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Older Adults' Trait Impressions of Faces Are Sensitive to Subtle Resemblance to Emotions

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Abstract

Younger adults (YA) attribute emotion-related traits to people whose neutral facial structure resembles an emotion (emotion overgeneralization). The fact that older adults (OA) show deficits in accurately labeling basic emotions suggests that they may be relatively insensitive to variations in the emotion resemblance of neutral expression faces that underlie emotion overgeneralization effects. On the other hand, the fact that OA, like YA, show a 'pop-out' effect for anger, more quickly locating an angry than a happy face in a neutral array, suggests that both age groups may be equally sensitive to emotion resemblance. We used computer modeling to assess the degree to which neutral faces objectively resembled emotions and assessed whether that resemblance predicted trait impressions. We found that both OA and YA showed anger and surprise overgeneralization in ratings of danger and naiveté, respectively, with no significant differences in the strength of the effects for the two age groups. These findings suggest that well-documented OA deficits on emotion recognition tasks may be more due to processing demands than to an insensitivity to the social affordances of emotion expressions.

Keywords

Aging; Face perception; Emotion resemblance; Overgeneralization; Trait impressions

Research investigating age-related changes in emotion recognition has yielded well-documented age-related declines in the ability to accurately label facial expressions of emotion. According to a meta-analysis of 17 studies, older adults (OA) are significantly worse than younger adults (YA) at labeling prototypical facial expressions (Ruffman et al. 2008). Nevertheless, OA do remain sensitive to emotion expressions, as they typically perform at above chance levels in labeling anger and other expressions, even when their accuracy is significantly lower than that of YA (Isaacowitz et al. 2007). Additionally, OA, like YA, consistently show a 'pop out' effect, whereby they are quicker to locate an angry schematic face among an array of neutral schematic faces than to locate a happy schematic face (Hahn et al. 2006; Mather and Knight 2006; Ruffman et al. 2009). Whereas it has been established that OA show deficits in labeling full-blown emotion expressions but retain sensitivity to intense anger expressions in an array of faces, research has not investigated age differences in responsiveness to more subtle emotion cues that have been shown to influence younger adults' trait impressions (e.g., Zebrowitz et al. 2010). The present study investigated

whether OA trait impressions are less responsive than those of YA to subtle variations in the resemblance of neutral expression faces to emotion expressions.

Our focus on trait impressions builds on research demonstrating that facial expressions of emotion not only communicate emotional states but also elicit trait impressions in YA, an effect that Secord (1958) dubbed *temporal extension*. For example, people displaying transient angry expressions are perceived to have stable traits associated with low warmth and high dominance (Knutson 1996; Montepare and Dobish 2003). In addition to perceiving people displaying emotion expressions to have traits congruent with those emotions, YA perceivers also show *emotion face overgeneralization* (Zebrowitz 1996, 1997), attributing emotion-related traits to people whose neutral expression facial structure resembles an emotion expression. For example, posed neutral expression faces that show more resemblance to an angry expression are perceived by YA as less likeable and trustworthy and more dominant, hostile, and threatening, with opposite impressions of neutral faces showing greater resemblance to a happy expression. These effects have been shown when resemblance to an angry expression is assessed by human raters (Montepare and Dobish 2003) or by objective computer methods (Said et al. 2009; Zebrowitz et al. 2010), as well as when resemblance is manipulated by lowering eyebrow height (Keating et al. 1981). The converse is also true, with neutral expression faces perceived as low or high in trustworthiness resembling angry or happy expressions, respectively (Oosterhof and Todorov 2008). According to the ecological theory of social perception (McArthur and Baron 1983; Zebrowitz et al. 2011), emotion overgeneralization is a byproduct of the adaptive value of responding appropriately to emotional expressions. For example, because the utility of responding appropriately to an angry person is highly important, falsely perceiving hostility when there is none present is less maladaptive than a failure to perceive it when warranted. Whereas it can be adaptive to respond appropriately to actual emotional expressions and to show false positives rather than miss an emotion, the adaptive value of emotion overgeneralization is uncertain, since it hinges on the question of whether people whose neutral expression faces resemble particular emotions actually have the associated traits. It is possible that some people do, but it is also likely that many do not. For example, people who look surprised due to large eyes are not necessarily naïve. In the absence of any information to indicate that the people posing neutral expression faces in the present study had personality traits that matched the emotion their faces resembled, we refer to the attribution of such traits as an emotion overgeneralization effect, not an accuracy effect.

