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# Variability of methacholine bronchoprovocation and the effect of inhaled corticosteroids in mild asthma

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# Abstract

**Background**—The methacholine challenge test quantifies airway hyper-responsiveness, which is measured by the provocative concentration of methacholine causing a 20% decrease in forced expiration volume in 1 second ( $PC_{20}$ ). The dose–response effect of inhaled corticosteroids (ICS) on  $PC_{20}$  has been inconsistent and within-patient variability of  $PC_{20}$  is not well established.

**Objectives**—To determine the effect of high- vs low-dose ICS on  $PC_{20}$  and within-patient variability in those with repeated measurements of  $PC_{20}$ .

**Methods**—A randomized, double-masked, crossover trial was conducted in patients with asthma on controller medications with  $PC_{20}$  of 8 mg/mL or lower (n = 64) to evaluate the effect of high-dose (1,000 µg/d) vs low-dose (250 µg/d) fluticasone for 4 weeks on  $PC_{20}$ . In addition, the variability of  $PC_{20}$  was assessed in participants who underwent 2 or 3  $PC_{20}$  measurements on the same dose of ICS (n = 27) over a 4-week interval.

**Results**—Because there was a significant period effect, dose comparison of the change in  $PC_{20}$  was assessed in the first treatment period. There was no significant difference in the change in  $PC_{20}$  for high- vs low-dose ICS (39% vs 30% increase, respectively; P = .87). The within- and between-participant variances for log  $PC_{20}$  were 0.84 and 0.96, respectively, with an intra-class correlation of 0.53, and 37% of participants had more than 2 doubling dose changes in  $PC_{20}$  in those with repeated measurements.

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A list of participants in the American Lung Association Asthma Clinical Research Centers is available in the "Acknowledgments" section.

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Supplementary Data

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**Conclusion**—The effect of ICS on  $PC_{20}$  is not dose dependent at fluticasone levels of 250 and 1,000  $\mu$ g/d. Interpersonal variability for  $PC_{20}$  is large. A lack of precise measurements should be taken into account when interpreting any change in  $PC_{20}$ .

# Introduction

The methacholine challenge test (MCT) is one of the most commonly used methods to assess and quantify airway hyper-responsiveness (AHR), a key feature of asthma. The result of this test is usually expressed as the provocative concentration of methacholine that causes a 20% decrease in forced expiration volume in 1 second (FEV<sub>1</sub>; PC<sub>20</sub>). PC<sub>20</sub> is frequently used as a tool for diagnosing asthma and as an entry criterion or outcome measurement in clinical research.<sup>1–3</sup> The authors previously reported that 23% of patients with physician-diagnosed asthma on active controller treatment have a negative MCT result and that the sensitivity of the test varies with race and presence of atopy.<sup>4</sup>

One reason for the decreased sensitivity of MCT may be the increasing use of high-potency inhaled corticosteroids (ICS) in the treatment of asthma. Although high-potency ICS can decrease AHR,<sup>5–8</sup> the results of studies investigating the dose–response relation between ICS and PC<sub>20</sub> have been inconsistent.<sup>8,9</sup> In addition, the American Thoracic Society guidelines<sup>10</sup> for MCT state that the 95% confidence interval for short-term (1- to 8-week) repeat testing for methacholine PC<sub>20</sub> is within  $\pm 1.5$  doubling doses. This information is based on older and smaller studies before the widespread use of high-potency ICS. Therefore, the authors investigated the effect of high- vs low-dose ICS on PC<sub>20</sub> in a randomized crossover trial and assessed the variability of the MCT for patients on a stable dose of high-potency ICS in a subgroup of those participants.

