

Original Research Article

No Associations Found between *PGBD1* and the Age of Onset in Japanese Patients Diagnosed with Sporadic Alzheimer's Disease

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Key Words

PGBD1 · Alzheimer's disease · Polymorphism · Age of onset

Abstract

Background/Aims: PiggyBac transposable element derived 1 (*PGBD1*) encodes a molecule involved in epigenetic mechanisms that have been implicated in Alzheimer's disease (AD), and recent genome-wide association studies and meta-analyses have indicated that a single nucleotide polymorphism (SNP), rs3800324, in *PGBD1* could be associated with AD and the age of onset. However, no Japanese patients were examined in these studies. The aim of the present study was to replicate the previous finding in Japanese AD cases. **Methods:** We performed a case-control study (211 cases and 156 controls) to investigate the association between *PGBD1* and Japanese AD using 4 tag SNPs including rs3800324. **Results:** Single SNP and haplotype analysis showed no association between AD and age of onset, whereas genotypic and allelic frequencies of the ϵ 4 of apolipoprotein E (*APOE*) showed an association with AD as expected. **Conclusion:** In Japanese AD, we observed no influence of *PGBD1*, as either a risk factor or a modifier, even though *APOE* was associated with AD in this population.

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Introduction

Several studies have investigated the etiology of Alzheimer's disease (AD) and conducted genetic analyses to identify specific candidate genes. Mutations in amyloid- β precursor protein (A β PP), presenilin-1 (PS1), and presenilin-2 (PS2) genes are well-known genetic causes for familial AD [1]. However, for sporadic and late-onset AD, although the ϵ 4 of apolipoprotein E (*APOE*) is a robust genetic risk factor for AD [2] with a gene-dose effect [3] and a relationship with the age of onset [3, 4], it cannot explain all aspects of the onset of AD [3]. Indeed, parallel to findings from genetic studies, environmental factors have also been implicated to play an integral role in AD. In addition to genetic candidates, environmental factors have also arisen as risk factors for AD, including diabetes mellitus [5], smoking [6], hypertension [7, 8], and low levels of nutritional antioxidants (e.g., vitamin C and E) [9]. Previously, epidemiological factors such as epigenetic mechanisms have also been suggested to play a role in AD [10, 11]. Epigenetics has been implicated in several complex diseases. Hypotheses involving epigenetic mechanisms, which do not involve alterations in the DNA sequence, have been suggested regarding changes in gene expression and protein activity or as additive factors for the genetic risk in AD patients. Previously, a relationship between DNA methylation, one type of epigenetic mechanisms, and AD was reported [12]. DNA methylation is also involved in retroposons, another type of epigenetic mechanisms [13, 14]. PiggyBac transposable element derived 1 (*PGBD1*) is a member of one of the families of transposases related to transposons, and belongs to the subfamily of PGBD genes. This gene product is specifically expressed in the brain, but its exact function is still unknown. A genome-wide association study of family-based case-control studies suggested that *PGBD1* is involved in AD [15, 16]. In addition, although weak (the minor allele was associated with an age of onset 1 year earlier than the major allele), a single nucleotide polymorphism (SNP) in *PGBD1*, rs3800324, has been shown to be associated with the age of onset in AD, as a risk modifier in a meta-analysis study [17]. However, no Japanese AD patients were examined in any of these studies. Thus, *PGBD1* may be related to an important aspect of the pathophysiology of AD involving an epigenetic mechanism, especially retroposons.

In the present study, we performed a case-control study to investigate whether *PGBD1* is associated with AD using Japanese common tag SNPs, and also if *PGBD1* could modulate the age of AD onset considering the *APOE* genotype status.

