

Published in final edited form as:

J Psychosom Res. 2013 October ; 75(4): 381–385. doi:10.1016/j.jpsychores.2013.07.014.

Emotion Suppression and Mortality Risk Over a 12-Year Follow-up

Benjamin P. Chapman, PhD, MPH¹, Kevin Fiscella, MD, MPH², Ichiro Kawachi, MD, PhD³, Paul Duberstein, PhD¹, and Peter Muennig, MD, MPH⁴

¹University of Rochester Medical Center, Department of Psychiatry, Laboratory of Personality and Development

²University of Rochester Medical Center, Departments of Family and Community and Prevention Medicine

³Harvard School of Public Health, Division of Social Epidemiology

⁴Department of Health Policy and Management, Mailman School of Public Health, Columbia University

Abstract

Objective—Suppression of emotion has long been suspected to have a role in health, but empirical work has yielded mixed findings. We examined the association between emotion suppression and all-cause, cardiovascular, and cancer mortality over 12 years of follow-up in a nationally representative US sample.

Methods—We used the 2008 General Social Survey-National Death Index (NDI) cohort, which included an emotion suppression scale administered to 729 people in 1996. Prospective mortality follow up between 1996 and 2008 of 111 deaths (37 by cardiovascular disease, 34 by cancer) was evaluated using Cox proportional hazards models adjusted for age, gender, education, and minority race/ethnicity.

Results—The 75th vs. 25th percentile on the emotional suppression score was associated with hazard ratio (HR) of 1.35 (95% Confidence Interval [95% CI] = 1.00, 1.82; $p = .049$) for all-cause mortality. For cancer and cardiovascular disease mortality, the HRs were 1.70 (95% CI = 1.01, 2.88, $p = 0.049$) and 1.47 (95% CI = .87, 2.47, $p = 0.148$) respectively.

Conclusions—Emotion suppression may convey risk for earlier death, including death from cancer. Further work is needed to better understand the biopsychosocial mechanisms for this risk, as well as the nature of associations between suppression and different forms of mortality.

Keywords

Emotion; Suppression; All-cause mortality; Cancer mortality; Cardiovascular disease mortality; General Social Survey

© 2013 Elsevier Inc. All rights reserved.

Address Correspondence to: Ben Chapman, PhD, MPH, University of Rochester Medical Center, Department of Psychiatry, 300 Crittenden, Rochester, NY 14642, fax: 585-273-1084, ben_chapman@urmc.rochester.edu.

Competing Interest : None to declare.

Publisher's Disclaimer: This is a PDF file of an unedited manuscript that has been accepted for publication. As a service to our customers we are providing this early version of the manuscript. The manuscript will undergo copyediting, typesetting, and review of the resulting proof before it is published in its final citable form. Please note that during the production process errors may be discovered which could affect the content, and all legal disclaimers that apply to the journal pertain.

Background

Emotion suppression, defined as a tendency to inhibit the expression of emotion (1), has long been suspected to influence health (2), with recent meta-analytic evidence linking suppression and chronic disease supportive of this long-held notion (3). Emotion suppression involves intentionally avoiding distressing feelings by thinking of other things or holding things in, while emotion repression is defined by lack of conscious awareness of negative emotion (4, 5).

Suppression is believed to operate on health first at a behavioral level, by inducing unhealthy coping behaviors such as over-eating as substitutes for healthy emotional expression (6). Second, at a physiological level, higher levels of autonomic reactivity to stress--measured both electrodermally and through blood pressure changes--have been reported among suppressors (7). Direct correlations between suppressive defensive styles and both catecholamines and glucocorticoids have also been reported (8, 9) and are reviewed in (10) and (11). In turn, neuroendocrine dysregulation, whether induced by stress processes or habitual health-damaging behaviors, has been implicated in the progression of a number of chronic diseases, and ultimately earlier death (12).

Epidemiologic evidence for links between suppression and mortality appeared initially in a Yugoslavian cohort study conducted in the 1970 by Grossarth-Maticeck (13). Specifically, a suppression-prone personality style called “anti-emotionality” predicted 10-year all-cause and Cardiovascular Disease (CVD) mortality (13). In other work, Grossarth-Maticeck noted associations between this personality style and cancer death (13, 14). Grossarth-Maticeck’s studies were subsequently the center of controversy around data collection and analysis (15–24), although he and his collaborator Hans Eysenck vigorously defended the results (25). “Type C” personality style—also defined by a tendency to suppress emotion—was linked to poor health outcomes in the 1980s (26–28). In the 1990s “Type D” personality, which involves affective distress in conjunction with social inhibition (presumably limiting emotional disclosure to others), was linked to both CVD death and numerous other health problems (29). Other studies on the suppression of anger in particular have noted increased all-cause mortality over 17 years in a US community sample (30), 6 years in Dutch CVD patients (31), and 8.5 years in a German sample (32). Yet there may also be a strong cultural contingency of suppression effects: in a Japanese community sample, lower levels of emotion suppression were linked to worse health (33), and in Japanese cancer patients, moderate, rather than high or low suppression levels were associated with survival (34). The construct of repression conveyed a survival advantage in male veterans of the American army followed for 16 years (35), hinting at differential acculturation and gender variability in mortality risk of constructs in the suppression/repression family (6).

