

Excess Infant Mortality Among Native Hawaiians: Identifying Determinants for Preventive Action

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As an indigenous population, Native Hawaiians share a similar sociopolitical status with American Indians and Alaska Natives and are recognized in the Native American Programs Act with access to some but not all of the assistance programs and tribal rights accorded to American Indians.¹ Although the health status of Native Hawaiians is relatively understudied, people of Native Hawaiian ancestry have been shown to suffer a greater burden of ill health, including higher rates of cardiovascular disease, diabetes, and obesity.^{2,3} Higher mortality rates are observed across the life span, including the most premature of deaths—among infants.⁴ Indeed, the infant mortality rate (IMR) is a sentinel population health metric because it reflects the cumulative health experience of women and families as well as society's ability to care for a most vulnerable and dependent subgroup. In 2002, the latest year with national estimates on Asian and Pacific Islander subgroups,⁵ the IMR among Native Hawaiians was the second highest (after African Americans) of any racial/ethnic group and 66% higher than for Whites.⁶ Previous studies have generally alluded to predominantly postneonatal determinants of excess infant death among Native Hawaiians, but all analyses rely on data from the 1980s or earlier.^{7–11} Advances in the application of analytic techniques to quantify components of disparities^{12,13} and new strategies to promote infant health (e.g., preconception health¹⁴ and safe sleep practices¹⁵) support the need for an updated examination to identify avenues for preventive action.

We examined the determinants and cause-specific sources of excess infant mortality among Native Hawaiians relative to Whites in a contemporary cohort of births in Hawai'i, where the majority of Native Hawaiians reside. According to the 2010 Census, 55% of those who report being of Native Hawaiian race, alone or in combination with other races, reside in Hawai'i.¹⁶

Objectives. We identified potential determinants and cause-specific sources of excess infant mortality among Native Hawaiians.

Methods. We compared infant mortality rates among Native Hawaiians and Whites by using data from the 2002 to 2009 Hawai'i State Linked Birth/Infant Death Cohort File. We evaluated the components of excess infant mortality by age and underlying cause of death as well as maternal sociodemographic, behavioral, and chronic condition disparities.

Results. The Native Hawaiian infant mortality rate was more than twice that for Whites (7.9 vs 3.5/1000 live births). Excess Native Hawaiian infant mortality was equally apportioned to neonatal and postneonatal deaths. Preterm-related causes of death accounted for 43.9% of the infant mortality disparity, followed by sudden unexpected infant death (21.6%) and injury (5.6%). In multivariable models, maternal educational inequality accounted for the largest portion of the neonatal mortality disparity (20.9%); younger maternal age (12.2%) and smoking (9.5%) were the only significant contributors to the postneonatal mortality disparity.

Conclusions. Addressing educational inequalities, promoting safe sleep practices, and reducing smoking among Native Hawaiian mothers would help to eliminate excess infant mortality. (*Am J Public Health.* 2013;103:e88–e95. doi:10.2105/AJPH.2013.301294)

METHODS

Data were from the Hawai'i State Linked Birth/Infant Death Cohort Files from 2002 to 2009. Deaths in the state are routinely linked to their corresponding birth certificates if the deceased were born in Hawai'i. Because infant deaths can occur up to 1 year after birth, deaths among infants born in one calendar year cohort can occur in the next calendar year. Thus, the 2002 to 2009 cohort file includes deaths in 2002 to 2010. Deaths that occur outside of Hawai'i are not included, and they are generally considered beyond the purview of Hawai'i state health efforts. During the study period (2002–2009), 99.7% (914/917) of infant deaths among those born in Hawai'i were matched. We analyzed data for births and linked infant deaths among those born to resident women of White ($n = 33\,683$ births; $n = 119$ deaths) or Native Hawaiian ($n = 40\,917$ births; $n = 323$ deaths) race. We selected Whites as the reference group because they have the lowest IMR of all major racial groups

in Hawai'i.¹⁷ In the race-coding conventions used in Hawai'i, a designation of White generally refers to single race, but a designation of Native Hawaiian may refer to persons with a combination of races in their makeup.¹⁸ The 8 data years from 2002 to 2009 provided the maximum statistical power for multivariable analyses without significant trending in mortality rates.

