REVIEW

Anterior versus posterior approach for treatment of thoracolumbar burst fractures: a meta-analysis

Gui Jun Xu · Zhi Jun Li · Jian Xiong Ma · Tao Zhang · Xin Fu · Xin Long Ma

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Abstract

Purpose To critically review and summarize the literature comparing the results of surgery via an anterior approach and that via a posterior approach for the treatment of thoracolumbar burst fractures to identify the better approach.

Methods In this meta-analysis, we conducted electronic searches of MEDLINE, EMBASE, the Cochrane Central Register of Controlled Trials and other databases using the search terms "thoracolumbar fractures", "anterior", "posterior", "controlled clinical trials". Relevant journals or conference proceedings were also searched manually. Data extraction and quality assessment were in accordance with Cochrane Collaboration guidelines. The analysis was performed on individual patient data from all the trials that met the selection criteria. Sensitivity analysis was performed when there was significant heterogeneity. Results were expressed as risk difference for dichotomous outcomes with 95 % confidence interval.

Results Four randomized clinical trials and three controlled clinical trials comparing the results of the anterior versus posterior approach in the treatment of

G. J. Xu and Z. J. Li equally contributed to the study.

X. Fu and X. L. Ma equally contributed to the corresponding author.

X. Fu · X. L. Ma

Department of Orthopedics, Tianjin Hospital, 406, Jiefang Nan Street, Hexi District, Tianjin 300211, China

thoracolumbar burst fractures were retrieved; these studies included 179 and 152 patients in the anterior and posterior approach groups, respectively. There were no differences in terms of neurological recovery, return to work, complications and Cobb angle between the two groups. The anterior approach was associated with longer operative time, greater blood loss and higher cost than the posterior approach.

Conclusions The posterior approach may be more effective than the anterior approach. However, more highquality, randomized controlled trials are required to compare these approaches and guide clinical decision-making. *Level of Evidence* Level II, therapeutic study. See the Guidelines for Authors for a complete description of level of evidence.

Keywords Anterior approach · Posterior approach · Thoracolumbar burst fractures · Meta-analysis

Introduction

The majority of spine fractures occur in the thoracolumbar region, with burst fractures accounting for 10–20 % of cases, presumably because in this region, the relatively immobile thoracic spine connects with the relatively mobile lumbar spine [1, 2]. The burst fracture is an injury characterized by anterior vertebral body height loss, fracture of the posterior wall or retropulsion bone fragments into the spinal canal [3]. The clinical features of burst fractures include acute back pain, restricted motion and neurological deficits such as motor or sensory changes and sphincter disturbances. Radiographic signs of vertebral instability include widening of the interspinous and interlaminar distances, translation of more than 2 mm, kyphosis

G. J. Xu · Z. J. Li · J. X. Ma · T. Zhang · X. L. Ma (⊠) Department of Orthopedics, Tianjin Medical University General Hospital, 154, Anshan Street, Heping District, Tianjin 300052, China e-mail: spiritseeker2008@gmail.com

of more than 20° , dislocation, height loss of more than 50 % and articular process fractures [4, 5].

Although many studies have obtained good results with nonoperative treatment of thoracolumbar burst fractures [6– 8], most authors agree that surgical treatment is required for symptomatic, unstable burst fractures [9, 10]. Surgical intervention can decompress neural elements, restore vertebral body height, correct angular deformity and stabilize the spine. Stabilization of these injuries has many advantages such as early mobilization and the potential for neurological improvement. However, controversies exist over the specific approach to be used for surgical treatment. Surgeons today have the option of either an anterior or a posterior approach. Some authors recommend posterior instrumentation with or without decompression because of the excellent results obtained in terms of spinal stability, anatomical alignment, postoperative neurological improvement and low patient morbidity [11–13]. In contrast, advocates of the anterior approach cite predictable decompression of the spinal canal, improvement in postoperative neurological function and no significant increase in surgical morbidity [14–16]. Both the anterior and posterior approaches are associated with satisfactory results as well as some complications. Treatment of thoracolumbar injury depends on many clinical factors such as patient age, extent of spinal canal compromise, sagittal index, anterior body height, degree of integrity of posterior elements and presence of neurological deficit [17]. All these parameters affect the decision regarding the surgical approach, with each parameter being an indication for one or the other approach. Therefore, no consensus has been reached about the ideal treatment approach [18, 19].

