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A role for XRCC2 gene polymorphisms in breast cancer risk and survival

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Abstract

Background—The XRCC2 gene is a key mediator in the homologous recombination repair of DNA double strand breaks. We hypothesised that inherited variants in the XRCC2 gene might also affect susceptibility to, and survival from, breast cancer.

Methods—We genotyped 12 XRCC2 tagging SNPs in 1,131 breast cancer cases and 1,148 controls from the Sheffield Breast Cancer Study (SBCS), and examined their associations with breast cancer risk and survival by estimating odds ratios (ORs) and hazard ratios (HRs), and their corresponding 95% confidence intervals (CIs). Positive findings were further investigated in 860 cases and 869 controls from the Utah Breast Cancer Study (UBCS) and jointly analysed together

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with available published data for breast cancer risk. The survival findings were further confirmed in studies (8,074 cases) from the Breast Cancer Association Consortium (BCAC).

Results—The most significant association with breast cancer risk in the SBCS dataset was the XRCC2 rs3218408 SNP (recessive model $p=2.3\times 10^{-4}$, MAF=0.23). This SNP yielded an OR_{rec} (95% CI) of 1.64 (1.25–2.16) in a two-site analysis of SBCS and UBCS, and a meta- OR_{rec} (95% CI) of 1.33 (1.12–1.57) when all published data were included. This SNP may mark a rare risk haplotype carried by 2 in 1000 of the control population. Furthermore, the XRCC2 coding R188H SNP (rs3218536, MAF=0.08) was significantly associated with poor survival, with an increased per-allele HR (95% CI) of 1.58 (1.01–2.49) in a multivariate analysis. This effect was still evident in a pooled meta-analysis of 8,781 breast cancer patients from the BCAC [HR (95% CI) of 1.19 (1.05–1.36), $p=0.01$].

Conclusions—Our findings suggest that XRCC2 SNPs may influence breast cancer risk and survival.

Keywords

Single nucleotide polymorphism; XRCC2; breast cancer risk; breast cancer survival

INTRODUCTION

Homologous recombination repair (HRR) of DNA double strand breaks (DSB) is a crucial cellular defence system to maintain genomic integrity. Unrepaired or incorrectly repaired DSB may give rise to chromosome aberrations, such as loss or gain of chromosome segments and chromosome translocations. These changes might lead to carcinogenesis by disruption of tumour suppressor genes and activation of proto-oncogenes.[1, 2] The involvement of the highly penetrant breast cancer genes BRCA1 and BRCA2 in the homologous recombination repair pathway[3–6] highlights the importance of this mechanism in breast cancer aetiology.

X-ray repair cross complementing gene-2 (XRCC2) possesses the ATP binding domains known as Walker motifs A and B, and is one of the RAD51 family of proteins that are implicated in DNA DSB repair.[7, 8] XRCC2-deficient cells show a greater than 100 fold reduction in HRR compared to XRCC2-proficient cells,[9] and demonstrate various forms of chromosomal instability that are often described in breast cancer.[7, 10–12, 8, 13] The restoration of RAD51 nuclear foci in XRCC2-deficient cells irrespective of RAD51 levels, [14, 15] and specific binding of XRCC2 to RAD51 family proteins[16] further suggests a non-redundant role of XRCC2 in normal HRR function. In addition, DNA damage caused by anticancer drugs and radiation has been documented to require XRCC2 for its repair in mammalian cells.[17–21] Several lines of evidence demonstrate that high levels of expression of XRCC3, another member of the RAD51 family of proteins, are associated with radio- and cytotoxic resistance in human tumour cell lines[22–24] suggesting that XRCC2 might also be relevant to the effects of tumour treatment.

It has been widely hypothesised that inherited variations in DNA sequence, such as single nucleotide polymorphisms (SNPs), may modulate DNA repair capacity, thus affecting individual susceptibility to cancer risk or survival. Earlier investigations concentrated on the XRCC2 missense SNP rs3218536 (R188H), for its association with breast cancer risk. An effect of this SNP has been largely ruled out by the Breast Cancer Association Consortium (BCAC) study,[25] although, due to its low minor allele frequency (MAF), small recessive effects could not be excluded. More recently, a few gene-based association studies have been carried out to evaluate XRCC2 SNPs in the context of breast cancer risk.[26–28] However, only one study examined the XRCC2 SNP associations with breast cancer

survival.[27] Therefore, we genotyped a comprehensive set of XRCC2 tagging SNPs (tSNPs) selected from resequencing data and tested the hypotheses that the XRCC2 germline variants captured by the tag SNPs affect breast cancer risk or survival.

