

## Community Diarrhea Incidence Before and After Rotavirus Vaccine Introduction in Nicaragua

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**Abstract.** We estimated the incidence of watery diarrhea in the community before and after introduction of the pentavalent rotavirus vaccine in León, Nicaragua. A random sample of households was selected before and after rotavirus vaccine introduction. All children < 5 years of age in selected households were eligible for inclusion. Children were followed every 2 weeks for watery diarrhea episodes. The incidence rate was estimated as numbers of episodes per 100 child-years of exposure time. A mixed effects Poisson regression model was fit to compare incidence rates in the pre-vaccine and vaccine periods. The pre-vaccine cohort ( $N = 726$ ) experienced 36 episodes per 100 child-years, and the vaccine cohort ( $N = 826$ ) experienced 25 episodes per 100 child-years. The adjusted incidence rate ratio was 0.60 (95% confidence interval [CI] 0.40, 0.91) during the vaccine period versus the pre-vaccine period, indicating a lower incidence of watery diarrhea in the community during the vaccine period.

### INTRODUCTION

Diarrheal diseases cause 1.3 million deaths of children each year and are an important cause of malnutrition, growth retardation, and developmental delays.<sup>1–3</sup> Although child deaths caused by diarrhea have declined substantially since the introduction of oral rehydration therapy, there has been little change in diarrhea incidence in low and middle income countries (LMICs) since the 1950s.<sup>4</sup>

Vaccines against rotavirus, one of the most common causes of diarrhea,<sup>5</sup> offer promise in reducing diarrhea incidence in LMIC settings. Rotavirus vaccines show excellent protection against rotavirus diarrhea in clinical trials conducted primarily in the industrialized world.<sup>6,7</sup> Evaluations of the vaccines' effectiveness in LMICs report less robust protection.<sup>8–11</sup> These evaluations have focused on the prevention of hospitalizations and severe diarrhea, although little is known about the vaccine's effect on diarrhea incidence in the community. However, the majority of rotavirus diarrhea cases are treated at home: 80–85% of cases do not receive treatment at any health facility.<sup>12</sup>

The goal of this study was to examine the childhood diarrhea incidence rate at the community level in Nicaragua, the first LMIC to introduce the pentavalent rotavirus vaccine (RV5, Rotateq<sup>®</sup>) in 2006. We examined watery diarrhea incidence in León, Nicaragua before and after RV5 introduction using active surveillance. Because rotavirus traditionally has a seasonal transmission pattern, we also examined changes in seasonal patterns of diarrhea incidence.

### METHODS

**Setting.** The study was performed in the municipality of León, Nicaragua's second largest city, with an estimated 2010

population of 192,164. Municipal water treatment plants have been in continuous use in León since the 1990s. Improvements in child mortality have been made in recent decades,<sup>13</sup> however, diarrhea remains among the most important causes of mortality and morbidity in children in León.<sup>14</sup>

Diarrhea incidence in Central America has two annual peaks; the dry season peak is attributed primarily to rotavirus infections, whereas the rainy season peak has traditionally been attributed to bacterial pathogens, such as enterotoxigenic *Escherichia coli* and *Shigella* spp.<sup>15–17</sup> Overall, rotavirus was considered one of the most common causes of diarrhea among Nicaraguan children before the immunization program, isolated among 28% of children receiving care for diarrhea in health facilities.<sup>18</sup>

To reduce the burden of rotavirus diarrhea, Nicaragua implemented a rotavirus immunization program with RV5 in October 2006. Nicaraguan infants are offered the vaccine at the age of 2, 4, and 6 months as part of the country's Expanded Program on Immunization (EPI).

**Sample selection.** Cohorts from two distinct periods were examined in the study: before implementation of the immunization program, or, "pre-vaccine" (December 2000 to March 2001, and July 2002 to January 2003; broken into two segments because of budgetary issues) and following implementation of the immunization program, or, "vaccine" (January 2010 to January 2011). The time periods were chosen to include an equal number of months during the rainy and dry seasons.

