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Statins Are Associated with Reduced Risk of Esophageal Cancer, Particularly in Patients with Barrett's Esophagus: A Systematic Review and Meta-Analysis

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Abstract

Background & Aims—Esophageal cancer is increasing in incidence in US, especially among patients with Barrett's esophagus (BE). Statins might prevent this cancer. We performed a systematic review with meta-analysis of studies that evaluated the effect of statins on the risk of esophageal cancer.

Methods—We conducted a systematic search of Medline, Embase, and Web of Science through August 2012. Studies were included if they evaluated exposure to statins, reported the development of esophageal cancer, and reported relative risks or odds ratios (ORs), or provided data for their estimation. Summary OR estimates with 95% confidence intervals (CIs) were calculated using the random-effects model. The analysis included 13 studies (including a post-hoc analysis of 22 randomized controlled trials) reporting 9285 cases of esophageal cancer among 1,132,969 patients.

Results—Meta-analysis of the studies showed a significant (28%) reduction in incidence of esophageal cancer among patients who took statins (adjusted OR, 0.72; 95% CI, 0.60–0.86), though there was considerable heterogeneity among studies. In analyzing a subset of patients known to have BE (5 studies, 312 esophageal adenocarcinomas [EAC] developed in 2125 patients), statins were associated with a significant (41%) decrease in the risk of EAC, after adjusting for potential confounders (adjusted OR, 0.59; 95% CI, 0.45–0.78) with consistent results

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among all studies. The number needed to treat with statins to prevent 1 case of EAC in patients with BE was 389.

Conclusions—Meta-analysis of existing observational studies indicates that statins protect against esophageal cancer and reduce the risk of EAC in patients with BE.

Keywords

Esophageal cancer risk; cholesterol-lowering drugs; HMG CoA reductase inhibitors; chemoprevention

INTRODUCTION

Esophageal cancer (EC) is the 6th most common cancer in men worldwide with a dismal 5-year survival rate.¹ While the incidence of esophageal squamous cell cancer (ESCC) is declining worldwide, the incidence of esophageal adenocarcinoma (EAC) has been rapidly rising, now forming majority of EC in the Western world. This increase may be partly attributable to the obesity epidemic.^{2, 3}

Barrett's esophagus (BE) is the most significant risk factor for development of EAC, and surveillance and early recognition of high grade dysplasia (HGD) and/or EAC may improve survival.^{4, 5} However, endoscopic surveillance of all patients with BE is expensive and not cost-effective.⁶ Hence, current strategies for improved management of EC are aimed at identifying patients at high risk for progression to EAC and identifying chemopreventive agents.

Statins, or 3-hydroxy-3-methylglutaryl coenzyme A (HMG-CoA) reductase inhibitors, used for primary and secondary prevention of cardiovascular diseases, may decrease the risk of cancers.⁷⁻⁹ *In vitro* and animal studies have shown that in addition to cholesterol reduction, statins have anti-proliferative, pro-apoptotic, anti-angiogenic and immunomodulatory effects, which prevent cancer development and growth.¹⁰ This effect has also been shown in EC cell lines.¹¹⁻¹⁴ Some recent observational studies have shown that statins may be associated with a lower risk of EC, particularly in patients with BE,¹⁵⁻¹⁷ whereas others have shown no beneficial effect.^{18, 19}

To better understand this issue, we performed a systematic review with meta-analysis of existing randomized controlled trials (RCTs) and observational studies that investigated the association between statins and risk of developing EC, in particular, the risk of development of EAC or progression of dysplasia in patients with BE.

METHODS

This systematic review was conducted following guidance provided by the Cochrane Handbook²⁰ and Kanwal and White²¹ and is reported according to the Meta-analysis of Observational Studies in Epidemiology (MOOSE) guidelines.²² The process followed *a priori* established protocol.

Data Sources and Search Strategy

A systematic literature search of Medline (1980 through August 31, 2012), Embase (1988 through August 31, 2012) and Web of Science (1993 through August 31, 2012) databases was conducted by two study investigators (S.S. and P.S.), independently, for all relevant articles on the effect of statin use on the risk of EC in all patients. Keywords used in the search included "HMG-CoA reductase inhibitor(s)", "statin(s)", "atorvastatin", "fluvastatin", "lovastatin", "pravastatin", "rosuvastatin," or "simvastatin," combined with

“cancer” or “neoplasm(s)” or “Barrett’s”. The title and abstract of studies identified in the search were reviewed to exclude studies that did not answer the research question. The full text of each remaining article was read to determine whether it contained information on the topic of interest. We also manually searched review article bibliographies and abstracts of gastrointestinal/oncological society meetings for the years 2003–2012, and consulted with experts in the field. When incomplete information was available, attempts were made to contact the corresponding authors of the studies for additional information.

Study Selection

Studies considered in this meta-analysis were either randomized controlled trials (RCTs) or observational studies that met the following inclusion criteria: (1) evaluated and clearly defined exposure to statins, (2) reported EC incidence and (3) reported relative risks (RR) (for cohort studies) or odds ratio (OR) (for case-control studies) or provided data for their calculation. Inclusion was not otherwise restricted by study size, language or publication type. Articles were excluded from the analyses if there was insufficient data for determining an estimate of RR/OR and a 95% confidence interval (CI), though these were included in the systematic review and described qualitatively. When there were multiple publications from the same population, only data from the most recent comprehensive report were included.

