

### NIH Public Access

Author Manuscript

Int J Cancer. Author manuscript; available in PMC 2014 February 18

Published in final edited form as: *Int J Cancer*. 2008 May 1; 122(9): 2062–2070. doi:10.1002/ijc.23344.

## Non-Hodgkin lymphoma and Obesity: a pooled analysis from the InterLymph consortium

Eleanor V. Willett<sup>1</sup>, Lindsay M. Morton<sup>2</sup>, Patricia Hartge<sup>2</sup>, Nikolaus Becker<sup>3</sup>, Leslie Bernstein<sup>4</sup>, Paolo Boffetta<sup>5</sup>, Paige Bracci<sup>6</sup>, James Cerhan<sup>7</sup>, Brian C.-H. Chiu<sup>8</sup>, Pierluigi Cocco<sup>9</sup>, Luigino Dal Maso<sup>10</sup>, Scott Davis<sup>11</sup>, Silvia De Sanjose<sup>12</sup>, Karin Ekstrom Smedby<sup>13</sup>, Maria Grazia Ennas<sup>14</sup>, Lenka Foretova<sup>15</sup>, Elizabeth A. Holly<sup>6</sup>, Carlo La Vecchia<sup>16</sup>, Keitaro Matsuo<sup>17</sup>, Marc Maynadie<sup>18</sup>, Mads Melbye<sup>19</sup>, Eva Negri<sup>20</sup>, Alexandra Nieters<sup>3</sup>, Richard Severson<sup>21</sup>, Susan L. Slager<sup>7</sup>, John J. Spinelli<sup>22</sup>, Anthony Staines<sup>23</sup>, Renato Talamini<sup>10</sup>, Martine Vornanen<sup>24</sup>, Dennis D. Weisenburger<sup>25</sup>, and Eve Roman<sup>1</sup> for the Interlymph Consortium

<sup>1</sup>Epidemiology and Genetics Unit, Department of Health Sciences, Seebohm Rowntree Building, University of York, YO10 5DD, UK <sup>2</sup>Division of Cancer Epidemiology and Genetics, National Cancer Institute, NIH, DHHS, Rockville, Maryland <sup>3</sup>Division of Clinical Epidemiology, German Cancer Research Centre, Heidelberg, Germany <sup>4</sup>Cancer Surveillance Program, 1441 Eastlake Avenue, Los Angeles, California <sup>5</sup>International Agency for Research on Cancer, Lyon, France <sup>6</sup>Department of Epidemiology and Biostatistics, University of California at San Francisco, San Francisco, California <sup>7</sup>Department of Health Sciences Research, Mayo Clinic College of Medicine, Rochester, Minnesota <sup>8</sup>Department of Preventive Medicine, Northwestern University Medical School, Chicago, Illinois <sup>9</sup>Department of Public Health, Occupational Health Section, University of Cagliari, Italy <sup>10</sup>Epidemiology and Biostatistics Unit, Aviano Cancer Centre, Aviano, Italy <sup>11</sup>Program in Epidemiology, Division of Public Health Sciences, Fred Hutchinson Cancer Research Center & Department of Epidemiology, School of Public Health and Community Medicine, University of Washington, Seattle, Washington <sup>12</sup>Epidemiology and Cancer Registry Unit, Catalan Institute of Oncology, Barcelona, Spain <sup>13</sup>Department of Medicine, Clinical Epidemiology Unit, Karolinska University Hospital, Stockholm, Sweden <sup>14</sup>Department of Cytomorphology, University of Cagliari, Italy <sup>15</sup>Cancer Epidemiology and Genetics, Masaryk Memorial Cancer Institute, Brno, Czech Republic <sup>16</sup>Istituto di Ricerche Farmacologiche "Mario Negri" and Istituto di Statistica Medica e Biometria, Universitá degli Studi di Milano, Milan, Italy <sup>17</sup>Aichi Cancer Center, Division of Epidemiology and Prevention, Nagoya, Japan <sup>18</sup>Registre des Hemopathies Malignes de Cote d'Or. U de Bourgogne EA4184. Dijon. France <sup>19</sup>Department of Epidemiology Research, Statens Serum Institut, Copenhagen, Denmark <sup>20</sup>Istituto di Ricerche Farmacologiche Mario Negri, 20156 Milan, Italy <sup>21</sup>Karmanos Cancer Institute and Department of Family Medicine, Wayne State University, Detroit, Michigan <sup>22</sup>Cancer Control Research Program, BC Cancer Agency, Vancouver, British Columbia, Canada <sup>23</sup>School of Public Health, Public Health University College, Dublin, Ireland <sup>24</sup>Pathology Tampere University Hospital, Tampere, Finland <sup>25</sup>Department of Pathology and Microbiology, University of Nebraska Medical Center, Omaha, Nebraska

#### Abstract

Nutritional status is known to alter immune function, a suspected risk factor for non-Hodgkin lymphoma (NHL). To investigate whether long-term over, or under, nutrition is associated with

Author for correspondence: Eleanor Willett, Epidemiology and Genetics Unit, Department of Health Sciences, Seebohm Rowntree Building, University of York, YO10 5DD, United Kingdom, Tel: 44(0)190 432 1892, Fax: 44(0)190 432 1899, eleanor.willett@egu.york.ac.uk.

NHL self-reported anthropometric data on weight and height from over 10000 cases of NHL and 16000 controls were pooled across 18 case-control studies identified through the International Lymphoma Epidemiology Consortium. Study-specific odds ratios (OR) were estimated using logistic regression and combined using a random-effects model. Severe obesity, defined as BMI of 40 kg m<sup>-2</sup> or more, was not associated with NHL overall (pooled OR=1.00, 95% confidence interval (CI) 0.70-1.41) or the majority of NHL subtypes. An excess was however observed for diffuse large B-cell lymphoma (pooled OR=1.80, 95% CI 1.24-2.62), although not all studyspecific ORs were raised. Among the overweight (BMI  $25-29.9 \text{ kg m}^{-2}$ ) and obese (BMI 30-39.9kg  $m^{-2}$ ), associations were elevated in some studies and decreased in others, while no association was observed among the underweight (BMI<18.5 kg  $m^{-2}$ ). There was little suggestion of increasing ORs for NHL or its subtypes with every 5 kg  $m^{-2}$  rise in BMI above 18.5 kg  $m^{-2}$ . BMI components height and weight were also examined, and the tallest men, but not women, were at marginally increased risk (pooled OR=1.19, 95% CI 1.06-1.34). In summary, whilst we conclude that there is no evidence to support the hypothesis that obesity is a determinant of all types of NHL combined, the association between severe obesity and diffuse large B-cell lymphoma may warrant further investigation.

#### Keywords

non-Hodgkin lymphoma; lymphoma; body mass index; weight; height; epidemiology

#### Introduction

Non-Hodgkin lymphomas (NHL) can arise following rare inherited disorders of the immune system, long-term immunosuppressive drug therapy and viral infections such as human immunodeficiency virus (HIV) and Epstein-Barr virus (EBV). In such instances, severe immunosuppression resulting from exposure usually leads to the development of specific NHL subtypes. For the majority of NHL however, the cause remains unknown but it is suspected that factors which affect the immune system are involved. In particular, it has been suggested that the degree of adiposity might be important since over (as well as under) nutrition can alter immune function (1;2). However, while several epidemiological studies have reported associations between excess weight and NHL (3–15) the evidence is far from conclusive (16–28). Here we present a pooled analysis of self-reported height and weight on over 10000 NHL cases and 16000 controls from 18 case-control studies identified through the International Lymphoma Epidemiology Consortium (InterLymph: http://epi.grants.cancer.gov/InterLymph/).

#### Materials and Methods

Through the InterLymph forum, 18 case-control studies of NHL with anthropometric data collected across 13 countries in parts of North America, Europe, and Japan between 1983 and 2004 were identified. Study designs are briefly outlined in Table 1, and more details are published elsewhere (3;8;14;18;24;29–36). Cases were identified using rapid ascertainment techniques, while controls were randomly selected from population registers (8 studies), outpatient clinics (3 studies) or inpatients (7 studies) hospitalised for a variety of non-neoplastic conditions such as circulatory, digestive or respiratory problems, or with traumatic or non-traumatic orthopaedic conditions. The appropriate ethical committees' approval was granted for each study and informed consent was given by all participants.

