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## Alcohol Consumption and Cardiovascular Mortality among U.S. Adults, 1987–2002

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#### Abstract

**Objectives**—To determine the association of alcohol consumption and cardiovascular mortality in the U.S. population.

**Background**—Alcohol consumption has been associated with a lower risk of cardiovascular disease in cohort studies, but this association has not been prospectively examined in large, detailed, representative samples of the U.S. population.

**Methods**—We analyzed nine iterations of the National Health Interview Survey, an annual survey of a nationally representative sample of U.S. adults between 1987 and 2000. Exposures of interest included usual volume, frequency, and quantity of alcohol consumption and binge drinking. Mortality was ascertained through linkage to the National Death Index through 2002. Relative risks were derived from random-effects meta-analyses of weighted, multivariable-adjusted hazard ratios for cardiovascular mortality from nine survey administrations.

**Results**—Light and moderate volumes of alcohol consumption were inversely associated with cardiovascular mortality. Compared with lifetime abstainers, summary relative risks were 0.95 (95% confidence interval [CI], 0.88–1.02) among lifetime infrequent drinkers, 1.02 (95% CI, 0.94–1.11) among former drinkers, 0.69 (95% CI, 0.59–0.82) among light drinkers, 0.62 (95% CI, 0.50–0.77) among moderate drinkers, and 0.95 (95% CI, 0.82–1.10) among heavy drinkers. The magnitude of lower risk was similar in subgroups of sex, age, or baseline health status. There was no simple relation of drinking pattern with risk, but risk was consistently higher among those who consumed 3 compared to 2 drinks per drinking day.

**Conclusions**—In nine nationally representative samples of US adults, light and moderate alcohol consumption were inversely associated with CVD mortality, even when compared with lifetime abstainers, but consumption above recommended limits was not.

#### Introduction

Alcohol consumption has been consistently associated with a lower risk of cardiovascular disease (CVD) in epidemiological studies (1,2), an association attributed in great part to the increase in high-density lipoprotein cholesterol (HDL-C) caused by alcohol consumption (3).

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However, a number of uncertainties about the association of alcohol consumption and CVD remain, punctuated by the absence of a long-term randomized controlled trial on CVD events with which to confirm the results of observational studies. These uncertainties include potentially diverse effects on coronary heart disease (CHD) and stroke (4), inclusion of former or occasional drinkers with long-term abstainers as a referent category (5), generalizability to the adult U.S. population (6), and the importance of drinking patterns in modifying the association (7). Measures of overall volume of alcohol consumption obscure the relative contributions of drinking frequency (how often alcohol is consumed), drinking quantity (how much alcohol is typically consumed on those days), and binge drinking (episodes of 5 or more drinks in a day), and their individual contributions to CVD risk have not been thoroughly investigated.

To evaluate the associations of alcohol consumption and drinking patterns with CVD, cerebrovascular, and CHD mortality in a nationally representative sample of U.S. adults, we aggregated data from nine administrations of the National Health Interview Survey (NHIS). The NHIS is an annual in-person survey conducted by the National Center for Health Statistics (NCHS) of the Centers for Disease Control and Prevention and has included detailed questions on alcohol consumption on selected surveys since 1987.

#### Methods

#### Study Sample

Since 1957, the NCHS has conducted annual cross-sectional household interview surveys of the noninstitutionalized civilian U.S. population. NHIS data are collected during a personal household interview conducted by interviewers employed and trained by the U.S. Bureau of the Census. The sampling plan follows a multistage area probability design, with oversampling of selected minority populations, including both black and Hispanic persons. Survey participation is voluntary, and the confidentiality of responses is assured by the Public Health Service Act.

The content of the survey is updated approximately every 10–15 years and has remained relatively unchanged since 1997. The NHIS questionnaire from 1982–1996 included a core set of basic health and demographic items and one or more added sets of questions on diverse health topics, but did not consistently include questions about health-related behaviors, such as alcohol consumption. Since 1997, the NHIS questionnaire has similar core questions that include demographic information, health status and limitations, health insurance, income and assets, and diverse health behaviors.