METHODS

Study populations

The Sheffield Breast Cancer study (SBCS) formed the discovery set for this study. The study characteristics and recruitment have been described in detail previously.[29, 30] Briefly, 1,266 female patients with histologically confirmed breast cancer were recruited from surgical outpatient clinics at the Royal Hallamshire Hospital, Sheffield and Rotherham District General Hospital between November 1998 and June 2002. Control subjects (n=1,270) were drawn from women aged 50–65 years who attended the mammography breast screening programme in Sheffield between October 2000 and August 2002. The eligibility criterion for controls was the absence of any evidence of breast malignancy. Study subjects were all resident in the Sheffield area and of Northern European ancestry. Tumour characteristics, such as histology, grade, lymph node status, estrogen receptor (ER) and progesterone receptor (PR) status and tumour size were retrieved by reviewing medical records and histopathology reports. Follow-up data on vital status was available until September 2009 through hospital records and the Trent Cancer Registry. All subjects gave informed consent for the collection of data and blood specimens, and approval for this study was obtained from South Sheffield Research Ethics Committee.

The Utah breast cancer study (UBCS) was used to replicate the findings in relation to breast cancer risk. Breast cancer cases (n=860) were drawn from high-risk cancer pedigrees, identified using a genealogical database (Utah Population Database, UPDB) linked to the Utah Cancer Registry.[31] Cases known to be due to BRCA1 or BRCA2 mutations were excluded. Controls (n=869) included unaffected family members and unrelated matched cancer-free controls. The latter were matched based on sex, birth year (within 5 years) and birth-place.[32]

Six studies from the BCAC (that had genotype data for XRCC2 R188H, and survival data available), were employed to verify the XRCC2 SNP associations with survival, including the Australian Breast Cancer Family Study (ABCFS, n=1,223), the Spanish National Cancer Centre Breast Cancer study (CNIO-BCS, n=190), the Hannover Breast Cancer Study (HABCS, n=598), the National Cancer Institute Breast Cancer Study in Poland (PBCS, n=1,507), Studies of Epidemiology and Risk Factors in Cancer Heredity (SEARCH, n=4,234) and the US Radiologic Technologist (USRT, n=322) Study.[33, 25] A brief description of the sources of breast cancer patients and the collection of clinical characteristics and vital status is given in the Supplementary Materials. Most breast cancer patients were prevalent cases (supplementary table 1). The majority of tumours were of ductal type, moderately or well differentiated, with low tumour stage (≤ 2), no lymph node involvement, and positive for ER and PR status (supplementary table 2).

Tagging reference panel and selection of tagging SNPs

The Polymorphism Discovery Resource (PDR90) is a mixed-ethnic population of 24 European, 24 African, 12 Mexican, 6 Native and 24 Asian Americans,[34] and has been used to discover genetic variants by thorough resequencing of all exons, conserved sequences, and over 1 kb of 5' upstream and 3' downstream regions of over 300 genes.[35] Using 4025 Yoruba-specific SNPs from HapMap Release 22 (i.e. those SNPs known to be polymorphic in only the Yoruban group), we identified and excluded 22 individuals likely to be of African genetic background. Data from the remaining 68 subjects (PDR68) were used

to select tSNPs for the XRCC2 gene. The GERBIL (genotype resolution and block identification using likelihood) software[36] was used to estimate haplotype frequencies of 61 common SNPs (MAF > 5%) across the XRCC2 gene region in PDR68. Two haplotype blocks were identified. A minimum allelic r^2 of 0.8 within blocks was employed to select tSNPs using the STATA programme, htSNP2.[37]

Genotyping and quality control

DNA samples were arrayed in 384-well plates, comprising equal numbers of cases and controls, together with duplicates (~10% of samples). Genotyping was carried out using the 5' nuclease (TaqMan) and SNPlex multiplex assays (Applied Biosystems, Foster city, CA). There was no significant deviation of genotype frequencies in controls from those expected under Hardy-Weinberg equilibrium (HWE) (see supplementary table 3), except for rs3218455 (empirical $p=1\times 10^{-4}$). No obvious clustering errors were found for rs3218455 by visual inspection of the cluster plots, and this SNP was included in the analysis to maintain tagging efficiency. Two SNPs were excluded based on a duplicate discordance of over 2.5%. Further to this, any SNP with a call rate of <80%, and any study subjects with > 50% missing genotypes, were also excluded from analyses. A summary of genotyping quality is given in supplementary table 3. Among these SNPs there were no differences in genotype missing rates between cases and controls, except for rs3218534 (p for fisher exact test=0.04, overall call rate of 95.9%).

Statistical analysis

The χ^2 goodness-of-fit test with 1 degree freedom was performed to examine the departure of genotype frequency from HWE among control subjects, and the empirical significance value was obtained using Monte Carlo permutation procedures. Associations between the XRCC2 SNPs and breast cancer risk under specified genetic models were evaluated by likelihood ratio tests (LRT), using the SNPAssoc 1.6-0 package in R 2.90.[38] SNP associations below the arbitrary Bonferroni threshold for multiple tests of 12 SNPs were selected for replication. P values shown in the tables are uncorrected for multiple testing. Haplotype analyses were performed using the SAS PROC HAPLOTYPE routine. Odds ratios (ORs) and corresponding 95% confidence intervals (CIs) for SNPs and haplotypes were derived from logistic regression models, with the most common genotype or haplotype used as the reference. The Genie software was also used to estimate ORs and 95% CIs, in order to account for the known relatedness amongst Utah subjects.[39, 40] Estimates and CIs were very consistent between the two methods.

