



# Internet-Based Treatment of Pathological Gambling with a Three-Year Follow-Up

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**Abstract.** Effective therapies for pathological gambling exist, but their use is limited to about 10% of the target population. In an attempt to lower the barriers for help, Internet-based cognitive behavioural therapy (ICBT) has been shown to be effective when delivered to a non-depressed sample with pathological gambling. This study sought to extend this finding to a larger, more representative population, and also test a model to predict responder status. Following advertisement, a total of 284 participants started an 8-week ICBT programme with minimal therapist contact via e-mail and weekly telephone calls of less than 15 min. The average time spent on each participant, including telephone conversations, e-mail, and administration, was 4 h. In addition to a mixed effects model to evaluate the effectiveness of the treatment, two logistic regression analyses were performed with the following eight pre-defined response predictor variables: work-life satisfaction, primary gambling activity, debts due to gambling, social support, personal yearly salary, alcohol consumption, stage of change, and dissociative gambling. ICBT resulted in statistically significant reductions in the scores of pathological gambling, anxiety, and depression as well as an increase in quality of life compared to pre-treatment levels. Follow-ups carried out in the treatment group at 6, 18, and 36 months indicated that treatment effects were sustained. Using the eight predictor variable model rendered an acceptable predictive ability to identify responders both at post-test ( $AUC = .72, p < .01$ ) and at 36-month follow-up ( $AUC = .70, p < .01$ ). We conclude that ICBT for pathological gamblers, even if depressed, can be effective and that outcome can partly be predicted by pre-treatment characteristics.  
*Key words:* gambling; Internet; treatment; prediction

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## Introduction

Since pathological gambling was first introduced in the *Diagnostic and statistical manual of mental disorders* (American Psychiatric Association, 1980), there has been a fast development in the availability of games, primarily via the Internet (Hodgins, Stea, & Grant, 2011). One example is Sweden which, in 2006, introduced a state-owned Internet poker site which is essentially open all the time. Nowadays you can, in principle, sit in your

underwear at home in the middle of the night with a whisky in your hand and gamble away all your savings. Social responsibility is said to be prioritised by the gaming operator, yet no peer-reviewed research has been published evaluating the effects of the preventive steps taken (e.g. warning messages) to ensure responsible gaming. The risks of developing pathological gambling habits are probably increased by the availability of online games, but this increased access has not been mirrored by increased access to psychological help in

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case gambling progresses into pathological gambling. This is unfortunate, since there are effective psychosocial treatments (Gooding & Tarrier, 2009; Pallesen, Mitsem, Kvale, Johnsen, & Molde, 2005). However, only about 10% of persons with pathological gambling seek treatment (Ladouceur, 2005). Besides a lack of therapists, long waitlists, and costs, it is likely that shame and stigma influence the decision to not seek help (Evans & Delfabbro, 2005). Therefore, there is a need to lower the barriers for seeking help. One way of achieving this objective is to offer help online.

However, the progress of online treatments for gambling problems has been slow. For instance, in a recent systematic review of Internet-based therapy studies on addictions, Gainsbury and Blaszczynski (2011) concluded that only one study has been carried out with acceptable methodology on pathological gambling. In that particular study, carried out by our research group, Carlbring and Smit (2008) randomised 66 participants to a treatment group or a waitlist control. The results were generally encouraging, with treatment gains maintained up to 36 months after treatment completion. However, the trial excluded all participants with a depression score exceeding 21 points on the self-rated version of the Montgomery-Asberg Depression Rating Scale (Svanborg & Asberg, 1994). Although the idea was to reduce the risk of including participants with suicidal ideation, since no face-to-face contact was offered, this procedure limits the ability to generalise, especially since more people were excluded ( $n = 158$ ) than included ( $n = 66$ ). It is uncertain if the results from that particular study can be generalised to the larger population of pathological gamblers. Consequently, the present study was designed to keep the exclusion frequency to a minimum by, for example, including participants regardless of their depression score.

Not everyone entering treatment for pathological gambling completes it and, of those who do, not everyone reaches high end-state functioning (Carlbring, Jonsson, Josephson, & Forsberg, 2010). Attempts have been made to try to predict who will drop out (Dowling, 2009) and who will benefit from treatment (Dowling, 2009; Dowling & Cosic, 2010; Dunn, Delfabbro, & Harvey, 2011; Raylu & Oei, 2007). In Internet-based programmes this is an appealing idea. If we knew who would

not likely benefit from a low-intensity intervention, we could override the stepped care approach in which a person receives the lowest service tier in the first instance, only stepping up to more intensive services as clinically required (Bower & Gilbody, 2005). Unfortunately, in order for patients to progress to the next level, a treatment failure needs to be recognised. If we had a way of predicting who would benefit from treatment, then those who would not benefit could skip the Internet-based treatment and go directly to a more intensive intervention.

