



Published in final edited form as:

Am J Obstet Gynecol. 2012 May ; 206(5): 447.e1–447.e6. doi:10.1016/j.ajog.2012.01.033.

Comprehensive analysis of *LAMC1* genetic variants in advanced pelvic organ prolapse

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Abstract

OBJECTIVE—We sought to comprehensively evaluate the association of laminin gamma-1 (*LAMC1*) and advanced pelvic organ prolapse.

STUDY DESIGN—We conducted a candidate gene association of patients (n =239) with stages III–IV prolapse and controls (n =197) with stages 0–I prolapse. We used a linkage disequilibrium (LD)-tagged approach to identify single-nucleotide polymorphisms (SNPs) in *LAMC1* and focused on non-Hispanic white women to minimize population stratification. Additive and dominant multivariable logistic regression models were used to test for association between individual SNPs and advanced prolapse.

RESULTS—Fourteen SNPs representing 99% coverage of *LAMC1* were genotyped. There was no association between SNP rs10911193 and advanced prolapse ($P=.34$). However, there was a trend toward significance for SNPs rs1413390 ($P=.11$), rs20563 ($P=.11$), and rs20558 ($P=.12$).

CONCLUSION—Although we found that the previously reported *LAMC1* SNP rs10911193 was not associated with nonfamilial prolapse, our results support further investigation of this candidate gene in the pathophysiology of prolapse.

Keywords

candidate gene; genetic epidemiology; *LAMC1*; laminin gamma-1; pelvic organ prolapse

Pelvic organ prolapse is a major public health issue for women because 40% of postmenopausal women are affected.¹ Symptomatic prolapse results in a marked reduction in quality of life because of pelvic discomfort and bladder and bowel dysfunction.² Although prolapse can be managed surgically, the recurrence rate is as high as 30%,³ and the annual health care costs of surgery alone total more than a billion dollars in the United States.⁴ Given the public health impact of prolapse, it is critical to understand the underlying pathophysiology of this disease to develop effective prevention strategies and specific treatment recommendations for high-risk women.

Currently, the underlying mechanisms of prolapse are poorly understood; however, a growing body of evidence supports a genetic predisposition.^{5–9} Epidemiologic studies suggest that a family history of prolapse is a significant risk factor for disease.^{6,8} A pedigree analysis of familial prolapse found that inheritance occurs in an autosomal dominant fashion with incomplete penetrance,⁵ and a linkage analysis of affected sibling pairs identified a predisposition gene for pelvic floor disorders on chromosome 9q21.¹⁰

Hypotheses regarding the pathophysiology of prolapse include abnormal synthesis or degradation of collagen¹¹ and elastin;^{12,13} thus, genetic mutations in extracellular matrix (ECM) pathways may result in the genesis and progression of prolapse. Nikolova et al⁹ identified a single-nucleotide polymorphism (SNP) (rs10911193) in the *LAMC1* gene encoding laminin, an ECM component in a study of familial prolapse. However, when studying nonfamilial prolapse in a case-control association study of 265 Caucasians and 146 African Americans, Chen et al¹⁴ reported that this laminin mutation and 2 additional functional SNPs in *LAMC1* (rs20563 and rs20558) were not associated with advanced prolapse. One limitation of the study by Chen et al may have been the smaller sample size.

Given the significant findings of *LAMC1* in familial prolapse by Nikolova et al,⁹ we felt that further investigation of *LAMC1* as a potential candidate gene for advanced pelvic organ prolapse was warranted in a larger cohort of unrelated individuals. In this study, we sought to perform a comprehensive analysis of the entire *LAMC1* gene, using a linkage disequilibrium (LD)-tagged SNP approach, in a larger case-control association study of non-Hispanic white women.

Materials and Methods

Study population

After institutional review board approval, we recruited subjects from the Division of Urogynecology at Duke University Medical Center for this case-control association study from February 2009 to December 2010.¹⁵ Patients were defined as nonpregnant women with stages III–IV prolapse based on the Pelvic Organ Prolapse Quantification (POP-Q) examination. Controls were women with stages 0–I prolapse who had no prior history of prolapse surgery. Women with a prior hysterectomy for nonprolapse indications were eligible to be controls.

