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Autism Risk Across Generations: A Population Based Study of Advancing Grandpaternal and Paternal Age

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Abstract

Context—Advancing paternal age has been linked to autism.

Objective—To further expand knowledge about the relation between paternal age and autism by studying the effect of grandfathers' age on childhood autism.

Design—Population-based multigenerational case-control study.

Setting—Nationwide Multi-Generation and Patient registers in Sweden.

Participants—We conducted a study of individuals born in Sweden since 1932. Parental age at birth was obtained for over 90% of the cohort. Grandparental age at the time of birth of the parent was obtained for a smaller subset (5,936 cases and 30,923 controls).

Main Outcome Measures—International Classification of Diseases (ICD) diagnosis of childhood autism in the Patient Registry.

Results—There was a statistically significant monotonic association between advancing grandpaternal age at the time of birth of the parent and risk of autism in grandchildren. Men who had a daughter when they were 50 or older were 1.79 times (95% confidence interval: 1.35-2.37, $p < 0.001$) more likely to have a grandchild with autism, and men who had a son when they were 50 or older were 1.67 times (95% confidence interval: 1.35-2.37, $p < 0.001$) more likely to have a grandchild with autism, compared to men who had children when they were 20-24, after controlling for birth year, sex, age of the spouse, family history of psychiatric disorders, highest family education and residential county. There was also a statistically significant monotonic association between advancing paternal age and risk of autism in the offspring. Sensitivity

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analyses indicated that these findings were not the result of bias due to missing data on grandparental age.

Conclusion—Advanced grandparental age was associated with increased risk of autism, suggesting that risk for autism could develop over generations. The results are consistent with mutations and/or epigenetic alterations associated with advancing paternal age.

Introduction

Autism is a neurodevelopmental disorder characterized by social deficiencies, language impairments and repetitive behavior patterns.¹ The disorder onsets early in life, has a high heritability, and is associated with a marked reduction in birth rates.²

During the last decade, evidence has accumulated suggesting that the offspring of older fathers have an increased risk of developing autism.³⁻⁶ A recent meta-analysis found that fathers aged 50 years and older were 2.2 times more likely to have a child diagnosed with autism compared to fathers younger than 30 years.⁵ Advanced paternal age has also been associated with other mental disorders such as schizophrenia,⁷⁻⁹ bipolar disorder,¹⁰ and general neurocognitive development in children.¹¹ The mechanism behind the paternal age effect on adverse neuropsychiatric outcomes is unknown. It has been suggested that *de novo* mutations occurring in the male germ cell line underlie the association.¹² In men, spermatogonial cells replicate every sixteenth day, resulting in approximately 200 divisions by the age of 20 years and 660 divisions by the age of 40.¹³ Each time the cell divides, the replication of the genome introduces the possibility of copy error mutations. In humans, it has been confirmed that sperm from older men have significantly more mutations.^{12, 14, 15} Levels of DNA proof-reading and repair enzymes decline as a function of advancing paternal age^{16, 17} and DNA fragmentation increases,¹⁸ further compromising the integrity of gene replication. Experiments based on mouse model related to advanced paternal age have confirmed that the offspring of older sires have a significantly increased risk of *de novo* copy number variants and several of these mutations involved genes previously linked to autism.¹⁹

Recently, several studies have reported that *de novo* mutations in autism pedigrees are predominantly paternal in origin and are significantly associated with advancing paternal age.²⁰⁻²³ An Icelandic study on individuals with sporadic schizophrenia or autism even showed that the rate of new mutations in relation to paternal age is two new mutations per year.²⁴ In addition, commentators have noted that the genetic architecture of neurodevelopmental disorders such as autism and schizophrenia is characterized by locus heterogeneity, variable expressivity of the same mutations and a cumulative impact on common biological pathways.^{25, 26} Thus, it is feasible the some paternal age-related *de novo* mutations may not result in adverse health outcomes in the offspring, but still contribute the overall burden of mutations inherited by subsequent generations. Thus, it would be predicted that both paternal and grandpaternal age could contribute to a cumulative ‘threshold’ of mutation that emerges in an increased risk of neurodevelopmental disorders such as schizophrenia and autism. By using the unique Swedish national registers we can test if the older the grandfather is when the parent is born, the greater the risk for autism in the grandchild and thus, further explore the paternal age effect.