In short, while studies on the mortality risk of emotion suppression have been suggestive, they have been far from definitive, underscoring the need for new data in broadly diverse population samples. Our primary aim was to examine whether emotion suppression was associated with death from any cause over a 12-year follow-up in a nationally-representative US sample. We also assessed links between emotion suppression and the two leading specific causes of death in the US, CVD and cancer. Finally, we explored whether suppression of anger in particular (30), or other indicators of more specific types of suppression, were linked to mortality.

Methods

Sample and Design

The General Social Survey (GSS) is an annual study of opinions and attitudes among the US public that is conducted by the National Opinion Research Center (NORC) at the University of Chicago. The survey uses a multi-stage probability sampling of non-institutionalized adults age 18 and over, with response rates from 70% to 82% in any given year (36). Interviews are conducted in person and involve a core set of questions asked every year (note that different people are included each year, so the survey is not a repeated-measures study). Each year, additional questions are also added for a representative subset of the panel. In 1996, the GSS Emotion Suppression Scale was administered to 737 respondents. Recently, GSS data from several years (including 1996) were linked to US National Death Index (NDI) records through 2008 (36), the standard national database for determining vital status in the US.

Measures

In addition to questions about social issues, the GSS records age, gender, race/ethnicity, and years of education on the basis of interviewer observation and subject report. The GSS' Emotion Suppression Scale (37) includes six items constructed using a face- and content-validity measurement approach (38). The scale items are: 1) "I keep my emotions to myself", 2) "I'm not afraid to let people know my feelings" (reverse scored), 3) "When I'm angry I let people know" (reverse scored), 4) "I often don't tell my friends something that I think will upset them", 5) "I try to be pleasant so that others won't get upset", and 6) "When I'm anxious I try not to worry anyone else". Responses are made on a 5-point Likert scale ranging from "Strongly Disagree" to "Strongly Agree", with items scored so that higher values indicated greater suppression. Construct validity evidence arises from associations between the scale and demographic and social factors (i.e., men suppress emotions more than women) (37, 39, 40). In our sample, Bentler's composite internal consistency reliability (41) was .70. Scale scores consisted of the mean across items (reverse coding the necessary items).

Vital status through 2008 was ascertained from the National Death Index (NDI). The validity of mortality records from the NDI is typically high, with true-positives achieved from social security numbers and the additional identifiers (used in the GSS matching process) reaching 99.8% (42). Cause of death was determined by collapsing International Classification of Disease-9 records into 285 mutually exclusive categories using the Clinical Classification Software System (CCS) (36). This was used to construct 3 outcomes: 1) death from any cause; 2) cancer death, consisting of any CCS cancer category; and 3) CVD death, including CCS categories pertaining to myocardial infarction, congestive heart failure, hypertension, coronary heart disease, and stroke. It was not possible to examine finer-grained categories due to prohibitively small numbers of deaths within each category. Further details on the GSS-NDI matching are available in (36).

Analysis

We examined the associations between emotion suppression and each mortality outcome using Cox proportional hazards models. A number of views exist on model building and confounder control. Most agree that crude associations reflect an indeterminate amount of negative (the unadjusted association is too large) or positive confounding (the unadjusted association is too small) and, in this case, provide biased estimates of the association in question (43). One view holds that each potential confounder should first be screened for bivariate associations with both a) the exposure, and b) the outcome, and included in the model only if such associations are apparent (44). The difficulty with this is that non-

significant associations may exist which still confound the estimate of interest substantially. As well, the bivariate associations between a confounder, exposure, and outcome may themselves be confounded in either direction by another variable, which cannot be detected in bivariate screening. As a result, every time a covariate is added, the covariance structure of the predictors can change, leading to data-driven searches over all possible subsets of confounders. A second alternative holds that any factors which lead to a 10% or greater change in the estimate of interest (in our case, the relative risk of emotion suppression) should be included (45). Some have pointed out that this depends on the order in which variables are entered, and that 10% changes may not place an adjusted estimate outside the 95% confidence interval of a less adjusted estimate, possibly lead to decisions based on random sampling variation (45). Yet another view, mainly used in clinical trials, holds that only factors associated with the outcome should be included because in some cases, this type of adjustment can increase the estimation efficiency (i.e., reduce the standard error) of the effect of interest (46). A fourth and final view holds that one should simply select covariates a priori, based on theory (particularly sociodemographic factors in health research) (47). Although this view too has its limitations, we prefer it in this case because there is sufficient theory and empirical results to guide a priori confounder selection. We thus controlled for the following pool of possible confounders, based on known associations with mortality risk and known (40) or potential associations with the suppression scale: age, sex, minority race, education level, and self-rated health. We did not include “mediators”—that is, variables resulting from emotional suppression, and preceding mortality on the causal pathway, since our goal was to estimate the total (i.e., direct plus indirect) mortality risk associated with suppression (47). Model diagnostics screened for violations of proportional hazards, interactions among predictors, and curvilinearity of associations (48).