# Methods

#### **Participant Selection**

The study was the second part of a 2-phase clinical study. The first phase was a crosssectional study of patients with asthma and nonasthmatic healthy controls to measure the sensitivity and specificity of the MCT, which is described elsewhere.<sup>4</sup> Participants in the second phase were recruited from participants with asthma who completed the first phase and who met the eligibility criteria. Inclusion criteria for the second phase were age 12 to 69 years, physician-diagnosed stable asthma, current treatment for asthma within the preceding 12 months with regular use of asthma controller medications, no asthma exacerbation during the prior 4 weeks, an FEV1 of at least 70% predicted before using a bronchodilator, a positive MCT result (methacholine PC20 8 mg/mL) during phase 1, successful completion of the fluticasone run-in period demonstrating adequate adherence, an Asthma Control Questionnaire (ACQ)<sup>11</sup> score lower than 1.5, using fewer than 16 puffs per week of a rescue  $\beta$  agonist during the final 2 weeks of the run-in period, and no asthma exacerbation during the previous 2 weeks. Study medication adherence was monitored through the use of diary cards filled out by the participants at each visit. In addition to medication use, participants recorded the morning peak expiratory flow rate using a Mini-Wright Peak Flow Meter (Clement Clarke International, Ltd, Harlow, Essex, United Kingdom), symptoms, and rescue medication use on the same diary card. Medication adherence was defined as self-report of

taking the full assigned dose of ICS. The study was approved by the institutional review board at each participating center and all participants provided written informed consent. The trial was registered at www.clinicaltrials.gov (identifier, NCT00705341).

#### **Study Design**

The study was conducted at 15 centers in the American Lung Association Asthma Clinical Research Centers from December 2008 until May 2010. The study was a randomized, double-masked, crossover trial designed to evaluate the effect of high- vs low-dose ICS on AHR. The participant enrollment and study schema are shown in Figures 1 and eFigure 1. The trial began with a 4-week run-in period on low-dose fluticasone (Flovent Diskus, 250  $\mu$ g/d; GlaxoSmithKline, Brentford, Middlesex, United Kingdom). After completing the runin period, eligible participants (N = 62) were randomly assigned to 1 of 2 treatment sequences: low-dose (LD) fluticasone (250  $\mu$ g/d) in the first 4-week period (period A) followed by high-dose (HD) fluticasone (1,000  $\mu$ g/d) in the second 4-week period (period B), or the reverse dose order. There was a 4-week washout interval between periods A and B during which the participants were maintained on LD fluticasone. The MCT was performed before and after each treatment period. A subset of study participants (n = 27) completed 2 (n = 6) or 3 (n = 21) MCTs during 3 consecutive visits at 4-week intervals while receiving a stable dose of 250  $\mu$ g/d of fluticasone, which provided an opportunity to assess MCT variability. The reasons for a missing MCT included an FEV1 less than 70% or a missed visit, which are documented in Figure 1.

#### Method of MCT

The 5-breath dosimeter method (Koko dosimeter, nSpire Health, Longmont, Colorado) was used according to published guidelines from the American Thoracic Society,<sup>10</sup> as previously described.<sup>4</sup> Briefly, the test sequence included 11 steps: diluent only, 0.03125, 0.0625, 0.125, 0.25, 0.5, 1.0, 2.0, 4.0, 8.0, 16.0, and 32.0 mg/mL. All medications and foods that might interact with methacholine were held before the test, including short-acting bronchodilators (6 hours), long-acting bronchodilators (24 hours), leukotriene modifiers (24 hours), antihistamines (48 hours), and theophylline (48 hours). Quality control measures were implemented centrally by the data coordinating center, including the calibration of each nebulizer cup to deliver 9  $\mu$ L ± 10% per 0.6-second actuation and certification of the technicians who performed the MCT.

#### **Statistical Analysis**

Comparisons between randomized groups were performed using Kruskal-Wallis tests for continuous variables and  $\chi^2$  tests or Fisher exact tests for categorical variables. For plots and summary statistics, PC<sub>20</sub> was imputed to be 48 mg/mL (1.5 times the highest dose inhaled) for those who did not exhibit a 20% decrease in FEV<sub>1</sub> at the maximum concentration tested (32 mg/mL). For all other analyses, multiple imputation<sup>12</sup> was used to adjust for the uncertainty in the unobserved PC<sub>20</sub> for these participants. Sensitivity analyses applying a range of imputation models were performed to assess the robustness of the estimates. Repeated measurement log-linear generalized estimating equation models were used to assess the dose effect on PC<sub>20</sub>. Variability of the MCT was measured using between-

and within-subject variances to construct an intraclass correlation coefficient (ICC) for participants on a stable ICS dose. *P* values less than .05 (2-sided) were considered statistically significant.

