

NIH Public Access

Author Manuscript

Cochrane Database Syst Rev. Author manuscript; available in PMC 2014 June 25.

Published in final edited form as:

Cochrane Database Syst Rev.; 5: CD007834. doi:10.1002/14651858.CD007834.pub2.

Iron for restless legs syndrome

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Abstract

Background—Restless legs syndrome (RLS) is a common neurologic syndrome and is associated with iron deficiency in many patients. It is unclear whether iron therapy is effective treatment for RLS.

Objectives—The objective of this review was to assess the effects of iron supplementation (oral or intravenous) for patients with RLS.

Search methods—We searched the Cochrane Central Register of Controlled Trials (CENTRAL), MEDLINE (Jan 1995 to April 2011); EMBASE (Jan 1995 to April 2011); PsycINFO (Jan 1995 to April 2011); and CINAHL (Jan 1995 to April 2011). Corresponding authors of included trials and additional members of the International Restless Legs Syndrome Study Group were contacted to locate additional published or unpublished trials.

Selection criteria—Controlled trials comparing any formulation of iron with placebo, other medications, or no treatment in adults diagnosed with RLS according to expert clinical interview or explicit diagnostic criteria.

Data collection and analysis—Two review authors extracted data and at least two authors assessed trial quality. We contacted trial authors for missing data.

Main results—Six studies (192 total subjects) were identified and included in this analysis. The quality of trials was variable. Our primary outcome was restlessness or uncomfortable leg sensations, which was quantified using the IRLS severity scale in four trials and another RLS symptom scale in a fifth trial. Combining data from the four trials using the IRLS severity scale, there was no clear benefit from iron therapy (mean difference in IRLS severity scores of -3.79,

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Contributions of Authors: Lynn Marie Trotti (LMT) conceived and co-ordinated the review; performed the search; screened the search results for acceptable trials; graded and extracted data from trials; entered and analyzed data; wrote the initial draft of the review; and incorporated feedback from the other authors into the final draft. Srinivas Bhadriraju (SB) screened the search results for acceptable trials; graded and extracted data from trials; and co-wrote the review. Lorne A. Becker (LAB) designed the review; resolved disagreements between LMT and SB regarding quality of trials; analyzed data; and co-wrote the review.

Declarations of Interest: None

Differences between Protocol and Review: A funnel plot was not constructed because we identified fewer than 10 studies for inclusion.

95% CI: -7.68 to 0.10, p = 0.06). However, the fifth trial did find iron therapy to be beneficial (median decrease of 3 points in the iron group and no change in the placebo group on a 10 point scale of RLS symptoms, p = 0.01). Quality of life was improved in the iron group relative to placebo in some studies but not others. Changes in periodic limb movements were not different between groups (measured in two studies). Objective sleep quality, subjective sleep quality and daytime functioning were not different between treatment groups in the studies that assessed them. The single study of subjects with end stage renal disease did show a benefit of therapy. Most trials did not require subjects to have co-morbid iron deficiency and several excluded patients with severe anemia. The single study that was limited to iron deficient subjects did not show clear benefit of iron supplementation on RLS symptoms. There was no clear superiority of oral or intravenous delivery of iron. Iron therapy did not result in significantly more side effects than placebo (RR 1.39, 95% CI 0.85 to 2.27).

Authors' conclusions—There is insufficient evidence to determine whether iron therapy is beneficial for the treatment of RLS. Further research to determine whether some or all types of RLS patients may benefit from iron therapy, as well as the best route of iron administration, is needed.

Medical Subject Headings (MeSH)

Iron [deficiency; *therapeutic use]; Iron, Dietary [therapeutic use]; Quality of Life; Randomized Controlled Trials as Topic; Restless Legs Syndrome [*drug therapy; etiology]; Severity of Illness Index; Sleep [physiology]; Treatment Outcome

MeSH check words

Humans

Background

Description of the condition

Restless legs syndrome (RLS) is a common condition consisting of four cardinal features: an urge to move the legs (which is usually but not invariably associated with unpleasant leg sensations); at least transient relief with movement; worsening with rest; and a predilection to occur during the evening and night. The diagnosis is one of clinical judgment, with no definitive confirmatory testing presently available. However, the presence of periodic limb movements, a response to dopaminergic medications, a positive family history, and the presence of sleep disruption can all provide supportive evidence for the diagnosis (Allen 2003). Periodic limb movements are stereotyped, repetitive flexions of the great toe and ankle that occur with a periodicity of between 5 and 90 seconds. Periodic limb movements are seen in 80 to 90% of patients with RLS (Montplaisir 1997, Trotti 2009), but they are not specific to RLS because they are also found in a variety of other sleep disorders and in aging in the absence of sleep disorders (Hornyak 2006). RLS is more common in women than in men (Berger 2004) and the prevalence of RLS increases with age (Bliwise 2006). RLS has a prevalence of 5 to 15% in populations of western European descent but lower prevalence in other investigated populations such as subsets of Asia (0.1 to 5%) (Stefansson 2007). RLS has long been suspected to have a genetic basis because of the high proportion of sufferers

who have a positive family history. Familial linkage studies have identified numerous potentially causative regions (Trotti 2008). Three recent genome-wide association studies have found a total of four single nucleotide polymorphisms that conferred increased risk for RLS or periodic limb movements, located in BTBD9 (on chromosome 6), in MEIS1 (on chromosome 2), in a region on chromosome 15 containing the genes MAP2K5 and LBXCOR1, and in PTPRD (on chromosome 9) (Stefansson 2007; Winkelmann 2007; Schormair 2008). Collectively, the population attributable risk for the variants at BTBD9, MEIS1, and MAP2K5/LBXCOR1 is approximately 70% (Winkelmann 2007).

Despite this genetic advance, the pathophysiology of RLS remains somewhat unclear. What is known about the functions of the recently implicated genes does not yet explain RLS pathology. From a clinical standpoint, a state of dopaminergic dysfunction is often suspected tocause RLS because patients with RLS frequently improve when givenlowdoses of dopaminergic medications (Allen 2004).

