Apolipoprotein E genotype, cardiovascular biomarkers and risk of stroke: Systematic review and meta-analysis of 14 015 stroke cases and pooled analysis of primary biomarker data from up to 60 883 individuals

Tauseef A Khan,¹[†] Tina Shah,²[†] David Prieto,¹ Weili Zhang,³ Jackie Price,⁴ Gerald R Fowkes,⁴ Jackie Cooper,⁵ Philippa J Talmud,⁵ Steve E Humphries,⁵ Johan Sundstrom,⁶ Jaroslav A Hubacek,⁷ Shah Ebrahim,¹ Debbie A Lawlor,⁸ Yoav Ben-Shlomo,⁸ Mohammad R Abdollahi,⁹ Arjen JC Slooter,¹⁰ Zoltan Szolnoki,¹¹ Manjinder Sandhu,¹² Nicholas Wareham,¹³ Ruth Frikke-Schmidt,¹⁴ Anne Tybjærg-Hansen,¹⁴ Gerda Fillenbaum,¹⁵ Bastiaan T Heijmans,¹⁶ Tomohiro Katsuya,¹⁷ Grazyna Gromadzka,¹⁸ Andrew Singleton,¹⁹ Luigi Ferrucci,²⁰ John Hardy,²¹ Bradford Worrall,²² Stephen S Rich,²³ Mar Matarin,²⁴ John Whittaker,¹ Tom R Gaunt,^{8,9} Peter Whincup,²⁵ Richard Morris,²⁶ John Deanfield,²⁷ Ann Donald,²⁷ George Davey Smith,^{8,28} Mika Kivimaki,²⁹ Meena Kumari,²⁹ Liam Smeeth,¹ Kay-Tee Khaw,³⁰ Michael Nalls,¹⁹ James Meschia,³¹ Kai Sun,³ Rutai Hui,³ Ian Day,^{8,9} Aroon D Hingorani^{2,29}* and Juan P Casas^{1,29}*

¹Faculty of Epidemiology and Population Health, London School of Hygiene and Tropical Medicine, London, UK, ²Centre for Clinical Pharmacology, Division of Medicine, University College London, London, UK, ³Key Laboratory for Clinical Cardiovascular Genetics, Fu Wai Hospital, Chinese Academy of Medical Sciences, Beijing, China, ⁴Centre for Population Health Sciences, University of Edinburgh, Edinburgh, UK, ⁵Centre for Cardiovascular Genetics, Division of Medicine, University College London, London, UK, ⁶Department of Medical Sciences & Uppsala Clinical Research Center, Uppsala University Hospital, Uppsala, Sweden, ⁷Centre for Cardiovascular Research and Institute of Clinical and Experimental Medicine, Prague, Czech Republic, ⁸MRC Centre for Causal Analyses in Translational Epidemiology, School of Social and Community Medicine, University of Bristol, Bristol, UK, ⁹Bristol Genetic Epidemiology Laboratories, School of Social and Community Medicine, University of Bristol, UK, ¹⁰Department of Intensive Care, University Medical Centre, Utrecht, The Netherlands, ¹¹Department of Neurology and Cerebrovascular Diseases, Pandy County Hospital, Gyula, Hungary, ¹²Department of Public Health and Primary Care, Strangeways Research Laboratory, Wort's Causeway, Cambridge, UK, ¹³MRC Epidemiology Unit, Institute of Metabolic Science, Addenbrooke's Hospital, Cambridge, UK, ¹⁴Department of Clinical Biochemistry, Rigshospitalet, Copenhagen University Hospital, Copenhagen, Denmark, ¹⁵Center for the Study of Aging and Human Development, Duke University Medical Center, Durham, NC, USA, 16 Molecular Epidemiology, Leiden University Medical Center, The Netherlands, ¹⁷Department of Clinical Gene Therapy, Osaka University Graduate School of Medicine, Suita, Japan, ¹⁸Institute of Psychiatry and Neurology, Second Department of Neurology, Sobieskiego, Warsaw, Poland, ¹⁹Laboratory of Neurogenetics, National Institute on Aging, US National Institute of Health, USA, ²⁰Longitudinal Studies Section, National Institute on Aging, US National Institute of Health, USA, ²¹Institute of Neurology, University College London, London, Watholian institute of Aging, US National institute of nearlin, USA, ²¹ Institute of Neurology, University concerciption, Eondon, Eondon, UK, ²²Department of Neurology, University of Virginia, USA, ²⁴Department of Clinical and Experimental Epilepsy, UCL Institute of Neurology, University College London, London, UK, ²⁵ Community Health Services, St George', University of London, London, UK, ²⁶UCL Department of Primary Care & Population Health, Royal Free Hospital, London, UK, ²⁷Vascular Physiology Unit, Institute of Neurology, University of Primary Care & Population Health, Royal Free Hospital, London, UK, ²⁷Vascular Physiology Unit, Institute of Neurology, University of Primary Care & Population Health, Royal Free Hospital, London, UK, ²⁷Vascular Physiology Unit, Institute of Neurology, University of Primary Care & Population Health, Royal Free Hospital, London, UK, ²⁷Vascular Physiology Unit, Institute of Primary Care & Population Health, Royal Free Hospital, London, UK, ²⁷Vascular Physiology Unit, Institute of Primary Care & Population Health, Royal Free Hospital, London, UK, ²⁷Vascular Physiology Unit, Institute of Primary Care & Population Health, Royal Free Hospital, London, UK, ²⁷Vascular Physiology Unit, Institute of Primary Care & Population Health, Royal Free Hospital, London, UK, ²⁷Vascular Physiology Unit, Institute of Primary Care & Population Health, Royal Free Hospital, London, UK, ²⁷Vascular Physiology Unit, Institute of Physiology Uni Child Health, University College London, London, UK, ²⁸School of Social and Community Medicine, University of Bristol, Bristol, UK, ²⁹Genetic Epidemiology Group, Department of Epidemiology and Public Health, University College London, London, UK, ³⁰Clinical Gerontology Unit, Addenbrooke's Hospital, Cambridge, UK and ³¹Mayo Clinic, Jacksonville, FL, USA

*Corresponding authors. E-mail: a.hingorani@ucl.ac.uk *E-mail: juan.pablo-casas@lshtm.ac.uk

Accepted 13 February 2013

Background At the *APOE* gene, encoding apolipoprotein E, genotypes of the $\varepsilon 2/\varepsilon 3/\varepsilon 4$ alleles associated with higher LDL-cholesterol (LDL-C) levels are also associated with higher coronary risk. However, the association of

† These authors contributed equally to the work.

APOE genotype with other cardiovascular biomarkers and risk of ischaemic stroke is less clear. We evaluated the association of *APOE* genotype with risk of ischaemic stroke and assessed whether the observed effect was consistent with the effects of *APOE* genotype on LDL-C or other lipids and biomarkers of cardiovascular risk.

- **Methods** We conducted a systematic review of published and unpublished studies reporting on *APOE* genotype and ischaemic stroke. We pooled 41 studies (with a total of 9027 cases and 61730 controls) using a Bayesian meta-analysis to calculate the odds ratios (ORs) for ischaemic stroke with *APOE* genotype. To better evaluate potential mechanisms for any observed effect, we also conducted a pooled analysis of primary data using 16 studies (up to 60883 individuals) of European ancestry. We evaluated the association of *APOE* genotype with lipids, other circulating biomarkers of cardiovascular risk and carotid intima-media thickness (C-IMT).
- **Results** The ORs for association of APOE genotypes with ischaemic stroke were: 1.09 (95% credible intervals (CrI): 0.84–1.43) for $\varepsilon 2/\varepsilon 2$; 0.85 (95% CrI: 0.78–0.92) for $\varepsilon 2/\varepsilon 3$; 1.05 (95% CrI: 0.89–1.24) for $\varepsilon 2/\varepsilon 4$; 1.05 (95% CrI: 0.99–1.12) for ε3/ε4; and 1.12 (95% CrI: 0.94–1.33) for $\varepsilon 4/\varepsilon 4$ using the $\varepsilon 3/\varepsilon 3$ genotype as the reference group. A regression analysis that investigated the effect of LDL-C (using APOE as the instrument) on ischaemic stroke showed a positive doseresponse association with an OR of 1.33 (95% CrI: 1.17, 1.52) per 1 mmol/l increase in LDL-C. In the separate pooled analysis, APOE genotype was linearly and positively associated with levels of LDL-C (*P*-trend: 2×10^{-152}), apolipoprotein B (*P*-trend: 8.7×10^{-06}) and C-IMT (P-trend: 0.001), and negatively and linearly associated with apolipoprotein E (*P*-trend: 6×10^{-26}) and HDL-C (*P*-trend: 1.6×10^{-12}). Associations with lipoprotein(a), C-reactive protein and triglycerides were non-linear.
- **Conclusions** In people of European ancestry, *APOE* genotype showed a positive dose-response association with LDL-C, C-IMT and ischaemic stroke. However, the association of *APOE* $\epsilon 2/\epsilon 2$ genotype with ischaemic stroke requires further investigation. This cross-domain concordance supports a causal role of LDL-C on ischaemic stroke.
- **Keywords** Stroke, lipids, apolipoprotein E, cardiovascular disease, systematic review, meta-analysis, biomarkers

Introduction

Worldwide, ischaemic stroke is the second leading cause of death after coronary heart disease (CHD) and is an important cause of disability, with a high burden of disease in low- and middle-income countries.^{1,2} Ischaemic stroke and CHD share several risk factors, including increasing age, high blood pressure, smoking and diabetes. However, although low-density lipoprotein cholesterol (LDL-C) is a known risk factor for CHD, its association with ischaemic stroke is uncertain,³ despite the fact that LDL-C-lowering statin drugs have been found to reduce both ischaemic stroke and CHD risk.⁴ This has prompted

speculation that the beneficial effect of statins on ischaemic stroke risk might be mediated through alternative pathways.⁵

Numerous studies have examined associations of common genetic variants of the apolipoprotein E (*APOE*) gene and cardiovascular outcomes, including ischaemic stroke and CHD. The *APOE* gene has two common non-synonymous polymorphisms, rs429358 and rs7412, that together give rise to three distinct *APOE* 'alleles' ($\varepsilon 2$, $\varepsilon 3$, $\varepsilon 4$); and these alleles in turn form six possible *APOE* 'genotypes' ($\varepsilon 2/\varepsilon 2$, $\varepsilon 2/\varepsilon 3$, $\varepsilon 2/\varepsilon 4$, $\varepsilon 3/\varepsilon 3$, $\varepsilon 3/\varepsilon 4$ and $\varepsilon 4/\varepsilon 4$).⁶

As confirmed by recent genome-wide association studies (GWAS), variants at the *APOE* locus are one

of the strongest signals for LDL-C, ^{7–10} and follow a positive dose-response association (when ordered $\epsilon 2/\epsilon 2$, $\epsilon 2/\epsilon 3$, $\epsilon 2/\epsilon 4$, $\epsilon 3/\epsilon 3$, $\epsilon 3/\epsilon 4$ and $\epsilon 4/\epsilon 4$).

