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ANTENATAL MAGNESIUM SULFATE FOR THE PREVENTION OF CEREBRAL PALSY IN PRETERM INFANTS <34 WEEKS' GESTATION: A SYSTEMATIC REVIEW AND META-ANALYSIS

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Abstract

We conducted a systematic review and meta-analysis of randomized controlled trials to determine whether magnesium sulfate administered to women at risk of preterm delivery before 34 weeks of gestation may reduce the risk of cerebral palsy in their children. Six trials involving 4796 women and 5357 infants were included. Antenatal magnesium sulfate was associated with a significant reduction in the risk of cerebral palsy (relative risk [RR], 0.69; 95% confidence interval [CI], 0.55–0.88), moderate or severe cerebral palsy (RR, 0.64; 95% CI, 0.44–0.92), and substantial gross motor dysfunction (RR, 0.60; 95% CI, 0.43–0.83). There was no overall difference in the risk of total pediatric mortality (RR, 1.01; 95% CI, 0.89–1.14). Minor side effects were more frequent among women receiving magnesium sulfate. In conclusion, magnesium sulfate administered to women at risk of delivery before 34 weeks of gestation reduces the risk of cerebral palsy.

Keywords

Magnesium sulfate; cerebral palsy; pediatric mortality; meta-analysis; systematic review; prematurity; preterm birth; neurologic handicap; infant development; preterm birth

INTRODUCTION

Cerebral palsy describes a group of disorders affecting the development of movement and posture, causing activity limitation, and which are attributed to non-progressive disturbances. Insults responsible for cerebral palsy are believed to have occurred during fetal development or infancy.¹ Cerebral palsy is the most prevalent chronic childhood motor disability with an estimated lifetime cost in 2003 of nearly \$1 million per person.² The prevalence of cerebral palsy reported in recent population-based studies ranges between 1.5 and 3.6 cases per 1000 live births.^{3–5} The United Cerebral Palsy Foundation estimates that

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nearly 800,000 children and adults of all ages in the United States have cerebral palsy.⁶ Secular trends in overall prevalence of cerebral palsy for over the last 40 years shows a modest increase in the frequency which has been attributed to a substantial increase in cerebral palsy in very low birth weight infants, which, in turn, is attributable to their increased survival resulting from improvements in neonatal intensive care.^{3, 5}

Preterm birth is a major risk factor for cerebral palsy, and the risk increases markedly with decreasing gestational age.⁷ Currently, infants born at less than 34 weeks of gestation constitute about 25% of all new cases of cerebral palsy.^{5, 8} Multiple pregnancy is associated with an increased risk of cerebral palsy. This has been attributed, at least in part, to preterm birth. However, fetal or neonatal death of a member of a multiple gestation, twin-to-twin transfusion syndrome, or intrapartum problems also contribute.^{9, 10}

Several observational studies have reported an association of antenatal treatment with magnesium sulfate for preterm labor or preeclampsia with a decreased risk of cerebral palsy in low birth weight or preterm infants. In 1995, Nelson and Grether¹¹ reported a case-control study investigating whether in utero exposure to magnesium sulfate were used to prevent convulsions in preeclampsia or if a tocolytic agent was associated with a lower prevalence of cerebral palsy in infants born weighing <1500 g. Children with cerebral palsy were less likely to have been exposed to magnesium sulfate than were control subjects (odds ratio, 0.14 [95% confidence interval, 0.05 to 0.51]) suggesting a protective effect of magnesium sulfate against cerebral palsy in these very low birth weight infants. Although some observational studies have reported similar findings,¹²⁻¹⁵ others have reported no association between administration of magnesium sulfate and the subsequent risk of cerebral palsy.¹⁶⁻²¹

In response to the conflicting evidence from observational studies, several randomized controlled trials of magnesium sulfate administered to mothers for fetal neuroprotection have been performed. A recent systematic review that assessed the effects of magnesium sulfate as a fetal neuroprotective agent when given to women considered at risk of preterm birth concluded that this therapy reduced the risk of cerebral palsy and substantial gross motor dysfunction in early childhood.²² This review, however, based its main conclusions on meta-analysis that included preterm infants less than 37 weeks of gestational age and did not perform any economic evaluation for estimating the cost-effectiveness of this intervention for the prevention of cerebral palsy.

We carried out a systematic review and meta-analysis of all available randomized controlled trials to determine the efficacy and safety of antenatal administration of magnesium sulfate to women at risk of preterm delivery before 34 weeks of gestational age for the prevention of cerebral palsy in their children.

MATERIALS AND METHODS

The systematic review was conducted following a prospectively prepared protocol and reported using the Quality of Reporting of Meta-analysis (QUOROM) guidelines for meta-analysis of randomized controlled trials.²³

Search

We searched PubMed, Embase, Cinahl, and Lilacs (all from inception to March 30, 2009), ISI Web of Science (<http://www.isiknowledge.com>) (1960 to March 30, 2009), the Cochrane Central Register of Controlled Trials (http://www.mrw.interscience.wiley.com/cochrane/cochrane_clcentral_articles_fs.html) (1960 to March 30, 2009), and Research Registers of ongoing trials (www.clinicaltrials.gov, www.controlled-trials.com, www.centerwatch.com,

www.actr.org.au, www.nrr.nhs.uk, and www.umin.ac.jp/ctr) using a combination of keywords and text words related to *magnesium*, *cerebral palsy*, and *neuroprotection*. Proceedings of the Society for Maternal-Fetal Medicine and international meetings on cerebral palsy, reference lists of identified studies, textbooks, previously published systematic reviews, and review articles were also searched. No language restrictions were applied. All searches were carried independently by the two authors and results were merged. For studies that resulted in multiple publications, the data from the publication with the largest sample size were used and supplemented if additional information appeared in the other publications.

Study selection

We included randomized controlled trials comparing magnesium sulfate with placebo or no magnesium sulfate for women at risk of preterm birth before 34 weeks of gestation, whose primary aim was to prevent cerebral palsy and other neurologic abnormalities in the unborn baby or if the primary aim was otherwise but data on cerebral palsy were reported for the infants. Quasi-randomized studies were excluded. We classified trials according to aim of the treatment with magnesium sulfate into two groups: “neuroprotection of the fetus” and “other aims”.

All published studies deemed suitable were retrieved and reviewed independently by the two authors to determine inclusion. Disagreements were resolved through consensus.

Outcome measures

The primary outcomes of interest were cerebral palsy and total pediatric mortality (fetal death + any mortality of live births that occurred during the two years of corrected age following the birth). We included total pediatric mortality because it is a competing outcome that would preclude the assessment of cerebral palsy. Pre-specified secondary outcomes for the neonate and infant included mild cerebral palsy, moderate or severe cerebral palsy, grade III or IV intraventricular hemorrhage, periventricular leukomalacia, Apgar score <7 at 5 minutes, neonatal seizures, respiratory distress syndrome, need for supplemental oxygen at 36 weeks, bronchopulmonary dysplasia, mechanical ventilation, necrotizing enterocolitis, substantial gross motor dysfunction, major neurologic disability, any neurological impairment, Bayley mental development index <70 and <85, Bayley psychomotor development index <70 and <85, blindness, and deafness. For the mother, secondary outcomes were death, cardiac or respiratory arrest, pulmonary edema, respiratory depression, hypotension, tachycardia, severe postpartum hemorrhage, cesarean delivery, and clinical and self-assessed maternal side effects of the infusion such as flushing, nausea or vomiting, sweating, problems at injection site, stopping of infusion because of adverse effects adverse effect, and any side effect.