We excluded women with a gynecological malignancy or inherited or acquired connective tissue diseases, including rheumatoid arthritis, systemic lupus erythematosus, scleroderma, polymyositis, dermatomyositis, or Marfan or Ehlers-Danlos syndromes from both groups. Sociodemographic data, past medical and obstetric history, and physical examination information were collected, which included age, race, ethnicity, parity, body mass index (BMI), and POP-Q points.

Although we recruited subjects of all ages with advanced prolapse, one approach commonly used in genetic epidemiology studies is to preferentially enroll younger subjects with more advanced disease to select for a high-risk group with a presumably higher genetic load.^{16,17} In this study, we preferentially recruited younger women with advanced prolapse for our patients, and for our control group, we preferentially recruited older women. Because the prevalence of prolapse increases with age, older controls help to ensure that these individuals are not at risk for developing prolapse. By selecting these more extreme cases, we attempted to enhance our ability to detect genetic differences between our 2 groups.

To avoid issues of confounding by race, also known as population stratification, we focused on non-Hispanic white women in this study. By focusing on non-Hispanic white women, our

goal was to further ensure that our cases and controls were derived from the same source population.¹⁸ In other words, we matched on race and ethnicity to minimize the risk of differences in the genetic background of our 2 cohorts.¹⁹ It is critically important to consider racial and ethnic groups in genetic epidemiologic studies because different racial/ethnic groups can have markedly different population frequencies of genetic variants as well as different linkage disequilibrium structures.

Genotyping and SNP selection

Genomic deoxyribonucleic acid was extracted from frozen stored peripheral venous whole-blood samples. To enable a more comprehensive evaluation of the *LAMC1* gene, we utilized an LD tagged SNP approach. A tagged SNP approach is an efficient methodology for comprehensive evaluation of a gene because most SNPs are in LD (ie, genetically linked and thus correlated) with other SNPs.^{20–22} That is, within each gene, there are groups of SNPs that are highly correlated with each other. Representative SNPs from each group (ie, tagged SNPs) can be assessed, which represents a detailed evaluation of the gene without having to genotype every SNP. Thus, a tagged SNP approach is comprehensive as well as efficient, in terms of time and cost.

We utilized the Tagger subsection of Haploview 4.2 software (Broad Institute, Cambridge, MA)²³ to identify tagged SNPs in *LAMC1* as well as in the 5 kb upstream and downstream of *LAMC1* and used a correlation (r^2) of greater than 0.8 and SNPs with a minor allele frequency (MAF) of greater than 0.05. We also included known and putative functional coding SNPs with an MAF greater than 0.05 from any given LD bin.

Genotyping was performed using the 7900 HT TaqMan system, which incorporates a standard PCR-based, dual-fluorescence, allele discrimination assay with a dual-laser scanner. Assays by Design and Assays on Demand were purchased through Applied Biosystems (Applied Biosystems, Foster City, CA). Quality control (QC) samples, composed of 12 reference controls, were included in each quadrant of the plate. All SNPs examined were successfully genotyped for 98% or greater of the individuals in the study. Error rate estimates for SNPs meeting QC benchmarks were 0.2%.

Statistical analysis

Sociodemographic data and physical examination findings for cases and controls were compared using a χ^2 test for discrete variables and a Student *t* test for continuous variables. A 2-sided alpha of 0.05 was considered significant. For all SNPs, deviation from Hardy-Weinberg equilibrium (HWE) was tested using a χ^2 test. Genotype clusters and genotyping calling for SNPs that were significantly out of HWE were re-reviewed and genotyping accuracy confirmed. To test for association of single SNPs with pelvic organ prolapse, we constructed logistic regression models and evaluated both unadjusted and adjusted models, using both additive and dominant models of genetic risk. The multivariable models adjusted for age, vaginal parity (dichotomous variable regarding any prior vaginal deliveries), and BMI.