In addition to LDL-C, APOE genotype has been associated with high-density lipoprotein cholesterol (HDL-C) and triglycerides, though the magnitude and dose-response relationships appear to differ from those described for LDL-C.¹¹ In addition, *APOE* genotype has been reported to be associated with some inflammatory biomarker such as C-reactive protein (CRP), mainly through the $\varepsilon 4$ allele,^{12–16} but its association with other inflammatory and coagulation biomarkers is less clear.^{17,18} Evidence from a large meta-analysis showed that APOE genotype follows a positive dose-response association with CHD concordant with that observed for LDL-C, supporting a causal role of LDL-C in CHD.¹¹ However, the association of APOE genotype with risk of ischaemic stroke is less clear and cannot be reliably extrapolated from the association with CHD. The largest published systematic review on APOE genotype and ischaemic stroke was inconclusive with only a tentative evidence for an association of the ε 4 allele with high risk of ischaemic stroke.¹⁹ Furthermore, this analysis was based on $\varepsilon 2$ and ε 4 carriers, whereas an analysis based on all six APOE genotypes is needed to estimate the genotype dose-response association. A similar limitation relates to the APOE association with carotid intima-media thickness (C-IMT), a non-invasive measure of atherosclerosis linked to both CHD and ischaemic stroke risk, that is used as a surrogate outcome in randomised trials.^{20,21}

Genome wide-association studies of ischaemic stroke and C-IMT have not identified any associations with genetic variation at the *APOE* locus.^{22,23} However, coverage of the single nucleotide polymorphisms (SNPs) contributing to or marking the APOE $\varepsilon 2/\varepsilon 3/\varepsilon 4$ genotype is incomplete in arrays used in GWAS, leading to uncertainty regarding the association of APOE genotype with C-IMT and stroke risk. We therefore undertook a systematic review and meta-analysis to examine the association between APOE genotype and ischaemic stroke risk, including twice as many cases as in the previous largest meta-analysis.¹⁹ The large data set allowed the evaluation of the risk of ischaemic stroke conferred by each of the six APOE genotypes individually, as well as the trend in risk across the genotype categories, which has not been possible in prior studies. In addition, to evaluate the potential mechanisms for any observed effect, we undertook a separate pooled analysis of primary data on the relationship between APOE genotype and a wide range of lipids and inflammation and coagulation markers as well as C-IMT.

Methods

Two separate meta-analyses were performed to test (i) the association of *APOE* genotype with cardiovascular traits and C-IMT, and (ii) association of *APOE* with

ischaemic stroke using both published and unpublished studies.

APOE genotype, cardiovascular traits and C-IMT

We developed a collaboration of 13 studies of European ancestry with information on the association of APOE genotypes with a wide range of blood lipids and apolipoproteins; inflammation, coagulation and metabolic markers; and C-IMT. The studies included were the English Longitudinal Study of Ageing (ELSA),²⁴ the Northwick Park Heart Study II (NPHS II),²⁵ the Whitehall II study (WHII),²⁶ the Edinburgh Artery Study (EAS),²⁷ the British Regional Heart Study (BRHS),²⁸ the British Women's Heart and Health Study (BWHHS),²⁹ the Caerphilly and Speedwell studies (CaPS),³⁰ the UCL Diabetes and Cardiovascular disease Study (UDACS) study,³¹ the Aspirin for Asymptomatic Atherosclerosis trial (AAAT),³² the European Atherosclerosis Research Study (EARS) I and II,³³ the Czech sub-study from the World Health Organization (WHO) Multinational Monitoring of Trends and Determinants in Cardiovascular Disease Study (Czech-MONICA)³⁴ and the Uppsala Longitudinal Study of Adult Men (ULSAM).³⁵ These data were supplemented by published data for the European Prospective Investigation of Cancer Norfolk study (EPIC-N),³⁶ Perth Carotid Ultrasound Disease Assessment Study (CUDAS)^{3'} and the Rotterdam Study.³⁸ Data request forms were based on a pre-specified analysis plan and did not include adjustment for lipid-lowering treatment. In a sub-sample of nine studies (11528 individuals), adjustment for age and gender for the APOE-LDL-C association was available. The complete data set comprised of 60883 individuals (Table 1).

Collectively the studies included information on several traits that were selected to interrogate the mechanisms by which variants at the APOE locus might exert their effects on ischaemic stroke. These were divided in two groups according to the level of evidence for their associations with APOE genotype. Traits with prior evidence for an association were lipids and apolipoproteins [total cholesterol, LDL-C, HDL-C, triglycerides (TG), apolipoprotein A1 (apoA1), apoB, apoE, and lipoprotein (a) (Lp(a)], inflammatory markers [CRP, fibrinogen, interleukin-6 (IL-6)] and C-IMT. Traits with unclear evidence for an association were coagulation factors [Factor VII, fibrinopeptide A, plasma viscosity, prothrombin fragment 1+2, tissue plasminogen activator (tPA), D-dimer, von Willebrand factor (vWF)), metabolic traits (homocysteine, insulin, glycated haemoglobin Alc (HbAlc), and creatinine], and some additional inflammatory markers (E-selectin, ICAM, VCAM and ferritin). In addition, we also included lifestyle factors (alcohol and BMI) and established risk factors for cardiovascular disease (blood pressure and glucose). For these traits, we did not anticipate an association

| | | , | | - | , , | | | | | | | | | | | |
|---|-----------------------|-----------------------|---------------------|----------------------|----------------------|----------------------|---------------------|----------------------|---------------------------------|---------------------|----------------------|----------------------|----------------------|-------------------------------|-------------------------|----------------------------|
| Study (ref) | ELSA ²⁴ | NPHS-II ²⁵ | WHIII ²⁶ | UDACS ³¹ | EAS ²⁷ | AAAT ³² | EARSI ³³ | EARSII ³³ | MONICA (CZECH) ³⁴ | ULSAM ³⁵ | BRHS ²⁸ | BWHHS ⁵⁴ | CaPS ^{30b} | EPIC Norfolk ³⁶ | CUDAS ^{37b} | Rotterdam ^{38b} |
| Study design ^a | Ρ | Ь | Р | CS | Ь | Р | CC | CC | Ρ | Р | Ь | Р | Р | Р | CS | Ρ |
| Sampling frame | Respondents of HSE | General practices | Workplace | Diabetic patients | General practices | General practices | Universities | Universities | Administrative districts | Health survey | General practices | General practices | General practices | General Practices | Community Population | Administrative district |
| N with DNA | 5274 | 2775 | 5500 | 575 | 940 | 2833 | 1881 | 779 | 2562 | 879 | 3947 | 3500 | 1500 | \sim 22 915 | 1111 | 5401 |
| % men | 48 | 100 | 77 | 59 | 50 | 28 | 51 | 100 | 46.5 | 100 | 100 | 0 | 100 | 46 | 50 | 59 |
| Year of DNA measurement ^c | 2004 | 1989–94 | 2002–04 | 2001-02 | 2004 | 1998 | 16-0661 | 1993–94 | 1997–98 & 2000– 01 | 2004 | 1998–2000 | 1999–2001 | 1993–94 | 1997–2000 | 1995–96 | 1992 |
| Baseline year | 1998, 1999, 2001 | 1989–94 | 1985–88 | 2001 | 1987 | 1998 | 1990–91 | 1993–94 | 1997 | 1970–73 | 1978–80 | 1999–2001 | 1979–83 | 1993 | 1995–96 | 1990–93 |
| Physical measu | res | | | | | | | | | | | | | | | |
| BMI | 5274 | 2682 | 5591 | 544 | 904 | | 1891 | 763 | 2562 | 874 | 3514 | | 1317 | | | |
| Systolic BP | 5274 | 2683 | 5592 | 547 | 903 | 2397 | 1891 | 765 | 2562 | 878 | 3511 | | 1331 | | | |
| Diastolic BP | 5274 | 2682 | 5592 | 547 | 901 | 2394 | 1891 | 765 | 2562 | 878 | 3511 | | 1330 | | | |
| Lifestyle measu | ures | | | | | | | | | | | | | | | |
| Alcohol | | 2685 | | | 904 | | 1888 | 764 | | | 3464 | | 1332 | | | |
| Blood measures | r A | | | | | | | | | | | | | | | |
| Total-C | 5274 | 2662 | 5573 | 547 | 903 | 2385 | 1891 | 753 | 2562 | 879 | 3497 | 3240 | 1296 | 22 915 | | |
| LDL-C | | 1677 | 5097 | 533 | 898 | | 1887 | 742 | | 874 | 3418 | 3161 | 1295 | 22 915 | | |
| HDL-C | 5274 | 1774 | | | 898 | | 1887 | 741 | 2562 | 877 | 3386 | 3236 | 1296 | 22 915 | | |
| Triglycerides | 5274 | 2664 | 5196 | 547 | 903 | | 1891 | 753 | 2562 | 879 | 2433 | 3239 | 1296 | | | |
| ApoAI | | 2287 | | | | | 1855 | 765 | | 357 | | | 1290 | | | |
| ApoB | | 2287 | | 383 | | | 1853 | 765 | | 357 | | | 1286 | | | |
| ApoE | 5195 | | | | | | 1881 | 765 | | | | | | | | |
| Lp(a) | | 2230 | | | 576 | | 1860 | | | 357 | | | 1249 | | | |
| Apo-AII | | | | | | | | | | | | | 1291 | | | |
| Homocysteine | | 1311 | | | | | | 757 | | | 3456 | | 1291 | | | |
| Glucose metabo | olism | | | | | | | | | | | | | | | |
| HbAlc | 5274 | | 5547 | 544 | | | | | | | 3471 | 3154 | | | | |
| Glucose | 5274 | | 4977 | 547 | 902 | | 1890 | 738 | 2562 | 878 | 3495 | 3221 | 1297 | | | |
| Insulin | | | 4635 | | | | 1766 | 707 | | 867 | 2410 | 3303 | 672 | | | |
| Coagulation fac | tors/markers: | | | | | | | | | | | | | | | |
| Viscosity | | | | | 863 | 2383 | | | | | | 3093 | 1286 | | | |
| Tp(a) | | 166 | | | 767 | | | | | | 3513 | 3250 | 1212 | | | |
| vWF | | 171 | | | 819 | | | | | | 3513 | 3251 | 1178 | | | |
| D-Dimer | | 171 | | | 816 | | | | | | 3509 | 3251 | 1178 | | | |
| Factor VII | | 2673 | | | 578 | | 1624 | | | | 3504 | 3234 | 934 | | | |
| Prothrombin | | 2658 | | | 572 | | | | | | | | 901 | | | |
| Fib peptide A | | 2658 | | | 899 | | | | | | | | | | | |