Our study used data from NHIS years that included alcohol consumption data (1987–1988, 1990–1992, and 1997–2000). Response rates for alcohol-related modules for adults included in these NHIS years range from 70% in 1999 to 87% in 1988. All NHIS data are publicly available (http://www.cdc.gov/nchs/about/major/nhis/quest\_data\_related\_doc.htm).

#### **Alcohol Consumption**

As noted, alcohol consumption was ascertained variably in the NHIS. Table 1 summarizes the methods used. In 1987 and 1992, a validated food frequency questionnaire (8) was administered that included separate quantity-frequency items for beer, wine, and liquor. In 1988 and 1997–2000, alcohol-related items included 1) intake of 12 or more drinks in one's lifetime, in any previous year, and in the past year, and 2) drinking frequency, drinking quantity, and binge drinking in the past year. In 1990, similar questions were asked but with a time frame of the last two weeks for current drinking; only individuals who reported that those two weeks were typical for consumption in the past year were included in these analyses. Quantity-frequency questions were also administered in 1991 as part of a Drug and

Alcohol Use module to adults aged 18–44 years, although intake in any previous year was not queried. Binge drinking was ascertained in 1988, 1991, and 1997–2000.

From these questions, we derived six alcohol consumption categories. Current abstainers were categorized as never drinkers (<12 drinks in one's lifetime), lifetime infrequent drinkers (>12 in one's lifetime, <12 drinks in any previous year), and former drinkers (>12 in one's lifetime, >12 drinks in a previous year). Current drinkers were categorized as light (current use of 3 drinks per week), moderate (current use of >3–7 drinks per week for women and >3–14 drinks per week for men), and heavy (current use of >7 drinks per week for women and >14 drinks per week for men). In 1987 and 1992, the three abstainer categories could not be differentiated and were combined. In 1991, former and lifetime infrequent drinkers were combined but could be distinguished from never drinkers.

#### **Mortality Ascertainment**

In 2004, the NCHS conducted an updated linkage of the NHIS to death certificate data found in the National Death Index (NDI). Linked mortality files include NHIS surveys from 1986 to 2000 with mortality follow-up from the month of interview through December 31, 2002. Mortality information is based upon the results from a probabilistic match between NHIS and NDI death certificate records (9). All NHIS participants 18 years and older were eligible for mortality follow-up. To minimize potential breaches of confidentiality, the NCHS developed a minor data perturbation plan; empirical studies suggest that the original and perturbed data files yield very similar descriptive statistics and model-based associations with mortality (10).

Deaths were coded for underlying cause using NCHS bridge codes spanning the Ninth and Tenth Revisions of the International Classification of Diseases (ICD-9, years 1979 to 1998; ICD-10, years 1999 to present) (11). We separately examined cardiovascular mortality (bridge codes 053 and 075, corresponding to ICD-10 I00–I99), CHD mortality (bridge code 058, ICD-10 I20–I25), and cerebrovascular mortality (bridge code 070, ICD-10 I60–I69).

#### Covariates

We included a core set of covariates available in all NHIS surveys that also assessed alcohol consumption. We categorized self-reported race/ethnicity as non-Hispanic white, non-Hispanic black, Hispanic, and other. We grouped educational attainment as less than high school, high school, and some college, and marital status as married or non-married. We categorized region of residence in four U.S. Census regions (Northeast, South, Midwest, and West). Urbanization was categorized as residence in a U.S. census-defined metropolitan statistical area (MSA) with a population of 1 million or more, a population of less than 1 million, or outside of a defined MSA. Smoking status was defined in five categories: never smokers, former smokers, and tertiles of intensity among current smokers. We calculated body-mass index (BMI) based upon self-reported height and weight and used standard World Health Organization categories of underweight, normal weight, overweight, and obese. Self-reported general health status was categorized as excellent, very good, good, and fair/poor/unknown.