In the present study, we examined whether OA trait impressions from faces would show anger, surprise, and happy emotion overgeneralization effects equivalent to those shown by YA. Given OA deficits in accurately labeling strong emotion expressions, one might expect OA to show little or no sensitivity to variations in the subtle resemblance of neutral expression faces to emotions. This would yield either no significant emotion overgeneralization or significantly weaker overgeneralization than that shown by YA, particularly in the case of anger overgeneralization for which OA show greater labeling deficits (mean effect size $r = .34$), than they do for surprise ($r = .07$) and happiness ($r = .08$) (Ruffman et al. 2008). On the other hand, the fact that OA do show some accuracy labeling emotions as well as sensitivity to anger in the emotion pop out paradigm suggests that they may show significant emotion overgeneralization.

Method

Participants

Sixteen young adult (YA) participants (8 men) and 16 older adult (OA) participants (8 men) completed the study. YA participants, aged 18–21 ($M = 18.8$, $SD = .91$), were recruited from a university and received course credit or \$15 payment. OA participants, aged 65–93,

($M = 75.6$, $SD = 7.3$), were recruited from the local community and were paid \$25. OA were screened for dementia using the Mini-Mental State Examination (Folstein et al. 1975), and all scored above 26 out of 30 ($M = 28.6$, $SD = 1.3$).

Control Measures

Measures of vision, affect, and cognitive function were administered to assess differences between our OA and YA samples and to determine whether any age differences in emotion overgeneralization could be attributed to differences in these more general processes. The results revealed several differences between OA and YA, most of which are consistent with previous aging research (see Table 2). OA performed worse than YA on the Reading the Mind in the Eyes Test (Eyes test; Baron-Cohen et al. 2001), which assessed the ability to read and label mental states from variations in the appearance of the eye region. OA also performed worse on the Benton Facial Recognition Test (Benton et al. 1983), which assessed the ability to match faces that were the same identity when presented with different orientations, expressions, and lighting conditions, although OA performance on the was in the normal range for normative samples of OA (e.g., Eslinger et al. 1985). In addition, visual acuity (Snellen Eye Chart) and contrast sensitivity (Mars Letter Contrast Sensitivity Test, Mars Perceptrix, Chappaqua, NY, USA), were lower for OA, but in the normal range for both age groups. There were no age differences in color vision (Ishihara's Tests for Color Deficiency, Ishihara 2010). Consistent with previous findings (Mather and Knight 2005, 2006), OA reported more positive affect than YA but did not differ in negative affect on a computerized version of the PANAS (Watson et al. 1988), which was administered because mood can affect emotion perception (Bouhuys et al. 1995). OA also scored higher than YA on the Shipley Vocabulary test (Shipley 1946), consistent with their higher education level and the maintenance of crystallized intelligence in older adulthood (Horn and Cattell 1967). The lower scores of OA than YA on a timed Pattern Comparison Task (Salthouse 1993) are consistent with decreases in processing speed in older adulthood (Salthouse 1996). Further, OA showed poorer executive control as evidenced by their worse performance on a short-form 48 item computerized version of the Wisconsin Card Sort Task to assess executive functioning (the Berg Card Sort Task (BCST; downloaded from <http://pebl.sourceforge.net/battery.html> and validated by Piper et al. 2012), although they showed equivalent levels of perseverative errors, which are the most diagnostic of executive dysfunction (Greve et al. 2005).

