

34. Brennan SL, Henry MJ, Nicholson GC *et al.* Socioeconomic status, obesity and lifestyle in men: the Geelong Osteoporosis Study. *J Mens Health* 2010; 7: 31–41
35. Thornton LE, Crawford DA, Ball K. Who is eating where? Findings from the socioeconomic status and activity in women (SESAW) study. *Public Health Nutr* 2011; 14: 523–531
36. Adams RJ, Appleton SL, Hill CL *et al.* Risks associated with low functional health literacy in an Australian population. *Med J Aust* 2009; 191: 530–534
37. Beauchamp A, Peeters A, Wolfe R *et al.* Inequalities in cardiovascular disease mortality: the role of behavioural, physiological and social risk factors. *J Epidemiol Community Health* 2010; 64: 542–548
38. Perna L, Thien-Seitz U, Ladwig K-H *et al.* Socio-economic differences in life expectancy among persons with diabetes mellitus or myocardial infarction: results from the German MONICA/KORA study. *BMC Public Health* 2010; 10.
39. Ward MM. Access to care and the incidence of end-stage renal disease due to diabetes. *Diabetes Care* 2009; 32: 1032–1036
40. Cunningham J. Socio-economic gradients in self-reported diabetes for Indigenous and non-Indigenous Australians aged 18–64. *Aust N Z J Public Health* 2010; 34: S18–S24
41. Pichler RH, de Boer IH. Dual renin–angiotensin–aldosterone system blockade for diabetic kidney disease. *Curr Diab Rep* 2010; 10: 297–305
42. Wolf G, Musch M, Müller N *et al.* Association between socioeconomic status and renal function in a population of German patients with diabetic nephropathy treated at a tertiary center. *Nephrol Dial Transplant* 2011; 26: 4017–4023
43. Overland J, Hayes L, Yue DK. Social disadvantage: its impact on the use of Medicare services related to diabetes in NSW. *Aust N Z J Public Health* 2002; 26: 262–265
44. Bayliss EA, Bhardwaja B, Ross C *et al.* Multidisciplinary team care may slow the rate of decline in renal function. *Clin J Am Soc Nephrol* 2011; 6: 704–710
45. Stewart JH, McCredie MR, Williams SM *et al.* The enigma of hypertensive ESRD: observations on incidence and trends in 18 European, Canadian, and Asian-Pacific populations, 1998 to 2002. *Am J Kidney Dis* 2006; 48: 183–191
46. Patch C, Charlton J, Roderick PJ *et al.* Use of antihypertensive medications and mortality of patients with autosomal dominant polycystic kidney disease: A population-based study. *Am J Kidney Dis* 2011; 57: 856–862

Received for publication: 27.1.2012; Accepted in revised form: 2.7.2012

Nephrol Dial Transplant (2012) 27: 4180–4188

doi: 10.1093/ndt/gfs021

Advance Access publication 19 March 2012

## Longer dialysis session length is associated with better intermediate outcomes and survival among patients on in-center three times per week hemodialysis: results from the Dialysis Outcomes and Practice Patterns Study (DOPPS)

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### Abstract

**Background.** Longer dialysis session length (treatment time, TT) has been associated with better survival among hemodialysis (HD) patients. The impact of TT on clinical markers that may contribute to this survival advantage is not well known.

**Methods.** Using data from the international Dialysis Outcomes and Practice Patterns Study, we assessed the association of TT with clinical outcomes using both standard regression analyses and instrumental variable approaches. The study included 37 414 patients on in-center HD three times per week with prescribed TT from 120 to 420 min.

**Results.** Facility mean TT ranged from 214 min in the

USA to 256 min in Australia–New Zealand. Accounting for country effects, mortality risk was lower for patients with longer TT {hazard ratio for every 30 min: all-cause mortality: 0.94 [95% confidence interval (CI): 0.92–0.97], cardiovascular mortality: 0.95 (95% CI: 0.91–0.98) and sudden death: 0.93 (95% CI: 0.88–0.98)}. Patients with longer TT had lower pre- and post-dialysis systolic blood pressure, greater intradialytic weight loss, higher hemoglobin (for the same erythropoietin dose), serum albumin and potassium and lower serum phosphorus and white blood cell counts. Similar associations were found using the instrumental variable approach, although the positive associations of TT with weight loss and potassium were lost.

**Conclusions.** Favorable levels of a variety of clinical markers may contribute to the better survival of patients receiving longer TT. These findings support longer TT prescription in the setting of in-center, three times per week HD.

**Keywords:** DOPPS; hemodialysis; outcomes; survival; treatment length

## Introduction

The morbidity and mortality rate of patients receiving three times per week hemodialysis (HD) remain unacceptably high [1]. Compared to ‘standard’ dialysis, daily in-center and long nightly home dialysis have been associated with better outcomes and quality of life in small cohorts of selected patients [2–4]. In the recent Frequent Hemodialysis Network trial, six times per week dialysis was associated with favorable outcomes compared to the standard regimen [5]. While extended dialysis schedules may lead to better clinical outcomes, logistical, financial, and other impediments remain for their use for the majority of HD patients.

Most HD patients worldwide receive conventional three times per week dialysis with a duration of <5 h [1]. Even in this setting, shorter dialysis session length (treatment time, TT) has been associated with worse survival [6–11]. While the association of TT with survival is independent of dialysis dose, most prior studies did not provide a mechanistic insight through assessment of the association of TT with clinical markers (e.g. hemoglobin, serum phosphorus, blood pressure) which may contribute to morbidity and mortality in this population. Furthermore, despite the wide range of adjustments and analytic techniques [11], prior studies failed to completely address differences in the health status of patients receiving longer versus shorter TT. The present study highlights international differences in TT, presents associations of TT with intermediate measures and applies an instrumental variable approach to account, in part, for unmeasured confounders that may bias the associations of TT with clinical outcomes.

## Materials and methods

### Data sources

The Dialysis Outcomes and Practice Patterns Study (DOPPS) is a prospective cohort study of in-center HD patients. Details on the study design have been described previously [12, 13].

The current study included data from seven countries (France, Germany, Italy, Japan, Spain, the UK and the USA) in DOPPS 1 (1996–2001) and from an additional five countries (Australia, Belgium, Canada, New Zealand and Sweden) in DOPPS 2 (2002–04) and 3 (2005–08). Selected data are presented within the geographic regions: North America (USA + Canada); Eur/ANZ (European countries + Australia and New Zealand) and Japan. Detailed case-mix and comorbid data were collected at study entry. Cause-specific mortality and hospitalization events were collected during the study follow-up. Informed patient consent was obtained as indicated in accordance with local requirements.

