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Meta-analysis of Marital Dissolution and Mortality: Reevaluating the Intersection of Gender and Age

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Abstract

The study of marital dissolution (i.e. divorce and separation) and mortality has long been a major topic of interest for social scientists. We conducted meta-analyses and meta-regressions on 625 mortality risk estimates from 104 studies, published between 1955 and 2011, covering 24 countries, and providing data on more than 600 million persons. The mean hazard ratio (HR) for mortality in our meta-analysis was 1.30 (95% confidence interval [CI], 1.23-1.37) among HRs adjusted for age and additional covariates. The mean HR was higher for men (HR, 1.37; 95% CI, 1.27-1.49) than for women (HR, 1.22; 95% CI: 1.13-1.32), but the difference between men and women decreases as the mean age increases. Other significant moderators of HR magnitude included sample size; being from Western Europe, Israel, the United Kingdom and former Commonwealth nations; and statistical adjustment for general health status.

Keywords

Meta-analysis; Meta-regression; Marital dissolution; Mortality; Gender; Age

Introduction

The association between marital status, health, and mortality was one of the first issues to be systematically studied by sociologists and demographers, dating back to Durkheim's classic study on suicide (Durkheim, 1951 [1897]). Over the years, numerous studies have examined this relationship, with many of them focusing on the risk of death among divorced and separated persons. The vast majority of these studies reported an increased risk of death for divorced and separated people. However, a few studies found no significant association (Burgoa et al., 1998; Goldman et al., 1995) and the magnitude of the relative risk varied

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substantially. Specifically, the wide range of mortality risks often varies by socio-demographic variables such as gender (Burgoa et al., 1998; Kolip, 2005; Kravdal, 2003; Zajacova, 2006), age (Breeze et al., 1999; Helweg-Larsen et al., 2003), geographical location (Artnik et al., 2006; Keller, 1969), and how well the study controlled for covariates (Bagiella et al., 2005; Moos et al., 1994).

Three main streams of research emerge from the literature: research into possible confounding, examinations of possible mediating mechanisms, and studies of possible moderating factors. We begin with a brief review of research on confounding factors and on potential mediators. The current trend in the literature, however, is towards identifying the role of moderating factors in the marital dissolution-mortality association (Hughes & Waite, 2009; Williams & Umberson, 2004). Since meta-analysis and meta-regression techniques are well suited for investigating these moderating factors, we focus the present paper on this latter research stream.

Research on confounding factors

First, some scholars have suggested that the association between marital dissolution and mortality may in fact be at least partly spurious. According to this line of reasoning, this association is largely the result of physical health, psychological health, economic, and behavioral factors. For example, Mastekaasa (1994) argued that low psychological wellbeing of either one or both partners may in fact be present for many years before marital dissolution and serve as one of the main causes for this dissolution. Parlin and Johnson (1977) and Rushing (1979) also argued that those who are psychologically unstable may be less able to sustain marriages. Similarly, Gottman and Levenson (1992) showed that lower physical health may predict marital dissolution. People who suffer from physical health problems are not only less likely to get married, but are also more likely to suffer from marital stress, marital problems, and eventually marital dissolution

Preexisting economic difficulties may also increase the likelihood of marital dissolution rather than being its result. Hansen (2005) for example, who conducted an eight year panel data study of unemployment and marital dissolution, found that unemployment leads to an increased risk of marital dissolution. Yet, as many other studies have previously shown, unemployment and economic adversity are also clearly associated with health problems (Alavinia & Burdorf, 2008; Bamba & Eikemo, 2009; Hammarstrom, 1994; Janlert, 1997; Jin et al., 1995; Murphy & Athanasou, 1999) and a higher risk of mortality (Ahs & Westerling, 2006; Costa & Segnan, 1987; Iversen et al., 1987; Martikainen, 1990; Moser et al., 1987; Tsai et al., 2004). Finally, a variety of behaviors which may be damaging to a person's health and increase the risk of mortality may also lead to marital tensions and higher marital dissolution rates. These include, among others, excessive substance abuse, alcoholism, smoking, reckless driving, inadequate diets, and eating disorders (Amato & Rogers, 1987; Mudar et al., 2001; Wyke & Ford, 2002; Yamaguchi & Kandel, 1997).

Research on mediation mechanisms

Previous studies suggested three predominant factors through which marital dissolution may lead to negative health consequences: social support, economic well-being, and physical and mental health. First, much of the sociological and public health literature has concluded that social support provides people with a variety of physical and mental health benefits that decrease mortality rates (e.g. Berkman, 1985; Blazer, 1982; Cornell, 1992; Litwak et al., 1989; Rosengren et al. 1993; Schaefer et al., 1981; Uchino et al., 1996). In her analysis of Japanese families, Cornell (1992) found that married women enjoy better relationships within the household and consequently lower mortality rates. Lye (1996) found that divorce weakens parent-child relations, and Kalmijn (2007) showed that when men divorce, they

tend to receive reduced social support from their children. These studies suggest that marriage provides people with important resources in the form of social support, which may serve as a buffer against health problems and mortality risks. When the marriage ends, this support system often weakens, and with it also the emotional support and regulation of health behaviors associated with being married (Ross et al., 1990; Umberson et al., 1992)

A second mediating factor examined in the marital dissolution literature is economic well-being. Both Lillard and Waite (1995) and Rogers (1995) showed that divorced and separated people die at a higher rate than married people, and that the relationship is explained partly by economic factors. Smock's (1993, 1994) findings suggested that financially women are more adversely affected by divorced. Duncan and Hoffman (1985) found that if women remarry, the negative effect of their previous economic loss is eliminated. McManus and DiPrete (2001), while acknowledging the economic disadvantages of women following divorce, showed that men also incurred a serious financial penalty from divorce due to the loss of a second income and the payment of child support. Such economic disadvantages may, with time, translate into diminished health coverage and cheaper and less healthy nutrition, which eventually lead to poorer health and higher mortality rates.

The third mediating factor emphasized in the marital dissolution literature relates to stress and mental health (Williams & Umberson, 2004). Some have argued for a *marital resource model*, suggesting that marriage increases people's psychological well-being (Cherlin et al., 1998; Gove, 1973). Kessler and Essex (1982), for example, found that married people are less susceptible to stressful events, an advantage which is lost when the marriage ends. Others suggested that marital dissolution itself may have a negative effect on mental health (Parlin & Johnson, 1977; Williams et al., 2008), which may lead to higher mortality rates. According to this *crisis model*, the stress associated with marital dissolution directly undermines health (Booth & Amaro, 1991; Williams et al., 1992).

Moderating Factors in the Marital Dissolution-Mortality Association

The contribution of meta-analysis and meta-regression techniques to our understanding of confounding and mediation mechanisms is limited. However, these methods are especially helpful in investigating moderating factors in the relationship of marital dissolution and mortality. In a recent meta-analysis of marital status (including marital dissolution) and mortality among the elderly, Manzoli et al. (2007) found that divorced and separated persons had a 16% higher risk of mortality than the risk of married persons. The inclusion criteria they used, however, were quite restrictive and their analysis leaves many questions unanswered. Their study examined only the elderly (65 and older), was limited to studies published between 1994 and 2005, and included relative risks (RRs) only if the original studies adjusted for age, sex, and additional covariates.

Additional meta-analyses are therefore needed to determine the magnitude of the association for younger age groups and to determine whether gender or geographic region become significant factors once these younger age groups are included in the analysis (both gender and region were non-significant in the Manzoli et al. study). Furthermore, Manzoli et al. only examined the effects of statistical adjustment and follow-up duration on the risk of mortality of married vs. unmarried individuals, a category in which divorced/separated, widowed, and never married individuals were joined together. Therefore, no conclusion could be drawn regarding marital dissolution and trends in mortality risk over time were not investigated.

Perhaps the most interesting finding in the analysis of Manzoli et al. (2007) was the lack of a statistically significant difference in the risk of mortality between divorced men and divorced women. This finding is especially surprising given that a host of previous studies

have established the existence of a gender difference in the risk; divorced men were often found to be at a higher risk for health problems and mortality than divorced women (e.g. Burgoa et al., 1998; Hajdu et al., 1995; Ikeda et al., 2007; Kolip, 2005; Lillard & Waite, 1995). This suggests a possible interaction effect between age and gender, where the gender differences in the risk for mortality following marital dissolution diminish at older ages. This supposition seems to be supported by the findings of a number of studies which showed that the risk of mortality following marital dissolution is higher at younger ages than it is at older ages (e.g. Hajdu et al., 1995; Kravdal, 2007). However, other researchers have suggested that because older adults often experience a pile-up effect of stressors, the negative health effects of marital dissolution should be greater for older compared to younger individuals (Ensel & Lin, 2000; Ensel et al., 1996; Williams & Umberson, 2004).

Another potentially important moderating factor is the recency of the marital dissolution. Williams and Umberson (2004) suggested that the health decline associated with marital dissolution should dissipate with time. Conversely, Hughes and Waite (2009) have found that those people who spent a greater portion of their lives divorced had more health problems. To date, very few studies have compared varying follow-up durations in order to empirically examine the moderating effects of marital dissolution recency on the risk of mortality. The few that have done so mostly found higher mortality rates when marital dissolution occurred more recently, but results were not consistent (Brockmann & Klein, 2002; Jenkinson et al., 1993). Furthermore, gender is potentially implicated with respect to the moderating effects of dissolution recency because men are more likely than women to remarry (Glick & Lin, 1987), and remarriage serves as a potential buffer against the long-term harmful effects of marital dissolution.

