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The Prevalence of Cognitive Impairment in Older Adults with Heart Failure

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Abstract

BACKGROUND—Our understanding of the prevalence of cognitive impairment (CI) in older adults with heart failure (HF) in a nationally-representative sample is limited.

Conflict of Interest

Author Contributions:

Gure: Study concept and design, data analysis and interpretation, manuscript preparation.

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Blaum: Data analysis and interpretation, manuscript preparation

Giordani: Data interpretation, manuscript preparation.

Koelling: Data interpretation and manuscript preparation.

Galecki: Data interpretation, manuscript preparation

Pressler: Data interpretation, manuscript preparation.

Hummel: Data interpretation and manuscript preparation.

Langa: Study concept and design, data interpretation, manuscript preparation.

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OBJECTIVES—We used a national probability sample to determine the prevalence of CI in older adults with HF.

DESIGN—Cross-sectional analysis of the 2004 wave of the nationally representative Health and Retirement Study linked to 2002–04 Medicare administrative claims

SETTING—United States, community-dwelling

PARTICIPANTS—6,189 respondents 67 years old.

MEASUREMENTS—An algorithm was developed using a combination of self- and proxy-report of a heart problem and the presence of 1 Medicare claim in administrative files using standard HF diagnostic codes. On the basis of the algorithm, 3 categories were created to characterize the likelihood of a HF diagnosis: 1) High or Moderate Probability of HF; 2) Low Probability of HF; and 3) Not a HF case. Cognitive function was assessed using a screening measure of cognitive function or by proxy rating. Age-adjusted prevalence estimates of CI were calculated for the highmoderate probability HF group, the low probability HF group, and the non-HF cases.

RESULTS—The prevalence of CI consistent with dementia in older adults with HF was 15%; while the prevalence of mild CI was 24%. The odds of dementiain those with HF was significantly increased, even after adjustment for age, education level, net worth and prior stroke (OR: 1.52; 95% CI: 1.14 - 2.02).

CONCLUSION—CI is common in older adults with HF and is independently associated with an increased risk for dementia. A cognitive assessment should be routinely incorporated into HF-focused models of care.

Keywords

Heart failure; aged; cognitive impairment; Medicare claims data

INTRODUCTION

Heart failure (HF) predominantly affects older persons in whom multiple complex chronic conditions often coexist (1–4). HF remains the leading cause of hospitalization in adults 65 years in the United States. Several initiatives exist to reduce the mortality, morbidity, and health care utilization of older adults with HF; however, these recent innovations in HF management have generally not translated into improved mortality and re-hospitalization rates(1, 5, 6). The reasons for unchanged outcomes may stem from the lack of attention given to prognostically important conditions, such as cognitive impairment (CI)(7, 8). Yet, our understanding of the prevalence of CI in older adults with HF is limited. HF patients, who have mild CI or dementia, may have an increased risk for re-hospitalization and mortality because of poor adherence to prescribed therapy due to the complexity of self-management tasks in HF management(9–11).

CI, including both dementia and mild CI is associated with cerebrovascular and cardiovascular disease and is linked to HF (6, 12, 13), but its prevalence and effects on morbidity and mortality are not well understood among those with HF. Prevalence estimates for CI among patients with HF vary widely from 5–75%, with most estimates falling between 25–46% (12–15). Prior work examining HF's association with CI has emphasized a case-control study design using non-representative samples, including inpatients and those with a range of etiologies of HF. Furthermore, the assessment of cognition has not been standardized; investigators have employed a myriad of tools ranging from cognitive screens such as the Mini-Mental State Exam to abbreviated neuropsychological test batteries(15–17). Because of these and other limitations, it is currently difficult to confidently assess the

true extent of CI among older adults with HF. Our research hypotheses are that 1) older adults with HF will have a higher prevalence of CI when compared to other heart condition subgroups; 2) HF is independently associated with a higher risk of CI.

We sought to determine the prevalence of CI among older adults with HF using a nationally representative sample, and also to determine whether HF is independently associated with CI.