TT was defined as the prescribed HD session length and was analyzed as a continuous variable (per 30 min longer TT) as well as a categorical variable (<200, 200–225, 226–250 and >250 min). Because TT for the great majority of patients was at exactly 30 min intervals, we used 180, 210, 240 and 270–300 min as respective surrogate names for the categories. Mortality included all-cause mortality, cardiovascular death and sudden death (mortality due to cardiac arrhythmia, cardiac arrest or hyperkalemia).

Hospitalizations included all-cause hospitalization, hospitalizations due to cardiovascular events and congestive heart failure or fluid overload. Intermediate outcomes included intradialytic weight loss, pre- and post-dialysis systolic blood pressure (SBP) and laboratory values [hemoglobin, white blood cell count (WBC), serum phosphorus, potassium, albumin and ferritin] measured at study enrollment. Sensitivity analyses were conducted using the delivered TT (i.e. the HD session length as actually received).

### Statistical analysis

Differences in patient characteristics across TT categories were assessed using a test for trends. Linear mixed models for continuous outcomes and the Generalized Estimating Equation method with logit link function for dichotomized outcomes were used to examine the associations between TT and patient intermediate outcomes. Models were adjusted for patient characteristics, DOPPS country and study phase and accounted for facility clustering, assuming a compound symmetry covariance structure. Cox models were used to estimate the associations of TT with mortality/hospitalization risk, were stratified by country and study phase and accounted for facility clustering using robust sandwich estimators. The proportional hazard assumption was tested and satisfied.

In order to partially account for patient-level unmeasured confounders which may impact the relationship between TT and outcomes, we also conducted a separate set of analyses applying an instrumental variable approach that used the dialysis facility as the instrument [14–17]. For patient intermediate outcomes, we conducted the standard two-stage least square instrumental variable method for continuous outcomes and an extended, two-stage instrumental variable method with a linear model as the first stage and logistic regression as the second stage for dichotomized outcomes. For risk of mortality/hospitalization, we used an extended instrumental variable approach that uses a linear model first stage and a Cox model second stage [18]. Since the *F*-statistic in all the first-stage models was > 25, we rejected the null hypothesis of weak instruments with the interpretation that the instrumental variable estimates are less biased [15, 19, 20].

A multiple imputation method was used to correct for potential biases that could be caused by missing values using the standard software IVEware [21]. All statistical analyses were performed using SAS software, version 9.2 (SAS Institute, Cary, NC). The authors have followed the STrengthening the Reporting of OBServational studies in Epidemiology (STROBE) Statement guidelines for reporting observational studies [22].

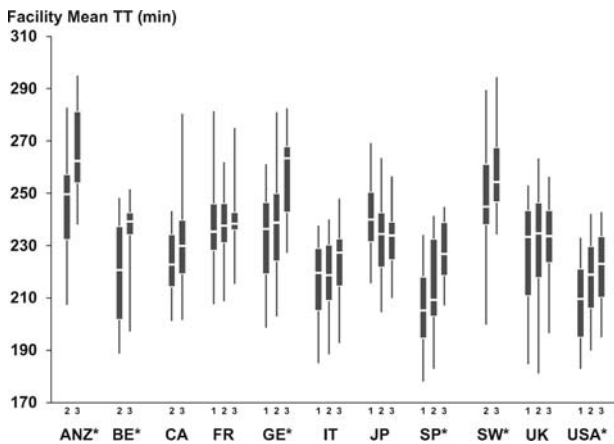
## Results

### Study sample

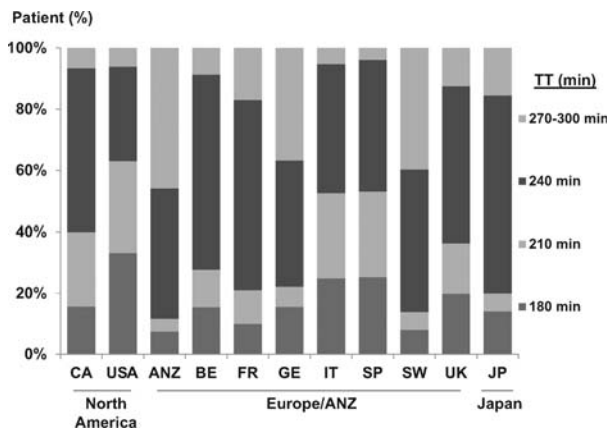
This study included 37 414 patients receiving three HD treatments per week with prescribed TT from 120 to 420 min at study enrollment; 15 442 patients were from 308 facilities in DOPPS 1, 11 553 patients were from 322 facilities in DOPPS 2 and 10 419 patients were from 300 facilities in DOPPS 3. The mean follow-up was 19 months. During the study period, 8961 patients died (mortality rate: 0.15/year).

### Distribution of TT across DOPPS countries and over time

Distributions of facility mean TT (FMTT) by DOPPS country and phase are presented in Figure 1. Large differences in FMTT were observed across countries ( $P < 0.001$ ), with the longest average FMTT ( $256 \pm 23$  min) in ANZ and the shortest in the USA ( $214 \pm 17$  min). Overall, FMTT increased over time from DOPPS 1 to DOPPS 3 ( $P < 0.001$ ). A significant increase in FMTT over time was found in ANZ, Belgium, Germany, Spain, Sweden and the USA ( $P < 0.05$ ), and a significant decrease was found in Japan ( $P = 0.01$ ). Distributions of patient prescribed TT in DOPPS 1–3 by country are presented in Figure 2. Significant differences in the distribution of TT were found



**Fig. 1.** Distribution of facility mean prescribed TT by DOPPS country and by phase. The box shows the 25th–75th and the whiskers the 5th–95th percentile ranges. \* $P < 0.05$  for increase over time. DOPPS Phase 1 (1996–2001), Phase 2 (2002–04) and Phase 3 (2005–08). ANZ, Australia and New Zealand; BE, Belgium; CA, Canada; FR, France; GE, Germany; IT, Italy; JP, Japan; SP, Spain; SW, Sweden; UK, United Kingdom; USA, United States of America.