NCHS queried physical activity and diet only in selected NHIS administrations. Leisuretime physical activity was assessed with a series of items on specific activities in 1990 and 1991 (12,13) and general items about moderate and heavy physical exertion in 1997–2000 (14); to combine these, activity was categorized in quartiles. Diet was assessed with a foodfrequency questionnaire in 1987 and 1992. As previously described (15), we used foods grouped as fruits and vegetables, meat/poultry/fish, dairy products, added fats, and desserts, all categorized in tertiles.

#### **Statistical Analysis**

We performed separate, similar analyses of alcohol consumption and mortality for each NHIS survey year using Cox proportional hazards regression, using appropriate sample weights provided by the NCHS. We set age as the metameter (16) and adjusted for sex and race in initial models, and additionally for marital status, education, region, urbanization, smoking, BMI, and general health status in multivariable models.

From the hazards ratios estimated in individual Cox proportional hazards models, we derived summary estimates of relative risk using random effects meta-analysis to combine across survey years (17,18). Because each survey provides estimates for the same underlying population (the non-institutionalized adult population of the U.S.), we also performed pooled analyses in sensitivity analyses.

For primary analyses, we set never drinkers as a referent; because this excluded data from 1987 and 1992, we included these years in additional analyses with all non-drinkers as a referent only after comparing the risk associated with former and lifetime infrequent drinking. We tested several domains of alcohol consumption, including average volume of intake, frequency and quantity of intake, and binge drinking.

We performed sensitivity analyses with additional adjustment for diet or physical activity where such data were available, with exclusion of individuals who died within the first two years (i.e., a two-year lag), and restricted to five years of follow-up for each survey (to reduce misclassification over longer follow-up). We performed prespecified analyses stratified by sex, age (<55 versus 55 and older), race (white versus non-white), and general health (excellent versus other); to maximize precision, we used all available NHIS years in stratified analyses. Given the possibility that low folate intake attenuates the lower risk of CHD associated with alcohol intake (19), we also examined risk stratified by follow-up that occurred prior and subsequent to the initiation of folate fortification in the U.S. in 1998 (20).

#### Results

A total of 245,207 participants were included in the nine iterations of the NHIS included in these analyses (Table 1), generally ranging from 20,000–40,000 participants per year. There were a total of 10,670 cardiovascular deaths during 1,987,439 person-years of follow-up, including 6,135 related to CHD and 1,758 to stroke. Baseline characteristics according to alcohol consumption in the pooled sample are shown in the Supplementary Tables.

#### Average Volume of Alcohol Consumption

Figure 1 shows adjusted hazard ratios for CVD mortality according to light, moderate, and heavy drinking at baseline in the individual NHIS surveys. In general, light and moderate drinking were associated with lower risk than abstention, with somewhat lower risk estimates for moderate drinking, while heavy drinking was not clearly associated with higher or lower risk.

Table 2 shows summary hazard ratios for CVD mortality according to alcohol intake across seven NHIS years with information on lifetime drinking habits. Risk tended to be similar among never drinkers and lifetime infrequent drinkers. A modestly higher risk was observed among former drinkers only in initial analyses and was eliminated by multivariable adjustment. In contrast, multivariable adjustment had relatively little effect upon risk estimates among current drinkers. There was no higher adjusted risk among former drinkers than never drinkers for CHD mortality (adjusted hazard ratio 1.00; 95% CI, 0.89–1.11) or cerebrovascular mortality (adjusted hazard ratio 1.06; 95% CI, 0.86–1.33).

Table 3 shows the association of average alcohol intake with CVD, coronary, and stroke mortality across all NHIS years, both overall and in prespecified subgroups. In general, the association of light and especially moderate drinking with lower risk was observed to similar degrees across all subgroups, although estimates tended to be further below 1 for coronary than stroke mortality, particularly for heavy drinking. Light and moderate drinking were also inversely associated with cardiovascular deaths not coded as ischemic and cerebrovascular, with hazard ratios of 0.73 (95% CI, 0.64–0.83) and 0.71 (95% CI, 0.59–0.85), respectively.