In each time period, a random sample of households was selected from the Health and Demographic Surveillance Site-León (HDSS-León). The HDSS-León provides longitudinal surveillance for a population-based sample of 28% of León's population, residing in 50 out of 208 randomly selected geographical clusters of equal population size<sup>19</sup> (Figure 1). Children < 5 years of age in each sampled household were eligible for inclusion in the study. An "open cohort" design was used: children were excluded from the study after their fifth birthday, and new infants born into the household or children moving into the household during the study periods were eligible for inclusion.

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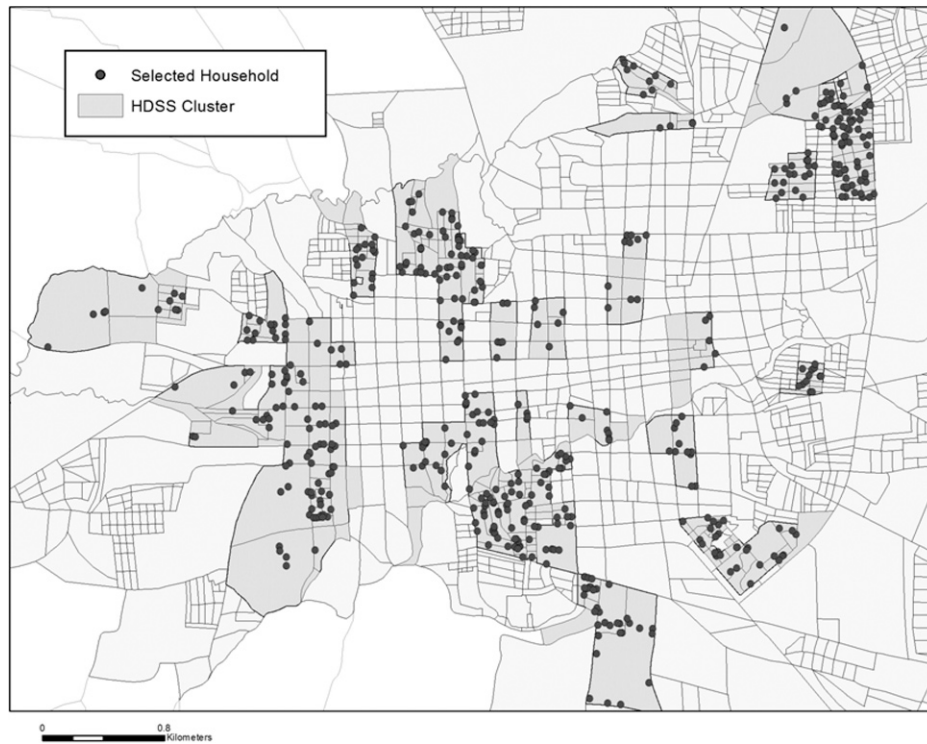


FIGURE 1. Map of León, Nicaragua: Geographic clusters and households included in HDSS-León surveillance site.

**Data collection and study instruments.** Female field workers visited sampled households every 2 weeks and performed household interviews. The interview was conducted with the mother or with the child's caretaker if the mother was unavailable. Quality control of the interviews and data collection was performed by the field supervisor with systematic and random evaluations.

The study instrument contained information on diarrhea episodes and clinical characteristics as reported by the mother or caretaker, family characteristics, household characteristics, breastfeeding history, and during the vaccine period only, a detailed rotavirus vaccine history taken from the child's immunization card.

At each interview, the field worker asked about all diarrhea episodes within the past 14 days, and collected information on the presence of vomiting, fever, and the presence of visible blood in the stool. A clinical definition of watery diarrhea suggestive of rotavirus infection was developed based on the literature, and defined as a diarrhea episode with a maximum of four or more bowel movements per 24 hour period, with the presence of vomiting or fever (or both), and without any blood in the stool.<sup>20–25</sup> An episode was considered to be new if the participant experienced at least 3 days without an episode before the onset of a new episode. In addition to diarrhea episodes, breastfeeding status, and rotavirus vaccine status were also updated at each household visit.