The results of the pathological gambling treatment prediction studies are mixed. Some studies have suggested that high pre-treatment alcohol use is a negative predictor, while others have failed to replicate that finding (Dowling, 2009). Other studies have found that there is a lower likelihood of treatment response if the participant is young, living alone, has a low income and a low readiness to change (Dowling, 2009). In a study by Dowling on females presenting with pathological gambling, no differences could be found between dropouts and completers (Dowling, 2009). One reason for the inconclusiveness could be that researchers use different definitions of what a dropout is (*cf.* Eysenbach, 2005) and different predictors. Another problem is that groups are generally small, making statistical methods like logistic regression less suitable. Finally, there are reports that there might be different predictors in face-to-face therapy and Internet-based treatment (Andersson, Carlbring, & Grimund, 2008; Nordgreen et al., 2012).

In this study, we investigated predictors of who will have a successful outcome following ICBT by the use of eight variables collected in the screening phase. Previous studies have reasoned that since pathological gambling and alcohol and substance abuse share many commonalities, it is reasonable to assume that predictors of outcome in those studies also could have some relation to pathological gambling (Dowling, 2009). Even if results are inconclusive in the gambling field as to what variables can be used to predict the outcome, we decided to use eight different predictors in different domains that both theoretically and clinically would be plausible predictors of outcome. First, we entered work-life satisfaction as suggested by Sander and Peters (2009).

Second, primary gambling activity was used as a predictor variable since severity of gambling and psychosocial problems have been associated with preferred form of gambling activity (Champine & Petry, 2010). Third, debts due to gambling were entered, because financial difficulties have been reported to be an important motivator (Suurvali, Hodgins, & Cunningham, 2010). The fourth variable was social support, as lower levels have been shown to correlate with therapy dropout (Dunn et al., 2012; Melville, Casey, & Kavanagh, 2007; Oakes et al., 2011). Fifth, personal yearly salary was used as a proxy for socioeconomic status which is a known risk factor for pathological gambling (Welte, Barnes, Wiczorek, Tidwell, & Parker, 2004). Sixth, alcohol consumption was included as suggested by Rash, Weinstock, and Petry (2011), since it has been shown that those who drink while gambling tended to gamble in more risky ways and experience more negative consequences of gambling (Cronce & Corbin, 2010). Hazardous alcohol use has previously been linked to continued difficulties with gambling over a 7-year period (Abbott, Williams, & Volberg, 2004). Seventh, in accordance with the transtheoretical model of behaviour change (Prochaska, DiClemente, & Norcross, 1992), the individual's readiness to act on a new healthier behaviour (stage of change) was included as suggested by Wohl and Sztainert (2011). Finally, dissociative gambling was included as suggested by Jacobs (1988). Although pathological gamblers might not experience dissociative symptoms at a higher rate than normal controls (Ledgerwood & Petry, 2006), it still can be a possible predictor of treatment outcome as it can differentiate different subtypes (Milosevic & Ledgerwood, 2010).

## Methods

### *Design*

Since the superiority of the treatment over a waitlist control had previously been demonstrated (Carlbring & Smit, 2008), this trial was designed as a non-comparative, single group study with measurements at baseline and at 3, 6, 18, and 36 months. The effect of the intervention was assessed using four criteria: Gambling (NORC DSM-IV Screen for gambling problems [NODS]; Gerstein et al., 1999), anxiety, depression (Hospital Anxiety and

Depression Scale [HADS]; Zigmond & Snaith, 1983), and overall life satisfaction (quality of life inventory [QOLI]; Frisch, Cornell, Villanueva, & Retzlaff, 1992). The study protocol was approved by the regional medical ethics committee (Vetting the ethics of research involving humans in Sweden) and written informed consent was obtained from the participants. The trial is registered in Clinical Trials with the identifier NCT01381250.

### *Participants and recruitment*

Advertisements in newspapers with the heading "Do you have trouble controlling your gambling?" and "Do you want to stop gambling?" were used to recruit participants (*cf.* Doiron & Nicki, 2007). Selection took place with a computerised screening interview consisting of the NODS (Gerstein et al., 1999), the HADS (Zigmond & Snaith, 1983), the QOLI (Frisch et al., 1992), and a number of additional questions related to gambling activities and demographics.

To be included in the study, participants had to meet the following criteria: (1) fulfil the DSM-IV-TR (American Psychiatric Association, 2000) criteria for pathological gambling according to the 1-month version of NODS, (2) live in Sweden, and (3) have gambled at least once in the past 30 days.

Of the 464 individuals who applied between May 2005 and May 2007, 316 fulfilled the criteria and 284 subsequently started treatments. The reasons for exclusion are specified in the CONSORT flowchart (see Figure 1). The mean age of the included 284 participants (81% male) was 32.2 (SD = 8.8; range: 18–62) and the mean age of onset of regular gambling was 23.7 (SD = 9.1). The mean duration of gambling problems was 5.7 years (SD = 5.5). The most prevalent type of game played was poker and video lottery terminals (VLTs). For a more detailed description, see Tables 1 and 2.