METHODS

Data

We used data from the 2004 wave of the Health and Retirement Study (HRS), a biennial, longitudinal survey of a nationally representative cohort of U.S. adults aged 51 years or older.(18) The HRS provides detailed self-report information on chronic diseases and task-specific disabilities.

Sample definition

Ninety percent of HRS respondents provided consent to link their Medicare claims to their HRS survey data. We viewed claims data two years prior to the 2004 interview to capture the greatest number of HRS respondents with HF in the Medicare claims data (19, 20). In the 2004 survey wave, 9,663 respondents were aged 67 and above, ensuring that each individual had 2 full years of Medicare eligibility during our surveillance period. The sample was further reduced to 8,207 after exclusion of those HRS respondents who were not continuously enrolled in Medicare Parts A and B. We excluded 1,650 respondents who were enrolled in Medicare managed care because these plans may not have complete data. Finally, we excluded 429 subjects who were in a nursing home at the time of the analysis. Thus, our analytic sample included 6,189 respondents, representative of approximately 23.4 million community-dwelling adults 67 years old with Medicare fee-for-service in the U.S. in 2004.

Definition of heart failure and other heart problems by self- or proxy-report

Self-or proxy-report heart problem including HF was determined based on survey responses to questions in the HRS 2004 core interview wave. All respondents and proxies were asked: "Has a doctor ever told you that you had a heart attack, coronary heart disease, angina, congestive heart failure (CHF), or other heart problems?" The response to this question was used to sort respondents into two mutually exclusive heart problem categories: 1) no HF or other heart problem; 2) heart problem(including HF).

Claims-based diagnosis of heart failure

HF was defined with Medicare claims by making use of the International Classification of Diseases-Ninth Revision-Clinical Modification (ICD-9-CM) diagnosis codes for primary and secondary diagnoses. Using Medicare Part A and Part B claim files, we determined the claims-based diagnosis of HF for the 2-year period prior to each respondent's HRS interview date (19). We employed ICD-9-CM coding algorithms to identify HF as a diagnosis in inpatient, outpatient and carrier files using codes 398.91, 402.01–402.91, 404.01–404.93, and 428.0–428.90 in a two-year period(19).

Heart Failure case-finding approach

Our prior work identified the test-operating characteristics of self- or proxy-report of HF by comparing to HF diagnosis codes in linked Medicare claims as a diagnostic gold standard (sensitivity: 25%; specificity: 99%)(21). Older adults in the HRS were more likely to report

HF if they had poorer health status. Under-reporting of HF was much more common this study, in particular within minority populations and in those with fewer Medicare claims, perhaps suggesting less severe disease. Therefore, we utilized the presence of a heart problem to better capture respondents who may be unaware of HF as a diagnosis. A newly-developed algorithm was derived from this work (see Figure 1). We utilized a combination of self- and proxy-report of a heart problem and the presence of 1 Medicare claim in inpatient, outpatient, and carrier files using standard HF diagnostic codes to increase our yield of HF cases. On the basis of the algorithm, we created 3 categories to characterize the likelihood of having a HF diagnosis: 1) High/Moderate Probability of HF; 2) Low Probability of HF; and 3) Not a HF case.

Survey-based assessment of cognitive performance

For self-respondents, the presence of dementia was assessed with a modified version of the Telephone Interview for Cognitive Status (TICS-m), a validated cognitive screening instrument patterned on the Mini-Mental State Examination(22). The TICS-m used in the HRS is a 27-point scale and has high sensitivity and specificity for cognitive impairment (CI) and dementia in community samples of older adults (22–24). The test items include: 1) immediate and delayed 10-word recall test to measure memory; 2) serial seven subtraction test to measure working memory; and 3) counting backward test to measure speed of mental processing (25, 26). Cut-scores to identify subjects with CI were based on findings from the Aging, Demographics and Memory Study(27). Based on prior work validating these cutpoints, cognitive function was categorized as normal (12 points) or CI (<12 points) with the CI category divided in to mild (7-11) and moderate/severe impairment (0-6)consistent with dementia (27). Cognitive categories for respondents with proxy informants were based on both the proxy and interviewer assessments of cognitive function. Proxies assessed memory and whether the respondent was cognitively impaired. Cognitive function scores were created based on the proxy interview ratings and the number of IADL limitations, resulting in a scale ranging from 0 to 11, with scores 3-5 indicating, mild cognitive impairment (CI) and scores 6-11 indicating CI consistent with dementia.