The potential role of cultural differences is also of interest as marital dissolution has been studied extensively across the globe. Some of the previous studies reported particularly high (two to five times higher) mortality hazard ratios following marital dissolution in Eastern European countries (Artnik et al., 2006; Dzurova, 2000; Malyutina et al., 2004), in Scandinavia (Rosengren et al., 1993; Rosengren et al., 1989; Villingshoj et al., 2006), and in Japan (Nagata et al., 2003). Yet other studies of Eastern European (Mollica et al., 2001) and Scandinavian (Nilsson et al., 2005; Nybo et al., 2003; Samuelsson & Dehlin, 1993) countries found either no association between marital dissolution and mortality, or even at times a negative association, where divorced people lived longer. While a pattern is difficult to discern, culture is clearly a potential moderator of the marital dissolution-mortality association. Theoretically, one may expect marital dissolution to have a more substantial negative impact in relatively conservative and traditional societies (such as East Asia or the Arab world), where divorce is looked upon unfavorably and the cultural stigma is stronger.

Hypotheses

The present meta-analysis contributes to the body of knowledge on marital dissolution by utilizing the heterogeneity of research settings found in the literature to assess the impact of potential moderators. Some, such as gender and age, are easier to evaluate within an individual study (although most previous studies still did not examine them). Others—such as marital dissolution recency and cultural differences—have rarely been addressed by individual studies and are much more easily examined by comparing across studies. Meta-analysis and meta-regression techniques are well-suited to this task, allowing more nuanced comparisons of various moderators. Following our theoretical discussion, we conduct tests of gender-age interactions, gender-recency interactions, geographic region, time period, and a number of specific study design characteristics. The research reviewed above suggests a number of hypotheses regarding potential moderating factors. We test five basic hypotheses:

*H*₁: The overall mortality hazard ratio of those who experienced marital dissolution will be higher than the mortality hazard ratio of married people.

*H*₂: The mortality hazard ratio associated with marital dissolution will be greater for younger age groups than for older ones.

*H*₃: Divorced men will have a higher mortality hazard ratio than divorced women. However, this difference will be most pronounced in younger ages, and in older ages the difference in risk will decline.

*H*₄: The harmful effects of marital dissolution will be most pronounced during the first few years following marital dissolution and will decrease thereafter. This decrease, however, will be more substantial for men than for women.

*H*₅: The harmful effects of marital dissolution will be more strongly felt by individuals in relatively traditional and conservative cultures.

Methods

Search strategy and coding procedures

In June 2005, we conducted a search of electronic bibliographic databases to retrieve all publications combining the concepts of psychosocial stress and all-cause mortality. We used 100 search clauses for Medline, 97 for EMBASE, 81 for CINAHL, and 20 for Web of Science (see Section 1 of Appendix for the full search algorithm used for Medline; information on the remaining search algorithms is available from authors upon request). We identified 1570 unique publications. Using these results as a base, we iteratively searched the bibliographies of eligible publications; the lists of sources citing an eligible publication; and the sources identified as “similar to” an eligible publication. We exhausted the literature after 9 iterations (the full description of this iterative search protocol is available from the authors upon request). We re-ran the electronic keyword searches in these databases and completed the search and coding stages in January 2012.

The electronic database searches were performed by a research librarian. Two authors trained in meta-analysis coding procedures independently determined publication eligibility and extracted the data from the articles. Data was entered into and publications were tracked throughout the process using basic spreadsheets (See Section 2 of Appendix for full list of variables for which data were sought). Any unpublished work encountered was considered for study inclusion. Although our search was conducted in English, we were able to locate and translate the relevant portions of 35 publications written in German, Danish, French, Spanish, Dutch, Polish, or Japanese. The most frequent reasons for study exclusion include the lack of an eligible psychosocial stress or social isolation measure, failure to report a ratio measure of mortality risk, and confounding in the dependent variable such that the outcome was not strictly all-cause mortality. Figure 1 summarizes the number of publications considered at each step of the search process. The full database contains 282 publications examining the effects of various stressful events on all-cause mortality. To evaluate coding accuracy we randomly selected and recoded 40 of these publications (including 446 point estimates). No major errors were found.

The present analysis uses the subset of articles ($n = 104$) that reported the effect of marital dissolution on all-cause mortality. 98 of these publications appeared in peer-reviewed journals; four in a book chapter; and one in an unpublished dissertation. One publication was translated from Spanish, two from Danish, and one from German in consultation with native speakers; the remaining 100 publications were in English (see Table 1).

Statistical methods and inclusion criteria

For the present analyses, a study was included if the outcome variable was all-cause mortality and a clear comparison was made between a group of people who were divorced or separated at baseline and another group who were married at baseline. No studies that used the general population were included in the analyses. In total, the 104 publications provided 625 point estimates for analysis. Cross-sectional studies were included in the analysis provided that the design of the study closely followed that of a longitudinal study. An indicator variable was created in order to examine the possible effects of this inclusion decision.

Statistical methods varied from study to study, necessitating the conversion of odds ratios, rate ratios, standardized mortality ratios, relative risks, and hazard ratios (HRs) into a common metric. All non-hazard-ratio point estimates were converted to hazard ratios—the most frequently reported type (See Section 3 of Appendix). As is standard practice, we used the standard errors reported in the publications to calculate the inverse variance weights. When not reported, standard errors were calculated using (1) confidence intervals, (2) t statistics, (3) χ^2 statistics, or (4) p-values. When upper-limit p-values were the only estimate of statistical significance available (e.g. in cases where we knew only that the p-value lay somewhere between .01 and .05), the midpoint of the upper and lower limits was used to estimate the p-value. In 222 cases (out of the 625 point estimates) no measure of statistical significance was reported and standard errors were estimated using multiple regression (see section 4 of Appendix). An indicator variable was created so analyses could be conducted both with and without data points where the standard error was estimated.

Many meta-analysts prefer to use only the most general point estimates reported in a given publication. While this strategy makes it easier to maintain independence between point estimates and makes the calculations of the inverse variance weights straight-forward, it also results in a substantial loss of information. We sought instead to maximize the number of point estimates analyzed, capturing variability both between and within each publication rather than just the former (For a similar analytic strategy see Roelfs et al., 2010; 2011; Roelfs et al., 2011; Shor et al., 2012). In cases where a given set of person-years was represented more than once, we utilized a variance adjustment procedure (See Section 5 of Appendix).

To control for time- and location-specific marital dissolution norms, we gathered data on the number of divorces per 1,000 persons, matched to the remaining data by country and baseline start year. Data were obtained primarily from the *United Nations Demographic Yearbook* from 1958, 1976, 1982, 1990, 1991, 1993-2000, 2002, 2003, 2005, and 2006. Additional data were obtained from the 1869, 1879, 1889, 1899, 1909, and 1920 Netherlands Censuses and from the 1997 *China Statistical Yearbook*.

Two measures of study quality were adopted to assess study bias. First, the 2 authors who performed the coding assigned a 3-level subjective rating to each publication. Publications were rated as low quality if they contained obvious reporting errors or applied statistical methods incorrectly (in these cases, data were coded only when sufficient information was available to extract corrected relative risk estimates). Publications were rated as high quality if models were well-specified (i.e. the correct model was used relative to the state of the art at the time of publication) and discussions and reporting of study results were detailed. Second, based on the results of a principal components factor analysis, we constructed a scale measure (continuous, range = 0 to 10) using the following: (a) the 5-year impact factor (ISI Web of Knowledge, 2009) of the journal in which the article was published (an impact factor of 1 was assigned when the impact factor was not available); and (b) the number of citations received per year since publication according to ISI Web of Knowledge. See

Section 6 of Appendix for additional information on the factor analysis. The Spearman correlation between the subjective rating and the factor-analysis-derived rating was low ($\rho = 0.173$; $p < 0.001$). The factor analysis further indicated that these two measures tapped different dimensions of quality.

Both Q -tests and I^2 measures were used to assess the presence and magnitude of heterogeneity in the data (Huedo-Medina, Sanchez-Meca, & Marin-Martinez, 2006). Q -test results from preliminary analyses revealed substantial heterogeneity across studies' effect sizes. In light of this all meta-analyses and meta-regression analyses were calculated by maximum likelihood using a random effects model. Analysis was performed with SPSS 19.0 using matrix macros provided by Lipsey and Wilson (2001). The possibility of selection and publication bias was examined using a funnel plot of the log HRs against sample size. Due to heterogeneity in the data, funnel plot asymmetry was tested using both Egger's test (Egger & Davey-Smith, 1998) and weighted least squares regressions of the log HRs on the inverse of the sample size (Moreno et al., 2009; Peters et al., 2006).