NHL diagnoses were pathologically confirmed and subsequently coded to the World Health Organisation (WHO) classification (37) (15 studies), the REAL classification (the 1999– 2002 Italian study), or Working Formulation (North Italy and UCSF). Cases with HIV were

excluded. Diagnostic codes from the different studies were combined as previously described (38). The analysis here considers specific B-cell subtypes of NHL (diffuse large B-cell lymphoma: ICDO3 codes 9679/3, 9680/3, 9684/3; follicular lymphoma: 9690/3, 9691/3, 9695/3, 9698/3; chronic lymphocytic leukaemia/small lymphocytic lymphoma: 9670/3, 9823/3; marginal zone lymphoma: 9689/3, 9699/3; mantle cell lymphoma: 9673/3; Burkitt lymphoma: 9687/3, 9826/3; and other unspecified B-cell lymphoma: 9671/3, 9728/3), and T-cell lymphomas as a whole (9700/3, 9701/3, 9702/3, 9705/3, 9708/3, 9709/3, 9714/3, 9716/3, 9717/3, 9718/3, 9719/3, 9729/3, 9827/3) as well as NHL in total (defined by the above ICDO3 codes and 9591/3, 9675/3, and 9727/3).

In all studies, information on anthropometrics, demographics, lifestyle, occupations and medical histories were collected by in-person or telephone interviews. For the purposes of the present analyses, anonymised data were provided and checked for inconsistencies before coding uniformly. Within each study, height in metres was categorised using sex-specific quintiles of the height distribution among controls, and data were then combined across studies to reflect the relative position, rather than the absolute value, of this variable. In the statistical analysis, the referent category for height was taken as the 3<sup>rd</sup> quintile, since this central group contains the median and has the narrowest range. Usual adult weight was requested in 10 studies (Nebraska, UCSF, SCALE and EpiLymph studies). Elsewhere different questions were used (weight at diagnosis/interview (HERPACC1, HERPACC2); one year (NCI-SEER, British Columbia, North Italy, Italy); two years (Mayo Phase 1); or five years (UK) prior to diagnosis/interview).

For the pooled analysis, body mass index (BMI) was computed by dividing weight in kilograms by the square of height in metres where each study's weight variable was considered at the closest time point prior to diagnosis/interview, or else the usual adult weight. BMI was grouped using the World Health Organisation categories of underweight (<18.5 kg m<sup>-2</sup>), normal (18.5–24.99 kg m<sup>-2</sup>), grade 1 overweight (25–29.99 kg m<sup>-2</sup>), grade 2 obese (30–39.99 kg m<sup>-2</sup>) and grade 3 obese (40 kg m<sup>-2</sup> or more) (39). For a person 1.7 m (5' 7") tall, these cut-off points relate to weights of 53 kg (118 lb), 72 kg (159 lb), 87 kg (191 lb), and 116 kg (255 lb) respectively. Socioeconomic status was defined by the level of education attained, except in British Columbia and the UK where self-reported household income and a census-based household deprivation indicator were used respectively; and no socioeconomic status information was collected in the Japanese studies (HERPACC1 and 2).

Statistical analysis followed similar methods to those employed in previous InterLymph pooling projects (40–44). Firstly, individual data were combined in an unconditional logistic regression model adjusting for study, age, sex, and race. To test for between-study heterogeneity, this model was compared using the likelihood ratio test with the model that included an additional term for interaction between the anthropometric variable and a variable indexing the studies. Heterogeneity was assumed to be present when the likelihood ratio test yielded a p-value<0.05. This flexible approach utilises all data and provides one statistic to test for heterogeneity. Where the likelihood ratio test was not statistically significant, the pooled adjusted OR and 95% CI computed from all individual data in an unconditional logistic regression model are presented.

Between-study heterogeneity was further examined among risk estimates at each category of the anthropometric variables. Study-specific odds ratios (OR) and 95% confidence intervals (CI) adjusted for sex, age, and race were computed using unconditional logistic regression (45). For each category of height or BMI, the study-specific ORs were combined using a random effects meta-analysis to produce a combined OR and corresponding 95% CI. The extent of heterogeneity for each category was indicated by Cochran's *Q*-statistic which was

considered statistically significant when p<0.10. The  $I^2$ -statistic was also reported to describe the percentage of total variation in the study-specific ORs which was due to heterogeneity (46).

Since the ORs were diverse across studies, a variety of approaches were applied to explore heterogeneity (47). To assess relative obesity within study populations rather than the absolute value, BMI was grouped into quintiles based on the control distributions within each study before combining the relative quintile groupings across studies; these analyses are not presented here since their findings were similar to those reported. Sensitivity analyses using various stratifications and subsets of data were also conducted (48). Study-specific ORs were combined by continent, study design and time period (corresponding to the original lymphoma classification used) as well as by level of participation. Given that the study-specific associations with BMI were heterogeneous in all analyses, forest plots with ORs pooled by continent were judged to be the most informative. Pooled ORs stratified by study design are also presented.

Within studies, analyses were performed separately for men, women, Caucasian subjects and persons aged 18 to 65. The resulting study-specific ORs were combined in a random-effects meta-analysis to examine heterogeneity. Potential confounding factors, such as smoking, alcohol and socioeconomic status, were assessed by comparing study-specific regression models with and without the confounding factor using the likelihood ratio test. A factor was considered a confounder when the likelihood ratio test was significant and the adjusted OR changed by more than 10%. Continuous variables corresponding to 10 cm increases in height and 5 kg m<sup>-2</sup> increases in BMI were created to assess trends. All analyses were conducted using Stata (49).

#### Results

The pooled dataset from the 18 case-control studies comprised anthropometric information from 10453 cases of NHL and 16507 controls. Most cases (85%) were diagnosed with a B-cell lymphoma, 5% with a T-cell lymphoma and for 11%, immunophenotype was not known. The three most common NHL subtypes were diffuse large B-cell lymphoma (DLBCL) (32%), follicular lymphoma (FL) (22%) and chronic lymphocytic lymphoma/ small lymphocytic lymphoma (CLL/SLL) (16%). A slightly higher proportion of cases were men (57%), 90% of all cases were Caucasian and the median age was 60 years. Cases tended to be older in age, of white race and of lower socioeconomic status than controls (data not shown).

Height distributions among male and female controls varied by study; for both sexes, the median height was highest in the American studies, generally decreased from Northern to Southern Europe, and was lowest in the two Japanese studies (data not shown). Among men, compared to the third quintile the odds ratio was increased in the highest quintile (OR=1.19, 95% CI 1.06–1.34), but was close to one in the lowest two quintiles (Supplementary Table 1). When examining trend within studies, no consistent population pattern emerged; most studies showed no evidence of a trend with 10 cm increases in height, six a significant positive trend, and two a significant negative trend (data not shown). Similar patterns were observed for the majority of NHL subtypes. Little association between height and NHL, or its subtypes, was observed among women (Supplementary Table 1).

Figure 1 gives the distribution of BMI among controls by study. Like height, studies conducted in the US had the greatest median BMI, and Japan the lowest. When BMI was classified using WHO categories, associations between BMI and NHL were heterogeneous between studies (likelihood ratio test:  $\chi^2$ =139.1, p<0.0001). Study-specific ORs showed that

the heterogeneity was most marked in Grade 1 overweight, where ORs ranged from 0.50 (95% CI 0.34-0.74) in EpiLymph Italy to 1.70 (95% CI 1.02-2.84) in EpiLymph Ireland and Grade 2 obese (ranging from OR=0.42, 95% CI 0.24–0.74 in EpiLymph Italy to OR=1.78, 95% CI 1.36–2.32 in UCSF) (Figures 2(b) and (c)). In the underweight and Grade 3 obese categories, where the numbers of subjects were small, ORs were also diverse (ranging from OR=0.27, 95% CI 0.03-2.34 in EpiLymph Ireland to OR=3.14, 95% CI 0.41-23.9 in EpiLymph Finland; and from OR=0.19, 95% CI 0.02-1.58 in EpiLymph Germany to OR=4.23, 95% CI 1.51–11.9 in UK, respectively) (Figures 2(a) and (d)). Trends with a 5 kg m<sup>-2</sup> increase in BMI above 18.5 kg m<sup>-2</sup> were significantly increased in two studies, significantly decreased in four studies and showed little effect in the remaining studies (Figure 3). ORs were pooled across North America, Northern Europe, Southern Europe and Japan. In North America, a homogeneous increased OR was suggested for Grade 1 overweight (Figure 2(b)) but no effect was found among Grade 3 obese (Figure 2(d)), and with the exception of the Californian study (UCSF), no significant positive trends were observed (Figure 3). Heterogeneity was still evident when the analyses were restricted to population-based studies conducted in the period 1998–2005; to those designed to code to the WHO classification; or to those where control participation rates were 70% or more. Similarly study-specific ORs were heterogeneous among men or women; subjects aged 18 to 65; or Caucasian subjects (data not shown).

