

# Population pharmacokinetic– pharmacodynamic analysis for eribulin mesilate- associated neutropenia

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## WHAT IS ALREADY KNOWN ABOUT THIS SUBJECT

- Eribulin mesilate is an inhibitor of microtubule dynamics that is approved for the treatment of late-stage metastatic breast cancer.
- One of the major dose-limiting effects of eribulin mesilate is neutropenia.

## WHAT THIS STUDY ADDS

- Eribulin-associated neutropenia was described using a semi-physiological population pharmacokinetic–pharmacodynamic model.
- Patient characteristics predictive for interindividual variability in pharmacodynamic parameters were identified, with albumin being a clinically important covariate.
- Eribulin dosing guidelines to optimize dose for neutropenia based on expected changes in model parameters were generated.

## AIMS

Eribulin mesilate is an inhibitor of microtubule dynamics that is approved for the treatment of late-stage metastatic breast cancer. Neutropenia is one of the major dose-limiting adverse effects of eribulin. The objective of this analysis was to develop a population pharmacokinetic–pharmacodynamic model for eribulin-associated neutropenia.

## METHODS

A combined data set of 12 phase I, II and III studies for eribulin mesilate was analysed. The population pharmacokinetics of eribulin was described using a previously developed model. The relationship between eribulin pharmacokinetic and neutropenia was described using a semi-physiological lifespan model for haematological toxicity. Patient characteristics predictive of increased sensitivity to develop neutropenia were evaluated using a simulation framework.

## RESULTS

Absolute neutrophil counts were available from 1579 patients. In the final covariate model, the baseline neutrophil count ( $ANC_0$ ) was estimated to be  $4.03 \times 10^9$  neutrophils  $l^{-1}$  [relative standard error (RSE) 1.2%], with interindividual variability (IIV, 37.3 coefficient of variation % [CV%]). The mean transition time was estimated to be 109 h (RSE 1.8%, IIV 13.9CV%), the feedback constant ( $\gamma$ ) was estimated to be 0.216 (RSE 1.4%, IIV 12.2CV%), and the linear drug effect coefficient (SLOPE) was estimated to be  $0.0451 \mu g l^{-1}$  (RSE 3.2%, IIV 54CV%). Albumin, aspartate transaminase and receipt of granulocyte colony-stimulating factor (G-CSF) were identified as significant covariates on SLOPE, and albumin, bilirubin, G-CSF, alkaline phosphatase and lactate dehydrogenase were identified as significant covariates on mean transition time.

## CONCLUSIONS

The developed model can be applied to investigate optimal treatment strategies quantitatively across different patient groups with respect to neutropenia. Albumin was identified as the most clinically important covariate predictive of interindividual variability in the neutropenia time course.

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## Introduction

Treatment options for metastatic breast cancer that increase overall survival are limited, and include anthracyclines and taxanes. However, patients may not always show adequate response to these agents or may develop resistance [1]. The anticancer drug eribulin mesilate (E7389) is an inhibitor of microtubule dynamics and is the first in a new class of anticancer drugs referred to as the halichondrins.

The pivotal phase III trial (EMBRACE) of eribulin mesilate showed promising results for women with heavily pretreated metastatic breast cancer, demonstrating a significant benefit in the improvement of median overall survival of 2.5 months compared with the treatment of physician's choice [2]. Based on this trial, eribulin mesilate has been approved by the Food and Drug Administration, European Medicines Agency and Japanese regulatory authorities. In the USA, eribulin mesilate is approved for the treatment of patients with metastatic breast cancer who have previously received at least two chemotherapeutic regimens for the treatment of metastatic disease. Prior therapy should have included an anthracycline and a taxane in either the adjuvant or metastatic setting.

One of the major dose-limiting adverse effects of eribulin mesilate is occurrence of neutropenia. It is therefore important that the incidence and dynamics of this clinically important adverse effect are well understood across the full population of patients who are being treated with eribulin mesilate.

Population pharmacokinetic–pharmacodynamic (PK-PD) modelling may be used to describe the time course of neutropenia. Friberg *et al.* [3] have described a semi-physiological PK-PD model for haematological toxicity, consisting of system-specific and drug-specific parameters

[3]. Parameter consistency has been shown across drugs, and interspecies scaling has also been demonstrated [4]. The model has also been demonstrated to have predictive value during early clinical drug development [5], and has successfully been applied for a range of different drugs [6]. Moreover, a population PK-PD modelling approach allows identification of patient characteristics that are predictive of interpatient variability in PK or PD [7, 8]. This concept may also be applied to identify patient characteristics predictive for increased sensitivity to develop haematological toxicity, in order ultimately to determine optimized dosing schedules in different patient groups [9].

The objectives of this analysis were as follows: (i) the development of a semi-physiological population PK-PD model for eribulin-associated neutropenia; (ii) univariate identification of covariates predictive for interindividual variability in pharmacodynamic model parameters; and (iii) development of a multivariate covariate model that is predictive of interindividual variability in model parameters describing the dynamics of the absolute neutrophil count time course. Ultimately, this model may be applied to optimize drug treatment further in patients treated with eribulin mesilate and to support further clinical development.

## Methods

### Clinical studies

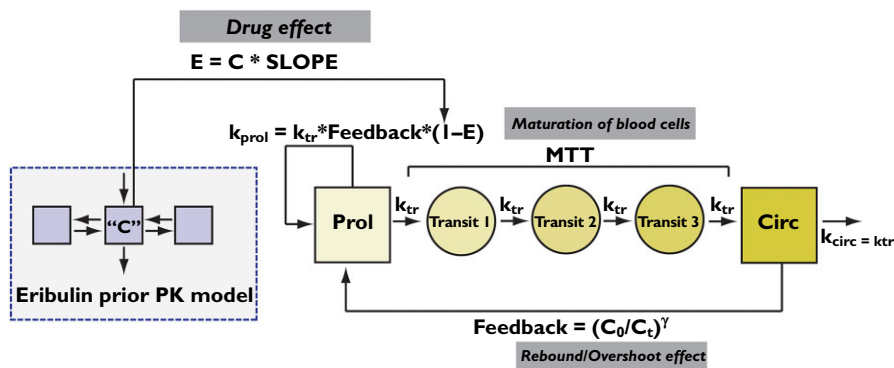
This analysis was conducted using a pooled data set of phase I, II and III studies of eribulin mesilate. An overview of the included studies is given in Table 1. The data set included 1579 individuals with 23 427 observed absolute neutrophil counts (ANC). Pharmacokinetic data were available for 428 patients (27%).