Patients and Methods

Our samples from sporadic Japanese AD patients (n = 211, male:female = 98:113) were obtained from the Department of Psychiatry, Juntendo University Hospital, Tokyo, Japan, and the Department of Psychiatry, Juntendo Koshigaya Hospital, Saitama, Japan. The control samples (n = 156, male:female = 74:82) were obtained from healthy volunteers from our hospital staff who had no history of dementia or other neuropsychiatric diseases. All AD cases were diagnosed according to the NINCDS-ADRDA criteria [18], and none had a familial history of AD. The mean age (\pm SD) of the AD group (68.1 \pm 10.8 years) was not significantly different from that of the control group (62.1 \pm 8.1 years) with the Student t test (t = 5.87, p > 0.05). Additionally, the distribution of males and females within the two groups was not significantly different (χ^2 = 0.05, p > 0.05). The previous study reported that the minor allele of rs3800324 in *PGBD1* is associated with an age of onset only 1 year earlier than the major allele [17]. Therefore, to accurately investigate the relationship between the age of onset and *PGBD1*, only the apparent age of onset (accurate within less than 1 year)

obtained from information from the patients' families (age when any criteria of NINCDS-ADRDA had appeared) from 67 present AD cases (60.5 ± 10.1 years; range 44–81) was used for analysis. An accurate age of onset, with an error less than 1 year, was not obtained from the remaining 144 patients, and these were excluded from the analysis regarding the age of onset.

Written informed consent was obtained from all participants after the procedures had been fully explained in detail to each patient and his/her family. The present study was carried out in compliance with the World Medical Association's Declaration of Helsinki and was approved by the Ethics Committee of the Juntendo University School of Medicine.

Genomic DNA was extracted from peripheral white blood cells using a QIAamp® DNA Blood Maxi kit (Qiagen, Courtaboeuf, France). For the selection of SNPs, tag SNPs for *PGBDI* [$r^2 > 0.8$, minor allele frequencies (MAF) > 0.05] were chosen from the International HapMap Project database (release 27, Phase II + III, February 2009, on NCBI B36 assembly, dbSNP b126) using the TAGGER algorithm with a successful TaqMan probe design. Additionally, a reported missense SNP, rs3800324 (G>A, Gly244Glu) showing a MAF > 0.05 within a Japanese population, was also chosen as a candidate SNP for *PGBDI* from the dbSNP database (<http://www.ncbi.nlm.nih.gov/SNP/>). Thus, the following 4 SNPs were investigated for *PGBDI*: rs2142731, rs3800324, rs3800327, and rs2142730 (the 'rs' notation in front of each SNP represents the identification from the US National Center for Biotechnology Information SNP cluster within the dbSNP database; <http://www.ncbi.nlm.nih.gov/SNP/>). All of these SNPs were analyzed using TaqMan® technology (Assay-by-Design™). Excluding the missense SNP rs3800324, all other SNPs were intronic tag SNPs (genomic map and structures of the human *PGBDI* gene, including the location of the SNPs, were shown in our previous study [19]). All investigated SNPs were typed using TaqMan technology with an ABI7500 system (Applied Biosystems, Foster City, Calif., USA). All probes and primers were designed by the Assay-by-Design service for Applied Biosystems. Polymerase chain reaction (PCR) was done with the standard PCR MasterMix reagent kit using a 4- μ l volume. Additionally, to ensure the quality of the results, we confirmed the SNPs from a few randomly chosen individuals using a direct DNA sequencing method to check for errors in the TaqMan method. All genotypes determined by direct sequencing were in agreement with the genotypes obtained by the TaqMan method for all investigated SNPs. Detailed information on PCR conditions and primer pairs is available upon request. *APOE* genotypes for all samples were determined according to a previous report [20].

Differences in age between healthy controls and patients were examined using two-tailed Student's t test, and the χ^2 test was used to examine differences in gender distribution. For the case-control association study, Hardy-Weinberg equilibrium (HWE) testing of the SNPs was done using SNPAllyze V7.0 Pro (Dynacom, Chiba, Japan). The HWE tests were carried out for all loci in patients and controls. Differences in genotypic and allelic frequencies were evaluated based on the case-control design using χ^2 analysis. Linkage disequilibrium (LD), denoted as D' , was calculated from the haplotype frequency using the expectation-maximization algorithm. The LD block was also identified using SNPAllyze software when D' was greater than 0.75. Case-control haplotype analysis was also performed using SNPAllyze software. Permutation analysis was used to determine the empirical significance and to calculate the p values based on 10,000 replications. The global p values represent the overall significance using the χ^2 test when the observed versus expected frequencies of all haplotypes are considered together. Individual haplotypes were tested for association by grouping all other haplotypes together and applying the χ^2 test with 1 d.f. All p values reported are two-tailed. Statistical significance was defined at $p < 0.05$.

