

CHEST

SLEEP DISORDERS

Esophageal Pressures, Polysomnography, and Neurobehavioral Outcomes of Adenotonsillectomy in Children

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Background: Esophageal pressure monitoring during polysomnography in children offers a gold-standard, "preferred" assessment for work of breathing, but is not commonly used in part because prospective data on incremental clinical utility are scarce. We compared a standard pediatric apnea/hypopnea index to quantitative esophageal pressures as predictors of apnearelated neurobehavioral morbidity and treatment response.

Methods: Eighty-one children aged 7.8 ± 2.8 (SD) years, including 44 boys, had traditional laboratory-based pediatric polysomnography, esophageal pressure monitoring, multiple sleep latency tests, psychiatric evaluations, parental behavior rating scales, and cognitive testing, all just before clinically indicated adenotonsillectomy, and again 7.2 ± 0.8 months later. Esophageal pressures were used, along with nasal pressure monitoring and oronasal thermocouples, not only to identify respiratory events but also more quantitatively to determine the most negative esophageal pressure recorded and the percentage of sleep time spent with pressures lower than -10 cm H_2O .

Results: Both sleep-disordered breathing and neurobehavioral measures improved after surgery. At baseline, one or both quantitative esophageal pressure measures predicted a disruptive behavior disorder (*Diagnostic and Statistical Manual of Mental Disorders, Fourth Edition*-defined attention-deficit/hyperactivity disorder, conduct disorder, or oppositional defiant disorder) and more sleepiness and their future improvement after adenotonsillectomy (each P < .05). The pediatric apnea/hypopnea index did not predict these morbidities or treatment outcomes (each P > .10). The addition of respiratory effort-related arousals to the apnea/hypopnea index did not improve its predictive value. Neither the preoperative apnea/hypopnea index nor esophageal pressures predicted baseline hyperactive behavior, cognitive performance, or their improvement after surgery. *Conclusions:* Quantitative esophageal pressure monitoring may add predictive value for some, if not all, neurobehavioral outcomes of sleep-disordered breathing.

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Abbreviations: AASM = American Academy of Sleep Medicine; DSM IV = *Diagnostic and Statistical Manual of Mental Disorders, Fourth Edition*; low Pes time = percentage of sleep epochs spent with most esophageal pressure swings more negative than -10 cm H_2 O; Pes nadir = most negative esophageal pressure swing recorded during a polysomnogram; RERA = respiratory effort-related arousal; SDB = sleep-disordered breathing

E sophageal pressure monitoring during sleep was recognized by 1982 to identify abnormalities in breathing that are more subtle than frank apneas and hypopneas, yet are still responsible for serious neurobehavioral and mental health consequences in children.¹ Esophageal pressure monitoring also offers a gold-standard assessment of upper airway resistance and work of breathing during polysomnography.² However, exactly how esophageal pressure monitoring should be used has not been well defined. Early literature on the upper airway resistance syndrome illustrated use of both quantitative measurements, such as the most negative esophageal pressure recorded during a polysomnogram, and respiratory effort-related arousals (RERAs) identified by more qualitative crescendo increases in esophageal pressure swings prior to arousals.^{1,3-7}

In practice, nasal pressure monitoring is simpler than esophageal pressure monitoring and has been adopted more commonly to identify rates of hypopneas and

RERAs,⁸ which in combination with rates of obstructive apneas offer an overall respiratory disturbance index with increased sensitivity to the subtlest forms of pediatric sleep-disordered breathing (SDB). Although one report has suggested that nasal pressure in comparison with esophageal pressure monitoring identifies RERAs equally well,⁹ with data derived from 15 adults rather than children, studies have rarely examined the clinical value of quantitative esophageal pressures, and the two monitoring approaches have not been compared for prediction of morbidity or treatment outcomes. More broadly, the degree to which any data beyond frank obstructive apneas effectively predict pediatric SDB-related morbidity or treatment response remains inadequately studied. Current guidelines for polysomnographic scoring of events more subtle than apneas have been established largely by consensus for the pediatric age group.⁸ Esophageal pressure monitoring is listed as "preferred" for identification of RERAs, but its unique potential as a more quantitative measure is not mentioned.