Hazard ratios tended to be numerically closer to 1 among minority populations, although light and moderate drinking were associated with significantly lower risk even among minority participants. This contrast was more pronounced for CHD mortality, as light and moderate drinking were associated with summary hazard ratios of 0.74 (95% CI, 0.67–0.82) and 0.66 (95% CI, 0.55–0.78), respectively, among non-Hispanic whites; the corresponding hazard ratios among members of minority populations were 0.85 (95% CI, 0.67–1.08) and 0.79 (95% CI, 0.56–1.10).

We performed several sensitivity analyses to test our findings (Table 2). Additional adjustment for physical activity and diet altered the observed summary estimates to a minimal degree. When a two-year lag was incorporated to eliminate early deaths, the observed hazard ratios were attenuated but remained statistically significant for light and moderate drinking. In contrast, restriction to five years of follow-up for each survey administration to minimize misclassification of alcohol intake over time tended to increase the lower risk associated with alcohol intake. Pooled analyses yielded roughly similar estimates, albeit with substantially tighter confidence intervals. In general, we found no material differences in CVD mortality associated with light or moderate drinking in analyses stratified by follow-up prior or subsequent to folate fortification in the U.S. (data not shown).

#### Frequency and Quantity of Alcohol Consumption

Figure 2 demonstrates the joint associations of drinking frequency and quantity with cardiovascular mortality compared with abstainers. There was not a simple clear pattern, although risk was consistently higher among those who consumed 3 compared to 2 drinks per drinking day. There was no consistent association of drinking frequency with risk. These general findings were similar in analyses of CHD rather than CVD mortality and among both men and women.

#### **Binge Drinking**

There was little overall association of binge drinking with CVD mortality. In analyses of NHIS years 1988, 1991, and 1997–2000, the multivariable-adjusted hazard ratios for CVD mortality (compared with long-term abstainers) were 0.65 (95% CI, 0.51–0.81) among drinkers who reported no binge drinking, 0.66 (95% CI, 0.46–0.95) among those who endorsed binge drinking <1 day per week, and 0.68 (95% CI, 0.44–1.06) among those who reported binge drinking 1 day per week, consistent with the lower overall risk associated with drinking.

After taking usual drinking habits into account, we found limited evidence that binge drinking might attenuate the lower risk of CHD mortality associated with light and moderate drinking among women. For example, among women who did not report binge drinking, light and moderate drinking together were associated with a summary hazard ratio for CHD mortality of 0.56 (95% CI, 0.41–0.76), compared with long-term abstention; the corresponding hazard ratio among binge drinkers was 1.23 (95% CI, 0.42–3.61). However, relatively few women reported binge drinking, and the difference in risk between light/

moderate drinkers according to binge drinking was not significant (p=0.17). Power was even more limited to assess the risk of binge drinking specifically associated with stroke mortality.

#### Discussion

In this analysis of nine nationally-representative samples of the U.S. population, comprising over 10,000 cardiovascular deaths, light and moderate alcohol consumption was associated with lower cardiovascular mortality, while heavy alcohol consumption was not.

Great interest remains in whether the previously observed relationships of alcohol consumption with lower rates of CVD or mortality are biased by choice of referent category (21,22). Because former drinkers may have ceased drinking for health-related reasons, they have typically been found to have higher rates of mortality in previous studies (23,24), although studies of CVD have not always demonstrated this increase in risk (25). Even less clear is the appropriate treatment of rare drinkers, as critics have variously suggested that they be fully excluded from the referent category of long-term abstainers (5) or replace abstainers as the referent (26). Our data demonstrate that, at least for CVD mortality, risk is similar among lifelong abstainers, lifelong rare drinkers, and former drinkers, and thus that the inverse associations of light and moderate drinking with risk of CVD mortality are robust with respect to the choice of referent category.

Existing data on alcohol use and CVD derive nearly exclusively from studies of limited generalizability to the full U.S. population, although some may be representative of specific U.S. communities (27,28). As a result, our results provide some of the strongest evidence to date that the observed associations can be generalized to the U.S. population, and are not limited to intensively-monitored cohorts of volunteers. However, the NHIS excludes institutionalized individuals, who are likely to include the frailest members of society for whom even limited alcohol consumption could be detrimental.

