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Long term associations between cholinesterase inhibitors and memantine use and health outcomes among patients with Alzheimer's disease

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Abstract

BACKGROUND—Randomized-controlled trials that examine the effects of Cholinesterase inhibitors (ChEI) and memantine on patient outcomes over long periods of time are difficult to conduct. Observational studies based on practice-based populations outside the context of controlled trials and open label extension studies that evaluate the effects of these medications over time are limited.

OBJECTIVES—To examine in an observational study (1) relationships between ChEI and memantine use and functional and cognitive endpoints and mortality in AD patients, (2) relationships between other patient characteristics on these clinical endpoints, and (3) whether effects of the predictors change across time.

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DESIGN—Multicenter, natural history study.

SETTING—Three university-based AD centers in the US.

PARTICIPANTS—201 patients diagnosed with probable AD with modified Mini-Mental State Examination scores of 30 or higher at study entry followed annually for 6 years.

MEASUREMENTS—Discrete-time hazard analyses were used to examine relationships between ChEI and memantine use during the previous 6 months reported at each assessment and time to cognitive (Mini-Mental State Examination, MMSE 10) and functional (Blessed Dementia Rating Scale, BDRS 10) endpoints and mortality. Analyses controlled for clinical characteristics including baseline cognition, function, and comorbid conditions, and presence of extrapyramidal signs and psychiatric symptoms at each assessment interval. Demographic characteristics included baseline age, sex, education, and living arrangement at each assessment interval.

RESULTS—ChEI use was associated with delayed time in reaching functional endpoint and death. Memantine use was associated with delayed time to death. Different patient characteristics were associated with different clinical endpoints

CONCLUSION—Results suggest long term beneficial effects of ChEI and memantine on patient outcomes. As for all observational cohort study, observed relationships should not be interpreted as causal effects.

Keywords

Alzheimer's disease; cholinesterase inhibitors; memantine; outcomes; longitudinal studies

INTRODUCTION

Since their introduction, cholinesterase inhibitors (ChEIs) and later the N-methyl-D aspartate receptor antagonist (memantine) have been shown in short-term clinical trials and longer-term open-label extension studies to stabilize or reduce the rate of decline in measures of cognitive function, activities of daily living, and behavior in some patients with Alzheimer's disease (AD).^[1-11] Most rigorous evidence of whether the effects of ChEIs and memantine are sustained over longer periods of time would come from long-duration, prospective, placebo-controlled trials. However, such trials are not only costly to conduct, there also are ethical concerns associated with exposing patients to placebo in trials of long duration because ChEIs and memantine have become standard of care for patients with AD. In the absence of these trials, observational studies based on practice-based populations may be one of the only ways to evaluate effects of these medications.^[12]

Several studies have assessed the effects of ChEIs and/or memantine treatment in real-world clinic settings.^[13-22] Results from these studies have been mixed. In one of the first observational studies on the effects of ChEI on patient outcomes, Doody and others found slowed decline in cognitive function after a year in patients treated with ChEIs compared to untreated patients.^[13] Comparing patients treated with ChEI or ChEI+memantine combination therapy of to an untreated earlier cohort of patients, Atri and others also reported slower decline in cognition and function in the treated group.^[17] Persistent treatment has been shown to be associated with slowed decline in cognition and function, but effects may be lost if treatment were disrupted.^[3, 18] On the other hand, also comparing a cohort of patients treated with ChEI with an earlier cohort of untreated patients, Lopez and others reported no association between ChEI use and time to cognitive and functional decline or to death, but significant delays in nursing home admission.^[14, 19]

In an earlier study using a large, multi-center cohort of patients with probable AD who were prospectively followed up to 6 years from early disease stages, we reported that patterns of ChEI and memantine use changed substantially over time and were consistent with practice guidelines of initiating ChEIs in mild to moderate AD and adding memantine in moderate to severe AD.^[23] In the current study, we take advantage of the availability of important clinical characteristics (e.g., comorbid conditions, psychiatric symptoms) that were not controlled for before, the long follow up period, and more current data, and further investigated the following questions: (1) are ChEI or memantine use associated with length of time to reach cognitive and functional outcomes and death, and (2) are these associations stable over time?