Method

Data source

By linking population-based Swedish longitudinal registers we compared the ages of parents and grandparents at offspring birth among individuals with or without childhood autism

diagnosis. The unique personal identification number assigned to each Swedish citizen at birth or upon arrival to the country (immigrants) enables linkage of national registers. The Patient Register includes practically all psychiatric inpatient discharge diagnoses in Sweden since 1973 recorded according to the International Classification of Diseases (ICD).²⁷⁻²⁹ The Patient Register also includes outpatient care in Sweden since 2001. The Swedish Multi-Generation Register contains information about biological parents of an index person and their birth dates.³⁰ A prerequisite for being included in the register is that the index person was born after January 1st 1932, and ever registered as living in Sweden after 1960. Ethical approval was given by the research ethics committee at Karolinska Institutet, Stockholm, Sweden.

Analytic cohort

We identified individuals diagnosed with childhood autism in the patient register (ICD-9 code 299.0 and ICD-10 code F84.0). We included diagnoses given at discharge from inpatient care since 1987 when the specific diagnostic code for childhood autism was first introduced and diagnoses given during outpatient care since 2001. Medical records are computerized and contain notations from psychiatrists, psychologists, neurologists, social workers, and nurses for inpatient and outpatient treatment. High validity of ICD diagnoses recorded in the Swedish patient register has been found by comparing diagnostic register code with medical records. The positive predictive value for most somatic and psychiatric diagnoses is about 85-95%³¹ and a medical record review substantiated the presence of DSM-IV autism in 94.3% of cases (83 of 88). The patient register was followed until December 31st 2009. Individuals that did not meet our criteria for autism were considered unaffected. Five unaffected individuals for each affected subject were frequency matched on gender and exact year of birth. Age data for parents and grandparents were linked to the study subjects. Ages of parents were defined as the parent's age at the time of the index person's birth. Ages of grandparents were defined as the grandparent's age at the time of the parent's birth. Birth dates were obtained from the Multi-Generation Register. We identified 9,868 individuals affected with childhood autism and 49,340 unaffected individuals. After linking ages of parents and grandparents, the final study sample consisted of 5,936 (60% of the initial sample) affected individuals and 30,923 (63 % of the initial sample) unaffected individuals with complete data on both maternal and paternal grandparents. A description of the study sample is outlined in Figure 1.

Covariates

A family history of psychiatric diagnosis was defined as having a parent or a grandparent with a diagnosis of schizophrenia, bipolar disorder or autism in the patient register, defined by ICD-8 and ICD 9 codes 295-299 (except 296,2 and 296B) and ICD-10 codes F21-29, F30-31 and F84. The selected diagnoses are possible confounders since they have been associated with paternal age in earlier studies. For the analyses without grandparental information, only information of parental history of psychiatric diagnosis were used. As a proxy measure of the socio-economic home environment of the grandchild, we examined parental education defined as highest achieved education level within each parental pair. Information about education level was retrieved from the longitudinal integration database for health insurance and labour market studies. Since coverage of outpatient care might vary across counties and place of residence might be a potential confounder, we also collected information about the probands' residential county from the Total Population Register. The distribution of covariates in the final data set is described in eTable 1. This table also shows the distribution of birth year divided into 10-year categories as well as the gender distribution where boys represented 72 % of our sample and 28 % of our study subjects were girls.

Statistical analysis

We estimated the relative risk of autism in offspring comparing different categories of parental and grandparental age by calculating the odds ratio (OR) and associated two-sided 95% confidence intervals (CI) using logistic regression. Ages were categorized into 5-year intervals with 20-24 years as the reference category. The analyses were performed in four steps. First, we adjusted only for birth year and sex (Model 1), followed by an analysis also adjusted for the age of each parent's or grandparent's partner/spouse (Model 2). We did not control for age of the maternal grandparents when analyzing paternal grandparents and vice versa, since we do not consider these ages to be directly correlated in the same manner as age of spouses. In Model 3, we added family history of psychiatric disorders and lastly, we included parental education as well as residential county (Model 4). Logistic regression analyses were performed in SAS (version 9.2; SAS Institute Inc, Cary, North Carolina), using Proc logistic. Statistical hypothesis testing was based on the 2-sided 5% level of significance. The models were evaluated for goodness-of-fit by visual inspections of the model residuals. We calculated variance inflation factors to check for collinearity between parental and grandparental age covariates and found no signs of such problems. Using a paternal age cutoff at 40 years we calculated the attributable risk (AR) by $P(EID) \cdot (RR-1) / RR$. In other words, we obtain estimates on the proportion of autistic children that could be avoided if fathers and grandfathers at ages 40 or older had their child before 40 years, assuming a causal effect.