A full-tagged SNP approach was conducted in an initial cohort of 396 non-Hispanic white women. Because 3 SNPs showed a nominal *P* value of .10, we expanded the final cohort to 436 women to increase the power. Given these independent effects of several SNPs, we also performed a haplotype analysis of all 2-way SNP combinations using the tool set of PLINK²⁴ to evaluate the possible association of a genetic background of more than 1 SNP for pelvic organ prolapse risk.

A power calculation was conducted for a nonfamilial case-control association using the QUANTO program (Gauderman and Morrison, <http://hydra.usc.edu/gxe>. Accessed Oct. 24,

2008).²⁵ A range of allele frequencies of 0.05–0.5 and varying effect sizes as well as dominant and additive models were considered. We recognize that these calculations are based on model parameters that we cannot know with certainty; however, we can derive a sense of the expected power for the proposed sample sizes. Given a sample size of 150 cases and 150 controls, we have greater than 80% power to detect an effect size of approximately 2.0 and higher for most allele frequencies and genetic models. We chose an effect size of 2.0 because we believed that this is a clinically meaningful difference to detect and represented a balance between clinical/biological significance and feasibility. All statistical analyses were conducted with SAS version 9.2 (SAS Institute, Cary, NC) unless otherwise noted.

Results

We enrolled a total of 436 women, of whom 239 were cases with advanced prolapse and 197 were controls with none to minimal prolapse. Cases were slightly younger (64.8 ± 10.3 vs 69.0 ± 10.2 years, $P < .001$) and more likely to have had 1 or more vaginal deliveries (96.6% vs 82.2%, $P < .001$) when compared with controls (Table 1). Cases and controls were similar in the Charlson co-morbidity index, smoking status, and BMI (27.8 ± 5.1 vs 28.5 ± 6.3 kg/m², $P = .18$). Of controls, 57.4% had stage 0 prolapse and 42.6% had stage I prolapse. For cases, a majority had stage III prolapse (94.6%) and 5.4% had stage IV prolapse.

We genotyped 15 SNPs based on our LD tagging selection. The Figure is a schematic of the *LAMC1* gene structure, which indicates exons, introns, and sites of translation initiation and termination. The genotyped SNPs are marked on the schematic with descriptions of the different types of SNPs, whether nonsynonymous coding, synonymous coding, or intronic. All SNPs were in Hardy-Weinberg equilibrium except rs17477864 ($P < .001$). This SNP was regenotyped and remained out of HWE and thus was not included in the final analysis. At an r^2 of 0.80, the remaining 14 SNPs provided 99% coverage of the gene.

In this cohort of non-Hispanic white women, there was no association between the previously reported *LAMC1* SNP rs10911193 with advanced pelvic organ prolapse, based on both the dominant and additive adjusted models (odds ratio [OR], 1.29; 95% confidence interval [CI], 0.77–2.17; $P = .34$) and OR, 1.23; 95% CI, 0.77–1.97; $P = .39$, respectively) (Table 2). However, there was a trend toward significance for the following SNPs in the dominant model: rs1413390, an intronic SNP (adjusted OR, 1.45; 95% CI, 0.92–2.29; $P = .11$), rs20558; a nonsynonymous coding SNP (adjusted OR, 1.44; 95% CI, 0.91–2.26; $P = .12$); and rs20563, a nonsynonymous coding SNP (adjusted OR, 1.44; 95% CI, 0.92–2.28; $P = .11$) (Table 2). These 3 SNPs were in perfect linkage disequilibrium with each other ($r^2 = 1.0$) (Figure). No other SNPs in *LAMC1* were associated with prolapse in either the dominant or additive adjusted models.

In haplotype analyses of all 2-SNP combinations, only 2 haplotypes were statistically significantly associated with prolapse (rs12739316_rs3768617, omnibus, $P = .005$; and rs10911211_rs12739316, $P = .04$). Interestingly, 3 haplotypes including the previously reported rs10911193 and each of the borderline significant SNPs from our study also showed a trend for significance in haplotype analyses (rs10911193_rs20563, omnibus, $P = .07$; rs10911193_rs20558, omnibus, $P = .07$; and rs10911193_rs1413390, omnibus, $P = .06$). Although none of the haplotype analyses would meet statistical significance after adjustment for multiple comparisons, these results continue to corroborate the potential role of these SNPs in prolapse.