Stimuli

The images used in the present study were 120 White young adult neutral expression faces (60 men), taken from a multi-race set of 360 faces previously used by Zebrowitz et al. (2010). These facial images were selected from four different databases: the University of Stirling PICS database (<http://pics.psych.stir.ac.uk/>), the AR face database (Martinez and Benavente 1998), NimStim, and 2 yearbooks (one high school and one university yearbook). Images were selected to have neutral expressions, no head tilt, no eyeglasses, and no facial hair. The images were color frontal facial photographs of people displaying neutral expressions and were cropped to remove clothing and non-facial cues. Images were sized to a resolution of 310×380 pixels, and displayed at a resolution of $1,280 \times 1,024$ pixels on a 17 inch CRT monitor, at an approximate size of 7.6×10.2 cm, subtending a visual angle of 7.2 by 9.7° with participants seated 60 cm away. To verify that faces had neutral expressions, four YA judges (2 men) provided smile ratings on a 5-point scale, with endpoints labeled no smile/big smile. Eleven faces (3 male) were removed from our analysis because their smile ratings were >2 , yielding a total $N = 109$. Although it is theoretically possible that other emotional expressions may have been present in the stimuli, we believe that it is unlikely that someone asked to pose a neutral expression would instead pose an

angry or surprised expression; thus any resemblance of the faces to those expressions would result only from facial structure.

Measures

Emotion Resemblance—The degree to which each generalization face resembled anger, surprise, and happy facial expressions was taken from the Zebrowitz et al. (2010) study, which used connectionist modeling to provide objective indices of emotion resemblance for each face. In this method, a computer network is trained to differentiate two categories of faces (e.g., an angry and neutral expression) based on their facial metrics. When subsequently tested on metrics from the faces rated in this study, the network indicates the probability that each face belongs to one of the trained categories (e.g., the probability that it is an angry face). This probability is then used to predict human judges' impressions of the faces. More details about the modeling procedure are provided in Zebrowitz et al. (2007, 2010).

Trait and Appearance Ratings—Participants rated each generalization face on four traits (hostile, trustworthy, naïve, and warm) and two appearance qualities (attractive and babyfaced) using 7-point scales with endpoints labeled (1) *not at all* (hostile, naïve, trustworthy, warm, attractive, or babyfaced) and (7) *very* (hostile, naïve, trustworthy, warm, attractive, or babyfaced). Although ratings of 'naïve' were not included in the Zebrowitz et al. (2010) study, we included it to assess surprise expression overgeneralization, both because Zebrowitz et al. found no significant surprise overgeneralization with the traits they used and also because previous research found that prototypical surprise expressions were associated with impressions of greater naiveté (Zebrowitz et al. 2007). Ratings of "warm" also were not included in the study by Zebrowitz et al. (2010). Rather, they had assessed "likeable". Because they found no significant effects of resemblance to happy expressions on ratings of the likeability of the Caucasian faces used in the present study, we substituted the trait rating "warm", which we thought might be more sensitive to emotion overgeneralization than the more subjective rating of "likeable".

Procedure

After obtaining informed consent participants were seated in front of a computer screen. They first completed a computerized version of the PANAS (Watson et al. 1988). Next, MediaLab software (Empirisoft, New York City, NY) was used to present the trait rating task. Before rating the faces, participants viewed an instruction screen for the rating task where they were told to rate each face based only on their first impression of the face, and explicitly told the following: "Do not be concerned about whether your judgments are right or wrong. We are simply interested in whether people show agreement in their first impressions of faces." Prior to rating the 120 faces for this study, participants rated a different set of 24 male faces on four trait scales (aggressive, dominant, manipulative, and trustworthy) for a separate study. Following these ratings, participants rated the faces included in this study with the following trait order: trustworthy, hostile, naïve, and warm. After the trait ratings, participants rated the 24 male faces for the other study followed by the 120 faces for this study on the appearance qualities of attractiveness and babyfacedness, with the order of these two qualities counterbalanced across participants. Faces for the present study were blocked by sex and the order of male and female faces was counterbalanced across participants. This procedure was employed so that variations in trait impressions would be less likely to reflect sex stereotypes, as we were interested in whether the first impressions of OA were responsive to variations in emotion resemblance within demographic groups, as has been previously shown for YA (Zebrowitz et al. 2010). Participants viewed each face for 4 s and made their rating after the face disappeared. The rating scale remained on the screen until the rating was made, at which time it disappeared

and a new face appeared. The entire rating portion lasted about an hour. Following this, participants completed the BCST and a computerized version of the Eyes Test and then completed the other control measures as well as a demographic and health questionnaire.