Sensitivity analyses

Diagnoses given during outpatient care have not been included in the Swedish Patient register until 2001. We therefore did sensitivity analyses including only inpatient data to examine potential differences between patients treated in inpatient care and those treated in outpatient care. These analyses included 1,845 cases (Model 1-3) and 1,843 cases (Model 4), respectively, with autism diagnoses assigned only during inpatient care.

We performed additional analyses on grandparental ages adjusting for ages of the parents to explore if this affected the results. We also wanted to investigate effects of potential truncation of parental ages after linkage of grandparental age data and did this by analyzing parental ages before the linkage and comparing the results with data from the main analysis (sample described in Figure 1). We could identify 9,221 affected and 44,232 unaffected individuals with parental age data, corresponding to 93% and 90% of the original samples. To address the issue of potential bias due to the different probabilities of being selected for the different sets of analysis, with and without requirements of valid grandparental data, we applied inverse probability weighting to the regression models with robust standard errors.

Results

Grandparental age and autism

The logistic regression analyses on grandpaternal ages are presented in Table 1. Analyses adjusted for birth year and sex showed a statistically significant association between older grandfathers and autism on both the maternal and the paternal sides. The risk of autism increased monotonically with advancing grandpaternal age. When adjusting for age of the spouse the effect was statistically significant across all age categories including maternal grandfathers 30 years and or older paternal grandfathers 25 years or older. The results remained after controlling for family history of psychiatric disorders and parental education. The highest risk was found in the oldest age categories in all 4 models. In the fully adjusted model, maternal grandfathers 50 years or older had an OR of 1.79 (CI= 1.35-2.37) compared to grandfathers aged 20-24 years, and paternal grandfathers 50 years or older had an OR of

1.67 (CI= 1.25-2.24). We found no trend indicating an increased risk for autism in grandchildren of women who gave birth at older ages (eTable 2).

Parental age and autism

Logistic regression analyses on parental ages (Table 2) suggested a statistically significantly risk increase of childhood autism in all paternal age categories 35 years or older in Model 1, again with a monotonic increase in risk with increasing age. After adjustments for age of the mother, the paternal age effect was found for fathers 30 years or older. Further adjustments did not have any evident effect on the point estimates or the confidence intervals. In the fully adjusted model (Model 4), the highest risk was found in men 50 years or older (OR= 2.26; 95% CI= 1.61-3.18). We found no evident trend for an increased risk of autism in offspring of older mothers after adjusting for father's age and the other covariates, with the exception of an excess of mothers aged 40 years and older in cases. The AR for grandfathers above the age 40 was estimated to between 3 % for both maternal and paternal grandfathers. For fathers, the corresponding proportion was 6 %."

Sensitivity analyses

The analyses conducted separately on inpatients showed similar age effects as in the main analysis (eTable 3). Although the confidence intervals were expectedly broader than in the main analyses, we identified a trend of increasing autism risk in grandchildren of older maternal and paternal grandfathers. Again, no such effects of advanced age were found for grandmothers. Similarly, the association between paternal age and autism was also evident in these analyses. Again, we could see an association between mothers aged 40 years or more, but we could not detect any overall trend between maternal age and autism.

Adjustments for parental ages did not have any major effect on the risk estimates in comparison to the main analyses (eTable 4); the statistically significantly increased risk for autism in the grandchildren of older grandfathers remained.

The analyses on all individuals with parental age data (eTable 5) showed similar associations between parental ages and autism as compared to the sample that required present grandparental age (Table 2). In addition, estimated odds ratios, associated confidence intervals and P-values were close to identical to our main results when using the inverse probability weighting procedure.