Analyses performed include meta-analyses of subgroups and multivariate meta-regression analyses. The following covariates were used in these analyses: (1) whether standard error was estimated (yes or no); (2) whether death rate was estimated (yes or no); (3) age of the publication, divided by 10; (4) age of the study, divided by 10; (5) age of the study, squared; (6) duration of the baseline period, in years; (7) years elapsed between the end of baseline and the beginning of follow-up; (8) maximum follow-up duration, in years; (9) whether a study used a longitudinal design; (10) whether study sample consisted of persons with previous stressful experiences or chronic health problems (yes or no); (11) proportion of respondents who were male; (12) mean age of sample at baseline, divided by 10; (13) age range of sample at baseline, divided by 10; (14) a series of interaction terms between gender, mean age, and follow-up duration; (15) a series of variables indicating whether gender, age, socioeconomic status, health, and other social characteristics were statistically controlled; (16) sample size, log transformed; (17) geographic region; (18) number of divorces per 1,000 population in corresponding nation-year; (19) subjective quality rating; and (20) the composite scale of study quality.

Results

Table 2 provides descriptive statistics on the 625 mortality risk estimates included in this study. Data were obtained from 104 studies published between 1955 and 2011, covering 24 countries, and representing more than 600 million people. Men and women are both well-represented in the dataset and 82.7% of the risk estimates came from study samples with mean ages greater than or equal to 40 years. The median of the maximum follow-up duration across all studies was 6.5 years. Of the HRs analyzed, Over 95% come from studies assigned a subjective quality rating of average or high; the mean 5-year impact factor was 3.59; and the mean number of citations received per year since publication was 2.07.

Table 3 presents the results of a number of meta-analyses. All analyses were stratified by the level of statistical adjustment of the risk estimate. Supporting our first hypothesis, persons who divorced or separated were significantly more likely to die than those who were married. The mean unadjusted HR was 1.51 (95% confidence interval [CI], 1.45-1.58; $n = 327$ HRs); age-adjusted HR, 1.49 (95% CI, 1.40-1.59; $n = 120$); and HR adjusted for age and additional covariates, 1.30 (95% CI, 1.23-1.37; $n = 178$). These results show that, in studies controlling for multiple covariates, marital dissolution is associated with a 30% higher risk of mortality. Table 3 also shows that the exclusion of HRs based on estimated death rates or the exclusion of HRs where the standard error was estimated does not substantively alter the mean HRs.

Subgroup Meta-analyses and Meta-regression Analyses

In the interest of presenting conservative results, from this point forward the discussion of Table 3 will focus only on HRs adjusted for age and additional covariates. Table 3 shows that marital dissolution was associated with decreased longevity for both genders. However, in accordance with our hypothesis, the magnitude of the association was greater for men (HR, 1.37; 95% CI, 1.27-1.49; n = 79 HRs) than for women (HR, 1.22; 95% CI, 1.13-1.32; n = 75). Table 4 presents the results of three meta-regression analyses, the first model consisting of all first order effects, the second model adding three interaction terms, and the third consisting of the final parsimonious model. All three models confirm that the risk of death for men who lost their spouse was substantially higher than the risk for women.

An interesting result comes from comparing groups by average age at baseline. There are no HRs adjusted for age and for additional variables for the age groups of 20-29 and 30-39. However, the unadjusted and age adjusted mean HRs for these two age groups suggest that the risk of mortality following marital dissolution in these younger age groups is higher than it is in older age groups. Indeed, in accordance with our hypothesis, marital dissolution is associated with increased mortality in almost all age groups, but there is a decrease in the magnitude of the association at older ages. In the 40 and 50 age group, divorced and separated persons had a 55% higher risk of death than married persons (HR, 1.55; 95% CI, 1.27-1.90; n = 13). The risk was still high for those aged 50 to 59 years (HR, 1.59; 95% CI, 1.34-1.89; n=13), but then decreased substantially for those aged 60 to 69 (HR, 1.36; 95% CI, 1.24-1.50; n=44), 70 to 79 (HR, 1.17; 95% CI, 1.01-1.36; n=17), and 80 or older (HR, 1.22; 95% CI, 1.14-1.31; n=90). The results of the meta-regression analysis (Model 3 of Table 4) reflect this downward trend among the latter four age groups (a 6% decrease for each additional 10 years; $p < .001$).

The impacts of gender and age on the magnitude of the HR are more complex than the meta-analyses can reveal. Models 2 (full model) and 3 (parsimonious model) of Table 4 both show a significant interaction effect between these two variables. In Model 3, the exponentiated regression coefficient for gender is 2.43 (95% CI, 2.02-2.91), 0.94 (95% CI, 0.92-0.96) for mean age, and 0.88 (95% CI, 0.85-0.91) for the interaction between gender and mean age. Taken together, these results tell us that the risk of death for men declines more rapidly with age than it does for women. By about 70 years of age, there is no longer a significant difference between men and women, with the mean HR for men falling below the mean HR for women at greater ages. However, by about 90 years of age there is little substantive difference remaining between persons who were divorced or separated and married persons. Figure 2 shows the gender-mean age interaction based on calculations from Model 3 of Table 4 (see Section 7 of Appendix for details).

The results reported in Table 3 show that there is no clear trend in HRs by follow-up duration. Model 2 of Table 4 confirms this pattern, as the exponentiated regression coefficients were not significant for follow-up duration ($p = 0.7470$) or for the gender-follow up interaction ($p = 0.0790$). The results in Table 3 show that marital dissolution was associated with substantially elevated mortality in studies conducted between 1950 and 1959 (HR, 1.62; CI, 1.18-2.22). The magnitude of the association declined in studies with a baseline year in subsequent decades. The mean HR was 1.21 (95% CI, 1.04-1.40) for studies with a baseline between 1960 and 1969, and 1.22 (95% CI, 1.14-1.31) for studies with a baseline between 1970 and 1979. However, beginning in 1980 the risk of mortality associated with marital dissolution began to increase again, with a mean HR of 1.35 (95% CI, 1.25-1.44) for studies with a baseline between 1980 and 1989 and to 1.42 (1.26-1.58) in studies with a baseline between 1990 and 1999. While these results suggest that the association between marital dissolution and mortality may be curvilinear across baseline

years (i.e. the age of the study), the results from Model 2 of Table 4 show no significant linear ($p = 0.4100$) or curvilinear ($p = 0.4180$) relationship.

With only two exceptions, the magnitude of the association between marital dissolution and mortality did not differ between regions of the world. The lack of statistical significance for most regions in Model 3 of Table 4 indicates that the magnitude of the mean HR was approximately the same in Scandinavia (the reference group); in the United States ($p = 0.6570$); in the Eastern European nations ($p = 0.3280$); in China, Japan, Taiwan, and Vietnam ($p = 0.9460$); and in Bangladesh, Brazil, and Lebanon ($p = 0.3080$). The mean HR was 15% lower in the United Kingdom, Canada, Australia, and New Zealand ($p < 0.0001$) and 21% lower in the Western European nations and Israel ($p < 0.0001$) when compared to the mean HR in Scandinavia. The results from Model 2 of Table 4 show that the number of divorces per 1,000 persons in a given nation-year also did not predict HR magnitude ($p = 0.9820$).

Table 4 shows that other significant predictors of HR magnitude include the indicator for whether the study sample consisted of persons with pre-existing health or stress conditions (a 14% decrease in the mean HR if so; $p = .0160$), the age range of the study sample (a 2% decrease in the mean HR for each additional 10 years of age range; $p = 0.0070$), and the log of the sample size ($p < 0.0001$). Table 4 also shows that HRs in studies where the standard error was estimated were somewhat lower when compared with studies where it was not estimated (a 9% decrease; $p = .0010$).

Analysis of Data Heterogeneity

The between-groups Cochrane's Q for the meta-analysis of all 625 HRs was statistically significant ($p < 0.0001$) and the I^2 statistic was high ($I^2 = 89.8$; 95% CI, 72.7-96.2), indicating that important moderating variables exist and supporting the decisions to use random effects models and conduct sub-group meta-analyses. Since the discussion of the meta-analyses focused on HRs adjusted for age and additional covariates, the corresponding heterogeneity test results were carefully examined. As shown in Table 3, the Q -tests for these sub-group meta-analyses were statistically significant for only two cases, the HRs from the 40-49.9 age group ($p = 0.0075$) and from the 2-year follow-up group ($p = 0.0080$). I^2 tests for these subgroups indicate heterogeneity was high for the 40-49.9 age group ($I^2 = 55.7$; 95% CI, 17.5-76.2) and the 2-year follow-up group ($I^2 = 63.2$; 95% CI, 21.1-82.9). The results from these three sub-group meta-analyses should be treated conservatively. However, all remaining subgroup analysis Q -tests and I^2 tests were non-significant, indicating that heterogeneity was adequately accounted for by the use of a random effects model.

Meta-regressions were also used to examine possible sources of heterogeneity in the data. The model fit statistics for Model 3 of Table 4 ($R^2 = 0.5301$; $p < 0.0001$ for Cochrane's Q of the model) indicate that this model captured a very substantial portion of the heterogeneity in the data. Nevertheless, the unexplained heterogeneity variance component for this and the other models shown in Table 4 remained highly significant (each $p < 0.001$), confirming the need to use a random effects model for all analyses.