### *Treatment*

The intervention was based on established CBT methods, as described in self-help books (Hodgins, 2002; Ladouceur & Lachance, 2006) and is described in more detail in Carlbring and Smit (2008). Briefly, the treatment can be described as an eight-chapter book on the Internet, with each of the chapters covering a specific topic. The modules include information and exercises and end with three to

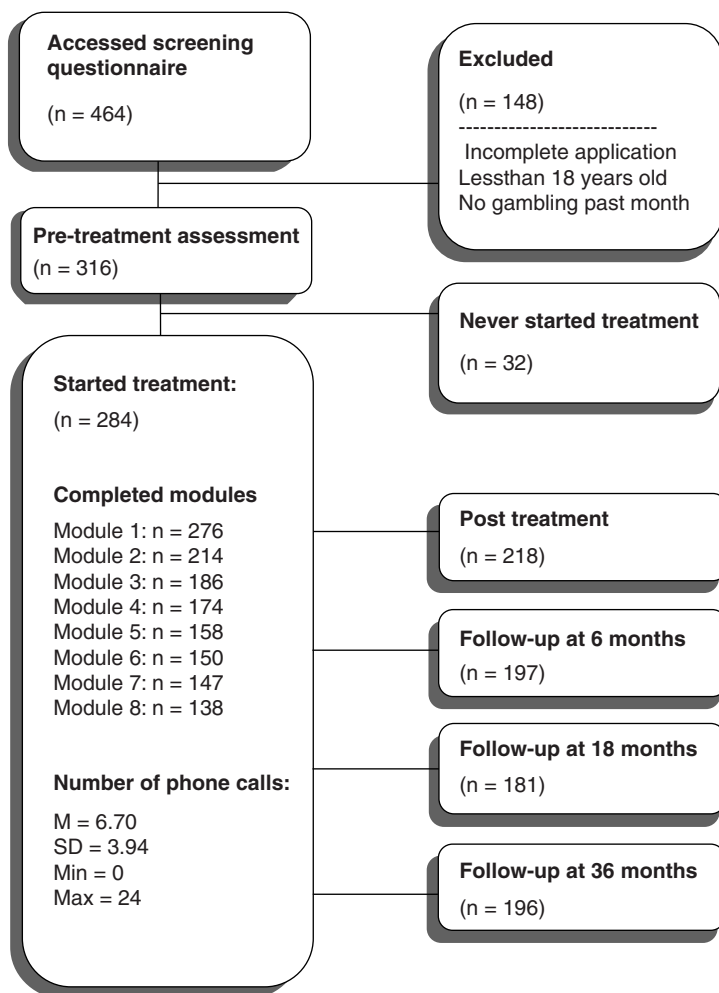


Figure 1. Participant flow, reasons for exclusion and number of participants providing data at different assessment points throughout the trial.

eight essay-style questions. Participants were asked to answer the questions and provide worksheets and report on the outcomes of different exercises. For each module they were required to post at least one message in an online discussion group about a predetermined topic. Feedback on homework assignments was usually given via e-mail within 24 h after participants had sent their answers. Once a week telephone calls were made by the therapists to each participant. The purpose was to provide positive feedback and encouragement as well as to answer any questions the subjects had regarding the modules. Each conversation lasted approximately 15 min and

a mean total of 6.7 ( $SD = 3.9$ ) calls were made during the eight weeks.

The therapists were four social workers with an additional 2-year basic training in CBT and motivational interviewing (MI; Miller & Rollnick, 2002). The mean total time per week spent on each participant in this study was approximately 20 min, including telephone calls, administration, and responding to e-mails.

### **Predictor variables**

A total of eight predictor variables were selected in order to predict responder status: (1) Work-life satisfaction was defined by the score in the corresponding question in the QOLI (Frisch et al., 1992) and subsequently

Table 1. Treatment outcome at pre-, post-, 6-, 18-, and 36-month follow-up for the collapsed group (n = 284) as well as the five different subtypes of primary problematic game using mixed effects model