Comment

In this comprehensive analysis of *LAMC1* genetic variants, we provide additional evidence that *LAMC1* is a potential candidate gene for nonfamilial, advanced pelvic organ prolapse.

Although we did not find an association between the previously reported SNP rs10911193 and advanced prolapse, we report on 3 *LAMCI* SNPs that show a trend toward significance, although they were not strictly statistically significant.

LAMCI encodes the γ 1 chain of laminin. Laminins are a heterogeneous group of ECM glycoproteins that form the major noncollagenous portion of the basement membranes.²⁶ There are at least 15 distinct forms of laminin, and these isoforms are comprised of 3 nonidentical chains: α , β , and γ chains. Nikovola et al⁹ first reported that the *LAMCI* SNP rs10911193 minor T allele was more prevalent among probands in a high-risk family with multigenerational early-onset prolapse and hernias when compared with the general population. Subsequently, Chen et al¹⁴ evaluated 3 *LAMCI* SNPs (rs10911192, rs20563, and rs20558) in 265 Caucasians (102 cases and 163 controls) and 146 African Americans (63 cases and 83 controls) and found no association with nonfamilial prolapse. However, their power to detect a significant association may have been limited by their sample size.

In our study, rs20563 and rs20558, both nonsynonymous, or missense, SNPs, showed a trend toward significance. The rs20563 G produces an amino acid substitution of isoleucine to valine, which is a fairly conservative change because both are amino acids with hydrophobic side chains. The rs20558 C results in a less conservative amino acid change from leucine to proline. The amino acid changes associated with rs20563 and rs20558 are predicted to have tolerated and benign effects on the protein function based on the Sorting Tolerant From Intolerant²⁷ and the Polymorphism Phenotyping v2²⁸ predictions in the Ensembl genome browser.^{29,30}

A major strength of this study is that we used a tagged SNP approach to comprehensively evaluate the entirety of the *LAMCI* gene. We evaluated a total of 14 SNPs, which provided 99% coverage of the gene. In contrast, Chen et al¹⁴ evaluated only 3 SNPs in *LAMCI*. Another critical factor is that we phenotyped cases and controls based on a standardized objective assessment, the POP-Q examination, and we selected more extreme phenotypes to try to concentrate the genetic effect. Another important aspect of this study is that we conducted a case-control association study, which evaluates unrelated individuals. It is critical to distinguish between studying a high-risk family vs unrelated individuals because the disease phenotype may be quite different.

A limitation of our study is that we were powered to detect an effect size of 2.0, which represented a balance between clinical/biological significance and feasibility; however, we did not have sufficient power to detect more modest effects (ie, an OR of <2.0). Thus, although we showed a trend toward significance for 3 SNPs, we were likely underpowered to show a significant association. Of note, effect sizes less than 2.0 have been shown to be clinically significant in other phenotypes; for example, the most consistent and recognized genetic locus for susceptibility to coronary artery disease, chromosome 9p21, shows effect sizes of 1.3–1.6.³¹ In this study, we focused on non-Hispanic white women to avoid population stratification;¹⁹ however, future genetic epidemiological studies should evaluate other races/ethnicities.

Another important issue in a case-control study is careful selection of the control subjects. In this study, we defined our cases and controls based on stage of prolapse. However, at this point, we do not know the ideal method for phenotyping pelvic organ prolapse. Should we define cases and controls based on physical examination and stage of prolapse, as we did in this study? Should we define prolapse based on the degree of bothersome symptoms, age at onset, or history of treatment, whether with surgery or a pessary? Should family history of prolapse be included? Also, it is possible that different pelvic floor disorders may share the same genetic susceptibility, for example, stress urinary incontinence (SUI) and prolapse. If

SUI and prolapse share the same genetic susceptibility, there is the potential to introduce misclassification bias if controls have SUI, even though they did not have prolapse. This point highlights the importance for future research into how best to phenotype prolapse, as well as the other pelvic floor disorders, and the importance of genetic and genetic epidemiological research of pelvic floor disorders.

