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## Days with Pain and Substance Use Disorders: Is There an Association?

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### Abstract

**OBJECTIVES**—We investigated possible associations between pain frequency and the five most common substance use disorders: alcohol abuse/dependence, cocaine abuse/dependence, methamphetamine abuse/dependence, opioid abuse/dependence, and marijuana abuse/dependence.

**METHODS**—We used data from the Rural Stimulant Study (RSS), a longitudinal (7 waves), observational study of at-risk stimulant users (cocaine and methamphetamine) in Arkansas and Kentucky (n=462). In fixed effects logistic regression models, we regressed our measures of substance use disorders on the number of days with pain in the past 30 days and depression severity.

**RESULTS**—Time periods when individuals had 1 to 15 days (OR=1.85, p<0.001) or 16+ days (OR=2.18, p<0.001) with pain in the past 30 days were more likely to have a diagnosis of alcohol abuse/dependence, compared to time periods when individuals had no days with pain. Compared to time periods when individuals had no pain days in the past 30 days, time periods when individuals had 16+ pain days were more likely to have a diagnosis of opioid abuse/dependence (OR=3.32, p=0.02). Number of days with pain was not significantly associated with other substance use disorders.

**DISCUSSION**—Pain frequency appears to be associated with an increased risk for alcohol abuse/dependence and opioid abuse/dependence in this population, and the magnitude of the association is medium to large. Further research is needed to investigate this in more representative populations and to determine causal relationships.

### Keywords

pain; alcohol abuse/dependence; cocaine abuse/dependence; methamphetamine abuse/dependence; opioid abuse/dependence; marijuana abuse/dependence; depression

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## 1. INTRODUCTION

Self-medication of distress, including physical pain, is often cited as a reason for substance abuse by patients and clinicians.<sup>1, 2</sup> This relationship between pain relief and substance abuse is most obvious for opioids, as individuals attempt to self-medicate their pain with opioid analgesics. Cannabinoids also have a pain modulating effect in animals and have been shown to relieve neuropathic pain in humans.<sup>3, 4</sup> Sativex, which contains tetrahydrocannabinol and cannabidiol, has recently been approved in Canada for the treatment of neuropathic pain. A review from the Institute of Medicine concluded “The available evidence from animal and human studies indicate cannabinoids can have a substantial analgesic effect.”<sup>3</sup>

On the other hand, the relationship between pain frequency and cocaine abuse, or pain frequency and methamphetamine abuse, or pain frequency and alcohol abuse, has received less attention, although there are theoretical reasons to expect a positive association. Alcohol may have analgesic effects,<sup>5</sup> although the evidence is not completely consistent. Cocaine is well known as a local anesthetic, is used in animal analgesia models,<sup>6</sup> and has been used as an intrathecal analgesic for over 100 years.<sup>7</sup> Amphetamine has a long record of use as a potentiator of opioid analgesia.<sup>8</sup> Further, the euphorogenic effects of opioids, alcohol, cocaine, marijuana, and methamphetamines are well known, and these substances might be used to self-medicate negative psychological sequelae of pain.

In studies from clinical settings, patients in substance use treatment have often been found to have higher rates of pain than the general population,<sup>9, 10</sup> and patients in treatment for pain have been found to have higher rates of SUDs than the general population.<sup>11</sup> One study found similar rates of substance abuse among methadone maintenance patients with and without pain but higher rates of mental health disorders among those with pain.<sup>12</sup> Analyses of survey data from the World Mental Health Surveys showed a non-linear relationship between alcohol use disorders and number of pain disorders in the U.S. sample; among individuals with 0 pain disorders, 1 pain disorder, and 2+ pain disorders rates of alcohol use disorders were 3.5%, 1.8%, and 3.9% respectively.<sup>13</sup> However, in another report from this study, among the U.S. sample with back pain, the rate of an alcohol use disorder was higher among those with back pain, compared to those with no back pain.<sup>14</sup> Similarly, those with arthritis were more likely to have an alcohol use disorder, compared to those with no arthritis.<sup>15</sup>

Generally these studies controlled for sociodemographic factors, but not mental health status, which is a potential limitation, as mental health status might be an important confounder in the relationship between pain and SUDs. Pain is associated with common mental health disorders, such as depression and anxiety,<sup>13, 16–20</sup> and common mental health disorders are associated with SUDs.<sup>21–24</sup> Thus, even if SUDs and pain were not causally associated, a statistical association between pain and SUDs might be expected, due to their common association with mental health disorders. This highlights the importance of controlling for mental health in models analyzing the association between pain and SUDs.