Data Abstraction and Quality Assessment

Data were independently abstracted onto a standardized form by two authors (S.S. and A.G.S.). The following data were collected from each study: study design, time period of study/year of publication, country of the population studied, primary outcome reported type, dose and duration of statin use, information source for exposure measurement, total number of persons in each group (exposed v. not exposed), OR and 95% CIs with and without adjustment for potential confounders, and potential confounders used for adjustment. Data on the following confounding risk factors for EC were extracted from each study: age, sex, race, cigarette smoker, body mass index (BMI), presence of BE (and if present, length of BE segment and grade of dysplasia), other medication use (aspirin/non-steroidal anti-inflammatory drugs [NSAIDs]/proton pump inhibitors [PPIs]) and alcohol use was assessed. Conflicts in data abstraction were resolved by consensus, referring back to the original article.

The methodologic quality of case-control and cohort was assessed by two authors independently (A.G.S. and P.S.) using the Newcastle-Ottawa scale.²³ In this scale, observational studies were scored across three categories: selection (4 questions) and comparability (2 questions) of study groups, and ascertainment of the outcome of interest (3 questions), with all questions with a score of one, except for comparability of study groups, where separate points were awarded for controlling age and/or sex (maximum two points). Any discrepancies were addressed by a joint reevaluation of the original article.

Outcomes assessed

The primary analysis focused on assessing the risk of all EC (EAC + ESCC) among users of statins. In addition, we also performed a subgroup analysis of studies restricted to patients with BE, to assess if statin use modifies the risk of progression to EAC (and/or HGD).

Data synthesis and Analysis

We used the random-effects model described by DerSimonian and Laird to calculate summary OR and 95% CI.²⁴ We assessed heterogeneity between study-specific estimates using 2 methods.^{21, 25} Firstly, the Cochran’s Q statistical test for heterogeneity, which tests the null hypothesis that all studies in a meta-analysis have the same underlying magnitude of

effect, was measured. Because this test is underpowered to detect moderate degrees of heterogeneity,²⁶ a p-value of <0.10 was considered suggestive of significant heterogeneity. Secondly, to estimate what proportion of total variation across studies was due to heterogeneity rather than chance, I^2 statistic was calculated. In this, a value of <30%, 30%–60%, 60%–75% and >75% were suggestive of low, moderate, substantial and considerable heterogeneity, respectively.^{21, 27} Once heterogeneity was noted, between-study sources of heterogeneity were investigated using subgroup analyses by stratifying original estimates according to study characteristics (as described above). In this analysis also, a p-value for differences between subgroups of <0.10 was considered statistically significant (i.e., a value of $p < 0.10$ suggested that stratifying based on that particular study characteristic partly explained the heterogeneity observed in the analysis). We explored potential causes of heterogeneity by stratifying for several methodological and clinical features of studies. This included study design (case-control v. cohort), method of determining exposure to statins (pharmacy prescription database v. patient interview v. electronic medical record [EMR]), study location (North America v. Europe), study setting (hospital-based v. population-based) and risk of bias in exposure and outcome assessment due to lack of sufficient information (abstract only v. fully published manuscript). The presence of publication bias was assessed quantitatively using Begg and Mazumdar adjusted rank correlation test (publication bias considered present if $p < 0.10$)²⁸ and qualitatively, by visual inspection of funnel plots of the logarithm of ORs versus their standard errors.²⁹ All P values were two tailed. For all tests (except for heterogeneity and publication bias), a probability level <0.05 was considered statistically significant.

Since outcomes were relatively rare, OR were considered approximations of relative risk (RR). Adjusted OR, where reported in studies, was used for analysis to account for confounding variables. We also calculated the number needed to treat with statins to prevent 1 case of EC in the general population, as well as in high-risk patients with BE, based on current estimates of incidence of EC.^{30–33}

All calculations and graphs were performed using Comprehensive Meta-Analysis (CMA) version 2 (Biostat, Englewood, NJ).

RESULTS

Search Results

Of a total 2336 unique studies identified using our search criteria, 13 studies fulfilled our inclusion criteria and were included in the meta-analysis (7 case-control, 5 cohort, 1 post-hoc analysis of 22 RCTs),^{15–19, 24, 34–41} 3 of which had been published only in the abstract form.^{34, 37, 39} Figure 1 summarizes the process of study identification, inclusion and exclusion. These cumulatively reported 9,285 cases of EC in 1,132,969 patients, primarily derived from a cohort of subjects seen in primary care clinics (with or without known esophageal or cardiovascular diseases). Five of these studies were performed in patients with known BE, and evaluated the risk of development of EAC (except 2 studies which combined risk of development of HGD and/or EAC or any dysplasia and/or EAC).^{15–18, 39} These cumulatively reported 312 cases of EAC in 2,125 patients with BE. Seven studies reported data from 3 population databases, hence only the most comprehensive reports from these were included, i.e., 1 study out of 2 from the QResearch database, UK,^{19, 42} 2 out of 3 studies from the General Practice Research Database, UK,^{34, 36, 43} and 1 out of 2 studies from the Veterans' Affairs national database, USA were included.^{15, 44}

Characteristics of Included Studies

The characteristics of these studies are shown in Table 1. The earliest study period began in 1983, and the latest ended in 2010. While studies in patients with BE reported development of EAC, studies on all cases of EC provided insufficient data on histological subtypes to study effects of statins on EAC and ESCC separately.

Quality of Included Studies

Seven studies were considered high-quality. Table 1 depicts the methodological quality of all studies. Most studies adjusted for the following confounders: age (10/13), sex (9/13), BMI (4/13), smoking (6/13), NSAID/aspirin use (8/13). In most studies, >90% of all patients were on PPIs. In all these studies, a temporal relation of development of EC to statins was established by excluding cases of EC developing prior to exposure to statins, though there was variable duration of minimum statin use before new incident EC cases were included for analysis (0 months–2 years, where reported).