Discussion

We can for the first time report that grandfather's age is associated with risk of childhood autism, independent of paternal or maternal age. We also confirm a statistically significant association between advanced paternal age and an increased offspring risk of autism. Associations between paternal age and autism has been reported in previous studies,³⁻⁶ including one using a 10-year Swedish birth cohort.⁵ We could also see some evidence of an association between maternal age and autism, in congruent with a novel meta-analysis.³²

A recent study reported an association between grandpaternal age and schizophrenia.³³ This association was exclusive for maternal grandfathers. There are, however, no reports of an association between paternal age and autism being transmitted to further generations. The only study that has, to our knowledge, looked at grandparents' ages in association to autism found no evidence of an effect of grandfathers' age but instead an association between age of maternal grandmothers and ASD. This study did, however, included 86 individuals with ASD as compared to the 5,936 individuals with the specific diagnosis childhood autism included in the present study.³⁴

Since autism is characterized by lower birth rates and high heritability,² it is puzzling that the disorder still exists and may even be increasing in prevalence.^{35, 36} The strong negative selection pressure should remove genes associated with this disorder from the gene pool promptly. One possible explanation to this paradox is that genetic variants increasing the risk for autism constantly arise as *de novo* mutations.³⁷ This hypothesis is supported by recent findings of *de novo* mutations in autism pedigrees being predominantly paternal in origin and significantly associated with advancing paternal age.²⁰⁻²⁴ Mendelian inheritance laws indicate that offspring who acquire a *de novo* autosomal mutation from their father's sperm should pass (on average) this mutation to half of their offspring. Age-related mutations in the male germ line could accumulate over several generations, and only influence the offspring's health after a certain "mutational threshold" has been breached^{12, 38}. Thus, paternal age and grandpaternal (both maternal and paternal grandfathers) may contribute to an offspring's mutational load resulting in an increased risk of disorders in the offspring. If this is true, autism should not only be associated with paternal age but also with grandpaternal age. In this study, we show for the first time that paternal age not only has an impact on autism risk in offspring but also affects subsequent generations.

Although the most favored hypothesis behind the paternal age effect suggests that the associations is caused by an increased rate of mutations in sperm of older men it has also been suggested that the paternal age effect is explained by men with mental or personality disorders being more likely to become fathers at older ages.^{39, 40} Although the coverage of family history of mental disorders might be incomplete and questionable in quality, the paternal age effect has been consistent in studies controlling for mental disorders in the parents. Also, within-family analysis of discordant siblings showed that the siblings affected with autism had older paternal age,⁵ supporting the notion of a causal association between paternal age and autism in the offspring. Petersen *et al.* showed that paternal age effect on schizophrenia was not present in later born children when father's age at the first born child was accounted for, opposing the hypothesis of *de novo* mutations.⁴⁰ By contrast, a Swedish study testing if the paternal age effect in autism only pertains to first-born children showed that the risk of autism increased in later born children irrespectively of when the father had his first child.⁵

Autism risk increased with advancing paternal and grandpaternal age. The risk estimates were very similar for maternal and paternal grandfathers. Although confidence intervals overlapped, the grandpaternal age risk estimates were lower than the risk estimates for paternal age which is consistent with the hypothesis of genetically mediated causes since the strength of a genetically mediated effect should reflect the genetic relatedness. On average, you inherit 50% of your genes from your parent and only 25 % from your grandparent so the overall mutational load would thus be higher if your father is old as compared to if your grandfather is old. Similarly, if the *de novo* mutation hypothesis is true the risk for autism would be greater in offspring of old fathers than in grandchildren of older men.

The coverage of data on parental age for cases and controls was very high (~90%). After linking grandparental ages approximately 60% of the initial sample remained. To collate a sample consisting of three generations, parents had to be born after 1932. Older parents are therefore more likely to be excluded than younger parents. Grandparental ages are not truncated but since parental and grandparental ages are correlated this should only result in an underestimation of the grandpaternal age effect as a result of parental age truncation. However the analyses conducted on the sample without requiring grandparental age data linkage suggested a similar paternal age effect as compared to the sample used for the main analyses. We therefore conclude that there was no major truncation of parental ages, which adds to the validity of our findings.

Although the proportion of girls with autism is relatively high in Sweden an extensive review including 29 studies showed the male/female sex ratio varied from 1.33 to 16.0.⁴¹ Since the sex ratio in the present study is well within this interval we do not consider the sex ratio to be atypical in the present study.

The main strength of this study is the inclusion of a large number of individuals diagnosed with childhood autism and available birth date information across three generations. The consistency in results between the main analyses and the analyses restricted to inpatients further strengthen our findings and the generalizability of the results.