Discussion

The results of the meta-analyses and meta-regressions show that the association between marital dissolution and mortality was not uniform across all subgroups and important moderators must be considered. In accordance with our hypothesis, among HRs adjusted for age and additional covariates, the risk of death for those who were divorced or separated was 30% higher than the risk among married persons. The association was greater for men

(an increased risk of 37%) than for women (an increase in risk of only 22%). These findings both expand and revisit the analysis of Manzoli et al. (2007), who conducted a meta-analysis of marital status and mortality among the elderly (65 and older) using studies published between 1994 and 2005. While Manzoli et al. found a 16% increase in mortality for elderly people who experienced marital dissolution, the current study found a more substantial increase in the risk (17% to 36% increase in risk, depending on the specific age group) for the elderly and an even higher risk (almost 60% increase in risk) for younger age groups. Furthermore, while Manzoli et al. found no differences across genders, our study shows that at least in younger age groups divorced and separated men have a higher risk of mortality than divorced or separated women.

Manzoli et al. argued that the typology of interpersonal ties has changed over the years, reflecting the cultural and socio-economic modifications that occurred in rapidly evolving societies (See also Henrard, 1996). By including in their analysis only studies that were published after 1994, Manzoli et al. were able to use only those studies conducted using longitudinal data but they could not evaluate the effect of these social trends on the magnitude of the risk. Our study, on the other hand, included study findings from earlier periods and found that the risk of mortality among divorced and separated persons has been relatively stable over time. Model 2 of Table 4 shows that the mean HRs did not significantly decline ($p = 0.4100$) with each additional 10 years that passed since baseline data collection.

Consistent with our theoretical suppositions, we found an interaction effect between gender and mean age. The mean HRs for both men and women declined as mean age increased, but more so for men than for women. Figure 2 shows a mean HR of 2.12 among samples of men with a mean age of 40 years. At the same mean age, the mean HR among women was 1.45. Among samples with a mean age of 70, the mean HR is 1.20 for both men and women. Among samples with a mean age of 90, the mean HR is not significantly different from 1.00 for women and is below 1.00 for men (calculated HR = 0.82). Some caution must be exercised when interpreting this finding. When the underlying death rates are very high in both the case and control groups (as is the case at older ages) ratio statistics such as the HR often lack statistical power to detect group differences. Given the magnitude of the age-effect, however, it is still likely that the observed age effects are substantively meaningful rather than merely a statistical artifact. Therefore, possible explanations for this finding need to be considered.

The pronounced gender difference in hazard ratios in the younger age groups may seem somewhat surprising given what we know about the disproportionate economic consequences of marital dissolution for men and women. A large body of research has established that women experience much larger reductions in income and standard of living following divorce than do men (e.g. Holden & Smock, 1991; McManus & DiPrete, 2001; Smock, 1994; Smock et al., 1999). Therefore one could have expected women's health, and their subsequent mortality risk, to be more adversely affected. However, consistent with previous research findings, our findings show that the mean mortality risk for middle-aged men was substantively higher than the risk for women.

One plausible explanation for the findings presented in Figure 2 might be that men experience a more dramatic decline in supportive social ties following divorce while women are better able to maintain their ties. Previous research has shown that in married couples women perform the majority of the work for maintaining parent-child relationships (Kaufman & Uhlenberg, 1998; Lye, 1996; Lye et al., 1995). Furthermore, in a recent study Kalmijn (2007) found that, in comparison to women, men experienced greater declines in social support from their children following marital dissolution. This effect was especially

strong if the marital dissolution occurred at an early stage. This finding may explain the interaction effect in Figure 2, where the difference in mean hazard ratios between men and women was much more pronounced at middle-age than at older ages. Social support appears to be a possible moderating factor in the gender differences in mortality following marital dissolution, but our data do not allow for a direct test of this hypothesis.

Another possible explanation for the findings presented in Figure 2 is that the deleterious effects of marital dissolution wane over time. Most marital dissolutions occur at a younger age, when the familial cell is relatively vulnerable (Fergusson, Horwood, & Shannon, 1984; Thornton & Rodgers, 1987). The relative instability of non-elderly couples may be the result of the pressures associated with child rearing, changes in the nuclear family structure when children leave the household, and the greater availability of alternate partners. One may thus argue that the lower excess risk of mortality following marital dissolution at older ages may simply reflect the effects of remarriage or habituation. According to this logic individuals who have lived without a partner for a long time may have adjusted to their status and found ways to compensate for the loss of social and economic support. This habituation hypothesis, however, is refuted by lack of significance for the interaction between gender and follow-up duration. In contrast to our hypothesis, the adverse effects of marital dissolution did not diminish over time for either men or women.

Finally, the analysis of differences in the risk of mortality by region failed to uncover differences in the effects of marital dissolution between most of the regions of the developed world. Mean HRs were approximately equal (ranging from 1.13 to 1.59; see Table 3) in Scandinavia; the United States; the United Kingdom, Canada, Australia, and New Zealand; East Europe; West Europe and Israel; and in China, Japan, Taiwan, and Vietnam. The magnitude of the effect only differed for the grouping of Bangladesh, Brazil, and Lebanon, though this apparent difference is not significant once other possible confounders are taken into consideration (see Table 4). Though the results from Table 4 show a significant reduction in the mean HR for the United Kingdom, Canada, Australia, and New Zealand and for West Europe and Israel, these results should be treated with caution. The lack of significance associated with the number of divorces per 1,000 people suggests that broad cultural differences are not strong predictors of differences in the relative mortality rate associated with marital dissolution.

Limitations

A major limitation of our study is one that is shared by many meta-analyses: reporting bias, or more specifically the non-reporting of non-significant findings, also known as the file-drawer effect (Berman & Parker, 2002; Egger & Davey-Smith, 1998). This tendency may lead to an over estimation of the mean HRs. Therefore, one should be especially careful in interpreting mean HRs which are relatively close to 1, even when these are significant (as is the case with some of the results in the current meta-analysis). A funnel plot of the log HRs against sample size appears somewhat asymmetric around the mean HR, suggesting the possibility of publication bias (Figure 3). Using Egger's test (Egger & Davey-Smith, 1998), there was significant evidence of publication bias ($p < 0.001$). Using Peters et al.'s test (Moreno et al., 2009; Peters et al., 2006) there was also evidence of publication bias among the unadjusted HRs ($p < 0.001$) and among HRs adjusted for age and additional covariates ($p = 0.006$), but not among HRs adjusted for age only ($p = 0.274$). The significant negative regression coefficients for the inverse of the sample size indicate that small studies with large HRs are missing from the analysis. Due to ongoing concerns with formal methods to correct for publication bias, such as the "trim and fill" method (Terrin, Schmid, Lau, & Olkin, 2003), we have not performed additional analyses using adjusted data. However, the nature of the bias is such that our results would tend to underestimate the mean HR rather

than overestimate it. The direction of the regression coefficients indicates that it is unlikely that publication bias caused an overestimation of the risk associated with marital dissolution.

A second limitation of the present study lies in the incomplete nature of the possible confounding and moderating variables included in the analysis. While the heterogeneity test results support the conclusion that a substantial portion of data variability has been accounted for, the continued need for random effects models indicates that important confounding and/or moderating factors were excluded from the models. A major goal of this research was to evaluate such factors, but additional research is clearly needed to further identify the sources of the data heterogeneity we observed.

A third limitation stems from the nature of the data. Most of the research on marital dissolution and mortality was conducted in the developed world. Relatively few publications used data from East Asia, the Middle East, and South America, and none used data from Africa. This limitation has two important consequences. First, the sample sizes in the developing world are small and any conclusions about the difference between the developed and the developing world should be made with caution. Second, since most of the results come from the developed countries they should not be extrapolated to populations in developing countries.

Conclusion

In conclusion, this study shows that marital dissolution is associated with a substantially increased risk of death among broad segments of the population. However, important moderators of this association – such as gender, age, and general health status – must be carefully considered in order to better understand this association. Future research should focus on understanding the health, socio-economic and behavioral factors through which this association is manifested, especially for younger men and women. Further research in developing countries is also needed to determine the magnitude of the risk. Since the majority of the world's population resides in developing nations, much work remains to be done.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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Research Highlights

- The risk of early mortality among the divorced and separated is 30% higher than the risk among the married
- Divorced men are more at risk than divorced women, but the differences between men and women decrease with age
- The harmful effects of marital dissolution are most pronounced during the first few years following the dissolution.

