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## Strong consistency of nonparametric Bayes density estimation on compact metric spaces with applications to specific manifolds

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## Abstract

This article considers a broad class of kernel mixture density models on compact metric spaces and manifolds. Following a Bayesian approach with a nonparametric prior on the location mixing distribution, sufficient conditions are obtained on the kernel, prior and the underlying space for strong posterior consistency at any continuous density. The prior is also allowed to depend on the sample size *n* and sufficient conditions are obtained for weak and strong consistency. These conditions are verified on compact Euclidean spaces using multivariate Gaussian kernels, on the hypersphere using a von Mises-Fisher kernel and on the planar shape space using complex Watson kernels.

## Keywords

Nonparametric Bayes; Density Estimation; Posterior consistency; Sample dependent prior; Riemannian manifold; Hypersphere; Shape space

## **1** Introduction

Density estimation on compact metric spaces, such as manifolds, is a fundamental problem in nonparametric inference on non-Euclidean spaces. Some applications include directional and axial data analysis, spatial modeling, shape analysis and dimensionality reduction problems in which the data lie on an unknown lower dimensional space. However, the literature on statistical theory and methods of density estimation in non-Euclidean spaces is still under-developed. Our focus is on Bayesian nonparametric approaches.

For nonparametric Bayes density estimation on the real line  $\Re$ , there is a rich literature, with Dirichlet process mixtures of Gaussian kernels providing a commonly-used approach (Escobar and West (1995)[6]) that leads to dense support (Lo (1984)[14]) and weak and strong posterior consistency (Ghosal *et al.* (1999)[8]). From the celebrated theorem of Schwartz (1965)[16], weak posterior consistency results when the true density  $f_0$  is in the Kullback-Leibler (KL) support of the prior, meaning that all KL neighborhoods around  $f_0$ are assigned positive probability. In general, it is quite difficult to show KL support for new priors for a density, though Wu and Ghosal (2008)[20] provide useful conditions for a class of kernel mixture priors, with Bhattacharya and Dunson (2010a)[2] extending these conditions to general compact metric spaces. It is widely accepted that weak consistency is an insufficient property when the focus is on density estimation. For example, if  $f_0$  is a density with respect to Lebesgue measure, weak consistency does not even ensure that the posterior assigns positive probability to the set of densities with respect to Lebesgue measure. Hence, it is important to provide stronger results.

Until very recently, essentially all the literature on theory of nonparametric Bayes density estimation focused on one-dimensional Euclidean spaces. An important development in multivariate Euclidean spaces is the article of Wu and Ghosal (2010)[21] who provide sufficient conditions for strong consistency in nonparametric Bayes density estimation from Dirichlet process mixtures of multivariate Gaussian kernels. However severe tail restrictions are imposed on the kernel covariance, which become overly restrictive when the data are very high dimensional. Also the theory developed in their paper is specialized and cannot be easily generalized to arbitrary kernel mixtures on more general spaces.

We are particularly interested in density estimation in the special case in which the compact metric space M corresponds to a Riemannian manifold. In order to extend kernel mixture models used in Euclidean spaces to manifolds *M*, the kernel needs to be carefully chosen. One approach is to introduce an invertible coordinate map between a subset of M and a Euclidean space (Hirsch (1976)[10]). Under such an approach, the density prior on M can be induced through a kernel mixture model in a Euclidean space. However, several major problems arise in using such an approach. Firstly, it is not possible to cover the entire manifold with a single smooth coordinate chart except for very simple manifolds, so unless the data are very concentrated one may obtain poor performance. Different local charts can be patched together to form an atlas, but this may introduce artifactual discontinuities in the resulting density. Because the coordinate map is not isometric, the geometry of the manifold can be heavily distorted. As good choices of coordinate frames necessarily depend on the observations, additional uncertainty is automatically induced. Due to these and other shortcomings of coordinate based methods, we focus on modeling approaches that are coordinate free in the sense that we build density models with respect to the invariant volume form on the manifold.

In Bhattacharya and Dunson (2010a)[2], a density model is presented on a general compact metric space with respect to any fixed base measure using a random mixture of probability kernels. Under mild conditions on the kernel and the mixing prior, it is shown that the prior probability of any uniform neighborhood of any continuous density  $f_0$  is positive and if  $f_0$  is positive everywhere, it lies in the KL support of the prior. This inturn implies posterior consistency in the weak sense. Density estimation on the planar shape space is presented as a special case. However strong posterior consistency is not addressed. The everywhere positivity restriction on the true density cannot be easily relaxed. Also besides the location mixing distribution, the only other parameter in the model is a scalar band-width. This restricts the flexibility when the sample size is small.

Focusing on kernel mixture priors for densities on a compact metric space M, in this article, we provide sufficient conditions on the kernel, prior and the underlying space to ensure strong consistency. Theorem 2 and Corollary 1 provide sufficient conditions to ensure that all total variation neighborhoods around  $f_0$  will be assigned probability converging to one as the sample size increases. The theoretical development relies on the method of sieves and exponentially consistent tests discussed in Barron (1989)[1]. However, applying this framework outside multivariate spaces is not standard and requires careful use of differential geometry. Through Theorem 1, we prove weak consistency for a bigger class of kernels than Bhattacharya and Dunson (2010a)[2]. The only requirement on the true density is that it is continuous everywhere. To illustrate the theory, we focus on density estimation on the unit hypersphere using von Mises-Fisher kernels and on the planar shape space using complex Watson kernels. In both these cases, it is shown that the kernels satisfy the sufficient conditions. The results also apply to Gaussian mixture densities on  $\Re^d$  whenever the true

density has compact support. In that case, a truncated and transformed Wishart prior on the covariance inverse, the transformation depending on the data dimension is shown to suffice as in Wu and Ghosal (2010)[21]. Appropriate kernel choices are presented on other manifolds such as axial spaces, Stiefel manifolds and Grasmannians which arise as generalisations of the two manifolds.

When the manifold is high-dimensional, priors satisfying conditions for strong consistency tend to put too little probability near bandwidths close to 0, which is undesirable for applications. A gamma prior on the inverse-bandwidth, for example, cannot be shown to satisfy the conditions. Hence, we extend the consistency results to cover cases with priors depending on the sample size *n*. Theorem 3 extends the Schwartz theorem to prove weak consistency, while Theorem 4 proves strong consistency using such priors. A gamma prior with scale decreasing with *n* at an appropriate rate satisfies the conditions for both weak and strong posterior consistency at an exponential rate. When using multivariate Gaussian mixtures, a truncated Wishart prior with hyper-parameters depending on *n*, is shown to work.

To mantain a free flow while reading, we put all the proofs together at the end in a section called Appendix.

## 2 Consistency theorems on compact metric spaces

#### 2.1 Weak posterior consistency

Let  $(M, \rho)$  be a compact metric space,  $\rho$  being the distance metric, and let X be a random variable on M (from some measurable space  $(\Omega, \mathcal{B}, Q)$ ). We assume that the distribution of X has a density with respect to some fixed finite base measure  $\lambda$  on M. The natural choice for such a  $\lambda$  when M is a Riemannian manifold is the invariant volume form. We are interested in modelling this unknown density via a flexible model. Let  $K(m; \mu, \kappa)$  be a probability kernel on M with location  $\mu \in M$  and other parameters  $\kappa \in N$ , N being a Polish space, that is, it is homeomorphic to a complete seperable metric space. In the special case, we choose  $N = (0, \infty)$  and then K may be called a location-scale kernel.

Given the parameters  $(\mu, \kappa)$ , *K* satisfies  $\int_M K(m; \mu, \kappa) \lambda(dm) = 1$ . Then a location mixture density model for *X* is defined as

$$f(m; P, \mathscr{K}) = \int_{\mathcal{U}} K(m; \mu, \mathscr{K}) P(d\mu)$$
(1)

with parameters *P* in the space  $\mathcal{M}(M)$  of all probability distributions on M and  $\mathcal{K} \in N$ . We call *P* the location mixing distribution. When  $N = (0, \infty)$ , we view  $\mathcal{K}^{-1}$  as the band-width of the kernel and hence call  $\mathcal{K}$  the *precision* or inverse band-width parameter. More generally  $\mathcal{K}$  comprises of many other parameters in different spaces determining the kernel shape, modality etc, and the precision is a particular function of  $\mathcal{K}$ . The upcoming consistecy theorems and examples will illustrate that function. Kernel mixture models are used routinely in Bayesian density estimation in Euclidean spaces, with Lennox *et al.*(2009)[13] applying such an approach to bivariate angular data and Bhattacharya and Dunson (2010a) [2], (2010b)[3] considering kernel mixtures on general metric spaces.

A prior  $\Pi_1$  on  $(P, \mathcal{K})$  induces a prior  $\Pi$  on the space of densities  $\mathcal{P}(M)$  on M through the model (1). Given a random realization  $X_1, \ldots, X_n$  of X, we can compute the posterior of f. The Schwartz theorem provides a useful tool in proving that the posterior assigns probability converging to one in arbitrarily small neighborhoods of the true density  $f_0$  as the sample size  $n \to \infty$ . Let  $F_0$  denote the probability distribution corresponding to  $f_0$ , let KL( $f_0; f) = \int_M$ 

 $f_0(m) \log\{f_0(m)/f(m)\}\lambda(dm)$  denote the KL divergence of another density *f* from  $f_0$ , and let  $K_{\varepsilon}(f_0)$  denote the KL neighborhood  $\{f \in \mathcal{P}(M): \operatorname{KL}(f_0; f) < \varepsilon\}$ .  $f_0$  is said to be in the KL support of  $\Pi$  or  $\Pi$  is said to satisfy the KL condition at  $f_0$  if  $\Pi\{K_{\varepsilon}(f_0)\} > 0$  for all  $\varepsilon > 0$ .

**Proposition 1 (Schwartz Theorem)**—If (1)  $f_0$  is in the KL support of  $\Pi$ , and (2)  $U \subset \mathcal{D}$ (*M*) is such that there exists a uniformly exponentially consistent sequence of test functions for testing  $H_0$ :  $f = f_0$  versus  $H_1$ :  $f \in U^c$ , then  $\Pi(U|X_1, ..., X_n) \to 1$  as  $n \to \infty$  a.s.  $F_0^{\infty}$ .

The posterior probability of  $U^c$  can be expressed as

$$\Pi(U^{c}|X_{1},\ldots,X_{n}) = \frac{\int_{U^{c}} \prod_{i=1}^{n} \frac{f(X_{i})}{f_{0}(X_{i})} \Pi(df)}{\int \prod_{i=1}^{n} \frac{f(X_{i})}{f_{0}(X_{i})} \Pi(df)}$$
(2)

Condition (1), known as the KL condition, ensures that for any  $\beta > 0$ ,

$$\liminf_{n \to \infty} \exp(n\beta) \int \prod_{i=1}^{n} \frac{f(X_i)}{f_0(X_i)} \Pi(df) = \infty \text{ a.s.}$$
(3)

while condition (2) implies that

$$\lim_{n\to\infty} \exp(n\beta_0) \int_{U^c} \prod_{i=1}^n \frac{f(X_i)}{f_0(X_i)} \Pi(df) = 0 \text{ a.s.}$$

for some  $\beta_0 > 0$  (depending on U) and therefore

$$\lim_{n\to\infty}\exp(n\beta_0/2)\Pi(U^c|X_1,\ldots,X_n)=0 \text{ a.s.}$$

Hence Proposition 1 provides conditions for posterior consistency at an exponential rate. Theorem 1 derives sufficient conditions on the kernel and the prior so that  $f_0$  is in the uniform support and hence KL support of  $\Pi$ . They are

A1 The kernel *K* is continuous in its arguements.

