Brain β -amyloid load approaches a plateau

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ABSTRACT

Objective: To model the temporal trajectory of β -amyloid accumulation using serial amyloid PET imaging.

Methods: Participants, aged 70–92 years, were enrolled in either the Mayo Clinic Study of Aging (n = 246) or the Mayo Alzheimer's Disease Research Center (n = 14). All underwent 2 or more serial amyloid PET examinations. There were 205 participants classified as cognitively normal and 55 as cognitively impaired (47 mild cognitive impairment and 8 Alzheimer dementia). We measured baseline amyloid PET-relative standardized uptake values (SUVR) and, for each participant, estimated a slope representing their annual amyloid accumulation rate. We then fit regression models to predict the rate of amyloid accumulation given baseline amyloid SUVR, and evaluated age, sex, clinical group, and *APOE* as covariates. Finally, we integrated the amyloid accumulation rate vs baseline amyloid PET SUVR association to an amyloid PET SUVR vs time association.

Results: Rates of amyloid accumulation were low at low baseline SUVR. Rates increased to a maximum at baseline SUVR around 2.0, above which rates declined—reaching zero at baseline SUVR above 2.7. The rate of amyloid accumulation as a function of baseline SUVR had an inverted U shape. Integration produced a sigmoid curve relating amyloid PET SUVR to time. The average estimated time required to travel from an SUVR of 1.5–2.5 is approximately 15 years.

Conclusion: This roughly 15-year interval where the slope of the amyloid SUVR vs time curve is greatest and roughly linear represents a large therapeutic window for secondary preventive interventions. *Neurology*[®] **2013;80:890-896**

GLOSSARY

AD = Alzheimer disease; **ADRC** = Mayo Alzheimer's Disease Research Center; **CI** = confidence interval; **CN** = cognitively normal; **ICC** = intraclass correlation; **MCI** = mild cognitive impairment; **MCSA** = Mayo Clinic Study of Aging; **PVC** = partial volume-corrected; **ROI** = region of interest; **SUVR** = standardized uptake values.

The most well-established biomarkers of Alzheimer disease (AD) at this time can be divided into 2 major categories: 1) measures of neuronal injury and degeneration¹⁻⁴ and 2) measures of brain amyloid beta (AB) deposition including CSF AB42^{1,2} and PET amyloid imaging.^{5–7} Some of us recently proposed a model describing the temporal evolution of AD biomarkers.⁸ Based on autopsy⁹ and in vivo amyloid PET studies^{10–12} available at the time we hypothesized that amyloid load follows a sigmoid trajectory as a function of time.⁸ A sigmoid magnitude vs time shape implies that the rate of amyloid accumulation in the brain initially accelerates then later decelerates. Our objective in this article was to test the hypothesis that amyloid PET load vs time can be reasonably modeled by a sigmoid function.

METHODS Participants. Participants were drawn from either the Mayo Clinic Study of Aging (MCSA) or the Mayo Alzheimer's Disease Research Center (ADRC). The MCSA (total current enrollment = 1,847) is a longitudinal population-based observational study of cognitive aging that was established in Olmsted County, Minnesota, in 2004. MCSA subjects were 70 years or older and did not have dementia at the time of enrollment. The ADRC (total current enrollment = 421) is a longitudinal observational sample of subjects recruited from the behavioral neurology practice at Mayo Clinic.

To be included, participants must have had 2 or more serial amyloid PET studies, and have clinical diagnoses of cognitively normal (CN), mild cognitive impairment (MCI), or AD dementia. Otherwise there were no other exclusion criteria. The amyloid PET studies

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were obtained over the period March 2006–July 2012. A total of 260 participants met these inclusion criteria (205 CN, 47 MCI, and 8 AD) and all were included in these analyses. Of these 260, 220 participants had 2 imaging visits, 38 had 3, and 2 had 4. Participants were categorized as impaired (defined as either MCI or AD dementia) or CN. All 205 CN participants were from the MCSA. The impaired group (n = 55) was composed of 41 MCI from the MCSA, 6 MCI from the ADRC, and 8 AD dementia from the ADRC.

Standard protocol approvals, registrations, and patient consents. These studies were approved by the Mayo institutional review board and informed consent was obtained from all participants or their surrogates.