Statistically significant between-study heterogeneity was also present for the three most common NHL subtypes (likelihood ratio tests for WHO BMI and DLBCL:  $\chi^2$ =104.2, p=0.002; FL:  $\chi^2$ =82.7, p=0.003; CLL/SLL:  $\chi^2$ =58.7, p=0.04). For these three subtypes, as for NHL as a whole, study-specific ORs varied around one in all WHO BMI groups, with tests for heterogeneity in the two-stage random effects model being significant among Grade 1 overweight and Grade 2 obese (DLBCL: Supplementary Figures 1(a)-(d); FL: Supplementary Figures 3(a)–(d); CLL/SLL: Supplementary Figures 4(a)–(d)). In the underweight and Grade 3 obese groups, the meta-analyses generally suggested that ORs were more homogeneous and the combined risk estimates were not significantly different from one. The pooled OR for DLBCL among Grade 3 obese was increased (OR=1.80, 95% CI 1.24–2.62, Q=16.7, p=0.40,  $I^2=4.4\%$ ), being elevated in North America and Northern Europe, but as with all analyses in this BMI group, study-specific risk estimates were diverse, based on small numbers of subjects, and with wide and overlapping confidence intervals (Supplementary Figure 1(d)). Like NHL as a whole, a 5 kg m<sup>-2</sup> increase in BMI did not consistently increase the risk of DLBCL (Supplementary Figure 2) or the other subtypes (data not shown). ORs for the rarer B-cell lymphomas and T-cell lymphoma were mostly not significantly different between studies, probably due to the small number of cases, and there was little suggestion of associations between these NHL subtypes and BMI (Supplementary Table 2).

Pooling data from studies with the highest WHO BMI prevalences of overweight/obese controls (EpiLymph Czech Republic, Nebraska, Mayo Phase 1, EpiLymph Italy, EpiLymph Germany, Italy-Aviano and Naples, and EpiLymph Finland) gave more homogeneous ORs (likelihood ratio test:  $\chi^2$ =32.3, p=0.12). Within this subset of seven studies, there was still little evidence that higher than average BMI increases the risk of NHL and its subtypes (Table 2). These findings were consistent when data were stratified by sex, age, or race.

#### Discussion

The present InterLymph analysis, which is based on 18 studies from 13 countries, found little evidence to support the hypothesis that excess weight-for-height is associated with NHL. A slightly increased OR amongst the tallest men was observed compared to those who were of mid-range height but no association was found among women. The large number of

subjects included in this analysis enabled examination of risks for subtypes of NHL. While findings for most were consistent with total NHL, an increased risk for DLBCL among persons with a BMI of 40 kg m<sup>-2</sup> or more was observed in a meta-analysis of study-specific ORs. For DLBCL, ORs were elevated with overweight/obesity in North America and amongst the most obese in Northern Europe, yet studies in either region did not show an increasing trend with a 5 kg m<sup>-2</sup> rise in BMI. Marked heterogeneity between studies was present for all categories of BMI, which remained when studies were combined by continent, study design, time period, WHO lymphoma classification used; and when data were restricted to men or women, persons aged 18 to 65, Caucasians alone or studies with participation rates of 70% or more. ORs were less heterogeneous amongst studies with the greatest proportions of controls with a BMI of 25 kg m<sup>-2</sup> or more. Of the seven studies in this subset, no effect of BMI on NHL risk was observed, and the lack of association with obesity was consistent across NHL subtypes, amongst men and women, and at age 45, 46–55, 56–65, and 66 years.

Six of the case-control studies included in this pooled analysis have previously published data on NHL and obesity (3;8;11;14;23;24;50) and a further 12 are included here for the first time. Apart from case-control studies, adiposity has been investigated in cohorts where height and weight were measured (9;10;12;13;25–27) or self-reported (5;15;20;21), and among persons with a hospital discharge for obesity (4;19;22). Cohort studies have the advantage of prospectively collected information, although not necessarily at a relevant time point. Positive associations with obesity have been reported for some cohorts (5;9;10;12;13;15), but not for others (4;19–22;25–27); and a further case-control investigation nested within a cohort reported a reduced risk based on measured height and weight (16). Only one additional study of case-control design- which is not part of the InterLymph consortium- has published its findings, observing an excess risk of NHL with obesity (6).

Hitherto only a few individual case-control studies and two cohort studies have considered lymphoma subtypes, proposing an association with excess adiposity for DLBCL, but less so for FL and CLL/SLL (8;11;14;15;21;22;24;50). A recent meta-analysis of published risk estimates suggested a slight increased risk of NHL, particularly DLBCL based upon data from both case-control and cohort studies (51). The pooled analysis presented here has the advantage of being less susceptible to positive publication bias since it is based on all studies within the InterLymph consortium that collected anthropometric information - around 40% of the data have not been presented before. Another advantage of pooling individual records is that it permits uniform categorization of data, as well as the assessment of the effects of potentially confounding factors. In this regard, adjustment for smoking and alcohol consumption did not greatly alter the risk estimates.

With respect to potential biases, participation rates were generally lower in controls than cases, and a particular concern is whether controls are representative of the populations from which cases were drawn. It is reassuring to note that pooling data from studies with control participation rates of 70% or more gave findings similar to those reported overall. Nonetheless, it is still possible that poor control participation could have influenced our findings since we cannot rule out the possibility that those with obesity-related health problems (e.g. type 2 diabetes, cardiovascular disease, respiratory difficulties, chronic musculoskeletal problems) may have been (more or) less likely to participate. If the latter applied, the increased risks in the highest BMI category could be an artefact of differential case-control participation.

The rapidly changing prevalence of obesity is a growing public health problem, and to further investigate the issue, age-standardised data calculated from height and weight

measurements were sourced from the World Health Organisation Global Database on Body Mass Index (http://www.who.int/bmi/). Interestingly, the relative order of the overweight  $(25-29.99 \text{ kg m}^{-2})$  /obesity (  $30 \text{ kg m}^{-2}$ ) prevalence across studies among our controls and that of the corresponding country-specific WHO BMI prevalence from around the year 2000 are not strongly correlated (Spearman's  $\rho$ =0.41, p=0.08) (Figure 4). WHO data place the USA, Germany and the UK at the top while our self-reported information rate the Czech Republic, USA and Italy as having the highest overweight/obesity prevalence. Whilst differences between our data and WHO are likely to be related to factors such as age, sex and time period, they serve to illustrate the rapidly changing patterns and wide variations that exist around the world.

Self-reports of anthropometric information is known to be inexact, with height tending to be overestimated and weight underestimated (52). The nature of individual misreporting is likely to be complex, being related not only to their actual size but also to other factors such as age and sex. In a cohort of British adults, for instance, where self-reported and measured data were compared, height was overestimated most by older people, shorter men and heavier women, while the greatest underestimation of weight was amongst heavier men and women (53). This tendency for people to report BMI closer to 'normal' may have diluted our odds ratios. It is also possible that weight loss associated with lymphoma may have influenced the recall of cases differently to that of controls. Because of this, at interview subjects were either asked to recall their usual weight or their weight at a specified times before diagnosis/interview, and restricting the analyses to the six studies (NCI-SEER, Mayo Phase 1, British Columbia, UK, North Italy and Italy) where data were requested at 1 or 5 years prior to diagnosis yielded similar results to the findings overall.

Whilst BMI derived from height and weight acts as an easily obtained estimate of adiposity, its use as a marker of obesity has several potential weaknesses. Across different ethnic groups, for example, a given BMI may not correspond to the same proportions of body fat (54). Moreover since the index was originally devised as a means of assessing average body composition among sedentary individuals of working age it may not truly reflect the degree of adiposity across the population as a whole. For instance, among the elderly where muscle mass may have started to decline, body fat mass may be underestimated by BMI whereas amongst athletes it may be overestimated. To account for the potential variation in BMI as a marker of body fat across different populations, we grouped our data according to studyspecific control distributions as well as WHO BMI categories. We also repeated the analyses restricting data to Caucasians, and to North American and Northern European studies combined. Sensitivity analyses were conducted too among persons aged 65 or less (71% of our subjects), and among those who were not regular heavy exercisers where this information was requested (NCI-SEER, British Columbia and HERPACC2). These additional investigations gave similar findings to the presented results. More specific estimates of adiposity may be derived from total body fat mass and, as a marker for abdominal fat distribution, waist-to-hip ratios, but such data were not obtained in the studies included here and have only rarely been investigated with respect to NHL elsewhere, showing little effect (15;21;26).