Table 1 Studies contributing to the analysis of APOE genotypes and cardiovascular biomarkers

(continued)

| · 🔾 | |
|--------|--|
| Ð | |
| Ē | |
| Ξ | |
| .н | |
| ÷ | |
| | |
| 5 | |
| ň | |
| \sim | |
| | |
| | |
| | |
| e | |
| | |
| 9 | |
| | |

| Table 1 C | ontinued | | | | | | | | | | | | | | | |
|--|---|-----------------------------|----------------------------|-----------------------------|---------------------|--------------------|---------------------|----------------------|---------------------------------|---------------------|--------------------|---------------------|---------------------|-------------------------------|--------------------------------|---|
| Study (ref) | ELSA ²⁴ | NPHS-II ²⁵ | WHIII ²⁶ | UDACS³¹ | EAS ²⁷ | AAAT ³² | EARSI ³³ | EARSII ³³ | MONICA (CZECH) ³⁴ | ULSAM ³⁵ | BRHS ²⁸ | BWHHS ⁵⁴ | CaPS ^{30b} | EPIC Norfolk ³⁶ | CUDAS ^{37b} Rotterdan | 100 and |
| Inflammation | markers | | | | | | | | | | | | | | | |
| Fibrinogen | 5274 | 2670 | 4879 | | 885 | 2185 | 1697 | | | 334 | 3511 | 3242 | 1284 | | | |
| I-CAM | | | | | 710 | | | | | | | | | | | |
| V-CAM | | | | | 709 | | | | | | | | | | | |
| E-selectin | | | | | 708 | | | | | | | | | | | |
| IL-6 | | | 4893 | 537 | 635 | | | | | | 2032 | 3253 | 767 | | | |
| CRP | 5274 | | 4927 | 536 | 611 | 2027 | - | 657 2. | 562 | | | 3156 | 3506 | | | |
| Ferritin | 5274 | | | | | | | | | | | | | | | |
| Renal function | ſ | | | | | | | | | | | | | | | |
| Creatinine | | 2664 | | 543 | | | | | | 861 | | 3240 | 1284 | | | |
| Macrovasculaı | function/stru | cture | | | | | | | | | | | | | | |
| Carotid IMT | | | 3890 | | 837 | | | | | 408 | | | | | 105 5401 | |
| Empty cell ^a ^a Study desi ^b From pub. | s denote da ign: P, pros lished data. | ata not apț șpective; Ct | olicable or S, cross se | - not availé ectional; C | able. C, case co | ntrol. | | | | | | | | | | |
| ^c Year of m | casurement | t is the yea | ar when b | lood for D | NA was c | obtained. T. | The cardiov | ascular bior | narker mea | surements | are in gen | ieral contei | mporaneou | s with this | time. | |

with APOE genotypes, but their assessment served to confirm protection from confounding of the genetic associations as a result of randomized inheritance of parental alleles (Mendel's second law).³⁹ Further details of the studies and measures are in Table 1 and Supplementary Table S1 (available as Supplementary data at *IJE* online).

Statistical analysis: APOE, cardiovascular traits and C-IMT

ApoB, apoE, plasma viscosity, Lp(a), homocysteine, TG, IL-6 and CRP exhibited a skewed distribution and were natural log transformed. In each study, we calculated the mean and standard deviation for each APOE genotype group, and used fixed effects meta-analyses to obtain the standardized mean difference (SMD), weighted mean difference (WMD) and their 95% confidence interval (CI), respectively, for each trait by genotype using the common $\varepsilon 3/\varepsilon 3$ genotype as the reference. Random effect meta-analysis was also conducted as a subsidiary analysis and is reported in Supplementary Table S2, available as Supplementary data at *IJE* online.⁴⁰ Data are presented in the order $\varepsilon_2/\varepsilon_2$, $\varepsilon_2/\varepsilon_3$, $\varepsilon_2/\varepsilon_4$, $\varepsilon_3/\varepsilon_3$, $\varepsilon_3/\varepsilon_4$ and $\varepsilon 4/\varepsilon 4$. I² (and its 95% CI) were calculated for each comparison.⁴¹ Because of the multiple comparisons (5 genotype comparisons for 32 traits) we also estimated corrected *P*-values for each estimate using the Sidak method.⁴² In order to explore the potential effect that age and gender may have on the association of APOE genotypes with LDL-C, we used a sub-sample of nine studies (11708 participants) and estimated the effect of the APOE genotypes with and without adjustment for age and gender. Given that studies included in the analysis span a long calendar period (1990 to 2004), we also stratified the studies in two equal time periods (before and after 1998) to explore the potential effect that year of collection of blood samples from which DNA was extracted may have on the APOE genotype-LDL-C association.

To evaluate a linear trend, meta-regression analysis was used on the summary estimated mean differences for each genotype using an additive model.^{43,44} The standard normal distribution was used to calculate P-values. This analysis was conducted using Stata software (version 11.2).

Meta-analysis of the of APOE genotype with ischaemic stroke

Search strategy

Following the HuGE Review Guidelines update,⁴⁵ we searched EMBASE and MEDLINE via PubMed up to March 2011 for articles in any language restricted to 'humans' using free text and MeSH expansion of the following terms: apoe, apo e, apolipoprotein e, brain infarction, cerebrovascular accident, cerebrovascular disease, and stroke. Additional relevant studies were identified through reference lists of published

meta-analyses, review articles and the studies included in the obtained papers.

Case-control and case-cohort studies of unrelated individuals were eligible if ischaemic stroke was verified by computerised tomography (CT) or magnetic resonance imaging (MRI), or if a physician confirmed the diagnosis of stroke, but not if self-reported. Studies reporting white-matter lesions and stroke in children were excluded. We used the largest data set for studies with serial publications and verified the decision with the study authors. We used the control group that most closely resembled the cases where more than one control group was studied. Two reviewers (T.A.K. and J.P.C.) reviewed the papers and consulted a third reviewer (A.D.H.) to resolve any disagreement.

Authors of identified studies were contacted to clarify information on genotype or stroke type where details were incomplete. We requested the following: (i) updated tabular data on six *APOE* genotypes by case-control status, (ii) blinding of lab staff, (iii) duplicate data in publications, (iv) type of stroke and (v) for studies with multiple ethnic groups, genotype counts were requested separately by ethnic group. Additionally we directly approached 88 studies that had reported on ischaemic stroke, to request information on *APOE* genotype, of which 38 replied (35 studies with ischaemic stroke and 3 with haemorrhagic stroke data).

Data were extracted into a pre-specified form and entered in Microsoft Excel. Details entered for each study were: study ID, author, year of publication, study location, study name, ethnicity, number of cases, source of cases and definition and exclusion criteria, control numbers, source of control and exclusion criteria, cases and controls matching variables (age and sex), genotype data on cases and controls, genotype frequencies, blinding of genotyping staff to identity of cases and controls, genotyping method, mean age of cases and controls. Any study showing results separately for more than one ethnicity has been counted as a different data set, so total studies represent total number of data sets.

A chi-square test on control group allele frequencies or whole study allele frequencies in cohort studies was used to assess departure from Hardy–Weinberg (HW) equilibrium.

Statistical analysis

We performed a Bayesian meta-analysis to estimate the unadjusted odds ratios (ORs) and 95% credible intervals (CrI) of ischaemic stroke for each genotype ($\varepsilon 2/\varepsilon 2$, $\varepsilon 2/\varepsilon 3$, $\varepsilon 2/\varepsilon 4$, $\varepsilon 3/\varepsilon 4$ and $\varepsilon 4/\varepsilon 4$) against the reference genotype $\varepsilon 3/\varepsilon 3$. This order was decided on the basis of association of *APOE* genotypes with LDL-C. A Bayesian model was used to estimate the five ORs simultaneously in one model using Markov Chain Monte Carlo techniques⁴⁶ (for details of the model see Supplementary

methods available at *IJE* online). The Bayesian approach has several advantages: It estimates simultaneously the five ORs mentioned above, taking into the fact that they are not independent from each other. It provides posterior credibility intervals that are more naturally interpreted than frequentist confidence intervals. It also allows for the use of the data coming from small studies without having to add any extra numbers to the situations when zero counts appear in an estimation of an OR.⁴⁷

To explore the trend of these ORs (if any) we estimated a meta-logistic regression model with ischaemic stroke as the outcome and the APOE genotypes as the explanatory variable (for details of the model see Supplementary methods available at *IJE* online). APOE genotypes were coded numerically using the mean difference in LDL-C between each genotype and the $\varepsilon 3/3$ group, as follows: $\varepsilon 2/2 = -0.756$ mmol/l, $\varepsilon 2/3 =$ $-0.481 \text{ mmol/l}, \ \epsilon 2/4 = -0.241 \text{ mmol/l}, \ \epsilon 3/3 = 0 \text{ mmol/l}$ (reference), $\varepsilon 3/4 = 0.161 \text{ mmol/l}$, $\varepsilon 4/4 = 0.312 \text{ mmol/l}$. This mean difference for LDL-C was derived from our APOE-LDL-C analysis from studies in European ancestry populations. The summary OR derived from the trend analysis could be indirectly interpreted as the effect of LDL-C on ischaemic stroke through the use of APOE $\varepsilon 2/\varepsilon 3/\varepsilon 4$ loci as the genetic instrument. To allow for a non-constant OR (equivalent to a non-linear trend in the log scale) we included the square of the explanatory variable (i.e. $LDL-C^2$) in the regression model. If the posterior 95% CrI of the coefficient of the quadratic term included the null value, we repeated the analysis simplifying the model by removing the quadratic term and estimating only a linear trend.