**Table 1**

Overview of clinical studies included in the analysis

Study	Phase	Indication	Subjects	Objective	Dose	Dose times	PK data	PD data	Reference
1	I	Solid tumours	33	MTD	0.25–1.4 mg m <sup>-2</sup> IV	Days 1, 8, 15 q 21	Yes	Yes	Goel <i>et al.</i> 2009
2	I	Solid tumours	21	MTD	0.25–4.0 mg m <sup>-2</sup> IV	Day 1 q 21	Yes	Yes	Tan <i>et al.</i> 2009
3	II	Breast cancer	104	ORR	1.4 mg m <sup>-2</sup> IV	Days 1, 8, 15 q 28	No	Yes	Vahdat <i>et al.</i> 2009
4	II	NSCLC	106	ORR	1.4 mg m <sup>-2</sup> IV	Days 1, 8, 15 q 28 or 1, 8 q 21	No	Yes	Spira <i>et al.</i> 2012
5	I	Solid tumours	6	ADME	1.4 mg m <sup>-2</sup> IV*	Days 1, 8 q 21*	Yes	No	Dubbelman <i>et al.</i> 2012
6	I	Solid tumours	17	Liver function	0.7–1.4 mg m <sup>-2</sup> IV	–	Yes	Yes	Devriese <i>et al.</i> 2012
7	I	Solid tumours	12	DDI	0.7–1.4 mg m <sup>-2</sup> IV	–	Yes	Yes	Devriese, Witteveen <i>et al.</i> 2013; Devriese, Mergui-Roelvink <i>et al.</i> 2013
8	I	Solid tumours	26	QTc prolongation	1.4 mg m <sup>-2</sup> IV	Days 1, 8 q 21	Yes	Yes	Lesimple <i>et al.</i> 2012
9	II	Prostate cancer	108	ORR	1.4 mg m <sup>-2</sup> IV	Days 1, 8 q 21	No	Yes	De Bono <i>et al.</i> 2012
10	II	Breast cancer	298	ORR	1.4 mg m <sup>-2</sup> IV	Days 1, 8 q 21	Yes	Yes	Cortes <i>et al.</i> 2010
11	III	Breast cancer	761	OS	1.4 mg m <sup>-2</sup> IV	Days 1, 8 q 21		Yes	Cortes <i>et al.</i> 2011
12	I	Solid tumours	15	MTD	0.7–2.0 mg m <sup>-2</sup> IV	Days 1, 8 q 21	Yes	Yes	Mukohara <i>et al.</i> 2012
13	II	Breast cancer	81	ORR	1.4 mg m <sup>-2</sup> IV	Days 1, 8 q 21	No	Yes	Aogi <i>et al.</i> 2012

\*First cycle a flat 2 mg dose of <sup>14</sup>C-eribulin mesilate on day 1 cycle 1 only; thereafter, 1.4 mg m<sup>-2</sup> on days 1 and 8. Abbreviations are as follows: ADME, absorption, distribution, metabolism, excretion study; DDI, drug–drug interaction study; IV, intravenous; MTD, maximal tolerated dose; NSCLC, nonsmall cell lung cancer; ORR, objective response rate; OS, overall survival; q, every; QTc, corrected QT interval.



**Figure 1**

Schematic representation of the semi-physiological model for haematological toxicity. C, drug plasma concentration; C<sub>0</sub>, circulating neutrophils at time = 0; C<sub>t</sub>, circulating neutrophils at time = t; E, drug effect; gamma, feedback constant; k<sub>circ</sub>, decay rate of circulating neutrophils; k<sub>prol</sub>, proliferation rate of neutrophils; k<sub>tr</sub>, maturation rate of neutrophils; MTT, mean transition time; PK, pharmacokinetics; SLOPE, linear drug effect coefficient.

All studies were approved by an Institutional Review Board or Independent Ethics Committee and conducted in accordance with International Conference on Harmonisation guidelines, the Declaration of Helsinki and good clinical practice.

### Estimation method and software

R (version 2.10) [10] was used for database handling and generation of diagnostic plots. Pirana (version 2.3.0) [11] was used for the model-building process. Parameter estimation and simulation of the nonlinear mixed effect models were performed using NONMEM (version 7.1.0) [12]. The use of the first-order conditional estimation with interaction method was preferred. However, if it proved not to be computationally feasible, the first-order estimation method was used.

### Pharmacokinetic data integration

A previously developed population PK model [unpublished data on file, Dr Z. Hussein (ziad\_hussein@eisai.net), Clinical Pharmacology & Translational Medicine, Eisai Limited, Hatfield, UK] including covariates was used to generate predicted PK profiles.

The PK model was systematically developed based on statistical significance of all included parameters, adequate goodness-of-fit plots, bootstrapping and visual predictive checks using both internal and external data sets. In summary, the PK model was a three-compartmental model with linear elimination. The covariates albumin, alkaline phosphatase (ALP) and bilirubin were related to clearance, the covariate dose was related to intercompartmental clearance, and fixed allometric scaling using bodyweight was included on all model parameters using exponents of 0.75 on clearances (CL, Q1 and Q2), and an exponent of 1 on volumes.

Using the previously developed PK model, either typical covariate-adjusted population parameters were used when no PK information was available or, alterna-

tively, when PK observations were available, empirical Bayes estimates were generated which described the individual PK profiles.

Subsequently, the individual or typical covariate-adjusted PK parameters were used as input for the sequential semi-physiological model for neutropenia.

The PK and sequential PK/PD modelling were based on the dose amount of the eribulin free base.

### Pharmacodynamic structural model development

The relationship between eribulin exposure and decrease in ANC was described using a semi-physiological lifespan model for haematological toxicity as described by Friberg *et al.* [3]. This model consists of drug-specific (e.g. linear drug effect coefficient [SLOPE] or concentration of 50% of the maximum effect [EC<sub>50</sub>] and maximum drug effect [E<sub>max</sub>]) and system-specific parameters [e.g. feedback parameter (γ) and the mean transition time (MTT)], and is schematically depicted in Figure 1. Observed absolute neutrophil counts were log-transformed prior to analysis to improve the symmetry of the residual error distribution. The structural model can be described using the following differential equations:

$$\frac{dPROL}{dt} = k_{prol} \times PROL \times (1 - E_{Eribulin}) \times \left(\frac{Circ_0}{Circ}\right)^\gamma - k_{tr} \times PROL \quad (1)$$

$$\frac{dTransit_1}{dt} = k_{tr} \times PROL - k_{tr} \times Transit_1 \quad (2)$$

$$\frac{dTransit_2}{dt} = k_{tr} \times Transit_1 - k_{tr} \times Transit_2 \quad (3)$$

$$\frac{dTransit_3}{dt} = k_{tr} \times Transit_2 - k_{tr} \times Transit_3 \quad (4)$$

$$\frac{dCirc}{dt} = k_{tr} \times Transit_3 - k_{circ} \times Circ \quad (5)$$

Here, the *Prol* compartment represents proliferative cells, the compartments *Transit*<sub>1-3</sub> represent transit compartments mimicking the cell maturation occurring with rate  $k_{tr}$ , and the *Circ* compartment represents the observed number of circulating neutrophils. The generation of new cells in the proliferating cells compartment is dependent on the following factors: (i) the number of cells in the compartment; (ii) a proliferation rate constant,  $k_{prol}$ , determining the rate of cell division; and (iii) a feedback mechanism from circulating cells ( $\gamma$ ). The drug concentration  $C$  in the central PK compartment is assumed to reduce the proliferation rate or induce cell loss. This is most commonly modelled using either a linear relationship (Eq. 6) or an  $E_{max}$  relationship (Eq. 7):

$$E_{Eribulin} = SLOPE \times C \quad (6)$$

$$E_{Eribulin} = \frac{E_{max} \times C}{EC_{50} + C} \quad (7)$$

Model discrimination was guided by the change in objective function value (OFV) between models, changes in interindividual variability (IIV) and residual unexplained variability (RUV), the magnitude of asymptotic relative standard errors (RSE), goodness of fit and residual error diagnostics. For hierarchical models, a difference in OFV of >7.88 ( $P < 0.005$ , d.f. = 1) was used to select the best model. The base model and final model were also evaluated using a visual predictive check. Parameter precision of the final model was evaluated using a nonparametric bootstrap analysis ( $n = 200$ ).