**Fig. 2.** Distribution of patient-level prescribed TT categories by DOPPS country. DOPPS Phase 1 (1996–2001), Phase 2 (2002–04) and Phase 3 (2005–08). ANZ, Australia and New Zealand; BE, Belgium; CA, Canada; FR, France; GE, Germany; IT, Italy; JP, Japan; SP, Spain; SW, Sweden; UK, United Kingdom; USA, United States of America.

across countries ( $P < 0.001$ ). The prevalence of TT  $< 200$  min was highest in the USA (33.1%) and lowest in Australia/New Zealand (7.5%).

#### Patient characteristics by TT categories

Patient characteristics by prescribed TT at study enrollment within each DOPPS region are shown in Table 1. Patients with longer TT were younger, more likely male, had longer end-stage renal disease (ESRD) duration and higher body weight ( $P < 0.001$  for all). The prevalence of comorbidities within each TT category varied across the DOPPS regions. In all regions, patients with longer TT had higher hemoglobin and serum albumin levels, were less likely to use a catheter as vascular access, had higher blood flow rates and were more likely to be treated with high-flux dialyzers ( $P \leq 0.01$  for all).

#### TT and mortality and hospitalization risk

Table 2 shows that longer TT was associated with lower mortality and hospitalization risk, both in unadjusted and adjusted standard regression models. Because most trends across categories were approximately linear, we also examined TT as a continuous variable. In Figure 3, the adjusted standard regression models show a significantly decreased risk of both mortality [hazard ratio (HR) = 0.94, 95% confidence interval (CI): 0.92–0.97] and hospitalization (HR = 0.97, 95% CI: 0.96–0.99) per 30 min longer TT. Results of the instrumental variable analyses yielded qualitatively similar estimates with (as expected) less precision.

A sensitivity analysis adjusting for patient height and target weight rather than body mass index was consistent (all-cause mortality HR = 0.94, 95% CI: 0.91–0.96). Stratifying by the median target weight (66 kg) provided similar results below (HR = 0.92, 95% CI: 0.89–0.95) and above (HR = 0.95, 95% CI: 0.92–0.98) the median ( $P$  for interaction = 0.42). Additionally, all-cause mortality results were consistent in models that excluded patients (i) using a catheter as vascular access [HR = 0.93 (95% CI: 0.90–0.95) per 30 min longer TT], (ii) with TT  $> 240$  min [HR = 0.95 (95% CI: 0.92–0.98) per 30 min longer TT] or (iii) who had been on dialysis for  $< 12$  months [HR = 0.94 (95% CI: 0.91–0.97) per 30 min longer TT]. In addition to prescribed TT, we evaluated the association of delivered (versus prescribed) TT with outcomes and found consistent results [all-cause mortality HR = 0.92 (95% CI: 0.89–0.95) per 30 min TT]. A sensitivity analysis adjusting for single pool Kt/V rather than blood flow rate attenuated the effect slightly, as expected [all-cause mortality HR = 0.96 (95% CI: 0.93–0.99) per 30 min longer TT], due to collinearity with TT.

A significant interaction effect ( $P < 0.001$ ) was found between TT and DOPPS regions (Figure 4). Longer TT was strongly associated with lower mortality in Japan [HR = 0.75 (95% CI: 0.69–0.81) per 30 min longer TT], still significantly associated with lower mortality in Eur/ANZ [HR = 0.94 (95% CI: 0.91–0.97) per 30 min longer TT], but not associated with mortality in North America [HR = 0.98 (95% CI: 0.95–1.02) per 30 min longer TT]. Similar TT effect on mortality was found in each region after adjusting for non-adherence with dialysis prescription (defined as any skipped dialysis session in the 30 days prior to DOPPS enrollment) [23]. Because Japanese patients receiving  $< 4$  h of TT may be a subset of the sickest patients, an analysis restricted to patients with TT  $\geq 240$  min was conducted and found similar results in Japan [all-cause mortality HR = 0.75 (95% CI: 0.64–0.88) per 30 min longer TT].

Overall, the protective effect of TT on mortality appeared to be most pronounced among patients with low blood flow rate ( $P = 0.02$  for interaction between TT and blood flow rate). However, this finding was likely confounded by region, as Japan had the lowest blood flow rates and the strongest association between TT and mortality. Within each of the three geographic regions, no significant interaction between TT and blood flow rate was found (all  $P > 0.1$ ).

**Table 1.** Patient characteristics by patient TT categories within each DOPPS region

		Patient TT				P-value <sup>a</sup>
		180 min	210 min	240 min	270–300 min	
Number of patients	All regions	8411 (22%)	7282 (19%)	16 795 (45%)	4926 (13%)	
	North America (%)	4924 (32%)	4587 (29%)	5124 (33%)	948 (6%)	
	Eur/ANZ (%)	2556 (17%)	2306 (15%)	7389 (49%)	2950 (19%)	
	Japan (%)	931 (14%)	389 (6%)	4282 (65%)	1028 (16%)	
Demographics						
Age (years)	North America	64.5	62.7	58.7	54.2	<0.0001
	Eur/ANZ	64.3	65.3	62.7	59.1	<0.0001
	Japan	66.3	64.2	60.5	56.5	<0.0001
Sex (male, %)	North America	45	52	63	80	<0.0001
	Eur/ANZ	53	52	59	70	<0.0001
	Japan	58	56	61	69	<0.0001
Vintage (years)	North America	1.9	2.3	2.4	3.7	<0.0001
	Eur/ANZ	2.2	3.8	3.7	5.1	<0.0001
	Japan	1.8	4.1	7.0	10.7	<0.0001
Target weight (kg)	North America	67.6	72.1	79.8	97.3	<0.0001
	Eur/ANZ	65.4	64.6	68.7	77.4	<0.0001
	Japan	50.7	51.3	52.8	55.6	<0.0001
Comorbidities						
Diabetes (%)	North America	47	52	53	58	<0.0001
	Eur/ANZ	28	26	29	34	<0.0001
	Japan	41	43	31	21	0.13
Hypertension (%)	North America	85	87	86	89	0.08
	Eur/ANZ	79	78	78	82	<0.0001
	Japan	75	71	67	56	0.04
Coronary artery disease (%)	North America	54	57	54	58	0.79
	Eur/ANZ	39	38	42	51	0.92
	Japan	27	31	29	25	0.83
Congestive heart failure (%)	North America	45	48	46	47	0.31
	Eur/ANZ	27	27	33	36	<0.0001
	Japan	26	23	17	13	0.36
Labs						
Hemoglobin (g/dL)	North America	10.7	10.9	11.0	11.6	<0.0001
	Eur/ANZ	10.5	11.2	11.2	11.5	<0.0001
	Japan	9.4	10.0	9.9	10.3	<0.0001
Albumin (g/dL)	North America	3.6	3.6	3.6	3.8	<0.0001
	Eur/ANZ	3.6	3.7	3.7	3.8	<0.0001
	Japan	3.6	3.7	3.8	3.9	<0.0001
Dialysis treatment						
Catheter use (%)	North America	42	39	42	29	<0.0001
	Eur/ANZ	31	20	23	14	<0.0001
	Japan	9	2	1	0	<0.0001
Blood flow (mL/min)	North America	363.4	383.9	384.1	414.0	<0.0001
	Eur/ANZ	273.1	307.4	301.9	308.6	<0.0001
	Japan	173.2	192.3	193.9	203.5	<0.0001
High-flux dialyzer use <sup>b</sup> (%)	North America	55	55	50	66	0.01
	Eur/ANZ	33	39	39	51	<0.0001
	Japan	62	74	73	76	<0.0001