We undertook several sensitivity analyses by re-estimating the logistic regression in different subsets of studies considered to be at lower risk of bias. These subsets were: studies with more than 100 cases, studies with control samples in HW equilibrium, studies with no zero cells for all APOE genotype categories, studies with genotyping staff blinded to case-control status, studies with imaging confirmed diagnosis of ischaemic stroke and studies that used new techniques for APOE genotyping (e.g. Taqman genotyping assay, One-tube assay, Heteroduplex LightTyper method, Pyrosequencing) that are less susceptible to genotyping miscalls.^{48,49} In order to evaluate the potential for publication bias (small-study bias) we, as a sensitivity analysis, re-estimated the effect of the APOE genotype on published versus unpublished studies, excluding small studies and re-quantifying the APOE-effect.

For the above-mentioned analyses we used fixedeffect models for the primary analysis; however, random effect models were also fitted and the results are reported as Supplementary material (available as Supplementary data at *IJE* online).

Priors

We used priors centred on the null that were non-informative in any specific direction but would cover a reasonable range of hypotheses by restricting their variance, e.g. an OR would have a prior distribution centred on 1, with a 95% probability of being between 0.03 and 30. For parameters in the log scale [log(OR) and log(odds)] we used normal distributions. For the standard deviation between studies in the random effects model we used a uniform prior between 0 and 1.5 times the observed standard deviation between the different studies. Bayesian analyses were performed using R (version 2.13) and JAGS software (version 2.2.0).

Results

APOE genotype, LDL-C and C-IMT

Twelve studies contributed data on APOE genotypes and LDL-C levels (Table 1). APOE genotypes exhibited a positive trend with LDL-C (47 771 subjects, P-trend: 2.1×10^{-49}). With $\varepsilon 3/3$ as reference, the mean differences (in mmol/l) were: -0.75 (95% CI: -0.86, -0.65) for $\varepsilon 2/2$, -0.48 (95% CI -0.50, -0.46) for $\varepsilon 2/3$, -0.24(95% CI: -0.24, -0.01) for $\varepsilon 2/4$, 0.16 (95% CI: 0.14, 0.18) for \$\varepsilon3/4\$ and 0.32 (95% CI: 0.26, 0.37) for \$\varepsilon4/4\$ (Figure 1a). There was a difference of 1.07 mmol/l between $\varepsilon 2/\varepsilon 2$ and $\varepsilon 4/\varepsilon 4$ individuals for LDL-C. These point estimates were used in the Bayesian meta-analysis to determine the coding of APOE genotypes and therefore to evaluate indirectly the effect of LDL-C on ischaemic stroke using a trend analysis. This positive dose-response between APOE genotypes and LDL-C remained unaltered in the sub-sample of studies that had provided data adjusted by age and gender (Supplementary Figure S1, available as Supplementary data at IJE online). Likewise, when the APOE-LDL-C analysis was stratified according to the time of DNA extraction, similar results were observed (Supplementary Figure S2, available as Supplementary data at IJE online).

Five studies contributed data on *APOE* genotypes and C-IMT (Table 1). We found a positive doseresponse association between *APOE* genotypes and C-IMT (11641 subjects, *P*-trend: 0.001) that strongly resembled the association between *APOE* genotype and LDL-C. There was a mean difference of 0.043 mm of C-IMT between $\epsilon 2/\epsilon 2$ and $\epsilon 4/\epsilon 4$ individuals (Figure 1b).

Association of APOE genotype with other lipids and apolipoproteins

Sixteen studies contributed data on *APOE* genotypes and cardiovascular phenotypes. *APOE* genotype exhibited a positive trend with total cholesterol (54377 subjects, *P*-trend: 2×10^{-152}) and apoB (6931 individuals, *P*-trend: 8.7×10^{-06}), similar to the association with LDL-C described above (Figure 2). We observed a negative trend of *APOE* genotype with levels of apoE (7841 individuals, *P*trend: 6×10^{-26}) and a weak negative trend association with HDL-C (50571 individuals, *P*-trend: 1.6×10^{-49}) and apoAI (6554 individuals, *P*-trend: 5.3×10^{-8}). Associations of *APOE* genotype with other lipids and lipoproteins were more complex in shape. There was a non-linear relationship between *APOE* genotype and triglycerides (27637 individuals), with both $\varepsilon 2/\varepsilon 2$ and $\varepsilon 4/\varepsilon 4$ individuals having higher concentrations of triglycerides, compared with the $\varepsilon 3/\varepsilon 3$ reference group. We also identified an association between *APOE* genotype and Lp(a) among 6272 individuals (Figure 2).

Association of APOE genotype with inflammation markers

We found a marked association of APOE genotype with CRP concentration (26401 individuals). With reference to the $\varepsilon 3/\varepsilon 3$ group, individuals with the $\varepsilon 4/\varepsilon 4$ genotype had the lowest concentration of CRP, whereas individuals with the $\varepsilon 2/\varepsilon 4$ genotype and the $\varepsilon 3/\varepsilon 4$ genotype had CRP levels approximately midway between the $\varepsilon_{3/\varepsilon_{3}}$ and $\varepsilon 4/\varepsilon 4$ groups. There was no difference in CRP levels among individuals with the $\varepsilon 2/\varepsilon 2$ or $\varepsilon 2/\varepsilon 3$ genotypes. We did not find any association of APOE genotype with another hepatocyte-derived inflammation marker, fibrinogen (25 961 individuals, P-trend: 0.09), with the macrophage/adipocyte-derived inflammation marker IL-6 (12117 individuals, P-trend: 0.9) (Figure 3), or with a range of other inflammation markers (Supplementary Figure S3. available as Supplementary data at IJE online).

Association of APOE with coagulation markers, metabolic traits and other variables related to cardiovascular risk

We found no evidence for an association of *APOE* genotype with the coagulation markers factor VII (12 547 individuals), fibrinopeptide A (3 557 individuals), D-dimer (8925 individuals) or tPA (58 908 individuals), the endothelial activation markers vWF (8932 individuals) and E-selectin (708 individuals), systolic blood pressure (28 334 individuals), BMI (25 916 individuals), alcohol consumption (11 037 individuals) or indices of glycaemic control including HbA1c (17 990 individuals), fasting glucose (25 781 individuals) and insulin (14 360 individuals) (Figure S4).

Mean difference (95%CI), I^2 (95% CI), unadjusted and Sidak-corrected *P*-values for each of the *APOE* genotype associations with cardiovascular traits are provided in Supplementary Table S2 (available as Supplementary data at *IJE* online).

APOE genotype and ischaemic stroke *Search results*

We screened 1395 abstracts from the primary search and identified 249 potentially eligible articles. Of these, 2 studies included self-reported stroke, 7 contained duplicate data and a further 163 reported incomplete data or were not ischaemic stroke, and were excluded, leaving 77 studies reporting stroke that



Figure 1 Association of APOE genotypes in studies of European ancestry individuals with (a) LDL-cholesterol (b) carotid intima-media thickness, (c) ischaemic stroke and (d) ischaemic stroke (trend analysis). Black boxes indicate summary estimates with their size proportional to weight. For (c) and (d), the x-axis is plotted in the log-scale with distance between APOE genotypes equal to mean difference of LDL-C in mmol/l. For (d), the effect estimate of trend analysis indicates an OR of 1.33 (95% CrI: 1.17, 1.52) per 1 mmol/l increase in LDL-C

were included in this analysis (Supplementary Figure S5, available as Supplementary data at *IJE* online). Of these, 64 included information on ischaemic stroke and 19 gave information on haemorrhagic stroke (there was some overlap as 16 studies gave information on both). Thirteen studies reported strokes that were not classified as ischaemic or haemorrhagic. Of these, 10 were studies in European ancestry subjects^{28,35,36,50–54} (2265 stroke events) exclusively and these were counted as ischaemic stroke since 80% of stroke in European ancestry is ischaemic in nature.² This took the total number of ischaemic stroke studies to 74 with total of 14015 cases and 77888 controls. This included data from 39 published studies 3635 cases and 42024 controls from previously unpublished studies from 35 from complete unpublished data sets plus data sets updated from previous published studies.

A description of study design, source of controls, use of neuroimaging, genotyping method, blinding, and number of cases and controls, observed and expected genotype and allele frequencies (by ethnic groups) is







Standardised Mean Difference

Figure 3 APOE genotypes and mean differences for inflammatory traits. The graphs are displayed in standardized scale to allow comparability and show standardised mean differences of biomarker levels with APOE genotypes with ϵ_3/ϵ_3 as reference. The values (on the right) correspond to absolute weighted mean difference. Black boxes indicate estimates proportional to counts and horizontal lines represent 95% CI. Ln, natural log transformed

given in Supplementary Tables S3, S4 and S5, available as Supplementary data at *IJE* online.

Association of APOE genotype with ischaemic stroke in European ancestry individuals

In 41 studies (14 prospective and 27 case-control studies), with a total of 9027 cases and 61 730 controls of European ancestry using the $\varepsilon 3/\varepsilon 3$ genotype as the reference group, the ORs for ischaemic stroke were 1.09 (95% CrI: 0.84, 1.43) for $\varepsilon 2/\varepsilon 2$, 0.85 (95% CrI: 0.78, 0.92) for $\varepsilon 2/\varepsilon 3$, 1.05 (95% CrI: 0.89, 1.24) for $\varepsilon 2/\varepsilon 4$, 1.05 (95% CrI: 0.99, 1.12) for $\varepsilon 3/\varepsilon 4$, and 1.12 (95% CrI: 0.94, 1.33) for $\varepsilon 4/\varepsilon 4$ (Figure 1c). We note the large uncertainty in OR estimates around the less common genotypes ($\varepsilon 2/\varepsilon 2$, $\varepsilon 2/\varepsilon 4$ and $\varepsilon 4/\varepsilon 4$), and in particular the $\varepsilon 2/\varepsilon 2$ genotype which only contributed 1% of the total cases.