Independent variables

The socio-demographic variables included in the analysis as independent variables were: age, race, gender, household net worth, and years of education. Self-reported chronic conditions included were stroke and diabetes. All measures were obtained from the HRS core interview in 2004.

Data analysis

All analyses were conducted using STATA/SE 11. To adjust for the complex sample design of the HRS, the differential probability of selection, and the probability of a respondent's consent to accessing their Medicare data, all analyses were appropriately weighted. Analyses were conducted to calculate population-based characteristics and the prevalence of CI based on sample data. Socio-demographic characteristics, TICS-m, and Proxy rating of cognitive status were compared across the 3 heart condition categories (not HF, low probability, and moderate/high probability HF groups) using adjusted Wald and Pearson chi-square test statistics, as appropriate. We compared prevalence estimates of cognitive function with categorical cognitive variables (normal, mild CI, and moderate/severe CI) by adjusted Wald chi-square statistical tests. Taylor series Linearization was used to estimate design-based standard errors for population estimates (incorporating the stratification and clustering of the HRS sample along with the weights) and compute 95% confidence intervals for population parameters.

We used multivariate logistic regression to identify covariates independently associated with the likelihood of moderate/severe cognitive impairment (CI), by heart disease category. Mild CI was combined with the normal cognition subgroup for these analyses. The final logistic models included the sociodemographic and health status factors that are associated with CI. Model 1 is an univariate analysis. Model 2 is age-adjusted only. Model 3 includes age, race, educational level, net worth, and self-reported prior stroke.

RESULTS

Sample Characteristics

Among the 6,189 respondents in the analytic sample, our case-finding algorithm identified 707 subjects (11.7% of the population) with a moderate/high probability of HF. This corresponds to approximately 2,700,000 Medicare fee-for-service benificiaries 67 years with HF in the U.S.A. in 2004. The mean age was 76.5 years old, 42% were male, 8% were Black, and 4% were Hispanic (Table 1). The mean number of years of education was 12.1. Approximately, 5,800 (92%) respondents were self-respondents; 436 (8%) were represented by a proxy-informant.

The subgroup with moderate/high probability of HF was older, with a mean age of 78.6 years compared to the low probability and no HF subgroups (p < .001). The moderate-high probability HF subgroup had a lower education level.

The majority of respondents in all heart condition categories were self-respondents; however, 8% of respondents utilized a proxy-informant. Twenty-percent of proxy-informants had a high likelihood of HF compared with 29% in the low-probability HF group and 51% in the no HF subgroup. The mean TICS-m score for those who could self-respond was 14.1±4.44. The high/moderate probability HF subgroup had the lowest mean TICS-m score 12.67 ±4.26 (low probability HF 13.68 vs. no HF 14.50; p < .001). There was no statistically significant difference in proxy rating score between the heart condition subgroups.

Prevalence of cognitive impairment (CI), by heart condition category

Theage-adjusted cognitive performance of respondents with HRS-Medicare data, by heart condition category is described in Table 1. The age-adjusted prevalence of moderate/severe CI consistent with dementia in the subgroup with the greatest probability of HF was 15% (95% confidence interval: 12-18%; p .001 when compared to no HF and low probability HF subgroups) and for mild CI it was 24% (95% confidence interval: 21-28%). The age-adjusted prevalence of mild CI among the different heart condition subgroups was not significantly different (Wald test p=0.11 comparing the low probability HF subgroup to the no HF group; p=0.34 when comparing the low probability HF group to the moderate/high probability HF subgroup.)