Risk of Esophageal Cancer – All Studies—Meta-analysis of all studies assessing the risk of EC revealed that the use of statins was associated with a significant 26% reduction in EC incidence (Figure 2A). The results showed considerable heterogeneity across all studies, with a Cochran's Q test $p < 0.01$, and the corresponding $I^2 = 80\%$. There was no evidence of publication bias, both quantitatively (Begg and Mazumdar's $p = 0.46$), and qualitatively, on visual inspection of the funnel plot. This risk reduction in EC incidence with statin exposure persisted after adjusting for potential confounders (adjusted OR, 0.72; 95% CI, 0.60–0.86) (Figure 2B), even though the results still showed substantial heterogeneity (Cochran's Q test $p < 0.01$, $I^2 = 74\%$).

On restricting analysis to only 7 high-quality observational studies (Newcastle-Ottawa score 8),^{15–19, 38, 39} use of statins continued to show a significant protective effect against development of EC (adjusted OR, 0.70; 95% CI, 0.56–0.88). This analysis revealed only moderate heterogeneity (Cochran's Q $p = 0.13$, $I^2 = 40\%$).

Sensitivity Analysis

Given significant heterogeneity in meta-analysis of all included, we performed sensitivity analysis to better understand the source of heterogeneity (Table 2). The method of documenting statin exposure (pharmacy record, EMR or patient interviews) showed a non-significant trend towards explaining heterogeneity in studies. There was no significant difference in the subgroups based on study design, study location, setting or publication type.

To assess whether any one study had a dominant effect on the summary OR, each study was excluded and its effect on the main summary estimate was evaluated. No study markedly affected the summary estimate or p-value for heterogeneity among the other summary estimates. Sufficient data were not available to perform stratified analyses based on age, gender, histologic type of EC or statin dose.

Risk of Esophageal Adenocarcinoma - Barrett's Esophagus—Meta-analysis of 5 studies in patients with BE revealed that use of statins was associated with a statistically significant 43% reduction in development of EAC and/or HGD (unadjusted OR, 0.57; 95% CI, 0.44–0.75) (Figure 3A). These results were very consistent across all included studies, with Cochran's Q test p value 0.54, and the corresponding I^2 statistic being 0%. There was no evidence of publication bias on visual inspection of the funnel plot. This remained after adjustment for potential confounders, including use of NSAIDs/aspirin and length and dysplasia status of BE (adjusted OR, 0.59; 95% CI, 0.45–0.78) (Figure 3B).

Subgroup Analysis

On pooled analysis of 2 studies, a greater protective effect against development of EAC with longer duration of statin use (>5 years) was noted (adjusted OR, 0.44; 95% CI, 0.24–0.78), as compared to shorter duration of statin use (<5 years) (adjusted OR, 0.55; 95% CI, 0.28–1.09), though this difference was not statistically significant.^{16, 17} Meta-analysis of the 2 studies that assessed the combined effect of statin and NSAID/aspirin use on the development of EAC in BE, demonstrated a marked 72% reduction in EAC incidence, which was higher than the use of either class of medications (adjusted OR, 0.28; 95% CI, 0.14–0.56).^{16, 17}

Number Needed to Treat

Due to significant heterogeneity between studies, a single summary estimate for number needed to treat with statins to prevent one case of EC could not be inferred. However, on restricting analysis to studies only on patients with BE, we could estimate the number needed to prevent one case of EAC. In a population-based study of patients with BE, the baseline histology at presentation showed no dysplasia in 73%, low-grade dysplasia in 13%, HGD in 3% and EAC in 9%.⁴⁵ Patients with BE have reported rates of EAC of 6.3, 3.3, 16.9 and 65.8 per 1,000 patient years in all patients with BE, patients with non-dysplastic BE, BE with low-grade dysplasia and BE with high-grade dysplasia, respectively.^{31–33} In these groups, the number needed to treat with statins to prevent one case of EAC per year, would be 389, 741, 146 and 39, respectively, considering the 41% risk reduction with statin use after adjusting for potential confounders.

DISCUSSION

Identification of potential agents for chemoprevention in EC is highly desirable. Aspirin, NSAIDs and PPIs may have some potential chemopreventive effects in patients with BE but they are not without significant side effects.^{46, 47} In this comprehensive meta-analysis of all existing studies in more than 1.13 million patients with 9,285 cases of EC, we found that use of statins is associated with a 28% reduction in the risk of EC, after adjusting for confounding variables, though there was substantial methodological heterogeneity in studies. This effect is more pronounced and consistent in patients with BE, in which statin use is associated with a 41% reduction in risk of neoplastic progression, after adjusting for potential confounders, including use of NSAIDs/aspirin and baseline BE segment length and dysplasia grade. This chemopreventive effect of statins in patients with BE is comparable to the 32% reduction in the risk of EAC seen in patients with BE with aspirin use.⁴⁶

The strengths of this analysis include the comprehensive and systematic literature search, consistency of association between statins and EAC in patients with BE between studies with different designs, its ability to partially evaluate the potential influence of measured confounders on the summary estimates and its evaluation of duration effect. The likelihood of important selection or publication bias in our meta-analysis is small. During the identification and selection process, we did not exclude any article because of methodological characteristics and when sufficient information was not available, we attempted to obtain unpublished results from the corresponding authors. Our results are consistent with a previous small systematic review⁴⁸ Alexandre et al included only 2 studies in patients with BE and 3 studies in general population cohorts. Their estimates for risk were 0.53 (95% CI, 0.36–0.78), and 0.86 (95% CI, 0.78–0.94); respectively.⁴⁸ Likewise, in another meta-analysis by Kuoppala et al, only 1 study on the effect of statins on esophageal cancer was included.⁸ In our analysis, we added numerous additional studies, including data from randomized trials, to bring the evidence base up-to-date, and provide a more robust pooled estimate on the effect of statins on EC risk. With the larger number of studies, we

were able to perform multiple subgroup analyses, evaluate heterogeneity and the presence of publication bias.