Studies that have compared the outcomes of thoracolumbar burst fractures treated via an anterior or a posterior approach have reported ambiguous results. Moreover, some doctors prefer a specific approach for a given fracture. Therefore, the objective of this study was to systematically review relevant randomized controlled trials to clarify the differences in these two approaches.

Materials and methods

Search strategy

Electronic searches of MEDLINE (1950–present), EM-BASE (1980–present), the Cochrane Central Register of Controlled Trials (most recent edition) and other internet databases were performed to identify trials according to the Cochrane Collaboration guidelines. We used the following search terms and different combinations of MeSH (Medical Subject Heading) terms and textual words: "thoracolumbar fractures", "anterior", "posterior", "controlled clinical trials". Manual searches, including those of reference lists of all included studies, were used to identify trials that the electronic search may have failed to identify. Two reviewers (Xu and Fu) independently assessed the titles and abstracts of all reports identified by the electronic and manual searches. There was no restriction on language. When inclusion was unclear based on abstracts, full-text articles were retrieved. Any disagreements were resolved through discussion.

Selection criteria and quality assessment

Trials with the following characteristics were included: (1) randomized, quasi-randomized or controlled clinical trials, (2) patients without confirmed pathological thoracolumbar burst fractures based on computed tomography (CT) and plain radiographs, (3) comparison of the anterior and posterior approaches of surgical management and (4) full-text articles. We excluded articles that were duplicate reports of earlier trials or post hoc analyses of randomized controlled trial data and articles whose full text we were unable to acquire. To assess the methodological quality of the included studies, we used a modification of the generic evaluation tool used by the Cochrane Bone, Joint and Muscle Trauma Group [20]. The methodological quality of each trial was scored from 0 to 24. Disagreements were resolved by consensus or consultation with the senior reviewer.

Data extraction

Two authors independently extracted data from the included articles. Information regarding the study design, patient demographics, inclusion and exclusion criteria, interventions, outcomes, follow-up duration and rate of lost to follow-up duration for each treatment group was extracted. Data were managed using the Review Manager (RevMan) 5.1 software (The Nordic Cochrane Centre, The Cochrane Collaboration, Copenhagen, Denmark). We attempted to contact authors for supplementary information when the reported data were inadequate.

Data analysis and statistical methods

The meta-analysis was undertaken using RevMan 5.1 for Windows (Cochrane Collaboration, Oxford, UK). We assessed the statistical heterogeneity of each study by using a standard Chi square test (statistical heterogeneity was considered significant at P > 0.1) and the I^2 statistic [21]. An I^2 value of 50 % or higher was considered to indicate substantial heterogeneity. When heterogeneity existed, pooled data were meta-analyzed using a random-effects model [22]. Otherwise, a fixed-effects model was used for the analysis. Risk difference (RD) and 95 % confidence interval (CI) were calculated for dichotomous outcomes,

while mean difference (MD) and 95 % CI were used for continuous outcomes.

Results

cases

Study characteristics

Figure 1 shows a flow chart of the study selection and inclusion process. The search strategy identified 679 citations, of which seven satisfied the pre-defined inclusion criteria for data extraction and meta-analysis [23-29] (Table 1). These studies involved 179 patients in the anterior approach group and 152 patients in the posterior approach group. The majority of the included trials were small studies with between 25 and 63 participants. Their quality-assessment scores ranged from 10 to 22. Three studies had a score of <14. The patients' characteristics were comparable within each study group. Individual patient data were available from these articles. This did not include data for those lost to follow-up. There were no between-group differences in gender, mechanism of injury, level of fracture or length of follow-up.