Published data based on a Medline search through April 2010 were incorporated into the meta-analysis for rs3218408. We also included data from the Cancer Genetic Markers of Susceptibility (CGEMS, <http://cgems.cancer.gov/data/>) genome-wide association study (GWAS) of Hunter *et al*[41], based on postmenopausal women (1,145 cases and 1,142 controls from the Nurses' Health Study, NHS). Preliminary analysis showed that OR estimates obtained from random and fixed effect models were similar, therefore meta-ORs for tSNPs were estimated and illustrated by forest plots using fixed effect models. The Cochran Q test and I^2 [42] were used to examine the homogeneity of ORs across different centres. All the meta-analyses were performed using the STATA command *metan*.[43]

For the analysis of associations with survival following breast cancer diagnosis, time at risk was defined as the interval between date of diagnosis and date of the last follow-up or death. The heterogeneity of the Kaplan-Meier (K-M) survival functions of genotypes for each SNP was assessed by the log-rank test. The Cox proportional hazard model was employed to estimate the hazard ratio (HR) for each tSNP in terms of genotype or alleles, adjusted for age at diagnosis, and taking into account time between study entry and diagnosis (left

censoring).[44] Lymph node status (categorical variable), grade (categorical variable) and tumour size (categorical variable; ≤ 2 cm, 2–5 cm, >5 cm) were included as covariates in the Cox models for those SNPs with significant associations. Pooled HR estimates were estimated by including study as a stratification variable. Schoenfeld residuals were used to assess the assumption of proportional hazards. All significance tests were two-sided and were performed using the Intercooled STATA 9.2 (College Station, TX), unless otherwise specified.

RESULTS

XRCC2 SNP associations with breast cancer risk

Twelve XRCC2 tSNPs, including one missense SNP rs3218536 (R188H), were successfully genotyped in 1,131 cases and 1,148 controls from the SBCS. The mean prediction accuracy achieved for ungenotyped SNPs based on these tSNPs was $> 90\%$, as assessed in PDR68. Of the 12 SNPs, 11 were found not to be associated with breast cancer risk under additive, codominant, dominant or recessive models (figure 1), with all ORs close to unity (see supplementary table 4). However, we observed statistical evidence of rs3218408 association with breast cancer risk under the codominant ($p_{LRT}=6.7\times 10^{-4}$) and recessive models ($p_{LRT}=2.3\times 10^{-4}$), and both of these p values were below a notional Bonferroni threshold of 4.2×10^{-3} for 12 SNPs (figure 1). The unadjusted OR for the recessive effect of rs3218408 was 1.92 (1.35–2.75), and remained similar after adjusting for age at diagnosis, age at menarche, age at first full term pregnancy and family history; 1.85 (1.25–2.72).

Haplotype analyses were performed in the two haplotype blocks as defined in Materials and Methods. All haplotypes with above 1% frequency were tested individually against the most common haplotype in the block, and rare haplotypes (frequency $< 1\%$) were grouped together. In Block 1 (rs3218556, rs3218536, rs3218534, rs3218501, and rs3218499) no haplotypes were associated with breast cancer risk relative to the most common haplotype (table 1). In Block 2 (rs3218455, rs3111465, rs3094406, rs3218408, rs3218400, rs2106776, and rs3218374), the combined group of rare haplotypes were associated with increased risk relative to the most common haplotype with an estimated OR (95% CI) of 1.68 (1.21–2.32) (table 1). On closer inspection, the inflated risk was primarily due to one rare haplotype that included the minor allele of rs3218408, together with the minor allele at rs3218374 (frequency of 0.89% in cases and 0.20% in controls), which was associated with an OR (95% CI) of 6.50 (1.88–22.48). This suggests that this very rare haplotype may harbour one or more susceptibility loci (table 1).

To confirm the single SNP finding for rs3218408, we genotyped it in 860 cases and 869 controls from the UBCS. The MAF of the G allele in UBCS controls was similar to that seen in the SBCS controls, 0.21 compared to 0.23 (supplementary table 3), and genotype frequencies were consistent with HWE ($p_{HWE}=0.98$). There was evidence of association between rs3218408 and breast cancer risk in the UBCS, although the inheritance pattern was more consistent with the additive and dominant models, in contrast to the codominant and recessive models that were indicated in the SBCS dataset (figure 2). However, there was no evidence of heterogeneity between studies for any model ($p_{het}=0.17-0.45$ and $I^2=0\%-46.6\%$; data not shown). In a joint analysis of the SBCS and UBCS, SNP rs3218408 showed statistically significant association with the susceptibility to breast cancer in the additive, recessive, and codominant models, with ORs (95% CIs) of 1.16 (1.04–1.29) [$p=6\times 10^{-3}$] per G allele, 1.64 (1.25–2.16) [$p=3.31\times 10^{-4}$] under a recessive model and 1.66 (1.26–2.19) [$p=3.18\times 10^{-4}$] for homozygosity of the minor allele. To further validate this finding, we extended the analysis to include publicly available data for rs3218408. These included data from Han *et al*[28] based on premenopausal women from the NHS and from the CGEMs GWAS of Hunter *et al*[41] based on postmenopausal women from the NHS. In addition, we

included data from the SEARCH study on the SNP rs3218499,[45] which is in high LD with rs3218408 ($r^2=0.97$ in our controls). All five studies consistently demonstrated a risk effect for the rare homozygotes, with meta-OR (95% CI) for the recessive model of 1.33 (1.12–1.57) [$p=0.001$] (figure 2) in a sample size of 5,518 cases and 5,890 controls.