METHODS

Sample

Data are drawn from the Predictors 2 cohort, consisting of patients recruited from Columbia University Medical Center, Johns Hopkins School of Medicine, and Massachusetts General Hospital. The study was approved by each local Institutional Review Board. The inclusion/ exclusion criteria have been fully described elsewhere.^[24-26] Briefly, subjects met DSM-III-R criteria for primary degenerative dementia of the Alzheimer type and NINDS-ADRDA criteria for probable AD. Enrollment required a modified Mini-Mental State Examination (mMMS) score 30, equivalent to approximately 16 on the Folstein Mini-Mental State Examination (MMSE).^[27, 28] Clinical diagnosis of AD has been confirmed in 93% of those with postmortem evaluation.^[26]

Study recruitment began in 1997, when widespread use of ChEIs began in the US, and is ongoing. After the baseline assessment, patients were followed annually. Those who missed a particular visit could respond at a subsequent visit. The cohort used in the current analysis included 201 patients who were followed for up to 6 years and provided data for 785 visits. Of these 201 patients, 13 had 6, 27 had 5, 37 had 4, 34 had 3, and 41 had 2 years of follow-up visits. 123 patients (61%) did not miss any visits, 15 (7%) missed 1, 19 (9%) missed 2, 22 (10%) missed 3, and the rest missed 4 or more visits. Median follow-up for the cohort was 4 years (mean=3.5, SD=2.0).

Measures

Clinical Endpoints—We used MMSE to assess patients' cognitive status and constructed a dichotomous variable indicating MMSE 10 at each visit. We used Blessed Dementia Rating Scale (BDRS) Parts I and II (Instrumental and Basic Activities of Daily Living) to assess patients' functional status and constructed a dichotomous variable indicating BDRS 10 at each visit. We chose these cutoff points because similar scores have been used as outcomes in many studies. Exploratory analyses of neighboring end points (i.e., MMSE 8, BDRS 8) did not change estimation results substantively. Patient deaths were typically reported by family members when we attempted to complete follow-up visits. For patients who could not be contacted, information on death was obtained through the National Death Index.

Main Independent Variables: ChEI and Memantine Use—All prescription and over the counter medication use during the previous 6 months were reported at each visit by the patient and informant on a medication acquisition form. Information reported included name of medication, number of days taking the medication, dosage, and number of pills per day. Because ChEIs have been shown to have similar efficacy despite slightly different pharmacological properties, we combined all ChEIs into one group. Because of the consistency of medication use reported in this sample, we constructed dichotomous variables indicating ChEI and memantine use during the 6 months prior to each assessment as our main independent variables instead of using more complex measures that take into account number of days of use or medication dosages.

Other Substantive Control Variables—To isolate the effects of ChEI and memantine use, we controlled for the following time-variant variables in the analysis. Columbia University Scale for Psychopathology in AD (CUSPAD), a semi-structured interview administered by physicians or trained research technicians, was used to measure psychotic symptoms.^[29] A modified Unified Parkinson's Disease Rating Scale (UPDRS) was used to measure presence of extrapyramidal signs (EPS).^[30, 31] Information on patients' living arrangements at each assessment was dichotomized as living at home or in a long-term care facility. The following time-invariant characteristics were included in the analysis: baseline cognition, function, number of comorbidities, baseline age, sex, education, and study site.^[26, 32] Because ChEIs and memantine were approved by the FDA for treatment of AD at different times, year of study entry was included to control for availability of medications on the market and any differences that could be related to different entry times into the study. Because most of the sample was white (n=188, 94%), ethnicity was not included in the models.