How the intervention might work

Iron deficiency has also been implicated in the pathophysiology of RLS, based on the clinical findings that patients with iron deficiency have a higher frequency of RLS (30%) than those without and that severity of iron deficiency correlates with the severity of RLS symptoms (Earley 2000). Treatment with dopamine agonists is sometimes complicated by a phenomenon known as augmentation, in which symptoms return with earlier time of onset, greater severity, and more extensive bodily involvement (Allen 2003); this treatment complication appears more likely to occur in those patients who have iron deficiency Trenkwalder 2008). As with dopamine, iron metabolism exhibits a circadian pattern similar to RLS symptoms; serum iron nadirs occur between 8 pm and midnight, in line with the peak time of RLS symptom severity (Earley 2000). Two of the most common causes of secondary RLS, pregnancy and end-stage renal disease, are also associated with iron deficiency (Allen 2004). Studies of iron in patients with RLS have demonstrated low cerebrospinal iron levels (Earley 2000), low iron concentrations in the substantia nigra on magnetic resonance imaging (Allen 2004), and decreased substantia nigra iron stores on autopsy specimens (Connor 2003). Additionally, iron has known interactions with dopamine, acting as a cofactor for tyrosine hydroxylase, the rate-limiting enzyme in dopamine production. Iron deficient rats have been shown to have altered dopamine functioning, with decreased D2 receptors, decreased dopamine transporter functioning, and elevated levels of extracellular dopamine in the caudate nucleus and putamen (Earley 2000; Nelson 1997).

Based on this presumed pathophysiology, iron therapy has been used for RLS patients, both those with and those without documented iron deficiency. Iron supplementation can be provided orally or intravenously, with several formulations available for each method of delivery. Intravenous iron provides the advantage of replenishing iron stores more quickly than oral therapy, which can take weeks to months, but intravenous therapy carries the added risk of severe complications such as anaphylaxis (Silverstein 2004). Current consensus guidelines for RLS put forth by the RLS Foundation's Medical Advisory Board recommend iron therapy for patients with a low ferritin (usually less than 20 nanograms per

milliliter) and consideration of iron therapy, on a case-by-case basis, for those with a ferritin level below 45 to 50 nanograms per milliliter (Silber 2004). The optimal timing to assess response to such therapy is not known.

Objectives

To evaluate the efficacy and safety of oral or parenteral iron for the treatment of RLS when compared with placebo or other therapies.

Methods

Criteria for considering studies for this review

Types of studies—We searched for all controlled trials investigating the treatment of RLS with oral or parenteral iron versus placebo, versus another drug, or with a nointervention control group. We planned to include controlled trials regardless of whether or not they were randomized or blinded. We planned to include parallel and cross-over trials, but not the second phase of cross-over trials because washout periods are not universally used or may not be long enough. Trials that allowed concurrentuse of other medications such as dopamine agonists, anticonvulsants or others were considered suitable for inclusion if they allowed equal access to such medications for patients in the iron and control groups.

Types of participants—We sought trials on adult patients (18 years or over) of either sex, in whom RLS was diagnosed according to expertclinical interview or to explicit diagnostic criteria such as those defined by the International Restless Legs Syndrome Study Group (IRLSSG) (Allen 2003). We planned to include trials on pregnant women or patients with renal disease (including patients with end stage renal disease who are on dialysis) and use these in the investigation of study heterogeneity.

Types of interventions—Therapy with any dose or regimen of oral or parenteral ironcontaining compounds compared with placebo, other drugs or no intervention.

Types of outcome measures

Primary outcomes: The primary outcome considered in this review was restlessness or unpleasant sensations as experienced subjectively by the patient. We did not pre-specify a rating scale or other measure of restlessness, although the majority of studies employed the International Restless Legs Syndrome Study Group severity scale (Walters 2003). This 10 question, 40 point scale addresses the severity and impact of RLS symptoms.

Secondary outcomes: Secondary outcomes included the following:

Efficacy related outcomes:

- 1. quality of life measures;
- 2. patient satisfaction with treatment;
- 3. PLM index (number of periodic limb movements per hour of sleep);

- 4. sleep quality (subjective and objective);
- 5. daytime functioning;
- 6. decreased occurrence of augmentation (according to the definition in Allen 2003).

Safety related outcomes:

- 1. adverse events during treatment;
- 2. discontinuation rate.

Search methods for identification of studies

Electronic searches—We searched the following databases: The Cochrane Movement Disorders Group's Trials Register (Jan 1995 to April 2011); MEDLINE (Jan 1995 to April 2011); EMBASE (Jan 1995 to April 2011); PsycINFO (Jan 1995 to April 2011); CINAHL (Jan 1995 to April 2011); the Cochrane Central Register of Controlled Trials (CENTRAL, *The Cochrane Library*, most recently available issue at the time of the search). The lower time period limit was set to 1995 because generally accepted diagnostic criteria did not exist until that date (see Walters 1995).

Our MEDLINE search used the Cochrane Highly Sensitive Search Strategy for identifying randomized trials (2008 revision) along with MESH headings and free-text searches designed to identify studies of iron administration and of patients with restless legs syndrome with high sensitivity (Appendix 1). The MEDLINE search was adapted as necessary for use in the other databases searched.

Searching other resources—We checked the reference lists of all included studies for other potentially relevant publications in any language. We contacted the corresponding authors of randomized controlled trials and selected members of the International Restless Legs Syndrome Study Group to attempt to find additional published and unpublished trials. We searched the US National Institutes of Health database of ongoing trials (at: clinicaltrials.gov) for ongoing or unpublished trials.

Data collection and analysis

Selection of studies—Two reviewers (LMT and SB) used the titles and abstracts of the identified citations to exclude trials that clearly did not meet the inclusion criteria of the review. If either reviewer thought that the trial might possibly meet the criteria, the full paper was obtained for further examination. The same two reviewers then reviewed articles that passed this initial screen to determine their fit with the review inclusion criteria. Authorship and results were not blinded. All types of disagreement were resolved by discussion and consensus.

Data extraction and management—We developed a data extraction form to aid in the collection of relevant and important data from included studies. As described above in'Criteria for considering studies for this review', we abstracted data on participants, interventions, and outcomes. We also abstracted on to this form the trial characteristics to be used in our assessment of methodological quality. In addition, we collected data on potential

effect modifiers, including co-morbidities such as end stage renal disease, baseline iron status, and pregnancy. Two reviewers (LMT and SB) extracted and cross-checked the data. We contacted study authors to obtain unpublished information, including outcome data not explicitly stated in the published papers. Disagreements concerning data extraction were resolved through the third reviewer (LAB).

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Assessment of risk of bias in included studies—Two reviewers (LMT and SB) assessed the methodological quality of each trial, using the domain-based evaluation outlined in the chapter'Assessing risk of bias in included studies' in the Cochrane Handbook for Systematic Reviews of Interventions version 5 (Higgins 2008). We used the kappa statistic to assess agreement among the reviewers regarding quality of the retrieved articles, resolving disagreements by discussion and consensus among all three reviewers.

Measures of treatment effect—Where appropriate, we combined treatment/control differences in the outcomes across studies using random-effects meta-analyses. We analyzed continuous data using weighted mean differences where possible and performed standardized mean difference analyses when different measurements were used among studies. We calculated risk ratios for dichotomous data using the Mantel-Haen-szel method and 95% confidence intervals. Where possible, we performed intention to treat or per protocol analysis to control attrition bias.