Our current knowledge of the genetic susceptibility of pelvic organ prolapse is limited, although research in this field is growing. Pelvic organ prolapse is considered a common complex disease, for which there are likely multiple genetic susceptibility markers. Historically, the field of human genetics was focused on Mendelian disorders, in which single-gene mutations with a high penetrance and large effect size result in rare diseases in families. There has been a paradigm shift to study common, complex diseases, in which there is a polygenic model of genetic risk, with multiple, often common, genetic polymorphisms conferring moderate risk in a general population, with concomitant gene-gene and gene-environment interactions. For prolapse, additional candidate genes that have been investigated include collagen type III alpha 1 chain (*COL3A1*),^{32,33} collagen type I alpha 1 chain (*COL1A1*),^{34,35} and matrix metalloproteinase-9 (*MMP9*).^{13,36}

Animal models have also highlighted the association of prolapse and defective ECM proteins, such as lysyl oxidase-like-1,^{12,37} fibulin-3,³⁸ and fibulin-5.³⁹ Although these findings highlight potential candidate genes for pelvic organ prolapse, we have yet to fully understand the pathophysiology of pelvic organ prolapse.

It is critically important to pursue future investigations into the genetics of pelvic organ prolapse because this research may uncover the underlying mechanisms for this highly prevalent and burdensome disease. These data will not only provide invaluable information regarding the pathophysiology of disease but will also help to identify genetically high-risk women to target for prevention strategies or specific treatment recommendations. *LAMC1* is an interesting candidate gene, and future large-scale investigations into this gene as well as other extracellular matrix genes are definitely warranted.

Acknowledgments

This study was supported by grant R03HD061766 from the *Eunice Kennedy Shriver* National Institute of Child Health and Human Development. J.M.W. is supported by grant K23HD068404 from the *Eunice Kennedy Shriver* National Institute of Child Health and Human Development.

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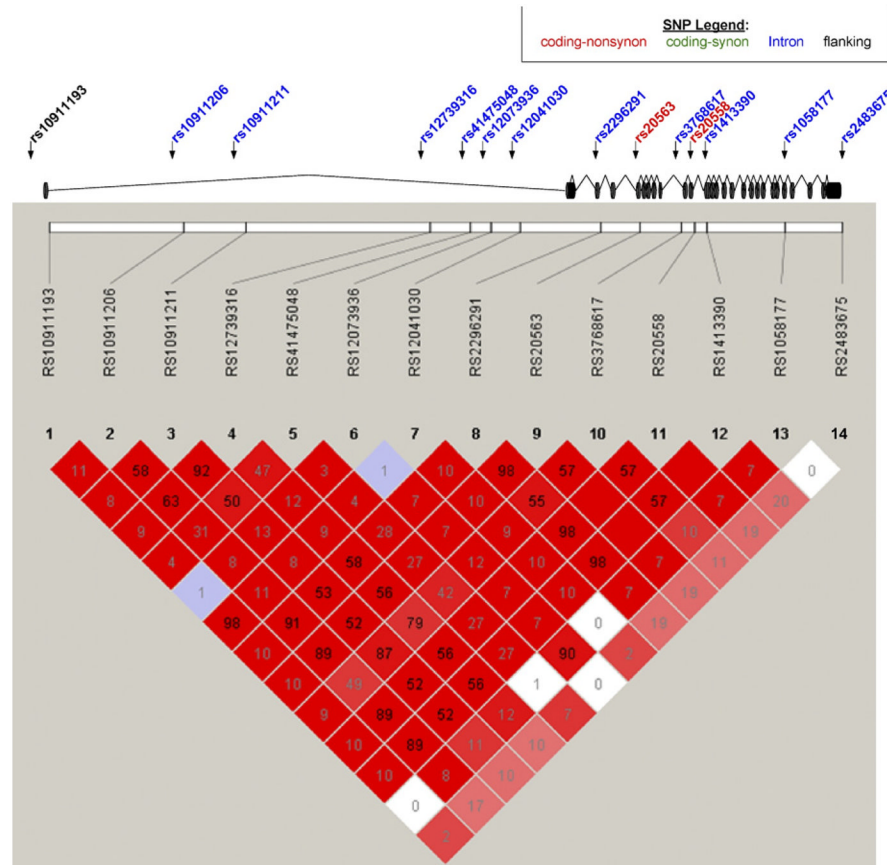