For any continuous function  $f: M \to \Re$  (written as  $f \in C(M)$ ), for any  $\varepsilon > 0$ , there exists a compact subset  $N_{\varepsilon}$  of N with non-empty interior, such that

- A2  $\sup_{m \in M, \ \mathcal{K} \in N_{e}} |f(m) \int_{M} K(m; \mu, \mathcal{K}) f(\mu) \lambda(d\mu)| < \varepsilon.$
- A3 For any  $\varepsilon > 0$ , the set  $\{F_0\} \times N_{\varepsilon}^o$  intersects with the (weak) support of  $\Pi_1$ . Here  $A^o$  denotes the interior of a set A.
- A4  $f_0$  is a continuous density.

**Theorem 1—**Under assumptions A1–A4, for any  $\varepsilon > 0$ ,

$$\Pi\left\{f\in\mathscr{D}(M):\sup_{m\in M}|f(m)-f_0(m)|<\varepsilon\right\}>0,$$

As a corollary, we obtain the KL property for the location-scale kernel, derived in Bhattacharya and Dunson (2010a)[2]. However unlike in there, we need not assume  $f_0$  to be positive everywhere.

When *U* is a weakly open neighborhood of  $f_0$ , condition (2) in Proposition 1 is always satisfied. Hence under assumptions A1–A4, weak posterior consistency at an exponential rate follows. Assumptions A1 and A2 impose some mild conditions on the kernel choice, which are easily satisfied by several parametric families. In particular, A2 implies that as a probability distribution on *M*,  $K(; \mu, \kappa)$  can be made arbitrarilly close in weak sense to the degenerate point mass at  $\mu$ , uniformly in  $\mu$ , for appropriate  $\kappa$  choice, thereby justifying the name 'location' for  $\mu$ . When the compact neighborhood  $N_e$  can be represented as the inverse image under some function  $\psi(\psi: N \rightarrow \Re^+)$  of some neighborhood around infinity, then  $\psi(\kappa)$  can be viewed as the precision parameter. We will provide examples of kernels on some non-Euclidean manifolds arising in shape and directional data analysis which satisfy A1 and A2. A common choice for  $\Pi_1$  satisfying A3 can be a full support product prior such as a Dirichlet process DP( $w_0P_0$ ) prior on *P* satisfying supp( $P_0$ ) = *M* and an independent everywhere positive density prior on  $\kappa$ .

#### 2.2 Strong consistency

When U is a total variation neighborhood of  $f_0$ , LeCam (1973)[12] and Barron (1989)[1] show that condition (2) of Proposition 1 will not be satisfied in most cases. In Barron (1989) [1] (also see Ghosal *et al.*(1999)[8]), a sieve method is considered to obtain sufficient conditions for the numerator in (2) to decay at an exponential rate and hence get strong posterior consistency at an exponential rate. This is stated in Proposition 2. In its statement, for  $\mathcal{F} \subseteq \mathcal{D}(M)$  and  $\varepsilon > 0$ , the  $L_1$ -metric entropy  $N(\varepsilon, \mathcal{F})$  is defined as the logarithm of the minimum number of  $\varepsilon$ -sized (or smaller)  $L_1$  subsets needed to cover  $\mathcal{F}$ .

**Proposition 2**—If there exists a  $\mathcal{D}_n \subseteq \mathcal{D}(M)$  such that (1) for *n* sufficiently large,  $\Pi(\mathcal{D}_n^c) < exp(-n\beta)$  for some  $\beta > 0$ , and (2)  $N(\varepsilon, \mathcal{D}_n)/n \to 0$  as  $n \to \infty$  for any  $\varepsilon > 0$ , then for any total variation neighborhood *U* of  $f_0$ , there exists a  $\beta_0 > 0$  (depending on *U*) such that

 $\limsup_{n\to\infty} \exp(n\beta_0) \int_{U^c} \prod_{1=1}^n \frac{f(X_i)}{f_0(X_i)} \Pi(df) = 0 \text{ a.s. } F_0^{\infty}. \text{ Hence if } f_0 \text{ is in the KL support of } \Pi, \text{ the posterior probability of any total variation neighborhood of } f_0 \text{ converges to 1 almost surely.}$ 

Theorem 2, which is the main theorem of this paper, describes a  $\mathcal{D}_n$  which satisfies condition (2). We assume that there exists a continuous function  $\varphi: N \to [0, \infty)$  for which the following assumptions hold.

A5 There exists positive constants  $\kappa_1$ ,  $a_1$ ,  $A_1$  such that for all  $\kappa$   $\kappa_1$ ,  $\mu$ ,  $\nu \in M$ ,

$$\sup_{m \in M, \mathscr{K} \in \varphi^{-1}[0,\kappa]} |K(m;\mu,\mathscr{K}) - K(m;\nu,\mathscr{K})| \le A_1 \kappa^{a_1} \rho(\mu,\nu).$$

A6 There exists positive constants  $a_2$ ,  $A_2$  such that for all  $\kappa_1$ ,  $\kappa_2 \in \varphi^{-1}$  [0,  $\kappa$ ],  $\kappa_1$ ,  $\kappa_1$ ,

$$\sup_{m,\mu\in M} |K(m;\mu,\mathcal{K}_1) - K(m;\mu,\mathcal{K}_2)| \le A_2 \kappa^{a_2} \rho_2(\mathcal{K}_1,\mathcal{K}_2),$$

 $\rho_2$  metrizing the topology of N.

- A7 For any  $\kappa$   $\kappa_1$ , the subset  $\varphi^{-1}[0, \kappa]$  is compact and given  $\varepsilon > 0$ , the minimum number of  $\varepsilon$  (or smaller) radius balls covering it (known as the *e*-covering number) can be bounded by  $(\kappa \varepsilon^{-1})^{b_2}$  for appropriate positive constant  $b_2$  (independent of  $\kappa$  and  $\varepsilon$ ).
- **A8** There exists  $a_3$ ,  $A_3 > 0$  such that the *e*-covering number of *M* is bounded by  $A_3e^{-a_3}$  for any e > 0.

For two positive sequences  $\{a_n\}$  and  $\{b_n\}$ ,  $\{a_n\}$  is said to be 'little-o' of  $\{b_n\}$ , written as  $a_n = o(b_n)$ , if the sequence  $\{a_n/b_n\}$  converges to 0 as  $n \to \infty$ .

**Theorem 2**—For a positive sequence  $\{\kappa_n\}$  diverging to  $\infty$ , define

$$\mathcal{D}_n = \{ f(P, \mathcal{K}) : P \in \mathcal{M}(M), \ \mathcal{K} \in \varphi^{-1}[0, \kappa_n] \}.$$

Under assumptions A5–A8, given any  $\varepsilon > 0$ , for *n* sufficiently large,  $N(\varepsilon, \mathcal{D}_n) \le C(\varepsilon)\kappa_n^{a_1a_3}$  for some  $C(\varepsilon) > 0$ . Hence  $N(\varepsilon, \mathcal{D}_n)$  is o(n) whenever  $\kappa_n$  is  $o(n^{(a_1a_3)^{-1}})$ .

As a corollary, we derive conditions on the prior  $\Pi_1$  on  $(P, \kappa)$  under which strong posterior consistency at an exponential rate follows.

Corollary 1—Under the hypothesis of Theorems 1 and 2 and

**A9**  $\Pi_1(\mathcal{M}(M) \times \varphi^{-1}(n^a, \infty)) < \exp(-n\beta)$  for some  $a < (a_1a_3)^{-1}$  and  $\beta > 0$ ,

the posterior probability of any total variation neighborhood of  $f_0$  converges to 1 a.s.  $F_0^{\infty}$ .

Theorem 1 ensures that  $f_0$  is in the KL support of  $\Pi$ . Theorem 2 and assumption **A9** ensure that  $\mathcal{D}_n$  satisfies conditions (1) and (2) of Proposition 2. Hence from the Proposition, the proof follows.

When we use a location-scale kernel, that is, when  $N = (0, \infty)$ , choose a prior  $\Pi_1 = \Pi_{11} \otimes \pi_1$ having full support, set  $\varphi$  to be the identity map, then a choice for  $\pi_1$  for which Assumption **A9** is satisfied is a Weibull density  $W eib(\mathcal{K}; \alpha, \beta) \propto \kappa^{\alpha-1} \exp(-\beta \kappa^{\alpha})$  whenever the shape parameter  $\alpha$  exceeds  $a_1a_3$  ( $a_1, a_3$  as in Assumptions **A5** and **A8**).

**Remark 1:** A gamma prior on  $\mathcal{K}$  satisfies the requirements for weak consistency but not **A9** (unless  $a_1a_3 < 1$ ). However that does not prove that it is not eligible for strong consistency because Corollary 1 provides only sufficient conditions. In Section 2.3, we prove it to be eligible as long as the hyperparameters are allowed to depend on sample size *n* in a suitable way.

When the underlying space is non-compact such as  $\Re^d$ , Corollary 1 applies to any true density  $f_0$  supported on a compact set, say M. Then the kernel can be chosen to have non-compact support, such as Gaussian, but to apply Theorem 2, we need to restrict the prior on the location mixing distribution to have support in  $\mathcal{M}(M)$ . We can weaken assumptions **A5** and **A6** to

- A5'  $\sup_{\kappa \in \varphi^{-1}[0, \kappa]} ||K(\mu, \kappa) K(\nu, \kappa)|| = A_1 \kappa^{a_1} \rho(\mu, \nu)$  and
- A6'  $\mu \in M \| K(\mu, \kappa_1) K(\mu, \kappa_2) \| \quad A_2 \kappa^{22} \rho_2(\kappa_1, \kappa_2) \forall \kappa_1, \kappa_2 \in \varphi^{-1} [0, \kappa] \text{ respectively.}$ Here  $\| f - g \|$  denotes the  $L_1$  distance between densities f and g. The proof of Theorem 2 can be easily modified to show consistency under the new assumptions and is left to the reader.

The multivariate Gaussian kernel can be represented as

$$K(m,\mu,\mathcal{K}) = (2\pi)^{-d/2} \det(\mathcal{K})^{1/2} \exp(-1/2(m-\mu)'\mathcal{K}(m-\mu))m, \mu \in \mathbb{R}^d, \mathcal{K} \in M^+(d),$$

 $M^{+}(d)$  being the space of all  $d \times d$  positive matrices. Hence  $\kappa^{-1}$  is the kernel covariance. It satisfies **A5**' and **A6**' as shown in Proposition 3. Here by  $\lambda_{1}(\kappa)$ , ...,  $\lambda_{d}(\kappa)$ , we denote the eigen-values of  $\kappa$  in increasing order.