The focus of this paper is the association between frequency of pain and the five most common substance use disorders (SUDs) in the U.S.:<sup>21</sup> alcohol abuse/dependence, cocaine abuse/dependence, methamphetamine abuse/dependence, opioid abuse/dependence, and marijuana abuse/dependence. Thus we hypothesized a positive association between pain frequency and each of the SUDs, as individuals try to self-medicate their physical pain, in much the same way they have been posited to self-medicate psychological symptoms.<sup>25</sup> In this paper we assessed the relationship between pain frequency and: (i) alcohol use disorders (i.e., abuse and dependence), (ii) cocaine use disorders; (iii) methamphetamine use

disorders, (iv) opioid use disorders, and (v) marijuana use disorders. We hypothesized that individuals with greater pain frequency would be more likely to have SUDs, and that the magnitude of the relationship would be especially strong for opioid abuse/dependence. We used data from the Rural Stimulant Study (RSS), a longitudinal (7 waves), observational study of at-risk stimulant users (cocaine and methamphetamine) living in Arkansas and Kentucky. In addition to its longitudinal design, the RSS has several important features that make it an excellent data source for this analysis. Because RSS respondents had high rates of SUDs, the study is adequately powered to investigate the association of pain with individual SUDs. Further, the RSS has measures of pain frequency, as well as excellent measures of the individual SUDs and depression, the most common mental health disorder.

## 2. METHODS

### 2.1 Sample

Data are from a natural history study of 462 at risk stimulant users (cocaine or methamphetamine use in the past 30 days) residing in rural counties of Arkansas and Kentucky<sup>26-31</sup>. Counties were classified as rural according to the U.S. Office of Management and Budget definition of a non-metropolitan county, or a county with a population of 50,000 or fewer persons. Participant eligibility criteria included being 18 years of age or older, using crack or powder form cocaine and/or methamphetamine by any route of administration in the past 30 days, receiving no formal drug abuse treatment within the past 30 days, and having a verifiable address within one of the study counties. Along with cocaine and methamphetamine, respondents had high rates of other SUDs (Table 1).

For the six counties, the 2000 United States Census data indicates a range of county characteristics in terms of socioeconomic status. The three Arkansas counties were 49–57% African-American, compared to 0–2% in Kentucky. The Arkansas and Kentucky counties had high rates of household incomes under \$10,000 (14–24%), and correspondingly high rates of families living below the federal definition of poverty (11–29%) and a wide range in overall employment rates (43–69%, lowest in Arkansas). The study was approved by the relevant institutional review boards and received a Certificate of Confidentiality from NIDA.

Participants were recruited using Respondent-Driven Sampling (RDS), a variant of snowball sampling.<sup>31-35</sup> Such non-probabilistic sampling methods are critical for recruiting community “hidden populations” such as illegal drug users or those with HIV. Theoretically, RDS can generate a sample that is much more representative of the hidden population under study than can snowball or targeted sampling.<sup>32</sup> One advantage of RDS over other targeted or referral sampling strategies is that initial “seeds” for sampling are not required to be random samples of the target population because RDS has been shown to “converge” to stable characteristics of the population following successive recruitment waves.<sup>33, 34, 36</sup>

In all counties, preliminary ethnographic methods were used to identify seeds who met study criteria.<sup>37</sup> Ethnographic methods included “hanging out” in propitious locations such as bars and county fairs, talking to community members about their knowledge of drug use, meeting with treatment providers, and handing out study “business” cards to anyone who knew drug users who might contact the study. Study seeds who completed the baseline interview were asked to give referral coupons to people they knew used drugs. If referrals resulted in study contact, the seeds received \$10 per contact for up to three contacts but up to six referrals were allowed. Subsequent participants also followed the same procedures. Confidentiality was maintained by requiring that potential study participants initiate study contact.

Recruitment was conducted between June 2003 and September 2004. Written informed consent was obtained prior to the baseline interview.

Trained research assistants conducted baseline interviews using computer-assisted personal interview technology on a laptop computer. At each six-month follow-up interview, the majority of questions contained in the baseline interview were repeated and urinalysis was conducted to help assure the veracity of self-reported drug use.<sup>38–40</sup> Extensive tracking information was obtained at the baseline interview and throughout all follow-up interviews so that participants could be re-located, culminating in 79% follow-up participation rate at the 36-month interview. The RSS sample is shown in table 1.

## 2.2 Measures

### 2.2.1 Dependent Variables

**Substance Abuse/Dependence Disorders:** The five individual SUDS (alcohol use disorders (i.e., abuse and dependence), cocaine use disorders, methamphetamine use disorders, opioid use disorders, and marijuana use disorders) were measured using the Substance Abuse Outcomes Module (SAOM), for the previous 6 months. The SAOM has undergone extensive reliability and validity examinations and demonstrates reasonable reliability (internal reliability coefficient alpha 0.58–0.90, test-retest reliability 0.56–0.99) and validity (concurrent validity generally 0.5–0.8, predictive validity 0.5–0.9).<sup>41</sup> Concurrent validity for the SAOM was based on longer key instruments such as a structured diagnostic interview for substance use disorders, the CIDI-SAM,<sup>42</sup> and the Addiction Severity Index (ASI).<sup>43</sup> The SAOM has shown a 90–93% agreement with the CIDI-SAM on DSM-IV substance use diagnosis (present/absent).<sup>41</sup>