Since the GSS emotion suppression measure is a continuum but defined in arbitrary Likert-scale type units, scores were scaled so that a one unit difference reflected the interquartile range, namely the difference between the 75th and 25th percentile of the distribution, or the interquartile range metric often used for exposure in epidemiologic research. This scaling provides an interpretable metric for “unitless scales” while keeping them continuous, i.e. not discarding information by actually categorizing into quartiles (48). Exploratory analyses also examined the mortality risk associated with response to the anger question (#3) and other individual items. Sensitivity analyses removed all deaths within the first year to exert additional control for baseline health.

Results

Table 1 shows the sample demographics, which were similar to the 1996 US population estimates from the decennial census (49). Mortality rates were in line with trends for that period reported by the Centers for Disease Control (50). The CVD death category was dominated by myocardial infarction, coronary atherosclerosis, and other or ill-defined heart disease (together, 58% of CVD deaths). Leukemia, lung, pancreatic, and colon cancer accounted for 47% of the cancer deaths. The GSS emotion suppression scale scores were roughly normally distributed, suggesting items succeeded in capturing information well over the range of suppression present in the population.

Table 2 shows the results of models for the first aim of quantifying the impact of emotion suppression on mortality risk. Emotion suppression at the 75th vs. 25th percentile conveyed a significant elevation in risk (HR = 1.35, or a 35% increase) of death from any cause. This was comparable to the increase in mortality risk observed for 3.1 years of life expectancy (HR = 1.35). The same difference in suppression was associated with a 70% increase in risk (HR = 1.70) of death from cancer (a 5.6 year difference in life expectancy). A non-significant elevation in risk was noted for CVD death. Removing deaths within the 1st year

amplified the effect of suppression (HRs (95% CIs) = 1.42 (1.04, 1.93), $p = 0.028$ for all-cause mortality; 2.08 (1.16, 3.75), $p = 0.013$ for cancer death). Suppression risk appeared proportional over time and a weak trend suggested suppression might confer greater risk for all-cause mortality in women ($p = 0.115$ for an interaction with gender), but this was not present for cancer or CVD death, and no other interactions or curvilinear associations were noted.

In exploratory analyses, those reporting higher anger suppression exhibited elevations in mortality risk (HR (95% CI) for 1 Likert point increase = 1.21 (1.02–1.43), $p = 0.029$ for all cause; 1.44 (1.03, 2.00), $p = 0.031$ for cancer; 1.43 (1.06, 1.91), $p = 0.018$ for CVD). Cancer death risk was significantly higher among those disagreeing more with “I’m not afraid to let people know my feelings” (that is, those reporting greater suppression; HR = 1.26, 95% CI = 1.04, 1.52, $p = 0.017$ for 1 Likert point shift).

Conclusions

Our analysis of a US nationally representative sample, followed for 12 years for mortality by cause of death, revealed significant associations between higher levels of emotion suppression and all-cause as well as cancer-related mortality. These findings have several implications. Theoretically, suppression is presumed to promote unhealthy behaviors as a substitute for appropriate emotional expression, and possibly engender neuroendocrine dysregulation (2, 51). However, whether any such biological costs are strong enough to ultimately influence mortality risk has been less certain. Our results contribute to the weight of evidence that the effects of suppression are detectable far down the progression of lifecourse health pathways, at their final common endpoint.

Psychosomatic theory and data have also suggested that emotion suppression may be implicated in cancer death, operating either through disease onset and/or course (14, 26, 52, 53), a hypothesis with which our findings are consistent. Grossarth-Maticeck defined suppression in terms of a “rationality-antiemotionality” personality tendency related to need for control, and reported that it had a potential role in cancer onset and/or survival (54). Observational studies subsequently lent some empirical support to the idea (55–57), although meta-analytic conclusions of a negligible association (58) sparked controversy further debate over the issue (59). Similarly, trials of supportive-expressive therapy, designed to reduce suppression, have yielded both positive (52, 53) and more ambiguous findings (60, 61) with respect to cancer survival (see also (62–64)). Thus, findings both with respect to incidence and death in the population (15–23), as well as survival in those with cancer (61) have been debated. Our results concern cancer mortality in the population—that is, both onset, i.e. incidence and course of the disease.