Analysis

We used discrete time hazard models to examine the relationships between ChEI and memantine use and time to cognitive and functional endpoints and mortality. We used binary, time-specific event indicators for BDRS, MMSE and death to reflect the first year a patient reached a clinical endpoint. Right-censoring occurred when a patient did not reach a clinical endpoint by their most recent assessment. Such patients would have a 0 on the time-specific event indicator for all assessment intervals up to the most recent assessment. Observations with missing outcomes were dropped from the analysis.

The estimated discrete-time hazard is the conditional probability that a patient will reach a clinical endpoint in an assessment interval, given that they did not reach the endpoint in an earlier interval. Although multiple observations from each patient may be correlated in longitudinal data, this conditional probability can be treated as if it came from a distinct, independent observation, and therefore not necessary to account for clustering effects within the individual.^[33] For each outcome, we first estimated a set of models that included only time effects (year=0, 1, 2... 5). This set of models describes for the entire sample the hazard profile for each clinical endpoint. If the estimated coefficients are approximately the same, the risk of reaching a clinical endpoint is unrelated to time and the hazard function is flat. If the estimated coefficient increases over time, the risk of reaching a clinical endpoint increases over time. We tested the constant hazard assumption using a likelihood-ratio test based on model deviance statistics.

Next, we examined bivariate relationship between each potential substantive covariate (e.g., ChEI use) and the clinical endpoints after controlling for time effects. These main effects models contain the underlying proportionality assumption that each covariate has the same effect on the outcome in every time period. We tested the proportionality assumption of each covariate using deviance statistics obtained from the unconstrained model that included the main effect of the covariate and an interaction term of the covariate and time in the constrained main effects model. If the unconstrained covariate model did not fit better than the main effects for that covariate was included in subsequent analyses. If the unconstrained

covariate model fits better than the main effects model, both the covariate and its interaction with time were included in subsequent analyses.^[34]

Finally, we estimated a full, multivariate model that included (1) all 6 binary time-effects indicators, (2) main effects for covariates that did not violate the proportionality assumption, and (3) main effects and interaction with time effects for covariate that violated the proportionality assumption. To account for possible clustering, e.g., within sites, robust standard errors were reported. The results of this full model allowed interpretation for each individual covariate effect on our clinical endpoints after adjusting for all other covariates.

RESULTS

Baseline Characteristics

At baseline, patients' average age was 76 (SD=8.1), 123 were female (61%), and 176 (88%) lived at home (Table 1). Patients were highly educated, with an average of 14 years of schooling (SD=3.1). Average MMSE was 22.0 (SD=3.4) and average BDRS was 3.6 (SD=2.1). 68 patients had psychotic symptoms (34%) and 29 had extrapyramidal signs (16%). Almost half of the patients (n=96, 48.4%) reported no comorbid conditions, a third (n=67, 34.1%) reported one, and the rest 38 patients (19%) reported two or more comorbid conditions.

Table 2 presents data on the number patients who were available for observation at each assessment interval, and the number and proportion of those who were taking ChEI or memantine. Because only patients with mild AD were included at study entry, four-fifths of the 201 patients (n=161) reported using ChEI at baseline, only 2% (n=4) reported using memantine, all of whom used it in combination with ChEI. By year 6, the proportion of patients who reported using ChEIs slightly decreased to 74.4%, the proportion of patients who reported using memantine increased to 48.8%, and the proportion of patients who reported taking either ChEIs or memantine remained relatively steady. Overall, 182 patients (91%) reported taking ChEIs at some point during the study, 19 (9.5%) were never treated with ChEI. 81 patients (40.2%) reported taking memantine at some point during the study, 120 (59.7%) were never treated with memantine. Patients who were treated with ChEIs at some point were followed for an average of 3.6 years (SD=2.0) compared to 2.1 years (SD=1.7) for those who were never treated with ChEIs. Patients who were treated with memantine at some point were followed for an average of 4.4 years (SD=1.5) compared to 2.9 years (SD=2.0) for those who were never treated with memantine.