Study quality assessment

We assessed study methodological quality using a modified scoring system proposed by Jadad et al.²⁴ which is based on three items: randomization, blinding, and follow-up. Points were awarded on the basis of the quality of randomization (2 points: computer-generated random numbers or similar; 1 point: not described; 0 points: quasi-randomized or not randomized [we excluded such studies]); double blinding (2 points: neither the person doing the assessments nor the study participant could identify the intervention being assessed; 1 point: not described; 0 points: no blinding or inadequate method), and follow-up (2 points: number or reasons for dropouts and withdrawals described, and assessment of primary outcomes in ≥95% of randomized fetuses; 1 point: number or reasons for dropouts and withdrawals described but assessment of primary outcomes in <95% of randomized fetuses; 0 points: number or reasons for dropouts and withdrawals not described). In addition, we

assessed concealment of allocation as follows: 2 points: adequate method (central randomization; or drug containers or opaque, sealed envelopes that were sequentially numbered and opened sequentially only after they have been irreversibly assigned to the participant); 0 points: no concealment of allocation or inadequate method or not described. Thus, the total score ranged from 0 (lowest quality) to 8 (highest quality). The methodological quality of included trials was assessed individually by the two reviewers who were not associated with any of the trials. When differences in scoring existed, a consensus was reached.

Data abstraction

One reviewer (AC-A) scanned abstracts and titles. Potentially relevant articles were acquired and data were extracted in duplicate from all reports and recorded on a piloted form independently by the two reviewers. There was no blinding by authorship. All outcome data were further verified with the original articles. Information was extracted on study characteristics (randomization procedure, blind assessment at baseline and follow up, follow up period, intention to treat analysis, and losses to follow up), participants (inclusion and exclusion criteria, numbers of mothers and infants in randomized groups, baseline characteristics, and country and date of recruitment), details of intervention (aim, loading dose, maintenance dose, median total dose, duration, and retreatment), and outcomes (number of outcome events including mortality and morbidity). The number of infants was used as the denominator for primary (cerebral palsy and pediatric mortality), neonatal, and infant neurodevelopmental outcomes. When maternal outcomes were presented, numerators and denominators were calculated based on the number of mothers. In an attempt to obtain additional data, we contacted three authors by e-mail of whom one responded. Disagreements in extracted data were resolved by discussion among reviewers.

Statistical analysis

Statistical analysis was done according to the guidelines of the Cochrane Collaboration.²⁵ We analyzed outcomes on an intend-to-treat basis. If this was not clear from the original article, then we carried out re-analysis where possible. If data for similar outcomes from two or more separate studies were available, we combined the data in a meta-analysis and calculated a summary relative risk (RR) with associated 95% confidence interval (CI). Heterogeneity of the results among studies was tested with the quantity I^2 , which describes the percentage of total variation across studies that is due to heterogeneity rather than chance.²⁶ A value of 0% indicates no observed heterogeneity whereas I^2 values of 50% or more indicate a substantial level of heterogeneity.²⁶ We planned to pool data across studies using a fixed-effects model if substantial statistical heterogeneity was not present. We used random-effects models to pool data across studies if the I^2 values were $\geq 50\%$.

We conducted sensitivity analyses to explore the robustness of findings for the primary outcomes according to statistical model (fixed-effects vs. random-effects), modified Jadad quality score (>4 vs ≤ 4), and completeness of follow-up of randomized fetuses ($\geq 95\%$ vs $<95\%$). Further subgroup analyses were planned to assess the primary outcomes by primary aim of the treatment with magnesium sulfate (neuroprotective vs other aims), median total dose of magnesium sulfate used (≤ 4 grams vs >4 grams), gestational age at trial entry (<34 , <32 , and <30 weeks), and plurality (singleton and multiple pregnancy). The meta-analysis by plurality of pregnancy, however, was not possible because there were not sufficient data available from the great majority of studies.

We assessed publication and related biases visually by examining the symmetry of funnel plots and statistically by using the Egger test.²⁷ The larger the deviation of the intercept of the regression line from zero, the greater was the asymmetry and the more likely it was that

the meta-analysis would yield biased estimates of effect. As suggested by Egger, we considered $P < 0.1$ to indicate significant asymmetry.

We also calculated the number needed to treat (NNT) for an additional beneficial outcome and the NNT for an additional harmful outcome with their 95% CIs for outcomes in which the treatment effect was significant at the 5% level (the 95% CI for the absolute risk difference did not include zero).²⁸ NNT was computed from the results of meta-analysis of relative risks as follows:

$$\text{NNT} = \left| \frac{1}{\text{Control group event rate} \times (1 - \text{relative risk})} \right|$$

In this review, NNT for an additional beneficial outcome is the number of women at risk of preterm delivery before 34 weeks of gestation who need to be treated with magnesium sulfate rather than with placebo to prevent one case of cerebral palsy. The NNT for an additional harmful outcome is the number of women at risk of preterm delivery before 34 weeks of gestation who need to be treated with magnesium sulfate rather than with placebo for one additional woman to be harmed by an adverse event.

We estimated the hypothetical impact of universal use of magnesium sulfate in American women at high risk of preterm delivery before 34 weeks of gestational age to prevent cerebral palsy in their children. Initially, we calculated the total number of new cases of cerebral palsy diagnosed each year in the US among infants born before 34 weeks of gestational age using published data on the total number of new cases of cerebral palsy recognized each year in the US and the percentage of infants with cerebral palsy born before 34 weeks of gestational age. Then, the total number of new cases of cerebral palsy among infants born before 34 weeks of gestational age was multiplied by the summary RR and the upper and lower bounds of its 95% CI obtained in our meta-analysis to estimate the hypothetical number of new cases of cerebral palsy with corresponding 95% CI that could be prevented annually using this intervention.

Finally, we performed an economic evaluation to calculate the incremental cost of preventing one case of cerebral palsy in the US through use of antenatal magnesium sulfate in women at risk of preterm delivery before 34 weeks of gestational age. First, we searched data published recently on the total cost per patient of administering magnesium sulfate as a tocolytic agent in the US. Then, the incremental cost of preventing one case of cerebral palsy (with corresponding 95% CI) by using antenatal magnesium sulfate was estimated by multiplying the total cost per patient of administering magnesium sulfate by the NNT for benefit and the upper and lower bounds of its 95% CI obtained in our analysis. All statistical analyses were performed with the StatsDirect version 2.7.2 (StatsDirect Ltd, United Kingdom).

RESULTS

The flow of the search is shown in Figure 1. Of the 331 potentially relevant citations identified, 5 studies (6 trials) published in 7 articles fulfilled the inclusion criteria after a detailed review of 92 studies.^{29–35} Because the study by Mittendorf et al^{29, 30} had two arms (tocolytic and neuroprotective), it was considered as two separate trials in this review. Of the other four studies included, three^{31, 33, 35} evaluated magnesium sulfate as an infant neuroprotective agent, and one³² assessed the efficacy of magnesium sulfate for preventing eclampsia. Eighty-seven studies were excluded, the main reason being the lack of data on cerebral palsy and/or pediatric mortality in randomized controlled trials on magnesium

sulfate for preventing preterm birth or eclampsia. Overall agreement on the inclusion of studies was 100% (κ , 1.00). The 6 trials included a total of 4796 women and 5357 infants.