Results

Inter-rater Agreement in Impressions

With one exception, Cronbach alphas revealed that the individual trait and appearance ratings of both OA and YA judges showed acceptable inter-rater reliabilities, indicating that OA, like YA, show consensual first impressions from faces (all alphas $>.7$; see Table 1). Given these inter-rater reliabilities, we computed mean ratings for each face across OA judges and across YA judges, and used face rather than rater as the unit of analysis, paralleling the analysis in the Zebrowitz et al. (2010) study. Also following the procedure of Zebrowitz et al. (2010), we created a danger composite by combining the hostile and reverse scored trustworthy ratings, which were highly correlated for both OA ($r(14) = .81, p < .0001$) and YA ($r(14) = .79, p < .0001$). The single exception to acceptable reliability was YA ratings of naïveté ($\alpha = .53$), which could not be raised by dropping any deviant participants. Despite this low alpha, we included YA naïve ratings in our analyses in order to compare effects with OA ratings. However, we recognize that finding weaker emotion overgeneralization effects for YA than OA impressions of naïveté could reflect YA lower reliability.

Emotion Overgeneralization Regression Analysis

Using face as the unit of analysis, we performed six separate regression analyses to predict OA and YA impressions (dangerous, naïve, warm) from resemblance to angry expressions, surprise expressions, and happy expressions as well as three other appearance variables. Specifically, to examine the effect of resemblance to emotion expressions on trait impressions with other appearance qualities controlled, we entered face sex (coded 0 for female, 1 for male), mean ratings of attractiveness and babyfacedness for each face by the relevant age group, along with the measures of how much each face resembled anger, happiness, and surprise, as determined by the connectionist model. We controlled other appearance qualities because each is correlated with emotion resemblance (Zebrowitz et al. 2007, 2010), and each has effects on trait impressions that parallel the predicted effects of emotion resemblance (Eagly et al. 1991; Kite et al. 2008; Montepare and Zebrowitz 1998). Therefore, controlling these appearance qualities ensures that differences in impressions of the faces reflects a response to their emotion resemblance, and not face sex, attractiveness, or babyfacedness. All variables were standardized and all regression coefficients reported below also are standardized. The overall R^2 values were significant for each of the six regressions: For ratings of danger, OA: $R^2 = .35, F(4,104) = 14.15, p < .001$; YA: $R^2 = .33, F(4,104) = 12.75, p < .001$; for ratings of warmth, OA: $R^2 = .43, F(4,104) = 19.47, p < .001$; YA: $R^2 = .35, F(4,104) = 13.74, p < .001$; and for ratings of naïveté, OA: $R^2 = .53, F(4,104) = 29.77, p < .001$; YA: $R^2 = .56, F(4,104) = 33.29, p < .001$.

Emotion Overgeneralization—Faces with greater resemblance to anger were rated higher in danger by both OA ($\beta_{OA} = .20, p = .045$) and YA ($\beta_{YA} = .22, p = .023$). They were also rated as less naïve by OA ($\beta_{OA} = -.14, p = .018$), whereas this effect was not significant for YA ($\beta_{YA} = -.06, p = .14$). Anger resemblance did not significantly predict warmth ratings for either OA ($\beta_{OA} = .03, p = .65$) or YA, although the latter showed a slight trend in the predicted direction ($\beta_{YA} = -.11, p = .13$). Faces with greater resemblance to surprise were rated higher in naïveté by both OA ($\beta_{OA} = .18, p < .001$) and YA ($\beta_{YA} = .10, p = .003$). Surprise resemblance did not predict ratings of warmth or danger (all β s $< .1$, all p s $> .25$). Resemblance to happiness was not a significant predictor of any ratings (all β s $< .07$, all p s

> .28), except for OA ratings of warmth, which showed a slight trend in the predicted direction ($r_{OA} = .09, p = .12$).