Throughout the study period, the proportion of patients who used specific ChEIs remained relatively stable. During the visits in which patients reported taking ChEIs, 61.7% reported taking donepezil, 13.6% galantamine, and 4.8% rivastigmine. We examined in detail the extent and consistency of using individual medications among those who reported taking the medication. Table 3 presents baseline data on the proportion of patients who reported taking each medication every day (180+ days during the previous 6 months) and the proportion of those who took effective dosages for each medication. At baseline, 80.6% of patients who were taking donepezil reported taking it every day, and 98.3% of those who were taking donepezil reported taking it at effective doses. Over the course of the study period, in 93.3% of the visits in which patients reported taking donepezil, they also reported that they took it every day during the previous 6 months, and almost all reported taking it at effective doses.

Unconditional Hazard Models

Figure 1 describes the hazard profile for the entire sample, i.e., the conditional probability that an individual would reach a clinical endpoint given that the individual did not reach the endpoint in an earlier interval. The hazard function for each outcome increased over time,

and was most pronounced in patients' function. Likelihood-ratio tests of the constant hazard assumption was rejected for all outcomes (all p<0.01).

Multivariate Discrete Time Hazard Analyses

Table 4 reports coefficient estimates of final multivariate models that examined time to reach functional and cognitive endpoints and mortality. Consistent with unconditional hazard models, estimated hazard for each outcome significantly increased over time. Year of study entry and site were not significantly associated with any outcome.

After controlling for patient characteristics, ChEI use was associated with longer time to reaching functional endpoint (estimated coefficient=-1.282, SE=0.433, computed odds ratio [OR]=0.278) and mortality (OR=0.409). Memantine use was associated with longer time to mortality (OR=0.443). (Full computations of odds ratios are available upon request.) Better baseline cognition was associated with longer time to reaching all endpoints. For each point improvement in baseline MMSE, estimated odds of reaching functional endpoint was 11% lower, of reaching cognitive endpoint was 21% lower, and of dying was 13% lower. Better baseline function also was associated with longer time to reaching functional endpoint and mortality. A one-point worsening in baseline BDRS almost doubled the estimated odds of reaching functional endpoint (OR=2.0). Similarly, a one-point worsening in baseline BDRS was associated with 36% increased odds of dying (OR= 1.36). Over time, increased risk of reaching functional endpoint and mortality from a one-point increase in baseline BDRS attenuated, but the effect was not statistically significant. Compared to patients without extrapyramidal signs, the odds of reaching functional endpoints for those with extrapyramidal signs was 5.6 times higher and the odds of dying was 2.3 times higher. At each assessment interval, compared to patients without psychiatric symptoms, the odds of reaching functional endpoints for those with psychiatric symptoms were significantly higher, with attenuating effects over time.

Secondary Sensitivity Analyses: Exploring Potential Association between Clinical Endpoints and Irregular ChEI Use

A small subsample of patients (n=19) in this study reported never been treated with ChEIs. Compared to patients who were treated with ChEIs at some point during the study, these patients were older, more likely to be female, with lower education, had shorter years of follow up, and had worse baseline MMSE scores. To test whether our results were sensitive to including patients who were never treated, we re-estimated our models using only patients who were treated with ChEIs at some point during the study. Results were substantively similar.

During the long course of the illness, it is likely that patients may change from treated to non-treated status or vice versa. The following treatment changes were observed in 56 patients: (1) 26 reported taking ChEIs at study entry and later discontinued; (2) 14 reported either taking ChEIs at study entry, discontinued at some point, and resumed treatment later, or were not taking ChEIs at study entry, initiated treatment at some point, and discontinued later; (3) 16 initiated and continued ChEI treatment after study enrollment. Because persistent treatment has been associated with slowed decline in cognition and function,^[3, 18] we explored patterns of discontinuation as related to reaching clinical outcomes in these 56 patients. Bivariate distributions between medication use and clinical endpoints did not show patients were taken off ChEIs as their conditions deteriorated, although power to detect differences was limited. Among the 81 patients who used memantine at some point during the study, the vast majority continued treatment once started, only 6 discontinued after one or two assessments, one of whom reported resuming at a later time.