#### Results

#### **Characteristics of Study Participants**

Sixty-two participants with mild asthma on a single controller medication (30% receiving ICS alone) or combination controller therapy (60%) were included (Table 1). The participants' overall asthma control at enrollment was suboptimal, with an average ACQ of 1.0 (n = 26 [42%] with a score 0.76, defined as well controlled).<sup>13</sup> Over the course of the entire trial, medication adherence was high, with slightly higher rates for patients randomized to HD and then LD ICS compared with those randomized to LD and then HD ICS (93% vs 90%, P < .0001); however, the magnitude of the difference was small. During the run-in phase and phase A, the adherence rates were 94% and 89%, respectively, and did not differ significantly between groups.

#### Influence of Dose of ICS on PC<sub>20</sub>

Participants were randomized to LD ICS followed by HD ICS (LD/HD, n = 30) or HD ICS followed by LD ICS (HD/LD, n = 32). The demographic and asthma characteristics were similar between these groups at randomization (Table 1).

There was a significant period effect (difference in the percentage change in PC<sub>20</sub> for HD and LD ICS depending on the order in which the doses were administered; test of interaction, P = .019; Table 2). When HD ICS was administered first, the PC<sub>20</sub> was slightly lower after HD ICS than after LD ICS (1.19 vs 1.25 mg/mL); conversely, in the other group, the PC<sub>20</sub> was higher after HD ICS than after LD ICS (4.31 vs 2.04 mg/mL). To remove the effect of the order in which the treatments were received and decrease the potential influence of treatment carryover, insufficient follow-up time, and loss to follow-up, the HD and LD MCT results were compared exclusively during period A, ignoring period B, and the results were analyzed as parallel groups. There was no difference in the percentage of participants with a positive MCT result at the end of period A (89% vs 88%, P > .99). In addition, there was no difference in the change in PC<sub>20</sub> for HD vs LD ICS (39% vs 30% increase, P = .87). Sensitivity analyses with a range of alternate PC<sub>20</sub> imputation structures produced similar results.

#### **MCT Variability**

The inter- and intra-subject variabilities of  $PC_{20}$  were evaluated during 3 consecutive visits for 27 participants who completed 2 (n = 6) or 3 (n = 21) MCTs after the initial positive MCT result at entry (methacholine  $PC_{20}$  8 mg/mL). These tests were conducted at 4-week intervals while participants received a stable dose of 250 µg/d of fluticasone. Over this 12week period across multiple MCTs, the within- and between-subject variances for log  $PC_{20}$ were 0.84 and 0.96, respectively, resulting in an ICC of 0.53. Furthermore, without change in the ICS dose, 74% of participants had more than 1 doubling dose change in their  $PC_{20}$ and 37% had more than 2 doubling dose changes (Fig 2).

#### Factors Associated with Increased Variability

The association of various participant characteristics at baseline with variance of PC<sub>20</sub> was examined. Older age, defined as older than 50 vs 50 years or younger, was associated with greater variance (difference in location 2.23, 95% confidence interval 1.62-2.59, P = .0001) and the presence of atopy was associated with lower variance (difference in location -2.18, 95% confidence interval -2.58 to -0.23, P = .0228). There was no significant difference in the variability for participants providing 2 and 3 measurements (P = .84). Other risk factors that were examined but did not have a significant impact on variability included sex, race, body mass index greater than 30 kg/m<sup>2</sup>, a family history of asthma, predicted FEV<sub>1</sub> greater than 400%, and an ACQ score lower than or equal to 0.76; all had minimal impact (absolute difference ranges 0.010-0.075, P = 0.55-0.98). Use of a prednisone burst (difference 0.36, P = .17) or an urgent care visit (difference -0.22, P = .20) in the past 12 months or peak flow greater than 400 (difference -0.23, P = .39) had a larger effect on variability but were not statistically significant. There was no significant difference in the variability but were same of enrollment (P = .74), but the sample size available for comparison was small.