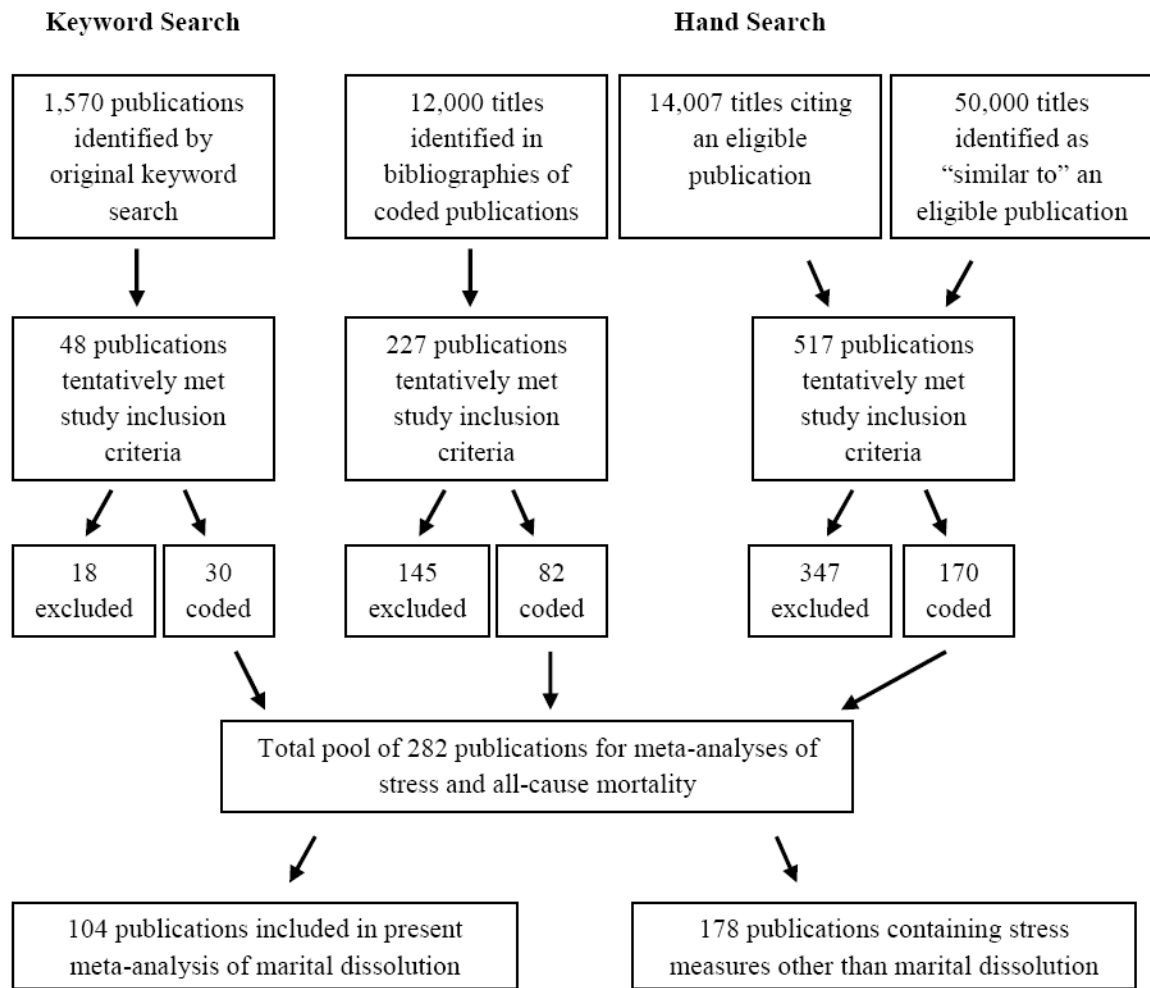


Figure 1. Flow Chart for Literature Search

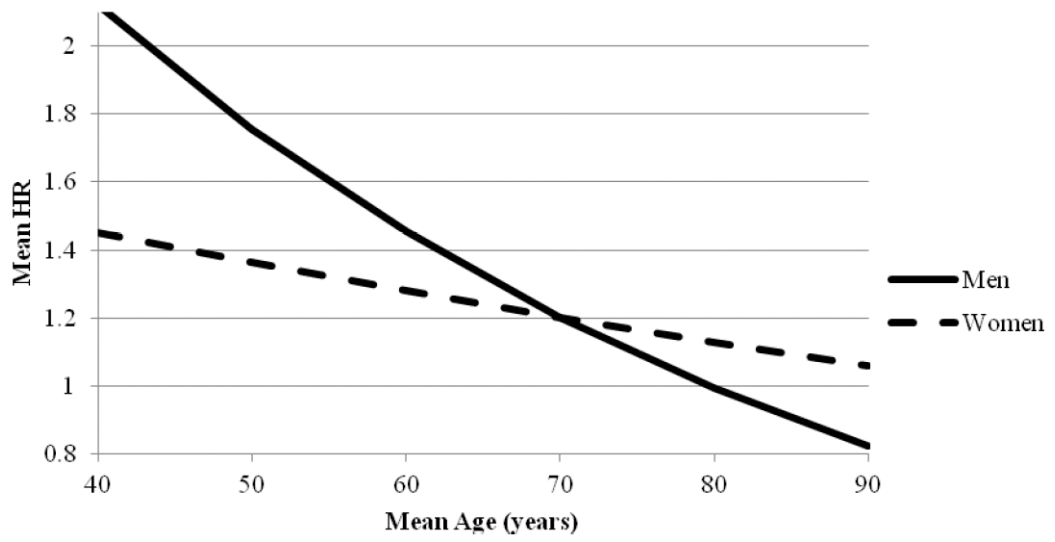


Figure 2. Mean HR by Mean Age and Gender Based on Model 3 of Table 4

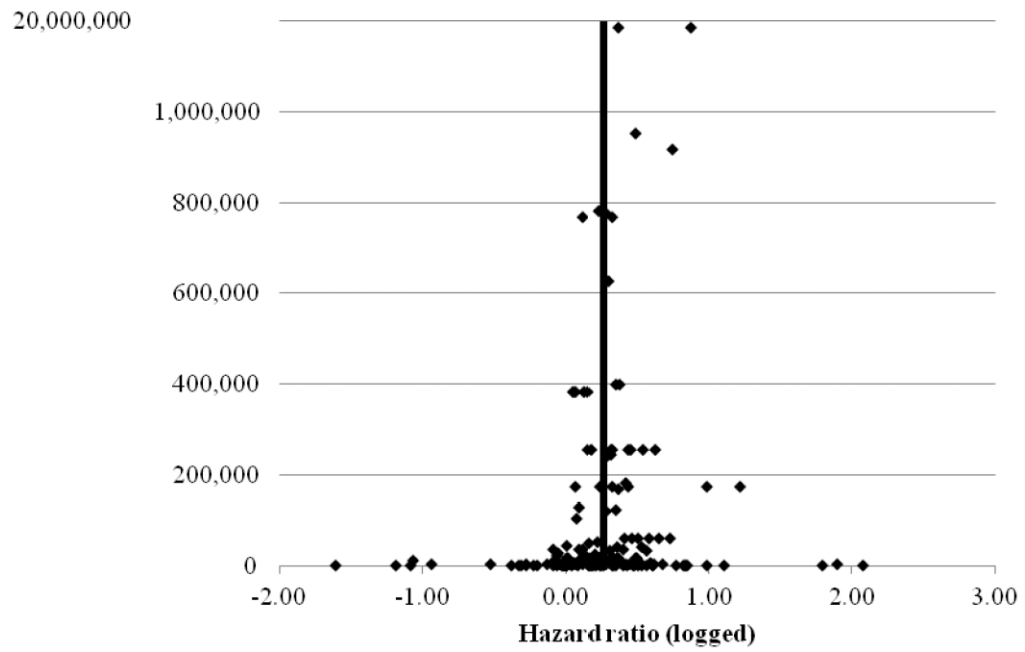


Figure 3. Funnel plot of hazard ratios (logged) versus sample size: hazard ratios statistically adjusted for age and additional covariates
Vertical line denotes the mean hazard ratio (logged) of 0.2619. Scale changes for samples 1,000,000 persons to provide greater resolution.

Table 1

Studies Included in the Analyses

Publication	Data Source	Country	Years	Sample Size	Mean HR	Number of HRs
Artnik et al. 2006	Census, 1991	Slovenia	1991-1998	1,927,000	1.88	2
Bagiella et al. 2005	EPESE	United States	1981-2002	14,456	1.19	6
Ben-Shlomo et al. 1993	Whitehall Study	United Kingdom	1967-1990	18,403	1.40	3
Berkson 1962	Census, 1950	United States	1949-1951	150,520,798	1.77	4
Berntsen 2011	Central Popul. Register	Norway	1971-2007	712,000	1.11	2
Breeze et al. 1999	OPCS-LS	United Kingdom	1971-1992	93,931	1.09	8
Brockman & Klein 2002	GSOEP	Germany	1984-1998	18,538	1.30	16
Brockmann & Klein 2004	GSOEP	Germany	1984-1998	12,484	1.24	8
Burgoa et al. 1998	Census, 1991	Spain	1991-1991	38,939,050	0.97	4
Cheung 2000	HALS	United Kingdom	1984-1997	3,378	1.19	2
Choi & Marks 2011	NSFH	United States	1987-2002	5,937	1.49	2
Comstock & Tomascia 1977	HCWC	United States	1963-1971	47,423	1.21	2
Doblhammer 2000	Census, 1981	Austria	1981-1997	1,254,153	1.35	1
Dupre et al. 2009	HRS	United States	1992-2006	8,715	1.27	1
Dzurova 2000	Census, 1990, 1995	Czech Republic	1990-1995	10,306,000	3.02	4
Ebrahim et al. 1995	BRHS	United Kingdom	1978-1990	7,735	1.46	6
Goldman & Hu 1993	Census, 1975	Japan	1975-1985	111,940,000	1.87	2
Goldman et al. 1995	NHIS-SOA	United States	1984-1990	7,478	1.00	2
Goodwin et al. 1987	NMTR	United States	1969-1982	25,706	1.27	1
Granlund et al. 2010	FilaBavi Field Laboratory	Vietnam	1999-2008	22,085	1.70	2
Grundy & Kravdal 2008	Census, 1980	Norway	1980-2003	1,530,101	1.76	6
Grundy & Tomassini 2010	ONS-LS	United Kingdom	1971-2001	41,341	1.25	2
Hajdu et al. 1995	Census, 1969-1991	Hungary, UK	1969-1991	61,786,277	1.72	40
Hank 2010	GSOEP	Germany	1984-2006	7,231	1.42	2
Hayward & Gorman 2004	NLSOM	United States	1966-1990	4,562	1.56	1
Helweg-Larsen et al. 2003	DANCOS	Denmark	1987-1999	6,693	1.82	1
Hemstrom 1996	Census, 1980	Sweden	1980-1986	1,896,626	1.55	4