We prospectively used both esophageal and nasal pressure monitoring during polysomnography, and paired results with intensive evaluations of behavior, cognition, mental health, and sleepiness before and after clinically indicated adenotonsillectomies. Approximately two-thirds of the children in this Washtenaw County Adenotonsillectomy Cohort II tolerated esophageal pressure monitoring for ≥ 2 h of sleep before and after surgery. The primary hypothesis was that quantitative esophageal pressures would help to predict baseline neurobehavioral morbidity and treatment outcomes. Secondary hypotheses were that RERAs and hypopneas would augment the predictive value of obstructive apneas.

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MATERIALS AND METHODS

Overview

Subjects were recruited from local otolaryngology practices. Staff helped to identify potential subjects, aged 3.0 to 12.9 years, who were scheduled for adenotonsillectomy for any clinical indication, but as usual,¹⁰ were not thought to need polysomnography prior to surgery. Sleep and neurobehavioral assessments were completed up to 3 days before adenotonsillectomy, and again about 6 months thereafter. A child psychiatrist, child psychologist, or behavioral developmental pediatrician interviewed each family prior to a full, nocturnal, laboratory-based polysomnogram. The next day, a multiple sleep latency test was administered. Between naps, children underwent neuropsychologic testing, and a parent completed behavioral rating scales and a standard socioeconomic survey.¹¹

As detailed online (e-Appendix 1), pediatric polysomnography followed standard recommendations,⁸ published after the start of this research protocol, except that piezoelectric strain gauges rather than inductance plethysmography were used to monitor thoracic and abdominal excursion. Both oronasal thermocouples and nasal pressure monitoring were used to assess airflow. Esophageal pressure was recorded through a water-filled, 6F pediatric feeding tube.^{12,13} Multiple sleep latency tests followed standard procedures,¹⁴ but nap opportunities were lengthened from the adult standard of 20 min to 30 min.^{15,16}

Scoring

All sleep studies were scored, or in a minority of cases, thoroughly rescored, by a single pediatric-experienced sleep and electroencephalography-registered chief technologist. Sleep staging followed standard criteria.8 Esophageal pressure swings were scored as the difference between peak inspiratory and peak expiratory readings. The technologist recorded the percentage of sleep epochs spent with most esophageal pressure swings more negative than -10 cm H_2O (low Pes time), and also the most negative esophageal pressure swing (Pes nadir) for each study. Obstructive apneas (2 breaths or longer), hypopneas, RERAs scored by nasal pressure or esophageal pressure, and central apneas were scored according to pediatric criteria recommended by the American Academy of Sleep Medicine (AASM) in 2007.8 The AASM-2007 apnea/hypopnea index was calculated as the number of apneas and hypopneas per hour of sleep, and the AASM-2007 respiratory disturbance index was calculated similarly, with the addition of RERAs. In multiple sleep latency tests, the mean sleep latency across all nap opportunities provided an objective measure of daytime sleepiness.17

Neurobehavioral Outcomes

Standardized, well-validated assessments were used to evaluate *Diagnostic and Statistical Manual of Mental Disorders*, *Fourth Edition* (DSM-IV) diagnoses, behavioral problems, and cognitive deficits long thought to represent the most important morbidity in childhood SDB.^{1,18-23} Psychiatric diagnoses were guided in large part by the Computerized Diagnostic Interview Schedule for Children–Parent,²⁴⁻²⁶ and the Children's Psychiatric Rating Scale.²⁷⁻²⁹ The final categorical diagnostic outcome variable was presence or absence of a DSM-IV-defined disruptive behavior disorder—attention-deficit/hyperactivity disorder, conduct disorder, or oppositional-defiant disorder—as concluded by the interviewing clinician. Behavioral outcome variables were generated by two validated parental rating scales for inattention and hyperactivity: the Conners' Parent Rating Scales³⁰ and the Child Symptom Inventory-4³¹ (or the Early Childhood Inventory-4³² for

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children between 3 and 5 years). A behavioral hyperactivity index³³ (mean 50; SD 10; with higher scores indicating more significant symptoms) was constructed from the average of the inattention and hyperactivity T scores produced by the two instruments. Finally, cognitive testing for a total of about 2 h included the NEPSY,³⁴ a developmental neuropsychologic test battery created for children ages 3 to 12 years of age: a memory and learning score and an attention/executive functions score were averaged to create a cognitive index (mean 100; SD 15; with higher scores indicating better performance).