PET methods. PET images were acquired using a PET/CT scanner. The ¹¹C Pittsburgh compound B–PET scan consisting of 4 5-minute dynamic frames was acquired from 40 to 60 minutes after injection. Image analysis was done using our in-house fully automated image processing pipeline.¹² A cortical amyloid PET retention ratio (SUVR) was formed by calculating the median uptake over voxels in the prefrontal, orbitofrontal, parietal, temporal, anterior cingulate, and posterior cingulate/precuneus regions of interest (ROIs) for each subject and dividing this by the median uptake over voxels in the cerebellar gray matter ROI of the atlas.¹³ The atlas and image registration steps were based on a 3D T1-weighted volume MRI scan.

The primary analyses were performed with MRI-based partial volume-corrected (PVC) amyloid PET. PVC was accomplished by creating voxel-wise tissue probability maps for each participant's MRI.¹² Each PET voxel was divided by the estimated proportion of gray plus white matter in the voxel. While we did employ PVC for all reported PET analyses, we found the effect of PVC to be negligible. Lin concordance coefficient between PVC and non-PVC amyloid PET SUVR was 0.95 at baseline and 0.94 for annual change in SUVR.

Statistical methods. We used 2-sample *t* tests and χ^2 tests to assess differences between the CN and impaired groups.

We estimated an intraclass correlation (ICC) to evaluate the measurement error associated with longitudinal amyloid PET SUVR measurements by fitting a random-intercepts plus random-slopes linear mixed effect model among CN participants, with time as a fixed effect. This model removes from the repeated measurements within each subject the variability due to a linear rate of change in SUVR and partitions the remaining variation as follows: 1) variability due to intersubject differences in baseline SUVR, denoted by $\sigma^2_{\rm baseline}$; 2) variability due to intersubject differences in rates of accumulation, denoted by $\sigma^2_{\rm rate}$; and 3) residual variability primarily consisting of measurement error, denoted by $\sigma^2_{\rm error}$. We report the ICC as $(\sigma^2_{\rm baseline} + \sigma^2_{\rm rate})/(\sigma^2_{\rm baseline} + \sigma^2_{\rm error})$ and report the residual standard error $\sigma_{\rm error}$ by itself and scaled by the median SUVR of 1.38, using this last quantity as our metric of relative measurement error.

For our primary analysis, using all available time points with amyloid PET SUVR measured, we fit a slope for each subject to represent the annual amyloid accumulation rate. We then used linear regression models to predict these subject-specific rates of amyloid accumulation given the participants' baseline amyloid PET SUVR. Age, sex, *APOE* genotype, and baseline clinical diagnosis were evaluated as covariates. Baseline amyloid PET SUVR was modeled with a restricted cubic spline with 4 knots to allow the accumulation rate to vary nonlinearly with baseline SUVR. The knots were placed at SUVR levels of 1.3, 1.5, 2.0, and 2.5 and chosen to allow for flexibility at SUVR that were interpretable and of interest. In our parameterization of the restricted cubic splines there is a coefficient representing the linear component of the predictor and 2 coefficients representing the nonlinear components. We tested linearity with an *F* test by testing the null hypothesis that both nonlinear coefficients were zero. We used a modified Euler method to solve the first-order differential equation for amyloid accumulation rate vs baseline amyloid PET SUVR as described by the fitted spline function and subject to the condition that time would be set to 0 for an SUVR of 1.3. In this process, we had fixed SUVR, and solved for times between those values. The resulting integral curve was smooth and continuous, with amyloid PET SUVR on the vertical axis, and predicted time from a value of 1.3 on the horizontal axis.

RESULTS Age, sex, or education did not differ between CN and impaired participants (table). As expected, the impaired group had a higher proportion of *APOE* ε 4 carriers and had a lower median Mini-Mental State Examination than the CN group. Baseline amyloid PET SUVR was higher in impaired participants (p < 0.001) and in *APOE* ε 4 carriers (p < 0.001). Age and sex were not associated with baseline amyloid load.

In our mixed effects model analysis, we observed a high ICC of 0.99. The residual standard error was found to be $\sigma_{error} = 0.038$. This value, when divided by the median SUVR among controls of 1.38, is approximately 3%. An ICC of 0.99, coupled with a relative measurement error of 3%, suggests that measurement error was quite small.

Including all participants, the rate of amyloid accumulation as a function of baseline amyloid PET SUVR had an inverted U shape (p < 0.001, test of nonlinearity) (figure 1). Rates of accumulation were low at low baseline amyloid PET SUVR. Rates increased to a maximum at baseline SUVR around 2.0 (95% confidence interval [CI] 1.9–2.1), above which rates declined—reaching amyloid accumulation rates near zero at baseline SUVR greater than 2.7.