Findings for CVD death were in a direction and of a magnitude consistent with prior work (13, 30, 32), though they did not reach statistical significance. Exploratory item analysis showed that increasing disagreement with the “When I’m angry I let people know” increased risk across all three outcomes, and increasing disagreement with “I’m not afraid to let people know my feelings” elevated cancer mortality risk. Both items make reference to others, underscoring the notion that emotional expression is fundamentally an interpersonal activity, although it can be done intrapersonally via writing exercises (65–67). Thus, these findings suggest that the capacity to reduce or relieve threatening or burdensome affects by disclosure to (and processing with) others may be related to the broad health-protective effects attached to social support (68). Persons who are reluctant to disclose their emotions to others may not elicit empathic responses from others. At the same time, disclosure is also incentivized or disincentivized by social environments, so suppression must be interpreted within a person’s particular social context and history.

Based on notions such as these, some have suggested interventions to reduce suppression. Studies of supportive-expressive therapy indicate that suppression and distress it invokes can be reduced (63, 64), while other work has shown salutary immune and neuroendocrine effects of expressive writing (65–67). Care must be exercised in deciding whether the present findings support the use of such interventions to explicitly reduce mortality. It is possible, for instance, that suppression is associated only with the onset of cancer, in which case initiating supportive-expressive therapy for those diagnosed with cancer may not be helpful. Conversely, disease prevention efforts will not be useful if suppression is implicated only in the course of cancer or other diseases, rather than its onset. However, treatments such as supportive-expressive therapy may yield improved quality of life by reducing distress regardless of whether or not they influence survival.

Conclusions must also be informed by a balanced analysis of other study strengths and limitations. The study involved a nationally representative sample, maximizing ecological validity for its target population, but there were only 34 cancer and 37 CVD deaths, pointing to the need for cautious conclusions and further replication. Nevertheless, our power was sufficient to detect modest to large associations for 2 of the 3 mortality outcomes. Future work might implement larger samples, which would also permit examination of finer-grained types of mortality and afford greater power for interactions with gender and social environmental moderators of suppression risk (6). Our item analysis was also exploratory, and intended to inform future work rather than formally test a priori hypotheses. We also examined suppression with a measure focusing on suppression of disclosure of emotion, and it is not clear to what extent findings may extend to repression. There is a clear conceptual distinction between suppression and repression, and based on that body of theory, one would not necessarily expect findings from one to reflect the other. Nevertheless, empirical overlap may occur due to similar measurement, i.e., scales that cannot completely distinguish the two processes. Suppression is also similar to, but distinct, from alexithymia which has also been linked to mortality (69, 70): the latter denotes an inability to label, verbalize or communicate one's emotional experience (71), rather than a willingness or desire to suppress it. The concept of *emotion inhibition* has also been treated virtually synonymously with suppression (6). We suspect that measures of these constructs may behave similarly in an epidemiologic context, although a conservative interpretation of our findings would restrict generalization to suppression per se. It is also unknown whether suppression is functioning as a proxy for some other psychosocial factor linked to mortality or for a health factor not captured by self-rated general health.

We also did not examine mechanisms for the association between suppression and mortality. An important future direction of research would entail a randomized trial of interventions to reduce emotional suppression to better define the causal link between emotional suppression, health, and mortality. Such a study would also permit a better understanding of the underlying behavioral and physiological (72) mechanisms linking emotional suppression to mortality. For instance, the suppression / repression family of constructs might result in persistent deleterious HPA activity, or persons with systems prone to such activity might also be prone to emotional suppression (8, 9). It is also possible that emotion suppression, particularly in extreme form, may reflect underlying psychopathology that is the true source of the mortality risk. Identification of specific mechanisms, e.g. distinguishing between intrapersonal emotional expression and interpersonal emotional expression, will facilitate design of interventional research to improve health.

Finally, we caution against interpretations of these findings—particular the cancer mortality risk—as indicative of an overly specific link between a particular psychological phenomenon, and a particular cause of death. This so-called “doctrine of specific etiology” arose from the work of Franz Alexander and undergirded earlier psychosomatic theory (13,

73), but posits a degree of specificity rarely observed, and our results are not particularly supportive of it. Similarly, generalizations to populations differing from the general US public, or beyond a 12-year follow-up span, must be resisted. Due to lower power, the study likely errors on the side of Type II error (i.e., missing an association) for associations of smaller magnitude. Study strengths included a nationally representative sample, a follow-up period of 12 years, and sufficient death rates to detect effects of a clinically meaningful size.

In conclusion, our findings suggest that emotion suppression warrants more detailed investigation as a possible mortality risk. It is a construct falling outside of many personality-based studies of longevity (74), yet as a coping or defensive process, might be considered an integral part of, rather than simply a product, of personality (1, 75). In addition to pursuing further evidence for basic mortality associations, future work can further delineate the biopsychosocial pathways through which inhibiting emotional expression leads to earlier death.