### 2.2.2 Independent Variables

**Pain Frequency:** The main independent variable of interest was the number of days in the past 30 days with pain. At each wave respondents were asked “How many days in the last 30 days have you had bodily pain (either recent or long-standing pain)?” To ease interpretation of the data and to allow for non-linear effects we coded this into three groups: 0 days, 1 to 15 days, and 15+ days. The RSS also contains data on the number of days that pain interfered with work activities, and number of days that pain interfered with social interactions; as might be expected these were highly correlated with the number of days with pain, and therefore not included in the final models. (The regression of number of days with pain on number of days that pain interfered with work activities pain had an  $R^2=0.41$  ( $p<0.001$ ), and the regression of number of days with pain on number of days that pain interfered with social activities had an  $R^2=0.39$  ( $p<0.001$ )). In preliminary models we also investigated the effects of pain severity (0–10 scale), but these effects were statistically not significant when controlling for number of days with pain and depression, and not included in the final models.

**Depression Severity:** Depression severity was measured with the PHQ-9.<sup>44</sup> We divide the full sample into four groups based on PHQ-9 scores: 0 to 4 (no depression), 5 to 9 (mild depression), 10 to 14 (moderate depression) and 15+ (moderately severe or severe depression).<sup>44</sup>

## 2.3 Analysis

We utilized five logistic regressions, one for each SUD as the dependent variable, and each regression included the number of days with pain, and depression severity, as independent variables. We utilized fixed effects logistic regressions, which we believe are particularly well-suited for statistical analysis of longitudinal, non-experimental data.<sup>45</sup> In these models,

if there are no missing data, the total number of observations in a regression equals the number of individuals multiplied by the number of time periods. In our case, while there was a theoretical possibility of 3234 observations, in actuality there were 2862 observations (89%). Fixed-effects models control for respondent characteristics that are stable over time, such as personality traits, thus possibly eliminating one potential source of omitted variable bias. This is done by including a separate dummy variable for each individual in the sample. As fixed effects models control for stable characteristics, variables such as gender and race are not (and cannot be) included in such models. Since the individual's level of education and income were also stable over the time period of our study, we did not include these variables in our models. The interactions between time and time-variant variables (pain frequency and depression) were also tested and were not included in the models due to non-significance.

We tested models that used both lagged predictors, (e.g., regressing six month substance use outcomes on baseline predictors, regressing 12 month substance use outcomes on six-month predictors, etc.) and models that used simultaneous predictors (e.g., regressing baseline outcomes on baseline predictors, etc.) These models produced similar results; in this paper we report the results from models using simultaneous predictors, that is regressing the baseline substance use disorder outcome on baseline predictors, regressing the 6 month substance use disorder outcome on predictors measured at 6 months, etc.

### 3. RESULTS

#### 3.1 The RSS Sample

Characteristics of the RSS sample at baseline are shown in Table 1. Thirty percent of the RSS sample had no days with pain in the past 30 days, 43.2% had 1 to 15 days, and 26.6% percent had 16+ days. Among those with 1 to 15 days with pain in the last 30 days, 59.6% reported a chronic source of pain, and among those with 16+ days, 81% reported a chronic source. While the RSS was designed to study at-risk stimulant users, the other SUDs were also common. At baseline 50.6% had alcohol abuse/dependence, 51.7% had cocaine abuse/dependence, 34.2% had methamphetamine abuse/dependence, 5.0% had opioid abuse/dependence, and 36.4% had marijuana abuse/dependence. Approximately half were from Kentucky, half from Arkansas. Mean age was 34, and reflecting gender differences in rates of substance abuse, the majority were male (58%). Almost all were either African-American (39.6%) or white (57.6%). Reflecting the disadvantaged population, only 40.7% had graduated high school and only about 3 in 10 were employed. At baseline, 26.6% percent had 16+ days with pain, while 43.2% had 1 to 15 days. Twenty-six percent had mild depression, 23.4% had moderate depression, and 15.4% had moderately severe or severe depression.

#### 3.2 Pain Frequency and SUDs

The results of our fixed effects logistic regression models are shown in Table 2, with each column representing a separate regression. Time periods when individuals had 1 to 15 days with pain (OR=1.85,  $p=0.0002$ ) or 16+ days (OR=2.18,  $p=0.0003$ ) were more likely to have a diagnosis of alcohol abuse/dependence, compared to time periods when individuals had no days with pain in the past 30 days. Compared to time periods when individuals had no pain days, time periods with 16+ pain days were more likely to have a diagnosis of opioid abuse/dependence (OR=3.32,  $p=0.02$ ); but the OR for 1 to 15 days, while large (OR=2.25) just missed being significant at the 0.05 level ( $p=0.06$ ) as a predictor of opioid abuse/dependence. Number of days with pain was not significantly associated with cocaine abuse/dependence, methamphetamine abuse/dependence, or marijuana abuse/dependence.