The association of cognitive impairment (CI) with heart failure

In a series of multivariate logistic regression models shown in Table 3 predicting the likelihood of moderate/severe CI, by heart condition category, HF was consistently associated with significantly increased odds of moderate/severe CI. Age had the greatest effect size in the model, as expected, given its strong association with mild CI and dementia. In the final model, the effect of HF on dementia was further attenuated with the addition of race, educational level, net worth, and self-reported prior stroke(OR: 1.52; 95% confidence interval: 1.14–2.02). Gender and self-reported diabetes were not significant predictors of moderate/severe CI in the final model. The low probability HF subgroup had higher odds of moderate/severe CI when age was the only covariate in the model. However, the odds were

not significantly different compared to the no HF subgroup in the full model (OR: 1.19; 95% confidence interval: 0.94–1.51).

DISCUSSION

Using a unique nationally-representative sample from the HRS- Medicare claims linked data, we found that approximately 40% of older adults 67 years old with HF are cognitively impaired, with 15% demonstrating moderate/severe impairment consistent with dementia. HF was independently associated with an increased risk for cognitive impairment (CI). Our prevalence estimates of CI fell within the range (25–46%) most commonly reported by prior studies, but the prevalence of moderate/severe CI consistent with dementia was lower than previously reported. We also found that HFhas a significant, independent association with dementia after adjustment for important sociodemographic and clinical predictors.

Comparing prevalence estimates across studies is difficult because of differences in patient samples (in HF type and study setting) and the large variability in cognitive tests utilized across studies. Our study's strengths are its use of a broad case-mix of HF cases by use of an algorithm which likely captures the mild-severe stages of HF, the presence of a comparison groups within the cohort, and its use of standardized cognitive measures that are validated from clinically diagnosed cognitively impaired cases from ADAMS, an epidemiologic substudy of dementia in the HRS(26). The TICS-m and informant assessment of cognitive performance discriminate well between demented and mild cognitively impaired cases (22).

Several limitations of our study warrant mention. First, both self-report and Medicare claims data have limited validity as accurate measurements of HF. Multiple factors can influence a patient's awareness of a heart condition such as illness severity, including socioeconomic background, race, patient health awareness gained through encounters with healthcare providers and effective self-management skills. The validity of Medicare administrative data is also highly dependent on many factors: ICD-9 coding practices, organizational culture, coder experience, and the type of administrative file utilized(28). For these reasons, our prevalence estimates may be an over- or underestimate of HF cases in the population. Using HRS-Medicare linked data in prior work, self-report prevalence of HF at 5% is an undercount; claims-based prevalence at 16%, an over count (21). Thus, our algorithm which combines self- and proxy-report of a heart problem with 1 Medicare diagnostic HF claim is likely more accurate than use of self- or proxy-report or claims data alone in this data source.

Second, medical records were not available to further adjudicate cases, ascertain HF severity, or determine the use of medications that might affect cognition. Third, approximately 11% of the HRS respondents did not consent to the linkage of their survey data to Medicare claims. These factors may introduce selection bias since the population enrolled in managed care programs differs from those with traditional Medicare fee-forservice plans with respect to health and socioeconomic status (29). Fourth, the cognitive measures utilized in our study may insufficiently probe the cognitive domains most commonly affected in HF patients leading to inaccuracies in our findings. Our estimates of moderate/severe impairment are lower than prior epidemiological reports likely due to the exclusion of nursing home residents.(30) Nonetheless, this work provides among the first estimates of global cognitive function in a nationally representative sample.

CI commonly co-occurs with HF in older adults, yet assessing cognition is not routinely incorporated into traditional HF disease management models of care. Awareness of cognitive status by healthcare providers, in particular HF care providers, is important for

setting expectations for patient participation in a HF treatment plan and achieving optimal care. We found an independent association of HF with the presence of moderate/severe CI-a cognitive performance level most consistent with dementia. CI is an under-recognized comorbid condition in community-dwelling older adults with HF which will have important effects on prognosis, physical function, quality of life, healthcare utilization, and mortality.

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Figure 1.