The characteristics of included studies are presented in Table 1. Two studies were performed in the United States,^{29, 35} one each in France,³³ Australia and New Zealand,³¹ and the remaining one was conducted in 19 countries across five continents.³² Trials included women with gestational ages <34 weeks²⁹ (165 infants), <33 weeks³³ (688 women), <32 weeks³⁵ (2444 infants), and <30 weeks³¹ (1255 infants). The Magpie trial³² reported data for 3283 infants whose mothers were randomized before delivery. Of these, 805 infants were randomized before 34 weeks of gestational age and 788 between 34 and 36 weeks. Data for the 805 infants randomized before 34 weeks of gestational age included in our meta-analyses were extracted from the review by Doyle et al.²² Multiple pregnancies were included in all trials. Two trials^{31, 33} included women at high risk of preterm delivery because of expected or planned birth within 24 hours. The trial by Rouse et al³⁵ included women at high risk for spontaneous delivery because of premature rupture of membranes or advanced preterm labor and women with indicated preterm delivery within 2 to 24 hours. The main reasons for preterm birth were preterm labor and premature rupture of the membranes (63 and 9% in the study by Crowther et al,³¹ 85 and 61% in the study by Marret et al,³³ and 10 and 87% in the study by Rouse et al,³⁵ respectively), followed by antepartum hemorrhage and chorioamnionitis (14 and 14% in the study by Crowther et al,³¹ and 19 and 11% in the study by Marret et al,³³ respectively). In the study by Crowther et al,³¹ 15% of women enrolled had preeclampsia. One study only included women in active preterm labor with cervical dilation \geq 4 cm (tocolytic arm) or $>$ 4 cm (neuroprotective arm).²⁹ The Magpie trial³² included only women with preeclampsia. Three trials excluded women with preeclampsia.^{29, 33, 35} The definition and diagnostic criteria of cerebral palsy were clearly reported in four trials.^{31, 32, 34, 35} The study by Mittendorf et al³⁰ did not report a definition or details of the diagnosis. A pediatrician made the diagnosis of cerebral palsy in all trials at a corrected age of at least 18 months^{30, 32} or 24 months.^{31, 34, 35}

The studies included a range of dosing schedules. In five trials,^{29, 31–33} women received a loading dose of 4 g of intravenous magnesium sulfate. In the trial by Rouse et al³⁵ women received a loading infusion of 6 g. There was no maintenance infusion of magnesium sulfate in two trials (neuroprotective arm of the study by Mittendorf et al²⁹ and the study by Marret et al³³). The maintenance doses of magnesium sulfate used in four trials were 1 g/h,³¹ 2 g/h,³⁵ 2–3 g/h (tocolytic arm of the study by Mittendorf et al. study²⁹), and either 1 g/h intravenously or 5 g every four hours intramuscularly.³² The maintenance infusion was continued until birth (if it occurred within 24 hours) or up to 24 hours in the study by Crowther et al,³¹ or until 24 hours in the Magpie trial.³²

In the study by Rouse et al,³⁵ if delivery had not occurred after 12 hours and was no longer considered imminent, the infusion was discontinued and resumed when delivery was deemed imminent again. If at least 6 hours had passed since the discontinuation of the study medication, another loading dose was given. Retreatment with magnesium sulfate was not allowed in three trials^{31–33} and in the neuroprotective arm of the study by Mittendorf et al.³⁰ There was no information on duration of maintenance infusion and whether retreatment was permitted in the tocolytic arm of the study by Mittendorf et al.³⁰ The median total dose of magnesium sulfate received by women in the magnesium group ranged between 4 and 49.8 g.

The quality assessment of trials is shown in Table 2. The overall quality of the trials was relatively good. All studies were randomized with the appropriate method to generate the sequence of randomization (computer-generated). Five trials, including the neuroprotective arm of the study by Mittendorf et al,³⁰ were double-blinded in which a placebo was used,

and the outcomes were evaluated by individuals who were blinded to treatment group allocation. The tocolytic arm of the study by Mittendorf et al,³⁰ was not double-blinded because the women were randomized to receive magnesium sulfate or other tocolytic therapy such as ritodrine, terbutaline, indomethacin, or nifedipine, although diagnosis of cerebral palsy was made by a pediatrician who was masked to the antenatal treatment. Four trials had adequate methods of allocation sequence concealment.^{31, 32, 34, 35} The tocolytic arm of the study by Mittendorf et al³⁰ had no concealment of allocation whereas the neuroprotective arm did not report on this item. Information on withdrawals and dropouts was available for four trials.^{31, 32, 34, 35} In the two trials performed by Mittendorf et al³⁰ there was no statement on numbers and reasons for withdrawals in each group. Three trials reported assessment of primary outcomes in 95% of randomized fetuses.^{31, 34, 35} The Magpie trial³² evaluated the primary outcomes in only 47.4% of fetuses of all gestational ages randomized before delivery. Four trials had a modified Jadad score of 7 or more.^{31, 32, 34, 35} The two trials performed by Mittendorf et al³⁰ had scores of 2 (tocolytic arm) and 4 (neuroprotective arm).

Primary outcomes

Six trials (5357 infants) reported cerebral palsy, total pediatric mortality, and the combined outcome of death or cerebral palsy. The risk of giving birth to an infant subsequently receiving a diagnosis of cerebral palsy was significantly lower in the group of women who received magnesium sulfate than among women who did not receive magnesium sulfate (3.9% vs 5.6%; RR, 0.69 [95% CI, 0.55–0.88]) (Figure 2) (Table 3). There was evidence of low statistical heterogeneity among the trials included in this analysis ($I^2=4.4\%$). The funnel plot of trials of antenatal magnesium sulfate in the prevention of cerebral palsy appeared symmetrical either visually or in terms of statistical significance (intercept, 0.23 [95% CI, -1.97 to 2.43]; $P = 0.79$) (Figure 3) suggesting that there was no evidence of either publication or related biases.

The number of women at risk for preterm delivery less than 34 weeks of gestation who needed to be treated with magnesium sulfate rather than with placebo to prevent one case of cerebral palsy in their children was 52 (95% CI, 31–154). The random-effects analysis of the primary outcome of cerebral palsy yielded effect sizes similar in magnitude and direction to those obtained from the fixed-effects analysis (Table 4).

Antenatal magnesium sulfate significantly decreased the risk of cerebral palsy in sensitivity analyses limited to the 4 trials with modified Jadad quality score >4 (RR, 0.69 [95% CI, 0.54–0.88]; $I^2=0.0\%$) and to the 3 trials with completeness of follow-up of randomized fetuses 95% (RR, 0.69 [95% CI, 0.54–0.88]; $I^2=0.0\%$). The RR for the subgroup of 4 trials (4,446 infants) whose primary aim was neuroprotection was 0.71 (95% CI, [0.55–0.91]; $I^2=25.2\%$) and for the subgroup of 2 trials (911 infants) whose primary aim was tocolysis or to prevent eclampsia was 0.37 (95% CI, [0.09–1.58]; $I^2=0.0\%$). The lowered risk of cerebral palsy was demonstrated even in the subgroup of 3 trials (3981 infants) that included women with gestational age <32 weeks at trial entry and in the subgroup of 4 trials (4610 infants) that used a median total dose of magnesium sulfate greater than 4 grams. No statistically significant differences between groups treated with antenatal magnesium sulfate and controls in cerebral palsy were seen in the subgroups of 2 trials that used a median total dose of magnesium sulfate 4 grams and that included women with gestational age <30 weeks at trial entry.

Three trials (4,387 infants) reported cerebral palsy according to severity.^{31, 34, 35} Moderate or severe cerebral palsy was significantly reduced in the group who received magnesium sulfate (2.1% vs 3.2%; RR, 0.64 [95% CI, 0.44–0.92]; $I^2=0.0\%$; NNT for benefit, 74 [95% CI, 41–373]). The risk of mild cerebral palsy was reduced by 26% in the group allocated

magnesium sulfate rather than placebo, although this did not achieve statistical significance (RR, 0.74 [95% CI, 0.52–1.04]).