### 3.3 Depression and SUDS

Compared to time periods when individuals had no depression, time periods when individuals had mild, moderate and moderately severe/severe depression were significantly associated with higher rates of cocaine abuse/dependence, and marijuana abuse/dependence. Results for the other SUDs were less consistent. While generally the odds ratios were positive, the only other significant findings were that periods with moderate depression were associated with opioid abuse/dependence (OR=2.67, p=0.04) and methamphetamine abuse/dependence (OR=2.35, p=0.006).

## 4. DISCUSSION

In separate analyses we investigated the relationship between pain frequency and the five most common SUDs. This allows us to investigate the relationship between a given SUD and pain frequency, and compare the magnitude of the associations across the five SUDs, which, to our knowledge, has never been done before. We found that the number of days with pain in the past 30 days was significantly associated with alcohol abuse/dependence and opioid abuse/dependence, and the strength of the association was moderate to strong (e.g., OR's of 1.85 to 3.32). Although we utilized longitudinal data, we emphasize that these results should be viewed as measures of association, and do not necessarily imply causality. The relationship between the number of days with pain in the past 30 days and cocaine abuse/dependence, methamphetamine abuse/dependence, and marijuana abuse/dependence were not statistically significant.

The results for alcohol use disorders are potentially important. Alcohol use disorders are the most common substance use disorders, occurring in 4 to 9% of the U.S. population in a given year,<sup>21, 23, 46, 47</sup> and cause substantial morbidity,<sup>48</sup> accounting for about 5% of all disability in Western industrialized countries.<sup>49</sup> The negative social and health consequences associated with alcohol use disorders are protean<sup>50</sup> and include increased suicidal behaviors,<sup>51, 52</sup> high rates of criminal justice involvement and violence,<sup>53</sup> and substantial medical/physical consequences.<sup>50, 54</sup> The medical consequences of alcohol use disorders, such as cirrhosis and premature death, are particularly high among Hispanics, Native Americans, and African Americans compared to whites.<sup>55, 56</sup>

Because of the high prevalence and societal costs of alcohol abuse/dependence, the identification of risk factors for these disorders is important, particularly if the risk factor identified is strongly associated with the disorder, occurs commonly, and is potentially modifiable. Notably, we found that the magnitude of the association between alcohol use disorders and pain frequency was comparable to, or even larger than, the magnitude of association between alcohol abuse/dependence and depression, a well established risk factor for alcohol use disorders. Further, chronic pain occurs commonly, significantly affecting approximately 37% of the general population.<sup>57</sup> By comparison, depressive disorders affect approximately 9% of the population in a given year.<sup>21, 23</sup> Given this, it is interesting to view our results in terms of the population-attributable risk for alcohol abuse/dependence (that is, the proportion of alcohol abuse/dependence disorders that can be attributed to various factors), which is a function of both the strength and the prevalence of the risk factor. Because pain disorders are more prevalent than depression, the population-attributable risk for alcohol abuse/dependence from pain may actually be as great, or greater, than population-attributable risk for alcohol abuse/dependence from depression.

Improved detection and treatment of chronic pain disorders might decrease the onset of alcohol abuse/dependence, and careful assessment and appropriate treatment of pain in individuals with alcohol abuse/dependence might facilitate the treatment of an existing disorder. Chronic pain can be successfully treated with medications,<sup>58</sup> cognitive-behavioral



therapy, and multidisciplinary pain treatment.<sup>59</sup> Treatment of chronic pain in the presence of SUDs is complicated,<sup>60</sup> but our findings suggest that among individuals with comorbid pain and SUDs concurrent treatment of chronic pain and SUDs might be more successful than treatment of only the SUD, or only the pain condition. Unfortunately, there is limited data on evidence based programs to simultaneously treat chronic pain and SUDs remain to be developed.<sup>61–64</sup>

Besides being a possible risk factor for alcohol abuse/dependence, the link between pain and alcohol abuse is important for other reasons. First, use of alcohol by patients on opioid therapy for chronic pain increases the risks of overdose and death.<sup>65</sup> Second, alcohol interferes with the efficacy of antidepressant treatment of depression,<sup>66</sup> and likely pain. Third, persistent pain reduces the effectiveness of treatment for alcohol disorders.<sup>67</sup>

We observed an interesting trend, to our knowledge not discussed in the literature previously. In models which contain measures of both pain frequency and depression as independent variables, depression is most strongly associated with cocaine abuse/dependence and marijuana abuse/dependence, while pain frequency is most strongly associated with alcohol abuse/dependence and opioid abuse/dependence. To a certain degree, this makes sense from a self-medication perspective. Although opioids are the oldest known anti-depressants, their analgesic effects are likely greater than their effects on mood, in most individuals. In contrast, the effects of cocaine on mood are likely greater than the analgesic effects. On the other hand, we expected a statistically significant positive association between marijuana abuse/dependence and number of days with pain, given the established analgesic properties of cannabinoids. However, not only was the relationship not statistically significant, the magnitude of the relationship was modest, with OR's ranging from 1.17 to 1.25.