**Proposition 3**—The multivariate Gaussian kernel satisfies  $\mathbf{A5}'$  with  $\varphi$  being the largest eigen-value function,  $\varphi(\mathcal{K}) = \lambda_d(\mathcal{K})$  and  $a_1 = 1/2$ . It also satisfies  $\mathbf{A6}'$  once we restrict *N* to be the space of all positive matrices with the least eigen-value being bounded below by some pre-specified positive constant, say,  $\lambda_1$ , i.e.,  $N = \{\mathcal{K} \in M^+(d) : \lambda_1(\mathcal{K}) \mid \lambda_1\}$ . The space  $M^+(d)$  (and hence *N*) satisfies **A7** while any compact subset M of  $\mathbb{R}^d$  satisfies **A8** from Theorem 2 with  $a_3 = d$ . Hence if

$$\Pi_1(\mathscr{M}(M) \times \{\mathscr{K} \in N: \lambda_d(\mathscr{K}) > n^a\}) < \exp(-n\beta)$$

for some a < 2/d and  $\beta > 0$ , and if  $f_0$  is in the KL support of  $\Pi$ , strong posterior consistency follows.

Theorem 4 in Ghosal *et al.*(1999)[8] provides sufficient conditions on  $f_0$  and  $\Pi_1$  for the KL condition to be satisfied while using a Gaussian mixture model in the univariate setting to model a compactly supported density. It assumes  $\Pi_1 = \Pi_{11} \otimes \pi_1$  with  $F_0 \in \text{supp}(\Pi_{11})$  and  $\infty \in \text{supp}(\pi_1)$ . The theorem can be extended to multivariate setting with the condition on  $\Pi_1$  relaxed to, for any  $\kappa > 0$ , there exists a  $\kappa \in M^+$  (*d*) with  $\lambda_1(\kappa) \quad \kappa$  such that  $(F_0, \kappa) \in \text{supp}(\Pi_1)$ . Therefore  $\lambda_1(\kappa)$  can be viewed as the kernel precision in this case. A full support product  $\Pi_1$  on  $\mathcal{M}(M) \times N$  will satisfy these requirements. Using a product prior, a choice for  $\pi_1$  for which strong consistency also follows can be the so called *truncated transformed Wishart* defined as follows. Set  $\kappa = \Lambda^a$  for any  $a \in (0, 2/d)$  with  $\Lambda$  following a Wishart distribution restricted to N. Then  $\kappa$  is said to follow a truncated transformed Wishart with transformation power a.

**<u>Remark 2</u>**: The truncation restriction on the space *N* is not undesirable, because for more precise fit, we are interested in low bandwidths and the least eigen-value of  $\mathcal{K}$  can be viewed as the inverse of the band-width. However, lower the transformation power, lower is the prior probability for high precisions which is undesirable when sample sizes are not high.

In Wu and Ghosal (2010)[21], strong consistency is proved in the special case of Dirichlet process Gaussian mixtures used to model density  $f_0$  having support as  $\Re^d$ . It requires *a* to be less than 1/d resulting in even smaller precision. In the next section, we prove that no transformation is required (*a* = 1) as long as the hyper-parameters are allowed to depend on the sample size appropriately.

#### 2.3 Consistency with sample size-dependent priors

When the dimension of the manifold is large, as is the case in shape analysis with a large number of landmarks, the constraints on the shape parameter in the proposed Weibull prior on the inverse bandwidth become overly-restrictive. In particular, for strong posterior consistency, the shape parameter needs to be very large in high-dimensional cases, implying a prior on the bandwidth that places very small probability in neighborhoods close to zero,

which is undesirable in many applications. By instead allowing the prior to depend on sample size *n*, we can potentially obtain priors that may have better small sample operating characteristics, while still leading to strong consistency. However, for *n*-dependent priors, the KL condition is no longer sufficient to ensure that (3) holds and hence the Schwartz theorem breaks down. In this section, we will modify the conditions and derive weak and strong consistency results for *n*-dependent priors.

As recommended in earlier sections, we let *P* and  $\mathcal{K}$  be independent under  $\Pi_1$ . Then, assuming  $P \sim \Pi_{11}$  is a constant prior, we focus on the case in which  $\mathcal{K}$  has a sample sizedependent prior density  $\pi_n$  with respect to some base measure  $\lambda_1$  on N,  $\mathcal{K} \sim \pi_n(\mathcal{K})\lambda_1(d\mathcal{K})$ . We pick  $\lambda_1$  to have full support. Depending on the context,  $\pi_n$  will refer to both the density and distribution of  $\mathcal{K}$ . Denote the resulting sequence of induced priors on  $\mathcal{D}(M)$  as  $\Pi_n$ . Theorem 3 proves weak posterior consistency under the following assumptions on the prior.

- A10 The prior  $\Pi_{11}$  contains  $F_0$  in its support.
- A11 For any  $\varepsilon > 0$ , for all  $\mathcal{K} \in N_{\varepsilon}$ ,

 $\liminf_{n \in \mathcal{M}} \operatorname{exp}(n\varepsilon) \pi_n(\mathscr{K}) = \infty.$ 

Here  $N_e$  is as defined in Assumption A2.

**Theorem 3**—Under assumptions A1 and A2 on the kernel, A4 on the true density  $f_0$ , and, A10 and A11 on the prior, the posterior probability of any weak neighborhood of  $f_0$  converges to one a.s.  $F_0^{\infty}$ .

The proof is immediate from the following two lemmas.

Lemma 1—Under assumptions A1–A2, A4 and A10–A11, for any e > 0,

$$\liminf_{n \to \infty} \exp(n\varepsilon) \int_{-1}^{n} \frac{f(X_i)}{f_0(X_i)} \Pi_n(df) = \infty$$
(4)

a.s.  $F_0^{\infty}$ .

**Lemma 2**—If there exists a uniformly exponentially consistent sequence of test functions for testing  $H_0$ :  $f = f_0$  versus  $H_1$ :  $f \in U^c$ , and  $\Pi_n(U^c) > 0$  for all n > C with C a sufficiently arge constant, then for some  $\beta_0 > 0$ ,

$$\lim_{n \to \infty} \exp(n\beta_0) \int_{U^c} \prod_{1}^n \frac{f(X_i)}{f_0(X_i)} \Pi_n(df) = 0$$

a.s.  $F_0^{\infty}$ .

The proof of Lemma 2 is related to that of Lemma 4.4.2. from Ghosh and Ramamoorthi (2003)[9] which is stated for a constant prior  $\Pi$  but with the set  $U^c$  depending on *n*, they call this  $V_n$ . There it is assumed that  $\lim \inf_{n\to\infty} \Pi(V_n) > 0$  but that is not necessary as long as  $\Pi(V_n) > 0$  for all large *n*. Lemma 1 is proved in the Appendix.

With a location-scale kernel, *N* being  $(0, \infty)$ , a gamma prior  $\pi_n(\mathcal{K}) \propto \exp(-\beta_n \mathcal{K}) \kappa^{\alpha-1}$ , a,  $\beta_n > 0$ , denoted by Gam $(a, \beta_n)$ , satisfies Assumption A11 on entire of *N*, as long as  $\beta_n$  is o(n).

With a multivariate Gaussian kernel on  $\Re^d$  with dispersion  $\kappa^{-1}$ , a Wishart prior on  $\kappa$ ,

$$\pi_n(\mathscr{K};\beta_n,q) = 2^{-dq/2} \Gamma_d(q/2) \beta_n^{dq/2} \exp(-\beta_n/2 \operatorname{Tr}(\mathscr{K})) \det(\mathscr{K})^{(q-d-1)/2}, q > d-1, \beta_n > 0,$$

denoted as Wish $(\beta_n^{-1}I_d, q)$  satisfies **A11** on entire  $M^+(d)$ , as long as  $\beta_n$  is o(n). Here  $\Gamma_d(.)$  denotes the *multivariate gamma function* defined as

$$\Gamma_d(q/2) = \int_{\mathcal{M}^+(d)} \exp(-\operatorname{Tr}(\mathscr{K})) \det(\mathscr{K})^{(q-d-1)/2} d\mathscr{K}$$

For strong consistency, we impose the following additional condition on  $\pi_n$ . Let  $a_1$  and  $a_3$  be as in Assumptions A5 (or A5') and A8 respectively.

**A12** For some  $\beta_0 > 0$  and  $a < (a_1 a_3)^{-1}$ ,

 $\lim_{n\to\infty}\exp(n\beta_0)\pi_n\{\varphi^{-1}(n^a,\infty)\}=0.$ 

This assumption is in place of A9 used for constant priors.

**Theorem 4**—Under Assumptions A1–A2 and A4–A8 and A10–A12, the posterior probability of any total variation neighborhood of  $f_0$  converges to 1 a.s  $F_0^{\infty}$ .

The proof is very similar to that of Corollary 1. This is because under assumptions A1–A2, A4 and A10–A11, the conclusion (4) of Lemma 1 holds. The other assumptions are to show that the  $L_1$ -metric entropy of  $\mathcal{D}_n$  is o(n) while  $\prod_n(\mathcal{D}_n^c)$  is exponentially small,  $\mathcal{D}_n$  being defined in Theorem 2. Under these assumptions, the proof of Proposition 2 goes through to prove strong consistency with sample size dependent priors. This is also mentioned in §5 of Ghosal *et al.*(1999)[8]. They require  $\liminf_{n\to\infty} \prod_n (K_e(f_0)) > 0$  in place of of the assumption  $\prod(K_e(f_0)) > 0$  for constant priors but this is only to ensure that (4) holds.

Again as in §2.2, we can weaken assumptions A5 and A6 to A5' and A6' respectively.

For a location-scale kernel, a Gam $(a, \beta_n)$  prior on precision  $\mathcal{K}$  satisfies **A12** when  $n^{1-a}$  is  $o(\beta_n)$  for some  $a \in (0, (a_1a_3)^{-1})$ . Hence, for example, we have weak and strong posterior consistency with  $\beta_n = b_1 n/\{\log(n)\}^{b_2}$  for any  $b_1, b_2 > 0$ .

For a multivariate Gaussian kernel, to satisfy assumption A6', we need to truncate the space N to {  $\kappa \in M^+(d)$ :  $\lambda_1(\kappa) = \lambda_1$ }, as proved in Proposition 3. Then we may set a truncated Wishart prior on  $\kappa$ , defined as

$$\pi_n(\mathscr{K}) = \frac{\exp(-\beta_n/2\mathrm{Tr}(\mathscr{K}))\det(\mathscr{K})^{(q-d-1)/2}}{\int_{A\in\mathbb{N}}\exp(-\beta_n/2\mathrm{Tr}(A))\det(A)^{(q-d-1)/2}dA}, \ \mathscr{K}\in\mathbb{N}.$$
(5)

Then for Assumption A12 to be satisfied, we require  $n^{1-a}$  to be  $o(\beta_n)$  for some  $a \in (0, (a_1a_3)^{-1})$ . This is shown in Proposition 4. Hence we have weak and strong posterior consistency once we set  $\beta_n = b_1 n/\{\log(n)\}^{b_2}$  for any  $b_1, b_2 > 0$ . Unlike in §2.2, we impose no transformation constraints, which is very helpful especially when sample sizes are not that high while the data dimensions are huge.