Antineoplastic Effects of Statins

In vitro and animal studies have shown that statins exert antineoplastic effects through both HMG-CoA reductase-dependent and HMG-CoA reductase-independent pathways (Figure 4). The primary mechanism of action of statins on cholesterol reduction is by competitive inhibition of HMG-CoA reductase. This prevents post-translational prenylation of the Ras/Rho superfamily, which are important mediators of cell growth, differentiation and survival.¹⁰

This anti-neoplastic effect of statins has also been demonstrated in human-derived EC cell lines through inhibition of Ras farnesylation, ERK and Akt signaling pathways as well as by inhibiting COX-2, which are expressed in both EAC and ESCC.^{11, 12, 14} Using three different EAC cell lines, Ogunwobi and colleagues have shown that statins inhibit cell proliferation and induce apoptosis in a concentration dependent manner, additive to inhibition of COX-2 with either NS-398 or celecoxib. Statins up-regulated the proapoptotic proteins Bad and Bax in these EAC cell lines.¹¹ These studies suggest that statins may offer exciting potential as chemopreventive agents in EC, especially, BE.⁴⁹ Unfortunately, there are no animal models that effectively mimic human disease in terms of risk factors, molecular pathogenesis and response to interventions.⁵⁰

Dose and Duration of Statin Use

A meta-analysis of effects of different doses of statins was not possible due to limited information from individual studies. Data from the QResearch database as well as the General Practice Research Database, as well as from a cohort of patients with BE demonstrate that higher dose of statins (simvastatin 40mg or equivalent) may be more beneficial for both EAC and ESCC, than low dose statin (simvastatin <40mg per day).^{17, 42, 43} In the individual patient data analysis of 5 RCTs conducted by the CTT collaboration, there was no significant difference in incidence of EC with 'more' or 'less' statin use (EC incidence, 'more' v. 'less': 20/19829 v. 21/19783).⁴⁰

In our analysis, we found that prolonged statin use conferred a greater protective effect of statins in patients with BE. While Nguyen et al¹⁵ reported that cumulative filled statin prescription for >12 months has a reduced risk of EAC compared to those with 12 month or those with no statin prescription, Beales et al¹⁷ reported that use of statins for more than 5 years was associated with a significantly lower incidence of EAC, as compared to 6 months-5 years of use [adjusted OR (95% CI) for statin use >5 yrs v. 2-5 yrs v. 6 months-2 yrs: 0.41 (0.15-0.85) v. 0.58 (0.27-1.43) v. 0.77 (0.29-1.87)]. Kastelstein et al and Friedman et al, however, did not find a significant difference in the effect of statins when used for more than 5 years or less.^{16, 35} Using time-varying analysis in the large QResearch database, Cox et al demonstrated that the reduction in risk of EC was apparent one to three years after starting statins and persisted during the first five years of treatment, and in women the risk returned to normal within the first year after stopping treatment and in men one to three years after stopping treatment.⁴²

Differences in Observational Studies and Clinical Trials

The chemopreventive effect of statins was seen primarily in observational studies which accounted for a large majority of the included EC cases (9,162 cases, 98.7%). RCTs included in the study did not demonstrate any significant chemopreventive effect of statins, though these accounted for a small minority of the included EC cases (123 cases, 1.3%). Importantly, the clinical trials included in the meta-analysis represented post-hoc analysis of

22 RCTs performed on the effect of statins on cardiovascular mortality. By design, the patients enrolled in these RCTs were at low risk for development of EC. Also, given the small number of cases of EC, the studies were not adequately powered to detect a significant difference in the placebo and statin group with regard to development of EC. Moreover, since the occurrence of EC was not the primary objective of these trials so patients were not routinely screened for development of EC; this might have affected the detection rate of EC. The follow-up duration in these RCTs was short. These factors may explain why current clinical trials of statins for prevention of cardiovascular mortality do not demonstrate a chemopreventive effect of statins against EC.

On the other hand, the chemopreventive effect of statins seen in observational studies may also represent an over-estimate its true effect. Observational studies lack the experimental random allocation of the intervention necessary to test exposure-outcome hypotheses optimally. Despite adjusting for numerous covariates, it is not possible to eliminate the potential of residual confounding. It is possible that the observed decreased risk of EC seen in statin users may relate to a healthy user bias.⁵¹ Healthier patients taking statins may be more likely to be prescribed preventive medications and/or be more compliant with preventive health measures, as compared to patients not taking statins. The latter poorly compliant patients may have other unhealthy lifestyle practices predisposing them to EC. However, given the strength and consistency of association especially in patients with BE, along with experimental and *in vitro* data that suggests biological plausibility of the beneficial effects of statins in preventing EAC, suggests that this is unlikely to be primary driver of results. Observational studies are inherently not able to definitely establish when the intervention may exert its influence, the minimum effective dose/duration and frequency of medication intake is required to achieve a benefit. Although we found a potential duration benefit, differing exposure definitions between studies make recommendations regarding dosing regimens problematic.