As mentioned, our results should be viewed as measures of association, and do not necessarily imply causality, due to possible complex, bi-directional effects. Besides self-medication, there are other possible explanations for our significant results. For example, because of potential hyperalgesia, opioids and alcohol may worsen pain over time, although this is controversial for opioids and to our knowledge, has never been investigated for alcohol.

Our use of longitudinal data, with multiple data points for each individual, allowed us to use fixed-effects models, which we view as a distinct advantage in our study. In particular, fixed-effects models control for all stable characteristics of the individual. This is important, as potentially important, stable factors, such as personality, and personality disorders have generally not been controlled for in previous analyses. Thus, the potential for biased coefficients resulting from omitted confounders is decreased.

Several limitations deserve discussion. First, the measures of pain in the RSS have not been validated. Second, we utilized a sample of individuals in rural Arkansas and Kentucky who were at risk for stimulant use disorders, and who had extensive use of other substances. The extent to which results from this population can be generalized to the U.S population is unknown, and our results need to be replicated and validated in other studies with more representative populations. We would note however that future studies that investigate these issues might also have to use non-representative samples, as statistical power considerations necessitate the investigation of outcomes that are relatively infrequent, such as opioid abuse/dependence, or methamphetamine abuse/dependence, in samples that are significantly enriched for these disorders. Third, while we controlled for depression severity, the RSS did not contain measures of other common mental health disorders, such as anxiety disorders. Fourth, our data do not allow us to determine whether and individual with an opioid use

disorder was using the opioid solely for the opioid use disorder, or for the opioid use disorder and pain.

## 5. CONCLUSIONS

Pain frequency appears to be associated with increased risk for both opioid abuse/dependence and alcohol abuse/dependence. The population-attributable risk for alcohol abuse/dependence from frequent pain may be greater than that associated with depression, as chronic pain is more prevalent than depression. Pain is also known to impede treatment for depression, so that risk may be magnified when both pain and depression are present.<sup>68</sup> Further research is needed to investigate the relationship between alcohol abuse/dependence and pain frequency in more representative samples.

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## Bibliography

- Franklin KB. Analgesia and abuse potential: an accidental association or a common substrate? *Pharmacol Biochem Behav.* 1998; 59:993–1002. [PubMed: 9586860]
- Jarcho JM, Mayer EA, Jiang ZK, et al. Pain, affective symptoms, and cognitive deficits in patients with cerebral dopamine dysfunction. *Pain.* 2012; 153:744–54. [PubMed: 22386471]
- Joy, JE.; Watson, SJ.; Benson, JA., editors. *Marijuana and medicine: assessing the science base.* Washington, DC: National Academy Press; 1999.
- Wilsey B, Marcotte T, Tsodikov A, et al. A randomized, placebo-controlled, crossover trial of cannabis cigarettes in neuropathic pain. *J Pain.* 2008; 9:506–21. [PubMed: 18403272]
- Perrino AC Jr, Ralevski E, Acampora G, et al. Ethanol and pain sensitivity: effects in healthy subjects using an acute pain paradigm. *Alcohol Clin Exp Res.* 2008; 32:952–8. [PubMed: 18445106]
- Pamplona FA, Vendruscolo LF, Takahashi RN. Increased sensitivity to cocaine-induced analgesia in Spontaneously Hypertensive Rats (SHR). *Behav Brain Funct.* 2007; 3:9. [PubMed: 17298672]
- Brill S, Gurman GM, Fisher A. A history of neuraxial administration of local analgesics and opioids. *Eur J Anaesthesiol.* 2003; 20:682–9. [PubMed: 12974588]
- Dalal S, Melzack R. Potentiation of opioid analgesia by psychostimulant drugs: a review. *J Pain Symptom Manage.* 1998; 16:245–53. [PubMed: 9803052]
- Sheu R, Lussier D, Rosenblum A, et al. Prevalence and Characteristics of Chronic Pain in Patients Admitted to an Outpatient Drug and Alcohol Treatment Program. *Pain Med.* 2008
- Mertens JR, Lu YW, Parthasarathy S, et al. Medical and psychiatric conditions of alcohol and drug treatment patients in an HMO: comparison with matched controls. *Arch Intern Med.* 2003; 163:2511–7. [PubMed: 14609789]
- Atkinson JH, Slater MA, Patterson TL, et al. Prevalence, onset, and risk of psychiatric disorders in men with chronic low back pain: a controlled study. *Pain.* 1991; 45:111–21. [PubMed: 1831555]
- Barry DT, Beitel M, Garnet B, et al. Relations among psychopathology, substance use, and physical pain experiences in methadone-maintained patients. *J Clin Psychiatry.* 2009; 70:1213–8. [PubMed: 19607760]
- Gureje O, Von Korff M, Kola L, et al. The relation between multiple pains and mental disorders: Results from the World Mental Health Surveys. *Pain.* 2007:12.
- Demyttenaere K, Bruffaerts R, Lee S, et al. Mental disorders among persons with chronic back or neck pain: Results from the World Mental Health Surveys. *Pain.* 2007; 129(3):332–42. [PubMed: 17350169]
- He Y, Zhang M, Lin EH, et al. Mental disorders among persons with arthritis: results from the World Mental Health Surveys. *Psychol Med.* 2008:1–12.