**Proposition 4**—For a positive sequence  $\{\beta_n\}$  diverging to infinity, Assumption A12 is satisfied for the Truncated-Wishart density sequence  $\pi_n$  in (5) if there exists an  $a \in (0, (a_1a_3)^{-1})$  for which  $\beta_n$  satisfies  $n^{1-a}/\beta_n \to 0$  as  $n \to \infty$ .

In the subsequent sections, we present kernel choices for density estimation on some specific non-Euclidean manifolds that arise in several applications. We illustrate how to apply Theorems 1, 2, 3 and 4 and obtain weak and strong posterior consistency.

#### 3 Application to unit hypersphere

Let *M* be the unit sphere  $S^d$  embedded in  $\Re^{d+1}$ . It is a compact Riemannian manifold of dimension *d* and a compact metric space under the chord distance  $\rho(u, v) = ||u-v||_2$ ,  $||.||_2$  denoting the  $L^2$ -norm. Spherical data on  $S^2$  arise in the context of directional data analysis. Most of the shape spaces are quotients of high dimensional spheres. Hence it is important to develop consistent inference procedures on this space, and very few results exist in the context of Bayesian nonparametrics.

To define a probability density model as in (1) with respect to the volume form V, we need a suitable kernel which satisfies the assumptions in Section 2. One of the most commonly used probability densities on this space is the von Mises-Fisher (vMF) density which is given by

$$vMF(m;\mu,\mathscr{K}) = c^{-1}(\mathscr{K})exp(\mathscr{K}m^{T}\mu), \ m,\mu \in S^{d}, \ \mathscr{K} \in [0,\infty),$$
(6)

with *c* being the normalizing constant which can be derived to be

$$\frac{2\pi^{d/2}}{\Gamma(\frac{d}{2})} \int_{-1}^{1} \exp(\mathscr{K}t) (1-t^2)^{d/2-1} dt.$$
(7)

The vMF density on  $S^1$  was first derived in Von Mises (1918)[18] and the density in case of  $S^2$  was given by Fisher (1953)[7]. Watson and Williams (1953)[19] generalized this distribution to  $S^d$  and examined many of its properties. It can be shown that the parameter  $\mu$  is the extrinsic mean (as defined in Bhattacharya and Patrangenaru (2003)[4]), and hence can be interpreted as the distribution location. The parameter  $\kappa$  is a measure of concentration, with  $\kappa = 0$  corresponding to the uniform distribution having constant density equal to  $1/\int_S dV(dm)$ . As  $\kappa$  diverges to  $\infty$ , the vMF distribution converges to a point mass at  $\mu$  in an  $L^1$  sense uniformly. This is proved in Theorem 5.

Theorem 5—The vMF kernel satisfies Assumptions A1 and A2.

Hence from Theorem 1, weak posterior consistency follows using the location mixture density model (1) with a Dirichlet Process prior on P and an independent gamma prior on  $\mathcal{K}$ . In the d=2 special case, Lennox *et al.*(2009)[13] proposed a closely related model but did not consider theoretical properties. Theorem 6 verifies the assumptions for strong consistency.

**Theorem 6**—With  $\varphi(\mathcal{K}) = \mathcal{K}$ , the vMF kernel on  $S^d$  satisfies assumption A5 with  $a_1 = d/2 + 1$  and A6 with  $a_2 = d/2$ . The compact metric-space ( $S^d$ ,  $\rho$ ) satisfies assumption A8 with  $a_3 = d$ .

As a result a Weibull prior on  $\mathcal{K}$  with shape parameter exceeding  $(d + d^2/2)^{-1}$  satisfies the condition of Corollary 1 and strong posterior consistency follows. The proofs of Theorems 5 and 6 use the following lemma which establishes certain properties of the normalizing constant.

**Lemma 3**—Define  $\tilde{c}(\mathcal{K}) = \exp(-\mathcal{K})c(\mathcal{K}), \mathcal{K} = 0$ . Then  $\tilde{c}$  is decreasing and for  $\mathcal{K} = 1$ ,

$$\tilde{c}(\mathcal{K}) \geq C\mathcal{K}^{-d/2}$$

for some appropriate positive constant C.

When *d* is large, as is often the case for spherical data, a more appropriate prior on  $\kappa$  for which weak and strong consistencies hold can be  $\text{Gam}(a, \beta_n)$  as mentioned at the end of §2.3.

It is easy to check that the vMF density is the conditional distribution, given ||X|| = 1, of a Gaussian random vector X on  $\Re^{d+1}$  with mean  $\mu$  and dispersion matrix  $\kappa^{-1}I_{d+1}$ . A more general family of distribution on  $S^d$  may be obtained as the conditional distribution, given ||X|| = 1, of a Normal X on  $\Re^{d+1}$  with mean  $\mu$  and dispersion matrix  $\kappa^{-1}$ ,  $\kappa$  in the space  $M^+$  (d+1) of  $(d+1) \times (d+1)$  positive matices. Then we obtain the *Fisher-Bingham* family of kernels. It can be interesting to show that the resulting kernel mixture satisfies the assumptions of Theorems 1, 2, 3 and 4 and obtain posterior consistency. We postpone that to later works.

A generalization of the sphere is the Stiefel manifold  $V_{d+1,k}$ , the space all k dimensional orthonormal frames in  $\Re^{d+1}$ . One can easily extend the vMF kernel to the so called Fisher kernel on this manifold and carry out density estimation. Again, proving that consistency holds is postponed for future works.

Another important manifold arising in axial data analysis is  $\mathbb{RP}^d$ , the space of all rays in  $\mathbb{R}^{d+1}$ . This manifold can be obtained as the quotient of  $S^d$  after identifying antipodal points p and -p as identical. In the next section, we illustrate density estimation on its complex analogue, the complex projective space. It is easy and simpler to obtain analogous results on the real version.

## **4 Planar Shape Space**

#### 4.1 Background

Let *M* be the planar shape space  $\sum_{2}^{k}$  which is defined as follows. Consider a set of *k* landmark locations, k > 2, on a 2D image, not all points being the same. We refer to such a set as a *k*-ad. The similarity shape of this *k*-ad is what remains after removing the Euclidean rigid body motions of translation, rotation and scaling. We use the following shape representation first proposed by Kendall (1984)[11]. Denote the *k*-ad by a complex *k*-vector

z in  $c^k$ . To remove the effect of translation from z, let  $z_c = z - \overline{z}$ , with  $\overline{z} = (\sum_{j=1}^k z_j)/k$  being the centroid. The centered k-ad  $z_c$  lies in a k-1 dimensional complex subspace, and hence we can use k-1 complex coordinates. The effect of scaling is then removed by normalizing

the coordinates of  $z_c$  to obtain a point *w* on the complex unit sphere  $C S^{k-2}$  in  $C^{k-1}$ . Since *w* contains the shape information of *z* along with rotation, it is called the preshape of *z*. The similarity shape of *z* is the orbit of *w* under all rotations in 2D which is

$$[w] = \{e^{i\theta}w: \theta \in (-\pi, \pi]\}.$$

This represents a shape as the set of all intersection points of a unique complex line passing through the origin with  $c_{s^{k-2}}$  and the planar shape space  $\sum_{2}^{k}$  is then the set of all such shapes. Hence  $\sum_{2}^{k}$  can be identified with the space of all complex lines passing through the origin in  $c^{k-1}$  which is the complex projective space and is a compact Riemannian manifold of dimension 2k - 4. The  $\sum_{2}^{k}$  can be embedded into the space of all order k - 1 complex Hermitian matrices via the embedding  $J([w]) = ww^*$ , \* denoting the complex conjugate transpose. This embedding induces a distance on  $\sum_{2}^{k}$  called the extrinsic distance which generates the manifold topology and is given by

$$d_{E}([u], [v]) = \left\| J([u]) - J([v]) \right\| = \sqrt{2(1 - |u^{*}v|^{2})} \quad ([u], [v] \in \sum_{1}^{k}).$$

For more details, see Bhattacharya and Dunson (2010a)[2] and the references cited therein.

#### 4.2 Density model

We define a location-mixture density on  $\sum_{2}^{k}$  as in (1) with respect to the Riemannian volume form V and the kernel being a complex Watson density. This complex Watson density was used in Dryden and Mardia (1998)[5] for parametric density modelling and is given by

$$CW(m;\mu,\mathcal{K}) = c^{-1}(\mathcal{K}) \exp\{\mathcal{K}(|z^*v|^2 - 1)\} \quad (m = [z], \mu = [v])$$
(8)

with 
$$c(\mathcal{K}) = \pi^{k-2} \mathcal{K}^{2-k} \left( 1 - \exp(-\mathcal{K}) \sum_{r=0}^{k-3} \frac{\mathcal{K}^r}{r!} \right).$$
 (9)

It is shown in Bhattacharya and Dunson (2010a)[2] that the complex Watson kernel satisfies assumptions A1 and A2 in §2. Using a Dirichlet Process prior on the location mixing distribution and an independent gamma prior on the precision parameter, Theorem 1 implies that the density model (1) has full support in the space of all positive continuous densities on -k

 $\sum_{2}^{k}$  in uniform and KL sense and hence the posterior is weakly consistent.

Theorem 7 verifies that the complex Watson kernel also satisfies the regularity conditions in A5 and A6.

**Theorem 7**—The complex Watson kernel  $CW(m; \mu, \kappa)$  on the compact metric space  $\sum_{2}^{k}$  endowed with the extrinsic distance  $d_E$  satisfies assumption A5 with  $a_1 = k - 1$  and A6 with  $a_2 = 3k - 8$ .

The proof uses Lemma 4 which verifies certain properties of the normalizing constant.

**Lemma 4**—Let  $c(\mathcal{K})$  be the normalizing constant for  $CW(\mu, \mathcal{K})$  as defined in (9). Then *c* is decreasing on  $[0, \infty)$  with

$$\lim_{\kappa \to 0} c(\mathscr{K}) = \frac{\pi^{k-2}}{(k-2)!} \text{ and } \lim_{\mathscr{K} \to \infty} c(\mathscr{K}) = 0.$$

If we define

$$\tilde{c}(\mathcal{K}) = 1 - \exp(-\mathcal{K}) \sum_{r=0}^{k-3} \frac{\mathcal{K}^r}{r!},$$

it follows that  $\tilde{c}$  is increasing with

$$\begin{split} &\lim_{\mathcal{K}\to 0} \tilde{c}(\mathcal{K}){=}0, \ \lim_{\mathcal{K}\to\infty} \tilde{c}(\mathcal{K}){=}1 \ and \\ &\tilde{c}(\mathcal{K}) \geq (k{-}2)!^{-1} \exp(-\mathcal{K})\mathcal{K}^{k{-}2}. \end{split}$$

The proof follows from direct computations.