Time	Mean estimates (SE) across all types of games	Significant pair wise comparisons between time points	Mean estimates (SE), type of game					p Differences	Significant differences between each type of game at each time point	p Interaction	
			Time	Poker (n = 102)	VLT (n = 76)	Casino (n = 34)	Betting (n = 43)				Bingo (n = 29)
Gambling (NORC DSM-IV Screen for gambling problems)											
Pre	8.1 (0.1)	Pre > all times	<0.01	8.2 (0.1)	8.5 (0.2)	8.0 (0.2)	8.0 (0.2)	7.7 (0.3)	0.042	VLT > Betting, Bingo	0.748
Post	1.8 (0.2)			2.0 (0.2)	2.3 (0.2)	1.8 (0.3)	1.8 (0.3)	1.4 (0.3)		VLT > Betting, Bingo	
6 mo	2.0 (0.2)			2.1 (0.2)	2.4 (0.2)	1.9 (0.3)	1.9 (0.3)	1.5 (0.3)		VLT > Betting, Bingo	
18 mo	2.1 (0.2)			2.2 (0.2)	2.5 (0.3)	2.0 (0.3)	2.0 (0.3)	1.7 (0.3)		VLT > Betting, Bingo	
36 mo	2.0 (0.2)			2.1 (0.2)	2.4 (0.3)	1.9 (0.3)	1.9 (0.3)	1.5 (0.3)		VLT > Betting, Bingo	
Anxiety (Hospital anxiety and depression scale)											
Pre	11.0 (0.3)	Pre > all times	<0.01	10.7 (0.4)	11.7 (0.4)	11.7 (0.6)	10.5 (0.6)	10.5 (0.7)	0.169	No difference	0.425
Post	7.2 (0.3)	Post > 6, 18		6.8 (0.4)	7.9 (0.5)	7.8 (0.6)	6.6 (0.6)	6.7 (0.7)		No difference	
6 mo	6.4 (0.3)			6.1 (0.4)	7.1 (0.5)	7.1 (0.6)	5.9 (0.6)	5.9 (0.7)		No difference	
18 mo	6.3 (0.3)			5.9 (0.4)	7.0 (0.5)	7.0 (0.7)	5.8 (0.6)	5.8 (0.7)		No difference	
36 mo	6.5 (0.4)			6.1 (0.4)	7.2 (0.5)	7.1 (0.7)	5.9 (0.6)	6.0 (0.7)		No difference	
Depression (Hospital anxiety and depression scale)											
Pre	9.3 (0.3)	Pre > all times	<0.01	8.8 (0.4)	9.3 (0.4)	10.6 (0.6)	9.3 (0.5)	8.6 (0.6)	0.061	Casino > Poker, VLT, Bingo	0.483
Post	5.5 (0.3)	Post > 6		5.0 (0.3)	5.5 (0.4)	6.8 (0.6)	5.6 (0.5)	4.8 (0.6)		Casino > Poker, VLT, Bingo	
6 mo	4.8 (0.3)			4.3 (0.4)	4.8 (0.4)	6.1 (0.6)	4.8 (0.5)	4.1 (0.6)		Casino > Poker, VLT, Bingo	
18 mo	5.0 (0.3)			4.5 (0.4)	5.0 (0.4)	6.3 (0.6)	5.0 (0.5)	4.3 (0.6)		Casino > Poker, VLT, Bingo	

Table 1. Continued

Time	Mean estimates (SE), type of game	p Difference (overall)	Mean estimates (SE), type of game			p Differences	Significant differences between each type of game at each time point	p Interaction
			Poker (n = 102)	VLT (n = 76)	Casino (n = 34)			
36 mo	5.0 (0.3)		4.6 (0.4)	5.1 (0.4)	6.4 (0.6)	5.2 (0.5)	4.4 (0.6)	
Overall life satisfaction (quality of life inventory)								
Pre	0.3 (0.1)	<0.01	0.4 (0.2)	0.2 (0.2)	-0.2 (0.3)	0.5 (0.2)	0.4 (0.3)	0.286
Post	1.3 (0.1)		1.5 (0.2)	1.3 (0.2)	0.9 (0.3)	1.5 (0.2)	1.5 (0.3)	
6 mo	1.6 (0.1)		1.7 (0.2)	1.5 (0.2)	1.2 (0.3)	1.8 (0.2)	1.7 (0.3)	
18 mo	1.6 (0.1)		1.7 (0.2)	1.5 (0.2)	1.2 (0.3)	1.8 (0.2)	1.7 (0.2)	
36 mo	1.4 (0.1)		1.5 (0.2)	1.3 (0.2)	1.0 (0.3)	1.6 (0.2)	1.5 (0.3)	

Note. VLT, video lottery terminal.