Analyses

Baseline and follow-up measures were compared initially using a nonparametric signed rank test for paired continuous data or the McNemar test for dichotomous data. Associations between baseline sleep and neurobehavioral morbidity measures were assessed with the nonparametric Spearman correlation coefficient ρ , or by Wilcoxon two-sample tests. Each association that showed statistical significance in these bivariate comparisons was then examined in a multiple regression of the neurobehavioral measure at baseline, or its change after adenotonsillectomy, on the baseline sleep measure. General linear models or logistic regression models were used as appropriate to take several potential confounders into account. The level of significance was set at P < .05.

RESULTS

Subjects

Subject recruitment (Fig 1) generated the sample of 81 children whose demographics and findings are summarized in Table 1. Clinically suspected SDB was an indication for adenotonsillectomy for 70 (86%), and only 18 (22%) had or also had chronic tonsillitis or throat infections. The baseline AASM-2007 apnea/ hypopnea index ranged from 0.20 (normal < 1) to 81.20 (severe). The AASM-2007 apnea/hypopnea index on average was seven times higher than the obstructive apnea index. The AASM-2007 respiratory disturbance index, which added RERAs to the apnea/hypopnea index, exceeded the latter on average by only 8.5%. The obstructive apnea index, apnea/hypopnea index, and respiratory disturbance index were strongly intercorrelated ($\rho = 0.70$ for the first and second variable, $\rho = 0.68$ for the first and third, and $\rho = 0.99$ for the second and third, all P < .0001).

At baseline, at least one disruptive behavior disorder was diagnosed in 28 of the 81 children (35%): 25 (31%) had attention-deficit/hyperactivity disorder, 12 (15%) had oppositional defiant disorder, and none had conduct disorder. The 81 subjects had somewhat elevated behavioral hyperactivity indices, 0.67 SD above age- and sex-specific norms. Cognitive indices reflected performance just better than average. The average mean sleep latency of 24 ± 6 min on the 30-min multiple sleep latency test was similar 899 families were approached about the research at a visit to an otolaryngologist



81 children (including 44 boys and 37 girls) tolerated esophageal pressure monitoring at baseline during at least 2 hours of sleep (mean of 563 ± 108 minutes)

FIGURE 1. Recruitment of 81 subjects who provided data for the current report. Among 147 children whose families agreed to participate in the study, comparisons between the 81 children whose full sleep and neurobehavioral outcome data (at baseline and follow-up) were available for analysis in this study, and the 66 children excluded because full data were not available, revealed that the following baseline variables showed no statistically significant difference (all P > .05, Wilcoxon signed rank test or $\dot{\chi}^2$, as appropriate): sex, socioeconomic status, BMI z score, total baseline polysomnogram recording time, sleep latency, sleep efficiency, arousal index, percent of time spent in each sleep stage, obstructive apnea index, apnea/hypopnea index, respiratory disturbance index, minimum oxygen saturation, percentage of sleep time spent with oxygen saturation < 90%, most negative recorded esophageal pressure (available for some time in n = 15 excluded children), percentage of sleep time spent with esophageal pressure more negative than $-10 \text{ cm H}_2\text{O}$, behavioral hyperactivity index, cognitive index, and mean sleep latency on the multiple sleep latency test. In contrast, the following variables did differ $(P{<}.05)$ at baseline between the two groups of 81 and 66 children: age $(7.8 \pm 2.8 \text{ y vs } 6.4 \pm 1.8 \text{ y})$, race (84% vs 67% white), total sleep time $(513 \pm 62 \text{ min vs } 534 \pm 46 \text{ min})$, and frequency of a disruptive behavior disorder diagnosis (35% vs 58%).

to that published previously for a series of children with obstructive sleep apnea, primary snoring, or no SDB. $^{15}\,$

Follow-up studies were performed 0.60 ± 0.07 years (7.2 months) after baseline studies (range 0.46 to 0.80 years). Postoperative apnea severity measures improved (Table 1), but on average still met diagnostic criteria for pediatric sleep apnea.³⁵ Respiratory event indices (Fig 2) improved more dramatically than did the minimum oxygen saturation or the two quantitative esophageal pressure measures (Fig 3). Neurobehavioral measures also improved, to an extent that ranged from robust (eg, 51% reduction in frequency of disruptive behavior disorders) to marginally nonsignificant (1 minute on the multiple sleep latency test).

