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## Risk of Injury Due to Alcohol – Evaluating Potential Bias Using the Case-Crossover Usual-Frequency Method

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### Abstract

**Background**—The usual-frequency case-crossover method, comparing exposure before an event with typical exposure of the same person, is widely used to estimate the risk of injury related to acute alcohol use. Prior results suggest that risk estimates might be biased upward compared with other methods.

**Methods**—Using data from 15 emergency-room studies in 7 countries, we compared the usual-frequency case-crossover method with case-control analysis, using non-injury patients as controls. Control-crossover analysis was performed to examine potential bias and to adjust risk estimates.

**Results**—The cross-study pooled odds ratio (OR) of injury related to drinking was 4.7 (2.6–8.5) in case-crossover analysis and 2.1 (1.6–2.7) in case-control analysis. A control-crossover analysis found an indication of bias (OR=2.2 [1.8–2.8]), which was larger among less frequent drinkers.

**Conclusion**—Findings suggest that the potential overestimation of injury risk based on the usual-frequency case-crossover method might be best explained by recall bias in usual-frequency estimates.

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The case-crossover design was introduced to evaluate the effect of transient exposures on the risk of acute outcomes.<sup>1,2</sup> One application has been in studies of emergency-room patients to estimate the risk of injury related to alcohol consumption. Unlike traditional case-control studies using non-injury emergency-room patients<sup>3</sup> or community samples<sup>4</sup> as

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controls, the case-crossover design compares injury patients' drinking before the event with their own alcohol intake during an earlier control period. Two commonly used approaches to assessing exposures for the control period are the usual frequency of drinking (e.g., within the last year)<sup>5-8</sup> and drinking during one or several matched prior intervals (e.g., the same time one week before injury).<sup>6,9-11</sup>

Results from previous studies, however, suggest that risk estimates produced using the usual-frequency case-crossover method might be biased upward when compared with other methods. For example, the usual-frequency method generated larger estimates than the matched interval case-crossover method in the two studies employing both approaches.<sup>6,9</sup> A usual-frequency case-crossover analysis on emergency-room data in 16 countries produced a pooled odds ratio (OR) of 5.7 (95% confidence interval (CI)= 4.0-8.0) for injury related to any drinking,<sup>7</sup> substantially larger than the pooled odds ratio of 1.6 (1.2-1.9) from a 5-country case-control study using non-injury patient controls.<sup>12</sup>

This study examines potential bias using the usual-frequency case-crossover method, compared with the case-control design for estimating risk of injury from drinking. A case-crossover analysis on control (i.e., non-injury) patients (called the control-crossover approach) is also conducted to inform this comparison. Because the excess risk with control-crossover is expected to be zero, the approach has been widely used in case-crossover studies as a validity check and to adjust biased estimates.<sup>9,13-18</sup>

## METHODS

Our analyses are based on data from the Emergency Room Collaborative Alcohol Analysis Project,<sup>12,19</sup> including 15 studies (each typically covering a city or region) in 7 countries. Data from probability samples of emergency-room patients admitted for injury or illness were collected using a similar methodology.<sup>20</sup> By using data from various regions around the world, findings from these analyses are more robust, with risk estimates examined in various socioeconomic and cultural conditions thought to be associated with alcohol use and injury, as well as varying levels of emergency-room utilization.

Cases and controls were defined as patients who reported their primary reason for visiting the emergency room as either injury or non-injury illness. Injury and non-injury patients were compared in case-control analysis, whereas only injury patients were used in the case-crossover analysis and only non-injury patients in the control-crossover analysis. For the injury and non-injury samples, exposure to acute drinking was defined as consumption of any alcoholic beverage by the patient during the six-hour period prior to either the injury or the illness event that led to their emergency room visit.