In order to examine age differences in the magnitude of the emotion overgeneralization coefficients, we performed planned comparisons, using an approach as described in Rosenthal and Rosnow (1984, p. 506) to examine whether two effects are significantly different from one another. This test involved converting effect sizes to Z-scores and testing whether the Z-scores of these two effects were significantly different. With one exception, these analysis revealed no significant age differences in emotion overgeneralization (all $t_s < 1.14, p_s > .3$). The exception was a marginally significant tendency for anger resemblance to have a stronger influence on YA than OA warmth ratings, $t(120) = 1.81, p = .072$. However, as noted above, resemblance to anger did not significantly influence warmth ratings for either age group, which calls for caution in interpreting this effect.

Other Appearance Qualities—We summarize here the effects of attractive and babyface ratings across the six regression analyses. Attractiveness predicted lower danger and higher warmth ratings for both OA and YA: danger: ($r_{OA} = -.59$ to $-.61$, all $p_s < .001$; $r_{YA} = -.54$ to $-.60$, all $p_s < .001$), warmth ($r_{OA} = .54$ to $.55$, all $p < .001$; $r_{YA} = .43$ to $.45$, all $p < .001$). Attractiveness also predicted lower ratings of naiveté by YA ($r_{YA} = -.10$ to $-.12$, all $p < .015$), with only nonsignificant trends for OA ratings ($r_{OA} = -.10$ to $-.13$, $p_s = .085-.14$). On the other hand, babyfacedness predicted higher naiveté ratings for both OA ($r_{OA} = .56$ to $.62$, all $p_s < .001$) and YA ($r_{YA} = .45$ to $.46$, all $p_s < .001$). Babyfacedness also predicted lower ratings of danger by YA ($r_{YA} = -.33$ to $-.40$, all $p_s < .002$) but not OA ($r_{OA} = -.10$ to $-.18$, $p_s = .10-.29$); and the same pattern was true for ratings of higher warmth ($r_{YA} = .25$ to $.28$, all $p < .002$; $r_{OA} = .02$ to $.03$, all $p_s > .5$).

Discussion

Our results reveal that the tendency for YA impressions of neutral expression faces to be influenced by their objective resemblance to emotion expressions extends to OA, with no significant age differences in the magnitude of these effects. Both OA and YA perceived greater danger in neutral expression faces that showed greater resemblance to anger, consistent with previous emotion overgeneralization results for YA (Zebrowitz et al. 2010). Both also perceived greater naiveté in faces that showed greater resemblance to surprise. In addition, OA perceived less naiveté in neutral expression faces that showed more resemblance to anger, consistent with evidence that people with highly-expressive angry expressions look less naïve (Zebrowitz et al. 2007). Finally, it is noteworthy that the effect of anger resemblance on impressions of danger and the effect of surprise resemblance on impressions of naiveté held true controlling for attractiveness and babyfacedness, which were associated with danger and naiveté ratings by OA and/or YA.

One may question whether these results should be construed as emotion overgeneralization or accuracy. Had we asked participants to tell us what emotion the faces resemble and their choices of anger, surprise, or happy matched the emotion identified by the computer modeling, then it would be appropriate to talk about accuracy. However our questions focused on trait impressions, which is why we frame our results in terms of emotion overgeneralization. For example, if participants rate faces that look subtly surprised as more naïve, we cannot call that accuracy unless we know that the people whose faces more resembled surprise are in fact more naïve. Nevertheless, it is true that participants' trait impressions reveal that they are detecting a subtle, objective resemblance of the faces to an emotion expression, and it is striking that OA do this to the same extent as YA.

Unlike resemblance to anger and surprise, resemblance to happy expressions did not influence the trait impressions of either YA or OA. This result may be due to the restriction of faces used in the present study to the Caucasian faces from Zebrowitz et al. (2010), who found that Caucasian judges showed happy emotion overgeneralization effects on impressions of the likeability of Black and Korean faces but not White faces. On the other hand, they showed stronger anger emotion overgeneralization effects on impressions of the danger of White than Black faces. The moderation of the overgeneralization effects by face race was attributed to the baseline level of impressions, with negative emotion resemblance affecting impressions of own-race faces that were rated relatively positively, and happy resemblance affecting impressions of other-race faces that were rated relatively negatively.