# Discussion

In the present study, the authors made several and clinically important observations about  $PC_{20}$  measured in MCT: (1) change to a dose of a high-potency ICS over a short period had little effect on AHR in participants already taking LD ICS, (2) there is a large amount of intra-subject variability with MCT in the same patients on a stable dose of ICS, and (3) age and atopy are associated with MCT variability.

The authors did not observe a dose-dependent change in  $PC_{20}$  when comparing LD with HD ICS in participants with asthma over a 4-week interval. This is in agreement with a metaanalysis including 11 studies in steroid-naive patients with mild asthma that did not show a dose-response effect,<sup>7</sup> but conflicts with a meta-analysis of 25 placebo-controlled trials evaluating the dose-response of ICS on AHR in patients with asthma with varying degrees of disease severity that concluded that higher doses of ICS produced greater improvement in AHR.<sup>8</sup> The lack of a consistent dose-dependent pattern of change in PC<sub>20</sub> in the present study could be because the treatment period with ICS was too short to affect AHR. The authors chose a treatment period of 4 weeks based on the fact that the initial ICS effect on PC<sub>20</sub> can be seen as soon as 1 week.<sup>14,16</sup> It is also possible the authors did not detect an effect of HD ICS on AHR in patients with mild persistent asthma owing to a ceiling effect, such as having limited room for improvement in those who have only mild disease. Similarly, the dose-dependent nature of the relation may be observable only for lower doses of ICS than were used in this study, with a plateau in the effect at levels of ICS at or above what was defined as a low dose. An ideal scientific design would have been to take the participants completely off their ICS during the run-in and washout periods. However, this was not possible because the present study population involved those with mild to moderate asthma in whom daily ICS treatment was indicated as standard of care. Moreover, the large degree of intra-subject variability may have prevented detection of a dose-dependent effect of ICS on PC<sub>20</sub>.

Although this study was not specifically designed to evaluate the reproducibility of  $PC_{20}$ measurements, the authors were able to make observations on the variability of  $PC_{20}$  in a subset of participants who underwent multiple MCTs on a stable low dose of ICS. A metaanalysis evaluated the reproducibility of measurements in bronchial challenge and found that the ICC was higher than 0.9 in short-term studies (<4 months), but many of the studies included were small.<sup>17</sup> However, a larger degree of intra-subject variability was seen in the present study, including 27 participants with asthma: the ICC was 0.53 and more than one third of participants had a change in their PC20 of at least 2 doubling doses. The authors speculate that this large degree of intra-subject variability of MCT may have contributed to the decreased sensitivity of MCT observed in the authors' previous study<sup>4</sup> and to the lack of a dose-dependent change in  $PC_{20}$  in the present study. This variability could be due to methodologic issues with the MCT. However, the authors believe this is unlikely because they had instituted quality control procedures with centralized certification and calibration of nebulizer cups and standardization of equipment and techniques to ensure that the MCT was performed in a standardized fashion. This variability also could be due to changes in the participants' clinical conditions. The authors could not detect a reason for the variation in PC20 from baseline characteristics; markers of asthma control and severity at baseline (ACQ score, FEV<sub>1</sub>, and peak flows) were similar for participants with high and low MCT variability. Although the present study included those with stable asthma and excluded those with poorly controlled asthma, many did not have well-controlled asthma. Therefore, dayto-day subclinical disease instability from various environmental exposures is a possible cause of this variability in PC20. Further exploration of the effect of season, which was not feasible in this study because of the small sample, also would be valuable. Further, it is likely that AHR is not a static characteristic of asthma and varies with time; after all, the periodic nature of the syndrome is a defining feature of asthma.