Publication	Data Source	Country	Years	Sample Size	Mean HR	Number of HRs
Henretta 2007	HRS	United States	1994-2002	4,335	1.14	2
Henretta 2010	HRS	United States	1994-2002	8,918	1.26	2
Horwitz & Weber 1974	Census, 1968	Denmark	1968-1968	1,710,800	2.11	14
Horwitz & Weber 1974	Census, 1968	Denmark	1968-1968	1,710,800	1.83	6
Hurt et al. 2004	HDSS	Bangladesh	1982-1998	14,803	1.47	4
Ikedo et al. 2007	JCCSECR	Japan	1988-1999	90,064	1.26	10
Iribarren et al. 2005	CARDIA Study	United States	1985-2000	5,115	1.68	2
Iwasaki et al. 2002	Komo-Ise Study	Japan	1993-2000	11,565	1.44	4
Jenkinson et al. 1993	ASSET	United Kingdom	1986-1990	1,376	0.70	3
Johansen et al. 1996	Danish Cancer Registry	Denmark	1968-1994	7,302	1.09	4
Johnson et al. 2000	NLMS	United States	1978-1989	281,460	1.33	16
Joung et al. 1996	Census, 1986	Netherlands	1986-1990	14,572,000	1.56	2
Kalediene et al. 2007	Census, 1989, 2001	Lithuania	1989-2001	3,670,000	1.43	4
Kaplan & Kronick 2006	NHIS	United States	1989-1997	80,018	1.25	1
Keller 1969	Original Data	United States	1959-1966	706	0.92	2
Kohler & Kohler (unpublished)	Danish Twin Registry	Denmark	1968-2000	7,093	1.20	4
Kohler & Preston 2011	Census, 1992	Bulgaria	1993-1998	1,976,644	1.56	4
Kolip 2005	Census, 1999	Germany	1999-2000	45,500,000	2.08	2
Kraus & Lilienfeld 1959	Census, 1950	United States	1949-1951	150,520,798	1.81	32
Kravdal 2003	Census, 1960-1990	Norway	1960-1999	31,998	1.09	3
Kravdal 2007	Census, 1980, 1990	Norway	1980-1999	4,091,000	1.48	8
Kroenke et al. 2006	Nurses' Health Study	United States	1992-2004	2,835	0.93	2
Kumle & Lund 2000	Census, 1970	Norway	1970-1989	1,338,716	1.34	8
Laaksonen et al. 2009	Population Register	Finland	1999-2004	506,000	1.51	1
Lillard & Pamis 1996	PSID	United States	1984-1990	4,092	1.09	4
Liu 2009	NHIS-LMF	United States	1986-2002	517,314	1.13	2
Liu & Sullivan 2003	Original Data	United States	1994-1998	646	1.52	2
Makuc et al. 1990	NHANES I	United States	1971-1984	3,181	1.10	2
Malyutina et al. 2004	Novosibirsk MONICA	Russia	1984-1998	11,404	2.35	6

Publication	Data Source	Country	Years	Sample Size	Mean HR	Number of HRs
Manor et al. 1999	ILMS	Israel	1983-1992	72,527	1.16	5
Manor et al. 2000	ILMS	Israel	1983-1992	79,623	1.17	4
Martelin 1996	Census, 1970	Finland	1970-1975	4,606,307	1.25	4
Martelin et al. 1998	Census, 1970, '75, '80, '85	Finland	1970-1990	4,606,307	1.10	4
Martikainen 1990	Census, 1980	Finland	1980-1985	4,779,535	1.91	2
Martikainen 1995	Census, 1980	Finland	1980-1985	4,779,535	1.44	2
Martikainen et al. 2005	Census, 1975, 1995	Finland	1975-2000	3,165,000	1.40	24
Matthews & Gump 2002	MRFIT	United States	1982-1990	12,336	1.30	2
Mellstrom et al. 1982	Census, 1968-1978	Sweden	1968-1978	7,914,000	1.34	2
Mendes DeLeon et al 1992	KRIS	Netherlands	1972-1982	3,365	1.21	3
Metayer et al. 1996	WDCDCS	United States	1990-1993	138	0.36	3
Mollica et al. 2001	Original Data	Croatia	1996-1999	529	0.80	1
Molloy et al. 2009	Scottish Health Survey	United Kingdom	1995-2006	7,788	1.62	4
Moos et al. 1994	DVAMC	United States	1986-1991	21,139	1.36	1
Nagata et al. 2003	Takayama Study	Japan	1992-1999	3,505	5.23	3
Nilsson et al. 2005	MPP	Sweden	1974-1992	53,111	0.77	2
Norekval et al. 2010	Original Data	Norway	1992-2008	145	3.20	1
Nybo et al. 2003	Census, 1998	Denmark	1998-2000	2,249	0.81	4
Ortmeyer 1974	Census, 1960	United States	1959-1961	180,671,000	1.72	4
Patel et al. 2010	SEER Program	United States	1992-2004	7,997	1.41	1
Rahman 1993	Matlab DSS	Bangladesh	1974-1982	64,263	1.75	14
Rahman 1997	Matlab DSS	Bangladesh	1974-1982	24,889	1.93	16
Regidor et al. 2001	Census, 1996	Spain	1996-1998	3,110,121	1.71	12
Rendall et al. 2011	SIPP	United States	1984-1999	175,007	1.33	4
Rosengren et al. 1989	GMPPT	Sweden	1970-1983	9,869	2.02	5
Rosengren et al. 1993	Original Data	Sweden	1983-1991	752	2.51	2
Samuelsson & Dehlin 1993	Original Data	Sweden	1922-1990	392	0.57	2
Sheps 1961	Census, 1950	United States	1949-1951	112,354,034	1.81	36
Shkolnikov et al. 2007	Census, 2001	Lithuania	2001-2004	3,481,295	1.69	4

Publication	Data Source	Country	Years	Sample Size	Mean HR	Number of HRs
Shurtleff 1955	Census, 1940, 1950	United States	1940-1951	150,520,798	1.89	22
Shurtleff 1956	Census, 1950	United States	1950-1951	150,520,798	1.72	8
Singh & Siabpush 2001	NLMS	United States	1979-1989	301,183	1.14	6
Singh & Siabpush 2002	NLMS	United States	1979-1989	300,910	1.34	3
Smith & Waitzman 1994	NHANES I	United States	1971-1984	20,729	1.41	8
Smith & Zick 1994	PSID	United States	1968-1987	2,604	1.09	4
Sorlie et al. 1995	NLMS	United States	1979-1989	530,507	1.38	24
Spence 2006	NLS-MW	United States	1967-2001	3,258	1.26	1
Tucker et al. 1996	Terman Life-Cycle Study	United States	1950-1991	1,077	1.72	8
Tucker et al. 1997	Terman Life-Cycle Study	United States	1950-1991	1,101	1.65	2
Va et al. 2011	SMHS	China	2002-2008	52,147	2.49	2
	SWHS	China	1996-2009	74,857	1.11	2
Vallin & Nizard 1977	Census, 1968	France	1968-1968	49,780,543	1.05	30
Van Poppel & Joung 2001	Census, 1860-1960	Netherlands	1860-1969	11,486,630	1.44	2
Villingshoj et al. 2006	Danish Cancer Registry	Denmark	1991-2002	770	2.49	5
Wang et al. 2011	SEER Program	United States	1992-2008	127,753	1.23	4
Xavier et al. 2010	Epidoso EES	Brazil	1991-2003	1,533	1.26	1
Zajacova 2006	NHANES I	United States	1971-1992	12,036	1.21	2
Zajacova & Hummer 2009	NHIS	United States	1986-2002	619,320	1.38	2
Zick & Smith 1991	PSID	United States	1971-1984	1,990	1.28	4

Table 2

Distribution of mortality risk estimates (n = 625) in the analysis by selected variables

Variable	Distribution	
Publication date	1955-1959	9.9%
	1960-1969	6.7
	1970-1979	9.0
	1980-1989	1.3
	1990-1999	30.7
	2000-2011	42.4
Level of statistical adjustment		
Unadjusted	52.3%	
Adjusted for age only	19.2	
Adjusted for age and additional covariates	28.5	
Gender	Women only	44.6%
	Men only	47.4
	Both genders	8.0
Mean age of study sample at baseline	< 20	0.8%
	20 – 29.9	5.8
	30 – 39.9	10.7
	40 – 49.9	23.3
	50 – 59.9	29.3
	60 – 69.9	9.6
	70 – 79.9	14.4
	80	6.1
Baseline start year	1860 – 1939	2.9%
	1940 – 1949	12.1
	1950 – 1959	7.4
	1960 – 1969	13.9
	1970 – 1979	27.1
	1980 – 1989	20.0
Stressed population?	1990 – 2002	16.6
	Yes	5.3%
	No	94.7
Longitudinal study design?	Yes	61.9%
	No	38.1
Region		
Scandinavia	19.5%	
United States	37.9	
UK, Canada, Australia, and New Zealand	7.7	
East Europe	7.5	