Table 2. *Factors associated with responder status at post-treatment (n = 218)*

	Number of individuals	Responder (%)	ORs, 95 CI ( <i>p</i> value)			
			Unadjusted		Adjusted	
<b>Work-life satisfaction</b>						
Very poor	46	30.4	Ref.	( <i>p</i> = 0.09)		( <i>p</i> = 0.42)
Poor	51	35.3	1.2	0.5–2.9	1.0	0.4–2.7
Moderate	67	37.3	1.4	0.6–3.0	1.3	0.5–3.2
Good	54	53.7	2.7	1.2–6.0	2.0	0.7–5.6
<b>Primary gambling</b>						
Poker	79	38.0	Ref.	( <i>p</i> = 0.38)		( <i>p</i> = 0.57)
VLT	56	32.1	0.8	0.4–1.6	0.5	0.2–1.2
Casino	25	36.0	0.9	0.4–2.3	1.1	0.3–3.2
Betting	33	51.5	1.7	0.8–3.9	0.8	0.3–2.3
Bingo	25	48.0	1.5	0.6–3.7	1.1	0.3–3.2
<b>Debts due to gambling</b>						
No	58	37.9	Ref.	( <i>p</i> = 0.80)		( <i>p</i> = 0.75)
Yes	158	39.9	1.1	0.6–2.0	0.9	0.4–1.9
<b>Social support</b>						
No	65	32.3	Ref.	( <i>p</i> = 0.24)		( <i>p</i> = 0.29)
Moderate	100	40.0	1.4	0.7–2.7	1.8	0.8–4.0
Yes	50	48.0	1.9	0.9–4.1	1.7	0.7–4.2
<b>Personal yearly salary (Swedish krona)</b>						
0–150,000	48	37.5	Ref.	( <i>p</i> = 0.85)		( <i>p</i> = 0.54)
151,000–210,000	52	44.2	1.3	0.6–2.9	1.5	0.5–4.0
211,000–300,000	57	36.8	1.0	0.4–2.2	0.9	0.3–2.3
301,000 or greater	53	41.5	1.2	0.5–2.6	0.7	0.3–2.1
<b>Number of standard drinks during a typical drinking day</b>						
0–2 drinks	51	54.9	2.7	1.4–5.4	2.7	1.2–6.3
3–4 drinks	48	35.4	1.2	0.6–2.5	1.3	0.6–2.9
5 drinks or more	103	31.1	Ref.	( <i>p</i> = 0.02)		( <i>p</i> = 0.06)
<b>Stage of change</b>						
Contemplation	173	37.0	Ref.	( <i>p</i> = 0.14)		( <i>p</i> = 0.31)
Action	38	50.0	1.7	0.8–3.5	1.5	0.7–3.5
<b>Dissociative gambling</b>						
Never	18	66.7	7.0	1.6–30.8	9.8	1.4–65.8
Sometimes	83	43.4	2.7	0.8–8.8	3.0	0.7–12.7
Often	93	33.3	1.7	0.5–5.8	1.9	0.5–8.2
Always	18	22.2	Ref.	( <i>p</i> = 0.03)		( <i>p</i> = 0.08)

Notes. *The total number of participants does not always add up to 218 for all factors owing to partly missing data. The goodness-of-fit test was positive (Hosmer and Lemeshow:  $p = .74$ ).*

categorised using Swedish norm data (Öst, Breitholtz, & Thulin, 1997). (2) Primary gambling activity was categorised by asking the participant (at screening) what game was most accountable for their problems. If two or more gambling types were reported, the one responsible for the majority of the problem was chosen. (3) Debts due to gambling were categorised as yes or no, regardless of the amount. (4) Social support was defined by the participant answering the question “Do you feel that you have someone available who can

give you the right type of support to help you cope with life’s stresses and problems?” (5) Personal yearly salary was divided into four categories, in an attempt to simulate no income, low income, moderate income, and high income. (6) The number of standard drinks during a typical drinking day was based on the corresponding question in the Alcohol Use Disorders Identification Test (Wennberg, 1996) (“How many drinks containing alcohol do you have on a typical day when you are drinking?”) (7) Stage of change was coded by

analysing the response to the Swedish version of the Readiness to Change questionnaire. (8) Dissociative gambling was coded by analysing the response to a modified version of Jacobs' question now reading "Have you ever completely lost track of time when gambling?" (Jacobs, 1988). Treatment response was defined as having a total of 0 points on the 1-month version of NODS and no incident of gambling over the past 30 days.

### **Statistical analysis**

To test for treatment outcome we used mixed-effects models as suggested by Gueorguieva and Krystal (2004). To investigate possible predictive factors associated with the outcomes, we used logistic regression analysis. Continuous variables were categorised to gain clarity of effect and clinical interpretation. First, we studied crude (unadjusted) association of each factor with the odds of response. Second, to study the adjusted associations for possible confounders, we used multivariable logistic regression models with all factors included. The associations are presented as odds ratios (ORs) with 95% confidence intervals (CIs). There was no transformation or centring with the pre-intervention measurements.

This analysis was done because in the analysis of longitudinal data, repeated observations for the same individual are correlated. If the repeated measures structure correlation is ignored, this may lead to imprecise variation estimates from the regression models, leading to incorrect statistical conclusions. Such correlation violates the assumption of independence necessary for more traditional, repeated-measures analysis and might lead to bias in regression parameters which could lead to wrong conclusions (Brown & Prescott, 1999; Gueorguieva & Krystal, 2004). Furthermore, mixed-effect models are more effective in accommodating missing data and the integration of time-varying factors, which are issues in this study.

When comparing the outcome for different subtypes of problematic games we used a covariance pattern model (Brown & Prescott, 1999), a special case of mixed-effects models. A separate model was estimated for each of the four outcome measures (HADS Anxiety, HADS Depression, NODS, and QOLI). The variance-covariance matrix was assumed to

be block diagonal but unstructured within a block defined by subjects. To study if groups differed across the time points, with respect to the outcome measures, we tested the interaction between time and group. We used the restricted maximum likelihood (REML) as our model estimation method and present estimated means and differences between treatments and their respective standard error (SE) means.