In conclusion, this pooled analysis of case-control studies from 13 countries, crossing 3 continents, did not find an association between NHL and increased BMI. ORs were raised in studies from some countries, namely the US, Canada and Northern European nations, but even within this group, heterogeneity was observed, questioning the validity of a combined odds ratio. The findings presented here were based on individual data from a large number of subjects enrolled in 18 studies, pooling of which were accomplished through the InterLymph consortium. Some of the advantages of this pooled analysis include information on confounders and NHL subtypes as well as data on height and weight, the constituent

components of BMI. One potential confounding factor not assessed here is diet but dietary data will be examined, in conjunction with BMI, in a future InterLymph pooled analysis. Such investigations may further elucidate whether NHL or its subtypes are associated with obesity per se.

#### Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

#### Acknowledgments

Studies that contributed data to this pooled analysis were supported by: NCI contracts PC65064, PC67008, PC67009, PC67010, PC71105 and the Intramural Research Program of the National Cancer Institute, National Institutes of Health (NCI-SEER study); American Institute for Cancer Research grant 99B083 (Nebraska study); grants CA92153 (Mayo Clinic study); grant CA50850 from NCI (USCF study); the Canadian Cancer Society through the National Cancer Institute of Canada, the Canadian Institutes for Health Research and the Chan Sisters Foundation (British Columbia study); the Leukaemia Research Fund (UK study); NIH grant 5RO1 CA69269-02, Swedish Cancer Society grant 02 661, Plan Denmark grant, The Danish National Research Foundation grant (SCALE); European Commission 5<sup>th</sup> Framework Program, Quality of Life (grant No. QLK4-CT-2000-00422); the Spanish Ministry of Health (grant No. 04-0091, CIBER 06/0073); La Fondation de France, Compagnia di San Paolo di Torino, Programma Oncologia 2001; the German Federal Office for Radiation Protection (grants No. StSch4261 and StSch4420); the Health Research Board (Ireland) (EpiLymph); the National Research Council (CNR) Applied Project "Clinical Applications of Oncological Research" and the Italian Association for Cancer Research (northern Italy study); grant CA51086 from NCI, the European Community (Europe Against Cancer Programme), and the Lega Italiana per la Lotta Contro i Tumori (Italy study); Grant-in-Aid for Scientific Research from the Ministry of Education, Science, Sports, Culture and Technology of Japan (HERPACC1 & 2).

#### References

- 1. Marti A, Marcos A, Martinez JA. Obesity and immune function relationships. Obes Rev. 2001 May; 2(2):131–40. [PubMed: 12119664]
- 2. Samartin S, Chandra RK. Obesity, overnutrition and the immune system. Nutrition Research. 2001; 21(1–2):243–62.
- Holly EA, Lele C, Bracci PM, McGrath MS. Case-control study of non-Hodgkin's lymphoma among women and heterosexual men in the San Francisco Bay Area, California. Am J Epidemiol. 1999 Aug 15; 150(4):375–89. [PubMed: 10453814]
- Wolk A, Gridley G, Svensson M, Nyren O, McLaughlin JK, Fraumeni JF, Adami HO. A prospective study of obesity and cancer risk (Sweden). Cancer Causes Control. 2001 Jan; 12(1):13– 21. [PubMed: 11227921]
- Calle EE, Rodriguez C, Walker-Thurmond K, Thun MJ. Overweight, obesity, and mortality from cancer in a prospectively studied cohort of U.S. adults. N Engl J Med. 2003 Apr 24; 348(17):1625– 38. [PubMed: 12711737]
- Pan SY, Johnson KC, Ugnat AM, Wen SW, Mao Y. Association of obesity and cancer risk in Canada. Am J Epidemiol. 2004 Feb 1; 159(3):259–68. [PubMed: 14742286]
- Bahl S, Cotterchio M, Kreiger N, Klar N. Antidepressant medication use and non-Hodgkin's lymphoma risk: no association. Am J Epidemiol. 2004 Sep 15; 160(6):566–75. [PubMed: 15353417]
- Cerhan JR, Bernstein L, Severson RK, Davis S, Colt JS, Blair A, Hartge P. Anthropometrics, physical activity, related medical conditions, and the risk of non-hodgkin lymphoma. Cancer Causes Control. 2005 Dec; 16(10):1203–14. [PubMed: 16215871]
- Oh SW, Yoon YS, Shin SA. Effects of excess weight on cancer incidences depending on cancer sites and histologic findings among men: Korea National Health Insurance Corporation Study. J Clin Oncol. 2005 Jul 20; 23(21):4742–54. [PubMed: 16034050]
- Rapp K, Schroeder J, Klenk J, Stoehr S, Ulmer H, Concin H, Diem G, Oberaigner W, Weiland SK. Obesity and incidence of cancer: a large cohort study of over 145,000 adults in Austria. Br J Cancer. 2005 Oct 31; 93(9):1062–7. [PubMed: 16234822]

- Willett EV, Skibola CF, Adamson P, Skibola DR, Morgan GJ, Smith MT, Roman E. Non-Hodgkin's lymphoma, obesity and energy homeostasis polymorphisms. Br J Cancer. 2005 Oct 3; 93(7):811–6. [PubMed: 16160698]
- Chiu BC, Gapstur SM, Greenland P, Wang R, Dyer A. Body mass index, abnormal glucose metabolism, and mortality from hematopoietic cancer. Cancer Epidemiol Biomarkers Prev. 2006 Dec; 15(12):2348–54. [PubMed: 17164355]
- Engeland A, Tretli S, Hansen S, Bjorge T. Height and body mass index and risk of lymphohematopoietic malignancies in two million Norwegian men and women. Am J Epidemiol. 2007 Jan 1; 165(1):44–52. [PubMed: 17041129]
- Chiu BC, Soni L, Gapstur SM, Fought AJ, Evens AM, Weisenburger DD. Obesity and risk of non-Hodgkin lymphoma (United States). Cancer Causes Control. 2007 May 7.
- 15. Lim U, Morton LM, Subar AF, Baris D, Stolzenberg-Solomon R, Leitzmann M, Kipnis V, Mouw T, Carroll L, Schatzkin A, Hartge P. Alcohol, Smoking, and Body Size in Relation to Incident Hodgkin's and Non-Hodgkin's Lymphoma Risk. Am J Epidemiol. 2007 Jun 27.
- Paffenbarger RS Jr, Wing AL, Hyde RT. Characteristics in youth predictive of adult-onset malignant lymphomas, melanomas, and leukemias: brief communication. J Natl Cancer Inst. 1978 Jan; 60(1):89–92. [PubMed: 272469]
- 17. Whittemore AS, Paffenbarger RS Jr, Anderson K, Lee JE. Early precursors of site-specific cancers in college men and women. J Natl Cancer Inst. 1985 Jan; 74(1):43–51. [PubMed: 3855486]
- Franceschi S, Serraino D, Bidoli E, Talamini R, Tirelli U, Carbone A, La Vecchia C. The epidemiology of non-Hodgkin's lymphoma in the north-east of Italy: a hospital-based case-control study. Leuk Res. 1989; 13(6):465–72. [PubMed: 2770331]
- Moller H, Mellemgaard A, Lindvig K, Olsen JH. Obesity and cancer risk: a Danish record-linkage study. Eur J Cancer. 1994; 30A(3):344–50. [PubMed: 8204357]
- Zhang S, Hunter DJ, Rosner BA, Colditz GA, Fuchs CS, Speizer FE, Willett WC. Dietary fat and protein in relation to risk of non-Hodgkin's lymphoma among women. J Natl Cancer Inst. 1999 Oct 20; 91(20):1751–8. [PubMed: 10528026]
- 21. Cerhan JR, Janney CA, Vachon CM, Habermann TM, Kay NE, Potter JD, Sellers TA, Folsom AR. Anthropometric characteristics, physical activity, and risk of non-Hodgkin's lymphoma subtypes and B-cell chronic lymphocytic leukemia: a prospective study. Am J Epidemiol. 2002 Sep 15; 156(6):527–35. [PubMed: 12226000]
- 22. Samanic C, Gridley G, Chow WH, Lubin J, Hoover RN, Fraumeni JF Jr. Obesity and cancer risk among white and black United States veterans. Cancer Causes Control. 2004 Feb; 15(1):35–43. [PubMed: 14970733]
- Bosetti C, Dal Maso L, Negri E, Talamini R, Montella M, Franceschi S, La Vecchia C. Re: Body mass index and risk of malignant lymphoma in Scandinavian men and women. J Natl Cancer Inst. 2005 Jun 1; 97(11):860–1. [PubMed: 15928310]
- 24. Chang ET, Hjalgrim H, Smedby KE, Akerman M, Tani E, Johnsen HE, Glimelius B, Adami HO, Melbye M. Body mass index and risk of malignant lymphoma in Scandinavian men and women. J Natl Cancer Inst. 2005 Feb 2; 97(3):210–8. [PubMed: 15687364]
- Lukanova A, Bjor O, Kaaks R, Lenner P, Lindahl B, Hallmans G, Stattin P. Body mass index and cancer: Results from the Northern Sweden Health and Disease Cohort. Int J Cancer. 2006 Jan 15; 118(2):458–66. [PubMed: 16049963]
- MacInnis RJ, English DR, Hopper JL, Giles GG. Body size and composition and the risk of lymphohematopoietic malignancies. J Natl Cancer Inst. 2005 Aug 3; 97(15):1154–7. [PubMed: 16077074]
- Fernberg P, Odenbro A, Bellocco R, Boffetta P, Pawitan Y, Adami J. Tobacco use, body mass index and the risk of malignant lymphomas--a nationwide cohort study in Sweden. Int J Cancer. 2006 May 1; 118(9):2298–302. [PubMed: 16331621]
- Samanic C, Chow WH, Gridley G, Jarvholm B, Fraumeni JF Jr. Relation of body mass index to cancer risk in 362,552 Swedish men. Cancer Causes Control. 2006 Sep; 17(7):901–9. [PubMed: 16841257]