Sleep Apnea Measures and Neurobehavioral Morbidity at Baseline

The 28 children with baseline disruptive behavior disorders, in comparison with remaining children, on average had a low Pes time of 61% rather than 44% (P = .010, Table 2). These 28 also trended toward higher obstructive apnea indices (P = .090),

Table 1—Demographic, Sleep, and Neurobehavioral Measures at Baseline and Follow-up After Adenotonsillectomy
for 81 Children

Variable	Baseline Mean (SD)	Follow-up Mean (SD)	Effect Size	P Value ^a
Demographics				
Age, y	7.8 (2.8)	8.4 (2.8)		
Sex, male, %	54			
Race, white, %	84			
Socioeconomic status ^b	2.4 (0.9)			
BMI z score	0.52 (1.36)	0.81 (1.04)	0.24	<.0001
Sleep				
Total recording time, min	597 (38)	598 (44)	0.02	.9489
Total sleep time, min	513 (62)	528 (57)	0.24	.0746
Sleep latency, min	26 (23)	23 (31)	0.13	.0193
Sleep efficiency, ^b %	86 (8)	88 (7)	0.30	.0118
Arousal index ^b	12 (6)	9 (3)	0.60	<.0001
Stage 1 sleep, %	12 (5)	9 (3)	0.62	<.0001
Stage 2 sleep, %	44 (7)	47 (7)	0.44	<.0001
Stage 3 sleep, %	25 (5)	23 (6)	0.25	.0105
REM sleep, %	19(5)	20(4)	0.20	.1198
Sleep apnea measures				
Obstructive apnea index ^b	0.97 (3.15)	0.17 (0.27)	0.47	<.0001
Apnea/hypopnea index ^b	6.9 (10.9)	1.9(1.7)	0.79	<.0001
Respiratory disturbance index ^b	7.5 (11.1)	2.1(1.8)	0.84	<.0001
Minimum oxygen saturation, %	90.1 (6.1)	92.2 (4.0)	0.42	<.0001
Sleep time with oxygen saturation < 90%, %	0.55 (2.92)	0.02 (0.16)	0.34	.0684
Minimum esophageal pressure, cm H ₂ O	-37(17)	-22(9)	1.11	<.0001
Sleep time with esophageal pressure more	50 (32)	27 (31)	0.74	<.0001
negative than $-10 \text{ cm H}_{2}O$, %				
Neurobehavioral measures				
Disruptive behavior disorder, ^b % with positive diagnosis	35	17		.0017
Behavioral hyperactivity index	57 (11)	53 (9)	0.39	<.0001
Cognitive index	103 (13)	112 (12)	0.73	<.0001
Mean sleep latency on multiple sleep latency test, min	24 (6)	25 (5)	0.23	.0516

REM = rapid eye movement.

^aNonparametric signed rank test for difference between baseline and follow-up (for continuous data), or the McNemar test (for disruptive behavior disorder categorical diagnosis at baseline vs follow-up).

^bIndex = events per hour of sleep, scored following American Academy of Sleep Medicine 2007 recommendations; disruptive behavior disorder = *Diagnostic and Statistical Manual of Mental Disorders, Fourth Edition*-consistent, psychiatrist- and scheduled interview-determined diagnosis of attention-deficit/hyperactivity disorder, oppositional defiant disorder, or conduct disorder; sleep efficiency = (total sleep time/total recording time); socioeconomic status = Hollingshead (1965) rank (1 = highest, 5 = lowest).

but the difference between group means was small $(1.00 \pm 2.66 \text{ vs } 0.96 \pm 3.40)$. Trends emerged for mean sleep latency (sleepiness) to be predicted by the previous-night obstructive apnea index, percent of sleep time with oxygen saturation < 90%, Pes nadir, and low Pes time. The behavioral hyperactivity index and cognitive index were not predicted by any previous-night SDB measures.