To evaluate the ability of the predictor model to correctly discriminate between responders vs. non-responders, we calculated the area under the receiver operating curve (ROC; trapezoid rule). An area under the curve (AUC) equal to 0.5 suggests no discrimination, 0.7 to <0.8 is acceptable, 0.8–0.9 is excellent, and >0.9 is outstanding discrimination (Hosmer & Lemeshow, 1989). We used the nonparametric method to calculate the SE for the AUC used in the CI and in the comparison of the ROC curves. This gives a conservative estimate of the SE, which implies that the risk of a false difference (type I error) in the evaluation and comparison of the ROC curves is low (Hanley & McNeil, 1982). Hosmer and Lemeshow's goodness-of-fit test was used to examine whether the adjusted models adequately fitted the data, and a  $p$  value >.05 indicates acceptable fit.

Analyses were done in SPSS 18.0 (SPSS, Inc., Chicago, IL, USA).

## **Results**

### **Treatment outcome**

The effects of the treatment are presented in three ways. First, we examine whether the treatment had any effect if type of gambling problem was not taken into account and the participants were analysed as a whole. Then the participants are divided into five subgroups based on the principal type of gambling problem and the analysis is repeated to see whether the results persist. Finally, we analysed whether there was any interaction between the type of gambling problem and the treatment outcome.

As shown in the first four columns of Table 1, there were statistically significant main effects of time on all outcome measurements when analysing the whole group, with statistically significant improvements in all four outcomes from pre- to post-intervention.



This suggests that the treatment might have a significant positive impact on the participants and that improvements were maintained after 6, 18, and 36 months.

One of the objectives of the present study was to see how results would differ from a previous study (Carlbring & Smit, 2008) by including participants with more severe depression (MADRS-S score of 21 or higher). For the participants with more severe depression, the changes in outcome scores (pre-compared to post-intervention) for the gambling-specific scale NODS ( $n(241) = .93$ ;  $p = .352$ ) and for the anxiety measure (HADS-A;  $n(241) = 1.58$ ;  $p = .115$ ) were similar to those for participants with less severe depression. However, and fully consistent with the regression to the mean effect, the more severely depressed group had a significantly larger change in depression score (HADS-D;  $n(241) = 3.04$ ;  $p = .003$ ) compared to the less severely depressed. The more severely depressed group also benefited more in terms of quality of life (QOLI;  $n(241) = 2.43$ ;  $p = .016$ ).

Since there was no significant interaction between time and gambling problem, the conclusion drawn is that there is no evidence that the type of gambling activity affects the treatment outcome (see Table 1). However, there is a significant main effect with regard to time on NODS. This is explained by the fact that there was a significant difference between those who gambled at VLTs, those who bet money on sports and those who played bingo.

This initial difference in the NODS score was then maintained throughout over all measuring points. For those who were primarily casino gamblers, there was a similar trend in depression scores. What was common throughout was that none of these significant differences between the types of gambling are clinically relevant, since all effect sizes were insignificant (Cohen's  $d < .20$ ).

### Prediction

As evident from Table 2, responder status at post-treatment was significantly associated with the background variables of alcohol consumption and dissociative gambling in the unadjusted analysis. Specifically, consumption of three or more standard drinks during a typical drinking day reduced the ratio of responder status to half compared with those who typically consume two drinks or less. Therefore, low alcohol consumption was associated with a higher response rate. In the adjusted analysis only trends were identified ( $p = .056$  and  $p = .075$ , respectively).

As evident from Table 3, three variables have a significant individual predictive ability: work-life satisfaction, number of standard drinks during a typical drinking day, and dissociative gambling. The type of primary problematic game showed only a trend ( $p = .054$ ). However, the AUC for each of these individual variables was relatively low (.58–.61).

Table 3. Evaluation of the ability to predict responder status at post-treatment for different predictor variables and models

Test result variable(s)	Area	SE <sup>a</sup>	Asymptotic sig. <sup>b</sup>	Asymptotic 95% CI	
				Lower bound	Upper bound
1. Work-life satisfaction	0.596	0.042	0.023	0.515	0.678
2. Primary gambling	0.582	0.042	0.054	0.500	0.663
3. Debts due to gambling	0.510	0.042	0.812	0.427	0.593
4. Social support	0.548	0.042	0.257	0.466	0.630
5. Personal yearly salary	0.526	0.042	0.547	0.442	0.609
6. Standard drinks	0.597	0.042	0.023	0.514	0.679
7. Stage of change	0.548	0.043	0.262	0.464	0.631
8. Dissociative gambling	0.609	0.041	0.010	0.528	0.691
Model with all eight predictors	0.718	0.038	0.000	0.643	0.793

Note. Results are expressed by the area under the ROC.

<sup>a</sup> Under the nonparametric assumption.

<sup>b</sup> Null hypothesis: true area = .5.