There was no overall difference in the risk of total pediatric mortality, including fetal mortality and under two years of corrected age mortality, between infants exposed to magnesium sulfate and those non exposed (15.1% versus 14.8%; RR, 1.01 [95% CI, 0.89–1.14; $I^2 = 38.9\%$) (Table 3). The lack of evidence for any overall effect of magnesium sulfate on total pediatric mortality was found in all sensitivity and subgroup analyses performed except one (subgroup of two trials by Mittendorf et al³⁰ with modified Jadad quality score = 4 (RR, 6.61; 95% CI, 1.12–36.11; $I^2=31.0\%$) (Table 4). The combined outcome of death or cerebral palsy was slightly lower for children in the magnesium sulfate group (19.0% versus 20.4%), although this was not a statistically significant difference.

Secondary outcomes

There were no significant differences between the groups in the risk of adverse neonatal outcomes (Table 5), although a non-significant increase was seen in the risk of necrotizing enterocolitis in the magnesium group compared with the control group (7.1% vs 5.9%; RR, 1.23 [95% CI, 0.98–1.54]; $I^2=0.0\%$)

With regard to infant neurodevelopmental outcomes, the risk of substantial gross motor dysfunction at the corrected age of 2 years was significantly lower among children whose mothers received magnesium sulfate than among children whose mothers did not receive magnesium sulfate (3 trials; 4387 infants; 2.6% vs 4.2%; RR, 0.60 [95% CI, 0.43–0.83]; $I^2=0.0\%$; NNT for benefit, 53 [95% CI, 32–146]). There were no significant effects of antenatal magnesium sulfate treatment on other infant neurodevelopmental outcomes such as major neurologic disability, any neurological impairment, Bayley mental and psychomotor development indexes, blindness, and deafness (Table 6).

There was no evidence of an effect of magnesium sulfate on the risk of maternal death (3 trials; 3,867 women), cardiac or respiratory arrest (3 trials; 3867 women), pulmonary edema (1 trial; 2,241 women), respiratory depression (2 trials; 3,303 women), severe postpartum hemorrhage (2 trials; 1626 women), and cesarean section (3 trials; 3,867 women) (Table 7). Compared with women receiving placebo, those exposed to magnesium sulfate had about a 50% increased risk of both hypotension and tachycardia (RR, 1.51 [95% CI, 1.09–2.09]; NNT for harm, 30 [95% CI, 17–156] and RR, 1.53 [95% CI, 1.03–2.29; NNT for harm, 28 [95% CI, 14–379], respectively). Maternal side effects secondary to study medication were significantly more common among women allocated magnesium sulfate rather than placebo, including flushing (58.4% vs 8.3%; NNT for harm, 2 [95% CI, 2–2]), nausea or vomiting (16.3% vs 3.9%; NNT for harm, 8 [95% CI, 7–10]), sweating (25.2% vs 3.4%; NNT for harm, 5 [95% CI, 4–5]), problems at injection site (37.6% vs 4.1%; NNT for harm, 3 [95% CI, 3–3]), stopping of infusion because of adverse effects (7.5% vs 2.6%; NNT for harm, 20 [95% CI, 16–29]), and any side effect (70.7% vs 17.6%; NNT for harm, 2 [95% CI, 2–2]). All funnel plots showed no asymmetry, either visually or in terms of statistical significance ($P>.10$ for all, by Egger test).

Impact and economic evaluation of the intervention

The United Cerebral Palsy Foundation has estimated that about 8,000 babies and infants are diagnosed with cerebral palsy each year in the US⁶ of which about 25% (n=2,000) are born before 34 weeks of gestational age.^{5, 8} Therefore, if all women who deliver before 34 weeks of gestation receive antenatal magnesium sulfate, the estimated number of new cases of cerebral palsy that hypothetically could be prevented annually is 620 (95% CI, 240–900).

The number of women at risk of preterm delivery before 34 weeks of gestation who needed to receive magnesium sulfate to prevent one case of cerebral palsy was 52 (95% CI, 31–154). A recent cost decision analysis from the US revealed that the total cost for administering magnesium sulfate as a tocolytic agent, including costs attributable to the evaluation and treatment of its adverse events, was \$197.90 per patient in 2005.³⁶ Thus, if magnesium sulfate was given to all women at risk of preterm delivery before 34 weeks of gestation, the incremental cost of preventing one case of cerebral palsy would be \$10,291 (95% CI, 6,135–30,477).

COMMENT

In this systematic review, we found persuasive evidence that magnesium sulfate administered to women at high risk of delivery before 34 weeks of gestation reduces the risk of cerebral palsy in their children. The evidence was strongest for the subgroup of trials that specifically evaluated the use of antenatal magnesium sulfate in preventing cerebral palsy. In addition, this therapy was also associated with a significantly decreased risk of moderate or severe cerebral palsy and substantial gross motor dysfunction without any statistically significant effect on the risk of total pediatric mortality. This last finding suggests the reduced risk for cerebral palsy does not appear to be due to selective mortality of magnesium sulfate-exposed infants. Overall, magnesium sulfate therapy had no significant effect on the risk of major maternal complications, adverse neonatal outcomes, and other infant neurodevelopmental outcomes, although a trend was found towards increased risk of necrotizing enterocolitis. About 70% of women reported any minor side effect associated with the infusion of magnesium sulfate mainly flushing, problems at injection site, and sweating.

The strength of our findings is based on its compliance with stringent criteria for performing a rigorous systematic review of randomized controlled trials, the inclusion of a relatively large number of women and infants in these meta-analyses, the relatively narrow confidence intervals obtained which made our results more precise, the evidence of clinical and statistical homogeneity in the results of the trials for cerebral palsy, the sensitivity analysis restricted to high quality trials that upheld our main results, and the symmetrical funnel plots that suggested absence of publication and related biases in our meta-analyses.

We were unable to determine whether antenatal magnesium sulfate is more effective to prevent cerebral palsy in infants of singleton pregnancies than in infants of multiple pregnancies. Specific data for the use of magnesium sulfate according to plurality of pregnancy were reported only by two authors.^{31, 35} Crowther et al³¹ found that magnesium sulfate had no a significant effect on the risk of cerebral palsy in both infants of singleton pregnancies (RR, 1.01 [95% CI, 0.61–1.68]) and infants of multiple gestations (RR, 0.52 [95% CI, 0.21–1.25]). Rouse et al³⁵ reported that magnesium sulfate reduced the risk of moderate or severe cerebral palsy in infants of singleton pregnancies (RR, 0.52 [95% CI, 0.27–0.98]) but not in infants of twin pregnancy (RR, 0.73; 95% CI, 0.27–1.92).

A recent meta-analysis²² reported similar findings to those found in our review. In fact, antenatal magnesium sulfate given to women at risk of preterm birth before 37 weeks of gestational age reduced the risk of cerebral palsy (RR, 0.68 [95% CI, 0.54–0.87]) and substantial gross motor dysfunction (RR, 0.61 [95% CI, 0.44–0.85]) in their child without a statistically significant effect on pediatric mortality (RR, 1.04 [95% CI, 0.92–1.17]). Notwithstanding, our findings need to be distinguished from those of this meta-analysis. We specifically evaluated the effect of antenatal magnesium sulfate in preventing cerebral palsy in preterm infants whose mothers were less than 34 weeks of gestational age at randomization. The meta-analyses by Doyle et al²² included 5,357 preterm infants whose

mothers were less than 34 weeks of gestation at randomization plus 788 infants from the Magpie trial³² whose mothers were between 34 0/7 and 36 6/7 weeks of gestation when treated with magnesium sulfate or placebo. In this last subgroup of preterm infants, cerebral palsy was diagnosed in 0 of 394 infants in the magnesium sulfate group (0.0%) and in 2 of 394 infants in the placebo group (RR, 0.20 [0.01–4.15]). Thus, there is no evidence that antenatal magnesium sulfate administered to women at risk of preterm delivery between 34 0/7 and 36 6/7 weeks of gestation decreases the risk of cerebral palsy in their infants. In addition, the Magpie trial was not primarily designed to evaluate the neuroprotective role of magnesium sulfate and had a low rate of follow-up. Finally, the systematic review by Doyle et al.²² did not perform an economic evaluation to assess the cost-effectiveness of this intervention for the prevention of cerebral palsy or evaluate some important neonatal and maternal outcomes such as necrotizing enterocolitis, maternal pulmonary edema, and some maternal side effects of the infusion (e.g. flushing, nausea or vomiting, sweating, and problems at injection site).