Our findings have added salience in light of the recent evidence that autism is associated with *de novo* and inherited mutations.^{20-24, 42-44} Considering the association between advanced paternal age and *de novo* copy number variants in an animal model,⁴⁵ we speculate that paternal age-related mutagenesis is associated with an increased risk of autism via two mechanisms. The offspring of older fathers may be at increased risk of acquiring *de novo* mutations, as previously speculated.⁴⁶ Considering our finding linking grandpaternal age and risk of schizophrenia, we propose that a proportion of age-related *de novo* mutations are phenotypically silent in the offspring, but can still influence risk of autism in subsequent generations, perhaps via the interaction with other susceptibility factors. This indirect mechanism is consistent with the evidence that some mutations associated with neurodevelopmental disorders can occur in apparently healthy individual.^{47, 48}

Age of parenthood is increasing in many societies and thus it is feasible that the incidence of paternal age-related disorders will increase over time. Our findings provide new information about the paternal age effect and its impact on future generations. Older men should not be discouraged to have children based on these findings but the results may be important in understanding the mechanism behind childhood autism and other psychiatric and neurodevelopmental disorders.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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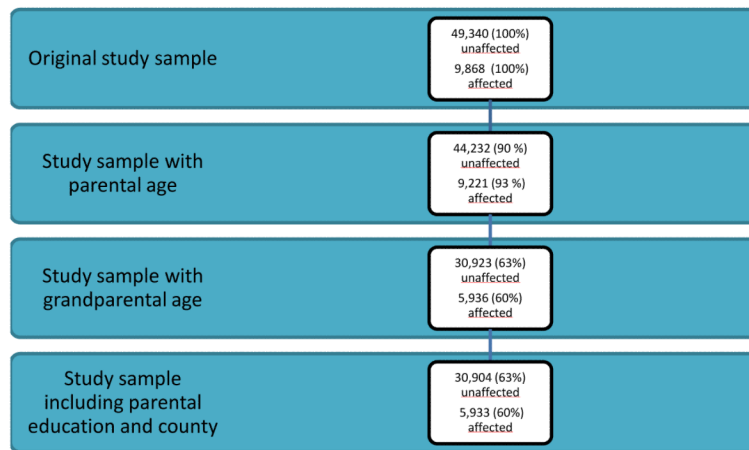


Figure 1.
Study sample

Table 1

Results from logistic regression analyses on grandpaternal ages and autism risk

	No of controls (%)		Model 1-3		Model 1		Model 2		Model 3		Model 4	
	No of controls (%)	Model 1-3	No of cases (%)	Model 1-3	OR	CI	OR	CI	OR	CI	OR	CI
Maternal grandfather age, years												
<20	675 (2.18)	122 (2.06)	0.96	0.79-1.18	0.91	0.74-1.12	0.91	0.73-1.12	0.90	0.73-1.11		
20-24	6721 (21.73)	1253 (21.11)	1.00		1.00							
25-29	9801 (31.69)	1787 (30.10)	0.98	0.91-1.06	1.07	0.98-1.17	1.07	0.98-1.17	1.08	0.99-1.18		
30-34	7082 (22.90)	1334 (22.47)	1.01	0.93-1.10	1.18	1.06-1.31	1.18	1.06-1.31	1.19	1.07-1.32		
35-39	3868 (12.51)	808 (13.61)	1.10	1.00-1.22	1.33	1.17-1.50	1.32	1.16-1.50	1.31	1.15-1.49		
40-44	1843 (5.96)	393 (6.62)	1.12	0.99-1.27	1.32	1.13-1.55	1.31	1.11-1.53	1.32	1.12-1.54		
45-49	666 (2.15)	154 (2.59)	1.22	1.01-1.46	1.39	1.12-1.73	1.37	1.10-1.71	1.34	1.07-1.67		
50	267 (0.86)	85 (1.43)	1.67	1.30-2.15	1.90	1.44-2.51	1.87	1.42-2.48	1.79	1.35-2.37		
Paternal grandfather age, years												
<20	702 (2.27)	123 (2.07)	0.96	0.79-1.18	0.88	0.72-1.09	0.88	0.72-1.09	0.91	0.73-1.12		
20-24	6293 (20.35)	1139 (19.19)	1.00		1.00							
25-29	9694 (31.35)	1793 (30.21)	1.02	0.94-1.11	1.09	1.00-1.19	1.09	0.99-1.19	1.10	1.00-1.20		
30-34	7046 (22.79)	1387 (23.37)	1.08	0.99-1.17	1.18	1.06-1.31	1.17	1.05-1.30	1.17	1.05-1.30		
35-39	4277 (13.83)	831 (14.00)	1.05	0.95- 1.16	1.17	1.04-1.33	1.17	1.03-1.32	1.15	1.02-1.31		
40-44	1971 (6.37)	405 (6.82)	1.11	0.98-1.26	1.26	1.08-1.47	1.24	1.06-1.46	1.23	1.05-1.44		
45-49	672 (2.17)	180 (3.03)	1.45	1.22-1.74	1.64	1.34-2.02	1.62	1.32-1.99	1.60	1.30-1.97		
50	268 (0.87)	78 (1.31)	1.56	1.20-2.02	1.76	1.32-2.35	1.72	1.29-2.30	1.67	1.25-2.24		