Results of data analysis

Radiographic evaluation

The Cobb angle was reported in all studies, except for the study by Stancic et al. [25]. Wood et al. [26] and Esses et al. [23] compared canal encroachment. At the final follow-up, canal encroachment had improved to a greater extent in the anterior approach group than in the posterior approach group (MD, -8.2; 95 % CI -17.33 to 0.92; P = 0.08; Fig. 2). However, no between-group difference was found in the Cobb angle (MD, -0.06; 95 % CI -0.62to 0.51; P = 0.84; Fig. 3).

Frankel score

Although five studies evaluated neurological recovery, we could obtain data from only four studies, in which a total of 194 patients (107 in the anterior approach group and 87 in the posterior approach group) were examined for neurological function using Frankel scores. We could not use the information from the study by Lin et al. [29], because we



Table 1 Characteristics of the included studies

Study	Group	Number	Age	Gender		Fractur	e level				Lost	Follow-up duration	QAS	
			(years)	Male	Female	~T10	T11	T12	L1	L2	L3~		(months)	
Lin et al. [29]	Anterior	32	38	14	18	0	3	10	14	5	0	0	45 (24–72)	22
	Posterior	32	39	16	16	0	2	9	17	4	0	0		
Stancic [25]	Anterior	13	36	7	6	0	0	1	10	2	0	0	>12	20
	Posterior	12	35	8	4	0	0	2	9	2		0		
Sasso et al. [28]	Anterior	40	37	29	11	0	38				2	3	31 (9-89)	13
	Posterior	13	34	10	3	0	0	4	6	2	1	0		
Wood et al.	Anterior	22	38	12	8	0	1	1	10	8	0	2	44 (24–108)	21
[26]	Posterior	21	42	13	5	0	0	4	11	3	0	3		
Esses [23]	Anterior	18	34	25	15	1	0	5	22	9	3	2	20 (12-34)	19
	Posterior	22										0		
Hitchon et al.	Anterior	38	42	26	12	0	1	13	18	6	0	0	33 (6–96)	10
[27]	Posterior	25	42	19	6	0	1	3	12	9	0	0		
Danisa et al.	Anterior	16	35	11	5	0	0	1	13	2	0	0	27 (6-54)	12
[24]	Posterior	27	38	19	8	0	0	8	16	3	0	0		



Fig. 2 Forest plot showing canal encroachment in the two groups

	а	nterior	posterior					Mean Difference	Mean Difference					
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Fixed, 95% C		IV, F	ixed, 95	% CI		
Danisa OA 1995	9.8	9.8	16	9.5	6.8	27	1.1%	0.30 [-5.14, 5.74]		-	_	-		
Esses SI 1990	6.7	8	16	9.6	6.8	22	1.4%	-2.90 [-7.74, 1.94]						
Hitchon PW 2006	4.5	9.3	38	9.8	9.4	25	1.4%	-5.30 [-10.02, -0.58]						
Lin B 2011	6.23	1.22	32	6.14	1.16	32	94.6%	0.09 [-0.49, 0.67]						
Sasso RC 2006	9.2	8.67	37	11.6	10.64	13	0.8%	-2.40 [-8.82, 4.02]						
Wood KB 2005	11.06	12.59	20	12.5	8.06	18	0.7%	-1.44 [-8.10, 5.22]			-	-		
Total (95% CI)			159			137	100.0%	-0.06 [-0.62, 0.51]			•			
Heterogeneity: Chi ² =	6.99, df =	= 5 (P =	0.22);	l² = 29%	5				H	10		10		
Test for overall effect:						-20	-10 ante	rior pos	terior	20				

Fig. 3 Forest plot showing the Cobb angle in the two groups

used the improvement in Frankel scores (by one or more grades) to calculate the effects of the two surgical approaches. No significant differences were detected between the anterior and posterior approach groups (RD, 0.06; 95 % CI -0.07 to 0.19; P = 0.38; Fig. 4). In addition, Lin et al. [29] did not find a significant difference between the anterior and posterior approach groups.