Variable	Distribution	
West Continental Europe, Israel		17.8
China, Japan, Taiwan, Vietnam		4.0
Bangladesh, Brazil, Lebanon		5.6
Follow-up duration	1 st Quartile	2.5 years
	Median	6.5
	3 rd Quartile	11.25
Number of divorces per 1,000 population	1 st Quartile	1.28
	Median	2.39
	3 rd Quartile	2.75
Death rate estimated?	Yes	35.5%
	No	64.5
Standard error estimated?	Yes	40.5%
	No	59.5

Table 3

Meta-Analyses of the All-cause Mortality Risk for Divorced/Separated Persons Relative to Married Persons ^a

	Unadjusted				Adjusted for Age Only				Adjusted for Age and Additional Covariates ^b			
	HR (95% CI)	P-value	N _{HR}	Q-test P-value	HR (95% CI)	P-value	N _{HR}	Q-test P-value	HR (95% CI)	P-value	N _{HR}	Q-test P-value
All available data	1.51 (1.45, 1.58)	<0.0001	327	0.2916	1.49 (1.40, 1.59)	<0.0001	120	0.9214	1.30 (1.23, 1.37)	<0.0001	178	0.9999
Excluding point estimates converted to a HR using an estimated death rate	1.47 (1.39, 1.56)	<0.0001	219	0.1231	1.62 (1.45, 1.80)	<0.0001	62	0.8821	1.32 (1.22, 1.43)	<0.0001	122	0.9999
Excluding HRs with estimated standard error	1.51 (1.42, 1.60)	<0.0001	141	0.0064	1.45 (1.36, 1.54)	<0.0001	92	0.6927	1.33 (1.27, 1.40)	<0.0001	139	0.5787
Excluding cross-sectional studies	1.35 (1.26, 1.44)	<0.0001	125	0.0213	1.45 (1.36, 1.54)	<0.0001	84	0.8193	1.30 (1.24, 1.36)	<0.0001	178	0.7570
Gender												
Women	1.42 (1.33, 1.51)	<0.0001	152	0.9999	1.26 (1.15, 1.39)	<0.0001	52	0.9999	1.22 (1.13, 1.32)	<0.0001	75	0.9999
Men	1.67 (1.57, 1.78)	<0.0001	151	<0.0001	1.71 (1.57, 1.86)	<0.0001	66	0.7347	1.37 (1.27, 1.49)	<0.0001	79	0.9683
Region												
Scandinavia	1.60 (1.44, 1.77)	<0.0001	46	0.5995	1.43 (1.25, 1.63)	<0.0001	24	0.5773	1.35 (1.24, 1.47)	<0.0001	52	0.9264
United States	1.56 (1.46, 1.66)	<0.0001	139	0.9999	1.54 (1.34, 1.76)	<0.0001	30	0.9991	1.27 (1.17, 1.39)	<0.0001	68	0.9999
UK, Canada, Australia, New Zealand	1.52 (1.37, 1.70)	<0.0001	32	0.9987	1.28 (1.04, 1.59)	0.0229	8	0.9357	1.13 (0.91, 1.40)	0.2643	8	0.9348
East Europe	1.86 (1.68, 2.06)	<0.0001	35	0.0242	2.13 (1.77, 2.55)	<0.0001	10	0.1376	1.59 (0.97, 2.62)	0.0675	2	0.6112
West Europe, Israel	1.07 (0.96, 1.18)	0.2129	35	<0.0001	1.39 (1.27, 1.53)	<0.0001	44	0.5403	1.25 (1.10, 1.41)	0.0005	32	0.4953
China, Japan, Taiwan, Vietnam	1.35 (0.95, 1.91)	0.0969	9	0.9668	1.63 (1.10, 2.41)	0.0150	3	0.0400	1.52 (1.24, 1.85)	0.0001	13	0.2371
Bangladesh, Brazil, Lebanon	1.70 (1.30, 2.21)	0.0001	31	0.9889	1	...	0.64 (0.36, 1.13)	0.1213	3	0.0783
Age												
20 – 29.9	1.88 (1.51, 2.33)	<0.0001	22	0.9960	0	0	...
30 – 39.9	2.16 (1.78, 2.61)	<0.0001	22	0.9988	3.39 (2.31, 4.98)	<0.0001	2	0.9142	1	...
40 – 49.9	1.84 (1.67, 2.03)	<0.0001	43	0.7092	2.21 (1.61, 3.03)	<0.0001	6	0.2111	1.55 (1.27, 1.90)	<0.0001	13	0.0075
50 – 59.9	1.59 (1.47, 1.73)	<0.0001	61	0.9848	1.98 (1.63, 2.41)	<0.0001	10	0.4287	1.59 (1.34, 1.89)	<0.0001	13	0.9733
60 – 69.9	1.61 (1.46, 1.77)	<0.0001	55	0.2319	1.44 (1.26, 1.66)	<0.0001	24	0.8945	1.36 (1.24, 1.50)	<0.0001	44	0.9999
70 – 79.9	1.33 (1.18, 1.50)	<0.0001	33	0.6953	1.33 (1.09, 1.62)	0.0043	10	0.9370	1.17 (1.01, 1.36)	0.0320	17	0.9939
80	1.20 (1.11, 1.29)	<0.0001	90	<0.0001	1.39 (1.30, 1.50)	<0.0001	68	0.9177	1.22 (1.14, 1.31)	<0.0001	90	0.9919

	Unadjusted				Adjusted for Age Only				Adjusted for Age and Additional Covariates ^b			
	HR (95% CI)	P-value	N _{HR}	Q-test P-value	HR (95% CI)	P-value	N _{HR}	Q-test P-value	HR (95% CI)	P-value	N _{HR}	Q-test P-value
Baseline start year												
1860 – 1939	0.63 (0.21, 1.92)	0.4159	2	0.3876	1.39 (1.19, 1.62)	<0.0001	16	0.8573	0	...
1940 – 1949	1.74 (1.56, 1.94)	<0.0001	74	0.9999	1.68 (1.09, 2.59)	0.0182	2	0.4332	0	...
1950 – 1959	1.75 (1.53, 2.01)	<0.0001	27	0.9955	1.64 (1.38, 1.95)	<0.0001	15	0.9585	1.63 (1.13, 2.37)	0.0095	4	0.8219
1960 – 1969	1.32 (1.22, 1.43)	<0.0001	66	<0.0001	1.34 (1.04, 1.72)	0.0220	7	0.7537	1.21 (1.00, 1.45)	0.0485	14	0.9724
1970 – 1979	1.47 (1.34, 1.61)	<0.0001	71	0.9797	1.36 (1.20, 1.54)	<0.0001	36	0.9941	1.22 (1.12, 1.34)	<0.0001	62	0.9999
1980 – 1989	1.61 (1.45, 1.78)	<0.0001	46	0.1023	1.41 (1.21, 1.66)	<0.0001	17	0.9902	1.33 (1.22, 1.45)	<0.0001	35	0.9988
1990 – 1999	1.44 (1.24, 1.69)	<0.0001	35	0.9977	1.71 (1.49, 1.97)	<0.0001	23	0.0006	1.42 (1.24, 1.62)	<0.0001	36	0.1681
2000 – 2002	1.80 (1.28, 2.55)	0.0008	6	0.8764	1.69 (1.26, 2.27)	0.0005	4	0.5760	0	...
Follow-up duration												
6 months or less	1.84 (1.61, 2.09)	<0.0001	25	0.9676	0.96 (0.71, 1.30)	0.7992	4	0.4967	0	...
1 year	1.40 (1.30, 1.49)	<0.0001	73	<0.0001	3.02 (2.27, 4.01)	<0.0001	4	0.8744	1	...
2 years	1.70 (1.50, 1.91)	<0.0001	32	0.9630	1.71 (1.38, 2.10)	<0.0001	8	0.0559	1.45 (1.17, 1.79)	0.0008	8	0.0080
3 years	1.67 (1.51, 1.85)	<0.0001	78	0.9999	1.72 (1.29, 2.28)	0.0002	4	0.4696	2.16 (0.72, 6.46)	0.1685	2	0.1439
4 years	2.45 (1.28, 4.67)	0.0066	2	0.5955	1.62 (1.26, 2.09)	0.0002	5	0.6290	0.37 (0.09, 1.56)	0.1765	3	0.8380
5 years	1.30 (1.05, 1.61)	0.0165	12	0.9981	1.43 (0.99, 2.06)	0.0541	3	0.5493	1.16 (0.90, 1.51)	0.2459	5	0.9310
6 years	1	...	1.45 (1.24, 1.71)	<0.0001	16	0.9083	1.25 (1.09, 1.43)	0.0014	21	0.9960
7 years	1.93 (1.23, 3.02)	0.0042	7	0.9075	1	...	1.44 (1.17, 1.77)	0.0006	13	0.3467
8 years	1.67 (1.36, 2.07)	<0.0001	31	0.9846	1	...	0.83 (0.55, 1.26)	0.3916	4	0.0781
9 years	0.95 (0.68, 1.33)	0.7728	3	0.9375	1.31 (0.82, 2.08)	0.2577	3	0.1159	1.29 (1.10, 1.52)	0.0020	17	0.9999
10 years	1.27 (0.86, 1.85)	0.2266	8	0.9998	1.40 (1.25, 1.57)	<0.0001	27	0.9796	1.24 (1.02, 1.51)	0.0312	9	0.5542
11 years	1.44 (1.08, 1.91)	0.0121	8	0.9394	1.34 (1.05, 1.70)	0.0166	13	0.9999	1.33 (1.03, 1.70)	0.0270	17	0.9222
12 years	0	...	1.37 (0.89, 2.10)	0.1503	2	0.3414	1.32 (1.08, 1.61)	0.0066	11	0.9915
13 years	1.29 (0.78, 2.14)	0.3155	2	0.8107	1.46 (0.98, 2.19)	0.0631	3	0.6348	1.43 (1.02, 2.01)	0.0356	4	0.6137
14 years	1.32 (1.04, 1.67)	0.0225	10	0.8015	1.28 (0.96, 1.72)	0.0939	7	0.7138	1.85 (1.36, 2.53)	0.0001	6	0.7085
15 years	0	...	1.19 (0.71, 1.98)	0.5067	2	0.4932	1.07 (0.89, 1.30)	0.4697	15	0.9278
16-20 years	1.57 (1.33, 1.86)	<0.0001	18	0.0003	1.41 (0.91, 2.18)	0.1279	3	0.5705	1.51 (1.26, 1.81)	<0.0001	11	0.9597
21-25 years	1.40 (0.80, 2.44)	0.2405	2	0.8241	1.22 (0.94, 1.58)	0.1427	5	0.9341	1.39 (1.19, 1.62)	<0.0001	14	0.7103