Acknowledgments

Support: US National Institutes of Health grants K08031328 (BC), 1RC2MD004768 (PM)

References

1. Vaillant, G. *Ego Mechanisms of Defense: A Guide for Clinicians and Researchers*. Washington, DC: American Psychiatric Press; 1992.
2. Duberstein PR, Masling JM. Psychodynamic perspectives on sickness and health. 2000;363.
3. Mund M, Mitte K. The costs of repression: A meta-analysis on the relation between repressive coping and somatic diseases. *Health Psychol*. 2011;10.1037/a0026257
4. Giese-Davis J, Spiegel D. Suppression, repressive-defensiveness, restraint, and distress in metastatic breast cancer: separable or inseparable constructs? *J Pers*. 2001; 69(3):417–49. [PubMed: 11478732]
5. Myers LB, Vetere A, Derakshan N. Are suppression and repressive coping related? *Pers Individ Dif*. 2004; 36(5):1009–13.
6. Consedine NS, Magai C, Bonanno GA. Moderators of the emotion-inhibition-health relationship: A review and research agenda. *Review of General Psychology*. 2002; 6(2):204–28.
7. Cramer, P. Stress, autonomic nervous system reactivity, and defense mechanisms. In: Hentschel, U.; Smith, G.; Dragus, JG.; Ehlers, W., editors. *Defense Mechanisms: Theoretical, Research, and Clinical Perspectives*. Elsevier; 2004. p. 325-50.
8. Giese-Davis J, DiMiceli S, Sephton SE, Spiegel D. Emotional expression and diurnal cortisol slope in women with metastatic breast cancer in supportive-expressive group therapy. *Biol Psychol*. 2006; 73(2):190–8. [PubMed: 16750288]
9. Giese-Davis JSS, Abercrombie HC, Duran REF, Spiegel D. Repression and high anxiety are associated with aberrant diurnal cortisol rhythms in women with metastatic breast cancer. *Health Psychology*. 2004; 23(6):645–50. [PubMed: 15546233]
10. Hentschel, U.; Smith, G.; Dragus, JG. *Defense Mechanisms: Theoretical, Research, and Clinical Perspectives*. Amsterdam: Elsevier; 2004. Defense mechanisms and their psychophysiological correlates.
11. Singer, JL., editor. *Repression and dissociation: Implications for personality theory, psychopathology, and health*. Chicago, IL: University of Chicago Press; 1995.
12. Korte SM, Koolhaas JM, Wingfield JC, McEwen BS. The Darwinian concept of stress: benefits of allostasis and costs of allostatic load and the trade-offs in health and disease. *Neurosci Biobehav Rev*. 2005; 29(1):3–38. [PubMed: 15652252]
13. Grossarth-Maticke R, Bastiaans J, Kanazir DT. Psychosocial factors as strong predictors of mortality from cancer, ischaemic heart disease and stroke: the Yugoslav prospective study. *J Psychosom Res*. 1985; 29(2):167–76. [PubMed: 4009517]