16. Gureje O, Simon GE, Von Korff M. A cross-national study of the course of persistent pain in primary care. *Pain*. 2001; 92(1–2):195–200. [PubMed: 11323140]
17. Banks SM, Kerns RD. Explaining high rates of depression in chronic pain: a diathesis-stress framework. *Psychol Bull*. 1996; 119:95–100.
18. Sareen J, Jacobi F, Cox BJ, et al. Disability and poor quality of life associated with comorbid anxiety disorders and physical conditions. *Arch Intern Med*. 2006; 166(19):2109–16.
19. Von Korff M, Crane P, Lane M, et al. Chronic spinal pain and physical-mental comorbidity in the United States: Results from the National Comorbidity Survey Replication. *Pain*. 2005; 113:331–9. [PubMed: 15661441]
20. McWilliams LA, Goodwin RD, Cox BJ. Depression and anxiety associated with three pain conditions: results from a nationally representative sample. *Pain*. 2004; 111:77–83. [PubMed: 15327811]
21. Grant BF, Stinson FS, Dawson DA, et al. Prevalence and co-occurrence of substance use disorders and independent mood and anxiety disorders: Results from the National Epidemiologic Survey on Alcohol and Related Conditions. *Arch Gen Psychiatry*. 2004; 61(8):807–16. [PubMed: 15289279]
22. Kessler RC, Nelson CB, McGonagle KA, et al. The epidemiology of co-occurring addictive and mental disorders: implications for prevention and service utilization. *Am J Orthopsychiatry*. 1996; 66(1):17–31. [PubMed: 8720638]
23. Kessler RC, Chiu WT, Demler O, et al. Prevalence, severity, and comorbidity of 12-month DSM-IV disorders in the National Comorbidity Survey Replication. *Arch Gen Psychiatry*. 2005; 62:617–27. [PubMed: 15939839]
24. Kessler RC. The epidemiology of dual diagnosis. *Biol Psychiatry*. 2004; 56:730–7. [PubMed: 15556117]
25. Khantzian EJ. The self-medication hypothesis of substance use disorders: A reconsideration and recent applications. *Harv Rev Psychiatry*. 1997; 4(5):231–44. [PubMed: 9385000]
26. Borders TF, Booth BM, Han X, et al. Longitudinal changes in methamphetamine and cocaine use in untreated rural stimulant users: racial differences and the impact of methamphetamine legislation. *Addiction*. 2008; 103:800–8. [PubMed: 18412758]
27. Booth BM, Leukefeld C, Falck R, et al. Correlates of rural methamphetamine and cocaine users: results from a multistate community study. *J Stud Alcohol*. 2006; 67:493–501. [PubMed: 16736068]
28. Falck RS, Wang J, Carlson RG, et al. Perceived need for substance abuse treatment among illicit stimulant drug users in rural areas of Ohio, Arkansas, and Kentucky. *Drug Alcohol Depend*. 2007; 91:107–14. [PubMed: 17604917]
29. Garrity TF, Leukefeld CG, Carlson RG, et al. Physical health, illicit drug use, and demographic characteristics in rural stimulant users. *J Rural Health*. 2007; 23:99–107. [PubMed: 17397365]
30. Siegal HA, Draus PJ, Carlson RG, et al. Perspectives on health among adult users of illicit stimulant drugs in rural Ohio. *J Rural Health*. 2006; 22:169–73. [PubMed: 16606430]
31. Wang J, Carlson RG, Falck RS, et al. Multi-sample standardization and decomposition analysis: an application to comparisons of methamphetamine use among rural drug users in three American states. *Stat Med*. 2007; 26:3612–23. [PubMed: 17243192]
32. Heckathorn DD, Semaan S, Broadhead RS, et al. Extensions of respondent-driven sampling: a new approach to the study of injection drug users aged 18–25. *AIDS Behavior*. 2006; 6:55–67.
33. Heckathorn DD. Respondent-driven sampling II: Deriving valid population estimates from chain-referral samples of hidden populations. *Soc Probl*. 2002; 49:11–34.
34. Heckathorn DD. Respondent-driven sampling: a new approach to the study of hidden populations. *Soc Probl*. 1997; 44:174–99.
35. Wang J, Falck RS, Li L, et al. Respondent-driven sampling in the recruitment of illicit stimulant drug users in a rural setting: findings and technical issues. *Addict Behav*. 2007; 32:924–37. [PubMed: 16901654]
36. Wang J, Carlson RG, Falck RS, et al. Respondent-driven sampling to recruit MDMA users: a methodological assessment. *Drug Alcohol Depend*. 2005; 78:147–57. [PubMed: 15845318]
37. Draus PJ, Siegal HA, Carlson RG, et al. Cracking the cornfields: recruiting illicit stimulant drug users in rural Ohio. *Sociol Q*. 2005; 46:165–89.