Our results also provide insight into the consistency of the association of alcohol consumption and risk of CVD across subgroups. In general, light and moderate alcohol consumption were associated similarly with lower risk of CVD in nearly all segments of the U.S. population, including those who report themselves to be in excellent health (in whom the likelihood of residual confounding seems smallest (29)). One relative exception to this finding was race/ethnicity, where the association appeared stronger among non-Hispanic whites. Limited data suggest the apparent U- or J-shaped associations of alcohol consumption with CVD and mortality observed in other groups may not extend to African-Americans, who comprise the largest minority population in these surveys (30,31). Interestingly, alcohol consumption was associated with a numerically higher risk of MI among black Africans but a lower risk among white and mixed-race Africans in the INTERHEART Africa study (32), suggesting that findings among African-Americans may reflect known ancestral variation in alcohol dehydrogenase 1B isoforms or other alcohol-metabolizing enzymes (33). Prospective studies of alcohol consumption and CVD in minority populations are needed to clarify this issue.

The association of alcohol consumption with risk of CHD mortality tended to be stronger than that for cerebrovascular mortality in our analyses. This observation mirrors that observed for incident CVD in several other cohorts, where the associations of alcohol intake with CHD (27,34,35) have been stronger and more linear than the corresponding associations even with ischemic stroke (36–38). In part, this may reflect the particularly strong contributions of hypertension and atrial fibrillation as risk factors for ischemic stroke, both of which are positively associated with at least heavy drinking (39,40). In the NHIS,

the weaker association with cerebrovascular mortality likely also reflects some contribution of hemorrhagic stroke, which tends to be positively associated with alcohol consumption (41), perhaps related to the antiplatelet activity of alcohol (42).

Important limitations of our study warrant discussion. As with any observational study, our results may be confounded by factors for which we did not fully adjust, although the range of covariates available in the NHIS is relatively broad. Any unadjusted or incompletely unadjusted factor would need to be associated with both alcohol consumption and CVD and not strongly associated with the covariates already included in our models. Previous work in the Behavioral Risk Factor Surveillance System, another nationally representative survey, suggests that adjustment for race and education – which were included in these analyses – may be particularly important in reducing residual confounding (43).

The assessment of alcohol consumption in the NHIS occurred at a single point in time. This may have led us to underestimate the strength of associations for which alcohol consumption has acute effects on risk, although alcohol consumption is a relatively stable behavior among light- and moderate-drinking adults in the U.S. (44). It is most likely to have weakened associations of CVD with drinking frequency, where the presumed mechanisms of action include short-term, completely reversible effects on factors like thrombotic potential.

Although all data were collected prospectively using standardized, well-recognized instruments, the NHIS does not include measurement of biological factors that could be used to validate self-reported alcohol consumption (e.g., HDL-C (45)) or examine potential pathways that mediate the observed relationship (e.g., adiponectin (46)). Information on individual beverage types was also not uniformly collected, although most evidence suggests that wine, beer, and spirits are similar in their relationships with HDL-C(47) and with risk of CVD (48).

Despite the large overall sample size of our analyses, we were nonetheless limited in power in some cases, such as when stratifying participants by binge drinking and usual drinking habits simultaneously. In part, this reflects the relative rarity of binge drinking among otherwise light drinking women, but it nonetheless limits our ability to estimate the risk associated with binge drinking with precision. Similarly, power was insufficient to examine ischemic and hemorrhagic strokes separately; the overall association observed here likely reflects lower risk of more common ischemic strokes and higher risk of hemorrhagic strokes. (49)

The NHIS obtains mortality data from the National Death Index, which derives cause-ofdeath information from state vital statistics offices and ultimately from death certificates. Although this methodology matches formal nosologist evaluations reasonably well (50), death certificates are nonetheless prone to misclassify CVD as a cause of death (51). Because this type of misclassification is likely random with respect to alcohol consumption, the true strength of the associations we evaluated is apt to be greater than that observed.