14. Grossarth-Maticek R. Psychosocial predictors of cancer and internal diseases. An overview. *Psychother Psychosom.* 1980; 33(3):122–8. [PubMed: 7384381]
15. Vetter H. Some observations on Grossarth-Maticek's data base. *Psychological inquiry.* 1991; 2(3): 286–7.
16. Van der Ploeg HM. What a wonderful world it would be: a reanalysis of some of the work of Grossarth-Maticek. *Psychological inquiry.* 1991; 2(3):280–5.
17. Temoshok L. Assessing the assessment of psychosocial factors. *Psychological inquiry.* 1991; 2(3): 276–80.
18. Spiegel D. Second thoughts on personality, stress, and disease. *Psychological inquiry.* 1991; 2(3): 266–8.
19. Schuler G, Fox BH. Questions About Grossarth-Maticek's Procedures and Results. *Psychological inquiry.* 1991; 2(3):257–61.
20. Lee P. Personality and disease: a call for replication. *Psychological inquiry.* 1991; 2(3):251–3.
21. Fox BH. Quandaries created by unlikely numbers in some of Grossarth-Maticek's studies. *Psychological inquiry.* 1991; 2(3):242–7.
22. Derogatis LR. Personality, stress, disease, and bias in epidemiologic research. *Psychological inquiry.* 1991; 2(3):238–42.
23. Amelang M. Tales from Crvenka and Heidelberg: what about the empirical basis? *Psychological inquiry.* 1991; 2(3):233–6.
24. Coyne JC, Stefanek M, Palmer SC. Psychotherapy and survival in cancer: the conflict between hope and evidence. *Psychol Bull.* 2007; 133(3):367. [PubMed: 17469983]
25. Eysenck HJ. Reply to criticisms of the Grossarth-Maticek studies. *Psychological Inquiry.* 1991; 2(3):297–323.
26. Fox BH, Temoshok L, Dreher H. Mind-body and behavior in cancer incidence. *Advances.* 1988; 5(4):41–6.
27. Butow PN, Hiller JE, Price MA, Thackway SV, Krickler A, Tennant CC. Epidemiological evidence for a relationship between life events, coping style, and personality factors in the development of breast cancer. *J Psychosom Res.* 2000; 49(3):169–81. [PubMed: 11110988]
28. McKenna MC, Zevon MA, Corn B, Rounds J. Psychosocial factors and the development of breast cancer: a meta-analysis. *Health Psychol.* 1999; 18(5):520–31. [PubMed: 10519468]
29. Mols F, Denollet J. Type D personality among noncardiovascular patient populations: a systematic review. *General hospital psychiatry.* 2010; 32(1):66–72.10.1016/j.genhosppsych.2009.09.010 [PubMed: 20114130]
30. Harburg E, Julius M, Kaciroti N, Gleiberman L, Schork MA. Expressive/suppressive anger-coping responses, gender, and types of mortality: a 17-year follow-up (Tecumseh, Michigan, 1971–1988). *Psychosom Med.* 2003; 65(4):588–97. [PubMed: 12883109]
31. Denollet J, Gidron Y, Vrints CJ, Conraads VM. Anger, suppressed anger, and risk of adverse events in patients with coronary artery disease. *Am J Cardiol.* 2010; 105(11):1555–60. S0002-9149(10)00101-3 [pii]. 10.1016/j.amjcard.2010.01.015 [PubMed: 20494661]
32. Sturmer T, Hasselbach P, Amelang M. Personality, lifestyle, and risk of cardiovascular disease and cancer: Follow-up of population based cohort. *BMJ: British Medical Journal.* 2006; 332(7554): 1359.
33. Terada K, Kawakami N, Inaba S, Takatsuka N, Shimizu H. Rationality/antiemotionality personality and selected chronic diseases in a community population in Japan. *J Psychosom Res.* 2000; 48(1):31–5. [PubMed: 10750627]
34. Hirokawa K, Nagata C, Takatsuka N, Shimizu H. The relationships of a rationality/antiemotionality personality scale to mortalities of cancer and cardiovascular disease in a community population in Japan. *J Psychosom Res.* 2004; 56(1):103–11. S0022399903000461 [pii]. 10.1016/S0022-3999(03)00046-1 [PubMed: 14987971]
35. Boscarino JA, Figley CR. The impact of repression, hostility, and post-traumatic stress disorder on all-cause mortality: a prospective 16-year follow-up study. *J Nerv Ment Dis.* 2009; 197(6):461–6.10.1097/NMD.0b013e3181a61f3e [PubMed: 19525749]