Theorem 8 verifies that assumption **A8** holds on  $\sum_{2}^{k}$ 

**Theorem 8**—The compact metric space ( $\sum_{2}^{k}$ ,  $d_{E}$ ) satisfies assumption A8 with  $a_{3} = 2k - 3$ .

As a result, Corollary 1 implies that strong posterior consistency holds with  $\Pi_1 = (DP)$  $(\omega_0 P_0) \otimes \pi_1$ , for Weibull  $\pi_1$  with shape parameter exceeding (2k - 3)(k - 1). Alternatively one may use a gamma prior on  $\mathcal{K}$  with inverse-scale increasing with *n* at a suitable rate and we have consistency using Theorems 3 and 4.

The complex Watson kernel is a special case of the *complex Bingham* kernel which has density proportional to  $\exp(z^*Az)$  with respect to the volume form. This kernel has location corresponding to the shape of a eigen-vector corresponding to the largest eigen-value of A. Since it has more parameters, we expect better fit in smaller samples. We will prove that weak and strong posterior consistency holds while using this kernel in a later work.

When the landmarks are obtained from a 3D object, it is more appropriate to carry out an affine shape analysis, that is identify two *k*-configurations as identical if they are related by an affine transformation. One can identify the resulting shape space with the Grassmanian manifold - the space of all 3D subspaces of  $\Re^{k-1}$ , a result of Sparr (1992)[17]. The

Grassmanian is an extension of the real projective space and hence one may consider (real) Bingham kernels and construct kernel mixture density models on this space.

## 5 Summary & Future Work

We consider kernel mixture density models on general compact manifolds and obtain sufficient conditions on the kernel, priors and the space for the density estimate to be strongly consistent. Thereby we extend the existing literature on strong posterior consistency on  $\Re^d$  using Gaussian kernels to more general non-Euclidean manifolds. The conditions are verified for specific kernels on two important manifolds, namely the hypersphere and the planar shape space. It is discussed how to extend the kernel choice on these manifolds and construct their counterparts on other manifolds arising as generalisations. The multivariate Gaussian mixture model with an appropriate truncated and transformed Wishart prior on the within cluster covariance inverse is also shown to satisfy the consistency conditions when used to model a compactly supported density on  $\Re^d$ . We also allow the prior to depend on the sample size and obtain sufficient conditions for weak and strong consistency, while expecting better small sample operating characteristics. As a result a truncated Wishart prior on the covariance inverse of a multivariate Gaussian kernel is shown to satisfy the requirements for strong consistency.

In later works we plan to prove the results for other kernels on additional manifolds arising in applications. We also plan to extend the results to cover densities with non-compact support, in particular  $\Re^d$ . Since most of the non-Euclidean manifolds arising in applications are compact, that is not a high priority.

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## 6 Appendix

## 6.1 Proof of Theorem 1

The proof runs on the lines of that of Theorem 1 in Bhattacharya and Dunson (2010a)[2].

#### Proof

First of all we show that the set

$$\left\{P \in \mathcal{M}(M): \sup_{m \in M, \mathcal{K} \in N_{\varepsilon}} |f(m; P, \mathcal{K}) - f(m; F_0, \mathcal{K})| < \varepsilon\right\}$$
(10)

contains a weakly open neighborhood of  $F_0$ ,  $F_0$  being the distribution corresponding to  $f_0$ . The kernel *K* being continuous from assumption A1, for any  $(m, \mathcal{K}) \in M \times N_{\mathcal{E}}$ ,

$$\mathscr{W}_{m,\mathscr{K}} = \{P: |f(m; P, \mathscr{K}) - f(m; F_0, \mathscr{K})| < \varepsilon/3\}$$

defines an open neighborhood of  $F_0$ . The mapping from  $(m, \mathcal{K})$  to  $f(m; P, \mathcal{K})$  is a uniformly equicontinuous family of functions on  $M \times N_{\mathcal{E}}$ , labeled by  $P \in \mathcal{M}(M)$ , because, for  $m_1, m_2 \in M$ ;  $\mathcal{K}_1, \mathcal{K}_2 \in N_{\mathcal{E}}$ ,

$$|f(m_1; P, \mathscr{K}_1) - f(m_2; P, \mathscr{K}_2)| \leq \int_M |K(m_1; \mu, \mathscr{K}_1) - K(m_2; \mu, \mathscr{K}_2)| P(d\mu)$$

and *K* is uniformly continuous on  $M \times M \times N_{\varepsilon}$ . Therefore there exists a  $\delta > 0$  such that  $\rho_{12}((m_1, \kappa_1), (m_2, \kappa_2)) < \delta$  implies that

$$\sup_{P} |f(m_1; P, \mathscr{K}_1) - f(m_2; P, \mathscr{K}_2)| < \varepsilon/3.$$

Here  $\rho_{12}$  denotes any distance on  $M \times N$  inducing the product topology. Cover  $M \times N_e$  by finitely many balls of radius  $\delta: M \times N_e = \bigcup_{i=1}^{J} B\{(m_i, \mathscr{K}_i), \delta\}$ . Let  $\mathscr{W}_e = \bigcap_{i=1}^{J} \mathscr{W}_{m_i, \mathscr{K}_i}$  which is an open neighborhood of  $F_0$ . Let  $P \in \mathscr{W}_e$  and  $(m, \mathscr{K}) \in M \times N_e$ . Then there exists a  $(m_i, \mathscr{K}_i)$  such that  $(m, \mathscr{K}) \in B\{(m_i, \mathscr{K}_i), \delta\}$ . Then  $|f(m; P, \mathscr{K}) - f(m; F_0, \mathscr{K})|$ 

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Bhattacharya and Dunson

 $\leq |f(m;P,\mathscr{K}) - f(m_i;P,\mathscr{K}_i)| + |f(m_i;P,\mathscr{K}_i) - f(m_i;F_0,\mathscr{K}_i)| + |f(m_i;F_0,\mathscr{K}_i) - f(m;F_0,\mathscr{K})| \\ \leq \frac{\varepsilon_1}{2} + \frac{\varepsilon_1}{2} + \frac{\varepsilon_2}{2} = \varepsilon.$ 

This proves that the set in (10) contains  $W_{\epsilon}$ , an open neighborhood of  $F_0$ .

For  $P \in W_{\epsilon}$  and  $\mathcal{K} \in N_{\mathcal{E}}$ ,  $\sup_{m} |f_{0}(m) - f(m; P, \mathcal{K})|$ 

$$\leq \sup_{m} |f_0(m) - f(m; F_0, \mathscr{K})| + \sup_{m} |f(m; F_0, \mathscr{K}) - f(m; P, \mathscr{K})| < 2\varepsilon$$

because of assumptions A2 and A4. Hence

$$\Pi\left\{f:\sup_{m\in M}|f(m)-f_0(m)|<2\varepsilon\right\}\geq \Pi_1(\mathscr{W}_{\varepsilon}\times N_{\varepsilon})>0$$

because of assumption A3 and the fact that int( $W_e \times N_e$ ) interects with supp( $\Pi_1$ ). This implies the KL property when  $f_0$  is strictly positive (and hence bounded below), as shown in Corollary 1 of Bhattacharya and Dunson (2010a) [2].

In case  $f_0$  is not bounded below, we use Lemma 4 in Wu and Ghosal (2008)[20] to get a continuous everywhere positive density  $f_1$  (depending on  $f_0$  and e) for which  $\Pi(K_{\varepsilon}(f_1)) \leq \Pi(K_{2\varepsilon+\sqrt{\varepsilon}}(f_0))$ . From what we have proved above,  $\Pi(K_{\varepsilon}(f_1)) > 0$  and as a result the KL condition follows for  $f_0$ .

## 6.2 Proof of Theorem 2

In this proof and the subsequent ones, we shall use a general symbol C for any constant not depending on n (but possibly on  $\varepsilon$ ).

#### Proof

Given  $\delta_1 > 0 \ (\equiv \delta_1(e, n))$ , cover *M* by  $N_1 \ (\equiv N_1(\delta_1))$  many disjoint subsets of diameter at most  $\delta_1: M = \bigcup_{i=1}^{N_1} E_i$ . Assumption **A8** implies that for  $\delta_1$  sufficiently small,  $N_1 \le C \delta_1^{-a_3}$ . Pick  $\mu_i \in E_i$ ,  $i = 1, ..., N_1$ , and define for a probability *P*,

$$P_n = \sum_{i=1}^{N_1} P(E_i) \delta_{\mu_i}, \ P_n(\mathbf{E}) = (P(E_1), \dots, P(E_{N_1}))^T.$$
(11)

Denoting the  $L_1$ -norm as ||.||, for any  $\mathcal{K}$  with  $\varphi(\mathcal{K}) = \kappa_n$ ,

$$\begin{split} \left\| f(P,\mathscr{K}) - f(P_n,\mathscr{K}) \right\| &\leq \sum_{i=1}^{N_1} \int_{E_i} \left\| K(\mu,\mathscr{K}) - K(\mu_i,\mathscr{K}) \right\| P(d\mu) \\ &\leq C \sum_i \int_{E_i} \sup_{m \in \mathcal{M}} |K(m;\mu,\mathscr{K})| - K(m;\mu_i,\mathscr{K})| P(d\mu) \end{split}$$
(12)

$$\leq C\kappa_n^{a_1}\delta_1. \tag{13}$$

The inequality in (13) follows from (12) using Assumption A5.

For  $\kappa$ ,  $\tilde{\kappa}$  in  $\varphi^{-1}$  [0,  $\kappa_n$ ],  $P \in \mathcal{M}(M)$ ,

$$\begin{split} \left\| f(P,\mathscr{K}) - f(P,\widetilde{\mathscr{K}}) \right\| &\leq C \sup_{\substack{m,\mu \in M \\ \mathcal{K} \neq n}} |K(m;\mu,\mathscr{K}) - K(m;\mu,\widetilde{\mathscr{K}})| \\ &\leq C \kappa_n^{a_2} \rho_2(\mathscr{K},\widetilde{\mathscr{K}}), \end{split}$$
(14)

the inequality in (14) following from Assumption A6. Given  $\delta_2 > 0 \ (\equiv \delta_2(e, n))$ , cover  $\varphi^{-1}$  $[0, \kappa_n]$  by finitely many subsets of diameter at most  $\delta_2$ , the number of such subsets required being at most  $C(\kappa_n \delta_2^{-1})^{b_2}$ , from Assumption A7. Call the collection of these subsets  $W(\delta_2, \delta_2)$ *n*).