The study was approved by the Institutional Review Boards of the Universidad Nacional Autónoma de Nicaragua, León and the University of North Carolina at Chapel Hill. Informed consent was obtained from a parent or legal guardian of each participant.

**Statistical analysis.** Mixed effects linear and logistic regression models were used to compare participant characteristics in the pre-vaccine and vaccine periods. The mixed effects

models allowed for possible correlation between multiple participants living in the same household. The incidence rate of watery diarrhea was estimated as the number of episodes per 100 child-years of exposure time, and was stratified by age group (0–23 months, 24–59 months). The exposure time was estimated at the time each child was followed during the study and at-risk for a watery diarrhea episode; days with watery diarrhea were not included in the exposure time. Official rainfall data from León's regional airport<sup>26</sup> were used to assign each month of the study as either rainy ( $\geq 6$  mm per month) or dry ( $< 6$  mm per month).

A mixed effects Poisson regression model was fit to compare incidence rates in the pre-vaccine and vaccine periods. The mixed effects model allowed for possible correlation between multiple participants living in the same household as well as multiple measurements for each individual, and also allowed for adjustment for potential confounders. Incidence rates were compared between the rainy and dry seasons. Furthermore, in the vaccine period, incidence rates were compared between fully or partially immunized children and unimmunized children. These results were stratified by age group because of potential differences in the vaccine's effect by age group.

## RESULTS

A total of 726 children were enrolled during the pre-vaccine period and followed for 306.9 years of exposure time; 826 children were enrolled in the vaccine period and followed for 610.6 years of exposure time. Characteristics of children during the pre-vaccine and vaccine periods are shown in Table 1. Households observed during the pre-vaccine period had fewer indoor toilets compared with the vaccine period (46.7% versus

TABLE 1  
Characteristics of the participants in the pre-vaccine and vaccine periods

Characteristic	Pre-vaccine N = 726	Vaccine N = 826	P value for difference*
<b>Personal characteristics</b>			
Age (mean in months [SD])	25.3 [17.3]	28.9 [16.7]	< 0.001
Sex, % male	51.2% (372/726)	50.4% (416/826)	0.73
Breastfeeding upon study enrollment	36.4% (247/679)†	34.0% (281/826)	0.34
Received at least one dose of rotavirus vaccine	0.0% (0/726)	67.3% (556/826)	< 0.001
<b>Household characteristics</b>			
Indoor municipal water source	94.1% (683/726)	97.0% (801/826)	0.05
Indoor toilet	46.7% (339/726)	80.1% (662/826)	< 0.001
Concrete or tile floor (versus dirt floor)	72.9% (529/726)	77.2% (638/826)	0.008
<b>Family characteristics</b>			
Mother completed fourth grade	82.4% (598/726)	90.1% (744/826)	0.007
Mother employed	33.3% (242/726)	40.0% (330/826)	< 0.001

\* P values based on mixed effects logistic or linear regression models with each characteristic as the outcome and an indicator for vaccine period as the sole covariate.

† 6% of respondents chose not to answer this question.

80.1%,  $P < 0.001$ ) and mothers of children in the vaccine period were slightly more likely to be employed (33.3% versus 40.0%,  $P < 0.001$ ) and have completed a basic level of education (82.4% versus 90.1%,  $P = 0.007$ ). In all, 78.1% of vaccine period children were eligible to receive the rotavirus vaccine based on having a birthday on or after August, 2006. Upon study entry, 67.3% of all children in the vaccine period had received at least one dose of the rotavirus vaccine; among children < 24 months of age, 86.6% had received at least one dose of vaccine, whereas among children 24 to 60 months of age, 53.2% of children had received at least one dose of vaccine.

The crude watery diarrhea incidence rate was 35.8 episodes per 100 child-years in the pre-vaccine period and 24.9 episodes per 100 child-years in the vaccine period (Table 2). Incidence rates stratified by age group are shown in Figure 2; among both age groups there was a lower watery diarrhea incidence rate in the vaccine period.