Measures

We examined IMRs and disparities between births to Native Hawaiian and White women by age at death (neonatal, 0–27 days; postneonatal, 28–364 days), underlying cause of death (by *International Classification of Diseases, 10th Revision* code¹⁹), and birth weight or gestational age categories. These are conventional techniques for infant mortality investigations to evaluate the summary contribution of various causes of death that occur at different ages; neonatal deaths are largely attributable to preterm birth and congenital anomalies, and postneonatal deaths tend to be more

sensitive to infant care practices and access to primary health care (e.g., sudden unexpected infant death [SUID], injury, and infection).^{20–23}

We determined underlying cause of death by a National Center for Health Statistics computer algorithm that evaluates all of the causes listed by the certifying physician.²⁴ We grouped underlying cause-of-death codes by common categories of infant death: preterm related,²⁵ congenital malformations (Q00–Q99), SUID (R95, R99, W75), injury (U01, V01–W74, W76–Y84), and infection (A00–B99, G00, G03, J00–J21, J40–J42). The definition of preterm-related causes of death followed the classification developed by the Centers for Disease and Control and Prevention (gestational age < 37 weeks and underlying cause-of-death code K550, P000, P010, P011, P015, P020, P021, P027, P070–P073, P102, P220–229, P250–279, P280, P281, P360–369, P520–523, or P77).^{25,26} The SUID category comprised sudden infant death syndrome, unknown cause, and accidental suffocation or strangulation in bed, because of evidence of shifting reporting practices between these codes.²⁷

Analyses

To evaluate the components of the infant mortality gap connected to low birth weight and preterm birth, we conducted a Kitagawa analysis to distinguish the contribution of maternal health (i.e., percentage low birth weight and preterm birth) versus access to risk-appropriate care (i.e., birth weight- and gestational age-specific mortality rates).²² This method, described in detail elsewhere,^{12,22,28} provides a decomposition of the absolute difference in the overall IMR into the differences in the proportion born at a given category of birth weight or gestational age and differences in the mortality rates among infants born at that same birth weight or gestational age category. Birth weight data were missing in 0.04% of births, but these cases accounted for 4.0% of deaths among Native Hawaiians and only 0.8% of White deaths; we therefore imputed birth weight with the hot-deck technique used by the National Center for Health Statistics.²⁶ Because of concerns regarding the validity of gestational age data derived from the last menstrual period,²⁹ the Hawai'i Department of Health prefers the clinical estimate,

which served as the basis of gestational age assessment for our analysis. We based our calculation on the last menstrual period when the clinical estimate was missing (0.2%).

In addition, we performed a multivariable regression analysis to estimate the contribution of racial differences in sociodemographic, behavioral, and medical risk factors to disparities in infant, neonatal, and postneonatal mortality rates. We based our selection of risk factors available on the birth certificate on previous associations with perinatal outcomes; these were maternal age, education, marital status, county of residence, parity, plurality, smoking during pregnancy, and any chronic condition (cardiac, lung, or renal disease; anemia; hypertension; diabetes; and hemoglobinopathy). Although we considered including the timing of prenatal care entry, our results indicated counterintuitive confounded effects, because women who enter care later in pregnancy tend to have an extended or nonpreterm delivery by definition. Missing data for any given variable generally represented less than 1%, cumulating to 1.4% overall; we placed missing values in the reference group (mean imputation) to preserve the full study population. We performed Oaxaca decomposition^{13,30–32} to assess the contribution of prevalence differences in each factor to infant mortality disparities. This approach generally requires the use of additive or linear models where the mean of the outcome equals the sum of the mean values of the predictors multiplied by their coefficients; thus we used ordinary least squares regression.³² For binary responses, this type of regression is known as a linear probability model and provides unbiased estimators of coefficients but requires robust standard errors to account for heteroskedasticity.³³ The overall crude disparity in infant mortality ($Y_H - Y_W$) is equal to the unexplained disparity from the adjusted model (β_H) plus the sum of the adjusted effects (β coefficients) multiplied by the prevalence difference for each factor ($\sum_i^k \beta_i(X_{iH} - X_{iW})$):

$$(1) Y_H - Y_W = \beta_H + \sum_i^k \beta_i(X_{iH} - X_{iW}),$$

where H = Hawaiian and W = White.