### Statistical model development

The IIV on structural population parameters was described using a log-normal distribution, as follows (Eq. 8):

$$P_i = P_{P,COV} \times \exp(\eta_i) \quad (8)$$

where  $P_i$  represents the individual parameter value,  $P_{P,COV}$  represents the covariate adjusted typical parameter value, and  $\eta_i$  is an independent random variable with a distribution of  $N(0, \omega^2)$ , where  $\omega$  represents the population variance for interindividual variability distribution estimated.

Residual error on ANC<sub>s</sub> was included as a proportional relationship, as follows:

$$ANC_{ij,pred} = ANC_{ij,obs} \times \exp(\varepsilon_{ij}) \quad (9)$$

where  $ANC_{ij,obs}$  represents the observed ANC for the  $i$ th individual and the  $j$ th observations,  $ANC_{ij,pred}$  represents the individual predicted ANC value (i.e. based on

equations 1–7) for the  $i$ th individual and the  $j$ th observations, and  $\varepsilon_{ij}$  represents the residual error distributed  $N(0, \Sigma^2)$ , where  $\Sigma^2$  represents the population variance for residual unexplained variability.

### Covariate model development

A covariate-screening step was performed for available covariates (Table 2) using plots of empirical Bayes estimates vs. covariates, also taking into account the clinical relevance and biological plausibility of parameter-covariate relationships, as identified in previous implementations of this model for other compounds. Covariates with an expected impact on model parameters were evaluated also when parameter-covariate plots were ambiguous, for instance, due to increased magnitudes of  $\eta$ -shrinkage. We then performed a univariate covariate analysis identifying potential covariates predictive of interindividual variability in model parameters. The clinical relevance of the parameters identified in the univariate analysis was supported by a simulation study, in which the incidences in grade 3 and grade 4 neutropenia were computed.

**Table 2**

Patient demographics of the pooled analysis data set

Description	Unit	Value
Sex (male/female)	<i>n</i>	229/1359
Ethnicity		
Caucasian	<i>n</i>	1253
Black/African American	<i>n</i>	83
Asian/Pacific Islander	<i>n</i>	23
Japanese	<i>n</i>	96
Other/unknown	<i>n</i>	133
Previous chemotherapy (yes/no)	<i>n</i>	1527/61
Previous Pt-containing chemotherapy (yes/no)	<i>n</i>	446/1142
Previous radiotherapy (yes/no)	<i>n</i>	1212/375
Previous hormonal therapy (yes/no)	<i>n</i>	923/761
Blood transfusions (yes/no)	<i>n</i>	161/1523
Received G-CSF (yes/no)	<i>n</i>	382/1302
Age [median (IQR)]	years	58.0 (49.0–66.0)
Bodyweight [median (IQR)]	kg	67.7* (59.0–77.6)
Body surface area [median (IQR)]	m <sup>2</sup>	1.73 (1.61–1.86)
Height [median (IQR)]	cm	162 (157–168)
Albumin [median (IQR)]	g dl <sup>-1</sup>	3.90† (3.6–4.27)
Alkaline phosphatase [median (IQR)]	IU l <sup>-1</sup>	118‡ (82.0–206)
Alanine transaminase [median (IQR)]	IU l <sup>-1</sup>	25.0 (17.0–39.6)
Aspartate transaminase [median (IQR)]	IU l <sup>-1</sup>	30.0 (22.0–46.0)
Bilirubin [median (IQR)]	mg dl <sup>-1</sup>	0.50§ (0.40–0.70)
Serum creatinine [median (IQR)]	mg dl <sup>-1</sup>	0.80 (0.68–0.91)
Lactate dehydrogenase [median (IQR)]	IU l <sup>-1</sup>	328¶ (211–486)
Platelets [median (IQR)]	10 <sup>9</sup> l <sup>-1</sup>	260 (209–326)
Protein [median (IQR)]	g dl <sup>-1</sup>	7.1 (6.7–7.5)

Abbreviations are as follows: G-CSF, granulocyte colony-stimulating factor; IQR, interquartile range; Pt, platinum-containing. Covariates were missing for <6% of subjects for the first or last observation over time per patient, and <13% for intermediate missing observations over time per patient. For <6% of subjects, all covariates were missing, when excluding lactate dehydrogenase (7.82%) and platelets (8.98%). The scaling value used in the covariate model deviated from the median and was set at \*70 kg, †4, ‡100, §2 and ¶238.

The effect of continuous covariates  $\text{cov}_{i \rightarrow n}$  and dichotomous covariates  $\text{cov}_{i \rightarrow m}$  on the population parameter  $P_p$  were included as follows (Eq. 10):

$$P_{p,\text{cov}} = P_p \times \prod_{i=1}^n \left( \frac{\text{cov}_i}{\text{median}(\text{cov}_i)} \right)^{\theta_i} \times \prod_{j=1}^m \theta_j^{\text{cov}_j} \quad (10)$$

where  $P_{p,\text{cov}}$  represents the covariates adjusted population parameter value,  $P_p$  represents the typical population parameter estimate, and  $\theta_{i \rightarrow n}$  and  $\theta_{m \rightarrow j}$  represent covariate effect parameters. Multilevel categorical covariates, such as race, were evaluated by estimating separate covariate effects for each category.

Covariates that showed a drop in OFV larger than 10.8 ( $P < 0.001$ ) when tested univariately, were added to the full model. Subsequently, covariates were deleted from the full model in a stepwise backward elimination procedure (again using an OFV difference of 10.8,  $P < 0.001$ ). A conservative  $P$  value of  $P < 0.001$  was used in order to take into account potential deviations from the nominal  $P$  value under the first-order estimation method.

### Evaluation of the impact of covariates on risk for neutropenia

In order to evaluate the clinical relevance of covariates identified as significant in the univariate analysis on  $\text{ANC}_0$ , and the covariates identified as significant in the final covariate model, simulations were performed by computing the incidences of grade 3 and 4 neutropenia for each of the parameter–covariate relationships separately. Simulations were conducted using the parameter estimates from the univariate runs. For each of these simulation scenarios, the parameter estimates obtained for each of the associated models were used.