In case-control analysis, adjusted ORs were estimated from unconditional logistic regressions comparing acute drinking between injury and non-injury patients, controlling for sex and age. For the usual-frequency case- or control-crossover analysis, risk estimates were generated comparing acute drinking of the injury or non-injury patients before the event to their usual frequency of drinking. The case or control-crossover analysis can be considered as a stratified self-matched case-control analysis with the OR generated from the Mantel-Haenszel estimator for dichotomous exposure<sup>21</sup> (any drinking versus none). Maclure<sup>1</sup> showed that for the usual-frequency case-crossover analysis, the OR can be derived by dividing the total expected unexposed periods of the acutely exposed cases by the total expected exposed periods of the acutely unexposed cases. The expected exposed periods were obtained from subjects' usual frequency of drinking during last 12 months. To be consistent with the acute drinking assessment, a 6-hour period was also used as the effect period for each occasion of any alcohol use in the control period. The expected unexposed

periods were derived by subtracting the expected number of exposed periods from the total number of possible 6-hour periods each year, defined as  $365 \cdot 3$ , which excluded one six-hour sleeping period each day.<sup>22</sup>

The control-adjusted case-crossover estimates were obtained by dividing ORs from the case-crossover by those from the control-crossover analysis. In addition to study-specific estimates, we also estimated the pooled overall effect using a meta-analytic method<sup>23</sup> of weighted average, taking into account the standard errors of the study-specific estimates. Only random-effect estimates are reported in Table 1, given that the homogeneity test yielded very small p-values.

## RESULTS

As shown in Table 1, the usual-frequency case-crossover method produced ORs at least 1.5 times as large as the corresponding estimates from case-control analyses for 11 out of the 15 emergency-room studies, and 2 times the case-control results in 8 studies. The pooled OR was 4.7 (95% CI= 2.6–8.5) from case-crossover analysis, compared with 2.1 (1.6–2.7) from case-control analysis. This trend with case-crossover estimates was also seen in the control-crossover analysis. ORs were greater than 1.5 for 13 studies and greater than 2 for 9 studies. The control-crossover pooled OR was 2.2 (1.8–2.8), substantially larger than the expected estimate of 1. After adjusting the case-crossover estimate, based on the control-crossover estimate, the pooled adjusted case-crossover OR was 2.1 (1.5–3.1) – very close to 2.1 (1.6–2.7) from case-control analysis. The across-study Pearson correlation between the case-control and the unadjusted and adjusted case-crossover estimates (logarithm of ORs) improved from 0.75 to 0.84.

To examine the potential source of upward bias in the case-crossover estimates, we performed separate control-crossover analyses for subsamples at various drinking-frequency levels. For each frequency level, with all studies combined, the control-crossover ORs were quite similar to the ratio between the observed and expected exposure prevalence (Table 2). ORs generally decreased monotonically, from the lowest to the highest usual frequency level. The largest bias was found among those reporting the least frequent drinking (OR=15 for 1–5 times last year), while the estimate for “daily drinkers” was close to 1. Control-crossover analyses were also performed for each of the six countries separately, (Italy was dropped due to a small sample size) although with a slightly modified usual-frequency measure reduced to four categories. As shown in the Figure, clear monotonic relationships between the control-crossover estimates and usual frequency of drinking were observed for four of the six countries. One country exhibited no relationship (Spain, with log ORs close to zero for each usual-frequency level) and one appeared to be variable but with log ORs all above zero (Poland).

## DISCUSSION

Our comparisons between the case-control and case-crossover analysis and the control-crossover estimates show that the usual-frequency case-crossover method apparently overestimates the risk of injury related to drinking. In particular, the observed monotonic relationship between control-crossover estimates and usual frequency levels (where larger biases were associated with less-frequent drinking) may be most plausibly explained by recall bias. Survey estimates assessing annual alcohol volume, particularly estimates derived from usual-frequency and quantity measures, have consistently been found to account for only a fraction of per capita consumption from sales data.<sup>24</sup> Our results suggest the bias is more likely among persons with a less consistent drinking pattern, who may thus have more

difficulty recalling their drinking over time – particularly over an assessment period as long as 12 months.