We initially fitted a model that included a quadratic variable (LDL-C²) to allow for a non-linear trend, but the coefficient for this variable had a posterior credibility interval that included the null and we decided to remove this term and estimate a simpler model only with a linear trend. The results of the linear trend of the log-odds of *APOE* $\varepsilon 2/\varepsilon 3/\varepsilon 4$ genotypes (coded in terms of the *APOE* effect on LDL-C) on ischaemic stroke are described in Figure 1d. In this regression model, the OR of ischaemic stroke per 1 mmol/l increase in LDL-C was 1.33 (95% CrI: 1.17, 1.52), equivalent to an OR of 0.75 (95% CrI: 0.66, 0.85) per 1 mmol/l reduction in LDL-C. An exploratory sensitivity analysis showed that this positive linear trend remained largely



Figure 4 Association of APOE genotypes with ischaemic stroke in all ethnicities with (a) ischaemic stroke, (b) ischaemic stroke (trend analysis). Black boxes indicate summary estimates with their size proportional to weight. The x-axis is plotted in the log-scale with distance between APOE genotypes equal to mean difference of LDL-C in mmol/l. For (b), the effect estimate of trend analysis indicates an OR of 1.39 (95% CrI: 1.25, 1.54) per 1 mmol/l increase in LDL-C

unaltered in all the subgroups (Supplementary Figure S6, available as Supplementary data at *IJE* online), with none of them showing strong evidence of a quadratic trend (the coefficient of the quadratic term had a posterior interval that included the null).

The random effects model showed very similar results to the fixed effects analysis despite the greater uncertainty in the effect estimates for less frequent *APOE* genotypes. The trend analysis was also concordant with the fixed effects model (Supplementary Figure S7, available as Supplementary data at *IJE* online).

Association of APOE genotype with ischaemic stroke in all ethnicities

In 74 studies (16 prospective and 58 case-control studies), with a total of 14015 cases and 77 888 controls in all populations using the $\varepsilon_3/\varepsilon_3$ genotype as the reference group, the ORs for ischaemic stroke were 1.05 (95% CrI: 0.87, 1.26) for $\varepsilon_2/\varepsilon_2$, 0.88 (95% CrI: 0.82, 0.94) for $\varepsilon_2/\varepsilon_3$, 1.06 (95% CrI: 0.93, 1.21) for $\varepsilon_2/\varepsilon_4$, 1.15 (95% CrI: 1.09, 1.21) for $\varepsilon_3/\varepsilon_4$, and 1.22 (95% CrI: 1.05, 1.41) for $\varepsilon_4/\varepsilon_4$ (Figure 4a).

A visual inspection of the trend analysis of the log-odds of *APOE* on ischaemic stroke indicated the presence of a positive dose-response over the greater part of the LDL-C range (indirectly measured by the *APOE* genotypes). Under the assumption of a common model across ethnic groups for the LDL-C-ischaemic stroke association, we fitted a linear model and estimate that the OR of ischaemic stroke for 1 mmol/l of increase in LDL-C was 1.39 (95% CrI: 1.25, 1.54) (Figure 4b). However, we also allow for the possibility of alternative

models (i.e. quadratic model), and these results are included in our sensitivity analysis (Supplementary Figure S8, available as Supplementary data at *IJE* online). The random effects model produced very similar findings to the fixed effects model but with wider credible intervals (Supplementary Figure S9, available as Supplementary data at *IJE* online).

Assessment of the cumulative evidence, using the Venice criteria, suggest that the epidemiological credibility for the *APOE*-ischaemic stroke association is strong in European population studies and moderate or weak in all other ethnic groups (Supplementary Table S6, available as Supplementary data at *IJE* online).

The results of the sensitivity analysis that excluded small studies and the of one that re-estimated the effect for published and unpublished studies independently did not differ from the overall result, reducing the possibility that small-study bias influenced our results (Supplementary Figures S6 and S8, available as Supplementary data at *IJE* online).

In 11 studies with a total of 759 cases and 22 193 controls, using the $\varepsilon_3/\varepsilon_3$ genotype as the reference, no clear effect for any of the *APOE* genotypes with haem-orrhagic stroke was observed (Figure S10, available as Supplementary data at *IJE* online).

Discussion

Using the largest collection of studies yet accumulated, our analysis of European populations confirmed the presence of a positive dose-response association between APOE genotype and LDL-C, and clarified the genetic effect on C-IMT, which followed the same dose-response trend observed for LDL-C. The pattern was concordant in the meta-analysis of European ancestry studies on ischaemic stroke that also showed a positive dose-response association when the APOE genotypes were ordered and coded to reflect their effects on LDL-C (Figure 1). The large degree of consistency across the different outcome domains evaluated. [LDL-C (intermediate phenotype), C-IMT (surrogate end-point) and ischaemic stroke (clinical event)] argues for a causal effect of LDL-C on ischaemic stroke, and provides an adequate explanation for the beneficial effects on ischaemic stroke observed with statins in randomized trials.

In this respect, it is important to note that the estimate derived from our regression model in European ancestry studies that used *APOE* genotypes as instrument for LDL-C showed an OR of ischaemic stroke of 0.75 (95% CrI: 0.66, 0.85) per 1 mmol/l LDL-C reduction, which is consistent with the rate ratio of 0.79 (95% CI: 0.74, 0.85) for ischaemic stroke for the same LDL-C reduction derived from randomized trials of statins.⁵⁵

The large size of the data set synthesised in this report that included 2392 more cases of ischaemic stroke (including 6362 cases from unpublished and updated collections) than previous meta-analyses,¹⁹ permitted for the first time a reliable evaluation of the risk of ischaemic stroke in each of the six APOE genotype groups separately. Also, using APOE genotype as an instrument for LDL-C enabled the presence of a positive dose-response trend of LDL-C on ischaemic stroke to be determined. This is the largest study examining association of APOE genotypes with C-IMT to date, and shows that the genotypes are associated with atherosclerosis in a positive dose-response manner. This was similar in magnitude to the results of a previous meta-analysis examining this association.²⁰ This current study also suggests that the APOE effect on ischaemic stroke is mediated via LDL-C, providing further evidence of a causal role for this lipid fraction, which has been a point of conflict over several decades of research.

The finding of a positive dose-response effect of LDL-C (indexed by *APOE* genotype) on ischaemic stroke in European ancestry individuals was supported by the results from our sensitivity analysis that was restricted to the subset of studies considered to be at low risk of bias according to study level characteristics, such as study size, outcome diagnosis, blinding of genotyping staff, or genotyping technique. However, it is important to note that despite the large sample size included in this analysis, we still observed considerable uncertainty on the effect estimates for ischaemic stroke for the less frequent *APOE* genotypes, in particular the $\varepsilon 2/\varepsilon 2$ genotype.

When we extended our analysis of the *APOE* loci on ischaemic stroke to genetic studies from all ethnic

groups, a similar dose-response was observed for most of the association. However, a slight increase on ischaemic stroke for the section of the regression analysis associated with lower LDL-C levels, indexed by the genotypes $\varepsilon 2/\varepsilon 3$ and $\varepsilon 2/\varepsilon 2$, was observed. Our sensitivity analysis, however, indicated that a potential explanation for this apparent increase in risk could be due to genotyping errors as showed by the analysis restricted to the subset of studies in HW equilibrium, with blinding of the genotyping staff or those who used recent genotyping techniques. In this subset, the slight increase in risk of ischaemic stroke for $\varepsilon 2/\varepsilon 2$ genotypes was not seen and instead a positive dose-response relationship emerged (Supplementary Figure S6, available as Supplementary data at IJE online). A concern that some of the APOE genotypes may be prone to miscalling has been previously shown by some researchers,⁴⁸ indicating that studies that used Hixson and Vernier's PCR-RFLP method of genotyping or similar methods^{56,57} are more likely to have an excess of $\varepsilon 2/\varepsilon 2$ genotypes.^{48,58} This limitation is overcome by the newer APOE genotyping techniques (e.g. LightTyper assay and Tagman).49

A major limitation to this study is that we could not reliably assess the possibility of differential effects on the sub-types of ischaemic stroke. It has been postulated that different ischaemic stroke sub-types exhibit different associations with cholesterol and APOE genotype, suggesting differences in the underlying pathologies.¹⁹ Although information on haemorrhagic stroke was limited in our study, it was concordant with the lack of effect seen in observational studies (see Supplementary data at IJE online). However, our haemorrhagic stroke analysis had considerably less statistical power than our ischaemic stroke analysis. Larger studies are needed to answer this question with more confidence.⁵⁹ The ε 4 allele already has a validated association with Alzheimer's disease;⁶⁰ however, misclassification of Alzheimer's with ischaemic stroke is very unlikely as the two diseases have very distinct clinical and imaging profiles, and thus we believe this misclassification did not play a role in our results.

Small-study bias can be a major limitation in any meta-analysis. The analysis of *APOE* genotypes and cardiovascular traits was entirely based on *de novo* data and included a sample size of 60 883 individuals. This avoids the impact of small-study bias which is mainly observed in meta-analyses based on published studies. With regard to the analysis of *APOE* genotypes with stroke, our two sensitivity analyses, i.e. excluding small studies and re-quantifying the *APOE*-effect and comparing published versus unpublished studies, did not show any evidence of smallstudy bias.

Mechanisms linking APOE genotype to ischaemic stroke risk

To evaluate the mechanism by which *APOE* genotype might alter ischaemic stroke risk, we undertook a

detailed analysis of the association of APOE genotypes with a wide range of potential intermediate phenotypes including lipid, inflammation, coagulation, endothelial cell activation and metabolic markers measured in population-based and cross-sectional studies from European ancestry. Since genotype is determined at random at conception, intermediate phenotypes residing off the causal pathway from SNPs to disease, should be balanced evenly among the different genotypic groups, as they are in a randomized trial. In contrast, biomarkers that mediate the effect of genomic variation on disease risks should differ by genotype and the shape of the associations should be concordant.^{61,62} We identified robust associations of APOE genotype with total-, LDL- and HDL-C and triglycerides as reported in a recent systematic review.¹¹ However, we extended the analysis to apolipoproteins E (the protein encoded by the APOE gene), apoA1, apoB, Lp(a), and 18 other biomarkers, including the well-studied inflammation marker CRP, among several thousand individuals. An exploratory, but biologically interesting finding was the fact that the shape of the APOE association with these traits varied considerably. Whereas a doseresponse association was clearly observed with total-C, LDL-C, apoB, apoE and HDL-C, the effect on other traits appeared to be allele-specific, with lower Lp(a) and higher apoA1 being associated with the $\varepsilon 2$ allele, and lower CRP with the ε 4 allele. Although it is possible that these effects are due to a different biological mechanism of the APOE $\varepsilon 2$ and $\varepsilon 4$ variants on lipoprotein particle metabolism, it is also possible that these allele-specific associations could be due to different LD patterns between rarer, untyped, functional SNPs and the two genetic variants that comprise APOE $\varepsilon 2/\varepsilon 3/\varepsilon 4$. It is often argued that except for rare variants, most of the observed phenotypic changes in lipoprotein biology can be explained by differences in the $\varepsilon 2/\varepsilon 3/\varepsilon 4$ genotypes. We believe this view is simplistic and does not take into account the recent discoveries at this locus. Whole genome and dense genecentric SNP arrays (e.g. the Illumina Cardiochip) have identified associations of SNPs in the APOE gene region with LDL-C^{8,10,63-65} and Alzheimer's disease that encompass non-coding variants in the flanking genes and appear to be independent of the rs429358 and rs7412 SNPs (the two SNPS that encode $\varepsilon 2/\varepsilon 3/\varepsilon 4$) in multivariate analysis.63 These include variants in PVRL2, APOC1 and TOMM40.66,67 This suggests possible functional roles for one or more of the flanking genes (and their encoded proteins), mechanisms relating to APOE expression (not just function), or both. However, the capture of both common and rare SNPs in the APOE gene cluster is not yet exhaustive, so it is uncertain which of the implicated variants is causal, which mark as yet unidentified causal variation or whether causal SNPs differ for the different associated traits and disease outcomes. Not only are low-frequency variants (<5% minor allele frequency)