Grade 3 toxicity was defined as a  $\text{ANC} < 1 \times 10^9 \text{ cells l}^{-1}$  for  $>7$  days. Grade 4 toxicity was defined as  $\text{ANC} < 0.5 \times 10^9 \text{ cells l}^{-1}$  for  $>7$  days. Patient cohorts were simulated based on the approved dosing regimen of  $1.4 \text{ mg m}^{-2}$  at day 1 and 8 for a 21 day treatment cycle. Body surface area (BSA) was simulated from the observed BSA distribution in the full data set, with a mean of  $1.57 \text{ m}^2$  and standard deviation of 0.22. The BSA was truncated for values within  $1\text{--}3 \text{ mg m}^{-2}$ . A patient cohort of 2000 patients was simulated for each scenario in order to obtain reliable 95% prediction intervals.

### Evaluation of the impact of covariates on dosing guidelines

In order to evaluate the impact of covariates identified on model parameters, we evaluated the dose adjustment necessary to match the nadir as predicted for a typical patient receiving a dose of  $1.4 \text{ mg m}^{-2}$ , for all relevant combinations of model parameters. This was done in the following steps: (i) definition of the final PK–PD model differential equations in the R-package deSolve; (ii) simulation of a

single ANC–time profile for a dose of  $1.4 \text{ mg m}^{-2}$  using the final covariate model PD parameters and a typical BSA of  $1.73 \text{ m}^2$ ; (iii) computation of typical ANC nadir value; and (iv) repeated simulations for different possible deviations in PD parameter values, optimizing dose to match the nadir value obtained in the typical patient, using the optimization function in R.

The resulting matrix of dose adjustments was then depicted graphically, and can be used for easy derivation of recommended dose adjustments, based on specific combinations of patient covariates.

In addition, we simulated ANC–time profiles using the recommended dose of  $1.4 \text{ mg m}^{-2}$  for the typical individual, for different combinations of deviations in parameter values. This was done to assess the expected impact of various covariate-induced changes in PD parameter values on ANC–time profiles.

### Evaluation of the relationship between exposure and risk for neutropenia

The simulation approach used to compute incidence of grade 3 and 4 neutropenia for separate univariate covariate models described above was also utilized to evaluate the relationship between eribulin exposure (in terms of area under the concentration–time curve, AUC), and the incidence of grade 3 and 4 neutropenia.

Simulations were performed using the base model for three dose levels ( $0.6\text{--}4 \text{ mg m}^{-2}$ ), with 200 patients per dose level in each data set. Different dose levels were used to obtain a wide range of AUC values. BSA values to calculate absolute dose amounts were simulated according to the same algorithm as described for the clinical evaluation of covariate relationships. For each simulated individual, it was determined whether grade 3 or 4 toxicity had occurred. Then the incidence of grade 3 and 4 neutropenia at different AUC levels was calculated for each data set.

## Results

### Base model development

*Model development* An overview of patient demographics in the final pooled data set is summarized in Table 2. The first-order estimation method was used because first-order conditional estimation with interaction was not computationally feasible. Both SLOPE and  $E_{\text{max}}$  models were evaluated. Parameter estimates for the SLOPE model had superior precision compared with the  $E_{\text{max}}$  model. Also, we identified a high correlation between the fixed effects of  $\text{EC}_{50}$  and  $E_{\text{max}}$ . Therefore, the SLOPE model was selected for further model development. Interindividual variability was estimated for all fixed effects. Estimation of interoccasion variability (IOV) and off-diagonal covariances in IIV were not computationally feasible. Residual unexplained variability was best described by a proportional relationship.

**Table 3**

Parameter estimates of the base model and final covariate model

Base model				
Parameter	Units	Estimate (RSE%)	$\eta$ -shrinkage (%)	
Baseline neutrophils (ANC <sub>0</sub> )	10 <sup>9</sup> l <sup>-1</sup>	3.97 (1.2)	–	
Mean transition time (MTT)	h	96.7 (5.9)	–	
Feedback ( $\gamma$ )	–	0.206 (2.8)	–	
Linear drug effect (SLOPE)	$\mu\text{g l}^{-1}$	0.0414 (4.4)	–	
Proportional residual error	CV%	49.7 (3.2)	–	
Between-subject variability				
Baseline neutrophils (ANC <sub>0</sub> )	CV%	37 (2.7)	12.3	
Mean transition time (MTT)	CV%	23.2 (20.1)	37.3	
Feedback ( $\gamma$ )	CV%	19.9 (25)	51	
Linear drug effect (SLOPE)	CV%	56.8 (3.7)	27.6	
Residual variability				
Proportional residual error	CV%	49.7 (3.2)	–	
Full covariate model				
Parameter	Units	Estimate (RSE%)	Bootstrap median (IQR)‡	$\eta$ -shrinkage (%)
Baseline neutrophils (ANC <sub>0</sub> )	10 <sup>9</sup> l <sup>-1</sup>	4.03 (1.2)	4.02 (3.99–4.06)	–
Mean transition time (MTT)	h	109 (1.8)	109 (107–110)	–
Albumin*		0.374 (23.6)	0.375 (0.342–0.412)	–
Bilirubin*		–0.046 (25.7)	–0.046 (–0.054 to –0.037)	–
Alkaline phosphatase*		–0.0337 (30.3)	–0.034 (–0.041 to –0.026)	–
Lactate dehydrogenase*		–0.0561 (20.5)	–0.056 (–0.063 to –0.048)	–
Received G-CSF†		0.883 (2.2)	0.884 (0.870–0.896)	–
Feedback ( $\gamma$ )	–	0.216 (1.4)	0.216 (0.214–0.218)	–
Linear drug effect (SLOPE)	$\mu\text{g l}^{-1}$	0.0451 (3.2)	0.045 (0.044–0.046)	–
Albumin*		0.763 (18.6)	0.771 (0.682–0.831)	–
Aspartate transaminase*		0.119 (24.4)	0.121 (0.099–0.139)	–
Received G-CSF†		1.3 (8.2)	1.310 (1.270–1.366)	–
Between-subject variability				
Baseline neutrophils (ANC <sub>0</sub> )	CV%	37.3 (2.6)	37.3 (36.6–37.9)	11.3
Mean transition time (MTT)	CV%	13.9 (6.7)	13.9 (13.3–14.4)	36.3
Feedback ( $\gamma$ )	CV%	12.2 (13)	12.2 (11.1–13.2)	54.7
Linear drug effect (SLOPE)	CV%	54 (4.3)	54.0 (52.5–55.6)	28.6
Residual variability				
Proportional residual error	CV%	49.6 (2.9)	49.5 (48.6–50.7)	–

\*Power covariate effect. †Proportional covariate effect. ‡Nonparametric bootstrap ( $n = 200$ ), median and interquartile range. Abbreviations are as follows: G-CSF, granulocyte colony-stimulating factor; IQR, interquartile range; IIV, interindividual variability (CV%); RSE, relative standard error (%), obtained from NONMEM covariance step; SLOPE, linear drug effect coefficient.