XRCC2 SNP associations with survival of breast cancer patients

Vital status post-diagnosis was available for 814 of the 1,131 SBCS breast cancer cases. The median follow-up was 11.12 years (range 0.98–40.59). Many of the cases were prevalent cases, with a mean of 4.05 and a median of 2.85 years between diagnosis and recruitment to the study. Table 2 summarises the survival data by SNP. There were statistically significant differences in the K-M survival functions among genotypes for rs3218536 ($p=4\times 10^{-6}$), rs3218534 ($p=0.0224$), rs3218455 ($p=2\times 10^{-6}$) and rs3218374 ($p=0.0248$), with the p values for rs3218536 and rs3218455 being below a Bonferroni correction threshold of 4.2×10^{-3} for 12 SNPs. After adjustment for age and accounting for left-censoring time, the adjusted HRs [aHRs (95% CIs)] of the homozygous minor allele genotypes were 4.26 (1.69–10.72) and 3.86 (1.76–8.47) for rs3218536 and rs3218455, respectively, whilst both rs3218534 and rs3218374 heterozygous genotypes had about 40% reduction in aHRs, compared to the homozygous genotypes for the common allele. There were allele-dosage effects for both rs3218536 and rs3218455, with aHRs (95% CIs) of 1.48 (1.04–2.13) [$p=0.032$] for rs3218536 and 1.51 (1.08–2.10) [$p=0.016$] for rs3218455 (table 2). The assumption of proportionality of the baseline hazards was valid for both SNPs ($p>0.05$). These two SNPs are correlated ($r^2=0.85$) and thus are likely to reflect the same underlying effect. The rs3218536 SNP, which causes the amino acid substitution R188H, was included in a model that further adjusted for lymph node status, grade and tumour size. The minor A allele was significantly associated with poor survival, with a aHR (95% CI) of 1.58 (1.01–2.49) [$p=0.046$].

We tested for replication of the rs3218536 association with survival in six studies from the BCAC, for which genotype data for rs3218536 and survival data were available. The distribution of age at diagnosis, time from diagnosis to recruitment, follow-up time, and available clinical characteristics for each study are shown in supplementary tables 1 and 2. Table 3 shows a summary of the HR estimates associated with rs3218536 by study. Four of the six replication studies demonstrated an increased hazard for the allelic effect of rs3218536, although individual study hazard ratios were not statistically significant. The pooled analysis of 8,781 breast cancer cases (including 1,414 deaths) showed that overall, each copy of the minor allele for rs3218536 was associated with a hazard ratio of 1.19 (1.05–1.36), $p=0.01$, after adjustment for age and study (table 3 and supplementary figure 1). Similar results were obtained if the analysis was restricted to European subjects [1,381 deaths out of 8,615; HR: 1.19 (1.04–1.37)]. The HR for the combined replication studies when SBCS data was excluded was 1.15 (1.00–1.33)[$p=0.05$].

Pooley *et al* (2008) reported that the rs3218536 SNP was specifically associated with receptor-positive tumours.[27] Therefore, we explored the hypothesis of a differential survival effect according to receptor status. However, the effect of rs3218536 on survival in the pooled dataset did not vary significantly according to ER or PR status ($p_{\text{interaction}}=0.16$ for PR; $p_{\text{interaction}}=0.61$ for ER; supplementary table 5).

DISCUSSION

In this study we successfully genotyped a total of 12 tSNPs in the XRCC2 gene and examined their associations with breast cancer risk and survival in the SBCS. The SNP rs3218408 was associated with breast cancer susceptibility. A rare haplotype, including the minor allele at rs3218408, was also identified as associated with breast cancer risk, but this

result will require very large samples to replicate, given the rarity of the haplotype. SNP rs3218408 was also associated with breast cancer risk in the UBCS data set. The data were most consistent with a recessive mode of inheritance.