The authors also observed that older patients with asthma had greater variance and those with atopy had lower variance in  $PC_{20}$ . The authors are not aware of any other studies that evaluated risk factors for variation in  $PC_{20}$  in 1 patient. Recent studies have identified various different phenotypes of asthma and those with adult-onset asthma may have increased risk of health care usage,<sup>18</sup> which could be due in part to the variability of AHR. The authors were surprised that those with atopy had less variability in the serial  $PC_{20}$  measurements. This observation is in contrast with studies showing the association of the presence of an allergic condition with  $AHR^{19}$  and exposure to allergens with worsening in  $AHR.^{20}$  Because this association analysis was performed only in a small group of patients, full assessment of the reproducibility of  $PC_{20}$  in patients with asthma and the risk factors associated with variation awaits further study involving more subjects with a wide range of methacholine sensitivity and asthma severity.

There are several limitations to this study. First, some participants did not complete all MCT measurements. Efforts were made to decrease the impact of missing data using robust analytic techniques (generalized estimating equation and multiple imputation) and the focus on only period A. Second, the treatment effect was greater during the second period for the 2 doses of ICS. This may be due to an insufficient duration for the treatment and/or the washout period. A washout time of 4 weeks for ICS was chosen based on studies reporting exacerbations occurring after this period (indicating disappearance of treatment effect).<sup>21,22</sup>

Nevertheless, a differential effect was observed due to the ordering of doses. The authors attempted to mitigate the influence of the period effect by considering only the first treatment period. However, the power to detect a difference in the change in  $PC_{20}$  with a parallel design is much less than with a crossover design. Third, the study included patients with mostly mild persistent asthma on long-term asthma treatment, so the result may not apply to those with more severe asthma or those who do not require controller medications.

In summary, in patients with physician-diagnosed asthma on controller medication, no dosedependent change was observed in  $PC_{20}$  between 250- and 1,000- $\mu$ g/d ICS doses of fluticasone. A large amount of intra-subject variability of MCT was observed in patients with asthma on a stable dose of ICS, which decreases the precision of estimates of change in  $PC_{20}$ . This low-precision  $PC_{20}$  measurement would need to be taken account when using  $PC_{20}$  values for asthma diagnosis or when using  $PC_{20}$  as an outcome measurement in clinical studies.

# **Supplementary Material**

Refer to Web version on PubMed Central for supplementary material.

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#### Figure 1.

Enrollment, randomization, and follow-up of study participants. ACQ, Asthma Control Questionnaire;  $FEV_1$ , forced expiration volume in 1 second; MCT, methacholine challenge test;  $PC_{20}$ , provocative concentration of methacholine causing a 20% decrease in forced expiration volume in 1 second.



#### Figure 2.

Change in provocative concentration of methacholine causing a 20% decrease in forced expiration volume in 1 second ( $PC_{20}$ ) between visits in a subset of participants who underwent a methacholine challenge test on a stable dose of inhaled corticosteroids. Twenty-seven participants who were assigned to receive low-dose and high-dose inhaled corticosteroids underwent 2 (n = 6) or 3 (n = 21)  $PC_{20}$  measurements at 4-week intervals while receiving 250 µg/d of fluticasone during the run-in phase (visits 1 to 2 [V1 to V2]), the low-dose treatment phase (visits 2 to 3 [V2 to V3]), and the washout period (visits 3 to 4 [V3 to V4]). The methacholine challenge test was performed at the end of each interval (V2, V3, and V4). This figure shows the change in  $PC_{20}$  (measured in doubling doses) between each visit and the maximal change for each participant.