**Parameter estimates and model evaluation** The parameter estimates of the base model are given in Table 3. Relatively high interindividual variability (56.8 coefficient of variation % [CV%]) was found on the linear drug effect (SLOPE). All parameters were estimated with adequate precision (RSE < 25%). Shrinkage was minimal for baseline neutrophil count (ANC<sub>0</sub>) and SLOPE (<27.6 CV%), but was high for MTT and  $\gamma$  (37.3 and 51.0%, respectively). The magnitude of residual unexplained variability (49.6 CV%) was high. No relevant trend in observed vs. predicted values was observed.

### Covariate analysis

**Screening and univariate testing of covariates** Available patient demographics and laboratory values were used for a graphical evaluation of each covariate vs. individual

parameter values from the base model. All parameter-covariate relationships that were selected for univariate testing are summarized in Table 4.

**Parameter estimates and model evaluation** In order to keep the full model development feasible in terms of runtime, the full model was constructed containing all statistically significant covariates, except for covariates related to ANC<sub>0</sub>, because these were considered of less importance. Covariates that were statistically significant (Table 4) were added to a full covariate model, and subsequently a model reduction step was performed (all covariates were significant with dOFV > 32, i.e. highly statistically significant).

The final reduced covariate PK-PD model included the following parameter-covariate relationships:

**Table 4**

Univariate analysis results for parameters-covariate combinations evaluated

Parameter	Covariate	Type	dOFV*	Covariate effect	RSE (%)	Significance
ANC <sub>0</sub>	Prior chemotherapy	Categorical	-3.031	0.909	4.8	-
	Sex (male)	Categorical	-62.321	1.27	3.0	P < 0.001
	Blood transfusion	Categorical	-24.642	1.31	5.8	P < 0.001
	Pt-containing chemotherapy	Categorical	-18.924	1.11	2.6	P < 0.001
	Bodyweight	Continuous	-0.556	0.0358	145.5	-
	Albumin	Continuous	-119.266	-0.482	17.7	P < 0.001
MTT	Prior chemotherapy	Categorical	-9.573	0.863	NE	-
	Albumin	Continuous	-290.1	0.48	8.6	P < 0.001
	Bilirubin	Continuous	-145.685	-0.0997	21	P < 0.001
	AST	Continuous	-183.549	-0.0993	15.3	P < 0.001
	ALT	Continuous	-61.198	-0.0547	30.7	P < 0.001
	G-CSF	Categorical	-240.285	0.824	2.7	P < 0.001
	Lactate dehydrogenase	Continuous	-188.125	-0.111	15	P < 0.001
	ALP	Continuous	-118.353	-0.0779	18.5	P < 0.001
γ	G-CSF	Categorical	-25.72	1.1	3.7	P < 0.001
SLOPE	Asian	Categorical	-9.348	1.11	3.3	-
	Japanese	Categorical	-0.596	1.02	3.7	-
	Prior chemotherapy	Categorical	-25.31	1.47	12.6	P < 0.001
	Albumin	Continuous	-52.166	0.433	34.4	P < 0.001
	Sex (male)	Categorical	-21.761	0.84	5.3	P < 0.001
	Adjusted Child-Pugh scale	Categorical	-0.001	1	NE	-
	Bilirubin	Continuous	-0.64	-0.0142	215.5	-
	AST	Continuous	-84.325	0.136	24	P < 0.001
	Bodyweight	Continuous	-11.188	-0.183	41.5	P < 0.001
	G-CSF	Categorical	-58.501	1.27	6.3	P < 0.001

\*Change from OFV of base model (OFV = -9315.974). Abbreviations are as follows: ALT, alanine aminotransferase; AST, aspartate aminotransferase; dOFV, change in objective function value; G-CSF, granulocyte colony-stimulating factor; NE, not estimated; Pt, platinum-containing; RSE, relative standard error; SLOPE, linear drug effect coefficient.

MTT–albumin, MTT–bilirubin, MTT–ALP, MTT–lactate dehydrogenase (LDH), SLOPE–albumin, SLOPE–aspartate transaminase (AST) and SLOPE–granulocyte colony-stimulating factor (G-CSF).

When comparing the base model and the final covariate model, IIV was reduced for MTT from 23.2 to 13.9%, for feedback from 19.9 to 12.2%, and for SLOPE from 56.8 to 54.0%. The parameter estimates of the final covariate model are given in Table 3. All parameters were estimated with adequate precision of typical estimates (RSE < 30.3%). Shrinkage was between 11.3 and 54.7%. No relevant trends in observed vs. predicted neutrophil counts were observed. The visual predictive check described the observed data adequately (Figure 2).

Below are the equations providing the population predicted parameter values for MTT and SLOPE in the final model:

$$\begin{aligned}
 \text{MTT} = P_{\text{MTT}} \times \left(\frac{\text{Albumin}}{4}\right)^{0.374} \times \left(\frac{\text{Bilirubin}}{2}\right)^{-0.046} \times \\
 \left(\frac{\text{ALP}}{100}\right)^{-0.0337} \times \left(\frac{\text{LDH}}{238}\right)^{-0.0561} \times (0.883)^{\text{G-CSF}[0,1]}
 \end{aligned}
 \tag{11}$$

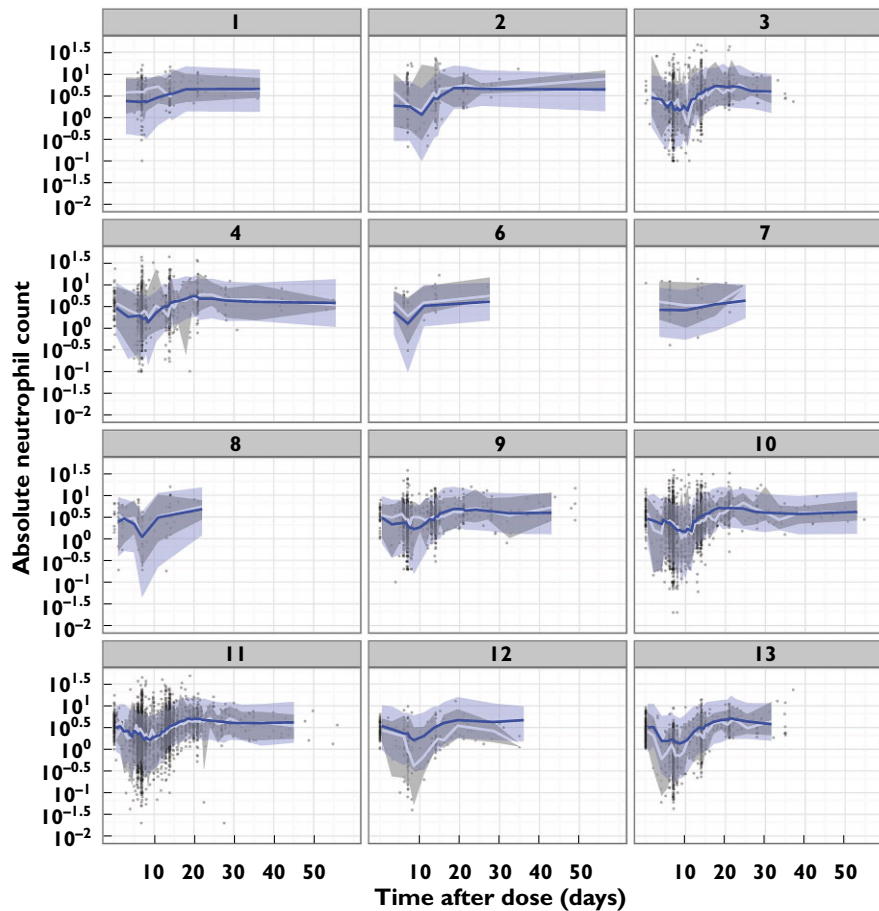
$$\begin{aligned}
 \text{SLOPE} = P_{\text{SLOPE}} \times \left(\frac{\text{Albumin}}{4}\right)^{0.763} \times \left(\frac{\text{AST}}{30}\right)^{0.119} \times (1.3)^{\text{G-CSF}[0,1]}
 \end{aligned}
 \tag{12}$$