Letting  $S_d = \{x \in [0, 1]^d : \Sigma x_i = 1\}$ ,  $S_d$  is compact under the  $L^1$ -metric  $(||x||_{L^1} = \Sigma |x_i|, x \in \mathbb{R})$  $\Re^d$ ), and hence given any  $\delta_3 > 0 (\equiv \delta_3(e))$ , can be covered by finitely many subsets of the cube  $[0, 1]^d$  each of diameter at most  $\delta_3$ . In particular cover  $S_d$  with cubes of side length  $\delta_3/$ *d* lying partially or totally in  $S_d$ . Then an upper bound on the number  $N_2 \equiv N_2(\delta_3, d)$  of such

$$L(S_d(1+\delta_3))$$

cubes can be shown to be  $\frac{\lambda(S_d(1+\delta_3))}{(\delta_3/d)^d}$ ,  $\lambda$  denoting the Lebesgue measure on  $\Re^d$  and  $S_d(r) = \{x \in [0, \infty)^d: \Sigma x_i \ r\}$ . Since  $\lambda(S_d(r)) = r^d/d!$ , hence

$$N_2(\delta_3, d) \leq \frac{d^d}{d!} \left(\frac{1+\delta_3}{\delta_3}\right)^d.$$

Let  $\mathcal{W}(\delta_3, d)$  denote the partition of  $S_d$  as constructed above.

Let  $d_n = N_1(\delta_1)$ . For 1  $i \quad N_2(\delta_3, d_n), 1 \le j \le C(\kappa_n \delta_2^{-1})^{b_2}$ , define  $\mathcal{D}_{ij} = \{ f(P, \mathcal{K}) : P_n(\mathbf{E}) \in \mathcal{W}_i, \ \mathcal{K} \in W_j \},\$ 

with  $W_i$  and  $W_j$  being elements of  $W(\delta_3, d_n)$  and  $W(\delta_2, n)$  respectively. We claim that this subset of  $\mathcal{D}_n$  has  $L^1$  diameter of at most e. For  $f(P, \mathcal{K})$ ,  $f(\tilde{P}, \tilde{\kappa})$  in this set,  $||f(P, \mathcal{K}) - f(\tilde{P}, \tilde{\kappa})||$ 

$$\left\| f(P, \mathcal{K}) - f(P_n, \mathcal{K}) \right\| + \left\| f(P_n, \mathcal{K}) - f(\tilde{P}_n, \mathcal{K}) \right\| + \\ + \left\| f(\tilde{P}_n, \mathcal{K}) - f(\tilde{P}, \mathcal{K}) \right\| + \left\| f(\tilde{P}, \mathcal{K}) - f(\tilde{P}, \tilde{\mathcal{K}}) \right\|.$$

$$(15)$$

From inequality (13), it follows that the first and third terms in (15) are at most  $C \kappa_n^{a_1} \delta_1$ . The second term can be bounded by

$$\sum_{i=1}^{d_n} |P(E_i) - \tilde{P}(E_i)| < \delta_3$$

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Bhattacharya and Dunson

and from the inequality in (14), the fourth term is bounded by  $C\kappa_n^{a_2}\delta_2$ . Hence the claim holds if we choose  $\delta_1 = C\kappa_n^{-a_1}$ ,  $\delta_2 = C\kappa_n^{-a_2}$ , and  $\delta_3 = C$ . The number of such subsets covering  $\mathcal{D}_n$  is at most  $CN_2(\delta_3, d_n)(\kappa_n\delta_2^{-1})^{b_2}$ . From Assumption **A8**, it follows that for *n* sufficiently large,

$$d_n = N_1(\delta_1) \le C \kappa_n^{a_1 a_3}.$$

Using the Stirling's formula, we can bound  $\log(N_2(\delta_3, d_n))$  by  $Cd_n$ . Also  $\kappa_n \delta_2^{-1}$  is bounded by  $C\kappa_n^{a_2+1}$ , so that  $N(e, \mathcal{D})$ 

$$C + C\log(\kappa_n) + Cd_n \leq C\kappa_n^{a_1a_3}$$

for *n* sufficiently large.

## 6.3 Proof of Proposition 3

Proof

In Lemma 5 of Wu and Ghosal (2010)[21], it is shown that

$$\left\|K(\mu,\mathcal{K}){-}K(\nu,\mathcal{K})\right\| \,\leq\, C\lambda_d(\mathcal{K})^{1/2}\rho(u,\nu)$$

for all  $\mu$ ,  $\nu \in \mathbb{R}^d$  and  $\kappa \in M^+(d)$ . This means that A5' is satisfied with  $a_1 = 1/2$ . Also by the geometry of  $\mathbb{R}^d$ , A8 is satisfied with  $a_3 = d$ .

To show A7, note that  $\varphi^{-1}[0, \kappa]$  is a subset of  $M^+(d)$  consisting of those positive matrices whose eigen values are bounded by  $\kappa$ . We equip  $M^+(d)$  with the  $L_2$  norm distance, i.e.

$$\rho_2(\mathcal{K}_1, \mathcal{K}_2) = \left\| \mathcal{K}_1 - \mathcal{K}_2 \right\|_2, \left\| \mathcal{K} \right\|_2^2 = \sum_{ij} \mathcal{K}_{ij}^2 = \operatorname{Trace}(\mathcal{K}^2)$$

and view it as a subset of M(d) - the space of all order *d* real matrices. Then  $\varphi^{-1}[0, \kappa]$  is contained in a ball of radius  $\sqrt{d\kappa}$  around the zero matrix. The *\varepsilon*-covering number of a such a ball is of the order  $(\sqrt{\kappa}/\varepsilon)^{d^2}$ . Hence **A7** is also satisfied.

Remains to check  $\mathbf{A6}'$ . Since  $\varphi^{-1}[0, \kappa]$  is a convex subset of M(d), use the Taylor's theorem to get

$$|K(x;\mu,\mathscr{K}_1) - K(x;\mu,\mathscr{K}_2)| \le \rho_2(\mathscr{K}_1,\mathscr{K}_2) \sup_{\mathscr{K} \in \varphi^{-1}[0,\kappa]} \left\| \frac{\partial}{\partial \mathscr{K}} K(x;\mu,\mathscr{K}) \right\|_2$$

for all  $x, \mu \in \mathbb{R}^d$ , and  $\mathcal{K}_1, \mathcal{K}_2 \in \varphi^{-1}$  [0,  $\kappa$ ]. This in turn implies that

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Bhattacharya and Dunson

$$\begin{split} & \left\| K(\mu, \mathcal{K}_1) - K(\mu, \mathcal{K}_2) \right\| \leq \\ \rho_2(\mathcal{K}_1, \mathcal{K}_2) \int_{\mathbb{S}^{d} \mathcal{K} \in \varphi^{-1}[0,\kappa]} \sup \left\| \frac{\partial}{\partial \mathcal{K}} K(x;\mu, \mathcal{K}) \right\|_2 dx. \end{split}$$

Some calculation will show that

$$\begin{split} \left\| \frac{\partial}{\partial \mathcal{K}} K(x;\!\mu,\mathcal{K}) \right\|_{2} &\leq \\ CK(x;\!\mu,\mathcal{K})(\sqrt{\sum_{1}^{d} \lambda_{j}^{-2}(\mathcal{K})} + \left\| x - \mu \right\|_{2}^{2}), \end{split}$$

*C* being some constant independent of  $x, \mu$  or  $\mathcal{K}$ . Since  $\varphi^{-1}[0, \kappa]$  consists of all positive matrices  $\mathcal{K}$  whose eigen values lie in  $[\lambda_1, \kappa]$ , this will imply that

$$\left\| K(\mu, \mathscr{K}_1) - K(\mu, \mathscr{K}_2) \right\| \le C \kappa^{d/2} \rho_2(\mathscr{K}_1, \mathscr{K}_2)$$

which means that A6' is also satisfied with  $a_2 = d/2$ . Here *C* denotes a constant independent of  $\mu$ ,  $\kappa_1$ ,  $\kappa_2$  or  $\kappa$ . The rest of the proof follows from Theorem 2 and Proposition 2.

## 6.4 Proof of Lemma 1

Proof

Fix  $\varepsilon > 0$ . Under assumptions A1, A2 and A4, it follows from the proof of Theorem 1 that there exists a weakly open neighborhood W of  $F_0$  (depending on  $\varepsilon$ ) such that  $K_{\varepsilon}(f_0)$  contains  $\{f(P, \kappa): P \in W, \kappa \in N_{\varepsilon}\}$ . Hence

$$\int_{1}^{n} \frac{f(X_{i})}{f_{0}(X_{i})} \Pi_{n}(df) \geq \int_{K_{\varepsilon}(f_{0})} \prod_{1}^{n} \frac{f(X_{i})}{f_{0}(X_{i})} \Pi_{n}(df)$$
  
$$\geq \int_{\mathscr{W} \times N_{\varepsilon}} \prod_{1}^{n} \frac{f(X_{i}; P, \mathscr{K})}{f_{0}(X_{i})} \pi_{n}(\mathscr{K}) \Pi_{11}(dP) \lambda_{1}(d\mathscr{K}).$$

By the law of large numbers, for any  $f \in K_{\mathcal{E}}(f_0)$ ,

$$\frac{1}{n} \sum_{i} \log\{(f_0/f)(X_i)\} \to \operatorname{KL}(f_0; f) < \varepsilon$$

a.s.  $F_0^{\infty}$  as  $n \to \infty$ . Therefore for any  $P \in W$  and  $\mathcal{K} \in N_{\mathcal{E}}$ ,

$$\lim_{n} \inf_{n} \exp(2n\varepsilon) \prod_{1}^{n} \frac{f(X_{i}; P, \mathscr{K})}{f_{0}(X_{i})} = \\\lim_{n} \inf_{n} \exp[n[2\varepsilon - (1/n)\sum_{i} \log\{f_{0}(X_{i})/f(X_{i}; P, \mathscr{K})\}]] = \infty \text{ a.s. } F_{0}^{\infty}.$$

Also from Assumption A11,  $\liminf_{n \to \infty} \exp(n\epsilon) \pi_n(\kappa) = \infty \forall \kappa \in N_{\epsilon}$  and hence

$$\liminf_{n} \operatorname{infexp}(3n\varepsilon) \prod_{1}^{n} \frac{f(X_{i}; P, \mathscr{K})}{f_{0}(X_{i})} \pi_{n}(\mathscr{K}) = \infty \text{ a.s. } F_{0}^{\infty}.$$

By Fubini-Tonelli theorem, there exists a  $\Omega_0 \subset \Omega$  with probability 1 such that for any  $\omega \in \Omega_0$ ,

$$\liminf_{n} \operatorname{infexp}(3n\varepsilon) \prod_{1}^{n} \frac{f(X_{i}(\omega); P, \mathscr{K})}{f_{0}(X_{i}(\omega))} \pi_{n}(\mathscr{K}) = \infty$$

for all  $(P, \mathcal{K}) \in \mathcal{W} \times N_e$  outside of a  $\Pi_{11}(dP) \otimes \lambda_1(d\mathcal{K})$  measure 0 subset. By Assumption A10, and since  $\lambda_1$  has full support and  $N_e$  has a non-empty interior,  $(\Pi_{11} \otimes \lambda_1)(\mathcal{K} \times N_e) > 0$ . Therefore using the Fatou's lemma, we conclude that

$$\lim_{n} \inf \exp(3n\varepsilon) \int \prod_{1}^{n} \frac{f(X_{i})}{f_{0}(X_{i})} \Pi_{n}(df) \geq \int \prod_{n}^{n} \frac{f(X_{i})}{f_{0}(X_{i})} \prod_{1} \frac{f(X_{i};P,\mathcal{K})}{f_{0}(X_{i})} \pi_{n}(\mathcal{K}) \right\} \prod_{1} (dP)\lambda_{1}(d\mathcal{K}) = \infty \text{ a.s. } F_{0}^{\infty}.$$

Since  $\varepsilon$  was arbitrary, the proof is completed.