Assessment of heterogeneity—We analyzed study heterogeneity using Cochrane's Q-test and I-square statistics.

We planned to investigate the following as potential sources of heterogeneity when sufficient data were available:

- different formulations of oral or parenteral iron;
- comparator;
- duration of treatment;
- patient characteristics, including pregnancy or end stage renal disease;
- baseline levels of ferritin or other documentation of iron deficiency.

Assessment of reporting biases—We planned to construct a funnel plot to identify possible publication bias.

Sensitivity analysis—We planned to conduct sensitivity analyses for high- versus lowquality trials and for parallel group versus cross-over trial designs. We prespecified subgroup analyses to examine the effects of different types of patients (e.g., pregnant women or patients with ESRD) and different formulations of iron (oral versus parenteral).

Results

Description of studies

See: Characteristics of included studies; Characteristics of excluded studies; Characteristics of ongoing studies.

The initial searches yielded fourteen potentially relevant trials, one of which (Birgegard 2010) was published during the preparation of this review and one (Allen 2009) was initially published only in abstract form. Of these, six studies met criteria for inclusion (Allen 2009, Davis 2000, Earley 2009, Grote 2009, Sloand 2004, Wang 2009), for a total of 192 subjects. Eight trials were excluded because of 1) lack of a non-iron control group (Birgegard 2010, Earley 2004, Earley 2005, O'Keeffe 1994, Simakajornboon 2003, Mohri 2008, Zilberman 2010), and/or 2) subject age younger than 18 (Simakajornboon 2003, Konofal 2008, Mohri 2008), and/or 3) study performed prior to standard RLS criteria definition in 1995 (O'Keeffe 1994). All six included studies were randomized, placebo-controlled, parallel group trials. In the included studies, RLS was defined using the 1995 IRLSSG criteria (Walters 1995) in one study (Sloand 2004), the 2003 IRLSSG criteria (Allen 2003) in three studies (Allen 2009; Grote 2009, Wang 2009), the Johns Hopkins telephone diagnostic interview for RLS in one study (Earley 2009), and by clinical interview by a neurologist in one study (Davis 2000). In this latter study, 24 of 28 subjects met all four diagnostic criteria for RLS using the 1995 criteria. One study (Grote 2009) included only subjects who were iron deficient (defined as a serum ferritin 45) and a second study included only those with low or lownormal iron stores (serum ferritins ranging from 15-75) (Wang 2009). One study (Sloand 2004) included only patients on dialysis (hemodialysis or peritoneal dialysis), while renal disease was an exclusion criteria for four of the remaining studies (Allen 2009, Earley 2009, Grote 2009, Wang 2009). No studies included pregnant women. Two studies administered oral iron (both as ferrous sulfate 325 mg twice a day) (Davis 2000, Wang 2009), while the remaining four used intravenous iron preparations, two of which were iron sucrose (Earley 2009, Grote 2009), one of which was iron dextran (Sloand 2004), and one of which was ferric carboxymaltose (Allen 2009). Our primary endpoint of restlessness or unpleasant sensations was evaluated in five of the included studies, using the International Restless Legs Syndrome Study Group severity scale (IRLS) in four studies (Allen 2009, Earley 2009, Wang 2009, Grote 2009) and a scale developed at the University of Rochester in the fifth (Sloand 2004).

In the first trial (Davis 2000), 28 patients were randomized to receive either placebo or oral ferrous sulfate liquid, 325 mg twice a day for 16 weeks, at which point subjects who wished to continue the study drug as monotherapy were followed for an additional ten weeks. Exclusion criteria included anemia with a hemoglobin less than 10, pregnancy, and current or recent treatment with iron sulfate. Iron status and renal function at baseline were not part of the inclusion or exclusion criteria. RLS symptoms were measured with two visual analog scales, the first which quantified the extent to which RLS symptoms interfered with sleep and the second which quantified how RLS symptoms affected each subject's life. In the second trial (Sloand 2004), 25 subjects with end-stage renal disease on dialysis (22 on hemodialysis, 3 on peritoneal dialysis) were randomized to receive either intravenous iron

dextran or placebo. Both interventions were given as a 15 mL test dose (30 mg of iron dextran) over three minutes followed by 1 hour of observation for adverse reactions; if no adverse reactions were noted, the remaining 485 mL (970 mg of iron dextran) was infused over three hours. RLS symptoms were measured at baseline, 1 week, 2 weeks, and 4 weeks after infusion using the University of Rochester RLS severity scale developed for this study. This three-question scale assesses frequency of RLS symptoms within the last two days, the amount of distress caused by RLS symptoms, and the duration of symptoms to generate a score from 0 to 10 point scale (with higher numbers corresponding to increased RLS severity).

Allen et al (Allen 2009) enrolled 46 subjects, of whom 43 received all doses of study medication. Subjects were randomized to receive either intravenous ferric carboxymaltose or placebo. Iron was delivered as 500 mg of ferric carboxymaltose in 100 mL of normal saline, given on day 0 and day 5. Normal saline was given using the same volume and schedule to subjects in the placebo group. Following day 28, subjects were assigned to receive ferric carboxymaltose or normal saline based on their serum ferritin and IRLS scores at day 28. Because the treatment assignation at day 28 was not randomized (although it remained blinded), we only considered data gathered before the treatment given on day 28. IRLS scores were measured at baseline, day 14, and day 28. Other outcome measures included the RLS Quality of Life scale (RLS-QoL), the Medical Outcomes Study sleep scale (a subjective measure of sleep quality), periodic limb movements (measured via leg actigraphy), the Clinical Global Inventory of Change, a patient global rating of change, and the Fatigue Severity Scale. The full manuscript (Allen 2011) was published during revision of this review, so data analysis was performed on pre-publication data provided by the authors. In the trial by Earley et al (Earley 2009), 21 subjects were randomized to receive either intravenous iron or placebo. A prespecified interim analysis was performed after 18 subjects had been enrolled, at which point the study was discontinued because of lack of effect of the intervention, so three of the randomized subjects did not complete the protocol. Intravenous iron was given as iron sucrose, 500 mL (containing 500 mg iron sucrose) over 10 hours, followed by a second infusion over 12 hours the following day. The same infusion schedule of saline was given to the placebo group. IRLS scores were measured at baseline and two weeks following intervention. Other outcome measures included a global rating scale of RLS severity, periodic limb movements during wake and sleep, changes in cerebrospinal fluid ferritin and transferrin levels, objective sleep quality (total sleep time and sleep efficiency during polysomnography), and changes in iron index within the substantia nigra on MRI.