We performed power calculations using the Power Calculator (<http://www.sph.umich.edu/csg/abecasis/CaTS/>). Power was calculated under the prevalence of 0.05 using an additive or a multiplicative model, based on allelic frequencies of the associated markers ranging from 0.22 (rs2142731) to 0.44 (rs3800327), with odds ratios (ORs) ranging from 1.057 (rs2142731) to 1.305 (rs2142730) for the SNPs investigated and an alpha level of 0.05.

To analyze interactions between SNPs in *PGBD1* and *APOE* genotypes, logistic regression analysis was performed using HealthSketch version 2.5 (Dynacom). The probability (P_c) of an individual being a case rather than a control is assumed to be affected by a set of SNPs, according to the logistic model. For example, $\text{logit}(P) = b_0 + b_1x_1 + b_2x_2$ for a single SNP. Here, we use a coding scheme $x_1 = -1, 0, 1$ and $x_2 = -0.5, 0.5, -0.5$ for genotypes homozygous for major alleles, heterozygous, and homozygous for minor alleles (for *APOE*, $\epsilon 4$ was assumed to be the minor allele, and $\epsilon 2$ and $\epsilon 3$ were assumed to be major alleles simultaneously), respectively, to represent an additive effect with x_1 and a dominant/recessive effect with x_2 . The weights were estimated with the maximum likelihood method and tested by comparison with the null hypothesis $\text{logit}(P_c) = b_0$ (constant). We performed a stepwise forward selection procedure with two purposes: the first was to test the relationships among multiple SNPs in *PGBD1* and the *APOE* genotype, and the second was to confirm the interaction among the SNPs if a significant association with $p < 0.05$ was shown with individual logistic models. Multiple regression analyses were performed to estimate the relationship among the age of onset, *APOE* status, and the 4 SNPs using SPSS software ver. 17.0 for Windows (IBM, Chicago, Ill., USA).

Results

Ultimately, 4 SNPs in *PGBD1* were genotyped, and the *APOE* genotype status was obtained in 211 patients with AD and 156 controls, with a genotyping completeness that ranged from 99.5 to 99.9%. Results for power analyses demonstrated that the power ranged from 14% (rs3800324) to 84% (rs2142730). No deviations from HWE in either patient or control samples were observed (all $p > 0.05$; table 1).

No SNPs in *PGBD1* showed any significant association between their allelic or genotypic frequencies and AD (table 1). $D' > 0.75$ was assumed to represent a strong LD, and results indicated that all investigated SNPs for each gene displayed a strong LD block in controls and patients with AD (all $D' > 0.95$) as we previously reported in Japanese controls [19]. Additionally, case-control haplotype association analyses using windows of 2, 3, or 4 SNPs were performed (minor haplotypes with frequencies less than 3% in both the AD cases and controls were omitted), and no haplotype analysis showed a significant association with AD (table 1). As expected, genotypic and allelic frequencies of *APOE* showed a significant association with AD, with an increased frequency of $\epsilon 4$ in AD (table 1).

We further evaluated the relationships and interactions among the 4 SNPs in *PGBD1* and the *APOE* genotype using logistic regression analysis, which can test for a combinatorial effect of multiple SNPs and confirm interactions among the 4 SNPs in *PGBD1* and the *APOE* genotype with a stepwise procedure. We did not identify any significant interactions between any combinations of SNPs. Again, only the additive and dominant/recessive models for *APOE* showed a significant association with the onset of AD (OR 2.11 and 1.96, respectively, all $p < 0.01$ based on the Wald test).

Multiple regression analyses did not show any relationship among the age of onset, *APOE* status, and/or the 4 SNPs in *PGBD1* (all $p > 0.05$).