Limitations

The included studies did not provide data based on histological subset of EC. It is well known that EAC and ESCC have differing risk factors, and different pathogenesis. Data from current studies are limited to infer a protective association between statin use and development of ESCC. In a study from UK using the General Practice Research Database, Bhutta et al showed an inverse association between statin use and development of ESCC (OR, 0.41; 95% CI, 0.21–0.80).⁴³ Additionally, all studies did not adjust for the same confounders. They generally failed to account for one or more of the following risk factors for EC: presence of BE including length of segment and presence of baseline dysplasia, concomitant use of NSAIDs/aspirin or PPIs, obesity and smoking status. However, using the adjusted data for any use of statins had little effect on the summary OR, suggesting that any difference attributable to using different confounders for adjustment is likely small. Another potential limitation that particularly applies to case-control studies evaluating EC is recall bias. However, in most studies pharmacy drug prescriptions information was used, and hence, the effects of this are likely minimal.

Implications for Practice

In patients with multiple risk factors for EAC (e.g., obese men with known long-segment, non-dysplastic BE or BE with low-grade dysplasia), statins may have a chemo-protective effect on dysplasia progression and development of EAC. Additionally, given significant potential benefit of statins in reducing risk of development of EAC in patients with BE with HGD, future prospective studies may assess whether statin use in these patients after therapeutic endoscopic intervention may prevent recurrence.

Conclusions

In summary, meta-analysis of all studies suggest that statin use is associated with a reduced risk of EC, with greater and most consistent benefit on the risk of EAC in patients with known BE. Longer duration of statin use, in combination with aspirin/NSAIDs may provide greater protective effect. Given the high mortality rates after a diagnosis of EC, these results support chemoprevention trials evaluating statins in populations at high risk of developing EAC.

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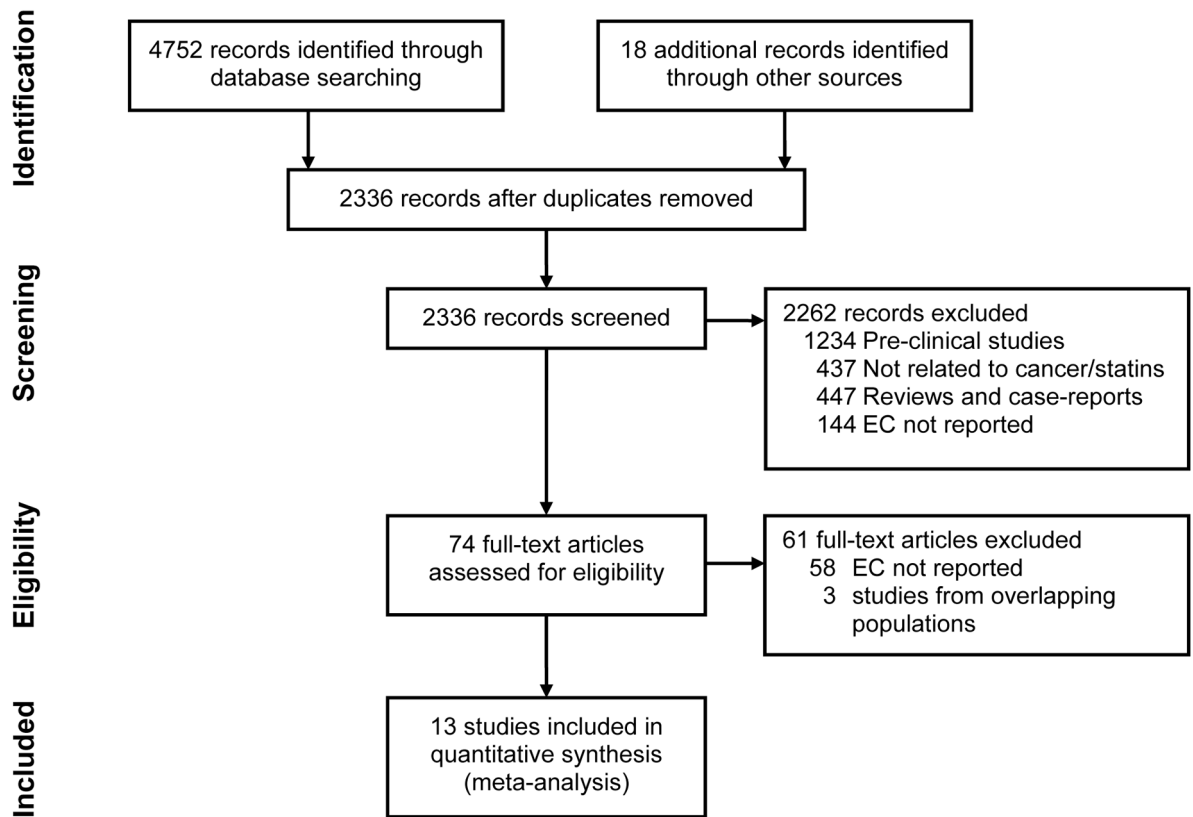
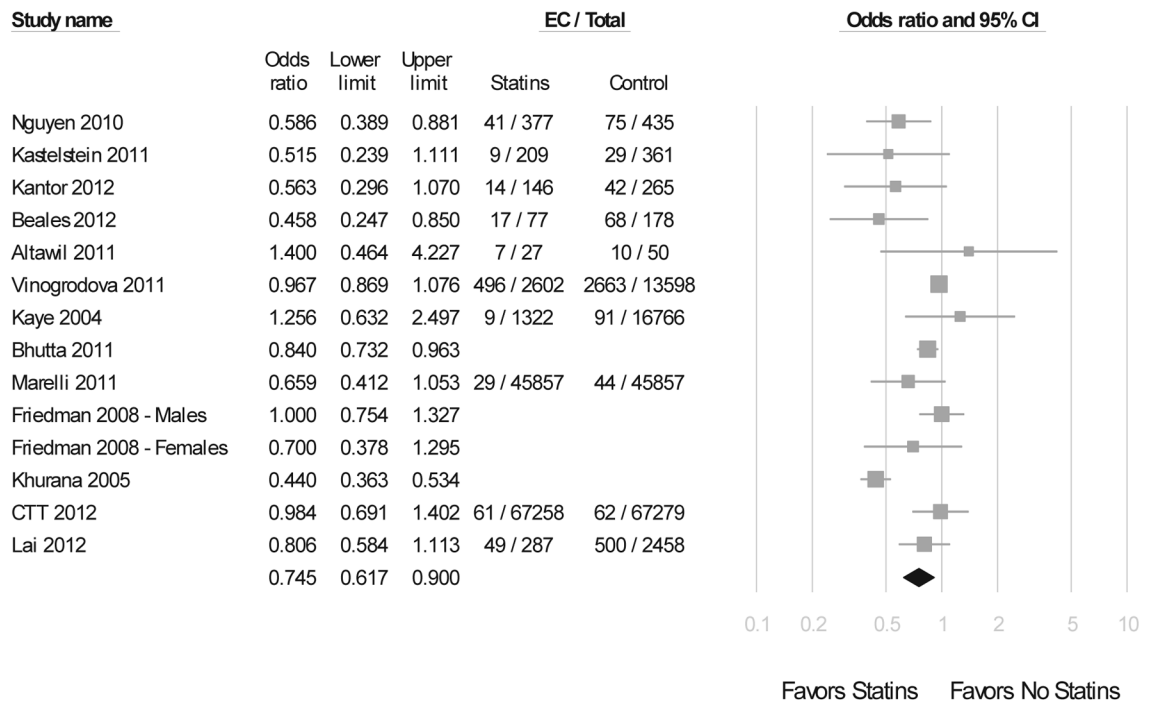