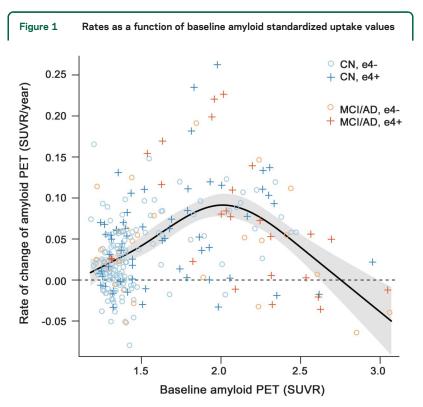
We evaluated age, sex, APOE genotype, and clinical diagnosis as potential covariates in the model predicting rate of accumulation as a function of baseline SUVR. Age was not associated with rate of amyloid accumulation. A trend for higher rates of amyloid accumulation among men (p = 0.06) was attenuated after including baseline amyloid PET SUVR in the model (p = 0.11). Similarly, a trend for higher rates of amyloid accumulation among impaired vs CN participants (p = 0.06) was attenuated after including baseline amyloid PET SUVR (p = 0.23) in the model. Rates of amyloid accumulation were greater in APOE ε 4 carriers than noncarriers (p < 0.001) and remained nearly significant after including baseline amyloid PET SUVR in the model (p = 0.05). There were no significant interactions with age, sex, APOE genotype, or clinical diagnosis and baseline amyloid PET SUVR on rate of amyloid accumulation. Therefore, our final model did not include any of these covariates.

In figure 2A, we plot the rate of amyloid accumulation per year as a function of baseline amyloid load in SUVR units. In figure 2B, we plot the integral with respect to time of the data in figure 2A, which models amyloid level in SUVR units as a function of

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Table Characteristics overall and by clinical diagnosis			
Characteristic	Overall (n = 260)	CN (n = 205)	MCI/AD (n = 55)
Age, y, median (IQR) (min, max)	79 (75, 83) (70, 94)	78 (75, 83) (71, 94)	81 (77, 83) (70, 92)
Male, n (%)	162 (62)	124 (60)	38 (69)
AD diagnosis, n (%)	8 (3)	-	8 (15)
Education, y, median (IQR) (min, max)	14 (12,16) (7, 20)	14 (12,16) (8, 20)	14 (12,16) (7, 20)
APOE ɛ4 positive, n (%)	87 (33)	62 (30)	25 (45)
MMSE, median (IQR) (min, max)	28 (27, 29) (6, 30)	28 (27, 29) (23, 30)	25 (24, 27) (6, 30)
Baseline amyloid SUVR, median (IQR) (min, max)	1.40 (1.31, 1.83) (1.18, 3.06)	1.38 (1.30, 1.62) (1.18, 2.95)	1.94 (1.40, 2.29) (1.22, 3.06)
No. of scans			
2	220 (85)	180 (88)	40 (73)
3	38 (15)	24 (12)	14 (25)
4	2 (1)	1 (0)	1 (2)
Years between first and last scan, median (IQR) (min, max)	1.3 (1.3, 1.5) (0.9, 5.2)	1.3 (1.2, 1.4) (0.9, 5.2)	1.4 (1.3, 2.0) (0.9, 4.1)
Annual change in amyloid SUVR, median (IQR) (min, max)	0.030 (0.002, 0.063) (-0.080, 0.263)	0.026 (0.001, 0.061) (-0.080, 0.263)	0.048 (0.003, 0.082) (-0.064, 0.226)

Abbreviations: AD = Alzheimer disease; CN = cognitively normal; IQR = interquartile range; MCI = mild cognitive impairment; MMSE = Mini-Mental State Examination; SUVR = standardized uptake values.