One possible alternative explanation for the larger case-crossover estimates observed is that, among non-injury emergency room patients, some illnesses (e.g., myocardial infarction<sup>25</sup>) may be related to alcohol use prior to the event. To the extent this is true, estimates from the control-crossover analyses represent at least some elevated risk of non-injury illness associated with acute drinking, and case-control analyses may thus underestimate risk of injury to some extent. However, the clear monotonic relationship between usual frequency and control-crossover estimates are stronger evidence for recall bias than an effect of drinking on non-injury illness.

The cross-study pooled estimate from case-control analysis is virtually the same as that from case-crossover analysis after adjusting for potential bias suggested by control-crossover analysis (OR=2.1 for both). Not all study-specific results are similar, however. For example, case-crossover adjusted ORs in the three Mexican studies are all larger than case-control estimates, suggesting some inconsistencies not caused by random error. There has been debate regarding whether the two designs are comparable, with the case-control study asking “why is the event happening to me versus another person,” whereas the case-crossover analysis asks “why now versus another time.”<sup>26</sup> Conversely, both case-control and case-crossover designs can be linked to an underlying cohort study,<sup>21,27</sup> and hypothetical examples have been constructed to show the equivalence of estimates from the two approaches if they follow the same assumed data-generating process.<sup>27</sup> While the results from the two approaches might not necessarily converge, large discrepancies between their estimates suggest bias; here, recall bias seems a plausible candidate. This potential bias is not restricted to the usual-frequency case-crossover design for estimating the alcohol-injury relationship; it would potentially apply to any exposure based on self-reported data over a long retrospective window.

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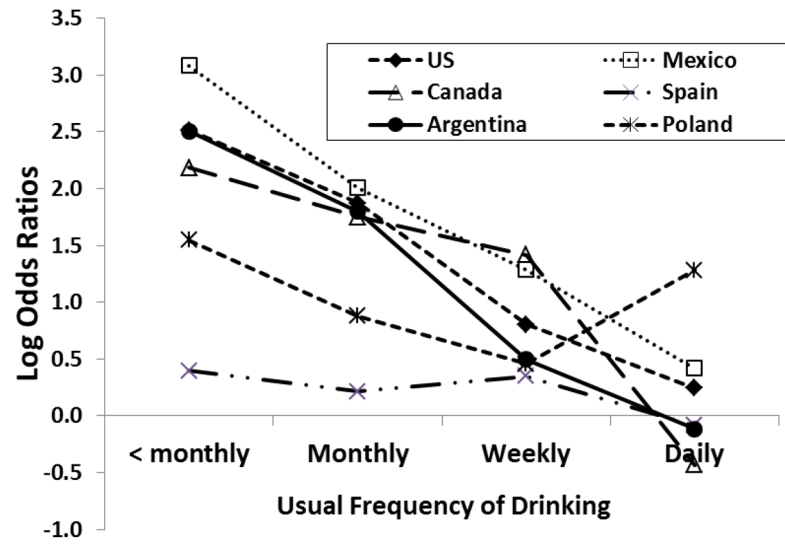
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**Figure.**  
Control-crossover estimates by usual frequency of drinking for 6 counties.

Table 1

Odds Ratios and 95% Confidence Intervals for Injury Related to Drinking from Case-control and Usual-Frequency Case-crossover Analysis, and ORs from Control-crossover Analysis, by Study Location and Pooled Estimates