The percentage of the disparity explained by each factor can be calculated as $\frac{\beta_i(X_{iH} - X_{iW})}{Y_H - Y_W} \times 100$. The contribution of racial differences in the effects of factors (interactions) could also be

evaluated, but we did not incorporate interactions because we found none to be significant. We conducted descriptive analyses in SAS version 9.2 (SAS Institute, Cary, NC) and the Oaxaca decomposition in Stata SE version 11.1 (StataCorp LP, College Station, TX) with the OAXACA add-on command.³⁴

RESULTS

The IMR among Native Hawaiians averaged 7.9 deaths per 1000 live births over the study period, more than twice the rate for Whites (3.5/1000; Table 1). With an average annual number of about 5000 births, the excess IMR difference among Native Hawaiians (4.4/1000) translates to about 22 deaths annually that would have been prevented if Native Hawaiians had the same risk of infant death as Whites. Although the relative infant mortality disparity was considerably greater at postneonatal than neonatal ages (relative risk = 4.0 vs 1.8), the absolute disparity was the same because of the greater overall risk of neonatal mortality. Thus, excess infant mortality among Native Hawaiians was equally apportioned to neonatal and postneonatal deaths. We detected no significant differences in infant mortality between those with full or partial Hawaiian ancestry and no significant trends over time for either Whites or Native Hawaiians (data not shown).

Among the major categories of cause of infant death, Native Hawaiian mortality exceeded that of Whites for preterm related, SUID, and injury (Table 2). Excess preterm-related deaths among Native Hawaiians accounted for 43.9% of the infant mortality disparity. SUID and injury accounted for another 21.6% and 5.6% of the infant mortality gap, respectively. Although too few White infant deaths from infection occurred to report a cause-specific IMR, a significant racial disparity could be detected (data not shown).

Consistent with the underlying cause-of-death results, about half of the infant mortality gap could be explained by the greater proportion of low birth weight and preterm birth among Native Hawaiians than Whites (Table 3). In particular, Native Hawaiians were 2.2 times as likely as Whites to be born at the smallest size (< 1000 g) and earliest gestational age (< 28 weeks), when the mortality risk is highest, and this disparity alone accounted for

TABLE 1—Native Hawaiian and White Infant Mortality Rates by Age at Death: Hawai'i, 2002–2009

Age at Death	Deaths		RR (95% CI)	Rate Difference (95% CI)	Excess Infant Deaths, ^c %
	Native Hawaiians, ^a No. (No./1000 Live Births)	Whites, ^b No. (No./1000 Live Births)			
Neonatal (0–27 d)	206 (5.0)	95 (2.8)	1.8 (1.4, 2.3)	2.2 (1.3, 3.1)	50.8
Postneonatal (28–364 d)	117 (2.9)	24 (0.7)	4.0 (2.6, 6.2)	2.1 (1.6, 2.7)	49.2
Total	323 (7.9)	119 (3.5)	2.2 (1.8, 2.2)	4.4 (3.3, 5.4)	100.0

Note. CI = confidence interval; RR = rate ratio.

^aAmong 40 917 births.

^bAmong 33 683 births.

^cCalculated as the age-specific rate difference divided by the total rate difference.

about 45% of the infant mortality gap and the majority of the total low birth weight–preterm birth contribution. In general we found no significant racial differences in the mortality rate within each low birth weight or preterm gestational age category (data not shown).

Native Hawaiian mothers were more likely than White mothers to have several socio-demographic and health risk factors: they were more likely to be adolescent and unmarried, to have lower levels of educational attainment and higher parity, to smoke, and to have a chronic condition (Table 4). In adjusted models, all covariates were significantly related to neonatal, postneonatal, or overall infant mortality, with the exception of chronic conditions. In total, these covariates explained

20.6% of the overall infant mortality disparity and a greater share of excess postneonatal than neonatal deaths (37.5% vs 4.2%). However, educational inequality explained 20.9% of the neonatal mortality gap (11.6% overall), but differences in parity and multiple births favored Native Hawaiians. Younger maternal age (12.2%) and higher smoking rates (9.5%) among Native Hawaiians were the only statistically significant contributors to the postneonatal mortality disparity.