38. Mieczkowski T, Newel R, Wraight B. Using hair analysis, urinalysis, and self-reports to estimate drug use in a sample of detained juveniles. *Subst Use Misuse*. 1998; 33:1547–67. [PubMed: 9657415]
39. Hser YI. Self-reported drug use: results of selected empirical investigations of validity. *NIDA Res Monogr*. 1997; 167:320–43. [PubMed: 9243568]
40. Hamid R, Deren S, Beardsley M, et al. Agreement between urinalysis and self-reported drug use. *Subst Use Misuse*. 1999; 34:1585–92. [PubMed: 10468109]
41. Smith GR, Burnam MA, Mosley CL, et al. Reliability and validity of the substance abuse outcomes module. *Psychiatr Serv*. 2006; 57:1452–60. [PubMed: 17035565]
42. Cottler LB, Robins LN, Helzer JE. The reliability of the CIDI-SAM: a comprehensive substance abuse interview. *Br J Addict*. 1989; 84:801–14. [PubMed: 2758153]
43. McLellan AT, Kushner H, Metzger D, et al. The Fifth Edition of the Addiction Severity Index. *J Subst Abuse Treat*. 1992; 9:199–213. [PubMed: 1334156]
44. Kroenke K, Spitzer RL, Williams JB. The PHQ-9: validity of a brief depression severity measure. *J Gen Intern Med*. 2001; 16:606–13. [PubMed: 11556941]
45. Allison, PD. *Fixed Effects Regression Methods for Longitudinal Data: Using SAS*. SAS Publishing; 2005.
46. Substance Abuse Mental Health Services Administration. Office of Applied Studies: Summary of Findings from the 2000 National Household Survey on Drug Abuse. Rockville, MD: Department of Health and Human Services; 2000.
47. Kessler RC, Berglund P, Demler O, et al. Lifetime prevalence and age-of-onset distributions of DSM-IV disorders in the National Comorbidity Survey Replication. *Arch Gen Psychiatry*. 2005; 62:593–602. [PubMed: 15939837]
48. Murray, C.; Lopez, A. *The Global Burden of Disease: A Comprehensive Assessment of Mortality and Disability from Diseases, Injuries, and Risk Factors in 1990 and Projected to 2020*. Boston: The Harvard School of Public Health on Behalf of the World Health Organization and the World Bank; 1996.
49. Murray CJ, Lopez AD. Evidence-based health policy--lessons from the Global Burden of Disease Study. *Science*. 1996; 274:740–3. [PubMed: 8966556]
50. Room R, Babor T, Rehm J. Alcohol and public health. *Lancet*. 2005; 365:519–30. [PubMed: 15705462]
51. Conner KR, Hesselbrock VM, Meldrum SC, et al. Transitions to, and correlates of, suicidal ideation, plans, and unplanned and planned suicide attempts among 3,729 men and women with alcohol dependence. *J Stud Alcohol Drugs*. 2007; 68:654–62. [PubMed: 17690798]
52. Borges G, Walters EE, Kessler RC. Associations of substance use, abuse, and dependence with subsequent suicidal behavior. *Am J Epidemiol*. 2000; 151:781–9. [PubMed: 10965975]
53. National Institute on Alcohol Abuse and Alcoholism. *Alcohol Research and Health*. 2001. Alcohol and Violence.
54. McGinnis JM, Foegen WH. Actual causes of death in the United States. *JAMA*. 1993; 270:2207–12. [PubMed: 8411605]
55. Caetano R. Alcohol-related health disparities and treatment-related epidemiological findings among whites, blacks, and Hispanics in the United States. *Alcoholism: Clinical and Experimental Research*. 2003; 27(8):1337–9.
56. NIAAA: Plan to Address Health Disparities Initiatives. [Web site]. 2006. Available at: [www.niaaa.nih.gov/aboutNIAAA/NIAAASponsoredPrograms/HealthDispintro.htm](http://www.niaaa.nih.gov/aboutNIAAA/NIAAASponsoredPrograms/HealthDispintro.htm)
57. Tsang A, Von Korff M, Lee S, et al. Common chronic pain conditions in developed and developing countries: gender and age differences and comorbidity with depression-anxiety disorders. *J Pain*. 2008; 9:883–91. [PubMed: 18602869]
58. Kroenke K, Krebs EE, Bair MJ. Pharmacotherapy of chronic pain: a synthesis of recommendations from systematic reviews. *Gen Hosp Psychiatry*. 2009; 31:206–19. [PubMed: 19410099]
59. Chou R, Huffman LH. Nonpharmacologic therapies for acute and chronic low back pain: a review of the evidence for an American Pain Society/American College of Physicians clinical practice guideline. *Ann Intern Med*. 2007; 147:492–504. [PubMed: 17909210]