Logistic regression of disruptive behavior disorder (present vs absent) on both low Pes time and the AASM-2007 apnea/hypopnea index showed that the former, but not the latter, retained statistical significance (OR = 1.023, 95% CI [1.005, 1.042], vs 0.969 [0.913, 1.028]). Thus, on average after accounting for the AASM-2007 apnea/hypopnea index, each 10 percentage point increase in low Pes time resulted in a 23% increase in the odds of having a disruptive behavior disorder. Similar results

104

emerged from logistic regression of disruptive behavior disorder on low Pes time and the obstructive apnea index (1.020 [1.003, 1.037] vs 0.940 [0.793, 1.113]). Logistic regression of disruptive behavior disorder on several potential confounders simultaneously—age, sex, white non-Hispanic vs other race/ethnicity, socioeconomic status, and BMI z score—in addition to low Pes time rendered the latter variable only marginally nonsignificant (1.015 [0.998, 1.032]). No other explanatory variable showed a significant association in this model.

Prediction of Neurobehavioral Improvement After Adenotonsillectomy

The 17 children whose disruptive behavior disorder was destined to resolve by follow-up, in comparison with the remaining 64 children, had increased

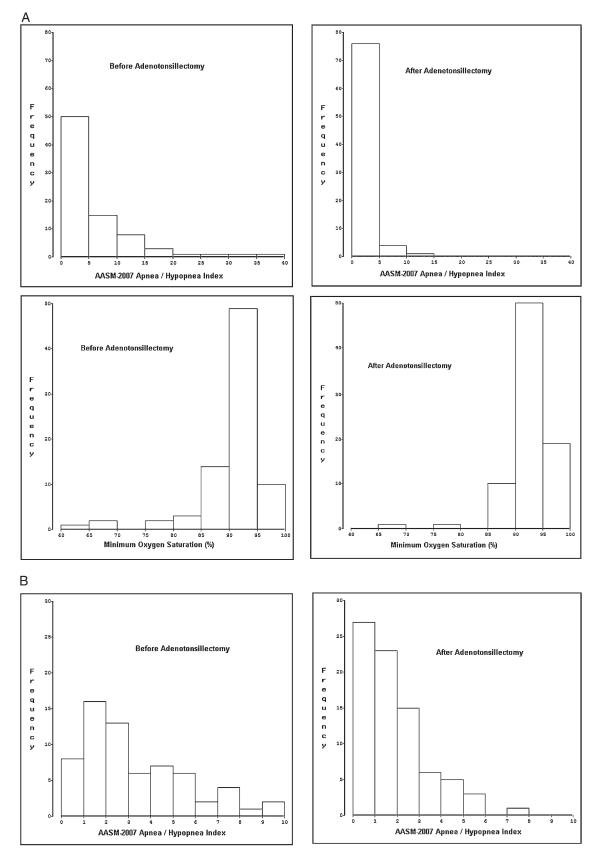


FIGURE 2. Distributions of two common measures of apnea severity during baseline and postadenotonsillectomy polysomnograms in 81 children. A, The AASM-2007 apnea/hypopnea index (first row) and minimum oxygen saturation (second row). One subject's apnea/hypopnea index (81.2) is not shown. B, Distributions in finer detail for AASM-2007 apnea/hypopnea indices of 10 or less. AASM = American Academy of Sleep Medicine.