36. Muennig P, Johnson G, Kim J, Smith TW, Rosen Z. The general social survey-national death index: an innovative new dataset for the social sciences. *BMC research notes*. 2011; 4:385.10.1186/1756-0500-4-385 [PubMed: 21978529]
37. Ross CE, Mirowsky J. Age and the balance of emotions. *Soc Sci Med*. 2008; 66(12):2391–400.10.1016/j.socscimed.2008.01.048 [PubMed: 18339465]
38. Anastasi, A.; Urbina, S. *Psychological Testing*. 7. Englewood Cliffs, NJ: Prentice Hall; 1997.
39. Mirowsky J, Ross CE. Sex differences in distress: real or artifact. *American Sociological Review*. 1995; 60(3):449–26.
40. Ferriss AL. Studying and measuring civility: a framework, trends, and scale. *Sociological Inquiry*. 2002; 72(3):376–92.
41. Bentler PM. Alpha, dimension-free, and model-based internal consistency reliability. *Psychometrika*. 2009; 74(1):137–43. [PubMed: 20161430]
42. Hermansen SW, Leitzmann MF, Schatzkin A. The impact on National Death Index ascertainment of limiting submissions to Social Security Administration Death Master File matches in epidemiologic studies of mortality. *Am J Epidemiol*. 2009; 169(7):901–8. kwn404 [pii]. 10.1093/aje/kwn404 [PubMed: 19251755]
43. Rothman, KJ.; Greenland, S.; Lash, TL. *Modern Epidemiology*. 3. Philadelphia, PA: Lippincott Williams & Wilkins; 2008.
44. Greenland S. Modeling and variable selection in epidemiologic analysis. *Am J Public Health*. 1989; 79(3):340–9. [PubMed: 2916724]
45. Maldonado G, Greenland S. Simulation study of confounder-selection strategies. *Am J Epidemiol*. 1993; 138(11):923–36. [PubMed: 8256780]
46. Piantidosi, S. *Clinical Trials: A Methodologic Perspective*. 2. New York: John Wiley & Sons; 2005.
47. Babyak MA. What you see may not be what you get: a brief, nontechnical introduction to overfitting in regression-type models. *Psychosom Med*. 2004; 66(3):411–21. [PubMed: 15184705]
48. Harrell, FE, Jr. *Regression Modeling Strategies*. New York: Springer-Verlag; 2001.
49. Census USBot. *Current Population Reports*. Washington, DC: 1996.
50. Center for Disease Control NCFHS. *Healthy People 2000 Review 1995/6*. Washington, DC: 1996.
51. Kreitler, S. *Defense Mechanisms and Physical Health*. In: Hentschel, UGS.; Draguns, mJG.; Ehlers, W., editors. *Defense Mechanisms*. Amsterdam: Elsevier; 2004.
52. Spiegel D, Bloom JR, Kraemer HC, Gotthel E. Effect of psychosocial treatment on survival of patients with metastatic breast cancer. *Lancet*. 1989; 2(8668):888–91.10.1016/S0140-6736(89)91551-1 [PubMed: 2571815]
53. Goodwin PJ, Leszcz M, Ennis M, Koopmans J, Vincent L, Guther H, et al. The effect of group psychosocial support on survival in metastatic breast cancer. *N Engl J Med*. 2001; 345(24):1719–26.10.1056/NEJMoa011871 [PubMed: 11742045]
54. van der Ploeg HM, Kleijn WC, Mook J, van Donge M, Pieters AM, Leer JW. Rationality and antiemotionality as a risk factor for cancer: concept differentiation. *J Psychosom Res*. 1989; 33(2): 217–25.10.1016/0022-3999(89)90049-4 [PubMed: 2724198]
55. Weihs KL, Enright TM, Simmens SJ, Reiss D. Negative affectivity, restriction of emotions, and site of metastases predict mortality in recurrent breast cancer. *J Psychosom Res*. 2000; 49(1):59–68. [PubMed: 11053605]
56. Watson M, Haviland JS, Greer S, Davidson J, Bliss JM. Influence of psychological response on survival in breast cancer: a population-based cohort study. *Lancet*. 1999; 354(9187):1331–6. [PubMed: 10533861]
57. Watson M, Homewood J, Haviland J, Bliss JM. Influence of psychological response on breast cancer survival: 10-year follow-up of a population-based cohort. *Eur J Cancer*. 2005; 41(12): 1710–4.10.1016/j.ejca.2005.01.012 [PubMed: 16098457]
58. Petticrew M, Bell R, Hunter D. Influence of psychological coping on survival and recurrence in people with cancer: Systematic review. *British Medical Journal*. 2002; 325:1066–75.10.1136/bmj.325.7372.1066 [PubMed: 12424165]