Grote et al (Grote 2009) randomized60subjects, all of whom were iron-deficient (based on a serum ferritin less than or equal to 45), to receive either intravenous iron or placebo. Iron was delivered as 200 mg of iron sucrose (10 mL of 20 mg/mL Venofer), given five times over a three week period. The placebo group received the same volume and dosing schedule of 0.9% NaCl. The IRLS scale was measured at baseline, week 3, week 7, week 11, 5 months, 8 months, and 12 months. Serum ferritin and the Epworth Sleepiness Scale were both measured as well.

Wang et al (Wang 2009) enrolled 18 subjects with low or low-normal ferritin levels (serum ferritin 15-75) to receive either oral iron (ferrous sulfate 325 mg bid) or matched placebo for twelve weeks. IRLS scores were measure datbase line, six weeks, and twelve weeks after intervention. Serum ferritin was also measured at 12 weeks, as was a single question regarding change in quality of life compared to prior to beginning the study medication (rated as improved, stayed the same, or worsened). Compliance (based on manual pill count) was measured at 6 and 12 weeks.

Risk of bias in included studies

Several of the studies had inadequate reporting, which limited our assessment of potential bias (Figure 1). Sequence generation was not clearly described in four of the studies (Allen 2009, Davis 2000, Earley 2009, Sloand 2004). Two studies were subject to possible unblinding because a study nurse or physician who interacted with subjects (although did not assess outcomes) were aware of allocation status (Allen 2009, Earley 2009). Incomplete outcome data were a potential problem in two studies (Davis 2000, Sloand 2004). In the study by Davis et al, 6 of 14 iron-treated subjects but only one of 14 placebo-treated patients withdrew before analysis at 14 weeks (the time of primary outcome measurement) and no imputation of missing values was performed. Sloand et al reported two drop-outs in the placebo group but reported outcome data for the entire group, without a clear discussion of how outcome data was imputed for the subjects who dropped out. It is difficult to fully rule out the possibility of selective outcome reporting in any of the trials without published protocols (Allen 2009, Davis 2000, Earley 2009, Sloand 2004, Wang 2009). Only one trial (Grote 2009) was registered in a publicly accessible trial registry (www.controlledtrials.com). This registration specified the same primary and secondary outcomes as reported in the published paper, but the date listed as the date on which the trial number was assigned was almost three years after the anticipated end date of the trial. Other biases were possible or probable in four studies. The intervention groups were imbalanced with respect to important potential confounders in one study (Sloand 2004), with mean baseline RLS severity scores being two points higher (out of a possible 10) in the placebo group and with duration of dialysis being longer in the placebo group (3.3 vs 2.5 years). Both of these could potentially indicate more severe disease in the placebo group, which might be more refractory to therapy. Two studies (Allen 2009, Grote 2009) were funded by the pharmaceutical companies selling the iron formulations tested in the studies. One study (Earley 2009) stopped early based on a pre-specified interim analysis of the data, which suggested that continuing the study to the planned enrollment targets was unlikely to show significant differences between the intervention groups. Kappa for agreement between two reviewers about risk of bias was 0.32.

Effects of interventions

The included studies did not all report the same outcome measures. For our primary outcome measure, subjective restlessness or leg discomfort experienced by the patient, data were available from five studies. Four of these studies (Allen 2009, Earley 2009, Grote 2009, Wang 2009)used the IRLS scale and the fifth(Sloand 2004) used a severity scale developed for use in the trial. Of the three studies using the IRLS and IV delivery of iron, Allen reported data for 9 and 23 days following completion of infusions, Earley reported

data for two weeks after infusion, and Grote for three, seven, and eleven weeks after infusion, as well as 5, 8, and 12 months after infusion. Wang reported changes in IRLS at 6 and 12 weeks after beginning twice daily oral iron. For the purpose of comparing changes in IRLS scores across studies, we analyzed changes in IRLS scores at 2 weeks in the Earley study, as this was the only time point available, 3 weeks in the Grote study (as this was the closest time point to that used the Earley study, which was also an intravenous iron study), 3 weeks (23 days) in the Allen study (to be consistent with the Grote study and because the majority of outcome measures were only measured at this time point), and 12 weeks in the Wang study (the duration of the intervention). Combining these four studies, there was no significant difference in change in IRLS scores between iron treatment groups and placebo (mean difference -3.79, 95% CI: -7.68 to 0.10, p = 0.06) (Figure 2). Standardized mean differences including data from the Sloand study could not be computed because the Sloand study reported results as medians rather than means, but this study did find a significant difference in change in RLS severity scores between the iron and placebo groups at 2 weeks (median decrease of 3 points in the iron group and no change in the placebo group, p =0.01).

Changes in quality of life with treatment were assessed in three studies (Allen 2009, Davis 2000, Wang 2009). There was no difference in change in quality of life between the treatment and placebo groups using a dichotomized measure of improved versus unchanged or worsened (RR 2.01, 95% CI: 0.54 to 7.45, $I^2 = 54\%$), but there was an improvement in quality of life measured as a continuous variable (SMD 0.72, 95% CI 0.21 to 1.24, $I^2 = 0\%$); each measurement of quality of life was available from two of the three trials. All studies assessed patient satisfaction with treatment, either directly or indirectly. In three of the studies, satisfaction was measured directly at a certain time point, when subjects were asked if they would prefer to remain in the study or leave the study to start different therapy for RLS (Allen 2009, Davis 2000, Earley 2009). In the remaining three studies, subject drop outs due to lack of efficacy of therapy were reported. Combining these two measures to determine the proportion of subjects who chose to leave the study due to lack of efficacy, there was no difference between iron and placebo groups (RR 0.62, 95% CI 0.30, 1.28, $I^2 =$ 53%). The Allen study was not included in this summary measure, because satisfaction was assessed after day 28 (at which time some subjects received additional, non-randomly allocated iron treatment). Periodic limb movements of sleep were measured in two studies (Allen 2009, Earley 2009), by actigraphy or polysomnography. The decrease after treatment was not significantly different between iron and placebo groups (SMD -0.19, 95% CI -0.70 to 0.32, $I^2 = 0\%$). Subjective sleep quality was assessed in two studies (Allen 2009, Davis 2000), using different measures. Sleep quality was not found to be different after treatment with iron versus placebo (SMD 0.19, 95% CI -0.58 to 0.95, $I^2 = 52\%$). Objective sleep quality, measured as sleep efficiency during polysomnography, was not significantly different between iron and placebo groups (-35.5 +/- 92.0 versus -41.4 +/-98.2) in the single study that measured it (Earley 2009). Daytime functioning was assessed in one study using the Epworth Sleepiness Scale, a subjective measure of propensity to fall asleep during routine activities (Grote 2009). Epworth scores were not given for the treatment groups, but were noted not to be significantly different from placebo at any time point. Changes in the rate of augmentation was not explicitly addressed in any of the included studies.