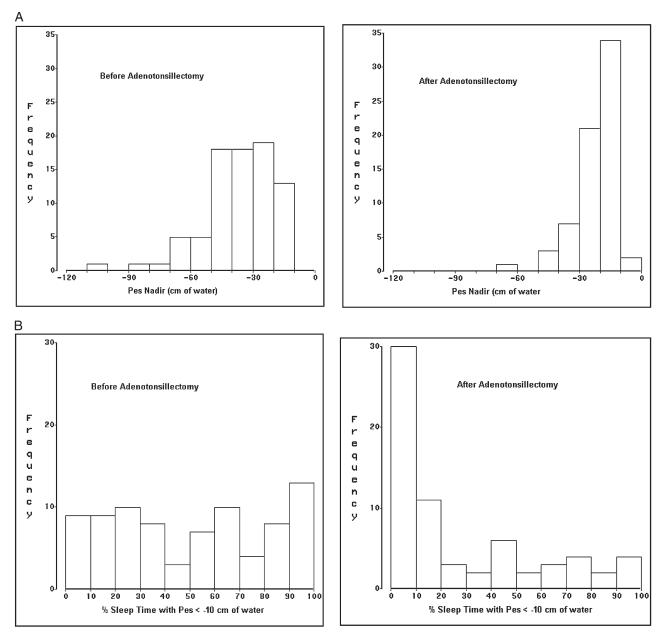


FIGURE 3. Distributions of two quantitative Pes measures during baseline and postadenotonsillectomy polysomnograms in 81 children. A, Pes nadir values before and after adenotonsillectomy. B, Percentage of sleep time spent with Pes more negative than $-10 \text{ cm H}_2\text{O}$ before and after adenotonsillectomy. Pes = esophageal pressure; Pes nadir = most negative esophageal pressure swing recorded during a polysomnogram.

baseline low Pes time (64% vs 47%, P = .021) and a trend toward lower Pes nadirs (-42 vs -35 cm H₂O, P = .087, Table 3). No other baseline SDB measures predicted which children would experience resolution. Improvement in mean sleep latency was predicted by the baseline obstructive apnea index and Pes nadir; low Pes time showed a trend. Improvements in behavior and cognition were not predicted significantly by any sleep apnea measure.

Logistic regression of disruptive behavior disorder resolution, from baseline to follow-up, on both low Pes time and the AASM-2007 apnea/hypopnea index showed that only low Pes time retained independent predictive value (OR = 1.023 [1.002, 1.045] vs 0.972 [0.908, 1.041]). Similar results emerged from regression on low Pes time and the obstructive apnea index (1.020 [1.001, 1.041] vs 0.943 [0.771, 1.153]). Regression of disorder resolution on low Pes time in addition to the potential confounders showed marginally nonsignificant predictive value for low Pes time (1.017 [0.997, 1.037]), but no other explanatory variable.

A general linear model with change in mean sleep latency as the outcome (follow-up minus baseline value)

Table 2—Unadjusted Associations Among 81 Children Between Sleep Apnea Measures at Baseline and Concurrent	1
Neurobehavioral Morbidity	

Apnea Measure	Disruptive Behavior Disorder, Present vs Absent		Behavioral Hyperactivity Index		Cognitive Index		Mean Sleep Latency on Multiple Sleep Latency Test	
	Wilcoxon Z	One-sided P Value	Spearman p	P Value	Spearman p	P Value	Spearman p	P Value
Obstructive apnea index ^a	-1.35^{b}	$.090^{\mathrm{b}}$	-0.16	.16	-0.066	.56	-0.19^{b}	.083b
Apnea/hypopnea indexª	0.0099	.50	-0.022	.85	-0.082	.47	-0.078	.49
Respiratory disturbance index ^a	0.0000	.50	-0.007	.95	-0.119	.29	-0.061	.59
Minimum oxygen saturation	-0.69	.24	-0.102	.37	0.011	.92	0.16	.15
Sleep time with oxygen saturation < 90%, %	0.66	.25	0.047	.68	-0.094	.40	-0.19^{b}	$.083^{b}$
Esophageal pressure nadir	-0.90	.18	-0.030	.79	-0.028	.80	0.19^{b}	$.092^{b}$
Sleep time with esophageal pressure below $-10 \text{ cm H}_2\text{O}, \%$	2.36 ^b	.010 ^b	0.097	.39	-0.176	.12	-0.19^{b}	.087 ^b

^aIndex = events per hour of sleep, scored following American Academy of Sleep Medicine 2007 recommendations. ^bResults that show statistical significance or trends.