The molecular mechanisms by which magnesium sulfate can prevent cerebral palsy in preterm infants are not fully understood. The most prevalent pathological lesion seen in cerebral palsy is periventricular white matter injury resulting from vulnerability of the immature preoligodendrocytes before 32 weeks of gestation.³⁷ Preoligodendrocytes are the precursors of myelinating oligodendrocytes, which constitute a major glial population in the white matter. Oxidative stress and excitotoxicity resulting from excessive stimulation of ionotropic glutamate receptors on preoligodendrocytes are the most prominent molecular mechanisms for periventricular white matter injury.^{38, 39} In addition, recent studies have identified functional N-methyl-D-aspartate (NMDA) glutamatergic receptors in oligodendroglial injury processes.^{40, 41} Antagonists of the NMDA receptors for glutamate are potent neuroprotective agents in several animal models of perinatal brain lesions.⁴² Growing evidence suggests that magnesium sulfate may reverse the harmful effects of hypoxic/ischemic brain damage by blocking the NMDA receptors and, as a calcium antagonist, hindering calcium influx into the cells.^{43, 44} In addition, magnesium sulfate could also reverse the destructive action of oxygen radicals and excitatory amino acids.^{45, 46}

The finding that antenatal magnesium sulfate significantly reduces cerebral palsy compared with placebo is a strong argument for its use in women at risk of preterm delivery before 34 weeks of gestation. Some limitations must be noted, however, when interpreting the results of this review. First, there were variations in the inclusion and exclusion criteria, gestational age at which magnesium sulfate was administered, loading and maintenance doses, duration of the intervention, and use of retreatment. In addition, the optimal time to administer magnesium sulfate was unclear. Second, two of the studies included in our review^{30, 32} made the assessment of cerebral palsy at 18 months of corrected age, and other³⁴ was made by parent telephone interview in 22% of surviving infants. For reasons yet poorly understood, many children, especially those born very preterm, exhibit a variety of neurologic abnormalities in the first 18 months of life that suggest the diagnosis of cerebral palsy which then can resolve later.⁴⁷ Therefore, it has been suggested that the diagnosis of cerebral palsy should be assigned cautiously before the age of 24 months (or 24 months after the expected date of delivery in the case of preterm infants), unless the disorder is exceptionally severe.⁴⁸ Third, not all of the trials reported included important secondary outcomes such as adverse neonatal and maternal outcomes, as well as infant neurodevelopmental outcomes. Finally, the results of the studies we summarized are directly applicable only to the women and infants included in those studies. Whether the results are applicable to women and infants with other clinical conditions is unknown.

The antenatal administration of magnesium sulfate to women at high risk of delivery before 34 weeks of gestation could be a cost-effective intervention. In 2003, Honeycutt et al²

estimated the average lifetime costs (direct and indirect) per person were \$921,000 for persons with cerebral palsy. We have estimated that if magnesium sulfate is administered to all women at high risk of preterm delivery before 34 weeks of gestation, the incremental cost of preventing one case of cerebral palsy would be \$10,291 (95% CI, \$ 6,135 to \$ 29,685). It should be noted that we did not estimate the cost-effectiveness of this intervention in terms of the marginal cost per quality-adjusted life year (QALY) gained.

Antenatal magnesium sulfate should be considered for use in women at high risk of delivery before 34 weeks of gestation, mainly in those with premature rupture of membranes, labor in active phase, and planned delivery within 24 hours. Loading and maintenance doses, and the duration of the treatment should not normally exceed 6 g, 1–2 g/h, and 24 hours, respectively. Further studies and/or meta-analysis of individual patient data from the available trials are required to assess the minimum effective dose of magnesium sulfate, the optimal time to administer it, the need for retreatment, the efficacy of intervention in singleton and multiple pregnancies, the cost-effectiveness of intervention, the effect on the risk of necrotizing enterocolitis, and the long term consequences of exposure for the women and their children.

In summary, our results suggest that antenatal magnesium sulfate could be used for the primary prevention of cerebral palsy in preterm infants less than 34 weeks of gestational age. However, since cerebral palsy is a result of multiple interacting risk factors rather than of a single cause, it is unlikely that antenatal magnesium sulfate administration alone can prevent all cases of this illness in preterm infants.

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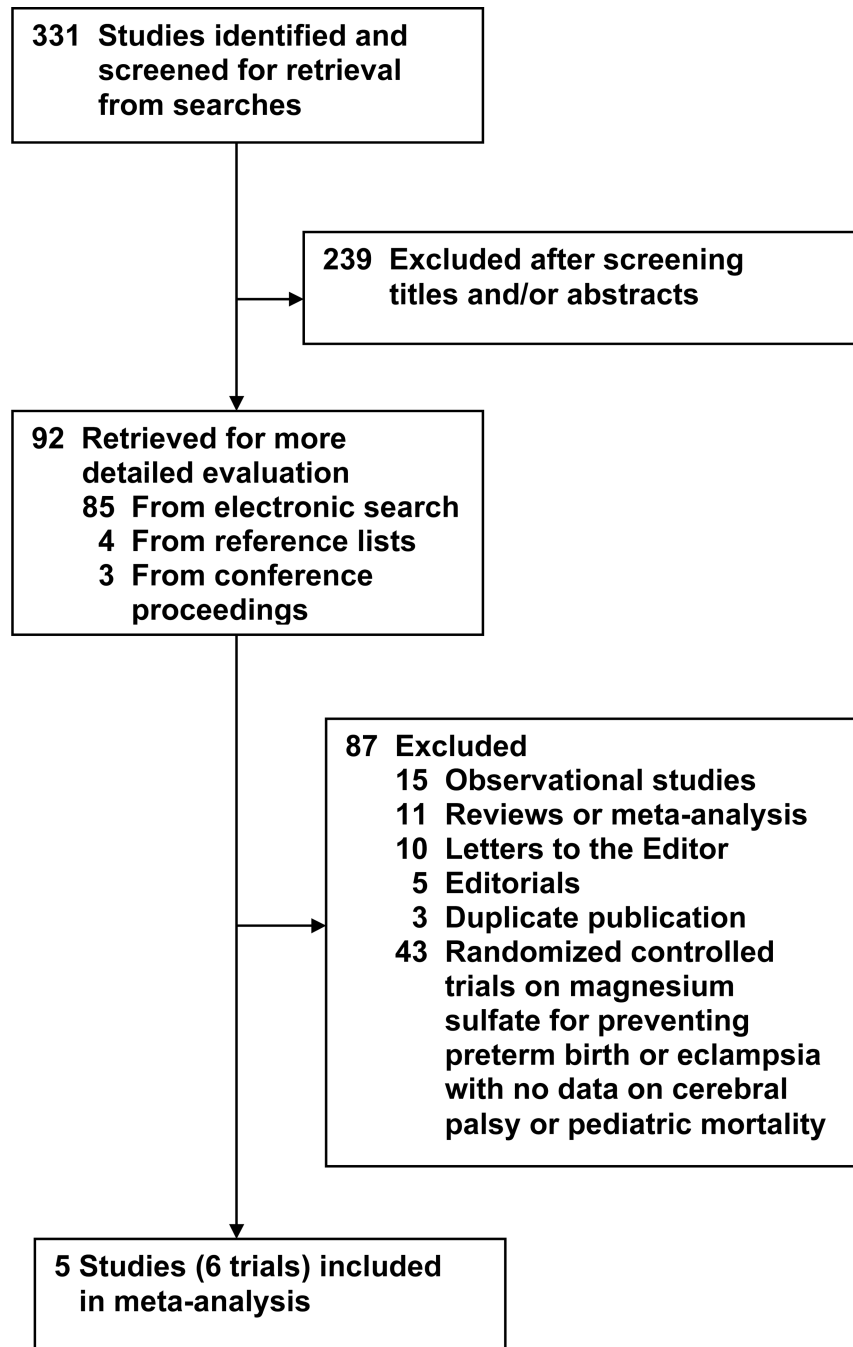