Four similar candidate gene case-control studies employing tag-SNP approaches have been published with respect to XRCC2 and breast cancer risk. Han *et al.* found the OR (95% CI) for the additive effect of rs3218408 was 0.98 (0.76–1.26) for premenopausal breast cancer risk in a sample size of 238 cases and 474 controls drawn from the predominantly Caucasian Nurses Health Study.[28] Using the genotype distributions supplementary to their publication, we estimated the OR (95% CI) for the recessive effect to be 1.26 (0.68–2.31). Pharoah *et al.* genotyped the rs3218499 SNP in 2176 cases and 2274 controls from the SEARCH study.[45] This SNP is correlated with rs3218408 with $r^2=0.97$, and yielded an OR (95% CI) of 1.14 (0.86–1.50) for the recessive model. Both of the above studies, together with GWAS data from the CGEMs NHS study, were incorporated into the meta-analysis shown in figure 2, resulting in an overall recessive OR (95% CI) of 1.33 (1.12–1.57). In another SEARCH study, Pooley *et al.* genotyped a panel of 8 SNPs in XRCC2, in 2,270 cases and 2,280 controls. However, none of these were any more strongly correlated to rs3218408 than the rs3218499 SNP included in our meta-analysis. Pooley *et al.* found a weak protective effect of rs3218536 (R188H), which was most significant in ER and PR positive tumours.[27] This effect was not seen in our SBCS data, nor was it reproduced in the large study done by the BCAC,[25] although we did observe a non-significant protective OR in SBCS receptor positive tumours [OR (95% CI) 0.80 (0.61–1.05); data not shown]. Haiman *et al.*[26] genotyped 24 XRCC2 SNPs, including rs3218408, in a multi-ethnic study of 2,093 cases and 2,303 controls, and none of the XRCC2 SNPs were associated with breast cancer under an additive model, although the genotype distributions were not available to allow assessment of any recessive effects and could therefore not be included in our meta analysis. The available data suggests a recessive mode of inheritance, although we are not able to rule out other models. While further studies are required to resolve this issue, it seems biologically plausible that homozygous deficiency of a protein involved in DNA repair might be associated with increased cancer risk.

We also examined the XRCC2 SNPs for their associations with overall survival in breast cancer patients. We observed that two XRCC2 SNPs (rs3218536 and rs3218455) were statistically associated with survival. These two SNPs are correlated ($r^2=0.85$) and are likely to be reflecting the same underlying effect. The effect of rs3218536 (R188H) remained significant in the multivariate analysis after adjustment for age at diagnosis, grade, lymph node, and tumour size, suggesting it may have an independent role in overall survival.

Our finding of a role for rs3218536 was not in accord with a recently published work, which found no statistically significant effect of the XRCC2 R188H SNP on breast cancer survival in 2,270 cases from the SEARCH study.[27] However, our pooled analysis of 7 datasets from the BCAC, including both the SBCS and a larger set of SEARCH cases (4,234), provided some support for an association. The minor allele of rs3218536 was associated with a 19% increased risk of death in a total of 8,781 case subjects ($p=0.01$). Further replications are needed to confirm this nominally significant result.

Despite the strength of the use of the large sample sizes, there are some limitations in this study. The risk association reported here was confined to the meta-analysis of the candidate gene studies and one available set of GWAS data; further replication by the incorporation of other relevant genome-wide association data would be beneficial. The use of the UBCS cases could lead to a biased estimate of risk, since they are drawn from high risk pedigrees. However, we see no evidence of this, since the estimate of the odds ratio from the meta-analysis remained the same when the UBCS data was excluded [recessive OR=1.33 (1.11–

1.59)]. In addition, the inclusion of prevalent cases may potentially bias our estimates of breast cancer risk and survival if certain genotypes favour long-term survival. However, we found no evidence that the risk SNP rs3218408 was associated with breast cancer survival, thus the risk estimate is unlikely to be biased. The inclusion of prevalent cases in the survival study may lead to bias in HR estimates. To minimise any potential bias, we employed the left-truncated Cox model. The use of this model with prevalent cases yields the same HR estimates as those found when only incident cases are included. [44]

Due to the role of BRCA1 and BRCA2 in DNA repair, a large number of candidate gene studies of DNA repair genes have been carried out, although few have been conclusively replicated. Genome-wide association and sequencing studies have however provided support for a role for DNA repair genes in breast cancer risk. RAD51L1 (RAD51B) is a member of the RAD51 family of DNA double strand break repair proteins and is associated with risk, [46] and a number of proteins that interact with BRCA1 and BRCA2, such as PALB2 and BRIP1, have been shown to be mutated, albeit rarely, in familial breast cancers (reviewed in [47]). More recently a variant near the MERIT40 gene, whose protein is a component of the BRCA1-A complex, has been shown to act as a modifier of risk in BRCA1 mutation carriers,[48] and to affect ovarian cancer risk.[49]