- Willett EV, Smith AG, Dovey GJ, Morgan GJ, Parker J, Roman E. Tobacco and alcohol consumption and the risk of non-Hodgkin lymphoma. Cancer Causes Control. 2004 Oct; 15(8): 771–80. [PubMed: 15456990]
- Besson H, Brennan P, Becker N, De SS, Nieters A, Font R, Maynadie M, Foretova L, Cocco PL, Staines A, Vornanen M, Boffetta P. Tobacco smoking, alcohol drinking and Hodgkin's lymphoma: a European multi-centre case-control study (EPILYMPH). Br J Cancer. 2006 Aug 7; 95(3):378–84. [PubMed: 16819547]
- Becker N, Deeg E, Nieters A. Population-based study of lymphoma in Germany: rationale, study design and first results. Leuk Res. 2004 Jul; 28(7):713–24. [PubMed: 15158093]
- 32. de Sanjose S, Shah KV, Domingo-Domenech E, Engels EA, Fernandez dS, Alvaro T, Garcia-Villanueva M, Romagosa V, Gonzalez-Barca E, Viscidi RP. Lack of serological evidence for an association between simian virus 40 and lymphoma. Int J Cancer. 2003 Apr 20; 104(4):522–4. [PubMed: 12584752]
- 33. Talamini R, Montella M, Crovatto M, Dal Maso L, Crispo A, Negri E, Spina M, Pinto A, Carbone A, Franceschi S. Non-Hodgkin's lymphoma and hepatitis C virus: a case-control study from northern and southern Italy. Int J Cancer. 2004 Jun 20; 110(3):380–5. [PubMed: 15095303]
- 34. Tajima K, Hirose K, Inoue M, Takezaki T, Hamajima N, Kuroishi T. A Model of Practical Cancer Prevention for Out-patients Visiting a Hospital: the Hospital-based Epidemiologic Research Program at Aichi Cancer Center (HERPACC). Asian Pac J Cancer Prev. 2000; 1(1):35–47. [PubMed: 12718687]
- 35. Suzuki T, Matsuo K, Ito H, Hirose K, Wakai K, Saito T, Sato S, Morishima Y, Nakamura S, Ueda R, Tajima K. A past history of gastric ulcers and Helicobacter pylori infection increase the risk of gastric malignant lymphoma. Carcinogenesis. 2006 Jul; 27(7):1391–7. [PubMed: 16400189]
- Spinelli JJ, Ng C, Weber JP, Connors JM, Gascoyne RD, Lai A, Brooks-Wilson A, Le ND, Berry B, Gallagher RP. Organochlorines and risk of non-Hodgkin lymphoma. Int J Cancer. 2007 In press.
- 37. Jaffe, E.; Harris, N.; Stein, H.; Vardiman, J. Tumours of the Haemopoietic and Lymphoid Tissues. IARC Press; 2001.
- 38. Morton LM, Turner JJ, Cerhan JR, Linet MS, Treseler PA, Clarke CA, Jack A, Cozen W, Maynadie M, Spinelli JJ, Costantini AS, Rudiger T, et al. Proposed classification of lymphoid neoplasms for epidemiologic research from the International Lymphoma Epidemiology Consortium (InterLymph). Blood. 2007 Mar 27.
- 39. WHO Expert Committee on Physical Status. Physical Status: The Use and Interpretation of Anthropometry: Report of a WHO Expert Committee. Geneva, Switzerland: World Health Organization; 1995. Report No.: 854
- 40. Morton LM, Hartge P, Holford TR, Holly EA, Chiu BC, Vineis P, Stagnaro E, Willett EV, Franceschi S, La Vecchia C, Hughes AM, Cozen W, et al. Cigarette smoking and risk of non-Hodgkin lymphoma: a pooled analysis from the International Lymphoma Epidemiology Consortium (InterLymph). Cancer Epidemiol Biomarkers Prev. 2005 Apr; 14(4):925–33. [PubMed: 15824165]
- Rothman N, Skibola CF, Wang SS, Morgan G, Lan Q, Smith MT, Spinelli JJ, Willett E, de Sanjose S, Cocco P, Berndt SI, Brennan P, et al. Genetic variation in TNF and IL10 and risk of non-Hodgkin lymphoma: a report from the InterLymph Consortium. Lancet Oncol. 2006 Jan; 7(1):27– 38. [PubMed: 16389181]
- Morton LM, Zheng T, Holford TR, Holly EA, Chiu BC, Costantini AS, Stagnaro E, Willett EV, Dal Maso L, Serraino D, Chang ET, Cozen W, et al. Alcohol consumption and risk of non-Hodgkin lymphoma: a pooled analysis. Lancet Oncol. 2005 Jul; 6(7):469–76. [PubMed: 15992695]
- 43. Wang SS, Slager SL, Brennan P, Holly EA, De SS, Bernstein L, Boffetta P, Cerhan JR, Maynadie M, Spinelli JJ, Chiu BC, Cocco PL, et al. Family history of hematopoietic malignancies and risk of non-Hodgkin lymphoma (NHL): a pooled analysis of 10 211 cases and 11 905 controls from the International Lymphoma Epidemiology Consortium (InterLymph). Blood. 2007 Apr 15; 109(8): 3479–88. [PubMed: 17185468]

- 44. Kricker A, Armstrong BK, Hughes AM, Goumas C, Smedby KE, Zheng T, Spinelli JJ, De SS, Hartge P, Melbye M, Willett EV, Becker N, et al. Personal sun exposure and risk of non Hodgkin lymphoma: A pooled analysis from the Interlymph Consortium. Int J Cancer. 2007 Aug 20.
- 45. Breslow, NE.; Day, NE. The Analysis of Case-Control Studies. Vol. 1. Lyon: International Agency for Research on Cancer; 1980. Statistical Methods in Cancer Research.
- 46. Higgins JP, Thompson SG. Quantifying heterogeneity in a meta-analysis. Stat Med. 2002 Jun 15; 21(11):1539–58. [PubMed: 12111919]
- Colditz GA, Burdick E, Mosteller F. Heterogeneity in meta-analysis of data from epidemiologic studies: a commentary. Am J Epidemiol. 1995 Aug 15; 142(4):371–82. [PubMed: 7625401]
- Delgado-Rodriguez M. Glossary on meta-analysis. J Epidemiol Community Health. 2001 Aug; 55(8):534–6. [PubMed: 11449009]
- 49. Stata Statistical Software: Release 9 [computer program]. College Station, Texas: Stata Corporation; 2005.
- Skibola CF, Holly EA, Forrest MS, Hubbard A, Bracci PM, Skibola DR, Hegedus C, Smith MT. Body mass index, leptin and leptin receptor polymorphisms, and non-hodgkin lymphoma. Cancer Epidemiol Biomarkers Prev. 2004 May; 13(5):779–86. [PubMed: 15159310]
- Larsson SC, Wolk A. Obesity and risk of non-Hodgkin's lymphoma: A meta-analysis. Int J Cancer. 2007 Apr 18.
- Gorber SC, Tremblay M, Moher D, Gorber B. A comparison of direct vs. self-report measures for assessing height, weight and body mass index: a systematic review. Obes Rev. 2007 Jul; 8(4):307– 26. [PubMed: 17578381]
- Spencer EA, Appleby PN, Davey GK, Key TJ. Validity of self-reported height and weight in 4808 EPIC-Oxford participants. Public Health Nutr. 2002 Aug; 5(4):561–5. [PubMed: 12186665]
- 54. WHO Consultation on Obesity. Obesity: preventing and managing the global epidemic: report of a WHO consultation. Geneva, Switzerland: World Health Organization; 2000. Report No.: 894
- 55. WHO Expert Committee on Physical Status. Physical Status: The Use and Interpretation of Anthropometry: Report of a WHO Expert Committee. Geneva, Switzerland: World Health Organization; 1995. Report No.: 854

Willett et al.