## 6.5 Proof of Proposition 4

#### Proof

For any a > 0,

$$\pi_{n}\{\varphi^{-1}(n^{a},\infty)\} = Pr(\lambda_{d}(\mathcal{K}) > n^{a}), \ \mathcal{K} \sim \pi_{n}$$

$$= Pr(\lambda_{d}(X) > n^{a} | \lambda_{1}(X) \ge \lambda_{1}), \ X \sim \text{Wish}(\beta_{n}^{-1}I_{d}, q)$$

$$\leq Pr(\lambda_{d}(X) > n^{a}) / Pr(\lambda_{1}(X) \ge \lambda_{1})$$

$$= Pr(\lambda_{d}(Z) > \beta_{n}n^{a}) / Pr(\lambda_{1}(Z) \ge \beta_{n}\lambda_{1}), \ Z \sim \text{Wish}(I_{d}, q),$$
(16)

the last identity following because X equals to  $\beta_n^{-1}Z$  in distribution. The numerator in (16) is less than  $Pr(\text{Tr}(Z) > \beta_n n^a)$ . The trace of Z follows a Gam(1/2, qd/2) distribution which has exponentially decaying tail. Hence the numerator is less than  $\exp(-C\beta_n n^a)$  for some C > 0 when *n* is sufficiently large.

Now we derive a lower bound for the probability in the denominator of (16). In Mallik (2003)[15], the joint density of  $\lambda_1(Z), \ldots, \lambda_d(Z)$  has been shown to be

$$f(x_1, \dots, x_d) = \frac{\left(\prod_{i=1}^d x_i\right)^{q-d} \exp(-\sum_{i=1}^d x_i) \prod_{1 \le i \le j \le d} (x_i - x_j)^2}{\prod_{i=1}^d (d-i)! (q-i)!}, 0 < x_1 < \dots < x_d < \infty.$$

Hence  $Pr(\lambda_1(Z) \quad \beta_n \lambda_1) = Pr(\lambda_j(Z) \quad \beta_n \lambda_1 \forall j) =$ 

$$C \int_{\beta_n \lambda_1 \le x_1 < \dots < x_d < \infty} (\prod_{i=1}^d x_i)^{q-d} \exp(-\sum_{i=1}^d x_i) \prod_{1 \le i < j \le d} (x_i - x_j)^2 \prod_{i=1}^d dx_i$$

$$\geq C \int_{\beta_n \lambda_1 \le x_1 < \dots < x_d < \infty} \exp(-2\sum_{i=1}^d x_i) \prod_{1 \le i < j \le d} (x_i - x_j)^2 \prod_{i=1}^d dx_i$$
(17)

for *n* sufficiently large. Integrate (17) by parts to get  $Pt(\lambda_1(Z) \quad \beta_n \lambda_1) = \exp(-C\beta_n)$  for appropriate C > 0 when *n* is sufficiently large. Hence there exists a C > 0 such that the ratio in (16) is less than  $\exp(-C\beta_n n^a)$ . If we pick *a* as in the Proposition, for any  $\beta_0 > 0$ , it follows that

$$\exp(n\beta_0)\pi_n\{\varphi^{-1}(n^a,\infty)\} < \exp\{-n(C\beta_n n^{a-1} - \beta_0)\}$$

which converges to zero because  $\beta_n n^{a-1}$  diverges to infinity. This verifies Assumption A12 with *a* as in the Proposition and  $\beta_0$  being any positive constant.

## 6.6 Proof of Theorem 5

#### Proof

Denote by *M* the unit sphere  $S^d$  and by  $\rho$  the chord distance on it. Express the vMF kernel as

$$K(m;\mu,\mathcal{K}) = c^{-1}(\mathcal{K}) \exp[K\{1-\rho^2(m,\mu)/2\}] \quad (m,\mu \in M; \mathcal{K} \in [0,\infty)).$$

Since  $\rho$  is continuous on the product space  $M \times M$  and c is continuous and non-vanishing on  $[0, \infty)$ , K is continuous on  $M \times M \times [0, \infty)$  and assumption A1 follows.

For a given continuous function f on  $M, m \in M, \mathcal{K} = 0$ , define

$$I(m, \mathscr{K}) = f(m) - \int_{M} K(m; \mu, \mathscr{K}) f(\mu) V(d\mu) = \int_{M} K(m; \mu, \mathscr{K}) \{ f(m) - f(\mu) \} V(d\mu).$$

Then assumption A2 follows once we show that

$$\lim_{\mathscr{K}\to\infty}(\sup_{m\in M}|I(m,\mathscr{K})|)=0.$$

To simplify  $I(m, \mathcal{K})$ , make a change of coordinates  $\mu \mapsto \tilde{\mu} = U(m)^T \mu$ ,  $\tilde{\mu} \mapsto \theta \in \Theta_d \equiv (0, \pi)^{d-1} \times (0, 2\pi)$  where U(m) is an orthogonal matrix with first column equal to *m* and  $\theta = (\theta_1, ..., \theta_d)^T$  are the spherical coordinates of  $\tilde{\mu} \equiv \tilde{\mu}(\theta)$  which are given by

$$\tilde{\mu}_j = \cos\theta_j \prod_{h < j} \sin\theta_h, \ j = 1, \dots, d, \quad \tilde{\mu}_{d+1} = \prod_{j=1}^d \sin\theta_j.$$

Using these coordinates, the volume form can be written as

$$V(d\mu) = V(d\tilde{\mu}) = \sin^{d-1}(\theta_1) \sin^{d-2}(\theta_2) \dots \sin(\theta_{d-1}) d\theta_1 \dots d\theta_d$$

and hence  $I(m, \mathcal{K})$  equals

$$\int_{\Theta_d} c^{-1}(\mathscr{K}) \exp\{\mathscr{K}\cos(\theta_1)\}\{f(m) - f(U(m)\tilde{\mu})\}\sin^{d-1}(\theta_1) \dots \sin(\theta_{d-1})d\theta_1 \dots d\theta_d$$
$$= c^{-1}(\mathscr{K}) \int_{\substack{\Theta_{d-1} \times (-1,1) \\ \sin^{d-2}(\theta_2) \dots \sin(\theta_{d-1})d\theta_2 \dots d\theta_d dt}} \exp(\mathscr{K}t)\{f(m) - f(U(m)\tilde{\mu})\}(1-t^2)^{d/2-1}$$
(18)

where  $t = \cos(\theta_1)$ ,  $\tilde{\mu} = \tilde{\mu}(\theta(t))$  and  $\theta(t) = (\arccos(t), \theta_2, ..., \theta_d)^T$ . In the integrand in (18), the distance between *m* and  $U(m) \tilde{\mu}$  is  $\sqrt{2(1-t)}$ . Substitute  $t = 1 - \kappa^{-1}s$  in the integral with  $s \in (0, 2^{\kappa})$ . Define

$$\Phi(s, \mathcal{K}) = \sup\{|f(m) - f(\tilde{m})|: m, \tilde{m} \in M, \rho(m, \tilde{m}) \le \sqrt{2\mathcal{K}^{-1}s}\}$$

Then

$$|f(m) - f(U(m)\tilde{\mu})| \le \Phi(s, \mathscr{K}).$$

Since *f* is uniformly continuous on  $(M, \rho)$ , therefore  $\Phi$  is bounded on  $(\Re^+)^2$  and  $\lim_{\kappa \to \infty} \Phi(s, \kappa) = 0$ . Hence from (18), we deduce that  $\sup_{m \in M} |I(m, \kappa)|$ 

$$c^{-1}(\mathscr{K})\mathscr{K}^{-1} \int_{\substack{\Theta_{d-1} \times (0,2\mathscr{K}) \\ \sin^{d-2}(\theta_2) \dots \sin(\theta_{d-1})d\theta_2 \dots d\theta_d ds \leq \\ C\mathscr{K}^{-d/2}\tilde{c}^{-1}(\mathscr{K}) \int_{0}^{\infty} \Phi(s, \mathscr{K}) e^{-s} s^{d/2-1} ds.}$$
(19)

From Lemma 3, it follows that

$$\limsup_{\mathscr{K}\to\infty}\mathscr{K}^{-d/2}\tilde{c}^{-1}(\mathscr{K})<\infty.$$

This in turn, using the Lebesgue Dominated Convergence Theorem implies that the expression in (19) converges to 0 as  $\mathcal{K} \to \infty$ . This verifies assumption A2.

## 6.7 Proof of Theorem 6

In the proof,  $B_d(r)$  denotes the ball of radius *r* around 0 in  $\mathbb{R}^d$ :

$$B_d(r) = \{x \in \Re^d : ||x||_2 \le r\}$$

and  $B_d$  refers to  $B_d(1)$ .

#### Proof

It is clear from (6) and (7) that the vMF kernel *K* is continuously differentiable on  $\mathbb{R}^{d+1} \times \mathbb{R}^{d+1} \times [0, \infty)$ . Hence

$$\sup_{m \in S^d, \mathcal{K} \in [0,\kappa]} |K(m;\mu,\mathcal{K}) - K(m;\nu,\mathcal{K})| \le \sup_{m \in S^d, x \in B_{d+1}, \mathcal{K} \in [0,\kappa]} \left\| \frac{\partial}{\partial x} K(m;x,\mathcal{K}) \right\|_2 \left\| \mu - \nu \right\|_2.$$

Since

$$\frac{\partial}{\partial x}K(m;x,\mathcal{K}) = \mathcal{K}\tilde{c}^{-1}(\mathcal{K})\exp\{-\mathcal{K}(1-m^Tx)\}m$$

its norm is bounded by  $\mathcal{K}\tilde{c}^{-1}(\mathcal{K})$ . Lemma 3 implies that this in turn is bounded by

$$\kappa \tilde{c}^{-1}(\kappa) \le C \kappa^{d/2 + 1}$$

for  $\kappa$  and  $\kappa$  1. This proves assumption A5 with  $a_1 = d/2 + 1$ .