Return to work

A total of 163 patients (84 in the anterior approach group and 79 in the posterior approach group) returned to work (Fig. 5). There was no significant between-group difference (RD, -0.02; 95 % CI -0.16 to 0.13; P = 0.8) in the number of patients who returned to work.

	anteri	or	poster	ior		Risk Difference	Risk Difference
Study or Subgroup	Events	Total	Events Total Weight			M-H, Fixed, 95% C	M-H, Fixed, 95% Cl
Danisa OA 1995	3	16	6	27	22.8%	-0.03 [-0.28, 0.21]	
Esses SI 1990	4	16	3	22	21.0%	0.11 [-0.14, 0.37]	- +
Hitchon PW 2006	13	38	11	25	34.3%	-0.10 [-0.34, 0.15]	
Sasso RC 2006	30	37	6	13	21.9%	0.35 [0.05, 0.65]	
Total (95% Cl)		107		87	100.0%	0.06 [-0.07, 0.19]	•
Total events	50		26				
Heterogeneity: Chi ² = 5	.91, df =	3 (P = 0	0.12); l ² =	49%			
Test for overall effect: $Z = 0.87$ (P = 0.38)							anterior posterior

Fig. 4 Forest plot showing Frankel scores in the two groups

	anteri	or	poster	ior	Risk Difference			Risk	nce		
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% C	I	М-Н,	5% CI		
Danisa OA 1995	11	16	16	27	25.5%	0.09 [-0.20, 0.39]					
Hitchon PW 2006	18	38	16	25	38.3%	-0.17 [-0.41, 0.08]					
Stancic MF	8	10	8	9	12.0%	-0.09 [-0.41, 0.23]					
Wood KB 2005	16	20	12	18	24.1%	0.13 [-0.15, 0.41]			-		
Total (95% CI)		84		79	100.0%	-0.02 [-0.16, 0.13]			•		
Total events	53		52								
Heterogeneity: Chi ² = 3	3.28, df =	3 (P = 0	0.35); l² =	8%			-				
Test for overall effect:		-1	-0.5 anter	u ior poa	0.5 aterior	I					

Fig. 5 Forest plot showing the number of patients who returned to work in the two groups

Operative data and cost

The differences in blood loss, operative time and cost between the two groups are shown in Figs. 6, 7, 8. Data on blood loss were available from three trials with a total of 145 patients (68 in the anterior approach group and 77 in the posterior approach group). There was significant heterogeneity ($I^2 = 82$ %) among these studies, and analysis with the random-effects model revealed significant between-group differences (MD, 310.67; 95 % CI 26.38–594.96; P = 0.03). Operative time was reported in three studies with 165 patients, and it significantly differed between the two surgical approaches (MD, 17.17; 95 % CI 5.78–28.57; P = 0.003). Only two studies reported the surgical cost for the two groups; the cost was higher in the anterior approach group than in the posterior approach group (MD, 15,530; 95 % CI 6,720-24,340; P = 0.0006).

Complications

The incidence of complications was documented in all studies with a total of 331 patients. There was significant

heterogeneity ($I^2 = 97 \%$) among the studies; however, this disappeared after the exclusion of two studies (Lin et al. [29] and Wood et al. [26]; $I^2 = 3 \%$, Fig. 9). We used the random-effects model for the analysis, and found no significant between-group difference in the incidence of complications (RD, -0.06; 95 % CI -0.15 to 0.02; P = 0.13). Among the reported complications, four cases of deep wound infections were recorded in three studies [24, 26, 27], and all four cases occurred in the posterior approach group.