Box-Whisker Plot of Body Mass Index among Controls by Study. Body mass index considered to be: Underweight if <18.5 kg m<sup>-2</sup>; Normal weight-for-height if 18.5–24.99 kg m<sup>-2</sup>; Grade 1 Overweight if 25–29.99 kg m<sup>-2</sup>; Grade 2 Obese if 30–39.99 kg m<sup>-2</sup>; and Grade 3 Obese if 40 kg m<sup>-2</sup> (55).

	Normal weight	Underweight		
Study	Cases/Controls	Cases/Controls		OR (95% CI)
North America				
NCI-SEER	165/146	7/3		1.82 (0.45, 7.29)
Nebraska	110/174	3/7		0.67 (0.17, 2.65)
Mayo Phase 1	152/153	4/4	!•	1.13 (0.28, 4.65)
UCSF	714/1453	18/31		1.05 (0.58, 1.90)
British Columbia	332/387	10/13	<b>_</b>	0.91 (0.39, 2.12)
Subtotal			$ \rightarrow $	1.03 (0.68, 1.56)
Northern Furope				
UK	420/615	7/21	<b>_</b>	0.50 (0.21, 1.19)
SCALEDenmark	516/632	15/29		0.63(0.33, 1.20)
SCALESweden	965/943	16/11		1.67 (0.76, 3.70)
EpiLymphIreland	44/90	1/6 —		0.27 (0.03, 2.34)
EpiLymphFinland	36/31	3/2		3.14 (0.41, 23.86)
EpiLymphGermany	223/284	8/12	<b>.</b>	0.97 (0.38, 2.45)
EpiLymphCzechRep	67/85	1/2		1.02 (0.09, 11.90)
Subtotal			$\diamond$	0.85 (0.54, 1.35)
Southern Europe	101/107	2/12	-	
EpiLymphSpain	121/195	3/12		0.40 (0.11, 1.48)
EpiLymphItaly	120/131	4/5		0.96 (0.25, 3.73)
N Italy	230/555	10/33		0.72 (0.35, 1.50)
Italy	94/182	1/4		0.45 (0.05, 4.16)
Subtotal				0.66 (0.38, 1.15)
Asia (Japan)			1	
HERPACC1	290/1515	28/140	<b></b>	1.04 (0.68, 1.60)
HERPACC2	135/691	10/79		0.64 (0.32, 1.28)
Subtotal			$\Rightarrow$	0.89 (0.57, 1.39)
Overall			<b>\$</b>	0.88 (0.71, 1.08)
NOTE: Weights are	from random effe	ctsanalysis		
			.1 .25 .5 11.52 5 5 10	

Study	Normal weight Cases/Controls	Grade 1 Overweight Cases/Controls		OR (95% CI)
North America NCI-SEER Nebraska Mayo Phase 1 UCSF British Columbia Subtotal	165/146 110/174 152/153 714/1453 332/387	178/157 149/217 197/190 451/771 323/307		1.00 (0.73, 1.38) 1.04 (0.75, 1.45) 1.08 (0.79, 1.47) 1.27 (1.09, 1.48) 1.15 (0.93, 1.44) 1.17 (1.05, 1.29)
Northern Europe UK SCALEDenmark SCALESweden EpiLymphIreland EpiLymphFinland EpiLymphGermany EpiLymphFrance EpiLymphCzechRe Subtotal	420/615 516/632 965/943 44/90 36/31 223/284 63/98 p67/85	287/373 414/390 782/797 57/67 27/24 176/306 55/67 89/137	** * * *	$\begin{array}{c} 1.11 \ (0.91, 1.36) \\ 1.21 \ (1.01, 1.46) \\ 0.86 \ (0.75, 0.99) \\ 1.70 \ (1.02, 2.84) \\ 0.76 \ (0.33, 1.73) \\ 0.66 \ (0.51, 0.86) \\ 1.18 \ (0.72, 1.92) \\ 0.74 \ (0.48, 1.15) \\ 0.98 \ (0.81, 1.19) \end{array}$
Southern Europe EpiLymphSpain EpiLymphItaly N Italy Italy Subtotal	121/195 120/131 230/555 94/182	139/197 71/139 135/456 67/200	++	$\begin{array}{c} 1.05 \ (0.76,  1.45) \\ 0.50 \ (0.34,  0.74) \\ 0.70 \ (0.54,  0.89) \\ 0.72 \ (0.49,  1.05) \\ 0.73 \ (0.55,  0.96) \end{array}$
Asia (Japan) HERPACC1 HERPACC2 Subtotal	290/1515 135/691	67/367 32/178	 ◆	0.95 (0.71, 1.27) 0.92 (0.61, 1.41) 0.94 (0.74, 1.20)
Overall <u>NOTE: Weights are</u>	e from random eff	ects analysis	.25 .5 11.52 3 5 1	0.95 (0.85, 1.07)

Study	Normal weight Cases/Controls	Grade 2 Obese Cases/Controls		OR (95% CI)
North America	165/146			0.01 (0.(1, 1.24)
NCI-SEEK	165/146	/////		0.91(0.61, 1.34) 1.26(0.05, 1.05)
Mayo Phase 1	152/153	109/122		1.30(0.93, 1.93) 1.24(0.87, 1.76)
LICSE	714/1453	113/136		1.24(0.87, 1.70) 1.78(1.36, 2.32)
British Columbia	332/387	129/108		1 38 (1 03, 1 86)
Subtotal	332/301	129/100	$\diamond$	1.34 (1.08, 1.65)
Northern Europe			i	
UK	420/615	100/102		1.37 (1.01, 1.86)
SCALEDenmark	516/632	112/120		1.08 (0.81, 1.44)
SCALESweden	965/943	204/223		0.81 (0.66, 1.01)
EpiLymphIreland	44/90	25/42		1.15 (0.62, 2.15)
EpiLymphFinland	36/31	10/16		0.63 (0.23, 1.72) 0.77 (0.54, 1.11)
EpiLymphGermany	223/284	08/101		0.77(0.54, 1.11) 0.58(0.22, 1.07)
EpiLymphCzechRe	03/98 n67/85	21/30	I	0.58(0.52, 1.07) 0.50(0.30, 0.85)
Subtotal	p07705	J 7 / / U		0.50(0.50, 0.05) 0.87(0.69, 1.09)
Subtotal			Ĩ	0.07 (0.0), 1.0))
Southern Europe				
EpiLymphSpain	121/195	50/87		0.83 (0.54, 1.27)
EpiLymphItaly	120/131	24/55	<u> </u>	0.42 (0.24, 0.74)
N Italy	230/555	51/110		1.06 (0.73, 1.53)
Italy	94/182	25/84		0.63 (0.38, 1.06)
Subtotal			$\sim$	0.73 (0.50, 1.05)
Asia (Japan)	200/1515	7/22		1 14 (0 50 2 (2)
HERPACCI HERPACC2	290/1515	//32		1.14(0.50, 2.62) 1.22(0.24, 4.25)
Subtotal	155/091	5/15		1.22(0.34, 4.33) 1.17(0.58, 2.33)
Subiolal				1.17 (0.36, 2.33)
Overall			•	0.97 (0.81, 1.15)
NOTE: Waishts	fuere new days - fu	anta amalanaia	Ţ	
NOTE: weights are	e from random eff	ects analysis		
		.1	.25 .5 11.52 3 5 1	0