OR:Odds ratios

CI: 95% Confidence Intervals

Model 1: Adj for birth year, sex

Model 2: Adj for birth year, sex, age of spouse

Model 3: Adj for birth year, sex, age of spouse, family history

Model 4: Adj for birth year, sex, family history, highest education, county

Model 1-3: controls=30923 cases=5936

Model 4: controls=30904 cases=5933

Table 2

Results from logistic regression analyses on parental ages and autism risk (sample with grandparental ages)

	No of controls (%) Model 1-3		No of cases (%) Model 1-3		Model 1		Model 2		Model 3		Model 4	
			OR	CI	OR	CI	OR	CI	OR	CI	OR	CI
Paternal age, years												
<20	244 (0.79)	49 (0.83)	1.11	0.81-1.53	0.98	0.70-1.37	0.97	0.69-1.36	1.00	0.71-1.40		
20-24	3452 (11.16)	650 (10.95)	1.00		1.00		1.00					
25-29	9685 (31.32)	1716 (28.91)	0.96	0.87-1.06	1.04	0.94-1.16	1.05	0.94-1.17	1.06	0.95-1.19		
30-34	9989 (32.30)	1836 (30.93)	1.02	0.92-1.12	1.14	1.02-1.29	1.15	1.02-1.30	1.18	1.04-1.33		
35-39	5236 (16.93)	1051 (17.71)	1.13	1.01-1.26	1.22	1.07-1.40	1.23	1.08-1.41	1.24	1.08-1.42		
40-44	1711 (5.53)	429 (7.23)	1.41	1.23-1.62	1.47	1.24-1.73	1.47	1.24-1.73	1.45	1.23-1.71		
45-49	460 (1.49)	149 (2.51)	1.82	1.49-2.24	1.87	1.49-2.34	1.85	1.47-2.31	1.83	1.46-2.30		
50	146 (0.47)	56 (0.94)	2.23	1.61-3.07	2.25	1.61-3.15	2.23	1.59-3.12	2.26	1.61-3.18		
Maternal age, years												
<20	826 (2.67)	191 (3.22)	1.19	1.01-1.42	1.23	1.02-1.48	1.22	1.02-1.47	1.15	0.96-1.39		
20-24	6255 (20.23)	1216 (20.49)	1.00		1.00		1.00					
25-29	11256 (36.40)	2009 (33.84)	0.93	0.86-1.01	0.89	0.81-0.97	0.89	0.82-0.97	0.93	0.85-1.01		
30-34	8726 (28.22)	1586 (26.72)	0.98	0.90-1.06	0.86	0.78-0.95	0.87	0.78-0.96	0.92	0.83-1.02		
35-39	3305 (10.69)	772 (13.01)	1.27	1.15-1.41	1.02	0.90-1.16	1.02	0.90-1.16	1.11	0.97-1.26		
40	555 (1.79)	162 (2.73)	1.59	1.32-1.92	1.13	0.92-1.40	1.13	0.92-1.40	1.26	1.02-1.56		

OR:Odds ratios

CI: 95% Confidence Intervals

Model 1: Adj for birth year, sex

Model 2: Adj for birth year, sex, age of spouse

Model 3: Adj for birth year, sex, age of spouse, family history

Model 4: Adj for birth year, sex, family history, highest education, county

Model 1-3: controls=30923 cases=5936

Model 4: controls=30904 cases=5933