Figure 1.
Study selection process.

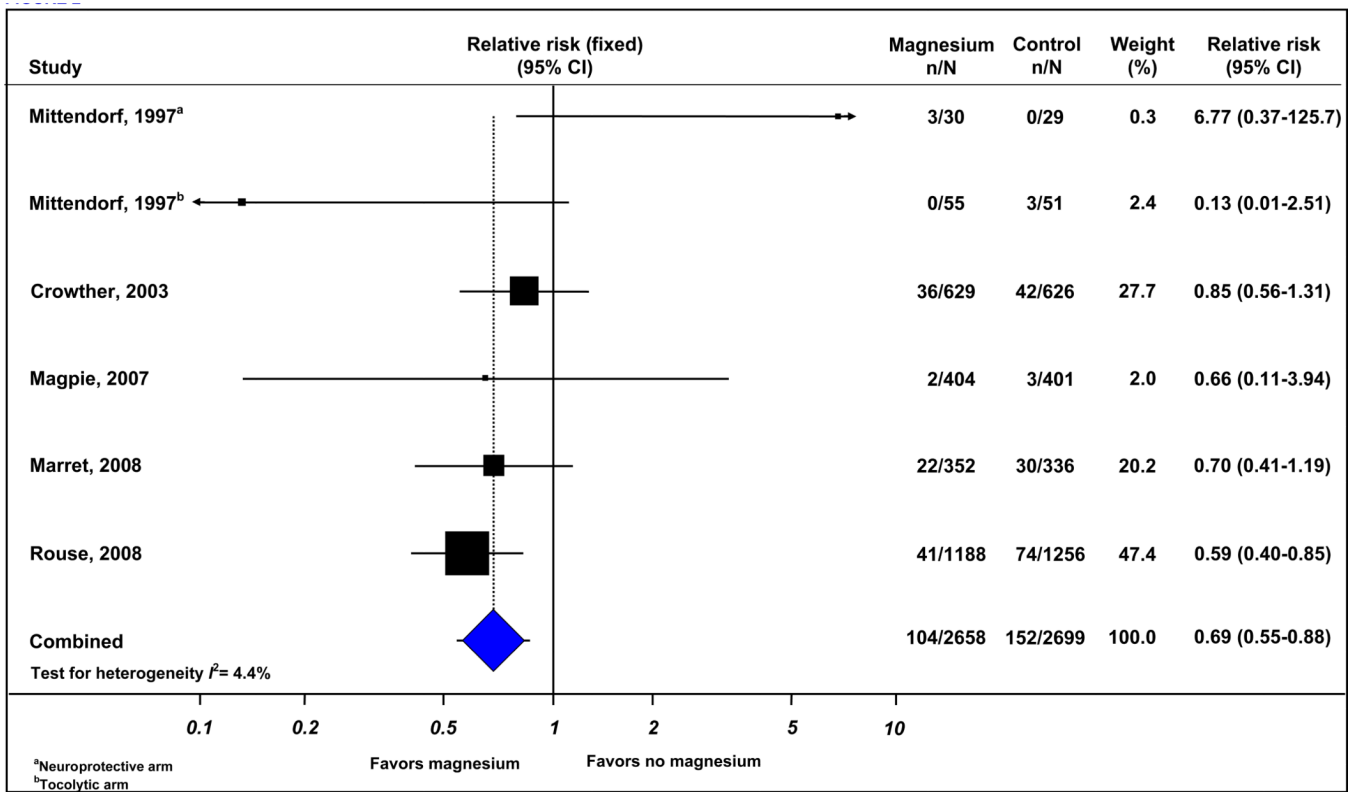


Figure 2.
Effect of magnesium sulfate on cerebral palsy.