<sup>a</sup>Test of trend adjusted for country and phase and accounted for facility clustering.

<sup>b</sup>High flux percent calculated after excluding those with missing flux information (22%).

### TT and intermediate outcomes

Associations between prescribed TT (both categorically and continuously) and intermediate outcomes are shown in Table 3. Longer TT was associated with levels of intermediate outcomes which are generally considered favorable, including higher hemoglobin [for a given erythropoietin (EPO) dose] and serum albumin, lower WBC and phosphorus. Longer TT was associated with greater weight loss and higher potassium levels in the standard regression models (perhaps due to unmeasured confounding by indication), but these associations were lost in the instrumental variable analysis (intended to lessen the

biases resulting from patient-level unmeasured confounders). The associations between TT and achievement of clinical practice targets (most recent at time of manuscript submission) are shown in Figure 5. These associations are in keeping with the findings in Table 3.

### Discussion

The present study examined a large cohort of patients receiving in-center, three times per week maintenance HD at 930 facilities in 12 countries participating in the



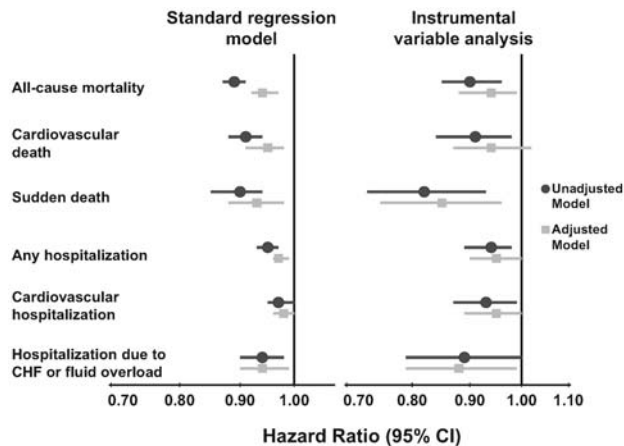
**Table 2.** Associations between prescribed TT categories and risk of mortality/hospitalization<sup>a</sup>

	Categorical TT (standard regression model)			
	180 min	210 min	240 min	270–300 min
All-cause mortality				
Unadjusted <sup>b</sup>	1.30 (1.22–1.40)	1.18 (1.11–1.26)	1.00 (reference)	0.78 (0.71–0.85)
Adjusted <sup>c</sup>	1.16 (1.07–1.24)	1.06 (0.99–1.13)	1.00 (reference)	0.90 (0.83–0.98)
Cardiovascular death				
Unadjusted <sup>b</sup>	1.29 (1.17–1.42)	1.18 (1.07–1.29)	1.00 (reference)	0.87 (0.77–0.98)
Adjusted <sup>c</sup>	1.18 (1.06–1.31)	1.06 (0.96–1.17)	1.00 (reference)	0.97 (0.86–1.10)
Sudden death				
Unadjusted <sup>b</sup>	1.30 (1.13–1.49)	1.10 (0.96–1.26)	1.00 (reference)	0.76 (0.63–0.91)
Adjusted <sup>c</sup>	1.19 (1.02–1.38)	0.99 (0.86–1.14)	1.00 (reference)	0.84 (0.70–1.01)
Any hospitalization				
Unadjusted <sup>b</sup>	1.14 (1.08–1.20)	1.01 (0.96–1.06)	1.00 (reference)	0.93 (0.89–0.98)
Adjusted <sup>c</sup>	1.10 (1.04–1.16)	1.00 (0.94–1.05)	1.00 (reference)	0.99 (0.94–1.05)
Cardiovascular hospitalization				
Unadjusted <sup>b</sup>	1.08 (1.01–1.16)	1.05 (0.98–1.13)	1.00 (reference)	0.99 (0.92–1.07)
Adjusted <sup>c</sup>	1.07 (1.00–1.15)	1.04 (0.97–1.11)	1.00 (reference)	1.01 (0.93–1.10)
Hospitalization due to CHF or fluid overload				
Unadjusted <sup>b</sup>	1.26 (1.11–1.43)	1.12 (0.99–1.27)	1.00 (reference)	0.92 (0.79–1.08)
Adjusted <sup>c</sup>	1.24 (1.09–1.42)	1.10 (0.97–1.24)	1.00 (reference)	0.94 (0.80–1.11)

<sup>a</sup>HRs (95% CI) shown for each outcome. CHF, congestive heart failure.

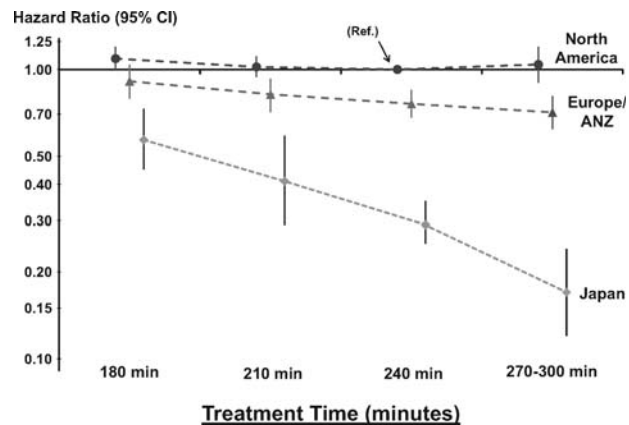
<sup>b</sup>Model stratified by country and study phase and accounted for facility clustering.