where  $P_{\text{MTT}}$  and  $P_{\text{SLOPE}}$  represent the typical population parameter values, and the depicted biochemical parameter values have the same units as defined in Table 2.

### Evaluation of the impact of covariates on risk for neutropenia

Covariates that were identified as significant in the univariate covariate analysis for ANC<sub>0</sub> and in the final covariate model were evaluated for their clinical relevance using a simulation study that quantified the change in the incidence of grade 3 and 4 neutropenia (Table 5).

The most pronounced effect was that of albumin, a covariate for ANC<sub>0</sub>, MTT and SLOPE. If albumin is decreased by 50%, changes in ANC<sub>0</sub>, MTT and SLOPE will increase grade 4 incidence by 23.3, 39.1 and 18.95%, respectively. The increases seen for hepatic function markers (AST, ALP and bilirubin) were minimal, with <4.4% increase in grade 4 neutropenia. For prior platinum-containing chemotherapy, changes were also minimal at <2.15%. Males showed a decreased incidence of 4.8% in neutropenia compared with females. Patients who received blood transfusions experienced an expected decrease in grade 4 neutropenia of 5.65%. Finally, patients who received G-CSF treatment showed an increased grade 4 incidence (3.5%) with respect to SLOPE.



**Figure 2**

Visual predictive checks of absolute neutrophil count ( $\times 10^9$  cells  $l^{-1}$ ) vs. time after dose for the final covariate model, by study number. The blue line and area represent the model-predicted median and 90% prediction interval. The white line and grey area represent the observed median and 90% prediction interval

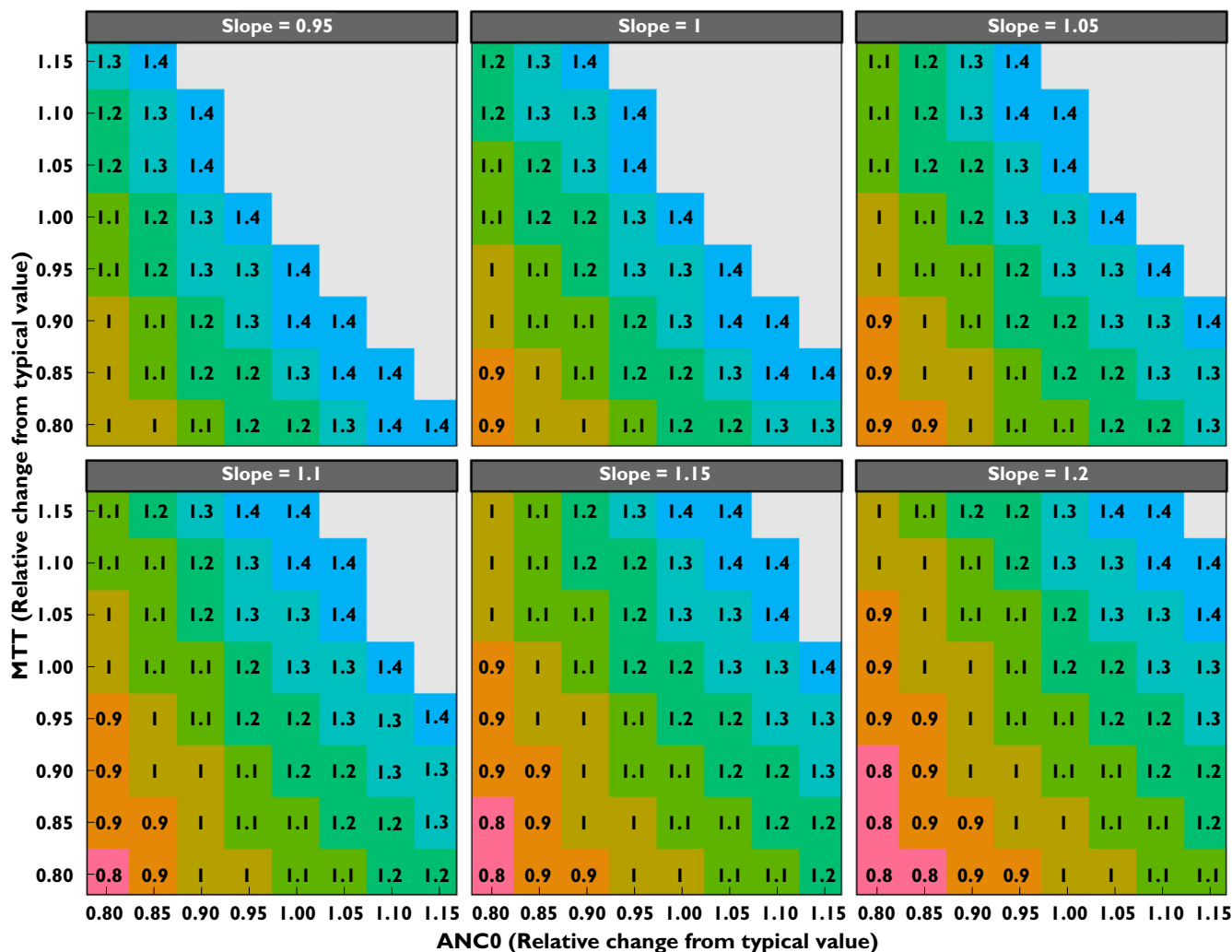
**Table 5**

Incidence of grade 3 and 4 neutropenia for different covariates, obtained using stochastic simulations ( $n = 2000$ ) using the registered eribulin dose of  $1.4 \text{ mg m}^{-2}$

Parameter	Covariate	Neutropenia incidence (%)		Change in neutropenia incidence* (%)	
		Grade 3	Grade 4	Grade 3	Grade 4
<b>Base model</b>		26.6	16.85	0	0
ANC <sub>0</sub>	Albumin 50% reduced	25.05	40.15	-1.55	23.3
	Received blood transfusion	19.95	11.20	-6.65	-5.65
	Received prior platinum-containing chemotherapy	25	14.70	-1.60	-2.15
	Sex (male)	21.6	12.05	-5.00	-4.80
MTT	Albumin 50% reduced	23.15	55.95	-3.45	39.10
	Alkaline phosphatase 50% reduced	25.65	13.55	-0.95	-3.30
	Bilirubin 50% reduced	25.45	13.15	-1.15	-3.70
	Received G-CSF	27.05	15.1	0.45	-1.75
SLOPE	Lactate dehydrogenase 50% reduced	27.3	16.50	0.70	-0.35
	Albumin 50% reduced	29.6	35.80	3.00	18.95
	Aspartate transaminase 50% reduced	24.65	12.45	-1.95	-4.4
	Received G-CSF	28.95	20.35	2.35	3.50

\*Compared with the base model incidences. Abbreviations are as follows: G-CSF, granulocyte colony-stimulating factor; SLOPE, linear drug effect coefficient.