DISCUSSION

In our updated examination of infant mortality among Native Hawaiians, we documented a large disparity between Native

Hawaiians and Whites, with infants born to Native Hawaiian mothers more than twice as likely to die as those born to White mothers in Hawai'i. This disparity translates to about 22 excess Native Hawaiian infant deaths per year in Hawai'i that would not occur if infants born to Native Hawaiian mothers had the same mortality risk as those born to White mothers. Although the relative disparity was greater for postneonatal than neonatal deaths, excess Native Hawaiian infant mortality was equally apportioned to neonatal and postneonatal deaths because of the greater overall risk of neonatal death. Although the authors did not expressly evaluate this, previous analyses showed an equal absolute disparity in both neonatal and postneonatal mortality between Native Hawaiians and Whites.^{6,9,10}

Consistent with the results for neonatal and postneonatal mortality, nearly half of all excess deaths were attributable to preterm-related causes (43.9%), followed by SUID (21.6%) and injury (5.6%). Previous studies also documented higher Native Hawaiian death rates from similar causes.^{10,11} The excess neonatal, preterm-related deaths among Native Hawaiians appeared to result from an excess of the most extremely preterm or low-birth weight births rather than differences in the gestational age– or birth weight–specific mortality risks. This suggests a role for disparities in maternal or women's health rather

TABLE 2—Cause-Specific Infant Mortality Rates Among Native Hawaiians and Whites: Hawai'i, 2002–2009

Underlying Cause of Death ^a	Deaths		RR (95% CI)	Rate Difference (95% CI)	Excess Infant Deaths, ^b %
	Native Hawaiians, No. (No./100 000 Live Births)	Whites, No. (No./100 000 Live Births)			
Preterm related ^c	133 (325.0)	45 (133.6)	2.4 (1.7, 3.4)	191.4 (123.8, 259.1)	43.9
Congenital anomalies ^d	27 (66.0)	25 (74.2)	0.9 (0.5, 1.5)	-8.2 (-46.5, 30.1)	-1.9
Sudden unexpected infant death ^e	58 (141.8)	16 (47.5)	3.0 (1.7, 5.2)	94.2 (51.0, 137.5)	21.6
Injury ^f	16 (39.1)	5 (14.8)	2.6 (1.1, 7.2)	24.3 (1.1, 47.4)	5.6
Infection ^g	12 (29.3)

Note. CI = confidence interval; RR = rate ratio. Ellipsis indicates number too small to report (< 5).

^aAccording to *International Classification of Diseases, 10th Revision (ICD-10)* categories.¹⁹

^bCalculated as the cause-specific rate difference divided by the total rate difference.

^cAccording to Centers for Disease Control and Prevention definition: gestational age < 37 weeks and underlying cause-of-death *ICD-10* code of K550, P000, P010, P011, P015, P020, P021, P027, P070–P073, P102, P220–229, P250–279, P280, P281, P360–369, P520–523, or P77.^{25,26}

^d*ICD-10* codes Q00–Q99.

^e*ICD-10* codes R95, R99, and W75.

^f*ICD-10* codes V01–W74 and W76–Y36.

^g*ICD-10* codes A00–B99, G00, G03, and J00–J42.

TABLE 3—Contribution of Low Birth Weight and Preterm Birth to the Native Hawaiian–White Infant Mortality Gap: Hawai‘i, 2002–2009