In conclusion, our study provides evidence supporting an association of the XRCC2 rs3218408 SNP with the risk of breast cancer. If further replicated, data for this SNP could be incorporated into risk models with other validated SNPs. Beyond this single SNP result, a significant haplotype association, incorporating the minor allele at this SNP, was also identified. However, much larger studies of multiple SNPs will be required to further investigate this potentially large effect. With respect to overall survival, we observed an association with the XRCC2 rs3218536 (R188H) SNP. An association with survival has also been reported for this SNP in pancreatic cancer patients, especially in those who received both chemotherapy and X-ray therapy.[50] With the observation that rs3218536 variant cells show increased resistance to cisplatin treatments compared to wild-type cells,[51] future studies are needed to evaluate rs3218536 in the context of chemotherapy, to determine whether there may be implications for treating breast cancer.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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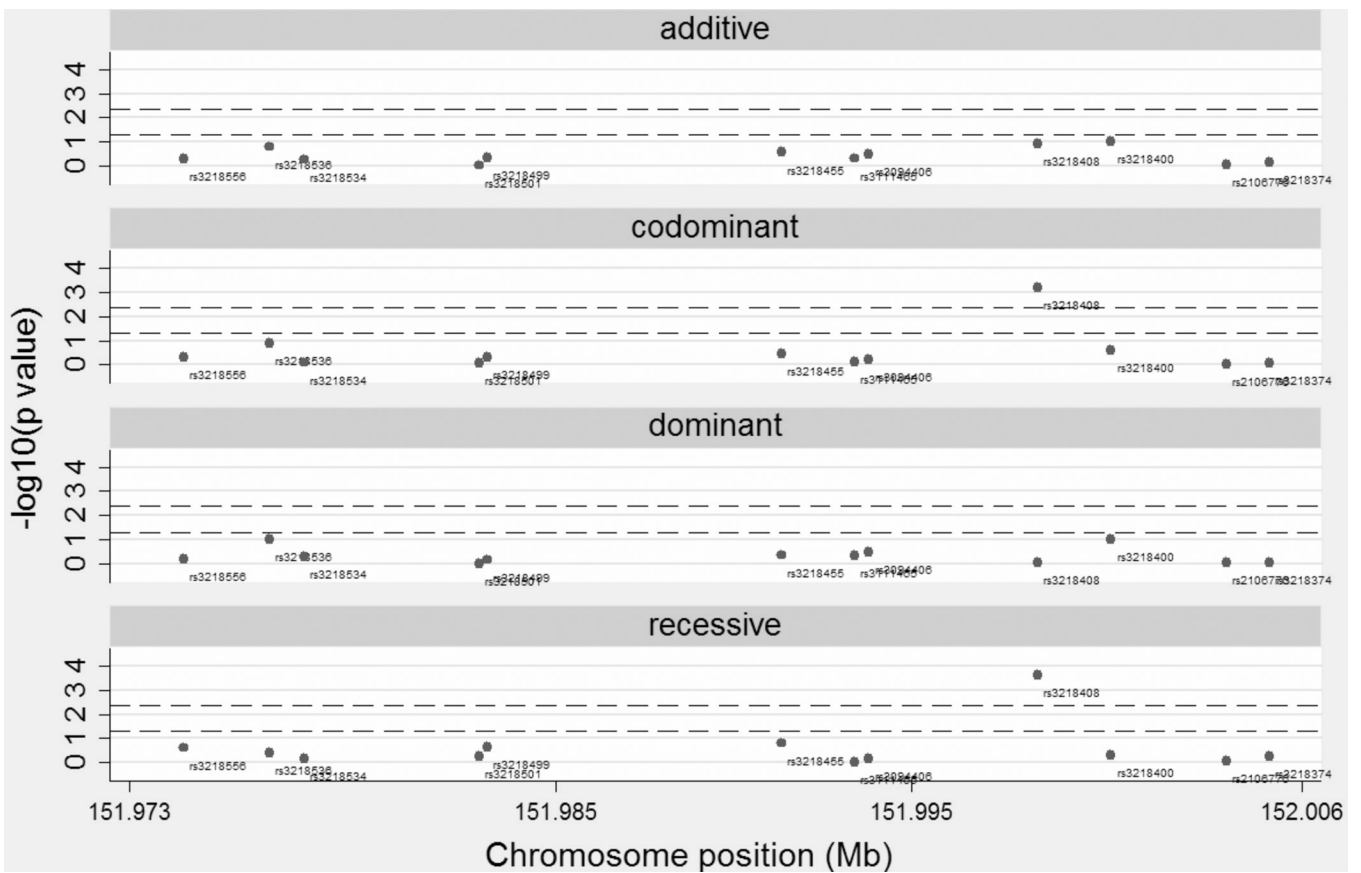


Figure 1. Negative $\log_{10} p$ values for additive, codominant, dominant and recessive models for the XRCC2 SNPs in the SBCS. For each model the upper dash line represents the Bonferroni threshold of 4.2×10^{-3} and the lower dash line represents the nominal significant value of 0.05. Chromosome positions are given in Mb and refer to NCBI build36/hg18 of the human genome