Adverse events were noted for five of the six studies. In four of the studies, data was reported (Davis 2000, Sloand 2004) or available from the authors (L. Grote, R. Allen) by number of subjects in each group to experience any adverse event. In the other (Earley 2009), adverse events were listed but it was not clear whether adverse events were unique occurences or whether an individual subject may have had more than one adverse event. For the purpose of our meta-analysis, therefore, only data from the first four studies were included. The risk of having any adverse event was no different between iron and placebo groups (RR 1.39, 95% CI 0.85 to 2.27, $I^2 = 51\%$). Gastrointestinal side effects (nausea/ vomiting, constipation, diarrhea, and abdominal pain) were considered as unique events (i.e., counted as separate occurences even if they occurred in the same patient) and were no more common in iron than placebo groups (RR 1.97, 95% CI 0.96 to 4.06, $I^2 = 2\%$). Subjects were no more likely to drop out of the study due to adverse events in the iron than the placebo group (RR 2.44, 95% CI 0.53 to 11.24, $I^2 = 0\%$); no drop outs due to adverse events occurred in three of the studies (Allen 2009; Earley 2009; Wang 2009).

Subgroup analyses by patient type were possible for the subgroups of ESRD and baseline iron deficiency. Only one study explicitly included patients with ESRD (Sloand 2004), and this study did show a significant improvement in RLS severity scores at 2 weeks. One study included only subjects who were iron deficient (as defined by a serum ferritin 45) (Grote 2009), with an additional study limited to those with low and low-normal iron store (as defined by serum ferritin levels between 15 and 75) (Wang 2009). Neither the Grote study alone nor both studies in combination found an improvement in IRLS scores in the treated group compared to placebo for our specified time points, although the Grote study did find a significant improvement at another time point (albeit not their prespecified primary endpoint). Two studies (Allen 2009, Davis 2000) reported that baseline iron stores were not different between subjects who responded to iron therapy and those who did not. Considering type of iron delivery (oral versus intravenous), the one study to use both oral iron and a measure of RLS severity (Wang 2009) found a significant improvement in RLS severity with treatment, while intravenous iron was associated with significant improvements in two of three studies (Allen 2009, Sloand 2004). Sufficient data were not available to conduct sensitivity analyses for low versus high quality trials nor parallel group versus crossover trials.

There was moderate heterogeneity for our primary outcome measure ($I^2 = 52\%$). Heterogenity decreased when only IV studies were considered together, suggesting that method of iron administration explained some of the observed heterogeneity. All studies used a placebo comparator, so this was not a source of heterogeneity. When studies reported multiple time points for outcome measurement, we chose the time point closest to that reported in other studies to minimize heterogeneity from duration of therapy. The IV studies measured outcomes at 2-3 weeks and the oral studies at 12 weeks, so heterogeneity caused by difference in time of outcome assessment cannot be separated from difference in method of administration. Baseline severity of PLMS might have contributed to heterogeneity, as PLMS greater than 15/hour were required for inclusion in Allen 2009 and Earley 2009 but PLMS causing significant sleep disruption were an exclusion criterion in Wang 2009. Considering only those studies in which subjects had low or low-normal iron status at

baseline resulted in greater heterogeneity ($I^2 = 77\%$), suggesting that baseline iron stores did not account for observed heterogeneity.

Discussion

There is insufficient evidence to conclude whether iron is beneficial for the treatment of restless legs syndrome. In our meta-analysis of the four studies that reported mean change in IRLS scores, the mean improvement in IRLS was 3.8 points higher in those subjects on iron than those taking placebo. This estimate is lower than the 6 point difference that has been proposed as a clinically meaningful improvement in IRLS scores (Trenkwalder 2007) and the 5.7 point difference observed with dopamine agonist therapy (Scholz 2011). However, due to the small number of included subjects (n = 139), the confidence interval around this estimate is broad and includes estimates that would be clinically meaningful. We were not able to include results from two of the available trials in this estimate, one of which found a significant improvement in subjective RLS symptoms. Additionally, several trials measured outcomes at multiple time points. For consistency, we used data from 2-3weeksafter IV iron and12 weeks after oral iron. However, the optimal time after iron administration to measure RLS-related outcomes is not known, and in the case of studies that showed significant benefit of iron at only some time points (e.g., the study by Grote et al, Grote 2009), this choice may have obscured a transient effect of iron.

While adverse events were no more common with iron therapy, there may be still be differences in tolerability between oral and intravenous iron for the treatment of RLS. Of the studies that could be included in our analysis of any adverse event, three used intravenous iron and one oral iron. The only one of these studies to show a significantly higher risk of adverse events in the treated group was the oral iron study. Tolerability may be influenced by factors other than route of administration (e.g., formulation of iron).

A lack of power in the primary studies may also have limited our ability to draw conclusions about the efficacy of iron therapy for RLS. Power analyses were only mentioned in four of the included studies (Allen 2009, Davis 2000, Grote 2009, Wang 2009). Of these, the studies by Allen and Grote had sufficient power (80-90%) to detect a difference in IRLS scores. However, the study by Davis et al was clearly underpowered, with only a 25% power to detect a difference in their primary outcome (improvement in average quality of sleep). The study by Earley et al may have been underpowered as well, as 36 subjects were originally planned to provide 80% power, but the study was stopped after only 18 subjects had completed the protocol due to apparent lack of effect. Furthermore, although randomized, the two treatment arms were unbalanced at baseline in the Sloand group (Sloand 2004), with more severely affected patients in the placebo group, which may have decreased the power to see an effect in this study.

The role of baseline iron deficiency in the response to iron therapy cannot be fully determined from the available evidence. Stratified response data by baseline iron status were not available for the majority of studies, although two studies found that responders and non-responders did not differ by baseline iron status. Additionally, several of the studies had exclusion criteria that would serve to exclude the most iron deficient subjects from

participating. The study by Wang et al excluded those with a percent saturation < 15% or anemia (defined as hemoglobin less than 11.1 in women and 14 in men), the study Earley et al excluded those with hemoglobin less than 12, and the study by Davis et al excluded those with hemoglobin less than 10. Given that iron deficiency is one of the most common causes of anemia, excluding anemic patients would serve to exclude a number of severely iron deficient patients.

Authors' Conclusions

Implications for practice

There is insufficient evidence to determine whether iron is effective therapy for restless legs syndrome or whether particular subsets of patients (e.g., those with end-stage renal disease on dialysis, those who are iron deficient) are more likely to benefit. Both oral and intravenous formulations of iron were reasonably well-tolerated in the included studies.