Variable	Birth Distribution			Infant Mortality Rate/1000 Live Births			Excess Infant Deaths Attributable to Distributional Differences ^a			
	Native Hawaiians, %	Whites, %	Prevalence RR (95% CI)	Prevalence Rate Difference (95% CI)	Native Hawaiians	Whites	RR (95% CI)	Rate Difference (95% CI)	No./1000	Total, %
Birth weight, g										
<1000	0.9	0.4	2.2 (1.8, 2.7)	0.5 (0.4, 0.6)	452.7	386.4	1.2 (0.9, 1.5)	66.4 (–31.8, 164.5)	1.9	45.4
1000–1499	0.7	0.6	1.3 (1.1, 1.5)	0.2 (0.0, 0.3)	55.9	65.7	0.9 (0.4, 1.7)	–9.7 (–52.8, 33.4)	0.1	2.2
1500–1999	1.4	1.2	1.1 (1.0, 1.3)	0.2 (0.0, 0.4)	27.4	19.1	1.4 (0.6, 3.3)	8.3 (–10.4, 27.0)	0.0	1.0
2000–2499	5.2	3.9	1.3 (1.2, 1.4)	1.3 (1.0, 1.6)	7.9	6.8	1.2 (0.5, 2.6)	1.1 (–4.7, 6.9)	0.1	2.3
Total low birth weight	8.3	6.2	1.3 (1.3, 1.4)	2.1 (1.7, 2.5)	61.5	39.1	1.6 (1.2, 2.0)	22.4 (10.8, 34.0)	2.2	50.9
Gestational age, wk										
<28	0.9	0.4	2.2 (1.8, 2.6)	0.5 (0.4, 0.6)	416.9	368.1	1.1 (0.9, 1.5)	48.8 (–44.3, 141.9)	2.0	45.9
28–31	0.9	0.7	1.2 (1.1, 1.5)	0.2 (0.1, 0.3)	31.7	48.0	0.7 (0.3, 1.4)	–16.3 (–48.2, 15.5)	0.1	1.7
32–33	1.3	1.0	1.3 (1.1, 1.5)	0.3 (0.1, 0.4)	36.0	8.9	4.1 (1.2, 13.6)	27.1 (8.3, 45.9)	0.1	1.5
34–36	8.3	6.4	1.3 (1.2, 1.4)	1.9 (1.5, 2.3)	5.9	5.1	1.2 (0.6, 2.4)	0.8 (–3.2, 4.8)	0.1	2.5
Total preterm	11.4	8.6	1.3 (1.3, 1.4)	2.9 (2.4, 3.4)	44.8	27.5	1.6 (1.3, 2.1)	17.3 (8.9, 25.7)	2.2	51.6

Note. CI = confidence interval; RR = rate ratio.

^aDetermined by the Kitagawa method, the racial difference in percentage born in a certain birth weight–gestational age category multiplied by the average of race-specific infant mortality rates for that category (excess deaths attributable to birth weight- or gestational age-specific mortality rate differences not shown because of lack of statistical significance).

than access differences to risk-appropriate neonatal care.

In our examination of the contribution of compositional differences in sociodemographic and medical risk factors, educational disparities accounted for the largest component of excess neonatal mortality (20.9%) and was likely underestimated because of possible mediators in the multivariable model (e.g., smoking, chronic conditions). Only a third of Native Hawaiian mothers but more than half of White mothers had attained postsecondary education. Greater educational attainment confers health advantages through knowledge and awareness of positive health behaviors as well as the economic resources to access health-promoting goods and services.³⁵ Efforts to reduce educational inequalities would address a fundamental social determinant of health and help to reduce a variety of health disparities, including those in neonatal mortality, possibly via improvements in women’s health status. Improving educational quality and hope for future career potential may also help to reduce adolescent childbearing,³⁶ a factor associated with excess postneonatal mortality among Native Hawaiians. Hawai‘i is among the 21 states and the District of Columbia to receive Race to the Top funding—President Barack Obama’s key educational reform initiative—with a goal of reducing Native Hawaiian educational achievement gaps.³⁷ In addition, the Affordable Care Act has funded models of maternal, infant, and early childhood home visiting that have shown an impact on child development and school readiness as well as on maternal education and vocational training.³⁸ This funding has expanded Hawai‘i home visiting services from a single model (Healthy Families America) operating in only 2 locations to a multi-model network serving at-risk communities in the entire state (A. H. H., personal communication with Home Visiting Unit, Hawai‘i State Department of Health, August 21, 2013).

Disparities in smoking also contributed significantly to excess postneonatal mortality. Smoke exposure both before and after delivery is a risk factor for SUID and increases susceptibility to respiratory infection.³⁹ Anonymous reporting in the Pregnancy Risk Assessment and Monitoring System reveals greater disparities in postpartum smoking, during a key exposure period for SUID and infection (22.0%

TABLE 4—Multivariable Decomposition of the Native Hawaiian-White Infant Mortality Disparity: Hawaii, 2002–2009