60. Morasco BJ, Duckart JP, Dobscha SK. Adherence to Clinical Guidelines for Opioid Therapy for Chronic Pain in Patients with Substance Use Disorder. *J Gen Intern Med.* 2011
61. Chelminski PR, Ives TJ, Felix KM, et al. A primary care, multi-disciplinary disease management program for opioid-treated patients with chronic non-cancer pain and a high burden of psychiatric comorbidity. *BMC Health Services Research.* 2005; 5:3. [PubMed: 15649331]
62. Currie SR, Hodgins DC, Crabtree A, et al. Outcome from integrated pain management treatment for recovering substance abusers. *J Pain.* 2003; 4:91–100. [PubMed: 14622720]
63. Ilgen MA, Haas E, Czyz E, et al. Treating chronic pain in veterans presenting to an addictions treatment program. *Cogn Behav Pract.* 2011; 18:149–60.
64. Rhodin A, Gronbladh L, Nilsson LH, et al. Methadone treatment of chronic non-malignant pain and opioid dependence--a long-term follow-up. *Eur J Pain.* 2006; 10:271–8. [PubMed: 15972261]
65. Braden JB, Russo J, Fan MY, et al. Emergency department visits among recipients of chronic opioid therapy. *Arch Intern Med.* 2010; 170:1425–32. [PubMed: 20837827]
66. Nunes EV, Levin FR. Treatment of depression in patients with alcohol or other drug dependence. *JAMA.* 2004; 291(15):1887–96. [PubMed: 15100209]
67. Caldeiro RM, Malte CA, Calsyn DA, et al. The association of persistent pain with out-patient addiction treatment outcomes and service utilization. *Addiction.* 2008; 103:1996–2005. [PubMed: 18855809]
68. Thielke SM, Fan MY, Sullivan M, et al. Pain limits the effectiveness of collaborative care for depression. *Am J Geriatr Psychiatry.* 2007; 15:699–707. [PubMed: 17670998]

**Table 1**

The Rural Stimulant Study (n=462): Baseline Descriptive Statistics

Variable	n (%)
State	
Arkansas	237 (51.3)
Kentucky	225 (48.7)
Sex	
Male	270 (58.4)
Female	192 (41.6)
Age (mean/SD/range)	34.1 ± 10.6 (18–61)
Race	
African-American	183 (39.6)
White	266 (57.6)
Hispanic/Latino	3 (0.6)
Native American	2 (0.4)
Other	8 (1.7)
Marital Status	
Single	215 (46.5)
Married	247 (53.5)
High school graduate?	
Yes	188 (40.7)
No	274 (59.3)
Employment Status	
Employed	142 (30.7)
Unemployed	320 (69.3)
Income	
less than \$10,000	108 (23.5)
\$10,000 or more	354 (76.5)
Days with pain in the past 30 days	
0 days	138 (30.1)
1–15 days	198 (43.2)
16+ days	122 (26.6)
Depression severity	
No depression	161 (34.8)
Mild depression	122 (26.4)
Moderate depression	108 (23.4)
Moderately severe or severe depression	71 (15.4)
<i>Dependent variables</i>	
Alcohol abuse/dependence (past 6 months)	234 (50.6)
Cocaine abuse/dependence (past 6 months)	239 (51.7)
Methamphetamine abuse/dependence (past 6 months)	158 (34.2)
Opioid abuse/dependence (past 6 months)	23 (5.0)

Variable	n (%)
Marijuana abuse/dependence (past 6 months)	168 (36.4)

**Table 2**  
Pain Frequency and Depression Severity as Predictors of Substance Use Disorders in Fixed Effects Models

Independent Variables	Alcohol abuse/dependence		Cocaine abuse/dependence		Methamphetamine abuse/dependence		Opioid abuse/dependence		Marijuana abuse/dependence	
	OR	p-value	OR	p-value	OR	p-value	OR	p-value	OR	p-value
<b>Days with bodily pain in past 30 days</b>										
1-15 days vs. 0 day	1.85	0.0002	1.25	0.24	1.44	0.14	2.25	0.06	1.25	0.23
16+ days vs. 0 day	2.18	0.0003	1.50	0.09	1.57	0.13	3.32	0.02	1.17	0.50
<b>Depression severity</b>										
Mild depression vs. No depression	1.18	0.33	1.91	0.0006	1.34	0.23	1.11	0.79	1.62	0.008
Moderate depression vs. No depression	1.34	0.19	2.71	<0001	2.35	0.006	1.31	0.57	2.38	0.0003
Moderately severe or severe depression vs. No depression	1.29	0.33	3.45	<0001	1.56	0.21	2.67	0.04	1.97	0.01