Study	Normal weight Cases/Controls	Grade 3 Obese Cases/Controls		OR (95% CI)
North America				
NCI-SEER	165/146	20/15		1.12 (0.55, 2.28)
Nebraska	110/174	11/15		1.20 (0.53, 2.72)
Mayo Phase 1	152/153	12/19		0.57 (0.27, 1.22)
UCSF	714/1453	5/9		1.14 (0.37, 3.47)
British Columbia	332/387	10/9		1.40 (0.56, 3.52)
Subtotal			$\diamond$	1.00 (0.69, 1.46)
Northern Europe				
UK	420/615	14/5		4.23 (1.51, 11.89)
SCALEDenmark	516/632	12/6	•	2.19 (0.81, 5.91)
SCALESweden	965/943	9/7		1.24 (0.46, 3.38)
EpiLymphIreland	44/90	1/2		1.26 (0.11, 14.57)
EpiLymphFinland	36/31	2/2		0.63 (0.07, 5.74)
EpiLymphGermany	223/284	1/6 —		0.19 (0.02, 1.58)
EpiLymphFrance	63/98	2/12		0.27 (0.06, 1.25)
EpiLymphCzechRep	67/85	1/4		0.30 (0.03, 2.81)
Subtotal			$\diamond$	1.00 (0.45, 2.19)
Southern Europe				
EpiLymphSpain	121/195	2/8		0.44 (0.09, 2.15)
EpiLymphItaly	120/131	2/5		0.41 (0.08, 2.16)
N Italy	230/555	1/2		1.26 (0.11, 13.96)
Italy	94/182	2/5		0.64 (0.12, 3.47)
Subtotal			$\diamond$	0.55 (0.23, 1.33)
Overall			$\diamond$	0.99 (0.70, 1.41)
NOTE: Weights are	from random eff	ects analysis		
			1 25 5 11 5 3 5 10	
			.1 .25.5 11.255 10	

#### Figure 2.

**Figure 2(a)**. Meta-analysis of the risk of NHL associated with BMI <18.5 kg m<sup>-2</sup> (Underweight) compared to BMI 18.5–24.99 kg m<sup>-2</sup> (Normal weight). Overall test for heterogeneity: Q=13.0, p=0.73; Variation in odds ratios (OR) attributable to heterogeneity:  $I^2=0.0\%$ . For continents: North America: Q=1.04, p=0.90,  $I^2=0.0\%$ ; Northern Europe: Q=7.87, p=0.25,  $I^2=23.7\%$ ; Southern Europe: Q=1.03, p=0.80,  $I^2=0.0\%$ ; Asia (Japan): Q=1.38, p= 0.24,  $I^2=27.5\%$ . Test for heterogeneity between continents: Q=1.82, p=0.61. Pooled odds ratios by study design were: Population-based studies: OR=0.91, 95% CI 0.68–1.21, Q=6.75, p=0.56,  $I^2=0.0\%$ ; Clinic-based studies: OR=0.92, 95% CI 0.65–1.31, Q=1.47, p=0.48,  $I^2=0.0\%$ ; Hospital-based studies: OR=0.67, 95% CI 0.39–1.17, Q=3.79, p=0.58,  $I^2=0.0\%$ . Test for heterogeneity between study designs: Q=1.04, p=0.59.

**Figure 2(b)**. Meta-analysis of the risk of NHL associated with BMI 25–29.99 kg m<sup>-2</sup> (Grade 1 overweight) compared to BMI 18.5–24.99 kg m<sup>-2</sup> (Normal weight). Overall test for heterogeneity: Q=60.0, p<0.001; Variation in odds ratios (OR) attributable to heterogeneity:  $I^2$ =70.0%. For continents: North America: Q=2.76, p=0.60,  $I^2$ =0.0%; Northern Europe: Q=25.0, p=0.001,  $I^2$ =72.1%; Southern Europe: Q=8.59, p=0.04,  $I^2$ =65.1%; Asia (Japan): Q=0.02, p=0.90,  $I^2$ =0.0%. Test for heterogeneity between continents: Q=23.4, p<0.001. Pooled odds ratios by study design were: Population-based studies: OR=0.97, 95% CI 0.82–1.14, Q=41.6, p<0.001,  $I^2$ =80.8%; Clinic-based studies: OR=0.99, 95% CI 0.82–1.20, Q=0.44, p=0.80,  $I^2$ =0.0%; Hospital-based studies: OR=0.91, 95% CI 0.72–1.16, Q=14.0, p=0.03,  $I^2$ =57.1%. Test for heterogeneity between study designs: Q=3.93, p=0.14.

**Figure 2(c)**. Meta-analysis of the risk of NHL associated with BMI 30–39.99 kg m<sup>-2</sup> (Grade 2 obese) compared to BMI 18.5–24.99 kg m<sup>-2</sup> (Normal weight). Overall test for heterogeneity: Q=59.7, p<0.001; Variation in odds ratios (OR) attributable to heterogeneity:  $I^2=69.8\%$ . For continents: North America: Q=8.18, p=0.08,  $I^2=51.1\%$ ; Northern Europe: Q=18.1, p=0.01,  $I^2=61.2\%$ ; Southern Europe: Q=7.88, p=0.05,  $I^2=62.0\%$ ; Asia (Japan): Q=0.01, p=0.93,  $I^2=0.0\%$ . Test for heterogeneity between continents: Q=25.4, p<0.001. Pooled odds ratios by study design were: Population-based studies: OR=1.06, 95%

CI 0.83–1.34, Q=41.3, p<0.001,  $I^2=80.7\%$ ; Clinic-based studies: OR=1.22, 95% CI 0.90– 1.67, Q=0.03, p=0.99,  $I^2=0.0\%$ ; Hospital-based studies: OR=0.77, 95% CI 0.60–0.98, Q=8.51, p=0.20,  $I^2=29.5\%$ . Test for heterogeneity between study designs: Q=9.81, p=0.007. **Figure 2(d)**. Meta-analysis of the risk of NHL associated with BMI 40 kg m<sup>-2</sup> (Grade 3 obese) compared to BMI 18.5–24.99 kg m<sup>-2</sup> (Normal weight). Overall test for heterogeneity: Q=21.9, p=0.15; Variation in odds ratios (OR) attributable to heterogeneity:  $I^2=26.8\%$ . For continents: North America: Q=2.89, p=0.58,  $I^2=0.0\%$ ; Northern Europe: Q=15.3, p=0.03,  $I^2=54.4\%$ ; Southern Europe: Q=0.69, p=0.88,  $I^2=0.0\%$ . Test for heterogeneity between continents: Q=2.91, p=0.23. Pooled odds ratios by study design were: Population-based studies: OR=1.33, 95% CI 0.88–2.00, Q=11.4, p=0.18,  $I^2=29.7\%$ ; Clinic-based studies: OR=0.57, 95% CI 0.26–1.22, No test for heterogeneity as only 1 study; Hospital-based studies: OR=0.51, 95% CI 0.25–1.05, Q=2.07, p=0.91,  $I^2=0.0\%$ . Test for heterogeneity between study designs: Q=8.41, p=0.015.

Study	OR (95% CI)
North America	
NCI-SEER -	0.98 (0.87, 1.11)
Nebraska	<ul> <li>➡</li> <li>1.11 (0.97, 1.26)</li> </ul>
Mayo Phase 1	1.02 (0.91, 1.15)
UCSF	➡ 1.29 (1.17, 1.42)
British Columbia	<ul> <li>➡</li> <li>1.11 (0.99, 1.24)</li> </ul>
Subtotal	♦ 1.10 (1.00, 1.22)
Northern Europe	
UK	★ 1.19 (1.07, 1.32)
SCALEDenmark	► 1.07 (0.97, 1.17)
SCALESweden	0.92 (0.85, 1.01)
EpiLymphIreland -	0.92 (0.71, 1.18)
EpiLymphFinland	0.83 (0.59, 1.18)
EpiLymphGermany	0.82(0.70, 0.95)
EpiLymphFrance	0.76 (0.61, 0.93)
EpiLymphCzechRep	0.71 (0.56, 0.89)
Subtotal	0.91 (0.80, 1.03)
Southern Europe	
EpiLymphSpain -	0.90 (0.76, 1.06)
EpiLymphItaly	$0.69\ (0.56,\ 0.86)$
N Italy	0.94 (0.80, 1.10)
Italy	0.83 (0.67, 1.02)
Subtotal	0.85 (0.75, 0.96)
Asia (Japan)	
HERPACC1	• 1.09 (0.90, 1.33)
HERPACC2 —	1.02 (0.76, 1.38)
Subtotal	> 1.07 (0.91, 1.26)
Overall	0.96 (0.89, 1.04)
NOTE: Weights are from random effects analysis	
.1 .25 .5	1 1.5 2 3 5 10

#### Figure 3.

Meta-analysis of the risk of NHL associated with 5 kg m<sup>-2</sup> increase in BMI above 18.5 kg m<sup>-2</sup> (Normal weight and above).