**Figure 3**

Relative change in mean transition time (MTT) vs. baseline absolute neutrophil count ( $ANC_0$ ), stratified by relative change in slope parameter, depicting the change in dose (in milligrams per square metre) needed to match the nadir obtained in the typical patient (i.e. when all parameters are 1) receiving the registered eribulin mesilate dose of  $1.4 \text{ mg m}^{-2}$ . Doses higher than  $1.4 \text{ mg m}^{-2}$  (grey areas) are not depicted because these have not been evaluated in clinical studies and also do not take into account other toxicities

### Evaluation of the impact of covariates on dosing guidelines

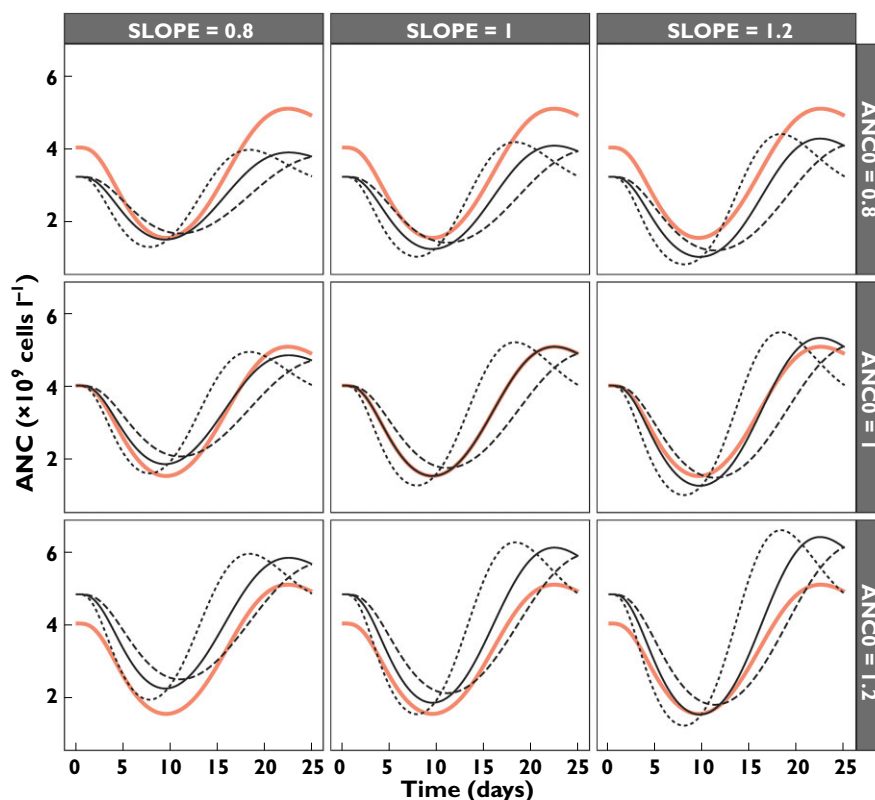
In Figure 3, dose adjustments are depicted for a range of relative deviations in relevant PD parameter values (SLOPE, MTT and  $ANC_0$ ), for which covariates were identified. Based on combinations of patient-specific covariate values, the expected deviations in change of model parameter values can be computed (e.g. using Eq. 10–12), and the predicted dose adjustment necessary to match the typical nadir value can be computed.

For instance, an individual with a  $ANC_0$  of  $3.2 \times 10^9 \text{ cells l}^{-1}$  (i.e. 20% reduction from typical value) and, based on a hypothetical patient laboratory values, an expected SLOPE of  $0.0496 \mu\text{g l}^{-1}$  (10% increase) and an expected MTT of 98.1 h (10% decrease), will lead to a recommended dose of  $0.9 \text{ mg m}^{-2}$ .

The quantitative relationship between the ANC–time profile for different parameter combinations is depicted in Figure 4.

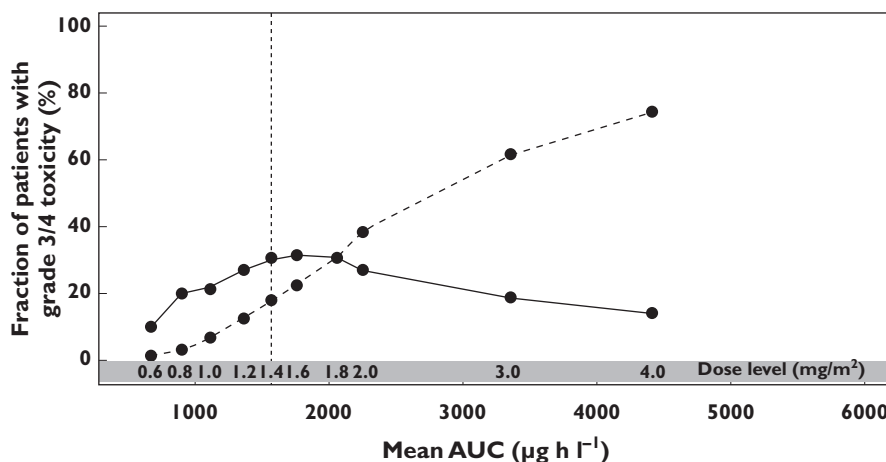
### Evaluation of the relationship between exposure and risk for neutropenia

In order to illustrate the direct relationship between exposure (i.e. AUC) and incidence of grade 3 or 4 neutropenia further, a simulation study was conducted using the base model (Figure 5). In this figure, the fraction of patients experiencing grade 3 or 4 neutropenia is depicted vs. exposure, demonstrating a clear exposure–response relationship. The incidence of grade 3 neutropenia gradually reaches a maximum near the approved dose levels of  $1.4 \text{ mg m}^{-2}$ , while grade 4 increases in a linear fashion for higher exposures.