under-represented on whole genome arrays but surprisingly, there is also sparse coverage of the common alleles in the APOE region (including the omission of rs7412 (Supplementary rs429358 and SNPs) Figure S11 and Supplementary Table S7, available as Supplementary data at *IJE* online). The low degree of LD at this locus also makes it difficult to impute which variants, means additional untyped SNP-disease/trait association signals may have been overlooked by the prior GWAS. Early re-sequencing studies and isoelectric focusing analysis of APOE isoforms⁶⁸⁻⁷³ indicate the existence of many lowfrequency, non-synonymous coding variants in this region, but these studies were narrowly focused on the APOE gene itself, rather than the flanking genes, and were neither exhaustive nor systematic.

The shape of the association of APOE genotype with an intermediate phenotype on the causal pathway should be consistent in the direction of the association with the disease end-point. Thus, the reported association of APOE genotype with CHD appears to be consistent with its effect on LDL-C. This is further supported by similar concordant associations of SNPs in the *HMGCR*, *PCSK9* and *SORT1* genes with both LDL-C and CHD risk.^{8,10,64,74} Taken together with data from observational studies and randomized trials, the findings endorse the role of LDL-cholesterol in the pathogenesis of CHD. The positive doseresponse association of APOE genotype with ischaemic stroke risk in European population studies could be consistent with an almost complete mediating effect of LDL-C because both associations are positive and linear. None of the other biomarkers studied, including other lipids and apolipoproteins, as well as Lp(a), or inflammation markers (including CRP) exhibited associations that were consistent in shape or direction with that seen for ischaemic stroke. Indeed, it is important to note that the $\varepsilon 4$ allele associated with high values of LDL-C and C-IMT as well as high stroke and CHD risk showed low CRP levels. Although these results appear to be in contradiction with observational studies, we considered that findings of APOE on CRP could be explained by differential pattern of LD, as described above. Moreover, recent large-scale Mendelian randomization studies have excluded CRP as causal factor in cardiovascular disease.⁷⁵

The concordance in shape of the *APOE* $\varepsilon 2/\varepsilon 3/\varepsilon 4$ with LDL-C and ischaemic stroke risk observed in European ancestry received further support from the *APOE* $\varepsilon 2/\varepsilon 3/\varepsilon 4$ effects on C-IMT (derived from European ancestry studies) that also revealed a positive dose-response association (Figure 1). However, to confirm that the positive dose-response association of *APOE* loci on ischaemic stroke is mainly due to its effect on LDL-C would require access to large-scale cohorts with genotype information, LDL-C measures and ischaemic stroke outcomes in sufficiently large numbers to undertake appropriate Mendelian randomization/instrumental variables analyses.

C-IMT represents the structural change in arterial intima that strongly associates with both cardiovascular risk factors (LDL-C) and cardiovascular events, and it is also used as an intermediate marker of atherosclerosis.²¹ Overviews of statin trials have noted a linear relationship between the degree of LDL-C lowering and the change in C-IMT during follow-up, with modest LDL-C reductions being associated with reduced progression and more extreme LDL-C reductions being associated with regression of C-IMT.⁷⁶ The linear dose-response association of APOE genotypes with C-IMT was concordant in shape with the association of APOE with LDL-C, and endorses C-IMT as a valid surrogate marker of LDL-C-mediated atherosclerosis.77,78 Similar to our study, a recent metaanalysis of APOE and C-IMT showed lower levels of C-IMT for $\varepsilon 2$ carriers and higher values for $\varepsilon 4$ carriers when compared with $\varepsilon_{3/3}^{79}$ In contrast, the only GWAS of C-IMT²³ reported a lack of association with APOE variant rs7412 that marks the ε 4 allele (P-value: 0.7). Our study has the advantage of reporting the full range of APOE genotypes that might better represent this association.

Conclusions

In the most comprehensive study to date, we found a positive dose-response association of *APOE* genotype with ischaemic stroke in European populations. Concordance between the effects of the *APOE* genotype on LDL-C, C-IMT and ischaemic stroke risk was observed. Although this argues for a causal role of LDL-C in ischaemic stroke, and also helps to clarify the benefit of statins in ischaemic stroke, our finding must be interpreted with caution and it does not exclude the possibility that additional risk factors may be involved.

The use of the custom genotyping array Metabochip,⁸⁰ which includes both *APOE* SNPs and a greater coverage around the *APOE* locus, will provide an opportunity to replicate our findings on C-IMT, but also to explore in further detail the allele-specific associations of *APOE* locus with certain cardiovascular traits.

Supplementary Data

Supplementary data are available at IJE online.

Funding

T.A.K. is funded by a Graduate Teaching Assistantship of London School of Hygiene & Tropical Medicine. A.D.H. was funded by a British Heart Foundation Senior Fellowship (FS/05/125). S.E.H, J.C. and P.J.T. are supported by the British Heart Foundation (PG2008/015). L.S. is funded by a Wellcome Trust Senior Fellowship in Clinical Science. T.R.G. was the

recipient of a British Heart Foundation Intermediate Fellowship FS/05/065/19497 and T.R.G., I.D., S.E., D.A.L. and Y.B.S. are the recipients of a British Heart Foundation project grant PG/07/131/24254. D.A.L., G.D.S., T.R.G. and I.D. work in a UK Medical Research Council-supported centre. M.R.A. as funded by postdoctoral support from the University of Bristol which funded some set-up costs of Bristol Genetic Epidemiology Laboratory (M.R.A., T.R.G., I.D.). RCUK funding to I.D. and G.D.S. supported a Collaboration Development Workshop in Cardiovascular Genetics held in Beijing in March 2010. Tamuno Alfred is thanked for statistical support. J.S. is funded by the Swedish Research Council (2007-5942) and the Swedish Heart Lung Foundation (20041151). J.A.H. is funded by Centre for Cardiovascular Research and Institute for Clinical and Experimental Medicine project 00023001 (IKEM CR). A.J.S. is supported by the NESTOR stimulation programme for geriatric research in The Netherlands (Ministry of Health and Ministry of Education), the Netherlands Organization for Scientific Research (NWO), the Netherlands Prevention fund, the municipality of Rotterdam, the International Alzheimer Research Foundation (IARF), and the Fund for Scientific Research (FWO-F), Flanders, Belgium. T.K. (Tomohiro Katsuya) was supported by Grants-in-Aid for Scientific Research (18590265, 18590811. 19650188. 21390223) from the Ministry of Education, Science, Sports and Culture of Japan. R.H. is supported by the grants from the Ministry of Science and Technology of China (2011CB503900, 2009DFB30050 and 2006CB503805). M.A.N., A.B.S., J.H., B.B.W., S.R., T.G.B., R.D.B., J.F.M.: ISGS and SWISS received funding from grants from the National Institute of Neurological Disorders and Stroke (US), R01 NS39987 and R01 NS42733. This work was supported in part by the Intramural Research Program of the National Institute on Aging, National Institutes of Health, Department of Health and Human Services; project number Z01 AG000954-06. This work used samples and clinical data from the NINDS Human Genetics Resource Center DNA and Cell Line Repository (http://ccr.coriell.org/ninds), human subjects protocol numbers 2003-081 and 2004-147. The inclusion of BLSA samples was supported in part by the Intramural Research Program of the National Institute on Aging, National Institutes of Health, Department of Health and Human Services; project number Z01 AG000015-50, human subjects protocol number 2003-078. The ISGS study was funded by a grant from the National Institute for Neurological Disorders and Stroke (US; R01 NS42733) (J Meschia, P.I.). The SWISS study was funded by a grant from the National Institute for Neurological Disorders and Stroke (US; R01 NS39987 (J Meschia, P.I.). This study utilized the high-performance computational capabilities of the Biowulf Linux cluster at the

National Institutes of Health, Bethesda, MD (http://biowulf.nih.gov).

Samples from the English Longitudinal Study of Ageing (ELSA) DNA Repository (EDNAR), received support under a grant (AG1764406S1) awarded by the National Institute on Aging (NIA). ELSA was developed by a team of researchers based at the National Centre for Social Research, University College London and the Institute of Fiscal Studies, UK. The data were collected by the National Centre for Social Research. Northwick Park Heart Study II (NPHS-II) was supported by the British Medical Research Council, the US National Institutes of Health (NHLBI 33014) and Du Pont Pharma, Wilmington, USA. The Whitehall II study (is supported by the US National Institutes of Health (NHLBl R01HL036310-20A2, NIA R01AG013196 and R01AG034454), the British Heart Foundation (RG/ 02/005) and the Medical Research Council, UK. Edinburgh Artery Study (EAS) was funded by the British Heart Foundation. The British Regional Heart Study (BRHS) is a British Heart Foundation Research Group and is supported by British Heart Foundation (RG/04/003). The views expressed in this publication are those of the authors and not necessarily those of the funding bodies. The British Women's Heart and Health Study (BWHHS) was funded by the Department of Health Policy Research Programme and the British Heart Foundation (BHF Project Grant No: PG/06/002 and DoH Project Grant No: 0090041). The Caerphilly Prospective Study (CaPS) was established by the former MRC Epidemiology Unit in Cardiff. Diabetes UK supported the creation of UDACS. The AAA Trial (AAAT) was funded by the British Heart Foundation, Chief Scientist Office in Scotland, Chest Heart and Stroke Scotland and Wellcome Trust. The EARSI study was supported by the European Community Concerted Action MRH4 COMAC Epidemiology, and the EARSII study was supported by the European Commission Contract BMH1-CT92-0206. Analyses in the Czech-MONICA study were supported by project grants No. 00023001 (I.K.E.M., C.R.) and 1M0510 (M.S.M.T., C.R.). The Uppsala Longitudinal Study of Adult Men (ULSAM) study was supported by the Swedish Research Council (2006–6555), Swedish Heart-Lung Foundation, Erik, Karin och Gösta Selanders Foundation, Thuréus Foundation, Lisa och Johan Grönbergs Foundation, Loo och Hans Ostermans Foundation and Uppsala University.