Studies	No. injured (No. exposed) <sup>a</sup>	No. not injured (No. exposed) <sup>a</sup>	Case-control ORs (95% CIs)	Case-crossover ORs (95% CIs)	Control-crossover ORs (95% CIs)	Case-crossover Adj ORs <sup>b</sup> (95% CIs)
San Francisco, CA, US	550 (191)	1,276 (199)	2.6 (2.1–3.3)	5.7 (4.6–7.0)	1.8 (1.5–2.1)	3.2 (2.4–4.1)
Contra Costa, CA, US	1,001 (190)	1,363 (194)	1.1 (0.9–1.4)	2.9 (2.4–3.4)	3.2 (2.7–3.8)	0.9 (0.7–1.1)
Kaiser hospital, CA, US	410 (34)	546 (39)	1.0 (0.6–1.7)	1.1 (0.8–1.6)	1.6 (1.2–2.3)	0.7 (0.4–1.1)
Jackson, MS, US	275 (62)	739 (48)	3.0 (1.9–4.5)	7.8 (5.4–11.1)	2.2 (1.6–3.0)	3.5 (2.2–5.6)
Santa Clara, CA, US	288 (44)	988 (62)	2.3 (1.5–3.6)	3.0 (2.1–4.4)	2.0 (1.5–2.7)	1.5 (0.9–2.4)
Mexico City, Mexico	1,609 (434)	533 (43)	3.0 (2.2–4.3)	24.0 (20.8–27.8)	4.1 (2.8–5.9)	5.9 (4.0–8.7)
Acapulco, Mexico	342 (99)	289 (30)	3.0 (1.9–4.8)	29.5 (21.8–39.9)	4.4 (3.0–6.5)	6.7 (4.1–10.9)
Pachuca, Mexico	668 (96)	727 (20)	4.1 (2.5–6.9)	18.1 (13.5–24.3)	3.4 (2.1–5.6)	5.3 (3.1–9.5)
Alberta, Canada	335 (96)	494 (44)	3.6 (2.4–5.3)	7.6 (5.7–10.0)	2.3 (1.7–3.2)	3.2 (2.1–5.0)
Quebec, Canada	263 (41)	387 (52)	1.0 (0.6–1.6)	3.6 (2.5–5.1)	3.5 (2.6–4.7)	1.0 (0.7–1.6)
Barcelona, Spain	1,683 (262)	679 (88)	1.1 (0.9–1.5)	1.1 (1.0–1.3)	1.0 (0.8–1.3)	1.1 (0.9–1.5)
Triste, Italy	303 (118)	157 (37)	1.8 (1.1–2.8)	2.9 (2.2–3.8)	1.4 (0.9–2.0)	2.1 (1.3–3.3)
Mar del Plata, Argentina	378 (85)	417 (44)	1.9 (1.2–2.8)	3.0 (2.3–3.8)	1.5 (1.1–2.1)	2.0 (1.3–3.0)
Warsaw, Poland	507 (45)	205 (6)	2.8 (1.1–6.8)	2.4 (1.7–3.3)	1.5 (0.6–3.8)	1.5 (0.6–4.0)
Sosnowiec, Poland	421 (67)	306 (19)	2.4 (1.4–4.2)	3.7 (2.8–4.9)	2.8 (1.6–4.7)	1.4 (0.7–2.5)
Pooled estimate <sup>c</sup>			2.1 (1.6–2.7)	4.7 (2.6–8.5)	2.2 (1.8–2.8)	2.1 (1.5–3.1)

<sup>a</sup>Injured and not injured patients are exposed to acute drink if they consumed alcohol 6-hour before their injury or illness events

<sup>b</sup>Adjusted ratios are case-crossover ORs divided by control-crossover ORs

<sup>c</sup>Pooled random effects from meta-analysis.



**Table 2**

Odds Ratios from Control-crossover Analysis as well as Observed and Expected Exposure Prevalence Separately for Usual Frequency Levels, all Emergency Room Studies Combined

Usual Frequency	No.	Observed Exposure prevalence before event (%)	Exposure prevalence expected from usual frequency (%)	Control-crossover ORs (95% CI)
1–5 times last year	1535	4.0	0.3	15.3 (11.9–19.8)
6–11 times last year	455	6.0	0.8	7.6 (5.2–11.2)
Once per month	616	9.7	1.1	9.8 (7.5–12.7)
2–3 times per month	797	11.4	2.7	4.6 (3.7–5.7)
1–2 times per week	1012	18.3	7.1	2.9 (2.5–3.4)
3–4 times per week	485	26.0	16.6	1.8 (1.4–2.2)
(Nearly) every day	1033	36.2	33.3	1.1 (1.0–1.3)