**Figure 4**

Absolute neutrophil count (ANC) vs. time for the typical patient receiving  $1.4 \text{ mg m}^{-2}$  (continuous red line), stratified for different relative deviations from the typical parameter values for SLOPE,  $\text{ANC}_0$  and MTT (continuous line, MTT = 1; dotted line, MTT = 0.8; dashed line, MTT = 1.2)



**Figure 5**

Simulation of exposure (mean AUC) vs. fraction of grade 3 (continuous line) and grade 4 neutropenia (dashed line) incidence after 200 simulations for different dose levels

## Discussion

Eribulin mesilate is an important drug for the treatment of heavily pretreated patients with metastatic breast cancer, as has been shown in the EMBRACE trial [2]. There is large

variability in the clinical condition of this patient population considering characteristics such as organ function, pretreatment, disease burden, age and co-morbidity. It is, therefore, of importance to understand quantitatively the occurrence and determinants of neutropenia, the major

dose-limiting adverse effect of eribulin mesilate. In this population PK-PD analysis, data from patients from all available clinical studies ( $n = 1579$ ) that were available at the time of analysis were included, in order to obtain quantitative insight into the occurrence of neutropenia in different groups of patients. The time course for eribulin-associated neutropenia was adequately described using the developed PK-PD model and could provide insight into the exposure–response relationship (Figure 5). System-specific parameters were found to be consistent with values reported previously for other compounds [3, 5]. Patient characteristics predictive of increased sensitivity to develop neutropenia were identified, and their impact was illustrated using a simulation framework.

We first performed a univariate covariate analysis to identify potential covariates predictive of IIV in model parameters (Table 4). The clinical relevance of the parameters identified in the univariate analysis on  $ANC_0$ , and the statistically significant covariates in the final covariate model, was supported by a simulation study in which the incidences of grade 3 and grade 4 neutropenia were computed (Table 5). Full model development was performed, including all identified covariates except for any covariates identified on  $ANC_0$ . We excluded covariates on  $ANC_0$  in order to keep runtimes manageable during model development. Moreover, unlike other model parameters, the  $ANC_0$  is a parameter which is known in individual patients prior to start of therapy and hence is of less importance. The final full covariate model included the effect of the covariates albumin, bilirubin, ALP and LDH on the parameter MTT, and the effect of the covariates albumin, AST and G-CSF treatment on the drug effect parameter SLOPE.

During full model building, G-CSF effect on  $\gamma$  was removed, ALT, AST and G-CSF effects on MTT were removed, and effects of prior chemotherapy and sex on SLOPE were removed. The removal of G-CSF effect on  $\gamma$  and MTT could be related to confounding effects with the retained effect of G-CSF on SLOPE. The removal of ALT and AST effect on MTT could be related to confounding effects because of the inclusion of related covariates ALP and bilirubin. The removal of prior chemotherapy could be related to the generally impaired health status of patients who have received prior treatments, but in addition the number of patients without pretreatment (4%) was also low. The effect of sex on the incidence of neutropenia was also found in other analyses of conducted clinical trials, but was, however, not retained in the final model. The effect of sex on baseline was not included by choice, whereas the effect on SLOPE dropped out due to lack of significance, possibly related to other confounding factors. Nonetheless, for topotecan and docetaxel, an effect of sex on  $ANC_0$  has also been reported [13].

Kloft *et al.* [13] have investigated patient-related predictors of interindividual variability on model parameters of the same semi-physiological population model as was

used in the present analysis, for four different anticancer agents.

With respect to the identified covariates in our final model on MTT, Kloft *et al.* [13] also identified bilirubin as a covariate on MTT, but did not identify albumin, ALP and LDH as covariates related to MTT. In addition, Kloft *et al.* [13] did identify LDH to be a covariate on  $\gamma$ , i.e. confirming the relevance of this covariate to its relationship with interindividual variation in the neutrophil time course.

The analysis by Kloft *et al.* [13] did not investigate eribulin but rather four other anticancer agents, including two taxanes (e.g. also mitotic inhibitors). In our analysis, we identified albumin, AST and G-CSF as significant parameters for SLOPE. The relevance of albumin was also found by Kloft *et al.* [13], and G-CSF was not included in their analysis.

The present analysis focused on the identification of determinants of neutropenia. Therefore, a previously developed PK model including covariates predictive for PK was used to generate the PK profiles as input for the PD model. For most of the patients in the present data set, no PK data were available, and only typical covariate adjusted PK parameters were used for these patients. It should be stressed that the influence of the population PK model used did not directly influence the outcome of the PD analysis, because the only aim of the PK model was to provide predictions of eribulin concentration–time profiles.

Due to the long runtimes, only the first-order estimation method could be used, which is regarded as a sub-optimal estimation method. The employment of data splitting to reduce runtimes and enable external validation of the results was considered; however, as the aim of this analysis was mainly the identification of determinants of neutropenia, the use of data splitting would decrease the informativeness of the covariate and thereby the power of the analysis to identify relevant covariates.

Overall, the developed PK-PD model described the observed data adequately (Figure 2), and parameters could be estimated with good precision, which was confirmed by the bootstrap analysis conducted (Table 3). In addition, the system-specific parameters identified in the semi-physiological model for haematological toxicity were in agreement with previously published values for other drugs [3].

Our analysis indicates that the overall disease state and/or liver function-related covariates (e.g. albumin, bilirubin, ALP, LDH and AST) are the most relevant predictors for interindividual variability in the time course of neutropenia, which is consistent with covariates identified by others [13]. Eribulin is moderately protein bound (49–65%) [14].

Using the developed PK-PD model, quantitative investigation of optimal treatment strategies across different patient groups with respect to neutropenia was performed. However, as some of the identified covariate values

are expected to be highly correlated (such as albumin and LDH), it is not feasible to provide concrete dosing guidelines based on specific deviations in covariate values, because of the large range of possible combinations. We therefore used a more general approach, in which we computed dose adjustments needed based on the magnitude of deviation as can be easily computed for specific patients and their associated combination of covariate values (Figure 3).

Figure 3 does not provide doses higher than the registered eribulin mesilate dose of  $1.4 \text{ mg m}^{-2}$ , because no doses higher than  $1.4 \text{ mg m}^{-2}$  were evaluated in either phase II or phase III clinical trials. In addition, this analysis did not take into account any other dose-limiting toxicities that may occur. Any dose increments should therefore be considered only when initial doses appear to be well tolerated.

In the registration trial for eribulin [2] and the associated summary of product characteristics for eribulin, several dose reductions have been suggested after the occurrence of adverse events. The suggested initial dose adaptations in the present analysis can potentially reduce the number of such dose reductions that are only implemented after the occurrence of an adverse event using the registered dose of  $1.4 \text{ mg m}^{-2}$ .

In conclusion, we have successfully described the exposure–response relationship for eribulin-associated neutropenia in a large number of patients. Moreover, we have provided a general approach to support dose adaptations in case of complex covariate models. The developed PK-PD model could be used to guide dose optimization, also with respect to infusion duration and timing of dose administration, with respect to eribulin-induced neutropenia.

## Competing Interests

All authors have completed the Unified Competing Interest form at [http://www.icmje.org/coi\\_disclosure.pdf](http://www.icmje.org/coi_disclosure.pdf) (available on request from the corresponding author) and declare: all authors had no support from any organization for the submitted work; A.D.R.H., J.H.B. and J.H.M.S. had no financial relationships with any organizations that might have an interest in the submitted work in the previous 3 years; the PhD programme of J.G.C.v.H. was financially supported by Eisai Ltd; A.G. and Z.H. are paid employees of Eisai Ltd; no other relationships or activities that could appear to have influenced the submitted work.

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