Discussion

Thoracolumbar burst fractures can be treated via different approaches: anterior, posterior or a combination of these two approaches in some special cases. Theoretically, the anterior approach offers some benefits such as better canal decompression [16, 30]. This approach provides better exposure of the fractured vertebrae, enabling a more thorough decompression [10]. In contrast, the posterior approach can only support indirect decompression [31]. Therefore, good canal remodeling after canal encroachment occurred in the

	anterior posterio			osterior			Mean Difference	Mean Difference					
Study or Subgroup	Mean SD Total Mea			Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Rando				
Danisa OA 1995	1,878	777	16	1,103	793	27	19.5%	775.00 [290.83, 1259.17]			→		
Lin B 2011	811.6	175.3	32	720	179.9	32	43.3%	91.60 [4.57, 178.63]					
Wood KB 2005	783.5	370.4	20	461.8	237.8	18	37.2%	321.70 [125.69, 517.71]					
Total (95% CI)			68			77	100.0%	310.67 [26.38, 594.96]					
Heterogeneity: Tau ² = Test for overall effect:	-1000 -500 (anterior	D 500 posterior	1000										

Fig. 6 Forest plot showing blood loss in the two groups

	anterior			ро	sterio	r		Mean Difference	Mean Difference				
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Fixed, 95% C		IV, F	ixed, 95	% CI	
Hitchon PW 2006	415	118	38	413	81	25	5.4%	2.00 [-47.15, 51.15]					
Lin B 2011	172.5	28.3	32	157.3	25.9	32	73.5%	15.20 [1.91, 28.49]			-	_	
Wood KB 2005	232.9	35.9	20	205	41.5	18	21.1%	27.90 [3.10, 52.70]				-	
Total (95% CI)			90			75	100.0%	17.17 [5.78, 28.57]			•		
Heterogeneity: Chi ² =	1.17, df	= 2 (P	= 0.56)); l² = 0%	%				100			50	100
Test for overall effect: $Z = 2.95$ (P = 0.003)									-100	ante	rior pos	terior	100

Fig. 7 Forest plot showing operative time in the two groups

	а	nterior		р	osterior			Mean Difference		Mean Difference				
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Fixed, 95% C	I	IV, I	Fixed, 95	% CI		
Danisa OA 1995	63.963	18.203	16	45.306	15.808	27	67.4%	18.66 [7.93, 29.39]			-	-		
Hitchon PW 2006	89.09	27.34	38	80.04	32.535	25	32.6%	9.05 [-6.38, 24.48]						
Total (95% CI)			54			52	100.0%	15.53 [6.72, 24.34]						
Heterogeneity: Chi ² = 1.00, df = 1 (P = 0.32); $ ^2 = 0\%$ Test for overall effect: Z = 3.45 (P = 0.0006)									-50	-25	0	25	 50	
										ante	mor pos	lenor		

Fig. 8 Forest plot showing surgical cost in the two groups

anterior			poater	ior		Risk Difference		Risk Difference				
Study or Subgroup	Events Total		Events	Total	Weight	M-H, Fixed, 95% C	l	М-Н,	Fixed, 9	95% CI		
Danisa OA 1995	0	16	4	27	19.7%	-0.15 [-0.31, 0.01]		-	╼┤			
Esses SI 1990	2	18	2	22	19.5%	0.02 [-0.17, 0.21]			-			
Hitchon PW 2006	2	38	5	25	29.6%	-0.15 [-0.32, 0.02]		-				
Sasso RC 2006	2	37	1	13	18.9%	-0.02 [-0.19, 0.14]			-			
Stancic MF	2	13	1	12	12.3%	0.07 [-0.18, 0.32]						
Total (95% CI)		122		99	100.0%	-0.06 [-0.15, 0.02]						
Total events	8		13									
Heterogeneity: Chi ² =	4.11, df =	4 (P = 0	0.39); l² =	3%			-		<u> </u>		_	
Test for overall effect:	Z = 1.52 (P = 0.1	3)				-1	-0.5 ante	v rior po:	0.5 sterior	1	

Fig. 9 Forest plot showing complications in the two groups

anterior approach group just as all the included studies concerning this parameter. Our meta-analysis showed that canal remodeling was better in the anterior approach group than in the posterior approach group at the final follow-up. Such remodeling has also been reported in other articles [32–34].