Figure 1.
Flowsheet summarizing study identification and selection.

Statins and Risk of Esophageal Cancer in All Patients - Unadjusted OR



Statins and Risk of Esophageal Cancer in All Patients - Adjusted OR

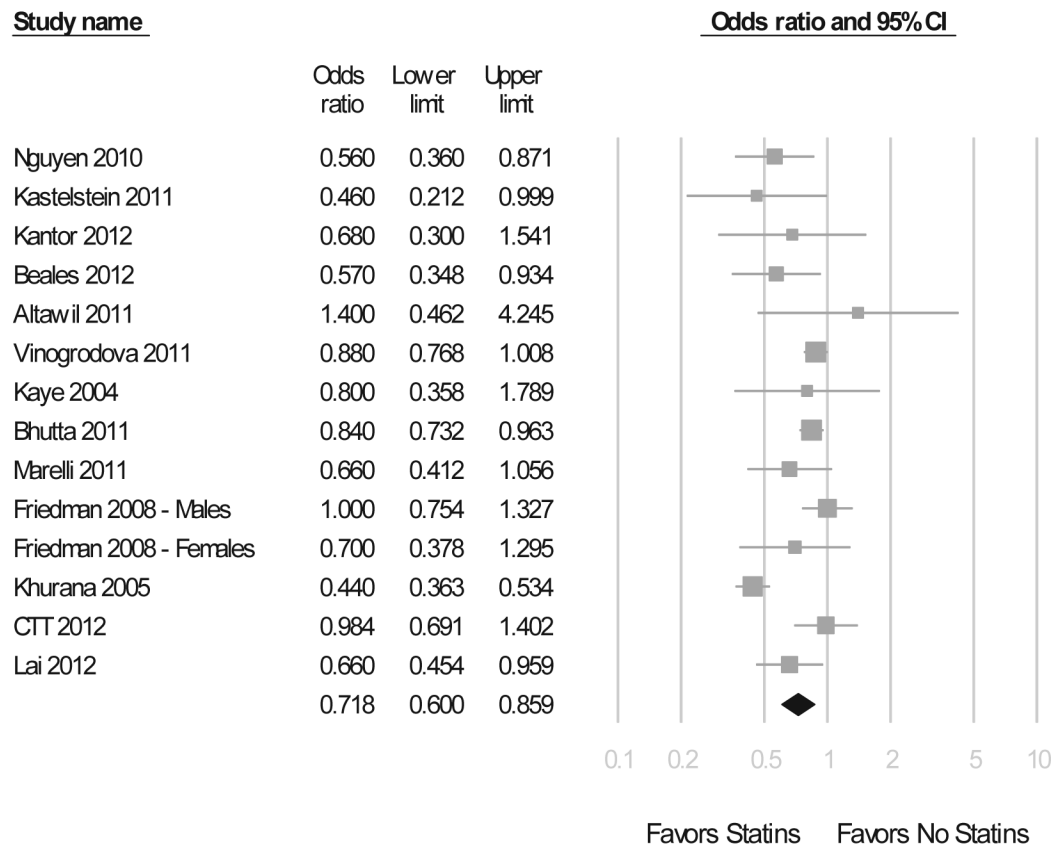
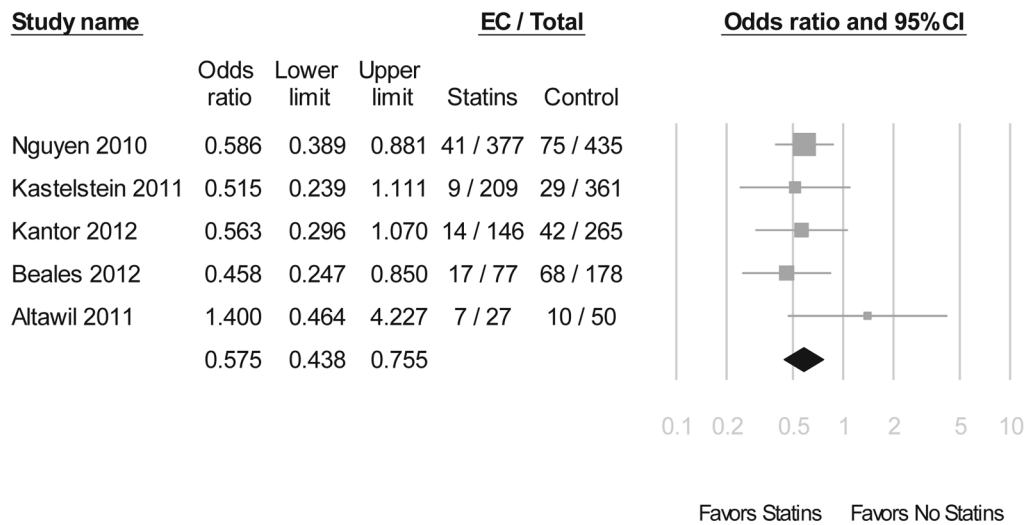