As evident from Table 3, the model with all eight variables resulted in an acceptable predictive ability to identify responders at post-treatment ( $AUC = .72$ ,  $p < .01$ ). However, similar results were obtained with a model with four variables (work-life satisfaction, primary gambling, dissociative gambling, and number of standard drinks during a typical drinking day;  $AUC = .69$ ), and a model with only two variables (dissociative gambling and number of standard drinks during a typical drinking day;  $AUC = .67$ ). Furthermore, there were no significant differences between the three models since the 95% CI for AUC all overlap. In conclusion, the number of standard drinks during a typical drinking day and dissociative gambling have the strongest association with responder status at post-treatment and have a marginal predictive ability in this sample ( $AUC = .67$ ).

Table 4 shows how responder status was associated with the prediction variables at 36-month follow-up. Only one variable, debts due to gambling, was significantly associated with a positive response. Both the unadjusted and the adjusted analyses showed that the response ratio was about twice as high for individuals with debt compared with the reference category of not having a debt. As shown in Table 5, the proportion of responders was relatively evenly distributed. No single variable had any significant predictive ability to discriminate between responder vs. non-responder at 36 months—not even debts due to gambling ( $AUC = .57$ ,  $p = .129$ ). However, a model with all eight variables together rendered an acceptable predictive ability to identify responders ( $AUC = .70$ ,  $p < .01$ ). Table 5 shows that an eight-variable model has a greater chance of identifying a responder.

## Discussion

The purpose of this study was twofold—to investigate whether the treatment was effective when participants with symptoms of depression were included and whether treatment response could be predicted by a set of pre-treatment characteristics. First, the intervention showed promise both immediately following the treatment termination and up to the 36-month follow-up. Second, it was possible to predict who will benefit from Internet-based treatment

both at post-test and at 36 months by using a model of eight background variables.

While the results point to the value of guided ICBT for pathological gambling, we cannot be certain since no control group was included. This is a weakness in the design as it opens up the possibility that the observed effects are better explained by regression to the mean or spontaneous remission. Indeed, there are reports of significant proportions of participants who naturally recover (Petry, 2005; Slutske, 2006). On the other hand, when the same treatment was compared to a waitlist control in a similar, controlled study (Carlbring & Smit, 2008), there were statistically significant changes as a result of the intervention. The changes observed in the present study are similar in magnitude to those seen previously for all four outcomes (gambling, anxiety, depression, and quality of life), indicating that the effects are probably not due to regression to the mean.

There was an important difference between the population in the previous study, which did not include subjects with a score higher than 21 on the MADRS-S, and the population in this study which did include such subjects. When these populations were analysed separately in the present study, the intervention had a similar effect in more severely depressed participants as for those with less severe depression for gambling and anxiety, but a greater effect for depression and quality of life. These results should be considered in light of the fact that they are not as statistically robust since they were shown in a subgroup of the main study population. Furthermore, the difference in depression score is likely to be an effect of regression to the mean, since these subjects had higher depression scores pre-intervention. Hence, future studies could investigate whether participants meeting DSM criteria for comorbid major depression really benefit more from the intervention.

Interestingly, there were significant differences in depression levels and NODS points between the subgroups of gamblers depending on what game they played, but that did not appear to affect outcome. Since type of game did not interact with treatment outcome, we conclude that no subgroup benefited more from the Internet-based treatment. This is also reflected by the overlapping ORs in the prediction analysis. It seems that other

Table 4. *Factors associated with responder status at 36-month follow-up (n = 196)*