Table 1. Distribution and statistical analysis of the *PGBD1* polymorphisms and their 2, 3, and 4 SNP-based haplotype analyses with *APOE* genotype distribution

a *PGBD1* polymorphisms and their 2, 3, and 4 SNP-based haplotype analyses

	Genotype frequency, n (%)			p value	HWE C*/AD**	Allele frequency n (%)		χ^2	p value	OR (95% CI)	Haplotype analysis (global p value)		
	A/A	A/G	G/G			A	G				2 SNP- based	3 SNP- based	4 SNP- based
rs2142731													
AD	128 (60.7)	74 (35.1)	9 (4.3)	0.858	0.472/0.831	330 (78.2)	92 (21.8)	0.095	0.758	1.057 (0.788–1.502)	0.718		
C	91 (58.3)	59 (37.8)	6 (3.8)			241 (77.2)	71 (22.7)						
rs3800324													
AD	91 (43.1)	95 (45.0)	25 (11.8)	0.930	0.857/0.900	277 (65.6)	145 (34.4)	0.146	0.703	1.062 (0.779–1.448)	0.525		
C	70 (44.9)	69 (44.2)	17 (10.9)			209 (67.0)	103 (33.0)						
rs3800327													
AD	62 (29.4)	112 (53.1)	37 (17.5)	0.781	0.602/0.330	236 (55.9)	186 (44.1)	0.425	0.515	1.103 (0.821–1.483)	0.405		
C	51 (32.7)	80 (51.3)	25 (16.0)			182 (58.3)	130 (41.7)						
rs2142730													
AD	102 (48.3)	95 (45.0)	14 (6.6)	0.118	0.871/0.254	299 (70.9)	123 (29.1)	2.781	0.095	1.305 (0.954–1.786)			
C	67 (42.9)	69 (44.2)	20 (12.8)			203 (65.1)	109 (34.9)						

b *APOE* genotype distribution

	Genotype frequency, n (%)						p value	Allele frequency, n (%)			χ^2	OR ¹ (95% CI)	p value
	$\epsilon 2/\epsilon 2$	$\epsilon 2/\epsilon 3$	$\epsilon 2/\epsilon 4$	$\epsilon 3/\epsilon 3$	$\epsilon 3/\epsilon 4$	$\epsilon 4/\epsilon 4$		$\epsilon 2$	$\epsilon 3$	$\epsilon 4$			
AD	1 (0.5)	4 (1.9)	3 (1.4)	117 (55.5)	66 (31.3)	20 (9.5)	0.0017	9 (2.1)	304 (72.0)	109 (25.8)	17.4	2.177 (1.477–3.215)	0.0001
C	4 (2.6)	5 (3.2)	0 (0)	106 (67.9)	39 (25.0)	2 (1.3)		13 (4.2)	256 (82.1)	43 (13.8)			

C = Controls. p value in bold indicate statistical significance ($p < 0.05$). ¹ $\epsilon 4$ vs. others.

Discussion

In the present study, *PGBD1*, a member of one of the families of transposases that may be involved in AD [15, 16], was investigated for an association with Japanese AD. No single SNP and/or haplotype in our case-control analysis of *PGBD1* showed an association with Japanese AD. In addition, although a missense SNP, rs3800324 in *PGBD1*, showed an association with the age of onset in AD as reported previously [17], the association of the same SNP in the present group of Japanese participants was not replicated. That previous study was performed with a sufficiently large number of cases (2,455 Caucasian cases) [17]. We are aware that the main limitation of our present study is the small number of participants, and we should carefully interpret the present negative findings. Our results may have involved a type II error (false negative) because although the influence of the $\epsilon 4$ allele in *APOE* in the present study showed statistical significance, the OR for the $\epsilon 4$ allele in the present AD cases (OR 2.177, 95% C 1.477–3.215) was lower than that observed in a previous study (OR 3.2, 95% CI 2.9–3.5) [21]. This was probably due to the small number of cases. However, although the $\epsilon 4$ genotypic and allelic frequencies were lower than expected, we observed a reproducible association in the present Japanese AD cases. Thus, we conclude that *PGBD1* did not affect the age of onset per se in Japanese AD patients to the same degree as the genetic influence of *APOE*. On the other hand, ethnic differences (Caucasian and Japanese) and differences in

the subtype of AD [late-onset AD in the previous study (mean age approx. 78 years) [17] and both early- and late-onset AD in the present study (mean age \pm SD 60.5 \pm 10.1 years, range 44–81, for the 67 cases with the known age of onset)] may also have led to the different results.

Here, we report our current data that will be used for a future meta-analysis study in Japanese AD patients and conclude that the influence of *PGBD1* as genetic risk factor in disease onset and for the age of onset as a modifier factor is not present in Japanese sporadic AD.

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Disclosure Statement

We have no potential conflicts of interest.

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