DISCUSSION

In this study, we assessed long term effects of ChEI and memantine on the natural history of AD outside the context of controlled trials and open label extension studies in order to provide a more 'real world' perspective. A large cohort of patients from early stages of AD was prospectively followed over 6 years. After controlling for important clinical characteristics, we found ChEI use was associated with longer time to reaching functional endpoints and death and memantine use was associated with delayed mortality. As for all observational cohort study, common caution in interpreting results is needed. In particular, the effects of ChEI and memantine use observed in this study should not be interpreted as causal effects.

A small number of studies have examined the relationship between ChEI use and mortality. An early randomized trial on tacrine use reported a trend toward lower mortality among those who received higher doses (>120mg/day) in open label follow up.^[5] Another study reported significantly lower 3-year mortality in tacrine users than nonusers.^[20] However, a more recent observational study comparing patients on ChEIs to a matched, historical sample of patients who were never exposed to ChEIs reported no association between ChEI use and cognitive and functional status at 1-year follow up or 3-year mortality.^[14] Another study reported ChEI use alone, or ChEI+memantine use were not associated with time to death.^[19]

Direct comparison of our results to these studies is difficult due to differences in study sample and methodology. Perhaps the closest study to ours is Lopez 2009, in which 443 patients with probable AD with at least one annual follow-up in a large urban ADRC were followed since 1997.^[19] Several differences in patient characteristics between these two studies should be noted. First, compared to the Lopez study, patients in our sample were on average 3 years older (mean age = 76 vs. 73), but with higher MMSE at study entry (mean MMSE=21.9 vs. 18.2). It is possible that patients in our sample had an earlier start in ChEI therapy, similar to those who participate in clinical trials, and that this may have led to stronger mortality effects observed in our study.^[35-37] Compared to the Lopez study, in which only 3% of subjects were never treated with ChEIs, 10% of our patients, the never treated group in our study was older while they were significantly younger in the Lopez study.

It is also important to note methodological differences between these studies. For example, while we included all recruited patients in our analyses, analysis in the Lopez et al study was restricted to patients with at least one follow-up. Differential rates of decline or differential effectiveness of medications in patients who did not have any follow-up data may have contributed to differences in study results. In several earlier studies, patients who were not treated with ChEI were derived from historical cohorts that may have confounding effects on the results.^[14, 17] It is likely that patients in the treated, more recent cohorts were diagnosed earlier than those in the untreated earlier cohorts. Improvements in medical care in more recent cohorts may have contributed to the relative slowing of disease progression in treated groups. These possibilities should be tested empirically in future studies.

While estimated associations between ChEI and memantine use and clinical endpoints are independent of patient characteristics that we controlled for in the models, unobserved characteristics that may differ between groups may remain and lead to possible biases in our results. It is possible that patients who took these drugs have naturally slower disease progression and their milder disease course was attributed to drug treatment. As in all observational studies, we cannot rule out the possibility of effects influencing observed

Several studies that examined the effects of ChEI+memantine combination therapy reported additional beneficial effects of combination therapy in slowed cognitive and functional decline and delayed nursing home placement.^[17, 19] In this study, the vast majority of patients who used memantine used it in combination with ChEI, with memantine monotherapy reported in only 20 visits from 14 patients. We therefore could not distinguish the effects of memantine monotherapy versus ChEI+memantine combination therapy. However, the statistically significant effects of both ChEI and memantine suggest beneficial effect of ChEI+memantine combination therapy in delaying mortality.

Our ability to estimate duration of medication use was limited by our data collection method. At each visit, patients were asked about medication use in the previous 6 months, however, follow-up was performed annually. Therefore, medication use was known for half of the interval between visits. This loss of information adds noise to estimation models. However, unless there were systematic biases in medication use or discontinuation at 6 month intervals that were not captured in the data, this loss of information should not bias our estimation results.