	Number of individuals	Responder (%)	ORs, 95 CI ( <i>p</i> value)			
			Unadjusted		Adjusted	
<b>Work-life satisfaction</b>						
Very poor	43	37.2	Ref.	( <i>p</i> = 0.39)		( <i>p</i> = 0.29)
Poor	45	37.8	1.0	0.4–2.4	0.8	0.3–2.1
Moderate	60	38.3	1.0	0.5–2.4	0.9	0.4–2.4
Good	48	52.1	1.8	0.8–4.2	1.9	0.7–5.1
<b>Primary gambling</b>						
Poker	75	38.7	Ref.	( <i>p</i> = 0.78)		( <i>p</i> = 0.40)
VLT	51	45.1	1.3	0.6–2.7	1.8	0.8–4.3
Casino	27	48.1	1.5	0.6–3.6	1.2	0.4–3.5
Betting	27	33.3	0.8	0.3–2.0	0.6	0.2–1.9
Bingo	16	43.8	1.2	0.4–3.7	1.5	0.4–5.8
<b>Debts due to gambling</b>						
No	51	29.4	Ref.	( <i>p</i> < 0.05)		( <i>p</i> < 0.05)
Yes	143	45.5	2.0	1.0–3.9	2.2	1.0–4.9
<b>Social support</b>						
No	59	35.6	Ref.	( <i>p</i> = 0.09)		( <i>p</i> = 0.20)
Moderate	89	38.2	1.1	0.6–2.2	1.1	0.5–2.3
Yes	45	55.6	2.3	1.0–5.0	2.2	0.9–5.7
<b>Personal yearly salary (Swedish krona)</b>						
0–150,000	47	40.4	Ref.	( <i>p</i> = 0.88)		( <i>p</i> = 0.90)
151,000–210,000	46	41.3	1.0	0.5–2.4	0.7	0.3–2.1
211,000–300,000	51	47.1	1.3	0.6–2.9	1.1	0.4–2.8
301,000 or greater	43	39.5	1.0	0.4–2.2	0.9	0.3–2.7
<b>Number of standard drinks during a typical drinking day</b>						
0–2 drinks	44	36.4	Ref.	( <i>p</i> = 0.49)		( <i>p</i> = 0.34)
3–4 drinks	39	35.9	1.0	0.4–2.4	1.0	0.3–2.7
5 drinks or more	98	44.9	1.4	0.7–3.0	1.7	0.7–3.8
<b>Stage of change</b>						
Contemplation	160	39.4	Ref.	( <i>p</i> = 0.16)		( <i>p</i> = 0.94)
Action	30	53.3	1.8	0.8–3.9	1.5	0.6–3.7
<b>Dissociative gambling</b>						
Never	14	42.9	1.2	0.3–5.4	1.7	0.2–12.3
Sometimes	76	38.2	1.0	0.3–3.1	1.2	0.3–4.9
Often	85	44.7	1.3	0.4–4.0	1.1	0.3–4.5
Always	16	37.5	Ref.	( <i>p</i> = 0.84)		( <i>p</i> = 0.95)

Notes. *The total number of participants does not always add up to 196 for all factors owing to partly missing data. Hosmer and Lemenshow:  $\chi^2 = .65$ .*

variables are more important for the outcome rather than the type of game played.

The results from the logistic regression suggested that it is possible to predict who will respond to ICBT both at post-intervention and at 36 months using a model with eight variables. The single most important background variables vary over time from alcohol consumption and dissociative gambling at post-intervention to whether one had a gambling debt or not at the 3-year follow-up. While the results are interesting, it is still unclear whether the model is valid for other

treatment modalities (e.g. group treatment and individual psychotherapy). The fact that having a gambling debt increased the ratio of later being a treatment responder is not consistent with what has been found previously. However, in a study by Ingle, Marotta, McMillan, and Wisdom (2008), having a small gambling debt was better than not having one. It is possible that the dichotomy in this study can account for the difference, but it is more likely that the two studies used different time frames. Consistent with the Ingle et al's study was that having

Table 5. Evaluation of the ability to predict responder status at post-treatment for different predictor variables and models

Test result variable(s)	Area	SE <sup>a</sup>	Asymptotic sig. <sup>b</sup>	Asymptotic 95% CI	
				Lower bound	Upper bound
1. Work-life satisfaction	0.558	0.045	0.189	0.471	0.646
2. Primary gambling	0.566	0.044	0.135	0.481	0.652
3. Debts due to gambling	0.567	0.044	0.129	0.482	0.653
4. Social support	0.577	0.044	0.084	0.490	0.664
5. Personal yearly salary	0.529	0.045	0.516	0.441	0.616
6. Standard drinks	0.552	0.044	0.242	0.466	0.638
7. Stage of change	0.532	0.045	0.478	0.444	0.619
8. Dissociative gambling	0.526	0.044	0.552	0.440	0.613
Model with all eight predictors	0.702	0.041	0.000	0.622	0.781

Note. Results are expressed by the area under the ROC.

<sup>a</sup> Under the nonparametric assumption.

<sup>b</sup> Null hypothesis: true area = .5.

social support seemed to have a positive influence.

A major issue in prediction analysis is how the outcome variable is defined. In this study, it was defined as having 0 points on the 1-month version of NODS and not having gambled at all over the past 30 days. It is possible that the results would have been different had the response been defined, for example, using the Jacobson and Truax formula for clinical significant improvement (Jacobson & Truax, 1991). Also, the results might be specific to different subgroups. This trial grouped together participants regardless whether they, for example, had any comorbid psychiatric disorder. Furthermore, it could be argued that other predictor variables could be important, such as sex, age, whether or not a person is gambling primarily for sensations seeking or to escape from anxiety, and the reason for seeking help (Dowling, 2009; Grant, Kim, Hollander, & Potenza, 2008; Jamieson, Mazmanian, Penney, Black, & Nguyen, 2011).

In summary, the results from this study generally provide evidence for the continued use and development of guided Internet-delivered self-help with telephone support for pathological gamblers. In addition, this study was probably the first to suggest a model for predicting treatment outcomes for pathological gambling by using guided Internet-based treatment. However, the results need to be replicated to ensure that this is not only sample-

specific associations. In addition, mechanisms of change needs to be addressed (*cf.* Andersson, Carlbring, Berger, Almlöv, & Cuijpers, 2009).

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