T.A.K., T.S., A.D.H. and J.P.C. designed the experiment and the analysis plan. T.A.K., T.S. and D.P. did the statistical analysis and drafted the report with A.D.H. and J.P.C. All authors provided critical revisions. All contributors had shared summary data and contributed to the interpretation of the results and to the redrafting of the report. All members of the coordinating centre contributed to the collection, standardization, analysis and interpretation of the data. T.A.K., T.S., D.P., J..PC. and A.D.H. had full access to all data in the study and had final responsibility to submit the report for publication. The study was conducted and analysed independently from its funders.

Conflict of interest: J.W. is 90% employed at GlaxoSmithKline while retaining a 10% appointment at London School of Hygiene & Tropical Medicine. A.D.H. is a member of the editorial board of *Drug and Therapeutics Bulletin* (a BMJ group publication) and has received honoraria for speaking at educational meetings most or all of which have been donated to charity. There are no other conflicts of interest.

KEY MESSAGES

- The six *APOE* genotypes when ordered from ε2/ε2, ε2/ε3, ε2/ε4, ε3/ε3, ε3/ε4 and ε4/ε4 have a linear association with LDL-C and CIMT thickness.
- The largest meta-analysis of *APOE* genotype with ischaemic stroke shows a positive linear association of increasing risk when ordered from $\varepsilon_2/\varepsilon_2$, $\varepsilon_2/\varepsilon_3$, $\varepsilon_2/\varepsilon_4$, $\varepsilon_3/\varepsilon_3$, $\varepsilon_3/\varepsilon_4$ and $\varepsilon_4/\varepsilon_4$ in European ancestry populations.
- The concordance in the dose-response associations of *APOE* genotype with LDL-C, CIMT and ischaemic stroke provides strong support for a causal role of LDL-C in ischaemic stroke.

References

- ¹ Lopez AD, Mathers CD, Ezzati M, Jamison DT, Murray CJL. Global and regional burden of disease and risk factors, 2001: systematic analysis of population health data. *Lancet* 2006;**367**:1747–57.
- ² Feigin VL, Lawes CMM, Bennett DA, Anderson CS. Stroke epidemiology: a review of population-based

studies of incidence, prevalence, and case-fatality in the late 20th century. *Lancet* 2003;**2**:43–53.

- ³ Lewington S, Whitlock G, Clarke R *et al.* Blood cholesterol and vascular mortality by age, sex, and blood pressure: a meta-analysis of individual data from 61 prospective studies with 55 000 vascular deaths. *Lancet* 2007;**370:**1829–39.
- ⁴ Baigent C, Keech A, Kearney PM *et al.* Efficacy and safety of cholesterol-lowering treatment: prospective

meta-analysis of data from 90,056 participants in 14 randomised trials of statins. *Lancet* 2005;**366**:1267–78.

- ⁵ Amarenco P, Steg PGG. The paradox of cholesterol and stroke. *Lancet* 2007;**370:**1803–04.
- ⁶ Mahley RW, Rall SC. Apolipoprotein E: far more than a lipid transport protein. *Annu Rev Genomics Hum Genet* 2000;1:507–37.
- ⁷ Teslovich TM, Musunuru K, Smith AV *et al.* Biological, clinical and population relevance of 95 loci for blood lipids. *Nature* 2010;**466**:707–13.
- ⁸ Aulchenko YS, Ripatti S, Lindqvist I *et al*. Loci influencing lipid levels and coronary heart disease risk in 16 European population cohorts. *Nat Genet* 2009;**41**:47–55.
- ⁹ Wallace C, Newhouse SJ, Braund P *et al*. Genome-wide association study identifies genes for biomarkers of cardiovascular disease: serum urate and dyslipidemia. *Am J Hum Genet* 2008;**82**:139–49.
- ¹⁰ Willer CJ, Sanna S, Jackson AU *et al.* Newly identified loci that influence lipid concentrations and risk of coronary artery disease. *Nat Genet* 2008;**40**:161–69.
- ¹¹ Bennet AM, Di Angelantonio E, Ye Z *et al.* Association of apolipoprotein E genotypes with lipid levels and coronary risk. *JAMA* 2007;**298**:1300–11.
- ¹² Grönroos P, Raitakari OT, Kähönen M *et al.* Relation of Apolipoprotein E Polymorphism to Markers of Early Atherosclerotic Changes in Young Adults. *Circulation* 2008;**72**:29–34.
- ¹³ Rontu R, Ojala P, Hervonen A *et al*. Apolipoprotein E genotype is related to plasma levels of C-reactive protein and lipids and to longevity in nonagenarians. *Clin Endocrinol* 2006;**64**:265–70.
- ¹⁴ Berrahmoune H, Herbeth B, Siest G, Visvikis-Siest S. Heritability of serum hs-CRP concentration and 5-year changes in the Stanislas family study: association with apolipoprotein E alleles. *Genes Immun* 2007;**8**:352–59.
- ¹⁵ Chasman DI, Kozlowski P, Zee RY, Kwiatkowski DJ, Ridker PM. Qualitative and quantitative effects of APOE genetic variation on plasma C-reactive protein, LDLcholesterol, and apoE protein. *Genes Immun* 2006;**7**:211–19.
- ¹⁶ Elliott P, Chambers JC, Zhang W *et al*. Genetic Loci associated with C-reactive protein levels and risk of coronary heart disease. *JAMA* 2009;**302:**37–48.
- ¹⁷ Marz W, Scharnagl H, Hoffmann MM, Boehm BO, Winkelmann BR, März W. The apolipoprotein E polymorphism is associated with circulating C-reactive protein (the Ludwigshafen risk and cardiovascular health study). *Eur Heart J* 2004;**25**:2109–19.
- ¹⁸ Tsoi L-M, Wong K-Y, Liu Y-M, Ho Y-Y. Apoprotein E isoform-dependent expression and secretion of pro-inflammatory cytokines TNF-alpha and IL-6 in macrophages. *ArchBiochem Biophys* 2007;460:33–40.
- ¹⁹ Sudlow C, Martínez González NA, Kim J, Clark C. Does apolipoprotein E genotype influence the risk of ischaemic stroke, intracerebral hemorrhage, or subarachnoid hemorrhage? Systematic review and meta-analyses of 31 studies among 5961 cases and 17965 controls. *Stroke* 2006; **37**:364–70.
- ²⁰ Paternoster L, Martínez González NA, Lewis S, Sudlow C. Association between apolipoprotein E genotype and carotid intima-media thickness may suggest a specific effect on large artery atherothrombotic stroke. *Stroke* 2008;**39**:48–54.
- ²¹ Polak JF, Pencina MJ, Pencina KM *et al.* Carotid-wall intima-media thickness and cardiovascular events. *N Engl J Med* 2011;**365:**213–21.

- ²² Bellenguez C, Bevan S, Gschwendtner A *et al.* Genome-wide association study identifies a variant in HDAC9 associated with large vessel ischaemic stroke. *Nat Genet* 2012;44:328–33.
- ²³ Bis JC, Kavousi M, Franceschini N *et al.* Meta-analysis of genome-wide association studies from the CHARGE consortium identifies common variants associated with carotid intima media thickness and plaque. *Nat Genet* 2011;**43**:940–47.
- ²⁴ Marmot M, Banks J, Blundell R, Lessof C. Health, wealth and lifestyles of the older population in England: The 2002 English Longitudinal Study of Ageing. *Institute for Fiscal Studies* 2003; Jan:357–74.
- ²⁵ Miller GJ, Bauer KA, Barzegar S *et al*. The effects of quality and timing of venepuncture on markers of blood coagulation in healthy middle-aged men. *Thromb Haemost* 1995;**73:**82–86.
- ²⁶ Marmot M, Brunner E. Cohort Profile: the Whitehall II study. Int J Epidemiol 2005;34:251–56.
- ²⁷ Fowkes FG, Housley E, Cawood EH, Macintyre CC, Ruckley CV, Prescott RJ. Edinburgh Artery Study: prevalence of asymptomatic and symptomatic peripheral arterial disease in the general population. *Int J Epidemiol* 1991; **20**:384–92.
- ²⁸ Walker M, Whincup PH, Shaper AG. The British Regional Heart Study 1975-2004. *Int J Epidemiol* 2004; 33:1185–92.
- ²⁹ Lawlor DA, Timpson N, Ebrahim S, Day INM, Davey Smith G. The association of oestrogen receptor alpha-haplotypes with cardiovascular risk factors in the British Women's Heart and Health Study. *Eur Heart J* 2006;**27**:1597–604.
- ³⁰ The Caerphilly and Speedwell Collaborative Group. Caerphilly and Speedwell collaborative heart disease studies. *J Epidemiol Commun Health* 1984;**38**:259–62.
- ³¹ Dhamrait SS, Stephens JW, Cooper JA *et al.* Cardiovascular risk in healthy men and markers of oxidative stress in diabetic men are associated with common variation in the gene for uncoupling protein 2. *Eur Heart J* 2004;**25:**468–75.
- ³² Price JF, Stewart MC, Deary IJ *et al.* Low dose aspirin and cognitive function in middle aged to elderly adults: randomised controlled trial. *BMJ* 2008;**337**:a1198.
- ³³ The European Atherosclerosis Research Study (EARS): design and objectives. *Int J Epidemiol* 1994;23:465–71.
- ³⁴ The World Health Organization MONICA Project (monitoring trends and determinants in cardiovascular disease): a major international collaboration. WHO MONICA Project Principal Investigators. J Clin Epidemiol 1988;**41**:105–14.
- ³⁵ Wohlin M, Sundström J, Lannfelt L *et al.* Apolipoprotein E epsilon4 genotype is independently associated with increased intima-media thickness in a recessive pattern. *Lipids* 2007;**42**:451–56.
- ³⁶ Ward H, Mitrou PN, Bowman R *et al.* APOE genotype, lipids, and coronary heart disease risk: a prospective population study. *Arch Int Med* 2009;**169**:1424–29.
 ³⁷ D. ih. ID. ID. C. Polence V. J. et al. Application F.
- ³⁷ Beilby JP, Hunt CCJ, Palmer LJ *et al.* Apolipoprotein E gene polymorphisms are associated with carotid plaque formation but not with intima-media wall thickening: results from the Perth Carotid Ultrasound Disease Assessment Study (CUDAS). *Stroke* 2003;**34**:869–74.
- ³⁸ Slooter AJ, Bots ML, Havekes LM *et al.* Apolipoprotein E and carotid artery atherosclerosis: the Rotterdam study. *Stroke* 2001;**32**:1947–52.