Our finding that OA trait impressions were sensitive to very subtle emotion information conveyed in neutral expression faces, with emotion overgeneralization effects equal to those shown by YA, stands in contrast to the well-documented OA deficits in labeling basic emotion expressions as well as more complex mental states and traits on the Eyes Test (e.g., Pardini and Nichelli 2009), which was replicated in the present study. This contrast suggests a dissociation in the processes that are engaged by traditional emotion labeling tasks and the Mind in the Eyes Test and those being tapped by our emotion overgeneralization task. The latter may engage processes more like those involved when angry faces in a neutral array “pop out” more than happy faces do for both OA and YA (Hahn et al. 2006; Mather and Knight 2006; Ruffman et al. 2009).

One possible explanation for the discrepancies between OA emotion labeling deficits and tasks that show intact emotional processing is that the labeling tasks involve controlled processing aimed at getting the ‘correct’ response, whereas the anger “pop out” effects and the emotion overgeneralization first impressions may both engage more automatic processing. Studies of YA neural activation during emotion labeling tasks support the argument that they involve controlled processing. Compared with passive viewing, which is an automatic processing task, emotion labeling yields a reduction in amygdala activation and an increase in prefrontal cortex activation, a signature of controlled processing (Hariri et al. 2000; Lange et al. 2003). Previous work has suggested that automatic, as compared to controlled, processing remains relatively intact in OA (Peters et al. 2007). As considerable research suggests the automatic nature of trait inferences (Bar et al. 2006; Dovidio et al. 1997; Rule and Ambady 2008; Willis and Todorov 2006), OA emotion overgeneralization may be due to intact automatic processing of subtle cues to emotion in their first impressions of faces. Although this provides a plausible explanation for our results, it would be useful to examine OA emotion overgeneralization using shorter stimulus presentations that more directly implicate automatic processing.

The distinction between automatic and controlled processing also may explain why previous research found that age differences in accurate emotion labeling explained age differences in the detection of deceit (Stanley and Blanchard-Fields 2008; Ruffman et al. 2011) and social gaffes (Halberstadt et al. 2011) despite our finding of no age differences in first impressions of neutral faces based on their resemblance to emotions. Like emotion labeling, detecting deception and gaffes may arguably involve greater controlled processing than our first impression task. Participants in the present study were explicitly told that there were no right or wrong answers and they should just give their first impression of each face. In contrast, participants in the study of social gaffe detection were asked to rate how appropriate behaviors seemed after being told that some were socially appropriate and others were socially inappropriate, and those in studies of deceit detection, were asked to rate how truthful statements seemed after being told that some were truthful and some were lies. These tasks make clear that there are correct and incorrect responses, which is likely to engage controlled processing. Indeed, detecting social gaffes involves determining what is

culturally appropriate and comparing behaviors to these expectations, while detecting deceit involves evaluating participants' stories as well as their behaviors to determine if they sound truthful.

It is also possible that stereotype threat undermines OA performance in emotion labeling and deceit and social gaffe detection. Although research on stereotype threat effects has focused on the tendency for OA memory to be adversely affected when tasks are presented in a way that raises performance anxiety due to beliefs about age-related declines in memory function (Chasteen et al. 2012), it is possible that perceptual processes like emotion, deceit, or social gaffe recognition are similarly diminished due to concerns about age-related declines in accuracy, which would not affect first impressions of faces when participants are explicitly told not to be concerned about whether their judgments are right or wrong. Although future research investigating these two possibilities would be worthwhile, our results clearly show that processing of subtle facial cues to emotion under conditions that do not engage performance anxiety is intact in OA.

Our evidence that OA are able to use threat-related information in their impressions of faces contrasts with previous evidence that OA show more positive impressions of faces than do YA, with OA rating faces as less threatening (Ruffman et al. 2006) as well as more trustworthy and less hostile (Zebrowitz et al. 2013). However, Zebrowitz et al. also found that OA and YA showed significant agreement with each other in their trait impressions, despite OA greater positivity. This suggests that OA and YA use some of the same processes in their trait impressions, and the results of the current study indicate that emotion overgeneralization is one of those processes.