Figure 2.

Figure 2A. Summary of unadjusted odds ratios assessing the risk of esophageal cancer with statin exposure in all included studies.

Figure 2B. Summary of adjusted odds ratios assessing the risk of esophageal cancer with statin exposure in all included studies.

Statins and Risk of Esophageal Adenocarcinoma in Patients with BE - Unadjusted OR



Statins and Risk of Esophageal Adenocarcinoma in Patients with BE - Adjusted OR

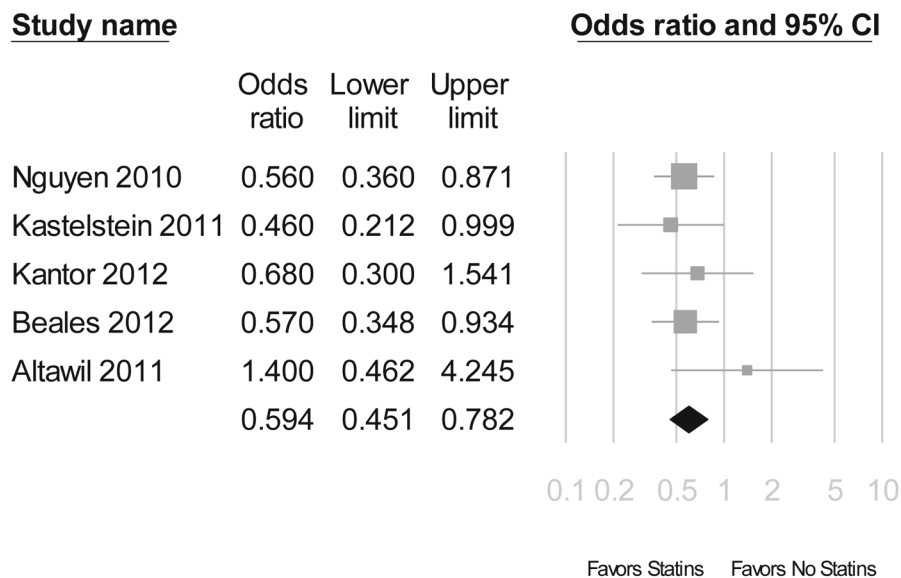


Figure 3.

Figure 3A. Summary of unadjusted odds ratios assessing the risk of esophageal adenocarcinoma and/or high grade dysplasia with statin exposure in patients with Barrett's esophagus.

Figure 3B. Summary of adjusted odds ratios assessing the risk of esophageal adenocarcinoma and/or high grade dysplasia with statin exposure in patients with Barrett's esophagus.