There are several important strengths of this study. First, there is the large number of well characterized AD patients who were prospectively followed from early disease stages and examined by the same measures, in the same memory disorders units, and by the same group of clinicians over multiple years with good follow-up rates. In contrast to AD clinical trial and open-label extension studies that typically included data of at most one or two years, duration of our study is substantially longer. Second, our study avoids the limitations of most clinical trials which often have stringent inclusion and exclusion criteria. The sample of patients included in this study and results obtained from these patients therefore reflect more real-world clinical practice. Third, this study addresses several methodological issues in longitudinal analyses. As data on clinical outcomes are often collected at discrete-time intervals (e.g., annually), the commonly used continuous-time methods (e.g., Cox models) are limited. We used the more appropriate discrete-time survival models, included both time-invariant and time-variant covariates, and relaxed the often violated proportional hazards assumption when appropriate.^[33, 34, 38] Therefore, our results go beyond findings from short-term clinical trials and provide evidence of the effects of cholinesteraseinhibitors and memantine that is fuller and more nuanced and have real-world clinical relevance in the treatment of patients with AD.

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Systematic review

Randomized-controlled trials that examine the effects of ChEI and memantine on patient outcomes over long periods of time are difficult to conduct. Observational studies are limited with mixed results.

Interpretation

Our results extend beyond the context of controlled trials and open label extension studies, as our sample has less stringent inclusion/exclusion criteria, is more current with substantially longer study duration. Our methodology allows a more nuanced analysis so results have more 'real-world' relevance in the treatment of patients with AD.

Future directions

We will consider instrumental variables or propensity scores approaches to address some study limitations, such as the possible that patients who took ChEI and memantine have naturally slower disease progression and their milder disease course was attributed to drug treatment, or the possibility of other variables influencing observed relationships between treatment and outcomes. We will extend our study from our current clinical sample to community samples.

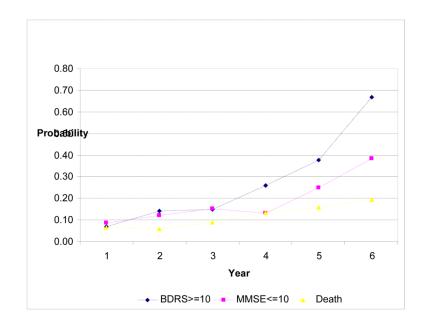


Figure 1. Unconditional Hazard Model for Functional, Cognitive, and Mortality Endpoints

	Table 1
Patient Characteristics at Baselin	e (n=201)

Variables	Mean (SD)	Range
Age	76.3 (8.0)	[49, 95]
Female (%)	61.2	
Years of schooling (%)		
12	42.3	
13-15	15.9	
16	41.8	
Living in a nursing home (%)	12.4	
Folstein Mini-Mental State Examination (MMSE)	21.9 (3.4)	[12, 30]
Modified Mini-Mental State Examination (mMMS)	37.2 (6.5)	[14, 52]
Blessed Dementia Rating Scale (BDRS)	3.6 (2.1)	[0, 10.5]
Presence of psychiatric symptoms (%)	33.8	
Presence of extrapyramidal signs (%)	15.9	
Number of comorbidities	0.8 (1.1)	[0, 8]
Site (%)		
Columbia	44.3	
Johns Hopkins	26.9	
Mass General	28.9	
Year of study entry (%)		
1997-1998	13.2	
1999	7.7	
2000	25.8	
2001	25.3	
2002	13.2	
2003-2007	4.9	

Table 2 Proportion of Patients who Reported Taking ChEI or Memantine at Each Visit