Standard error

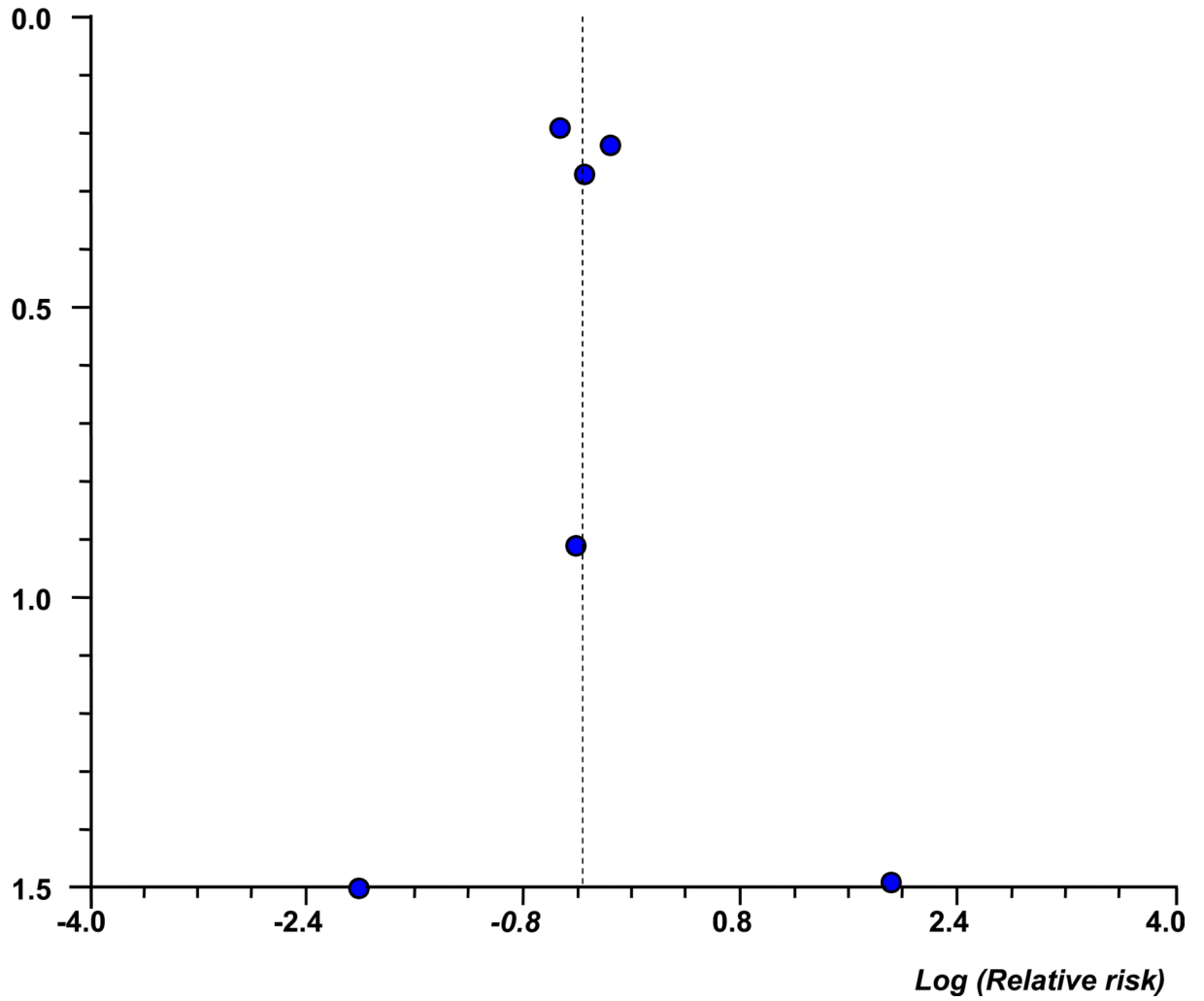


Figure 3.
 Funnel plot of trials of antenatal magnesium sulfate in the prevention of cerebral palsy.
 Circles indicate log (relative risks) from trials included in meta-analysis.

TABLE 1

Characteristics of studies included in the systematic review

Study, year	Location	Inclusion/exclusion criteria	No of infants			Magnesium sulfate		Cerebral palsy	Health professional who made the diagnosis and age at diagnosis
			Magnesium group	No magnesium group	Loading dose (g)	Maintenance dose	Median total dose (g) received by women in magnesium group		
Mittendorf et al. ^{29, 30} 1997	United States, Single center	1. Neuroprotective arm Inclusion: women with single or twin pregnancy in preterm labor between 25 and 33 weeks of gestation with or without premature rupture of the membranes, reassuring fetal assessment, and cervical dilatation >4 cm. Exclusion: preeclampsia, infection	30	29	4	None	4	Not reported	Developmental pediatrician after 18 months of corrected age
		2. Tocolytic arm Inclusion: women with single or twin pregnancy in preterm labor between 25 and 33 weeks of gestation with or without premature rupture of the membranes, reassuring fetal assessment, and cervical dilatation 4 cm. Exclusion: preeclampsia, infection	55	51	4	2–3 g/h; Duration not reported	49.8	Not reported	Developmental pediatrician after 18 months of corrected age
Crowther et al. ³¹ 2003	Australia and New Zealand, 16 centers	Inclusion: women with single, twin, triplet or quadruplet fetuses at risk of preterm delivery before 30 weeks' gestation because of planned or expected birth within 24 hours. Exclusion: women in second stage of labor, if they had received magnesium sulfate in this pregnancy, and contraindications to magnesium sulfate	629	626	4	1 g/h until birth (if occurred within 24 hours) or up to 24 hours	≈10.5	Abnormalities of muscle tone and loss of motor function. Mild cerebral palsy: disability in ambulant children that interfered only slightly with normal daily activities; moderate cerebral palsy: children attempting to walk at two years with or without appliances; Severe cerebral palsy: children permanently non ambulant	Developmental pediatrician and psychologist at 24 months of corrected age
Magpie ³² 2007 ^a	19 countries across five	Inclusion: women with singleton or multiple pregnancy with	404	401	4	Either 1 g/h intravenously	18.0 ^b	Severe cerebral palsy: not walking or unlikely to walk	Pediatrician at 18 months of

Study, year	Location	No of infants			Magnesium sulfate		Cerebral palsy	Health professional who made the diagnosis and age at diagnosis
		Magnesium group	No magnesium group	Loading dose (g)	Maintenance dose	Median total dose (g) received by women in magnesium group		
		<p>Inclusion/exclusion criteria</p> <p>and at high risk for spontaneous delivery because of premature rupture of the membranes, advanced preterm labor with dilatation of 4 to 8 cm and intact membranes, or indicated preterm delivery anticipated within 2 to 24 hours. Exclusion: delivery anticipated within less than 2 hours, cervical dilatation >8 cm, rupture of the membranes before 22 weeks, unwillingness of the obstetrician to intervene for the benefit of the fetus, major fetal anomalies or death, maternal hypertension or preeclampsia, maternal contraindications to magnesium sulfate, and receipt of intravenous magnesium sulfate within the previous 12 hours.</p>						

^aData for children whose mothers were less than 34 weeks of gestation and undelivered at randomization extracted from Doyle et al.'s review²²

^bFor women of all gestational ages and undelivered at randomization

^cIf delivery had not occurred after 12 hours and was no longer considered imminent, the infusion was discontinued and resumed when delivery was deemed imminent again. If at least 6 hours had passed since the discontinuation of the study medication, another loading dose was given

TABLE 2
Modified Jadad score for assessment of methodological quality of included studies

Item	Mittendorf, 1997 ^{29,30}		Crowther, 2003 ³¹	Magpie, 2007 ³²	Marret, 2007 ^{33,34}	Rouse, 2008 ³⁵
	Tocolytic arm	Neuroprotective arm				
Randomization	Yes	Yes	Yes	Yes	Yes	Yes
Method to generate randomization clear and appropriate	Yes	Yes	Yes	Yes	Yes	Yes
Double blind	No	Yes	Yes	Yes	Yes	Yes
Methods for blinding appropriate	No	Yes	Yes	Yes	Yes	Yes
Method of allocation concealment ^a	No	Unreported	Adequate	Adequate	Adequate	Adequate
Description of withdrawal or dropout	No	No	Yes	Yes	Yes	Yes
Completeness of follow-up of randomized fetuses (%) ^b	Unreported	Unreported	98.9	47.4 ^c	98.5	95.6
Total score	2	4	8	7	8	8

Yes=1 point, No=zero points; Scores: 0=lowest quality, 8=highest quality

^a Adequate=2 points, no concealment of allocation or inadequate method or unreported=zero points

^b Follow-up 95%=1 point, follow-up <95% or unreported=zero points

^c For fetuses of all gestational ages and undelivered at randomization

TABLE 3

Effect of magnesium sulfate on cerebral palsy and pediatric mortality

Outcome	No of trials	Number of events/total number		Relative risk (95% CI)	I ² (%)
		Magnesium	No magnesium		
Cerebral palsy	6 ^{30-32, 34, 35}	104/2658	152/2699	0.69 (0.55-0.88)	4.4
Moderate/severe cerebral palsy	3 ^{31, 34, 35}	45/2169	72/2218	0.64 (0.44-0.92)	0.0
Mild cerebral palsy	3 ^{31, 34, 35}	54/2169	74/2218	0.74(0.52-1.04)	0.0
Total pediatric mortality	6 ^{29, 31, 32, 34, 35}	401/2658	400/2699	1.01 (0.89-1.14)	38.9
Fetal mortality	5 ^{29, 31, 34, 35}	17/2254	22/2298	0.78(0.42-1.46)	0.0
Under two years of corrected age mortality	5 ^{29, 31, 34, 35}	217/2254	220/2298	1.00(0.84-1.19)	47.3
Death or cerebral palsy	6 ^{30-32, 34, 35}	505/2658	551/2699	0.92(0.83-1.02)	43.3

CI, confidence interval

TABLE 4

Sensitivity and subgroup palsy and pediatric mortality analyses of metaanalysis on effect of magnesium sulfate on cerebral