Ascertainment of HF cases in HRS-Medicare data via a newly-developed algorithm, weighted % (95% Confidence Interval) Wtd.= weighted Sd=standard deviation CI=cognitive impairment 95% CI=95% confidence interval Prob.=probability

Characteristics of Older Adults 67 years with HRS Data Linked to Medicare Claims (n=6,189)

Characteristics N (wtd. % or mean) weighted n	Not a HF case 3,749 (60.5%) 14,000,000	Low Probability HF 1,733 (27.8%) 6,500,000	Moderate/High Probability HF 707 (11.7%) 2,700,000	Total 6,189 23,365,184	p-value
Age mean (sd)	75.6 (6.51)	77.4 (6.75)	78.6 (6.86)	76.5 (6.71)	< .001
Race/ethnicity n (wtd. %)					
Hispanic	229(4)	75 (3)	37 (3)	341 (4)	.37
Black	435 (8)	211(7)	85 (9)	731 (8)	
White/Other	3,085 (88)	1,447 (89)	585 (88)	5,117 (88)	
Gender n (wtd. %)					<.001
Male	1,414 (38)	828 (48)	334 (47)	2,576 (42)	
Fenale	2,335 (62)	905 (52)	373 (53)	3,613 (58)	
Years of education					<.001
Mean (sd)	12.27 (3.21)	11.92 (3.32)	11.45 (3.32)	12.07 (3.26)	
Net worth (quartile) N (wtd. %)					<.001
\$61,000	931 (23)	503 (27)	263 (36)	1,697 (26)	
61,001–203,400	917 (24)	476 (27)	198 (28)	1,591 (25)	
203,401-512,000	966 (27)	405 (24)	140 (19)	1,511 (25)	
512,000	935 (26)	349 (22)	106 (17)	1,390 (24)	
Stroke n (wtd. %)	246 (6)	232 (14)	151 (21)	629 (10)	<.001
Diabetes n (wtd. %)	593 (15)	418 (24)	234 (.32)	1,245 (20)	< .001
Characteristics N (row, col %) Weighted n	Not a HF case 3,749 (60.5%) 14,000,000	Low Probability HF 1,733 (27.8%) 6,500,000	Moderate/High Probability HF 707 (11.7%) 2,700,000	Total 6,189 23,365,184	p- value
Self -respondents	3,526 (61, 93)	1,607 (28, 91)	620 (11, 85)	5,753 (92)	<.001
Prove reservations	13,000,000 223 (51-7)	5,900,000 17672990	2,300,000 87 (21-15)	21,000,000 436 (8)	
ι τυχу respondence	(1, 11) (27)	120 (22, 2)	0/ (41, 17)	(a) nc+	

Characteristics N (wtd. % or mean) weighted n	Not a HF case 3,749 (60.5%) 14,000,000	Low Probability HF 1,733 (27.8%) 6,500,000	Moderate/High Probability HF 707 (11.7%) 2,700,000	Total 6,189 23,365,184	p-value
	1,000,000	600,000	400,000	1,900,00	
27-point TICS-m score, Overall mean (sd)	14.5 (4.34)	13.7 (4.57)	12.7 (4.26)	14.1 (4.44)	<.001
11-point Proxy assessment of memory Overall mean (sd)	5.09 (3.38)	5.04 (3.32)	5.69 (3.20)	5.20 (3.33)	.41
*Cognitive categoryn,(weighted %; 95% CI)					<.001
Normal	2,698 (71; 70–73)	1,126 (68; 66–70)	411 (61;57–65)	4,235 (69; 68–71)	
Mild CI	778 (21; 19–22)	419 (22; 22–24)	184 (24;21–28)	1,381 (21; 20–23)	
Moderate/Severe CI	271 (8;7–9)	188 (10; 9–11)	112 (15; 12–18)	571 (9; 8–10)	
* Adjusted for age					

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Table 2

Predictors of Moderate/Severe CI, Logistic Regression, Odds Ratio (95% Confidence Interval), n=6,189

Heart Condition Status	Univariate	Age-adjusted	Multivariate adjustment [*]
Not Heart Failure	Ref.	Ref.	Ref.
Low Probability of HF	1.55 (1.24–1.99)	1.33 (1.05–1.68)	1.19 (0.94–1.51)
Moderate/High probabilityof HF	2.50 (1.99–3.14)	1.95 (1.50–2.54)	1.52 (1.14–2.02)

 * Covariates are age, race, educational level, net worth, and self-reported prior stroke