In summary, in aggregate analyses of nine nationally representative surveys of the U.S. population conducted between 1987 and 2000, light and moderate alcohol consumption were inversely associated with cardiovascular mortality, with little difference in risk among lifelong abstainers, lifelong rare drinkers, and former drinkers. The magnitude of lower risk was generally strongest for CHD mortality and among non-Hispanic whites. These data bolster previous epidemiological studies that have found lower rates of incident CVD among moderate drinkers, but also provide cautionary evidence that drinking above recommended limits eliminates this risk reduction. Although randomized trials to evaluate these relationships may yet be conducted to test these issues definitively, clinicians and patients at

this time must continue to make informed, individualized, and collaborative decisions about the safety of alcohol consumption.

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#### Abbreviations

CVD	cardiovascular disease
HDL-C	high-density lipoprotein cholesterol
CHD	coronary heart disease
NHIS	National Health Interview Survey
NCHS	National Center for Health Statistics
NDI	National Death Index
ICD-9/10	Ninth and Tenth Revisions of the International Classification of Diseases
CI	confidence interval

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#### Cardiovascular Mortality According to Alcohol Intake and NHIS Year

Figure 1.

Cardiovascular mortality according to alcohol intake and NHIS year. Adjusted hazard ratios are shown relative to abstainers from alcohol. T-bars indicate 95% confidence intervals.

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#### Figure 2.

Cardiovascular mortality according to quantity and frequency of alcohol consumption. Adjusted summary hazard ratios are shown relative to abstainers from alcohol.

Year	N	<b>CVD Deaths</b>	How Assessed	Time Frame	Participants	Former Drinking	<b>Binge Drinking</b>	Adult Response Rate
1987	21396	1865	FFQ	Past year	IIA	No	No	86%
1988	42144	3478	QF	Past year	IIA	Yes	Yes	87%
1990	29527	2352	QF	Past 2 weeks1	IIA	Yes	No	83%
1991	19956	85	QF	Past year	Age 18-44 years	$\rm Yes^2$	Yes	75%
1992	11076	631	FFQ	Past year	IIF	No	No	86%
1997 <sup>3</sup>	33477	888	QF	Past year	IIA	Yes	Yes	80%
1998	29761	639	QF	Past year	All	Yes	Yes	74%
1999	28206	427	QF	Past year	IIF	Yes	Yes	70%
2000	29664	305	QF	Past year	IIF	Yes	Yes	72%
I Only inc	lividuals	who reported tha	t their use in the pa	ist 2 weeks was ty	pical of their use in	the past year were inc	luded in these analy	ses.
<sup>2</sup> Former	and lifetin	me infrequent dri	inking were combir	ned in 1991.		4		

 $^3$  The NHIS administered a similar questionnaire from 1997 through 2000.

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# Table 2

Adjusted summary hazard ratios (and 95% confidence intervals) for cardiovascular mortality among NHIS participants in 1988, 1990, 1991, and 1997–2000, according to baseline volume of alcohol consumption.