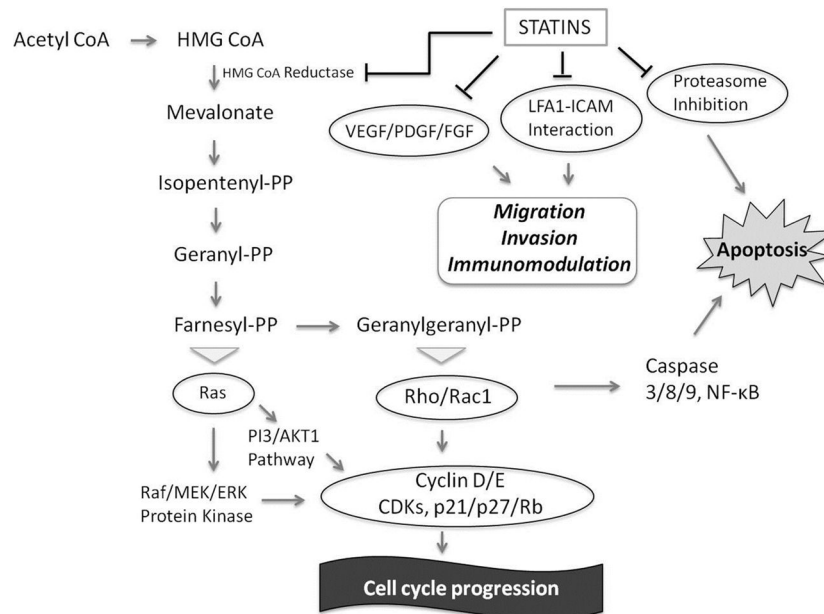


Figure 4. Mechanism of anticancer effects of statins

In vitro and animal studies have shown that statins exert antineoplastic effects through both HMG-CoA reductase-dependent and HMG-CoA reductase-independent pathways. The primary mechanism of action of statins on cholesterol reduction is by competitive inhibition of HMG-CoA reductase, the rate limiting step in cholesterol biosynthesis, blocking the conversion of HMG-CoA into mevalonate. Consequently, statins can inhibit several of the downstream products of the mevalonate pathway including the generation of isoprenoids. This prevents post-translational prenylation of small signaling G-proteins of the Ras/Rho/Rac superfamily, which are important mediators of cell growth, differentiation and survival.¹⁰ They also exert pro-apoptotic effects through regulation of Rho and RAF-mitogen activated protein kinase 1-extracellular regulated kinase (MEK-ERK) pathway through a HMG-CoA reductase dependent mechanism, by activating caspases and decreasing Bcl-2.^{52, 53} Statins inhibit the activation of the proteasome pathway, limiting the breakdown of cyclin-dependent kinase inhibitors p21 and p27, thus allowing these molecules to exert their growth-inhibitory effects.⁵⁴ In addition, they also exert an anti-inflammatory and immunomodulatory effects, modifying the cell adhesion cascade through both HMG-CoA reductase-dependent and HMG-CoA reductase-independent effects.¹⁰

Table 1

Characteristics of included studies assessing the risk of esophageal cancer with statin use

Study	Design	Location	Setting	Time period	Exposure assessment	Total subjects	EC cases	Variables adjusted for	Study quality (NOS) ⁴			Overall quality score (maximum 9)
									Selection	Comparability	Outcome/Exposure	
<i>Studies on patients with BE</i>												
Nguyen et al ⁴⁴	C-C	USA	Population-based	2000–2004	Pharmacy	812	116	1–3,6,10	***	**	***	8
Kastelein et al ¹⁶	Cohort	Netherlands	Hospital-based	2003–2010	Patient interview	570	38 [‡]	1,2,6,8,9	****	**	***	9
Kantor et al ¹⁸	Cohort	USA	Hospital-based	1983–2009	Patient interview	411	56	1,2,5,6	***	**	***	8
Beales et al ¹⁷	C-C	UK	Hospital-based	NR	Patient interview	255	85	1,2,4–6,11,12	****	**	***	9
Altawil et al [abstr] ³⁹	Cohort	USA	Hospital-based	2004–2010	EMR	77	17 [¶]	NR	***	**	***	8
<i>Studies on all patients with esophageal cancer</i>												
Kaye et al ³⁶	C-C	UK	Population-based	1990–2002	Pharmacy	18,088	100	NR	***	*	**	6
Vinogradova et al ¹⁹	C-C	UK	Population-based	1998–2008	Pharmacy	16,200	3,159	1,2,4,5,6,7,10	***	**	***	8
Bhutta et al [abstr] ³⁴	C-C	UK	Population-based	2000–2008	Pharmacy	21,475	4,242	1,2,4–7,11	**	**	**	6
Marelli et al ³⁸	Cohort	USA	Population-based	1990–2009	EMR	91,714	73	1–3,5,6	****	**	***	9
Friedman et al ³⁵	Cohort	USA	Population-based	1994–2003	Pharmacy	361,859	68	NR	****	-	***	7
Khurana et al [abstr] ³⁷	C-C	USA	Population-based	1998–2004	NR	484,226	659	1,5,7,12	*	*	**	4
Lai et al	C-C	Asia	Population-based	2000–2009	NR	2,745	549	1,2,6,12	***	*	*	5
CTT ⁴⁰	Post-hoc analysis of RCTs	Europe, Australia North America	Hospital-based	2012 [‡]	Variable	134,537	123	1,2,4,5,10	N/A	N/A	N/A	N/A

[‡] -EAC or HGD (in patients with Barrett's),

[¶] - all dysplasia or EAC,

[#] - EAC + squamous cell Ca,

[‡] - year of publication;

⁴ - Study quality assessment of observational studies performed using the Newcastle-Ottawa scale (each asterisk represents if individual criterion within the subsection were fulfilled) [C-C = Case-control, DM = Diabetes mellitus, NSAIDs = Non-steroidal anti-inflammatory drugs, NR = Not reported, NOS = Newcastle Ottawa Score, RCT = Randomized Controlled Trials]; 1=age, 2=sex, 3=face, 4= BMI, 5=smoking, 6=NSAIDs/aspirin, 7=DM, 8=BE length, 9=BE histology, 10=other comorbidities, 11=other medications, 12=alcohol use

Table 2

Subgroup analysis of all studies

Subgroups	No. of studies	Unadjusted OR	95% CI	Adjusted OR	95% CI	Heterogeneity between groups (p)*
EC type	EAC + ESCC	0.80	0.65–1.00	0.75	0.61–0.93	0.18
	EAC (patients with BE)	0.57	0.44–0.75	0.59	0.45–0.78	
Study design	Case-control	0.71	0.54–0.94	0.66	0.51–0.85	0.19
	Cohort	0.76	0.60–1.00	0.79	0.62–1.02	
	RCT (post-hoc analysis)	0.98	0.69–1.40	NR	NR	
Publication type	Full text	0.80	0.68–0.94	0.78	0.68–0.89	0.73
	Abstracts	NR	NR	0.70	0.39–1.24	
Study location	USA	0.67	0.48–0.92	0.68	0.48–0.95	0.30
	Europe	0.85	0.70–1.03	0.82	0.72–0.93	
Study setting	Population-based	0.76	0.61–0.95	0.71	0.57–0.88	0.82
	Hospital-based	0.70	0.47–1.03	0.74	0.53–1.05	
Method of ascertainment to exposure	Pharmacy	0.88	0.77–1.01	0.85	0.77–0.94	0.10
	EMR	0.81	0.42–1.56	0.81	0.42–1.55	
	Interview	0.51	0.35–0.75	0.56	0.39–0.82	

* - for adjusted OR; EAC – Esophageal adenocarcinoma, ESCC – Esophageal squamous cell carcinoma, N/A – Not applicable, NR – Not reported