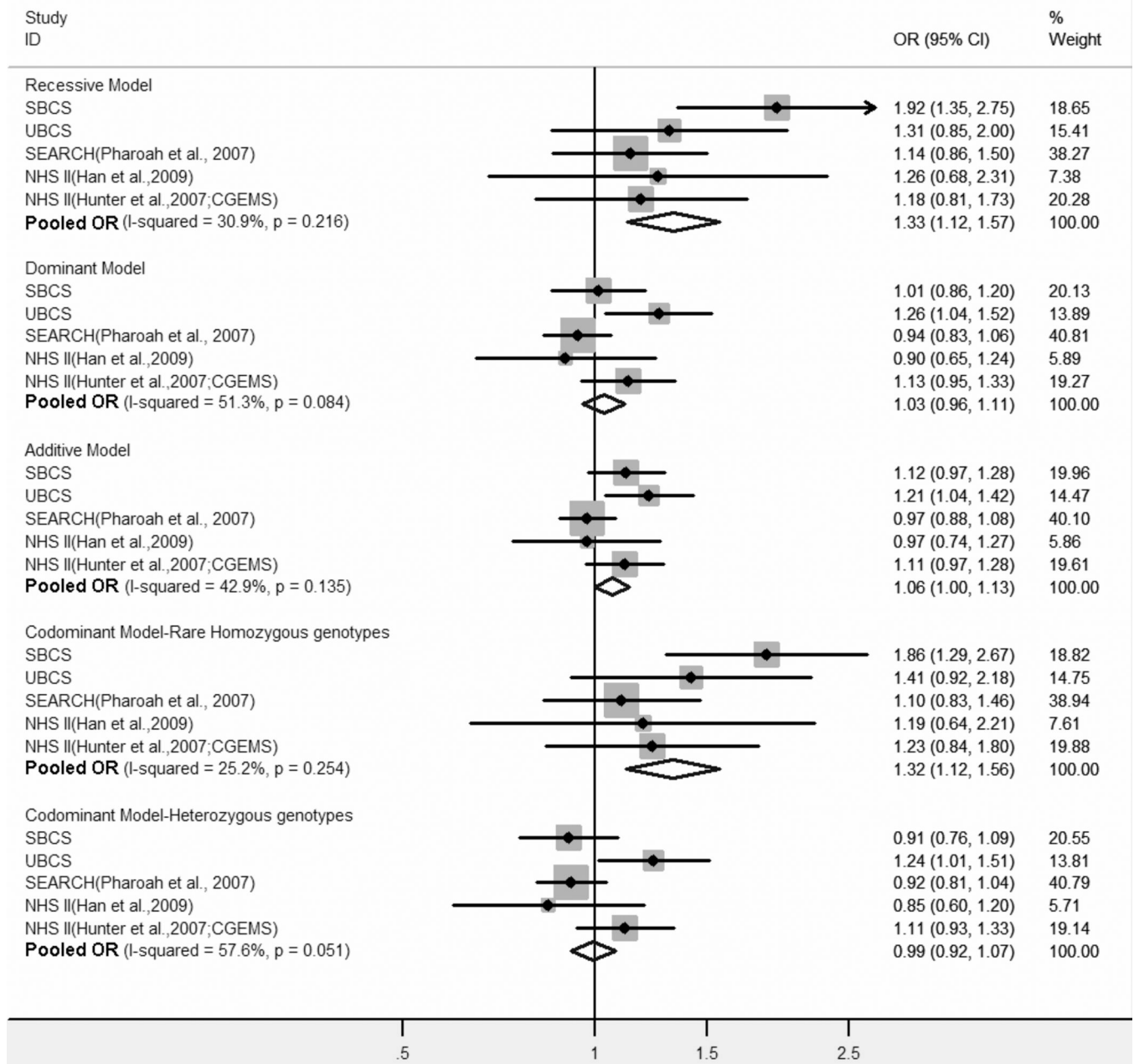


Figure 2. Meta-analysis of the association of rs3218408 with breast cancer risk. Due to the lack of rs3218408 genotype data in the SEARCH study, data for the highly correlated SNP, rs3218499, was used. Fixed effect estimates are shown, with p value for homogeneity in parenthesis

Table 1

Estimated breast cancer odds ratios (ORs) for haplotypes in the XRCC2 gene in SBCS

| Haplotypes* | Case frequency (%) | Control frequency (%) | OR (95% CI) |
|---------------------|--------------------|-----------------------|--------------------------|
| Block1 [†] | | | |
| 1-1-2-1-1 | 987.98 (43.68) | 1024.45 (44.62) | 1 |
| 1-1-1-1-2 | 524.75 (23.2) | 510.97 (22.25) | 1.07 (0.92–1.24) |
| 1-1-1-1-1 | 393.56 (17.4) | 368.08 (16.03) | 1.12 (0.94–1.33) |
| 1-2-1-1-1 | 178.07 (7.87) | 203.09 (8.85) | 0.9 (0.72–1.13) |
| 2-1-1-1-1 | 85.94 (3.8) | 93.79 (4.09) | 0.94 (0.69–1.3) |
| 1-1-1-2-1 | 80.14 (3.54) | 80.29 (3.5) | 1.03 (0.75–1.42) |
| rare haplotypes | 11.57 (0.51) | 15.33 (0.67) | 0.77 (0.34–1.77) |
| Block2 [‡] | | | |
| 1-1-1-1-1-1-2 | 959.74 (42.43) | 1004.07 (43.73) | 1 |
| 1-1-1-2-1-2-1 | 506.51 (22.39) | 501.74 (21.85) | 1.05 (0.90–1.23) |
| 1-1-1-1-2-2-1 | 230.88 (10.21) | 221.96 (9.67) | 1.10 (0.89–1.35) |
| 2-1-1-1-1-2-1 | 175.95 (7.78) | 197.57 (8.60) | 0.91 (0.73–1.13) |
| 1-2-2-1-1-1-1 | 102.96 (4.55) | 108.17 (4.71) | 0.99 (0.73–1.33) |
| 1-1-2-1-1-1-1 | 95.57 (4.23) | 96.37 (4.20) | 1.06 (0.76–1.47) |
| 1-1-1-1-1-2-1 | 85.98 (3.8) | 99.69 (4.34) | 0.89 (0.64–1.23) |
| rare haplotypes | 104.42 (4.61) | 66.43 (2.89) | 1.68 (1.21–2.32) |
| § { 1-1-1-2-1-1-2 | 20.24 (0.89) | 4.56 (0.20) | 6.50 (1.88–22.48) |
| Others | 84.17 (3.72) | 61.87 (2.69) | 1.37 (0.96–1.97) |

* 1 represents the common allele and 2 the minor allele. “Rare haplotypes” indicates those with frequency below 1%.