and both Pes nadir and AASM-2007 apnea/hypopnea index as explanatory variables showed that the Pes nadir retained significance ($\beta = -0.091$, SE = 0.043, t = -2.10, P = .039) whereas the AASM-2007 apnea/hypopnea index did not ($\beta = -0.00050$, SE = 0.067, t = -0.01, P = .99). Similarly, regression of mean sleep latency change on both Pes nadir and the obstructive apnea index showed that Pes nadir retained a trend toward association ($\beta = -0.076$, SE = 0.040, t = -1.91, P = .060) whereas the obstructive apnea index did not ($\beta = -0.076$, SE = 0.040, t = -1.91, P = .060) whereas the obstructive apnea index did not ($\beta = 0.141$, SE = 0.212, t = 0.67, P = .51). Regression of mean sleep latency

change on baseline Pes nadir and the potential confounders showed predictive value for Pes nadir $(\beta = -0.094, SE = 0.033, t = -2.88, P = .0052; model R^2 = 0.18)$ but no other explanatory variable.

DISCUSSION

This sample of 81 children studied intensively just before and about 6 months after adenotonsillectomy shows that quantitative esophageal pressures during baseline polysomnography can help to identify children who have disruptive behavior disorders, and

 Table 3—Unadjusted Associations Among 81 Children Between Sleep Apnea Measures at Baseline and Changes in

 Neurobehavioral Outcomes (Follow-up Minus Baseline)

	Resolution of Disruptive Behavior Disorderª		Change in Behavioral Hyperactivity Index		Change in Cognitive Index		Change in Mean Sleep Latency on Multiple Sleep Latency Test	
Baseline Apnea Measure	Wilcoxon Z	One-sided P Value	Spearman p	P Value	Spearman p	P Value	Spearman p	P Value
Obstructive apnea index ^b	-1.12	.13	0.125	.27	0.076	.50	0.22°	.045°
Apnea/hypopnea index ^b	0.15	.44	0.049	.66	0.077	.49	0.15	.19
Respiratory disturbance index ^b	0.28	.39	0.029	.80	0.11	.33	0.14	.22
Minimum oxygen saturation	-1.02	.16	0.095	.40	0.013	.91	-0.14	.22
Sleep time with oxygen saturation < 90%, %	0.11	.46	0.054	.63	0.070	.54	0.17	.12
Esophageal pressure nadir	-1.37°	$.087^{\circ}$	0.076	.50	0.090	.42	-0.30°	.0059°
Sleep time with esophageal pressure below $-10 \text{ cm H}_2\text{O}, \%$	2.08°	.021°	-0.104	.36	0.095	.40	0.21°	.055°

"Children whose disruptive behavior disorder at baseline resolved by follow-up (coded as 1, n = 17) were compared with all other children (coded as 0, n = 64).

^bIndex = events per hour of sleep, scored following American Academy of Sleep Medicine 2007 recommendations.

eResults that show statistical significance or trends.

children destined to benefit after surgery from resolution of those psychiatric disorders or improvement in excessive daytime sleepiness. In contrast, baseline esophageal pressure measures did not help to predict concurrent parent-rated hyperactive behavior; a composite cognitive measure of attention/executive functioning and memory/learning; or their improvement after surgery. Whenever SDB measures did predict neurobehavioral morbidity or treatment outcomes, quantitative esophageal pressures appeared to have more robust and independent utility than did standard rates of apneic events. Esophageal pressure monitoring was tolerated by most research volunteers in the larger cohort, as in the past,¹³ and again proved safe and easy to accomplish by sleep technologists. These observations, in perhaps the largest prospective sample of children studied to date with a gold-standard polysomnographic measure of work of breathing, provide insight into the strengths as well as limitations of laboratory-based sleep studies that are commonly recommended,^{36,37} if not always obtained,^{10,38} prior to treatment of suspected SDB in childhood.

Neurobehavioral problems are some of the main morbidities associated with childhood SDB.35,36 These problems show substantial improvement after adenotonsillectomy, but prior studies without quantitative esophageal pressure monitoring have largely failed to show associations between baseline apnea severity measures and neurobehavioral response to surgery. Current results are also notable because they suggest that considerable effort to identify subtle apneic events in children may not always improve clinical use. In our study, the pediatric obstructive apnea index showed predictive value as good or better than that of the AASM-2007 pediatric apnea/hypopnea index, which identified 7 times as many respiratory events. In contrast, among adults, hypopneas in comparison with apneas predict daytime sleepiness with equal strength, if less reliably.³⁹ Our data further suggest that a respiratory disturbance index, including RERAs scored using both nasal and esophageal pressure monitoring,⁸ may identify negligible numbers of respiratory events beyond those captured by a highly sensitive pediatric apnea/hypopnea index derived from nasal pressure monitoring and 2-breath hypopnea durations. The AASM-2007 pediatric hypopnea and RERA scoring rules, which by necessity are based more on consensus than outcome-based evidence,⁸ may merit revision as more data become available.