To verify A6, given  $\kappa_1$ ,  $\kappa_2 = \kappa$ , use the inequality,

$$\sup_{m,\mu\in S^d} |K(m;\mu,\mathscr{K}_1) - K(m;\mu,\mathscr{K}_2)| \le \sup_{m,\mu\in S^d,\mathscr{K}\leq\kappa} \left|\frac{\partial}{\partial\kappa} K(m;\mu,\mathscr{K})\right| |\mathscr{K}_1 - \mathscr{K}_2|.$$

By direct computations, one can show that

$$\begin{split} &\frac{\partial}{\partial \mathcal{H}} K(m; \mu, \mathcal{H}) = -\frac{\partial}{\partial \mathcal{H}} \tilde{c}(\mathcal{H}) \tilde{c}^{-2}(\mathcal{H}) \exp\{-\mathcal{H}(1-m^{T}\mu)\} \\ &- \tilde{c}^{-1}(\mathcal{H}) \exp\{-\mathcal{H}(1-m^{T}\mu)\}(1-m^{T}\mu), \\ &\frac{\partial}{\partial \mathcal{H}} \tilde{c}(\mathcal{H}) = -C \int_{-1}^{1} \exp\{-\mathcal{H}(1-t)\}(1-t)(1-t^{2})^{d/2-1} dt, \\ &\left|\frac{\partial}{\partial \mathcal{H}} \tilde{c}(\mathcal{H})\right| \leq C \tilde{c}(\mathcal{H}). \end{split}$$

Therefore, using Lemma 3,

$$\left|\frac{\partial}{\partial \mathscr{K}}K(m;\mu,\mathscr{K})\right| \le C\tilde{c}^{-1}(\mathscr{K}) \le C\tilde{c}^{-1}(\kappa) \le C\kappa^{d/2}$$

for any  $\mathcal{K} = \kappa$  and  $\kappa = 1$ . Hence **A6** is verified with  $a_2 = d/2$ .

Finally to verify **A8**, note that  $S^d \subset B_{d+1} \subset [-1, 1]^{d+1}$  which can be covered by finitely many cubes of side length  $\varepsilon/(d+1)$ . Each such cube has  $L_2$  diameter  $\varepsilon$ . Hence their intersections with  $S^d$  provides a finite  $\varepsilon$ -cover for this manifold. If  $\varepsilon < 1$ , such a cube intersects with  $S^d$  only if it lies entirely in  $B_{d+1}(1+\varepsilon) \cap B_{d+1}(1-\varepsilon)^c$ . The number of such cubes, and hence the  $\varepsilon$ -cover size can be bounded by

$$C\varepsilon^{-(d+1)}\{(1+\varepsilon)^{d+1} - (1-\varepsilon)^{d+1}\} \le C\varepsilon^{-d}$$

for some C > 0 not depending on e. This verifies **A8** for appropriate positive constant  $A_3$  and  $a_3 = d$ .

## 6.8 Proof of Lemma 3

Proof

Express  $\tilde{c}(\mathcal{K})$  as

$$C \int_{-1}^{1} \exp\{-\mathscr{K}(1-t)\} (1-t^2)^{d/2-t} dt$$

and it is clear that it is decreasing. This expression suggests that

$$\begin{split} \tilde{c}(\mathscr{K}) &\geq C \int_{0}^{1} \exp\{-\mathscr{K}(1-t)\}(1-t^{2})^{d/2-1} dt \\ &\geq C \int_{0}^{1} \exp\{-\mathscr{K}(1-t^{2})\}(1-t^{2})^{d/2-1} dt \\ &= C \int_{0}^{1} \exp\{-\mathscr{K}u\} u^{d/2-1}(1-u)^{-1/2} du \\ &\geq C \int_{0}^{1} \exp\{-\mathscr{K}u\} u^{d/2-1} du \\ &= C \mathscr{K}^{-d/2} \int_{0}^{K} \exp\{-v\} v^{d/2-1} du \\ &\geq C \{\int_{0}^{1} \exp(-v) v^{d/2-1} dv\} \mathscr{K}^{-d/2} \end{split}$$

if κ 1.

## 6.9 Proof of Theorem 7

Proof

Express the complex Watson kernel as

$$K(m;\mu,\mathscr{K})=c^{-1}(\mathscr{K})\exp\left(\frac{-\mathscr{K}}{2}d_{E}^{2}(m,\mu)\right).$$

Given  $\mathcal{K} = 0$ , define

$$\varphi(t) = \exp\left(\frac{-\mathscr{K}}{2}t^2\right), t \in [0, \sqrt{2}].$$

Then  $|\varphi'(t)| \leq \sqrt{2}\mathcal{K}$ , so that

$$|\varphi(t)-\varphi(s)| \leq \sqrt{2}\mathcal{K}|s-t|, s,t \in [0, \sqrt{2}]$$

NIH-PA Author Manuscript

Bhattacharya and Dunson

which implies that

$$|K(m;\mu,\mathscr{K}) - K(m;\nu,\mathscr{K})| \leq c^{-1}(\mathscr{K}) \sqrt{2} \mathscr{K} |d_E(m,\mu) - d_E(m,\nu)|$$
  
$$\leq \sqrt{2} \mathscr{K} c^{-1}(\mathscr{K}) d_E(\mu,\nu).$$
(20)

For  $\kappa$ , from Lemma 4, it follows that

$$\mathcal{K}c^{-1}(\mathcal{K}) \le \kappa c^{-1}(\kappa) = \pi^{2-k} \kappa^{k-1} \tilde{c}^{-1}(\kappa)$$
$$\le \pi^{2-k} \kappa^{k-1} \tilde{c}^{-1}(1)$$

provided  $\kappa$  1. Hence for any  $\kappa$  1,

$$\sup_{\mathcal{K}\in[0,\kappa]}\mathcal{K}c^{-1}(\mathcal{K})\leq C\kappa^{k-1}$$

and from inequality (20),  $a_1 = k - 1$  follows.

By direct computation, one can show that

$$\frac{\partial}{\partial \mathscr{K}} K(m;\mu,\mathscr{K}) = \pi^{k-2} \exp\left\{-\frac{1}{2} \mathscr{K} d_{E}^{2}(m,\mu) - \mathscr{K}\right\} \times c^{-2}(\mathscr{K}) \mathscr{K}^{2-k} \left[\sum_{r=k-1}^{\infty} \frac{\mathscr{K}^{r-1}}{r!} \{k - 2 - \frac{r}{2} d_{E}^{2}(m,\mu)\}\right].$$
<sup>(21)</sup>

Denote by *S* the sum in the second line of (21). Since  $d_E^2(m,\mu) \leq 2$ , it can be shown that

$$|k-2 - \frac{r}{2}d_{E}^{2}(m,\mu)| \leq \begin{cases} k-2 & \text{if } k-1 \leq r \leq 2k-4, \\ r-k+2 & \text{if } 2k-3 \leq r, \end{cases}$$

so that

$$\begin{split} |S| &\leq (k-2) \sum_{r=k-1}^{2k-4} \frac{\mathscr{H}^{r-1}}{r!} + \sum_{r=2k-3}^{\infty} \frac{\mathscr{H}^{r-1}}{r!} (r-k+2) \\ &= (k-2) \mathscr{H}^{k-2} \sum_{r=0}^{k-3} \frac{\mathscr{H}^{r}}{(r+k-1)!} + \mathscr{H}^{2k-4} \sum_{r=0}^{\infty} \frac{\mathscr{H}^{r}}{(r+2k-3)!} (r+k-1) \\ &\leq C \mathscr{H}^{k-2} e^{\mathscr{H}} + \mathscr{H}^{2k-4} e^{\mathscr{H}}. \end{split}$$

Plug the above inequality in (21) to get

$$\left|\frac{\partial}{\partial\mathcal{H}}K(m;\mu,\mathcal{H})\right| \le Cc^{-2}(\mathcal{H})\mathcal{H}^{2-k}\exp\left\{-\frac{1}{2}\mathcal{H}d_{E}^{2}(m,\mu)\right\}(C\mathcal{H}^{k-2}+\mathcal{H}^{2k-4})$$
$$\le Cc^{-2}(\mathcal{H})(C+\mathcal{H}^{k-2}).$$
(22)

For  $\kappa$  and  $\kappa$  1, using Lemma 4, we bound the expression in (22) by

$$Cc^{-2}(\kappa)(C+\kappa^{k-2}) = C\kappa^{2k-6}\tilde{c}^{-2}(\kappa)(C+\kappa^{k-2})$$
  
$$\leq C\kappa^{2k-6}\tilde{c}^{-2}(1)(C+\kappa^{k-2}) \leq C\kappa^{3k-8}$$
(23)

for  $\kappa$  sufficiently large. Since *K* is a continuously differentiable in  $\mathcal{K}$ , from (23) it follows that there exists  $\kappa_1 > 0$  such that for all  $\kappa = \kappa_1$ ,  $\mathcal{K}_1$ ,  $\mathcal{K}_2 = \kappa$ ,

$$\sup_{\substack{m,\mu\in\sum_{2}^{k}\\ \leq C\kappa^{3k-8}|\mathcal{K}_{1}-\mathcal{K}_{2}|}} |K(m;\mu,\mathcal{K}_{1})-K(m;\mu,\mathcal{K}_{2})| \leq \sup_{\substack{m,\mu\in\sum_{2}^{k},\mathcal{K}\in[0,\kappa]\\ \leq C\kappa^{3k-8}|\mathcal{K}_{1}-\mathcal{K}_{2}|} \left|\frac{\partial}{\partial\mathcal{K}}K(m;\mu,\mathcal{K})\right| |\mathcal{K}_{1}-\mathcal{K}_{2}|$$

This proves Assumption A6 with  $a_2 = 3k - 8$ .

#### 6.10 Proof of Theorem 8

In the proof,  $C_i$ , i = 1, 2, ... denote positive constants possibly depending on k.

#### Proof

The preshape sphere  $C s^{k-2}$ , as a real manifold, can be identified with the real unit sphere  $S^{2k-3}$ . Endow it with the chord distance induced by the  $L^2$ -norm

$$\left\| u \right\|_{2} = \sqrt{\sum_{i=1}^{k-1} |u_{i}|^{2}} (u = (u_{1}, \dots, u_{k-1})^{T}).$$

Then from Theorem 6, it follows that given any  $\delta > 0$ ,  $C S^{k-2}$  can be covered by finitely many subsets of diameter less than or equal to  $\delta$ , the number of such subsets being bounded

by  $C_1 \delta^{-(2k-3)} + C_2$ . The extrinsic distance  $d_E$  on  $\sum_{k=2}^{k} c_{k}$  can be bounded by the chord distance on  $c s^{k-2}$  as follows. For  $u, v \in c s^{k-2}$ ,

$$\begin{aligned} \left| u - v \right|_{2}^{2} &= 2 - 2 \operatorname{Re}(u^{*}v) \geq 2 - 2 |u^{*}v| = 2(1 - |u^{*}v|) \\ &\geq (1 + |u^{*}v|)(1 - |u^{*}v|) = \frac{1}{2} d_{r}^{2}([u], [v]). \end{aligned}$$

Hence  $d_{\varepsilon}([u], [v]) \leq \sqrt{2} ||u-v||_2$ , so that given any  $\varepsilon > 0$ , the shape image of a  $\delta$ -cover for  $c s^{k-2}$  with  $\delta = \varepsilon / \sqrt{2}$  provides an  $\varepsilon$ -cover for  $\sum_{2}^{k}$ . Hence the  $\varepsilon$ -covering size for  $\sum_{2}^{k}$  can be bounded by  $C_1 \varepsilon^{-(2k-3)} + C_2$ .