Maternal Characteristic	Covariate Distribution ^a		Adjusted Linear Parameter Estimates, Deaths/1000 Live Births			Disparity Explained by Covariate Distributional Differences ^b		
	Native Hawaiians, %	Whites, %	Total Infant Mortality, b (SE)	Neonatal Mortality, b (SE)	Postneonatal Mortality, b (SE)	Total Infant Mortality, % (SE)	Neonatal Mortality, % (SE)	Postneonatal Mortality, % (SE)
Race								
Native Hawaiian	100.0	0.0	3.46* (0.65)	2.12* (0.54)	1.34* (0.37)			
White (Ref)	0.0	100.0						
Maternal age, y								
< 20	15.8	3.8	2.34* (1.17)	0.49 (0.97)	1.85* (0.66)	3.2 (4.4)	-5.5 (7.2)	12.2* (5.0)
20–24	31.5	25.1	-0.43 (0.73)	-0.70 (0.60)	0.28 (0.41)			
25–34 (Ref)	42.8	53.5						
≥ 35	10.0	17.6	1.48 (0.88)	1.80* (0.73)	-0.32 (0.34)			
Education								
< High school	14.7	6.7	4.28* (1.22)	2.72* (1.01)	1.56* (0.69)	11.6* (4.0)	20.9* (6.6)	1.9 (4.4)
High school	52.3	34.7	1.29 (0.84)	1.82* (0.70)	-0.53 (0.48)			
Some college	20.7	25.8	1.31 (0.87)	1.50* (0.72)	-0.18 (0.50)			
≥ College (Ref)	12.3	32.8						
Marital status								
Married (Ref)	41.8	81.8						
Unmarried, father listed	50.0	14.3	0.12 (0.72)	-0.31 (0.59)	0.43 (0.41)	7.0 (6.2)	4.6 (10.1)	9.6 (7.0)
Unmarried, father not listed	8.2	3.9	6.08* (1.25)	4.86* (1.03)	1.21 (0.71)			
County of residence								
Oahu (Ref)	62.0	68.8						
Hawaii	20.0	12.5	-0.02 (0.78)	0.07 (0.65)	-0.09 (0.44)	0.1 (1.5)	0.2 (2.4)	0.0 (1.7)
Maui	12.5	13.6	-1.68 (0.87)	-0.75 (0.72)	-0.93 (0.49)			
Kauai	5.5	5.1	-2.92* (1.27)	-2.28* (1.05)	-0.64 (0.72)			
Parity								
First child	35.6	46.5	1.60* (0.65)	2.06* (0.54)	-0.46 (0.37)	-6.3* (2.3)	-16.4* (3.8)	4.2 (2.7)
Second or third child (Ref)	47.2	45.4						
≥ fourth child	17.3	8.1	-1.06 (0.91)	-1.49* (0.76)	0.43 (0.52)			
Plurality								
Singleton (Ref)	97.3	96.4						
Multiple	2.7	3.6	16.28* (1.63)	15.34* (1.35)	0.94 (0.93)	-3.5* (0.8)	-6.4* (1.5)	-0.4 (0.5)
Smoking								
No (Ref)	88.9	96.7						
Yes	11.1	3.3	2.82* (1.10)	0.18 (0.91)	2.64* (0.62)	5.0* (2.5)	0.6 (3.5)	9.5* (3.6)

Continued

TABLE 4—Continued

Chronic condition							
No (Ref)	78.3	90.7	1.18 (0.78)	1.11 (0.65)	0.07 (0.44)	3.4 (2.6)	6.2 (4.3)
Yes	21.8	9.3				20.6* (8.3)	4.2 (3.4)
Total explained							37.5* (9.6)

^aAll covariates differed significantly by race as determined by the χ^2 test; $P < .01$.

^bObtained from Oaxaca decomposition derived from adjusted coefficients multiplied by covariate distributional differences (summed over multiple category variables); positive numbers indicate a Native Hawaiian disadvantage that helps to explain disparity; negative numbers indicate a Native Hawaiian advantage that would serve to increase disparity if the covariate distribution were equalized to that of the White population; total percentage explained is the sum over all covariates, which equals the difference between the unadjusted and adjusted racial disparity divided by the unadjusted racial disparity.

* $P < .05$.

among Native Hawaiians vs 8.0% among Whites in 2009).⁴⁰ The State of Hawai'i has implemented aggressive population-based tobacco control policies, including a public smoking ban in 2006 and large increases in cigarette taxes.⁴¹ Even since 2009, the last year of our study, per-pack taxes have risen 60 cents to \$3.20, and only 3 other states have a higher tobacco tax (New York, Rhode Island, and Connecticut).⁴¹ This combination of state policies,⁴² along with the Affordable Care Act's requirement that Medicaid cover comprehensive smoking cessation treatments (both counseling and medication) for pregnant women,⁴³ may help to reduce maternal smoking.