OA performed worse than YA not only on the Reading the Mind in the Eyes Test, but also on several other control tasks, effects that replicated previous research findings as discussed above and demonstrated that our OA sample was typical of healthy community dwelling individuals. Yet, despite somewhat poorer face recognition ability, visual acuity, and contrast sensitivity, OA sensitivity to subtle facial resemblances to emotion was equal to that of YA. OA also reported more positive affect on the PANAS, a positivity effect that has been associated with OA greater avoidance of negative stimuli (Isaacowitz et al. 2006; Mather and Carstensen 2005). However, any such tendency did not impair their sensitivity to subtle facial cues to negative emotion, as OA were as likely as YA to form more negative impressions of faces that showed more resemblance to anger.

It should be noted that our failure to find any significant age differences cannot readily be attributed to low power to detect them, since using face as the unit of analysis ($N = 109$) provided reasonable power. Of course, it is possible that a larger sample of faces would allow us to detect small age differences in emotion overgeneralization. Specifically, we found nonsignificant trends suggesting that OA tended to associate happiness resemblance with greater warmth, and YA tended to associate anger resemblance with less warmth. It is possible that with a larger sample, these effects could become significant. However, the effects most likely to achieve significance with a larger sample of faces were the tendencies for OA to show stronger emotion overgeneralization for impressions of naiveté, which was likely due to lower reliability of YA ratings of this trait. It is also notable that using only younger faces provided a liberal test of age differences, since previous research has provided evidence for an own-age bias in emotion recognition (Malatesta et al. 1987), as well as other face perception tasks (Anastasi and Rhodes 2005; Bartlett and Fulton 1991; Voelkle et al. 2011). Despite the fact that own-age biases might weaken OA emotion overgeneralization when judging younger faces, they still showed effects equal to YA.

Conclusions

OA and YA showed equivalent emotion overgeneralization in trait impressions of neutral expression faces that varied subtly in their resemblance to anger or surprise, and these effects were independent of attractiveness, babyfacedness, and face sex. The anger emotion overgeneralization shown by OA suggests that any age-related reductions in the processing of negative material does not impair OA processing of subtle facial cues to negative emotions. More generally, our findings suggest that well-documented OA deficits on emotion recognition tasks may be due more to deficits in labeling expressions than to insensitivity to their social affordances, such as danger. This demonstrates the importance of the task context in assessing effects of aging, consistent with an ecological approach to aging (see Isaacowitz and Stanley 2011).

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Table 1
Alphas for YA and OA agreement in face ratings

	Age group	
	YA	OA
Danger	0.85	0.86
Trustworthy	0.81	0.78
Hostile	0.76	0.79
Naïve	0.56	0.73
Warm	0.82	0.84
Attractive	0.86	0.88
Babyfaced	0.83	0.85

Table 2

OA and YA scores on control measures

Measure	Older adults (N = 16)		Younger adults (N = 16)		t value	p value
	M	SD	M	SD		
Snellen visual acuity (denominator)	31.25	9.22	19.38	5.25	4.48	<.001
Mars letter contrast sensitivity	1.64	.11	1.74	.04	3.02	.005
Ishihara's test for color deficiency	13.88	.50	13.88	.34	0	1.00
Benton facial recognition test	45.50	3.74	47.63	3.54	1.65	.109
Pattern comparison test	29.06	5.56	40.13	6.70	5.08	<.001
Shipley vocabulary test	35.13	4.37	31.88	2.87	2.49	.019
PANAS positive affect	3.57	.52	2.77	0.82	3.24	.003
PANAS negative affect	1.47	0.34	1.65	.41	1.31	.20
Mind in the Eyes Test	21.38	4.62	26.31	5.39	2.78	.009
BCST correct responses	24.50	9.53	34.94	5.23	3.84	.001
BCST perseverative errors	7.06	5.72	5.81	.98	.86	.396
BCST non-perseverative errors	16.44	11.55	7.3	5.25	2.90	.007
BCST total errors	23.50	9.52	13.06	5.24	3.84	.001
BCST trials to complete first category	12.31	11.59	10.81	5.14	.473	.639
Level of education ^a	4.31	1.58	2.63	.50	4.07	<.001

^aLevel of Education was coded for highest level attained: (1) no high school diploma, (2) high school diploma, (3) some college, (4) Bachelor's degree, (5) some graduate work, (6) Masters degree, (7) Doctorate degree