[†] the order of SNPs in block 1 are rs3218556, rs3218536, rs3218534, rs3218501, and rs3218499.

[‡] The order of SNPs in block 2 are rs3218455, rs3111465, rs3094406, rs3218408, rs3218400, rs2106776, and rs3218374.

§ the bracket indicates the subdivision of the rare haplotypes into 1112112 and the rest.

Table 2

Associations of XRCC2 SNPs with survival in SBCS breast cancer patients

| SNPs | Genotypes* | No. Total/Death [†] | Log-rank p value | aHR [‡] (95% CI) |
|-----------|--------------|------------------------------|--------------------------|---------------------------|
| rs3218556 | CC | 721/178 | 0.1295 | 1 |
| | CT | 63/20 | | 1.63 (0.97–2.75) |
| rs3218536 | GG | 652/165 | 4×10⁻⁶ | 1 |
| | GA | 110/30 | | 1.19 (0.76–1.89) |
| | AA | 7/5 | | 4.26 (1.69–10.72) |
| | Per A allele | | | 1.48 (1.04–2.13) |
| rs3218534 | CC | 259/71 | 0.0224 | 1 |
| | CT | 396/93 | | 0.60 (0.42–0.85) |
| | TT | 150/41 | | 0.74 (0.46–1.19) |
| | Per T allele | | | 0.79 (0.62–1.01) |
| rs3218501 | CC | 721/193 | 0.2316 | 1 |
| | CG+GG | 54/8 | | 0.60 (0.26–1.35) |
| rs3218499 | GG | 445/122 | 0.7159 | 1 |
| | GC | 284/68 | | 0.87 (0.61–1.24) |
| | CC | 39/9 | | 0.80 (0.35–1.84) |
| rs3218455 | TT | 648/165 | 2×10⁻⁶ | 1 |
| | TC | 113/29 | | 1.16 (0.73–1.83) |
| | CC | 11/7 | | 3.86 (1.76–8.47) |
| | Per C allele | | | 1.51 (1.08–2.10) |
| rs3111465 | GG | 600/155 | 0.7031 | 1 |
| | GA+AA | 51/12 | | 0.95 (0.48–1.88) |
| rs3094406 | CC | 508/130 | 0.1816 | 1 |
| | CG+GG | 115/36 | | 1.49 (0.97–2.26) |
| rs3218408 | TT | 442/121 | 0.8815 | 1 |
| | TG | 248/61 | | 0.90 (0.62–1.30) |
| | GG | 66/16 | | 0.87 (0.47–1.63) |
| rs3218400 | CC | 586/155 | 0.6323 | 1 |
| | CA+AA | 190/47 | | 1.04 (0.71–1.51) |
| rs2106776 | CC | 177/53 | 0.1076 | 1 |
| | CT | 313/69 | | 0.79 (0.51–1.22) |
| | TT | 128/38 | | 1.26 (0.77–2.06) |
| rs3218374 | CC | 235/67 | 0.0248 | 1 |
| | CG | 365/89 | | 0.59 (0.41–0.86) |
| | GG | 157/42 | | 0.78 (0.49–1.24) |
| | Per G allele | | | 0.82 (0.64–1.05) |

* genotypes were grouped if the number of deaths were less than 3

[†] total numbers differ between SNPs due to missing genotypes[‡] adjusted for age (continuous variable) and left-censoring at recruitment

Table 3

Association of rs3218536 with breast cancer survival by study*

| Study | No. Total | No. Deaths | HR [†] (95% CI) | p value |
|---------------|-------------|-------------|--------------------------|-------------|
| ABCFS | 1223 | 270 | 1.14 (0.81–1.60) | 0.45 |
| CNIO-BC S | 190 | 6 | 1.03 (0.14–7.67) | 0.97 |
| HABCS | 598 | 86 | 0.80 (0.39–1.62) | 0.53 |
| PBCS | 1507 | 209 | 1.21 (0.83–1.77) | 0.33 |
| SBCS | 707 | 139 | 1.48 (1.04–2.13) | 0.03 |
| SEARCH | 4234 | 700 | 1.16 (0.97–1.39) | 0.11 |
| USRTS | 322 | 4 | 1.72 (0.17–17.86) | 0.65 |
| Pooled | 8781 | 1414 | 1.19 (1.05–1.36) | 0.01 |

* the number of cases do not correspond to supplementary table 1 due to some missing genotype data.

[†] adjusted for age and left-censoring at recruitment; pooled estimate stratified by study