Implications for research

Larger, high quality studies are needed to determine whether iron therapy is an effective therapy for RLS. Future studies should also be designed to identify specific patient characteristics (e.g., baseline iron stores, renal function) that may be related to response to treatment. If iron ultimately proves to be beneficial in RLS patients, the question of which formulation of iron is most effective and best tolerated will also need to be answered.

Acknowledgments

We wish to thank the investigators of the studies included in this review, especially Richard Allen, Bradley Davis, and Ludger Grote, who provided us with unpublished data or manuscript clarifications. We also wish to thank Jan Ulfberg and his coauthors for providing us access with their then-unpublished manuscript comparing oral and intravenous iron in blood donors.

Sources of Support

Internal sources

• No sources of support supplied

External sources

- Jazz Pharmaceuticals Fellowship Training Grant in Sleep Medicine, USA. Partial funding for LM Trotti
- PHS Grant KL2 RR025009 from the Clinical and Translational Science Award program, National Institutes of Health, National Center for Research Resources, USA. Partial funding for LM Trotti

Characteristics of Studies

Characteristics of included studies [ordered by study ID]

Allen 2009

Methods	Randomized, controlled trial
Participants	46 adult men and women with RLS, enrolled from multiple sites
Interventions	Intervention: ferric carboxymaltose 500 mg in 100 mL normal saline on days 0 and 5 Placebo: 100 mL normal saline with same dosing schedule

Outcomes	IRLS, RLS-QoL, CGI-1, PGI-1, MOS sleep scale, FSS, PLMS measured by actigraphy	
Notes		
Risk of bias	-	
Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	study randomized but method of sequence generation not specified
Allocation concealment (selection bias)	Low risk	
Blinding (performance bias and detection bias) All outcomes	Unclear risk	Study nurse not blinded and interacted with subjects (not an outcome assessor)
Incomplete outcome data (attrition bias) All outcomes	Low risk	
Selective reporting (reporting bias)	Unclear risk	no study protocol available
Other bias	High risk	funded by pharmaceutical company

Davis 2000

Methods	Randomized, controlled trial
Participants	28 adult men and women with symptomatic RLS currently on treatment, enrolled from a single neurology clinic (USA)
Interventions	Intervention: Ferrous sulfate liquid 325 mg bid. Placebo: a 50:50 mixture of water and 2% carboxymethlycellulose with a "slightly disagreeable taste, similar to liquid ferrous sulfate" Study drug continued for 16 weeks, then extended an additional 10 weeks if patients preferred study drug to their previous RLS medication
Outcomes	Visual analog scale (VAS) of the extent that RLS interferes with sleep; VAS of how RLS affects life, ferritin, GI side effects
Notes	Power for primary outcome was 25%
Risk of bias	

Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	block randomization but method not specified
Allocation concealment (selection bias)	Low risk	
Blinding (performance bias and detection bias) All outcomes	Low risk	
Incomplete outcome data (attrition bias) All outcomes	High risk	6 of 14 treated pts left before wk 14, which was primary endpoint; this data only reported for the people who remained
Selective reporting (reporting bias)	Unclear risk	no study protocol available
Other bias	Low risk	

Earley 2009

Methods	Randomized, controlled trial	
Participants	18 adult men and women with RLS and PLMS >= 15/hr, enrolled from a single medical center (USA)	
Interventions	Intervention: Iron sucrose, given as a 500 mL infusion (containing 500 mg iron sucrose) over 10 hours, then another 500 mL over 12 hours the following day. Placebo: NaCl, given in the same infusion volume & with the same dosing schedule	
Outcomes	IRLS (a secondary outcome), PLMS/hr by 2nd night PSG, ferritin, global rating scale (7 pt scale of symptom severity completed by the patient), sleep diary, PLMW on SIT, change in CSF ferritin, change in MRI iron index in substantia nigra	
Notes		
Risk of bias		
Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	randomized 2:1 but doesn't specify method of sequence generation
Allocation concealment (selection bias)	Low risk	
Blinding (performance bias and detection bias) All outcomes	Unclear risk	Nurse & MD knew treatment status; they were not the outcome assessors but did interact with some subjects
Incomplete outcome data (attrition bias) All outcomes	Low risk	
Selective reporting (reporting bias)	Unclear risk	no study protocol available
Other bias	High risk	Stopped early due to data-dependent process (i.e., a stopping rule)

Grote 2009

Methods	Randomized, controlled trial	
Participants	60 adult men and women meeting all 4 RLS diagnostic criteria who had IRLS scores >= 10 and s-ferritin < 45, recruited from three separate hospitals (Sweden)	
Interventions	Intervention: 200 mg iron sucrose IV (10mL of 20mg/mL iron (III) as iron (III)-hy-droxide sucrose complex) given 5 times over 3 weeks Placebo: NaCl given with same timing as intervention	
Outcomes	IRLS (BL, wk 3, wk 7, wk 11, mo 5, mo 8, mo 12, with difference at wk 11 set as primary outcome), ferritin, ESS	
Notes		
Risk of bias		
Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Low risk	
Allocation concealment (selection bias)	Low risk	

Blinding (performance bias and detection bias) All outcomes	Low risk	
Incomplete outcome data (attrition bias) All outcomes	Low risk	Intention to treat analysis with last observation carried forward
Selective reporting (reporting bias)	Unclear risk	Study protocol registered but registration date appears to be after study completion
Other bias	High risk	Funded by pharmaceutical company

Sloand 2004

Methods	randomized, placebo-controlled trial	
Participants	25 adult men and women with ESRD (on either hemodialysis or peritoneal dialysis) who also had RLS by IRLSSG criteria, recruited from the dialysis units of a single medical center (USA)	
Interventions	Active: IV iron dextran given as: 30 mg test dose over 3 min, followed by 970 mg (in 485 mL volume) over 3 hrs if no reaction to the test dose Placebo: NaCl given in the same volume & and with the same timing as the intervention group	
Outcomes	ferritin, Hct, GI side effects, U of R severity score (10 pt scale where 0=none, 10=bad) at 0, 1, 2, 4 wks	
Notes		
Risk of bias		
Bias	Authors' judgement	Support for judgement
Random sequence generation (selection bias)	Unclear risk	method of randomization not specified
Allocation concealment (selection bias)	Low risk	
Blinding (performance bias and detection bias) All outcomes	Low risk	
Incomplete outcome data (attrition bias) All outcomes	Unclear risk	two placebo patients dropped out after infusion, but Figure 1 reports data on all 14 placebo patients
Selective reporting (reporting bias)	Unclear risk	no study protocol available
Other bias	Unclear risk	baseline imbalance of 2 points in RLS severity score; there was also a broader range of scores in the treated group; imbalance in duration of dialysis