Figure 1. The *LAMCI* gene structure is displayed diagrammatically with exons represented as *black discs*. Genotyped SNPs are noted with *arrows* and are color coded as follows: coding-nonsynonymous (*red*); coding-synonymous (*green*); intronic (*blue*); and flanking (*black*). The *bottom panel* illustrates an LD graphic with r^2 values and squares without any value indicates perfect LD. *LAMCI*, laminin gamma-1; *LD*, linkage disequilibrium; *SNP*, single-nucleotide polymorphism.

TABLE 1

Sociodemographic characteristics and physical examination findings for cases and controls

Characteristic	Controls (n =197)	Cases (n =239)	P value
Age, y	69.0 ±10.2	64.8 ±10.3	<.001
Vaginal parity	162 (82.2)	230 (96.6)	<.001
Charlson comorbidity index	0.48 ±0.90	0.42 ±0.83	.51
Current smoker	8 (4.1)	16 (6.8)	.26
Body mass index, kg/m ²	28.5 ±6.3	27.8 ±5.1	.18
Prolapse stage			
0	113 (57.4)	0 (0.0)	
I	84 (42.6)	0 (0.0)	
II	0 (0.0)	0 (0.0)	
III	0 (0.0)	226 (94.6)	
IV	0 (0.0)	13 (5.4)	

Data presented as mean ±SD or n (%).

Logistic regression results for both the additive and dominant adjusted models for *LAMC1* SNPs adjusted for age, BMI, and vaginal parity

TABLE 2

SNP	Minor allele (frequency)	Dominant adjusted model		Additive adjusted model	
		OR (95% CI)	P value	OR (95% CI)	P value
rs10911193	T (0.11)	1.29 (0.77–2.17)	.34	1.23 (0.77–1.97)	.39
rs1413390	G (0.45)	1.45 (0.92–2.29)	.11	1.22 (0.90–1.65)	.21
rs20558	T (0.45)	1.44 (0.91–2.26)	.12	1.19 (0.88–1.61)	.25
rs20563	A (0.46)	1.44 (0.92–2.28)	.11	1.20 (0.89–1.63)	.23
rs10911206	A (0.48)	1.37 (0.85–2.20)	.19	1.16 (0.85–1.57)	.35
rs2296291	C (0.45)	1.36 (0.85–2.16)	.20	1.14 (0.84–1.56)	.39
rs12041030	G (0.11)	1.36 (0.80–2.29)	.25	1.28 (0.80–2.07)	.30
rs12739316	T (0.41)	0.79 (0.51–1.23)	.30	0.82 (0.60–1.11)	.20
rs3768617	T (0.42)	0.83 (0.53–1.31)	.42	0.84 (0.61–1.16)	.29
rs2483675	A (0.22)	1.15 (0.76–1.74)	.51	1.17 (0.82–1.68)	.38
rs10911211	T (0.40)	0.89 (0.57–1.39)	.62	0.90 (0.66–1.24)	.52
rs41475048	T (0.25)	1.04 (0.69–1.56)	.86	1.08 (0.77–1.50)	.66
rs1058177	C (0.09)	1.01 (0.58–1.76)	.97	1.01 (0.58–1.76)	.97
rs12073936	G (0.08)	1.00 (0.56–1.78)	.99	0.96 (0.55–1.68)	.88

BMI, body mass index; *CI*, confidence interval; *LAMC1*, laminin gamma-1; *OR*, odds ratio; *SNP*, single-nucleotide polymorphism.