On the other hand, our findings do suggest that quantitative esophageal pressures can supplement qualitative data commonly recorded from nasal and oral thermocouples, nasal pressure, and chest and abdominal excursion. Prolonged increased work of breathing, without discrete apneic events, can still be associated with sleepiness that improves with continuous positive airway pressure.^{1,4,40} Indeed, computeridentified respiratory cycle-related EEG changes, thought to represent nonvisible inspiratory microarousals outside regions of apnea or hypopnea, appear to worsen with increasing apnea severity,^{41,42} correlate with esophageal pressure swings,⁴³ diminish with treatment of SDB,⁴¹ and explain sleepiness beyond that accounted for by the apnea/hypopnea index.⁴⁴

Our study has several strengths, including the sample size given relatively intense evaluations of sleep, mental health, behavior, and cognition; the follow-up rate at 7 months (95%); and the rigorous scoring procedures. However, this study also has limitations. Most, but not all, children tolerated esophageal pressure monitoring, and subjects with and without complete data were not identical on every measure (Fig 1). Assessment of esophageal pressures by additional approaches not planned for this study, for example, focus on the ratio for any individual between esophageal pressure during sleep and quiet wakefulness, still merit exploration and could further improve predictive use. Nasal pressure was recorded with a transducer made for this purpose (e-Appendix 1), but high frequency filtering built into the transducer to eliminate vibration from snoring could have impeded optimal identification of flow limitation. Finally, only neurobehavioral outcomes were examined in our study, though these are among the most salient³⁶ and sensitive to subtle SDB that characterizes many children treated for the condition.³³

Our study does confirm meaningful neurobehavioral improvement after adenotonsillectomy, but the nonrandomized design cannot prove surgery was the cause. The ongoing randomized Childhood Adenotonsillectomy Study may provide more conclusive evidence of cause-and-effect,45 but does not monitor esophageal pressures or score RERAs. If the Childhood Adenotonsillectomy Study also finds significant neurobehavioral improvement after surgery, the field may still be left with dual imperatives to improve understanding of mechanisms by which SDB affects brain function, and to devise measures that allow more informative and predictive diagnostic assessments. In the meantime, our results suggest that esophageal pressure monitoring should be considered in clinical sleep laboratory evaluations when a child with suspected SDB also has a disruptive behavior disorder or excessive daytime sleepiness that could motivate a decision to pursue treatment. Use of quantitative esophageal pressure monitoring is unlikely to have significant negative impact,¹³ but could provide the most useful polysomnographic predictors of whether such a child will experience neurobehavioral benefit from adenotonsillectomy.

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Dr Chervin: contributed to designing and conducting the study, managing and analyzing data, and drafting or editing the manuscript.

Dr Ruzicka: contributed to designing and conducting the study, managing and analyzing data, and drafting or editing the manuscript. *Dr Hoban:* contributed to designing the study, interpreting sleep studies, and drafting or editing the manuscript.

Ms Fetterolf: contributed to designing the study, scoring sleep studies, and drafting or editing the manuscript.

Dr Garetz: contributed to designing and conducting the study, and drafting or editing the manuscript.

Mr Guire: contributed to designing the study, managing and analyzing data, and drafting or editing the manuscript.

Dr Ďillon: contributed to designing and conducting the study, and drafting or editing the manuscript.

Dr Felt: contributed to designing and conducting the study, and drafting or editing the manuscript.

Dr Hodges: contributed to designing and conducting the study, and drafting or editing the manuscript.

Dr Giordani: contributed to designing and conducting the study, and drafting or editing the manuscript.

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Other contributions: This work was performed at the University of Michigan, Ann Arbor, MI.

Additional information: The e-Appendix can be found in the "Supplemental Materials" area of the online article.

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