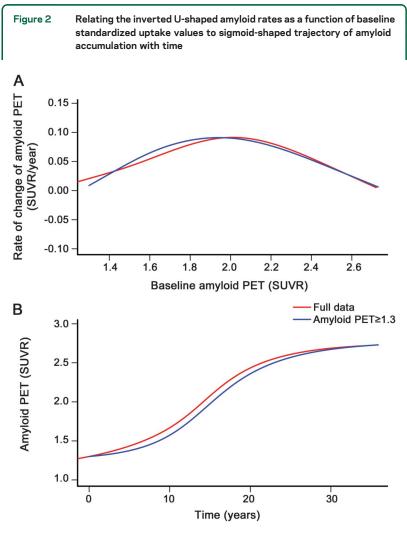


Scatterplot of annual rate of change of amyloid PET standardized uptake values (SUVR) vs baseline amyloid PET SUVR for all subjects (n = 260). The estimated mean slope from a linear regression model is shown with 95% confidence limits. Age, sex, diagnosis, and APOE genotype were not significant covariates in the model and are therefore not included here. Cognitively normal (CN) participants who are APOE ϵ 4– are represented by light blue circles, CN participants who are APOE ϵ 4– are represented by dark blue plus signs, impaired participants who are APOE ϵ 4– are represented by light orange circles, and impaired participants who are APOE ϵ 4+ are represented by light orange plus signs. AD = Alzheimer disease; MCI = mild cognitive impairment.

increasing time in years. These plots are provided based on fits among all participants (n = 260, in red) and also among only those participants whose baseline amyloid PET was 1.3 SUVR or greater (n = 212, in blue). However, truncating the lower limit at 1.3 had little effect on either the plots of rate of amyloid accumulation vs baseline amyloid SUVR, or of amyloid SUVR vs time (figure 2, A and B). For both the lower truncated (at 1.3 SUVR) and nontruncated analyses, the upper SUVR in the model was truncated at 2.7 due to sparseness of the data at SUVR values higher than this.

The findings above imply a sigmoidal relationship between amyloid load and time. This sigmoidal relationship is not directly tied to a specific age—1 subject might start increasing on the curve at 70, while another might start increasing on the curve at 80. Consequently, the x axes in figures 2B and 3 are in units of time, not age. Our data suggest that the midpoint in the sigmoid function (i.e., where rates of accumulation reach a maximum and beyond which begin to decelerate) occurs at amyloid PET SUVR around 2.0 (95% CI 1.9–2.1) and that amyloid load approaches a horizontal asymptote (i.e., plateaus) at SUVR somewhat above 2.7 (figure 2).

Symmetric bracket points in the sigmoid curve around the midpoint of 2.0 where the rate is maximum, and over which the function is fairly linear, occur at SUVR of roughly 1.5 and 2.5. In figure 3, we provide estimates of the time (in years) required to travel from a baseline amyloid PET SUVR of 1.5 to a series of higher values among those participants whose baseline SUVR was greater than or equal to 1.3. The average



(A) The rate of amyloid accumulation per year as a function of baseline amyloid load in standardized uptake value (SUVR) units. The red plot in A is the same as the plot in figure 1, except for the upper bound truncation (see below). (B) The integral with respect to time of the data in (A) which models amyloid level in SUVR units as a function of increasing time in years. These plots are provided for all participants (n = 260, in red) and for only those participants whose baseline amyloid PET was 1.3 SUVR or greater (n = 212, in blue). For both the lower truncated (at 1.3 SUVR) and nontruncated analyses, the upper SUVR in the model was truncated at 2.7.

estimated time required to travel from an SUVR of 1.5– 2.5 is approximately 15 years.

For reference purposes, we plotted within-participant trajectories (spaghetti plots) of amyloid accumulation as a function of age among all 195 participants (figure 4).

DISCUSSION Ideal conditions for modeling AD biomarker trajectories require collecting a complete data battery at multiple time points in many individuals over the entire course of the disease. Given that the course of the disease evaluable by currently available biomarkers may span 30 or more years it will take decades to collect such an idealized dataset. An alternative, however, is to piece together shorter time interval data in many participants to create plausible long-term disease models. This is only possible, however, if the individual shortinterval series of within-subject observations can be pieced together in a way that places each subject's observations in an appropriate order along a continuous measure of pathophysiologic severity. We believe that indexing individual participants by amyloid PET SUVR provides a valid solution to this problem in the current context where we are evaluating an amyloid biomarker. While each individual subject's PET trajectory was assumed to be linear, this does not preclude making nonlinear inferences about amyloid deposition vs time over the duration of the disease. The median time from first to last scan was 1.3 years, which represents a tiny fraction of the total duration of the disease. As the sampling interval becomes shorter, all nonlinear functions can be approximated locally by a linear fit.