<sup>c</sup>Model stratified by country and study phase, adjusted for age, sex, race, time on dialysis, BMI, 13 summary comorbid conditions, residual kidney function, prescribed blood flow rate and catheter use and accounted for facility clustering.



**Fig. 3.** Association between prescribed TT (per 30 min longer) and risks of mortality and hospitalization. Adjusted model: adjusted for age, sex, race, time on dialysis, BMI, 13 summary comorbid conditions, residual kidney function, prescribed blood flow rate and catheter use, stratified by country and phase of study and accounted for facility clustering. CHF, congestive heart failure.

DOPPS (1996–2008). TT prescription varied across countries, with the longest average TT in ANZ ( $255 \pm 41$  min) and the shortest in the USA ( $212 \pm 32$  min). These large differences must be interpreted along with consideration of other clinical practices, such as the use of high-flux dialyzers and delivered dialysis dose (which were both higher in North America). Overall, prescribed TT increased over the study period. The mean TT reported for US DOPPS participants is consistent with those recently reported by two large US dialysis organizations [10, 11]. The trend toward longer TT we observed in the USA was also reported by a US dialysis organization between 1996



**Fig. 4.** Association between prescribed TT and mortality by region. Interaction between TT and region ( $P < 0.0001$ ). Longer TT was associated with lower mortality in Eur/ANZ [HR = 0.94 (95% CI: 0.91–0.97) per 30 min TT,  $P = 0.0002$ ] and Japan [HR = 0.75 (95% CI: 0.69–0.81),  $P < 0.0001$ ] but not in North America [HR = 0.98 (95% CI: 0.95–1.02),  $P = 0.28$ ]. Model was adjusted for age, sex, race, time on dialysis, BMI, 13 summary comorbid conditions, residual kidney function, prescribed blood flow rate and catheter use, stratified by study phase and accounted for facility clustering. The chosen reference category was for North American patients with prescribed TT at 240 min.

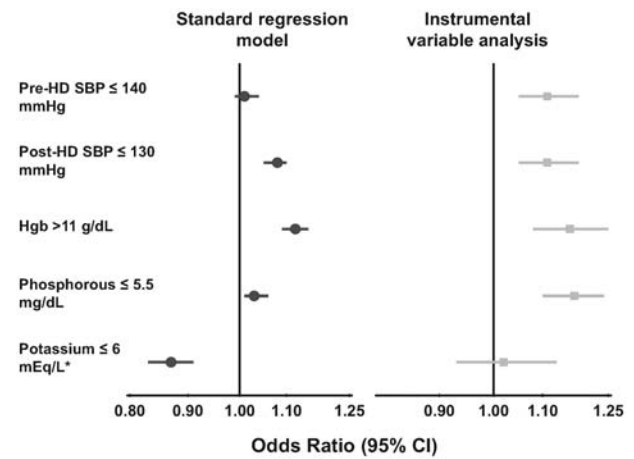
and 2008 [10]. However, the increase in TT over time in both our study (from  $208 \pm 32$  to  $221 \pm 31$  min) and that publication (from  $201 \pm 61$  to  $213 \pm 59$  min) was relatively small and may not have had an impact on clinical outcomes. In fact, among US DOPPS participants in 2005–08, only 12% had a TT  $>250$  min, while 23% were dialyzed for  $<200$  min.

In the present study, patients with longer TT had lower risk of all-cause and cardiovascular mortality. A new interesting finding is the strong association between longer

**Table 3.** Associations between prescribed TT and intermediate outcomes<sup>a</sup>

	Categorical TT (standard regression model)					Continuous TT	
	180 min	210 min	240 min	270-300 min	Standard regression model (per 30 min)	Instrumental variable approach (per 30 min)	
Weight loss (kg)	-0.61 (-0.70 to -0.52)	-0.21 (-0.29 to -0.12)	0.00 (reference)	0.38 (0.28 to 0.47)	0.26 (0.23 to 0.29)	-0.01 (-0.08 to 0.06)	
Pre-HD SBP (mmHg)	0.50 (-0.29 to 1.29)	0.45 (-0.31 to 1.22)	0.00 (reference)	0.21 (-0.67 to 1.09)	-0.29 (-0.56 to -0.01)	-1.35 (-2.09 to -0.61)	
Post-HD SBP (mmHg)	1.32 (0.52 to 2.12)	0.38 (-0.40 to 1.15)	0.00 (reference)	-2.18 (-3.07 to -1.28)	-1.01 (-1.28 to -0.73)	-1.49 (-2.27 to -0.70)	
Hemoglobin (g/dL)	-0.24 (-0.29 to -0.19)	-0.07 (-0.12 to -0.02)	0.00 (reference)	0.19 (0.13 to 0.24)	0.11 (0.09 to 0.13)	0.13 (0.08 to 0.19)	
Albumin (g/dL)	-0.02 (-0.04 to -0.01)	-0.01 (-0.03 to 0.01)	0.00 (reference)	0.05 (0.03 to 0.07)	0.02 (0.01 to 0.02)	0.05 (0.03 to 0.08)	
WBC (1000/mL)	0.10 (0.02 to 0.19)	0.05 (-0.03 to 0.13)	0.00 (reference)	-0.06 (-0.15 to 0.03)	-0.04 (-0.07 to -0.01)	-0.09 (-0.15 to -0.03)	
Ferritin (ng/mL)	-18.8 (-33.3 to -4.3)	-15.6 (-29.0 to -2.3)	0.00 (reference)	1.7 (-13.6 to 17.0)	6.7 (1.7 to 11.7)	17.7 (-2.4 to 37.8)	
Phosphorous (mg/dL)	0.05 (-0.01 to 0.11)	0.00 (-0.06 to 0.06)	0.00 (reference)	-0.03 (-0.09 to 0.04)	-0.04 (-0.06 to -0.02)	-0.16 (-0.22 to -0.11)	
Potassium (mEq/L) <sup>b</sup>	-0.11 (-0.13 to -0.08)	-0.03 (-0.06 to -0.01)	0.00 (reference)	0.10 (0.07 to 0.13)	0.05 (0.04 to 0.06)	0.00 (-0.03 to 0.02)	

<sup>a</sup>Estimate (95% CI) shown is the difference in each outcome associated with prescribed TT categories, compared to the reference category. Models adjusted for country and study phase, age, sex, race, time on dialysis, BMI, 13 summary comorbid conditions, residual kidney function, prescribed blood flow rate and catheter use and accounted for facility clustering.  
<sup>b</sup>Model also adjusted for dialyze K.