Table 3 shows the adjusted incidence rate ratios (IRR<sub>a</sub>) of watery diarrhea in the vaccine versus pre-vaccine periods, as estimated by the mixed effects Poisson regression model. In the model, we adjusted for sex, breastfeeding, maternal education, maternal employment status, living in a household with a dirt floor, indoor municipal water source, season, and indoor toilet. Among all ages, the IRR<sub>a</sub> was 0.60 (95% confidence interval [CI] 0.40, 0.91) in the vaccine period as compared with the pre-vaccine period ("overall effect"). Among children < 24 months of age the IRR<sub>a</sub> was 0.63 (95% CI 0.37, 1.05), whereas among children between 24 and 59 months of age the IRR<sub>a</sub> was 0.56 (95% CI 0.32, 1.07). Among unimmunized children only, the IRR<sub>a</sub> was 0.27 (95% CI: 0.13, 0.60) in the vaccine period as compared with the pre-vaccine period; this result is consistent with an indirect effect of the

immunization program, but could also be caused by other differences between the pre-vaccine and vaccine period. During the vaccine period, there were no significant differences in watery diarrhea incidence rate observed between immunized or partially unimmunized children as compared with unimmunized children, stratified by age group ("direct effect"). In both time periods, there was also no difference in diarrhea incidence detected between the rainy seasons as compared with the dry season.

## DISCUSSION

Using active community surveillance in León, Nicaragua, we observed a 40% lower incidence rate of watery diarrhea episodes suggestive of rotavirus infection in the vaccine period as compared with the pre-vaccine period. This reduction may be attributable to an overall protective effect of the immunization program on immunized and unimmunized children. The findings correlate with an observed decline in rotavirus prevalence among children seeking care for diarrhea in primary care clinics in León following the vaccine's introduction (14.0% in the pre-vaccine period versus 3.5% in the vaccine period).<sup>27</sup>

The reduction in incidence observed among unimmunized children provides evidence for herd protection conferred by rotavirus immunization programs, as noted in prior studies.<sup>28</sup> Surprisingly, the reduction in incidence among unimmunized children (IRR = 0.27, 95% CI 0.13, 0.60) was greater than

TABLE 2 Watery diarrhea in pre-vaccine and vaccine periods		
	Pre-vaccine	Vaccine
Number of households	411	530
Number of child participants	726	826
Numbers of children with watery diarrhea episodes	91	117
Total number of watery diarrhea episodes	110	152
Person-years in the study	306.9	610.6
Overall watery diarrhea incidence (episodes per 100 child-years)	35.8	24.9

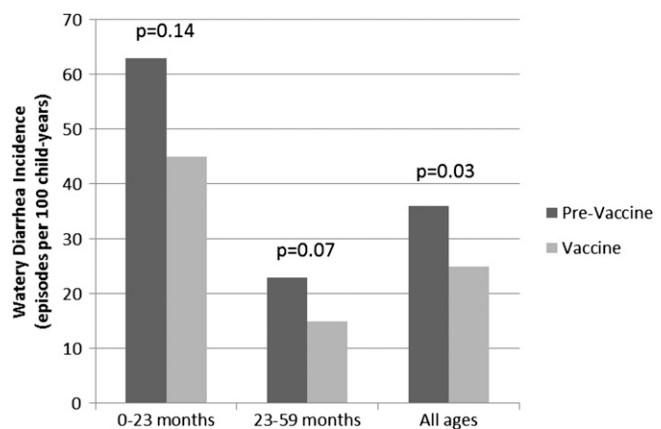


FIGURE 2. Watery diarrhea incidence by age group.