Other preventive opportunities suggested by our results include safe sleep campaigns to reduce excess SUID and improved messaging and health care access to reduce deaths from injury and infection. Analyses of the Hawai'i Pregnancy Risk Assessment and Monitoring System have confirmed that Native Hawaiian mothers are more likely to report using non-supine infant sleep positions (side or stomach), which increase SUID risk.⁴⁴ Soft bedding and bed sharing are additional risks that can be addressed with education on positioning through health care provider training, media campaigns, and the provision of bedside bassinets, which can support breastfeeding while minimizing the risks of bedsharing.^{45–47} In response to findings from the Hawai'i Child Death Review team, a culturally appropriate video on safe sleep practices was developed for use in the Special Supplemental Nutrition Program for Women, Infants, and Children and community health clinics. Additional avenues for SUID reduction are establishment of a safe sleep committee and laws requiring the provision of safe sleep counseling and education by all health care providers to expectant mothers and to families in postpartum units and during pediatric visits.⁴⁸

Several programs in Hawai'i could disseminate these messages and provide resources or referrals for preventing SUID, injury, and parental smoking. The Hawai'i home visiting models implemented through Affordable Care Act funding³⁸ provide screening, services, and counseling to reduce environmental risks for child maltreatment, such as establishing a medical home, providing connections to community resources, and promoting positive

parenting practices. The Native Hawaiian Health Care Systems were created by the Native Hawaiian Health Care Act of 1988⁴⁹ to provide essential educational, enabling, and health care services to Native Hawaiians in the State of Hawai'i. They operate on 5 of the Hawaiian islands and are an important service point for Native Hawaiian families that could be used to deliver messages and resources to promote smoking cessation and prevent SUID and other injury deaths.⁵⁰

Limitations

The advantage of capturing nearly all births and deaths in Hawai'i vital records was tempered by limitations in the measurement of certain risk factors on the birth certificate. For example, medical risk factors such as hypertension and diabetes are underreported on birth certificates,^{51,52} and others (e.g., maternal obesity) are not collected on the birth certificate currently used in Hawai'i. Thus, the impact of chronic conditions on infant mortality was likely underestimated and consequently their contribution to disparities as well. Other analyses have shown a greater burden of chronic conditions and obesity (not measured here) among Native Hawaiians,^{2,3,53–56} as well as a connection to infant mortality.⁵⁷ Linkages between vital records and discharge data would help to reduce this knowledge gap but will require the state vital statistics department to complete linkages with personally identifiable information. Additional variables not available on the birth certificate or medical record, such as stress, social support, poverty, and unintended pregnancy, could be explored in analyses of the Pregnancy Risk Assessment and Monitoring System.

Finally, our analysis was only representative of the disparity in Hawai'i. Although the majority of Native Hawaiians live in Hawai'i,¹⁶ a national analysis could be conducted in the coming decade when all states have implemented the 2003 revision of the birth certificate, which collects multiple race and other variables in a comparable format. Granular racial/ethnic data that identifies Native Hawaiians separately from other Pacific Islander groups is required by all federally sponsored surveys in section 4302 of the Affordable Care Act and marks a major advance in the ability to identify, monitor, and address racial/ethnic

disparities with appropriate specificity and uniformity.⁵⁸ Nonetheless, small sample sizes still impede the examination of Native Hawaiians in many national health surveys.

Conclusions

Our findings regarding the determinants of excess infant mortality among Native Hawaiians suggest several avenues for preventive action and further research. Reducing educational inequality, perhaps through reform initiatives and home-visiting programs, would go furthest in addressing a fundamental cause of multiple health outcomes. Additional strategies to address specific health risks related to postneonatal mortality, particularly safe sleep practices and smoking, are also warranted. Further efforts to link birth certificate and hospital discharge records and to analyze Pregnancy Risk Assessment and Monitoring System data would yield more information on the contribution of additional risk factors to preterm birth and neonatal mortality, such as chronic conditions, obesity, and other social determinants beyond education. ■

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Contributors

A. H. Hirai conceptualized and designed the study, conducted analyses, and led the writing. D. K. Hayes assisted with data acquisition and, with M. M. Tauli'i and L. J. Fuddy, contributed knowledge and information on relevant public health programs for practice-based implications. G. K. Singh provided analytic support and content expertise. All authors interpreted results and critically revised the article.

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Human Participant Protection

Ethics review was not required because the data did not include personal identifiers.

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