Wang 2009

Methods	Randomized, controlled
Participants	18 adult men and women with RLS who had ferritin 15-75 and IRLS scores ≥ 11
Interventions	Intervention: ferrous sulfate 325 mg bid in non-descript capsules Placebo: lactose, appearanced matched Study drug continued for 12 weeks

	Outcomes	IRLS (BL, 6 wks, 12 wks), ferritin (BL, 12 wks), single question on change in QOL at 12 wks (improved, same, worsened), compliance at 6 & 12 wks (by manual pill count)	
	Notes	authors raise issue of iron causing stool discoloration and therefore inadvertent unblind-ing, although this is speculative	
	Risk of bias		
Bias Authors' judgement Support for judgement		Support for judgement	
	Random sequence generation (selection bias)	Low risk	
	Allocation concealment (selection bias)	Low risk	
	Blinding (performance bias and detection bias) All outcomes	Low risk	
-	Incomplete outcome data (attrition bias) All outcomes	Low risk	
	Selective reporting (reporting bias)	Unclear risk	no study protocol available
	Other bias	Low risk	

Characteristics of excluded studies [ordered by study ID]

Study	Reason for exclusion
Birgegard 2010	No control group
Earley 2004	No control group
Earley 2005	No control group
Konofal 2008	Included children
Mohri 2008	No control group; included children
O'Keeffe 1994	Performed prior to 1995 (non-standard diagnostic criteria for RLS used); no control group
Simakajornboon 2003	No control group; included children
Zilberman 2010	No control group

Characteristics of ongoing studies [ordered by study ID]

Earley NCT00685815

Trial name or title	Intravenous Iron Metabolism in Restless Legs Syndrome	
Methods Double-blind, randomized controlled trial		
Participants 36 adults with RLS		
Interventions 500 mg of IV ferric carboxymaltose given on two consecutive		
Outcomes RLS symptoms		
Starting date	November 2006	
Contact information Christopher J. Earley, MD, PhD, Johns Hopkins Universit		

Notes

registered at www.clinicaltrials.gov

Data And Analyses

Comparison 1 Change in IRLS scores

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Change in IRLS severity scale score	4	139	Mean Difference (IV, Random, 95% CI)	-3.79 [-7.68, 0.10]

Comparison 2 Subgroup analyses of IRLS scores

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Subgroup IV iron	3	121	Mean Difference (IV, Random, 95% CI)	-2.45 [-5.59, 0.68]
2 Subgroup iron deficient	2	78	Mean Difference (IV, Random, 95% CI)	-4.93 [-12.50, 2.65]

Comparison 3 Quality of Life

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Quality of Life (dichotomous measure)	2	39	Risk Ratio (M-H, Random, 95% CI)	2.01 [0.54, 7.45]
2 Quality of Life (continuous measure)	2	64	Std. Mean Difference (IV, Random, 95% CI)	0.72 [0.21, 1.24]

Comparison 4 Patient satisfaction with treatment

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Withdrew to use other RLS meds	5	142	Risk Ratio (M-H, Random, 95% CI)	0.62 [0.30, 1.28]

Comparison 5 Change in PLMS

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Change in PLMS	2	60	Std. Mean Difference (IV, Random, 95% CI)	-0.19 [-0.70, 0.32]

Comparison 6 Subjective sleep quality

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Change in Subjective Sleep quality	2	64	Std. Mean Difference (IV, Random, 95% CI)	0.19 [-0.58, 0.95]

Comparison 7 Number with adverse events

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Adverse events	4	156	Risk Ratio (M-H, Random, 95% CI)	1.39 [0.85, 2.27]

Comparison 8 Gastrointestinal side effects

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Gastrointestinal side effects	5	174	Risk Ratio (M-H, Random, 95% CI)	1.97 [0.96, 4.06]

Comparison 9 Drop out due to adverse event

Outcome or subgroup title	No. of studies	No. of participants	Statistical method	Effect size
1 Drop out due to adverse event	3	113	Risk Ratio (M-H, Random, 95% CI)	2.44 [0.53, 11.24]

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Analysis 1.1 Comparison 1 Change in IRLS scores, Outcome 1 Change in IRLS severity scale score

Review: Iron for restless legs syndrome

Comparison: I Change in IRLS scores

Outcome: I Change in IRLS severity scale score

Study or subgroup	Experimental		Control		Mean Difference	Weight	Mean Difference
	Z	Mean(SD)	Z	Mean(SD)	IV,Random,95% CI		IV,Random,95% CI
Allen 2009	24	-8.9 (8.52)	61	-4 (6.11)	•	30.7 %	-4.90 [-9.28, -0.52]
Earley 2009	Ξ	-10.1 (5.1)	7	-12 (11.5)	•	13.5 %	1.90 [-7.14, 10.94]
Grote 2009	29	-8.7 (7.9)	31	-7.3 (7.7)	•	33.1 %	-1.40 [-5.35, 2.55]
Wang 2009	Ξ	-10.3 (7.4)	7	-1.14 (5.64)		22.7 %	-9.16 [-15.21, -3.11]
Total (95% CI) Heterogeneity: Tau ² = Test for overall effect:	75 = 7.83; Chi ² = 6.20 Z = 1.91 (P = 0.0 ¹	, df = 3 (P = 0.10); 56)	64 ² =52%			100.0 %	-3.79 [-7.68, 0.10]
lest tor subgroup unit	srences: INUL applic	able					
					-		
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				Favours e	xperimental Favours contr	0	

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Analysis 2.1 Comparison 2 Subgroup analyses of IRLS scores, Outcome 1 Subgroup -- IV iron

Review: Iron for restless legs syndrome

Comparison: 2 Subgroup analyses of IRLS scores

Outcome: I Subgroup – IV iron

Study or subgroup	Experimental		Control		Mean Difference	Weight	Mean Difference
	Z	Mean(SD)	Z	Mean(SD)	IN,Random,95% C	-	IV,Random,95% CI
Allen 2009	24	-8.9 (8.52)	61	-4 (6.11)		40.8 %	-4.90 [-9.28, -0.52]
Earley 2009	Ξ	-10.1 (5.1)	7	-12 (11.5)	-	→ II.3 %	1.90 [-7.14, 10.94]
Grote 2009	29	-8.7 (7.9)	31	-7.3 (7.7)		47.9 %	-1.40 [-5.35, 2.55]
Total (95% CI)	64		57		¢	100.0 %	-2.45 [-5.59, 0.68]
Heterogeneity: Tau ² =	1.27; Chi ² = 2.36	h = 2 (P = 0.31);	$ ^2 = 5\% $				
Test for overall effect:	Z = 1.54 (P = 0.1)	2)					
Test for subgroup diffe	rences: Not applic	cable					
				-		-	
				01-) -5 0 5	01	
				Favours e	xperimental Favours	s control	

