvement of fit over model with the same time trend in excess relative risk (δ) for each endpoint.