Overall test for heterogeneity: Q=87.5, p<0.001; Variation in odds ratios (OR) attributable to heterogeneity:  $I^2=79.4\%$ . For continents: North America: Q=15.5, p=0.004,  $I^2=74.1\%$ ; Northern Europe: Q=37.4, p<0.001,  $I^2=81.3\%$ ; Southern Europe: Q=5.32; p=0.15;  $I^2=43.6\%$ ; Asia (Japan): Q=0.12, p=0.73,  $I^2=0.0\%$ . Test for heterogeneity between continents: Q=29.0, p<0.001. Pooled odds ratios by study design were: Population-based studies: OR=1.02, 95% CI 0.92–1.13, Q=57.7, p<0.001,  $I^2=86.1\%$ ; Clinic-based studies: OR=1.04, 95% CI 0.94–1.14, Q=0.34, p=0.84,  $I^2=0.0\%$ ; Hospital-based studies: OR=0.85, 95% CI 0.79–0.92, Q=6.09, p=0.41,  $I^2=1.4\%$ . Test for heterogeneity between study designs: Q=23.4, p<0.001.



#### Figure 4.

Comparison of Control and WHO Overweight/Obesity Prevalences by Study. Overweight (BMI 25–29.99 kg m<sup>-2</sup>) and Obesity (BMI 30 kg m<sup>-2</sup>) prevalence from the World Health Organisation (WHO) Global Database on Body Mass Index (http:// www.who.int/bmi/). WHO prevalence was derived from the most recent published age- and sex- standardised BMI data calculated from height and weight measured in clearly defined population samples; these data were largely from around the year 2000. The relative order of control overweight/obesity prevalences across studies was not similar to that from data reported on the WHO Global database for BMI (Spearman's  $\rho$ =0.41, p=0.08).

~
~
_
_
_
U .
-
~
-
<u> </u>
_
_
_
$\sim$
0
_
-
-
0
_
_
-
<u> </u>
10
CO L
0
~
_
_
<b>_</b>

Table 1

s.
ysi
lal
ar
led
00
d o
the
11.
ed
pn
ncl
S 1.
die
stu
ol
ntr
ŝ
se-
ca
of
ics
ist
ter
rac
hai
$\mathcal{O}$

Study	Location	Year of Diagnosis	Age Range	Cases (N=10453)	Controls (N=16507)				Reference
				N	Participation (%)	Source	N	Participation (%)	
NCI-SEER	Detroit, Michigan; Iowa; Los Angeles, California; Seattle, Washington, USA	1998–2001	20-70	527	76	<65 years RDD; 65+ years random selection from centers for Medicare and Medicare and Medicare Services, stratified by study area, age, sex and race	468	52	(8)
Nebraska NHL Study	Nebraska, USA	1999–2002	20–75	387	74	RDD, frequency matched by age and sex	535	78	(14)
Mayo Clinic Phase 1	Iowa, Wisconsin, Minnesota, USA	2002–2005	18+	499	66	Frequency matched by age, sex and county of residence	499	02	n/a
UCSF	San Francisco, USA	1988–1995	21–74	1304	72	RDD, frequency matched by age, sex, and county of residence	2402	78	(3)
British Columbia Study	Vancouver and Victoria, British Columbia, Canada	2000–2004	20-82	828	78	Random selection from Client Registry of the Ministry of Health, frequency matched by age, sex and region	848	46	(36)
UK	Yorkshire, Lancashire, South Lakeland and parts of Southwest England	1998–2003	16–69	833	70	Random selection from general practice lists, individually matched by age, sex and region of residence	1141	69	(29)
SCALE	Denmark and Sweden	2000-2002	18–74	3055	81	Random selection from population register, frequency matched by sex and age	3187	71	(24)
EpiLymph Ireland	Six hospitals on the East Coast of the Republic of Ireland	2001–2003	18–80	135	90	Hospital controls matched by age $(\pm 5$ years), sex and study region	208	75	(30)
EpiLymph Finland	Finland	2001–2003	18–80	87		Hospital controls matched by age $(\pm 5)$	75		n/a

**NIH-PA** Author Manuscript

_
_
U
~
~
-
<u> </u>
a
-
$\mathbf{O}$
$\sim$
~
5
5
<u>u</u>
_
<u> </u>
<u> </u>
()
~
0
-
_
<u> </u>
<b>—</b>

Study	Location	Year of Diagnosis	Age Range	Cases (N=10453)	Controls (N=16507)				Reference
				Z	Participation (%)	Source	Z	Participation (%)	
						years), sex and study region			
EpiLymph Germany	Ludwigshafen/Upper Palatinate, Heidelberg/Rhine- Neckar Couny, Würzburg/Lower Frankonia, Hamburg, Bielefeld and Munich	1999–2002	18-80	496	88	Random selection from population register, individually matched by sex, age and study region	710	44	(31)
EpiLymph France	Amiens, Dijon and Montpellier	2000-2003	18-80	206	16	Hospital controls matched by age $(\pm 5$ years), sex and study region	276	74	(30)
EpiLymph Czech Republic	1 centre in Czech Republic	2001–2003	18–80	195	06	Hospital controls individually matched by age ( $\pm 5$ years), sex and study region	304	60	(30)
EpiLymph Spain	Barcelona, Tortosa, Reus and Madrid	1998–2002	18–80	428	82	Hospital controls matched by age (±5 years), sex and study region	631	96	(32)
EpiLymph Italy	2 centres in Italy	1998-2004	1880	222	93	Random selection from population census list, matched by age $(\pm 5 \text{ years})$ , sex and study region	336	66	(30)
Northern Italy	Aviano & Milan	1983–1992	17–79	429	>97	Patients admitted for acute, nonneoplastic, nonimmunologic conditions in the hospitals where cases diagnosed	1157	>97	(18)
Italy	Aviano & Naples	1999–2002	18-84	225	97	Hospital controls, frequency matched by age (in 5-year bands), sex and study centre to cases of lymphohematopoietic meoplasms including NHL and hepatocellular carcinoma	504	16	(33)
HERPACCI	Aichi Cancer Centre, Nagoya, Japan	1988–2000	18–79	416	66≈	Random sample of patients not diagnosed with cancer, individually	2260	66≈	(34;35)

	Participation (%)		66≈
	Z		966
	Source	matched by age and sex	Random sample of patients not diagnosed with cancer, individually matched by age and sex
Controls (N=16507)	Participation (%)		66≈
Cases (N=10453)	Z		181
Age Range			18–79
Year of Diagnosis			2001–2004
Location			Aichi Cancer Centre, Nagoya, Japan
Study			HERPACC2

(35)

Reference

**NIH-PA** Author Manuscript

**NIH-PA** Author Manuscript

**NIH-PA** Author Manuscript

**NIH-PA Author Manuscript** 

# Table 2

Number of cases and controls, pooled odds ratios and 95% confidence intervals for body mass index by all NHL subtypes and the three most common NHL subtypes in studies with the highest prevalence of overweight/obese controls<sup>a</sup>.

$BMI^b$	Controls (N=2963)	NHL <sup>C</sup> (N=2108)	$OR^d$	95% CI	DLBCL <sup>c</sup> (N=659)	$OR^d$	95% CI	FL <sup>c</sup> (N=457)	$OR^d$	95% CI	CLLL/SLL <sup>d</sup> (N=381)	$OR^{\mathcal{C}}$	95% CI
WHO catego	<i>ry</i> (kg $m^{-2}$ ):												
<18.5	36	24	0.85	0.50 - 1.44	8	1.09	0.48 - 2.49	5	0.96	0.36–2.59	3	1.47	0.40-5.42
18.5-24.99	1040	802	1		273	1		182	1		121	1	
25-29.99	1213	776	0.79	0.69-0.90	222	0.68	0.56 - 0.84	149	0.71	0.56 - 0.91	163	0.96	0.74 - 1.25
30-39.99	566	403	0.84	0.72 - 0.99	111	0.72	0.56 - 0.93	100	0.96	0.73 - 1.26	85	1.09	0.80 - 1.49
40	56	31	0.63	0.40 - 0.99	14	0.94	0.51 - 1.72	5	0.54	0.21 - 1.40	2	0.37	0.09 - 1.61
Missing	52	72			31			16			7		
Test for heten	rogeneity <sup>e</sup>		$\chi^2 = 32.2$	p=0.12		$\chi^2 = 25.6$	$p{=}0.27$		χ <sup>2</sup> =24.7	p=0.05		$\chi^2 = 17.8$	$p{=}0.16$

Grade 3 Obese.

<sup>d</sup>Odds ratios and 95% confidence intervals adjusted for study, sex, age, and race were estimated using unconditional logistic regression.

 $e^{r}$  Test for heterogeneity was conducted by testing for evidence of interaction between BMI and studies using the likelihood ratio test.

 $f_{
m Confidence}$  interval estimated using exact methods.