Subgroup	No of trials	Relative risk (95% CI)	I ² (%)
CEREBRAL PALSY			
Statistical model			
Fixed effects	6 ^{30-33,35}	0.69 (0.55-0.88)	4.4
Random effects	6 ^{30-33, 35}	0.70 (0.54-0.90)	NA
Modified Jadad quality score			
>4	4 ^{31-33,35}	0.69 (0.54-0.88)	0.0
4	2 ^{30a,b}	0.95(0.02-44.91)	71.1
Completeness of follow-up of randomized fetuses			
95%	3 ^{31,34,35}	0.69 (0.54-0.88)	0.0
<95% or unreported	3 ^{30,32}	0.83 (0.28-2.43)	43.5
Aim of the treatment			
Neuroprotective	4 ^{30a,31,33,35}	0.71 (0.55-0.91)	25.2
Other aims	2 ^{30b, 32}	0.37(0.09-1.58)	0.0
Median total dose of magnesium sulfate			
4 grams	2 ^{30a, 33}	1.37(0.18-10.70)	56.5
>4 grams	4 ^{30b,31,32,35}	0.67(0.51-0.88)	0.0
Gestational age at trial entry			
<34 weeks	6 ^{30-33,35}	0.69 (0.55-0.88)	4.4
<32 weeks	3 ^{31,32,35}	0.69(0.52-0.91)	0.0
<30 weeks	2 ^{31,32}	0.86(0.56-1.31)	0.0
PEDIATRIC MORTALITY			
Statistical model			
Fixed effects	6 ^{29,31-33,35}	1.01 (0.89-1.14)	38.9
Random effects	6 ^{29,31-33,35}	0.99(0.82-1.19)	NA
Modified Jadad quality score			
>4	4 ^{31-33,35}	0.99(0.87-1.11)	27.9
<4	2 ^{29a,b}	6.61 (1.21-36.11)	31.0

Subgroup	No of trials	Relative risk (95% CI)	I ² (%)
Completeness of follow-up of randomized fetuses			
95%	3 ^{31,34,35}	0.94(0.79–1.12)	40.7
<95%	3 ^{29,32}	1.10(0.93–1.30)	49.5
Aim of the treatment			
Neuroprotective	4 ^{29a,31,33,35}	0.95(0.80–1.12)	19.6
Other aims	2 ^{29b,32}	2.83(0.21–38.80)	72.9
Median total dose of magnesium sulfate			
4 grams	2 ^{29a,33}	0.88(0.57–1.35)	0.0
>4 grams	4 ^{29b,31,32,35}	1.01 (0.80–1.28)	59.3
Gestational age at trial entry			
<34 weeks	6 ^{29,31–33,35}	1.01 (0.89–1.14)	38.9
<32 weeks	3 ^{31,32,35}	1.03(0.83–1.28)	65.5
<30 weeks	2 ^{31,32}	0.97(0.67–1.41)	83.0

CI, confidence interval; NA, not applicable

^aNeuroprotective arm of study by Mittendorf et al.^{29, 30}

^bTocolytic arm of study by Mittendorf et al.^{29,30}

TABLE 5

Effect of magnesium sulfate on neonatal outcomes

Outcome	No of trials	Number of events/total number		Relative risk (95% CI)	I ² (%)
		Magnesium	No magnesium		
Intraventricular hemorrhage (all grades)	5 ^{30,31,33,35}	467/2254	493/2298	0.96(0.86–1.08)	20.1
Grade III/IV intraventricular hemorrhage	4 ^{30,31,35}	74/1902	91/1962	0.83(0.61–1.11)	0.0
Periventricular leukomalacia	5 ^{30,31,33,35}	71/2254	76/2298	0.93(0.68–1.28)	0.0
Apgar score <7 at 5 minutes	3 ^{31,33,35}	351/2169	351/2218	1.03(0.90–1.18)	7.3
Neonatal seizures	3 ^{31,33,35}	55/2169	70/2218	0.80(0.56–1.13)	0
Respiratory distress syndrome	2 ^{33,35}	730/1540	779/1592	1.01 (0.85–1.19)	65.8
Need for supplemental oxygen at 36 weeks	2 ^{31,33}	220/981	195/962	1.12(0.95–1.32)	23.1
Bronchopulmonary dysplasia	1 ³⁵	213/1188	218/1256	1.03(0.87–1.23)	NA
Mechanical ventilation	3 ^{31,33,35}	1381/2169	1446/2218	0.99(0.89–1.09)	82.1
Necrotizing enterocolitis	3 ^{31,33,35}	155/2169	131/2218	1.23(0.98–1.54)	0.0

CI, confidence interval; NA, not applicable

TABLE 6

Effect of magnesium sulfate on infant neurodevelopmental outcomes

Outcome	No of trials	Number of events/total number		Relative risk (95% CI)	I ² (%)
		Magnesium	No magnesium		
Substantial gross motor dysfunction	3 ^{31,34,35}	56/2169	94/2218	0.60 (0.43–0.83)	0.0
Major neurologic disability	2 ^{31,32}	93/1033	85/1027	1.09(0.83–1.43)	15.3
Any neurological impairment	2 ^{31,32}	198/1033	194/1027	1.02(0.86–1.20)	0.0
Bayley mental development index <85	3 ^{31,34,35}	639/2169	660/2218	1.00(0.91–1.09)	0.0
Bayley mental development index <70	1 ³⁵	165/1188	171/1256	1.02(0.84–1.24)	NA
Bayley psychomotor development index <85	1 ³⁵	299/1188	315/1256	1.00(0.88–1.15)	NA
Bayley psychomotor development index <70	1 ³⁵	134/1188	144/1256	0.98(0.79–1.23)	NA
Blindness	2 ^{31,34}	2/981	2/962	0.97(0.14–6.90)	0.0
Deafness	2 ^{31,34}	8/981	11/962	0.51 (0.05–4.96)	58.7

CI, confidence interval; NA, not applicable

TABLE 7

Effect of magnesium sulfate on maternal outcomes

Outcome	No of trials	Number of events/total number		Relative risk (95% CI)	I ² (%)
		Magnesium	No magnesium		
Death	3 ^{31,33,35}	0/1917	1/1950	0.32(0.01–7.92)	0.0
Cardiac or respiratory arrest	3 ^{31,33,35}	0/1917	0/1950	Not estimable	NA
Pulmonary edema	1 ³⁵	8/1096	3/1145	2.79(0.74–10.47)	NA
Respiratory depression	2 ^{31,35}	41/1631	31/1672	1.31 (0.83–2.07)	0.0
Hypotension	2 ^{31,33}	80/821	52/805	1.51 (1.09–2.09)	3.6
Tachycardia	1 ³¹	56/535	36/527	1.53(1.03–2.29)	NA
Severe postpartum hemorrhage	2 ^{31,33}	28/821	26/805	1.06(0.63–1.79)	0.0
Cesarean section	3 ^{31,33,35}	822/1917	834/1950	1.00(0.93–1.07)	21.6
Clinical and self-assessed maternal side effects of the infusion					
Flushing	3 ^{31,33,35}	1119/1917	162/1950	7.56(3.39–16.88)	93.8
Nausea or vomiting	3 ^{31,33,35}	312/1917	76/1950	4.60(1.54–13.75)	91.5
Sweating	2 ^{31,35}	411/1631	57/1672	6.37(1.96–20.68)	94.6
Problems at injection site	2 ^{31,35}	614/1631	68/1672	9.12(7.19–11.57)	0.0
Stopping of infusion because of adverse effects	2 ^{31,35}	123/1631	44/1672	2.81 (2.01–3.93)	0.0
Any side effect	3 ^{31,33,35}	1356/1917	343/1950	5.05(2.06–12.39)	98.3

CI, confidence interval; NA, not applicable