- ³⁹ Davey Smith G, Ebrahim S. Mendelian randomization: prospects, potentials, and limitations. *Int J Epidemiol* 2004;**33**:30–42.
- ⁴⁰ Egger M, Davey Smith G, Altman D. Systematic Reviews in Health Care: Meta-analysis in Context. Oxford: Wiley-Blackwell, 2001, p. 512.
- ⁴¹ Higgins JPT, Thompson SG, Deeks JJ, Altman DG. Measuring inconsistency in meta-analyses. *BMJ* 2003; **327:**557.
- ⁴² Abdi H. The Bonferonni and Šidák Corrections for Multiple Comparisons. In: Salkind NJ (ed.). *Encyclopedia* of Measurement and Statistics. Thousand Oaks, CA: Sage, 2007, pp. 1–9.
- ⁴³ Higgins JPT, Thompson SG. Controlling the risk of spurious findings from meta-regression. *Stat Med* 2004;23: 1663–82.
- ⁴⁴ Knapp G, Hartung J. Improved tests for a random effects meta-regression with a single covariate. *Stat Med* 2003;**22**: 2693–710.
- ⁴⁵ Sagoo GS, Little J, Higgins JPT. Systematic reviews of genetic association studies. Human Genome Epidemiology Network. *PLoS Med* 2009;6:e28.
- ⁴⁶ Gilks W, Richardson S, Spiegelhalter D. Markov Chain Monte Carlo in Practice. Abingdon: Chapman and Hall/ CRC/Taylor and Francis, 1995.
- ⁴⁷ Warn DE, Thompson SG, Spiegelhalter DJ. Bayesian random effects meta-analysis of trials with binary outcomes: methods for the absolute risk difference and relative risk scales. *Stat Med* 2002;**21**:1601–23.
- ⁴⁸ Bolla MK, Wood N, Humphries SE. Rapid determination of apolipoprotein E genotype using a heteroduplex generator. J Lipid Res 1999;**40**:2340–45.
- ⁴⁹ Abdollahi MR, Guthrie PAI, Davey Smith G, Lawlor DA, Ebrahim S, Day INM. Integrated single-label liquid-phase assay of APOE codons 112 and 158 and a lipoprotein study in British women. *Clin Chem* 2006;**52**: 1420–23.
- ⁵⁰ van Vliet P, Mooijaart SP, de Craen AJM, Rensen PCN, van Heemst D, Westendorp RGJ. Plasma levels of apolipoprotein E and risk of stroke in old age. *Ann NY Acad Sci* 2007;**1100**:140–47.
- ⁵¹ Zhu L, Fratiglioni L, Guo Z *et al*. Incidence of dementia in relation to stroke and the apolipoprotein E epsilon4 allele in the very old. Findings from a population-based longitudinal study. *Stroke* 2000;**31**:53–60.
- ⁵² Haan MN, Mungas DM, Gonzalez HM, Ortiz TA, Acharya A, Jagust WJ. Prevalence of dementia in older latinos: the influence of type 2 diabetes mellitus, stroke and genetic factors. J Am Ger Soc 2003;51:169–77.
- ⁵³ Fillenbaum GG, Blazer DG, Burchett BM, Saunders AM, Taylor DH. Apolipoprotein E epsilon4 and risk of mortality in African American and white older community residents. *The Gerontologist* 2002;**42**:381–86.
- ⁵⁴ Lawlor D, Bedford C, Taylor M, Ebrahim S. Geographical variation in cardiovascular disease, risk factors, and their control in older women: British Women's Heart and Health Study. J Epidemiol Community Health 2003;57: 134–40.
- ⁵⁵ Cholesterol Treatment Trialists' Ctt Collaborators. The effects of lowering LDL cholesterol with statin therapy in people at low risk of vascular disease: meta-analysis of individual data from 27 randomised trials. *Lancet* 2012; 6736:1–10.
- ⁵⁶ Hixson JE, Vernier DT. Restriction isotyping of human apolipoprotein E by gene amplification and cleavage with HhaI. J Lipid Res 1990;**31**:545–48.

- ⁵⁷ Crook R, Hardy J, Duff K. Single-day apolipoprotein E genotyping. J Neurosci Methods 1994;**53:**125–27.
- ⁵⁸ Appel E, Eisenberg S, Roitelman J. Improved PCR amplification/Hhal restriction for unambiguous determination of apolipoprotein E alleles. *Clin Chem* 1995;**41**:187–90.
- ⁵⁹ Ebrahim S, Sung J, Song Y-M, Ferrer RL, Lawlor DA, Davey Smith G. Serum cholesterol, haemorrhagic stroke, ischaemic stroke, and myocardial infarction: Korean national health system prospective cohort study. *BMJ* 2006; **333:**22.
- ⁶⁰ Mahley RW, Weisgraber KH, Huang Y. Apolipoprotein E4: a causative factor and therapeutic target in neuropathology, including Alzheimer's disease. *Proc Natl Acad Sci* U S A 2006;**103**:5644–51.
- ⁶¹ Drenos F, Talmud PJ, Casas JP *et al.* Integrated associations of genotypes with multiple blood biomarkers linked to coronary heart disease risk. *Hum Mol Genet* 2009;**18**:2305–16.
- ⁶² Lawlor DA, Davey Smith G, Harbord R, Timpson N, Day I, Ebrahim S. Clustered environments and randomized genes: a fundamental distinction between conventional and genetic epidemiology. *PLoS Med* 2007;4:e352.
- ⁶³ Talmud P, Drenos F, Shah S *et al.* Gene-centric association signals for lipids and apolipoproteins identified via the HumanCVD BeadChip. *Am J Hum Genet* 2009;**85**: 628–42.
- ⁶⁴ Kathiresan S, Melander O, Guiducci C *et al.* Six new loci associated with blood low-density lipoprotein cholesterol, high-density lipoprotein cholesterol or triglycerides in humans. *Nat Genet* 2008;**40**:189–97.
- ⁶⁵ Kathiresan S, Willer CJ, Peloso GM *et al.* Common variants at 30 loci contribute to polygenic dyslipidemia. *Nat Genet* 2009;**41:5**6–65.
- ⁶⁶ Roses AD, Lutz MW, Amrine-Madsen H *et al*. A TOMM40 variable-length polymorphism predicts the age of late-onset Alzheimer's disease. *Pharmacogenomics J* 2010; 10:375–84.
- ⁶⁷ Abraham R, Moskvina V, Sims R *et al*. A genome-wide association study for late-onset Alzheimer's disease using DNA pooling. *BMC Med Genomics* 2008;1:44.
- ⁶⁹ Stengård JH, Frikke-Schmidt R, Tybjaerg-Hansen A, Nordestgaard BG, Sing CF. Variation in 5' promoter region of the APOE gene contributes to predicting ischaemic heart disease (IHD) in the population at large: the Copenhagen City Heart Study. *Ann Hum Genet* 2007;**71**: 762–71.
- ⁷⁰ Stengård JH, Clark AG, Weiss KM *et al.* Contributions of 18 additional DNA sequence variations in the gene encoding apolipoprotein E to explaining variation in quantitative measures of lipid metabolism. *Am J Hum Genet* 2002; **71**(3):501–17.t.
- ⁷¹ Mahley R. Apolipoprotein E: cholesterol transport protein with expanding role in cell biology. *Science* 1988;240: 622–30.
- ⁷² van den Maagdenberg AM, Weng W, de Bruijn IH *et al.* Characterization of five new mutants in the carboxylterminal domain of human apolipoprotein E: no cosegregation with severe hyperlipidemia. *Am J Hum Genet* 1993; **52:**937–46.
- ⁷³ Mailly F, Xu CF, Xhignesse M *et al*. Characterization of a new apolipoprotein E5 variant detected in two French-Canadian subjects. *J Lipid Res* 1991;**32**:613–20.

- ⁷⁴ Wellcome Trust Case Control Consortium. Genome-wide association study of 14000 cases of seven common diseases and 3000 shared controls. *Nature* 2007;447: 661–78.
- ⁷⁵ C Reactive Protein Coronary Heart Disease Genetics Collaboration. Association between C reactive protein and coronary heart disease: mendelian randomisation analysis based on individual participant data. *BMJ* 2011;**342:**d548.
- ⁷⁶ Amarenco P, Labreuche J, Lavallée P, Touboul P-J. Statins in stroke prevention and carotid atherosclerosis: systematic review and up-to-date meta-analysis. *Stroke* 2004;35: 2902–09.
- ⁷⁷ Amato M, Montorsi P, Ravani A *et al*. Carotid intimamedia thickness by B-mode ultrasound as surrogate of coronary atherosclerosis: correlation with quantitative

coronary angiography and coronary intravascular ultrasound findings. *Eur Heart J* 2007;**28**:2094–101.

- ⁷⁸ de Groot E, Hovingh GK, Wiegman A *et al*. Measurement of arterial wall thickness as a surrogate marker for atherosclerosis. *Circulation* 2004;**109(Suppl. 1)**:III33–38.
- ⁷⁹ Paternoster L, Martinez-Gonzalez NA, Charleton R, Chung M, Lewis S, Sudlow CLM. Genetic effects on carotid intima-media thickness: systematic assessment and meta-analyses of candidate gene polymorphisms studied in more than 5000 subjects. *Circulation. Cardiovascular genetics* 2010;**3**:15–21.
- ⁸⁰ Voight BF, Peloso GM, Orho-Melander M *et al.* Plasma HDL cholesterol and risk of myocardial infarction: a mendelian randomisation study. *Lancet* 2012;**380**: 572–80.