#### Table 1

#### Participant characteristics

	Total (N = 62)	Order of ICS	
		HD/LD $(n = 32)$	LD/HD (n = 30)
General characteristics			
Male patients, n (%)	22 (35)	13 (41)	9 (30)
Race/ethnicity, n (%) <sup><math>a</math></sup>			
White	35 (56)	19 (59)	16 (53)
African American	19 (31)	7 (22)	12 (40)
Latino/Hispanic	7 (11)	6 (19)	1 (3)
Other	1 (2)	0 (0)	1 (3)
Age at enrollment (y), mean ± SD	$39.2\pm15.4$	$40.6 \pm 15.3$	$37.8 \pm 15.7$
Former smoker, n (%)	6 (10)	4 (13)	2 (7)
BMI (kg/m <sup>2</sup> ), mean ± SD	$27.7\pm6.2$	$28.2\pm 6.6$	$27.2\pm5.9$
Other conditions, n (%)			
Eczema	6 (10)	4 (13)	2 (7)
Sinusitis	7 (11)	3 (9)	4 (13)
Hay fever/rhinitis	39 (63)	20 (63)	19 (63)
GERD	14 (23)	7 (22)	7 (23)
Food allergy	10 (16)	6 (19)	4 (13)
Family history of asthma, n (%)	40 (65)	19 (59)	21 (70)
Atopy			
Positive allergy skin test results, median (range)	3 (0–14)	3 (0–14)	4 (0–14)
1 positive allergy skin test reaction, n (%)	49 (82)	23 (72)	26 (93)
Lung function, mean $\pm$ SD			
Percentage of predicted $FEV_1$ (before BD)	$86.6\pm12.7$	$89.4 \pm 12.9$	$83.6 \pm 11.9$
Percentage of predicted FVC (before BD)	$94.4 \pm 11.9$	$96.3 \pm 12.4$	92.4 ± 11.1
Asthma characteristics			
Age of onset (y), median (range)	15.5 (1–53)	15.5 (1–53)	15.5 (1–52)
Asthma treatment at enrollment, n (%)			
ICS/LABA combination	37 (60)	22 (69)	15 (50)
ICS (no combination)	19 (30)	9 (28)	10(33)
No ICS	6 (10)	1 (3)	5 (17)
Anti-leukotriene	13 (21)	4 (13)	9 (30)
Asthma control			
Use of daily SABA, n (%)	12 (19)	8 (25)	4 (13)
ACQ, mean ± SD	$0.6\pm0.4$	$0.7\pm0.4$	$0.6\pm0.5$
Urgent care visits in 12 mo, n (%)	15 (24)	8 (25)	7 (23)
Prednisone bursts in 12 mo, n (%)	12 (19)	7 (22)	5 (17)

Abbreviations: ACQ, 7-item Asthma Control Questionnaire; BD, bronchodilator; BMI, body mass index; FEV1, forced expiration in 1 second; FVC, forced vital capacity; GERD, gastroesophageal reflux disease; HD/LD, high-dose and then low-dose inhaled corticosteroids; ICS, inhaled corticosteroids; LABA, long-acting  $\beta$  agonist; LD/HD, low-dose and then high-dose inhaled corticosteroids; SABA, short-acting  $\beta$  agonist.

 $^a\mathrm{Race/ethnicity}$  was determined by self-identification to best descriptive category.

.

#### Table 2

Summary of PC<sub>20</sub> measurements at each follow-up visit

	Randomization	After period A	After washout	After period B
HD then LD ICS $(n = 32)$	on LD ICS (run-in)	after HD ICS	on LD ICS (washout)	after LD ICS
PC <sub>20</sub>				
Ν	32	28	27	26
Geometric mean (raw PC <sub>20</sub> range)	0.90 (0.11-7.13)	1.19 (0.10–26.92)	0.64 (0.008–14.49)	1.25 (0.008-48.00)
$PC_{20} > 32 mg/mL^{a}$	0	0	0	1
Missing	0	4	5	5
LD then HD ICS $(n = 30)$	on LD ICS (run-in)	after LD ICS	on LD ICS (washout)	after HD ICS
PC <sub>20</sub>				
Ν	29	23	23	20
Geometric mean (raw PC <sub>20</sub> range)	1.56 (0.09–7.16)	2.04 (0.16-48.00)	1.39 (0.14–30.55)	4.31 (0.32–48.00)
$PC_{20} > 32 mg/mL^{a}$	0	2	0	4
Missing	1	5	7	6

Abbreviations: HD, high-dose; ICS, inhaled corticosteroid; LD, low-dose; PC<sub>20</sub>, provocative concentration of methacholine causing a 20% decrease in forced expiration in 1 second.

 $^{a}$ The number of participants who did not achieve PC<sub>20</sub> at 32 mg/mL is listed and the multiple imputation was used in the calculation of the geometric mean of PC<sub>20</sub> to include data from as many participants as possible in the analysis.