A different approach to AD biomarker modeling that we and others have pursued previously is to index participants by degree of cognitive impairment.^{14–17} This approach is flawed, however, because all cognitive tests have floor and ceiling effects. Cognitive function is also mediated by cognitive reserve, rendering the relationship between biomarker measures of underlying brain pathophysiology and cognition variable.^{18–20}

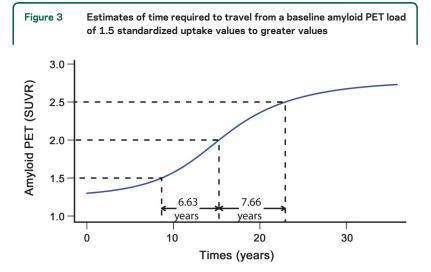
Age likewise is not a reasonable way to measure AD pathophysiologic severity. In a given elderly cohort, some participants will have entered the AD pathophysiologic pathway and some will not. Among those who have, individual participants will have entered at different ages. Therefore combining different participants who are indexed by age along the x axis does not model amyloid load as a function of disease severity and may obscure fundamental relationships. For example, while we see evidence of acceleration then deceleration in rates of amyloid accumulation when participants are indexed along the disease severity pathway by baseline amyloid SUVR (figures 1 and 2A), this relationship is not evident when participants are indexed by age, as in figure 4.

Of the major AD biomarkers, only those that measure brain amyloid are specific for AD pathophysiology.²¹⁻²⁴ Biomarkers of AD neurodegeneration are also abnormal in non-AD conditions including head trauma and stroke.3 This specificity of amyloid PET allowed us to assess the relationship between baseline amyloid SUVR and rate of change in participants who we were more confident were in the AD pathophysiologic path. Based on recent studies favoring lenient amyloid PET cutpoints^{23,25} to detect earliest evidence of brain amyloidosis, we used a cutpoint of 1.3 SUVR to select participants (n = 212) for this subanalysis; however, the results in figures 2 and 3 were not noticeably different with higher cutpoints of 1.4-1.5. In addition, as illustrated in figure 2, results were only slightly different when all participants vs only those with baseline SUVR greater than 1.3 were included.

Our data in figure 4 are similar to other reports where amyloid PET magnitude is plotted as a function

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Predicted times to reach amyloid PET values starting at amyloid PET of 1.5

Time (years)	
0.00	
1.79	
3.19	
4.40	
5.53	
6.63	
7.77	
8.99	
10.40	
12.09	
14.29	
17.51	

This is illustrated in the plot of amyloid PET load vs time in participants whose baseline standardized uptake value (SUVR) was greater than or equal to 1.3 (n = 212, in blue from figure 2B).

of age.^{7,26–29} The fact that, unlike some others,^{6,26,29,30} we did not see an overall increase in baseline SUVR or rate of amyloid accumulation with age may be due to the absence of individuals younger than 70 years in our dataset. Recent autopsy and amyloid PET data^{6,28,29} indicate that the late 50s is the age when some subjects, particularly *APOE* ε 4 carriers,^{6,26} first demonstrate significant amyloid pathophysiology.

Baseline amyloid PET SUVR and rates of amyloid accumulation were greater in *APOE* ε 4 carriers and in impaired participants. These data were largely consistent with existing literature.^{6,23,26,27,29,31–33} However, with baseline amyloid PET SUVR in the model, the rate of amyloid accumulation was not associated with clinical group and, while close, this association did not meet significance at p < 0.05 for *APOE* either. This implies that the rate of amyloid accumulation is more closely related to amyloid load than to *APOE* ε 4, cognitive impairment, or age.

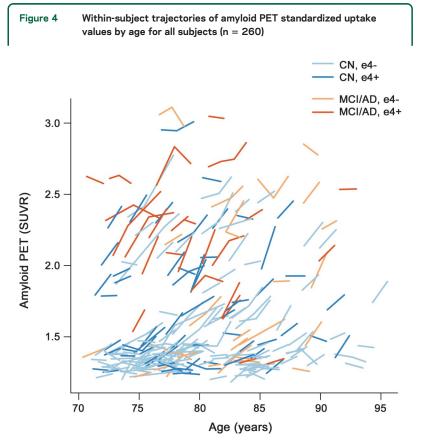
Some of our subjects with higher amyloid PET SUVR were impaired while others were not. This is consistent with the concept that the relationship between a given amount of amyloid deposition and clinical impairment is indirect. Cognitive impairment is a direct consequence of neurodegeneration while amyloid deposition is an earlier, "upstream" pathophysiologic event.11,34-37 Moreover, the ultimate clinical result of a given amount of amyloid can be modified by many factors. For a given amount of amyloid, cognitive decline may occur sooner in participants with risk-enhancing exposures such as comorbid brain pathologies, low education, or risk amplification genes. Likewise, decline may be delayed in participants with risk-reducing exposures such as high education¹⁸⁻²⁰ or protective genes.