		V0	lume of Alcohol Co	nsumption		
	Never Drinker	Lifelong Infrequent Drinker	Former Drinker	Light	Moderate	Heavy
CVD Deaths <sup>1</sup>	2754	1538	1700	1131	697	354
Person-Years	305966	242849	174714	520562	238498	94815
Age-Sex-Race Adj.	1.00	$0.98\ (0.91 - 1.05)$	1.15 (1.06–1.24)	0.66 (0.56–0.77)	0.58 (0.47–0.72)	1.07 (0.94–1.22)
Multivariable Adj. <sup>2</sup>	1.00	0.95 (0.88–1.02)	1.02 (0.94–1.11)	0.69 (0.59–0.82)	0.62 (0.50-0.77)	0.95 (0.82–1.10)
MV+Activity <sup>3</sup>	1.00	0.96 (0.89–1.03)	1.02 (0.94–1.11)	0.71 (0.61–0.83)	0.65 (0.53–0.79)	0.96 (0.83–1.11)
Pooled MV HR <sup>4</sup>	1.00	0.96 (0.89–1.03)	1.02 (0.94–1.10)	0.71 (0.65–0.78)	0.70 (0.63–0.78)	0.95 (0.82–1.10)
Two-Year Lag <sup>5</sup>	1.00	$0.96\ (0.89{-}1.04)$	1.03 (0.94–1.13)	0.72 (0.62–0.84)	$0.74\ (0.63-0.86)$	1.04 (0.90–1.21)
I Hazard ratios derive I	from random-effect	ts meta-analyses of NHIS years 19	88, 1990, 1991, and	1997–2000; 1987 ar	id 1992 did not inclu	de information with which to identify former drinkers
<sup>2</sup> Multivariable-adjuste	ed hazard ratios adju	usted for age, sex, race, smoking, 1	marital status, educa	tion, region, urbaniz:	ation, smoking, BMI	, and general health status.
$\frac{3}{3}$ Analyses adjusted for	r covariates in mult	ivariable models and physical acti-	vity assessed in 1990	), 1991, and 1997–20	000.	
<sup>4</sup> Hazard ratios derive <sup>1</sup>	from pooled (rather	t than meta-analyzed) analyses of i	individual NHIS adn	ninistrations.		
5 Lagged analyses excl	lude individuals wh	to died within the first two years al	fter administration o	f the respective NHI	Š	

# Table 3

Adjusted summary hazard ratios (and 95% confidence intervals) for cardiovascular, coronary heart disease, and stroke mortality among NHIS participants in 1987, 1988, 1990–1992, and 1997–2000, according to baseline volume of alcohol consumption.

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			The second second second	
	Non-Drinkers	Light	Moderate	Heavy
Cardiovascular Deaths	7400	1804	936	530
Person-Years	856659	677069	310062	143649
Multivariable Adjusted <sup>1</sup>	1.00	$0.76\ (0.68-0.85)$	0.67 (0.59–0.77)	0.89 (0.73–1.10)
Multivariable +Activity/Diet <sup>2</sup>	1.00	0.77 (0.69–0.85)	$0.69\ (0.61-0.80)$	0.90 (0.73–1.10)
5-Year Follow-Up <sup>3</sup>	1.00	$0.71 \ (0.60 - 0.84)$	0.65 (0.55–0.77)	0.83 (0.66–1.04)
Men	1.00	0.78 (0.69–0.90)	$0.70\ (0.61{-}0.81)$	0.95 (0.76–1.19)
Women	1.00	0.74 (0.66–0.83)	$0.70\ (0.60-0.81)$	0.88 (0.72–1.07)
White	1.00	$0.75\ (0.68-0.84)$	0.67 (0.58–0.78)	0.90 (0.73–1.10)
Not White	1.00	0.80 (0.65–0.99)	$0.76\ (0.61-0.94)$	0.92 (0.70–1.21)
Age <55 Years	1.00	0.76 (0.60–0.97)	$0.62\ (0.45-0.86)$	0.91 (0.60–1.40)
Age 55 Years	1.00	0.77 (0.69–0.86)	0.71 (0.63–0.80)	0.89 (0.76–1.06)
Excellent Health	1.00	0.70 (0.56–0.88)	0.69 (0.57–0.83)	0.92 (0.66–1.28)
Not Excellent Health	1.00	0.72 (0.64–0.80)	0.61 (0.52-0.72)	0.82 (0.65–1.03)
Coronary Heart Disease Deaths	4217	1057	553	308
Coronary Mortality	1.00	0.75 (0.66–0.84)	0.67 (0.57–0.79)	$0.80\ (0.61{-}1.05)$
Stroke Deaths	1241	288	135	94
Stroke Mortality	1.00	$0.80\ (0.61{-}1.05)$	0.76 (0.58–0.99)	1.25(0.92 - 1.70)

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 $^3$ Truncated analyses limited follow-up to five years following administration of the respective NHIS.