**Fig. 5.** Association between 30 min longer prescribed TT and achievement of clinical targets. Clinical targets are based on the Kidney Disease Outcomes Quality Initiative clinical practice guidelines for cardiovascular disease in dialysis patients [24], bone metabolism and disease in chronic kidney disease [25] and anemia [26]. Model adjusted for age, sex, race, time on dialysis, BMI, 13 summary comorbid conditions, residual kidney function, prescribed blood flow rate, catheter use, country and study phase and accounted for facility clustering; \*model also adjusted for dialyze K.

TT and lower risk of sudden death, which remained after adjusting for patient comorbidities (like diabetes and atrial fibrillation) that are risk factors for sudden death [27]. It is likely that the smaller plasma dialyze electrolyte gradients, less dramatic volume shifts and less sympathetic hyperactivity during longer dialysis sessions may contribute to the lower risk of sudden death.

Patients treated with frequent HD may experience lower mortality [5]. The National Cooperative Dialysis Study is the only randomized controlled trial conducted among patients on three times per week HD that assessed the impact of TT on outcomes. Despite a trend toward higher hospitalization risk observed in the short TT arm, no effect of TT on mortality was found [28]. However, the trial was terminated early and thus did not test effect of TT on mortality. Two cohort studies from the early 1990s also failed to find any association between TT and mortality [29, 30]. It is likely that other clinical practices (e.g. dialyzer type) in use at the time these studies were conducted were different than current practices; their findings may not be applicable to the current HD population. Several observational studies have reported higher mortality risk for patients receiving shorter TT [6-11]. Our findings also indicate a higher mortality risk, especially sudden death, among DOPPS participants receiving shorter TT; these findings were confirmed in instrumental variable analyses based on the premise that patients are ‘assigned’ to dialysis facilities that prescribe different TTs on average. While Table 1 shows that patients prescribed a longer TT are generally healthier overall, the instrumental variable analysis results reduce the biases resulting from unmeasured patient-level confounding and still show a significant survival benefit of longer TT.

As reported in a prior DOPPS analysis [8], the association of TT with mortality differed across geographic

region, being the strongest in Japan, intermediate in Eur/ANZ and no longer significant in North America. This finding is consistent with a recent analysis of US HD patients that reported no improved survival for patients with TT >4 h [10]. Since TT >4 h is relatively uncommon in North America, we conducted a sensitivity analysis among patients with TT ≤4 h and found very consistent results (overall findings and regional differences). Our sensitivity analyses indicate that the variability in the association of TT with mortality across regions is not explained by differences in blood flow or vascular access and suggest that other factors may play a role. Other differences that may vary across regions and impact the association of TT with mortality potentially include both patient characteristics and dialysis practices, and additional study is warranted.

Better control of anemia, blood pressure, fluid overload and phosphorus levels as well as improved nutrition, left ventricular function and quality of life have been reported in small cohorts of patients receiving daily in-center and long nightly dialysis [2–4]. Improved blood pressure and phosphorus control decreases in left ventricular mass and improvement in physical health were recently reported among Frequent Hemodialysis Network participants randomized to frequent dialysis [5].

Our study demonstrates an association between longer TT and better intermediate outcomes (with ‘better’ referring to generally accepted clinical targets). Longer dialysis sessions provide greater clearance of both small and larger molecules [31]. Greater clearance may, for example, improve inflammatory status as indicated by the lower WBC count. This may, in turn, improve anemia control and lower EPO requirements. Finally, the longer dialysis sessions allow for slower ultrafiltration rates and tolerance of greater fluid removal, leading to improved control of hypertension [32, 33]. This is indicated in instrumental variable analyses where longer TT was associated with lower SBP levels both before and after dialysis as well as with better achievement of current clinical guidelines for BP control [34]. In support of this finding, longer TT was also associated with lower risk of hospital admission for fluid overload or congestive heart failure, presumably due to volume overload. These data confirm the findings of improved volume control with the change to longer TT in a study of 17 patients published in the 1980s [35].

Overall, it is reasonable to postulate that the improvement of one or more of these clinical markers may contribute to better survival for patients receiving longer dialysis sessions. To our knowledge, our results provide support indicating that several pathophysiological mechanisms may link longer TT with longer survival and fewer hospitalizations.

A strength of the current study is that it applies both standard regression and instrumental variable approaches. The latter uses the dialysis facility as an instrument to lessen confounding by indication caused by unmeasured patient-level confounders [14–17]. Both types of analyses yielded generally corroborative associations between TT and patient outcomes. Results of prior patient-based

studies may have been biased by differences between patients receiving long versus short TT that may not have been completely taken into account, despite extensive model adjustments. For example, only more adherent patients may be willing to undergo the longer sessions; these patients are likely more adherent with medication prescription and dietary restrictions and may survive longer. On the other hand, patients who are sicker may be prescribed longer dialysis sessions. The instrumental variable methodology partially addresses this issue and is being applied to several fields of medical research [18, 36–40].

The established DOPPS infrastructure and representative sampling approach across 12 countries represents another strength, while raising regional differences in the association of TT with survival as a topic for further study. The extensive DOPPS data set allowed us to describe the association between delivered TT with outcomes, yielding very similar results as the prescribed TT analyses.

The limitations of the study are related to its observational design. Despite the extensive adjustments and the use of an instrumental variable approach, the potential for residual confounders remain and our results do not prove a causal effect between longer TT and better clinical outcomes.

Facilities delivering longer dialysis sessions may face higher costs [41], and shortening dialysis treatments may be associated with cost savings in certain payment environments [42]. The current US Centers for Medicare & Medicaid Services clinical performance measures and the planned Quality Incentive Program are based on delivered dialysis dose rather than TT [43]. Therefore, the pressures that dialysis providers in the USA will be facing with the implementation of the bundled ESRD prospective payment system [44] may incentivize shorter dialysis sessions as long as adequate urea clearance is provided. These incentives contrast with other countries, such as the Japanese reimbursement structure that favors at least 4 h of TT in all but the sickest patients and the German Qualitäts-sicherungs-Richtlinie Dialyse that bases reimbursement, in part, on achieved TT of at least 4 h [45]. Of note, TT in Germany has risen dramatically as this financial incentive has been implemented (W. Kleophas, personal communication).