In figure 2, it is evident that average rates decline monotonically with baseline SUVR for baseline SUVR greater than 2.0. The downward trend in rate above baseline SUVR of 2.0 is primarily driven by the 47 subjects with baseline SUVR between 2.0 and 2.7, not the few subjects with baseline SUVR greater than 2.7. Specific absolute SUVR will likely vary with different PET ligands and different PET image analysis approaches.^{22,23,29,38} We believe, however, that our conclusions about the biology of AD should not be affected by methodologic variation.

An ICC of 0.99 coupled with a relative measurement error of 3% suggests that the precision of our rate measurements seems more than adequate. Questions may also arise concerning bias in rate measures in impaired subjects. But of the 51 subjects with baseline SUVR greater than 2.0, 25 were CN and 26 were impaired. Thus, the observed decline in rates of amyloid accumulation with baseline SUVR above 2.0 can not be attributed to bias from rate measures only in impaired subjects in this range.

Rates of change were negative in all 4 participants with baseline SUVR greater than 2.7 (figure 1), which might imply that absolute amyloid load declines over time at very high baseline levels. Because the CI modeling rate as a function of baseline SUVR includes 0 at the highest SUVR (figure 1) though, our data do not support a conclusion other than that rates seem to approach zero at the upper range of observed amyloid load. An accumulation rate of 0 by definition implies a plateau in absolute amyloid load.

A sigmoid-shaped trajectory of brain amyloid burden as a function of time could be attributed to 2 different effects. One relates to sensitivity limits of any measurement technique at extremes. Floor (detection) and ceiling (saturation) measurement sensitivity effects would impart a sigmoid shape to a data distribution. A second more profound explanation is that this is a fundamental biologic phenomenon. A sigmoid-shaped function with respect to time would be



Cognitively normal (CN) participants are represented by blue lines and impaired participants by orange lines. Lighter colored lines represent APOE ϵ 4– participants and darker colors represent APOE ϵ 4+ participants. AD = Alzheimer disease; MCI = mild cognitive impairment.

consistent with reaching a state of equilibrium at high amyloid load.⁹

The sigmoidal nature of amyloid accumulation with time also has implications for formulating treatment strategies.^{39,40} The data in figure 3 imply a roughly 15-year interval from an SUVR of 1.5 to 2.5. This period where the slope of the amyloid SUVR vs time curve is greatest and roughly linear represents a large therapeutic window for secondary preventive interventions. However, our data also imply that therapeutic interventions designed to reduce the rate of new amyloid deposition (note the distinction from removing previously deposited amyloid) may be less effective in patients who have already reached plateau levels of amyloid deposition.

AUTHOR CONTRIBUTIONS

Clifford R. Jack: drafting/revising the manuscript, study concept or design, analysis or interpretation of data, contribution of vital reagents/tools/ patients, acquisition of data, statistical analysis, study supervision, obtaining funding. Heather Wiste: drafting/revising the manuscript, analysis or interpretation of data, statistical analysis. Timothy G. Lesnick: drafting/revising the manuscript, study concept or design, analysis or interpretation of data, statistical analysis. Stephen D. Weigand: drafting/revising the manuscript, analysis or interpretation of data, statistical analysis. David S. Knopman: drafting/revising the manuscript, study concept or design, analysis or interpretation of data, acquisition of data. Prashanthi Vemuri: drafting/revising the manuscript, analysis or interpretation of data. Vernon Shane Pankratz: drafting/revising the manuscript, study concept or design, analysis or interpretation of data, statistical analysis. Matthew L. Senjem: drafting/revising the manuscript, analysis or interpretation of data, contribution of vital reagents/tools/patients, acquisition of data, statistical analysis. Jeffrey L. Gunter: drafting/revising the manuscript, analysis or interpretation of data, statistical analysis. Michelle M. Mielke: drafting/revising the manuscript, analysis or interpretation of data. Val J. Lowe: drafting/revising the manuscript, study concept or design, analysis or interpretation of data, contribution of vital reagents/tools/patients, acquisition of data, study supervision, obtaining funding. Bradley F. Boeve: analysis or interpretation of data, acquisition of data. Ronald C. Petersen: drafting/revising the manuscript, obtaining funding.

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