TABLE 3

Adjusted incidence rate ratios\* (IRR<sub>a</sub>) of watery diarrhea in the vaccine versus pre-vaccine periods

Age group	IRR <sub>a</sub>	95% confidence intervals
0–24 months	0.63	0.37, 1.05
24–59 months	0.56	0.32, 1.07
All ages	0.60	0.40, 0.91

\* Adjusted by sex, breastfeeding, household water source, indoor toilet, dirt floor, season, maternal education, and maternal employment.

among the population as a whole (IRR = 0.60, 95% CI 0.40, 0.91). An additional analysis found no differences in IRR among unimmunized children living in neighborhoods with low versus high vaccine coverage. Although our findings were unexpected, a pattern of greater reductions in rotavirus diarrhea incidence among age groups with low versus high vaccine coverage was previously observed in a United States study of rotavirus diarrhea hospitalizations in the years before and after vaccine introduction.<sup>29</sup> In our study of a population with high vaccine coverage, it is possible that our sample size of unimmunized children was not large enough to detect a true difference in incidence among unimmunized children as compared with immunized children in the vaccine period. Alternatively, it is conceivable that rotavirus immunization could increase the risk of infection by a different diarrheal pathogen, which would provide the unimmunized children an advantage within this highly immunized population; however, there is no biological explanation for this alternate hypothesis.

We did not observe a statistically significant direct effect of the vaccine following introduction, which may be the result of strong herd protection effects afforded by the immunization program in a setting with high vaccine coverage. A similar phenomenon was observed by Ali and others<sup>30</sup> in an analysis of herd immunity conferred by killed oral cholera vaccines. To further investigate the direct effect of the rotavirus vaccine, we examined the dose effect by estimating age-stratified incidence rates among those who had received one, two, or three doses of vaccine. Among all age groups, the incidence rate was lower among those who had received the full course of vaccine than those who had received only two doses. However, among infants, the incidence rate among those receiving the full course was no different than those receiving just one dose. This could be explained by different exposures to enteric pathogens among infants who had received one dose as compared with those who had completed the full course; the majority of infants who had received just one dose were two to three months of age, were breastfed, and were less likely to interact with their environment as compared with older infants, who had received the full course.

This study highlights the importance of health and demographic surveillance systems in assessing disease rates in a population after an intervention. As the majority of diarrhea cases are typically treated at home, they would not be captured by more commonly used hospital-based surveillance systems.

Our study has several limitations. First, as an observational study, we cannot prove that differences in incidence rates between the pre-vaccine and vaccine periods were caused by the immunization program and not by other factors. We did observe certain improvements in living standards between the two time periods. Although we included these socioeconomic and household variables in our regression model, we did not find them to be associated with watery diarrhea episodes. It is possible, however, that factors associated with improvements

in living standards that we did not measure may have contributed to the overall decrease in diarrhea incidence rate observed in the vaccine period. Second, this report does not include laboratory data on the causes of diarrhea episodes, but instead used a clinically defined watery diarrhea definition based on characteristics of rotavirus diarrhea reported in the literature. Clearly, this definition is not as specific as laboratory analysis, but allows a narrowing of diarrhea cases that are more likely to be caused by rotavirus infection, and was feasible in a community setting. As this study was conducted in the community, the definition relied on the mother's or caretaker's reports of fever during the episode. Prior studies have shown that mothers detect fevers in their children with a high sensitivity, although with a lower specificity than when a thermometer is used by a health provider.<sup>31,32</sup> Therefore, our definition may have overestimated the incidence of episodes. However, if the rate of overestimation was similar between the pre-vaccine and vaccine periods, then our estimated IRR would be biased towards the null, suggesting the true reduction in incidence associated with the introduction of the rotavirus vaccine may be even greater than indicated by this analysis. Finally, as rotavirus transmission can vary from year to year, we cannot rule out that differences in watery diarrhea incidence noted in the two time periods may be caused by such year-to-year variability.

In summary, the changes in watery diarrhea incidence observed in the two time periods support an overall and indirect effect of the rotavirus immunization program at the community level. Future studies should attempt to identify the etiologies of diarrhea that persist in the community following vaccine introduction, to better design prevention programs and further reduce the burden of childhood diarrhea.

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