In the absence of randomized controlled clinical trials, we encourage health care providers to take into account findings from observational studies as well as supportive principles of dialysis, when making decisions regarding the duration of dialysis sessions. Similarly, policy makers and developers of quality measures worldwide may consider the current evidence about the duration of dialysis session when creating policies or guidelines that may affect TT.

In summary, our study confirms generally favorable clinical outcomes with longer TT and demonstrates associations of longer TT with better anemia, phosphorus and blood pressure control indicating possible mechanisms for improved clinical outcomes. These findings support longer TT prescription in the setting of three times per week HD.

**Acknowledgments.** This paper received editorial support from Jennifer McCready-Maynes, an employee of Arbor Research Collaborative for Health. The DOPPS is administered by Arbor Research Collaborative for Health and is supported by scientific research grants from Amgen (since 1996), Kyowa Hakko Kirin (since 1999, in Japan), Sanofi/Genzyme (since 2009), Abbott (since 2009), Vifor Fresenius Renal Pharma (since 2011) and Baxter (since 2011), without restrictions on publications.

**Conflict of interest statement.** F.T., J.Z., A.K., F.P., R.P. and B.R. are employees of Arbor Research Collaborative for Health which is supported by scientific research grants from Amgen (since 1996), Kyowa Hakko Kirin (since 1999, in Japan), Sanofi/Genzyme (since 2009), Abbott (since 2009) and Baxter (since 2011), without restrictions on publications. F.T. was supported by Award Number K01DK087762-01A1 from the National Institute Of Diabetes And Digestive And Kidney Diseases. Y.L., R.S., J.B. and T.A. have no disclosures to make. P.K. is on advisory boards for Fresenius and Baxter as well as Amgen and Genzyme.

(See related article by Eloot *et al.* Less water for haemodialysis: is multiple pass the future pace to go? *Nephrol Dial Transplant* 2012; 27: 3975–3978.)

## References

- US Renal Data System. *USRDS 2010 Annual Data Report: Atlas of Chronic Kidney Disease and End-Stage Renal Disease in the United States*. Bethesda, MD: National Institutes of Health, National Institute of Diabetes and Digestive and Kidney Diseases, 2010
- Suri RS, Nesrallah GE, Mairra R *et al.* Daily hemodialysis: a systematic review. *Clin J Am Soc Nephrol* 2006; 1: 33–42
- Walsh M, Culleton B, Tonelli M *et al.* A systematic review of the effect of nocturnal hemodialysis on blood pressure, left ventricular hypertrophy, anemia, mineral metabolism, and health-related quality of life. *Kidney Int* 2005; 67: 1500–1508
- Charra B. Fluid balance, dry weight, and blood pressure in dialysis. *Hemodial Int* 2007; 11: 21–31
- Chertow GM, Levin NW, Beck GJ *et al.* In-center hemodialysis six times per week versus three times per week. *N Engl J Med* 2010; 363: 2287–2300
- Held PJ, Levin NW, Bovbjerg RR *et al.* Mortality and duration of hemodialysis treatment. *JAMA* 1991; 265: 871–875
- Lowrie EG, Li Z, Ofsthun N *et al.* Measurement of dialyzer clearance, dialysis time, and body size: death risk relationships among patients. *Kidney Int* 2004; 66: 2077–2084
- Saran R, Bragg-Gresham JL, Levin NW *et al.* Longer treatment time and slower ultrafiltration in hemodialysis: associations with reduced mortality in the DOPPS. *Kidney Int* 2006; 69: 1222–1228
- Marshall MR, Byrne BG, Kerr PG *et al.* Associations of hemodialysis dose and session length with mortality risk in Australian and New Zealand patients. *Kidney Int* 2006; 69: 1229–1236
- Miller JE, Kovesdy CP, Nissenson AR *et al.* Association of hemodialysis treatment time and dose with mortality and the role of race and sex. *Am J Kidney Dis* 2010; 55: 100–112
- Brunelli SM, Chertow GM, Ankers ED *et al.* Shorter dialysis times are associated with higher mortality among incident hemodialysis patients. *Kidney Int* 2010; 77: 630–636
- Young EW, Goodkin DA, Mapes DL *et al.* The Dialysis Outcomes and Practice Patterns Study (DOPPS): an international hemodialysis study. *Kidney Int* 2000; 57 (Suppl 74): S74–S81
- Pisoni RL, Gillespie BW, Dickinson DM *et al.* The Dialysis Outcomes and Practice Patterns Study: design, data elements, and methodology. *Am J Kidney Dis* 2004; 44 (Suppl 2): S7–S15
- Angrist JD, Imbens GW, Rubin DB. Identification of causal effects using instrumental variables. *J Am Stat Assoc* 1996; 91: 444–455
- Wooldridge JM. *Introductory Econometrics*, 4th edn. Mason, OH: South-Western College Publishing, 2002. chapter 15
- Newhouse JP, McClellan M. Econometrics in outcomes research: the use of instrumental variables. *Annu Rev Public Health* 1998; 19: 17–34
- Greenland S. An introduction to instrumental variables for epidemiologists. *Int J Epidemiol* 2000; 29: 722–729
- Pisoni RL, Arrington CJ, Albert JM *et al.* Facility hemodialysis vascular access use and mortality in countries participating in DOPPS: an instrumental variable analysis. *Am J Kidney Dis* 2009; 53: 475–491
- Stock JH, Wright JH, Yogo M. A survey of weak instruments and weak identification in generalized method of moments. *J Bus Econ Stat* 2002; 20: 518–529
- Burgess S, Thompson SG. CRP CHD Genetics Collaboration. Avoiding bias from weak instruments in Mendelian randomization studies. *Int J Epidemiol* 2011; 40: 755–764. doi:10.1093/ije/dyr036
- Raghunathan TE, Solenberger PW, Van Hoewyk J. IVEware: Imputation and Variance Estimation Software. Survey Methodology Program, Survey Research Center, Institute for Social Research, University of Michigan, 2002
- von Elm E, Altman DG, Egger M *et al.* The Strengthening of Reporting of Observational Studies in Epidemiology (STROBE) statement: guidelines for reporting observational studies. *Lancet* 2007; 370: 1453–1457
- Saran R, Bragg-Gresham JL, Rayner HC *et al.* Nonadherence in hemodialysis: associations with mortality, hospitalization, and practice patterns in the DOPPS. *Kidney Int* 2003; 64: 254–262
- K/DOQI Workgroup. K/DOQI clinical practice guidelines for cardiovascular disease in dialysis patients. *Am J Kidney Dis* 2005; 45 (4 Suppl 3): S1–153
- Kidney Disease: Improving Global Outcomes (KDIGO) CKD-MBD Work Group. KDIGO clinical practice guideline for the diagnosis, evaluation, prevention, and treatment of Chronic Kidney Disease-Mineral and Bone Disorder (CKD-MBD). *Kidney Int Suppl* 2009, S1–130
- KDOQI Workgroup. KDOQI clinical practice guidelines and clinical practice recommendations for anemia in chronic kidney disease. *Am J Kidney Dis* 2006; 47 (5 Suppl 3): S11–145
- Genovesi S, Valsecchi MG, Rossi E *et al.* Sudden death and associated factors in a historical cohort of chronic haemodialysis patients. *Nephrol Dial Transplant* 2009; 24: 2529–2536
- Lowrie EG, Laird NM, Parker TF *et al.* Effect of the hemodialysis prescription of patient morbidity: report from the National Cooperative Dialysis Study. *N Engl J Med* 1981; 305: 1176–1181
- Owen WF, Jr, Lew NL, Liu Y *et al.* The urea reduction ratio and serum albumin concentration as predictors of mortality in patients undergoing hemodialysis. *N Engl J Med* 1993; 329: 1001–1006
- Held PJ, Port FK, Wolfe RA *et al.* The dose of hemodialysis and patient mortality. *Kidney Int* 1996; 50: 550–556
- Eloot S, Van Biesen W, Dhondt A *et al.* Impact of hemodialysis duration on the removal of uremic retention solutes. *Kidney Int* 2008; 73: 765–770
- Charra B, Caemard M, Laurent G. Importance of treatment time and blood pressure control in achieving long-term survival on dialysis. *Am J Nephrol* 1996; 16: 35–44
- McGregor DO, Buttmore AL, Nicholls MG *et al.* Ambulatory blood pressure monitoring in patients receiving long, slow home haemodialysis. *Nephrol Dial Transplant* 1999; 14: 2676–2679
- National Kidney Foundation. K/DOQI clinical practice guidelines for chronic kidney disease: evaluation, classification, and stratification. *Am J Kidney Dis* 2002; 39 (2 Suppl 1): S1–266
- Wizemann V, Kramer W. Short-term dialysis—long-term complications. Ten years experience with short-duration renal replacement therapy. *Blood Purif* 1987; 5: 193–201
- Stukel TA, Fisher ES, Wennberg DE *et al.* Analysis of observational studies in the presence of treatment selection bias: effects of invasive cardiac management on AMI survival using propensity score and instrumental variable methods. *JAMA* 2007; 297: 278–285
- Schneeeweiss S, Seeger JD, Landon J *et al.* Aprotinin during coronary-artery bypass grafting and risk of death. *N Engl J Med* 2008; 358: 771–783