Interestingly, although some authors have reported that compared to the posterior approach, the anterior approach provides better decompression of thoracolumbar fractures, which facilitates neurological recovery [35], others have found no significant differences [26, 28]. The relationship between the extent of canal encroachment and neurological function has been studied; although concerns have been raised regarding inadequate spinal canal remodeling after treatment via the posterior approach, there is no evident association between the percentage of canal encroachment and clinical symptoms [36, 37].

Generally, the most important purpose of the surgical management of thoracolumbar fractures is to minimize the change in the patients' lives. Our review showed that although the anterior approach was associated with better canal remodeling, it was not associated with a greater improvement in Frankel scores or a higher incidence of return to work. This is similar to reports that recovery of neurological function did not depend on the extent of spinal decompression and canal encroachment [37, 38]. Among the seven included studies, only Wood et al. [26] reported pain scores on a 10-point visual analog pain scale; the results showed that there was no significant difference in pain reduction between the two groups.

On critical evaluation of the included trials, we found that Lin et al. [29] reported many more complications in the anterior approach group than in the posterior group, including twenty-seven cases of hemopneumothorax, two cases of respiratory tract infection, three cases of intercostal neuralgia and thirteen cases of abdominal distension and constipation. In the study conducted by Wood et al. [26], there were seventeen "events" in the posterior approach group, including six cases of instrument removal, two cases of wound dehiscence, two cases of instrumentation/bone failure, two urinary tract infections, two cases of instrument breakage, one deep wound infection, one case of pseudarthrosis and one case of seroma. Because these two studies markedly differed from the other included studies, sensitivity analysis after the removal of these studies showed a significant reduction in the heterogeneity among studies, from I^2 97–3 %. After the exclusion of these two studies, our results showed that the incidence of complications did not differ between the two approaches. Deep wound infection is a catastrophic complication that necessitates implant removal. Only three studies reported this complication, and all four of the reported cases were in the posterior approach group, which suggests that deep wound infection is more common after surgery via a posterior approach. Unfortunately, we do not know why the smaller incision of the posterior approach was associated with a higher incidence of deep wound infection.

In addition, we found significant differences in operative time, blood loss and cost between the two groups. The anterior approach group was associated with longer operative times, greater blood loss and higher costs; thus, use of the posterior approach could potentially decrease the risks associated with long operative times and greater blood loss and transfusion.

We acknowledge some limitations of the literature and our review. First, although several relevant trials have been published, the majority were small and of low quality. Few comparative trials satisfied our inclusion criteria, including three articles with quality-assessment scores of below 14. As a result, we could not perform subgroup analysis because the information required was unavailable. Second, the heterogeneity of the study populations in terms of complications and therapeutic options posed additional challenges in evaluating the individual therapeutic options. This clinical heterogeneity, combined with the small sample sizes of the included studies, resulted in high I^2 values for our pooled results for complications, blood loss and Cobb angle. Third, the use of variable outcome measures and suboptimal reporting, often at nonstandardized intervals, further undermined informed decision-making. Lastly, we must mention that the selection of the surgical approach for the treatment of thoracolumbar fractures should be individualized because many factors influence this choice.

Conclusion

The results of this review showed that in the surgical management of thoracolumbar burst fractures, the anterior approach was not significantly superior to the posterior approach in terms of recovery of neurological function and return to work and that it was disadvantageous in terms of operative time, blood loss and cost. To some extent, the posterior approach was better than the anterior approach. However, the quality of the studies included was not satisfactory. Therefore, selection of the appropriate approach must be made cautiously and on a case-by-case basis. More high-quality, randomized controlled trials are required to guide the selection of the surgical approach in patients with thoracolumbar burst fractures.

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Conflict of interest Each author certifies that he has no commercial associations that might pose a conflict of interest with the submitted article.

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