59. Watson MD-HJ, Haviland J, Bliss J. Psychological coping and cancer: Study results should not have been dismissed. *British Medical Journal*. 2003; 326:598. [PubMed: 12637410]
60. Kissane D, Li Y. Effects of supportive-expressive group therapy on survival of patients with metastatic breast cancer: a randomized prospective trial. *Cancer*. 2008; 112(2):443–4. author reply 4. 10.1002/cncr.23179 [PubMed: 18058897]
61. Spiegel D, Butler LD, Giese-Davis J, Koopman C, Miller E, DiMiceli S, et al. Effects of supportive-expressive group therapy on survival of patients with metastatic breast cancer: a randomized prospective trial. *Cancer*. 2007; 110(5):1130–8. 10.1002/cncr.22890 [PubMed: 17647221]
62. Spiegel D. Mind matters in cancer survival. *Jama*. 2011; 305(5):502–3. 10.1001/jama.2011.69 [PubMed: 21285429]
63. Giese-Davis J, Koopman C, Butler LD, Classen C, Cordova M, Fobair P, et al. Change in emotion-regulation strategy for women with metastatic breast cancer following supportive-expressive group therapy. *J Consult Clin Psychol*. 2002; 70(4):916–25. 10.1037//0022-006X.70.4.916 [PubMed: 12182275]
64. Classen C, Butler LD, Koopman C, Miller E, DiMiceli S, Giese-Davis J, et al. Supportive-expressive group therapy and distress in patients with metastatic breast cancer: a randomized clinical intervention trial. *Arch Gen Psychiatry*. 2001; 58(5):494–501. 10.1001/archpsyc.58.5.494 [PubMed: 11343530]
65. Campbell RS, Pennebaker JW. The secret life of pronouns: flexibility in writing style and physical health. *Psychol Sci*. 2003; 14(1):60–5. [PubMed: 12564755]
66. Rosenberg HJ, Rosenberg SD, Ernstoff MS, Wolford GL, Amdur RJ, Elshamy MR, et al. Expressive disclosure and health outcomes in a prostate cancer population. *International journal of psychiatry in medicine*. 2002; 32(1):37–53. [PubMed: 12075915]
67. Petrie KJ, Fontanilla I, Thomas MG, Booth RJ, Pennebaker JW. Effect of written emotional expression on immune function in patients with human immunodeficiency virus infection: a randomized trial. *Psychosom Med*. 2004; 66(2):272–5. [PubMed: 15039514]
68. Thoits PA. Mechanisms linking social ties and support to physical and mental health. *J Health Soc Behav*. 2011; 52(2):145–61. 10.1177/0022146510395592 [PubMed: 21673143]
69. Grandi S, Sirri L, Wise TN, Tossani E, Fava GA. Kellner's emotional inhibition scale: a clinimetric approach to alexithymia research. *Psychother Psychosom*. 2011; 80(6):335–44. 10.1159/000328576 [PubMed: 21829045]
70. Kauhanen J, Kaplan GA, Cohen RD, Julkunen J, Salonen JT. Alexithymia and risk of death in middle-aged men. *J Psychosom Res*. 1996; 41(6):541–9. [PubMed: 9032717]
71. Honkalampi K, Lehto SM, Koivumaa-Honkanen H, Hintikka J, Niskanen L, Valkonen-Korhonen M, et al. Alexithymia and tissue inflammation. *Psychother Psychosom*. 2011; 80(6):359–64. 10.1159/000327583 [PubMed: 21829048]
72. Muennig P, Sohler N, Mahato B. Socioeconomic status as an independent predictor of physiological biomarkers of cardiovascular disease: Evidence from NHANES. *Preventive medicine*. 2007; 45(1):35–40. [PubMed: 17521717]
73. GrossarthMaticke R, Eysenck HJ, Vetter H. Personality type, smoking habit and their interaction as predictors of cancer and coronary heart disease. *Pers Individ Dif*. 1988; 9(2):479–95. Eseer Sene.
74. Chapman BP, Roberts B, Duberstein P. Personality and longevity: knowns, unknowns, and implications for public health and personalized medicine. *Journal of aging research*. 2011; 201110.4061/2011/759170
75. Haan, N. *Coping and Defending*. New York: Academic Press; 1977.

Table 1

Descriptive Statistics: Demographics and Cause of Deaths

Variable	Mean / N	SD / %
Age in 1996	44.0	16.5
Education (Years)	13.5	2.9
Female Gender	395	54%
White	593	81%
Black	102	14%
Other Minority	34	5%
Emotion Suppression *	2.9	.6
Self-Rated Health **	2.0	.8
Death, Any Cause, by 2008	111	15%
CVD Death by 2008	37	5%
Cancer Death by 2008	34	5%

Note. N = 729.

* Mean score on scale items, ranging from 1 to 5. A 1-unit increase corresponds to 1 Likert scale point (i.e., from 1, strongly disagree to 2, disagree) increase in agreement on average to admission of emotion suppression behavior.

** Mean score on scale of 1 (excellent) 2 (good) 3 (fair) 4 (poor).

Table 2

Multivariate Survival Models

	All-Cause Mortality (N=111)			Cancer Mortality (N=34)			CVD Mortality (N=37)		
	HR	95% CI	P	HR	95% CI	P	HR	95% CI	P
Age (Decades)	1.98	[1.73, 2.26]	<.001	2.19	[1.70, 2.81]	<.001	2.36	[1.84, 3.02]	<.001
Female Gender	.78	[.54, 1.14]	.219	.64	[.32, 1.27]	.202	1.04	[.54, 2.01]	.897
Black Race/Ethnicity	1.06	[.60, 1.88]	.339	.62	[.18, 2.13]	.446	.73	[.24, 2.19]	.571
Other Race/Ethnicity	2.95	[1.03, 8.47]	.043	2.75	[.34, 22.13]	.340	3.60	[.44, 29.44]	.231
Education (Years)	1.00	[.88, 1.06]	.881	.98	[.87, 1.10]	.680	.93	[.83, 1.05]	.241
Self-rated health	1.55	[1.20, 1.99]	<.001	1.56	[1.00, 2.46]	.050	1.35	[.87, 2.11]	.184
Emotion Suppression	1.35	[1.00, 1.82]	.049	1.70	[1.00, 2.88]	.049	1.47	[.87, 2.47]	.148

Note. HR = hazard ratio, 95% CI = 95% Confidence Intervals. Emotion suppression scaled by interquartile range, so HR corresponds to risk at 75th vs. 25th percentile of distribution. Self-rated health on 1–4 scale. #Estimate presented is from first five years of follow-up during which hazards are proportional, after which male-female (for all cause mortality) and minority-white (for CVD) risk ratios diminish then reverse. Inference for based on likelihood ratio test.