				Year			
	0	1	7	3	4	w	9
Number of patients at each visit	201	159	127	98	93	99	43
Taking ChEI							
Number of patients	161	131	66	81	70	48	32
Proportion of patients	80.1	82.4	78.0	82.7	75.3	72.7	74.4
Taking Memantine							
Number of patients	4	13	19	32	41	34	21
Proportion of patients	2.0	8.2	15.0	32.7	44.1	51.5	48.8
Taking both ChEI and Memantine							
Number of patients	4	12	16	29	33	30	19
Proportion of patients	2.0	7.5	12.6	29.6	35.5	45.5	44.2
Did not take either ChEI or Memantine							
Number of patients	40	27	25	14	15	14	6
Proportion of patients	19.9	17.0	19.7	19.9 17.0 19.7 14.3 16.1	16.1	21.2	20.9

Table 3

Extent and Consistency of ChEI and Memantine Use by Visit among Patients Who Reported Taking the Medication

	Prevalence at Baseline Visit		Study Period Prevalence From All Visits		
Generic name	Percent taking medication for 180+ days	Percent taking effective dose ^a	Percent taking medication for 180+ days	Percent taking effective dose ^{<i>a</i>}	
Donepezil	80.6	98.3	93.3	98.9	
Galantamine	83.3	61.1	88.7	73.6	
Rivastigmine	77.8	77.8	81.1	78.9	
Memantine	88.9 88.9		86.1	96.4	

^aEffective dose per day for Donepezil is defined to be 5mg or 10mg, for Galantamine 16-24mg, Rivastigmine 6-12mg, and for Memantine 10mg or 20mg. 201 patients were included at baseline and contributed to 785 observations for the longitudinal sample throughout the study period.

Table 4 Multivariate Results of Discrete-Time Hazard Models of Time to reach Functional and Cognitive Endpoints and Mortality^a

	Functional Endpoint (BDRS 10)	Cognitive Endpoint (MMSE 10)	Death
	Coef.	Coef.	Coef.
	(SE)	(SE)	(SE)
ChEI use	-1.282 ***	-0.379	-0.894 **
	(0.433)	(0.573)	(0.387)
Memantine use	-0.274	0.500	-0.813 *
	(0.388)	(0.498)	(0.450)
Baseline MMSE	-0.095 *	-0.241 ***	-0.142 ***
	(0.056)	(0.074)	(0.051)
Baseline BDRS	0.707 ***	0.068	0.307 **
	(0.196)	(0.238)	(0.136)
Baseline BDRS x year	-0.085	-0.041	-0.023
Busenne BBRS x year	(0.055)	(0.070)	(0.037)
Presence of EPS	1.727 ***	0.836	0.832 **
	(0.483)	(0.552)	(0.405)
Presence of psychiatric symptoms	1.071	2.258 ***	0.758
The second s	(0.772)	(0.863)	(0.616)
Presence of psychiatric symptoms x year	-0.072	-0.526 **	-0.109
resence of psychiatric symptoms x year			
Baseline comorbidities	(0.209) -0.023	(0.249) -0.288	(0.170) -0.332
Baseline comorbidities	(0.376)	-0.288 (0.465)	-0.332 (0.287)
Baseline comorbidities \times year	-0.036	0.140	0.099
Baseline conforbidities × year	(0.107)	(0.139)	(0.070)
Baseline age	-0.093 ***	-0.116 ***	0.011
č	(0.023)	(0.026)	(0.019)
Female	-0.533	0.384	-0.152
	(0.805)	(0.855)	(0.636)
female \times year	0.357	0.191	-0.031
·	(0.217)	(0.248)	(0.173)
Education	-0.066	0.148	-0.074
	(0.108)	(0.134)	(0.087)
Education × year	0.030	-0.021	0.022
	(0.030)	(0.040)	(0.025)
Living in a nursing home	0.609 *	0.072	0.247
	(0.364)	(0.526)	(0.375)

^aAll models controlled for indicators of assessment interval, year of study entry, and site. Robust standard errors were reported.

* p<.10
** p<.05
*** p<.01