	Unadjusted			Adjusted for Age Only			Adjusted for Age and Additional Covariates ^b				
	HR (95% CI)	P-value	N _{HR}	HR (95% CI)	P-value	N _{HR}	HR (95% CI)	P-value	N _{HR}	Q-test P-value	Q-test P-value
More than 25 years	1.06 (0.68, 1.67)	0.7888	6	1.41 (1.11, 1.79)	0.0042	10	1.37 (1.04, 1.80)	0.0240	6	0.7444	0.6711

^a All meta-analyses calculated by maximum likelihood using a random effects model. Numbers reported are the mean hazard ratio (HR), 95% confidence interval (CI), number of hazard ratios included in the mean HR calculation (N_{HR}), and the P-value from Cochran's Q-test for data heterogeneity. Ellipses indicated when n = 1 and meaningful mean HR could not be calculated.

^b The number and type of covariates varies between studies

Table 4

Multivariate Meta-Regression Analyses Predicting the Magnitude of the Effect of Marital Dissolution on Mortality ^a

Variable	Model 1: All Predictors Except Interaction Terms	Model 2: All Predictors Including Interaction Terms	Model 3: Parsimonious Model <i>b</i>
Constant	2.46 (1.59, 3.80) [p<0.001]	1.77 (1.15, 2.70) [p=0.009]	1.67 (1.36, 2.04) [p<0.001]
Standard error estimated (0 = No; 1 = Yes)	0.90 (0.83, 0.97) [p=0.010]	0.90 (0.83, 0.97) [p=0.005]	0.91 (0.86, 0.96) [p=0.001]
Death rate estimated (0 = No; 1 = Yes)	1.03 (0.96, 1.10) [p=0.384]	1.04 (0.97, 1.10) [p=0.283]	...
Publication Age (in 10 year increments)	1.01 (0.98, 1.05) [p=0.485]	1.02 (0.98, 1.06) [p=0.334]	...
Study age (in 10 year increments)	1.00 (0.98, 1.02) [p=0.744]	0.98 (0.93, 1.03) [p=0.410]	...
Study age * Study age	...	1.00 (1.00, 1.00) [p=0.418]	...
Baseline length (years)	1.00 (0.99, 1.00) [p=0.231]	1.00 (0.99, 1.00) [p=0.359]	...
Years between baseline and start of follow-up	1.00 (0.96, 1.04) [p=0.958]	1.00 (0.96, 1.03) [p=0.883]	...
Follow-up duration (years)	1.00 (0.99, 1.00) [p=0.646]	1.00 (0.99, 1.01) [p=0.747]	...
Longitudinal study design (0 = No; 1 = Yes)	1.02 (0.89, 1.17) [p=0.792]	1.02 (0.90, 1.17) [p=0.724]	...
Stressed population (0 = No; 1 = Yes)	0.86 (0.73, 1.00) [p=0.056]	0.86 (0.74, 1.00) [p=0.057]	0.86 (0.76, 0.97) [p=0.016]
Gender (Proportion of sample that is male; Range: 0-1)	1.17 (1.11, 1.24) [p<0.001]	2.41 (2.00, 2.90) [p<0.001]	2.43 (2.02, 2.91) [p<0.001]
Mean age (in 10 year increments)	0.88 (0.86, 0.90) [p<0.001]	0.93 (0.91, 0.96) [p<0.001]	0.94 (0.92, 0.96) [p<0.001]
Age range (in 10 year increments)	0.98 (0.96, 1.00) [p=0.021]	0.98 (0.96, 1.00) [p=0.011]	0.98 (0.97, 1.00) [p=0.007]
Interactions			
Gender * Follow-up duration	...	1.00 (0.99, 1.01) [p=0.790]	...
Gender * Mean age	...	0.88 (0.85, 0.91) [p<0.001]	0.88 (0.85, 0.91) [p<0.001]
Controls (0 = No; 1 = Yes)			
Gender	0.93 (0.81, 1.08) [p=0.353]	0.93 (0.82, 1.07) [p=0.323]	...
Age	1.03 (0.96, 1.12) [p=0.388]	1.03 (0.96, 1.10) [p=0.487]	...
Other demographics	1.01 (0.91, 1.13) [p=0.802]	1.02 (0.92, 1.13) [p=0.755]	...
SES	0.96 (0.85, 1.08) [p=0.470]	0.96 (0.86, 1.07) [p=0.439]	...
General health status	0.90 (0.79, 1.03) [p=0.121]	0.91 (0.80, 1.03) [p=0.125]	0.89 (0.81, 0.97) [p=0.010]
Smoking/Other health behaviors	0.96 (0.80, 1.16) [p=0.691]	0.94 (0.79, 1.12) [p=0.517]	...
Chronic condition	1.11 (0.93, 1.34) [p=0.255]	1.11 (0.94, 1.32) [p=0.230]	...
Psychiatric characteristics	1.08 (0.66, 1.75) [p=0.768]	1.12 (0.70, 1.78) [p=0.644]	...
Social ties	0.96 (0.86, 1.07) [p=0.495]	0.97 (0.88, 1.07) [p=0.558]	...
Previous stress	1.01 (0.90, 1.15) [p=0.819]	1.02 (0.91, 1.14) [p=0.782]	...
Log of sample size	1.02 (1.00, 1.04) [p=0.022]	1.02 (1.01, 1.04) [p=0.006]	1.02 (1.01, 1.04) [p<0.001]
Regions			

Variable	Model 1: All Predictors Except Interaction Terms	Model 2: All Predictors Including Interaction Terms	Model 3: Parsimonious Model ^b
Scandinavia	Reference	Reference	Reference
United States	1.03 (0.89, 1.18) [p=0.705]	1.03 (0.91, 1.18) [p=0.618]	1.02 (0.95, 1.09) [p=0.657]
UK, Canada, Australia, New Zealand	0.88 (0.79, 0.98) [p=0.020]	0.88 (0.79, 0.97) [p=0.009]	0.85 (0.78, 0.93) [p<0.001]
East Europe	1.08 (0.95, 1.23) [p=0.221]	1.06 (0.94, 1.19) [p=0.376]	1.04 (0.96, 1.14) [p=0.328]
West Europe, Israel	0.78 (0.71, 0.85) [p<0.001]	0.78 (0.71, 0.85) [p<0.001]	0.79 (0.74, 0.85) [p<0.001]
China, Japan, Taiwan, Vietnam	0.96 (0.79, 1.16) [p=0.674]	0.97 (0.80, 1.16) [p=0.704]	1.01 (0.86, 1.18) [p=0.946]
Bangladesh, Brazil, Lebanon	1.05 (0.76, 1.45) [p=0.764]	1.13 (0.83, 1.53) [p=0.444]	1.16 (0.87, 1.55) [p=0.308]
Number of divorces per 1,000 population	1.00 (0.95, 1.05) [p=0.994]	1.00 (0.95, 1.05) [p=0.982]	...
Subjective quality assessment (Range: 1-3)	1.02 (0.95, 1.10) [p=0.576]	1.02 (0.95, 1.10) [p=0.573]	...
Scale measure of study quality (Range: 0-10)	0.98 (0.95, 1.01) [p=0.125]	0.97 (0.95, 1.00) [p=0.074]	...
<i>R</i> ²	.4714	.5474	.5301
Variance Component	.0471 [p<.001]	.0374 [p<.001]	.0386 [p<.001]

^a All meta-regressions calculated by maximum likelihood using a random effects model (n=625). Number reported is the exponentiated regression coefficient (95% confidence interval) [p-value]. Ellipses indicate when a variable was not entered into a model.

^b Obtained using backwards elimination, p>.10 to exit.