38. Brookhart MA, Rassen JA, Wang PS *et al.* Evaluating the validity of an instrumental variable study of neuroleptics: can between-physician differences in prescribing patterns be used to estimate treatment effects?. *Med Care* 2007; 45 (10 Suppl 2): S116–S122
39. Ramirez SP, Albert JM, Blayney MJ *et al.* Rosiglitazone is associated with mortality in chronic hemodialysis patients. *J Am Soc Nephrol* 2009; 20: 1094–1101
40. Tentori F, Albert JM, Young EW *et al.* The survival advantage for haemodialysis patients taking vitamin D is questioned: findings from the Dialysis Outcomes and Practice Patterns Study. *Nephrol Dial Transplant* 2009; 24: 963–972
41. Hirth RA, Held PJ, Orzol SM *et al.* Practice patterns, case mix, Medicare payment policy, and dialysis facility costs. *Health Serv Res* 1999; 33: 1567–1592
42. Held PJ, García JR, Pauly MV *et al.* Price of dialysis, unit staffing, and length of dialysis treatments. *Am J Kidney Dis* 1990; 15: 441–450
43. H. R. 6331: Medicare Improvements for Patients and Providers Act of 2008. United States, 2008, pp. 60–67
44. Medicare Coverage for End-Stage Renal Disease Patients. *Compilation of the Social Security Laws Including the Social Security Act, as Amended, and Related Enactments Through January 1, 2009*. [http://www.ssa.gov/OP\\_Home/ssact/title18/1881.htm#ft576](http://www.ssa.gov/OP_Home/ssact/title18/1881.htm#ft576) (1st December 2011, date last accessed)
45. NN: Qualitäts sicherungs-Richtlinie Dialyse (German). *Bundesanzeiger* 2006; 58: 115a

Received for publication: 23.8.2011; Accepted in revised form: 16.1.2012

*Nephrol Dial Transplant* (2012) 27: 4188–4196

doi: 10.1093/ndt/gfs351

Advance Access publication 7 August 2012

## Exploring the relationships between patient characteristics and their dialysis care experience

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### Abstract

**Background.** Previous studies have shown that it is possible for patient experience to be influenced by factors that are not attributable to health-care. Therefore, if patient experience is to be used as an accurate indicator of clinical performance, then it is important to understand its determinants.

**Methods.** We used data from 840 dialysis patients who completed a validated patient experience survey. We created a potential theoretical framework based on available clinical knowledge to hypothesize the relationships between 13 demographic, socio-economic and health status factors and three outcome measures: global rating of the dialysis centre and the patient experience with the nephrologist's and nurses' care. The theoretical framework guided the selection of confounding variables for each determinant, which were then entered as terms in multivariable linear regression models.

**Results.** Patients who were of older age, of non-European decent, and who had a lower educational level, lower

albumin level, with better self-rated health and who were without co-morbidities reported higher global ratings with the dialysis centre than their counterparts. Past myocardial infarction and better self-rated health were found to be determinants of a more positive experience while in the nephrologist's care. A more positive experience with nurses' care was associated with factors including older age, Dutch origin background, lower educational level, lower albumin levels and better self-rated health.

**Conclusions.** Several characteristics of dialysis patients influence the way they rate and experience their care. When using the patient experience and ratings as indicators of clinical performance, they should be adjusted for such factors as identified in our study. This will facilitate a meaningful comparison of dialysis centres, and enable informed decision making by patients, insurers and policy makers.

**Keywords:** case-mix adjustment; health-care outcome assessment; health-care quality indicators; patient satisfaction; renal dialysis