Analysis 2.2 Comparison 2 Subgroup analyses of IRLS scores, Outcome 2 Subgroup -- iron deficient

Review: Iron for restless legs syndrome

Comparison: 2 Subgroup analyses of IRLS scores

Outcome: 2 Subgroup – iron deficient

Study or subgroup	Experimental		Control		Mean Difference	Weight	Mean Difference	
	Z	Mean(SD)	Z	Mean(SD)	IV,Random,95% CI		IV,Random,95% CI	
Grote 2009	29	-8.7 (7.9)	31	-7.3 (7.7)		54.5 %	-1.40 [-5.35, 2.55]	
Wang 2009	Ξ	-10.3 (7.4)	7	-1.14 (5.64)		45.5 %	-9.16 [-15.21, -3.11]	
Total (95% CI)	40		38			100.0 %	-4.93 [-12.50, 2.65]	
Heterogeneity: Tau ² =	= 23.32; Chi ² = 4.4	H3, df = 1 (P = 0.04	l); l ² =77%					
Test for overall effect:	Z = 1.28 (P = 0.20)	(0						
Test for subgroup diffe	erences: Not applic	cable						
					-	_		
				ī	0 -5 0 5 1	0		
				Favorine	exnerimental Eavours con	trol		

Review: Iron for restless legs syndrome

Comparison: 3 Quality of Life

y or subgroup Experimental Control M- n/N n/N H.Random.95% H.Random.95% H.Random.95% H.Random.95% $ -$	ly or subgroup Experimental Control $M_{\rm eff}$ (Risk Ratio $M_{\rm eff}$ (Risk	come: I Quality of	f Life (dichotomous measu	ure)			
n/N n/N <th>n/N n/N <t< th=""><th>or subgroup</th><th>Experimental</th><th>Control</th><th>Risk Ratio</th><th>Weight</th><th>Risk Ratio</th></t<></th>	n/N n/N <t< th=""><th>or subgroup</th><th>Experimental</th><th>Control</th><th>Risk Ratio</th><th>Weight</th><th>Risk Ratio</th></t<>	or subgroup	Experimental	Control	Risk Ratio	Weight	Risk Ratio
vis 2000 $7/8$ $8/13$ 698% $[.42 [0.86, 2.35]$ ang 2009 $7/1$ $1/7$ $1/7$ $1/7$ $1/2$ $[0.69, 2.87]$ and 2009 $7/1$ $1/7$ $1/7$ $1/7$ $1/7$ $1/2$ $[0.54, 7.45]$ 1 (95% CI) $1 9 (Control)events: 14 (Experimental), 9 (Control)ogeneticy: Tau2 = 0.57; Chi2 = 2.18, df = 1 (P = 0.14), l2 = 54\%or overall effect. Z = 1.04 (P = 0.30)or overall effect. Z = 1.04 (P = 0.30)or subgroup differences: Not applicable$	vis 2000 $7/8$ $8/13$ 69.8% $1.42 [0.86, 2.35]$ ang 2009 $7/11$ $1/7$ $1/7$ 30.2% $4.45 [0.69, 2.887]$ 1 (95% CI) 19 20 30.2% $2.01 [0.54, 7.45]$ Vertes: 14 (Experimental), 9 (Control) ogeneity: Tau ² = 0.57; Chi ² = 2.18, df = 1 (P = 0.14); l ² = 54\% or overall effect. Z = 1.04 (P = 0.30) or subgroup differences: Not applicable 0.05 0.2 1 5 20 Favours control favours experimental		N/u	N/n	H,Random,95% Cl		H,Kandom,95% Cl
ang 2009 $7/11$ $1/7$ 30.2 % 4.45 [0.69, 28.87]and 2009 $7/11$ $1/7$ 30.2 % 4.45 [0.69, 28.87]averts: 14 (Experimental), 9 (Control) 200 100.0 % 2.01 [0.54 , 7.45]ogeneity: Tau ² = 0.57; Chi ² = 2.18, df = 1 (P = 0.14); l ² = 54% 100.0 % 2.01 [0.54 , 7.45]or overall effect: Z = 1.04 (P = 0.30) 100.0 % 2.01 [0.54 , 7.45]or subgroup differences: Not applicable 100.0 % 2.01 [0.54 , 7.45]	ang 2009 $7/11$ $1/7$ $1/7$ 30.2% $4.45 [0.69, 28.87]$ d (95% CI) 19 20 100.0 % 2.01 [0.54, 7.45] events: 14 (Experimental), 9 (Control) ogeneity: Tau ² = 0.57; Chi ² = 2.18, df = 1 ($P = 0.14$); $l^2 = 54\%$ or overall effect: Z = 1.04 ($P = 0.30$) or subgroup differences: Not applicable 0.05 0.2 1 5 20 Favours control Favours experimental	avis 2000	7/8	8/13		69.8 %	1.42 [0.86, 2.35]
I (95% CI)1920100.0 %2.01 [0.54, 7.45]events: 14 (Experimental), 9 (Control) ogeneity: Tau ² = 0.57; Chi ² = 2.18, df = 1 ($P = 0.14$); l ² = 54% or overall effect: Z = 1.04 ($P = 0.30$) or subgroup differences: Not applicable2.01 [0.54, 7.45]	I (95% CI) 19 20 100.0 % 2.01 [0.54, 7.45] events: 14 (Experimental), 9 (Control) ogeneity: Tau ² = 0.57; Chi ² = 2.18, df = 1 (P = 0.14); l ² = 54% 000.0 % 2.01 [0.54, 7.45] or evenal effect. Z = 1.04 (P = 0.30) or evenal effect. Z = 1.04 (P = 0.30) 0.05 0.2 1 or subgroup differences: Not applicable 0.05 0.2 1 5 20 Favours control Favours control Favours experimental	ang 2009	7/11	1/7	•	30.2 %	4.45 [0.69, 28.87]
events: 14 (Experimental), 9 (Control) ogeneity: Tau ² = 0.57; Chi ² = 2.18, df = 1 (P = 0.14); l ² =54% or overall effect: Z = 1.04 (P = 0.30) or subgroup differences: Not applicable	events: 14 (Experimental), 9 (Control) ogeneity: Tau ² = 0.57; Chi ² = 2.18, df = 1 ($P = 0.14$); $l^2 = 54\%$ or overall effect: $Z = 1.04$ ($P = 0.30$) or subgroup differences: Not applicable 0.05 0.2 1 5 20 Favours control Favours control Favours experimental	l (95% CI)	19	20	•	100.0 %	2.01 [0.54, 7.45]
	0.05 0.2 I 5 20 Favours control Favours experimental	events: 14 (Experim ogeneity: Tau ² = 0.1 or overall effect: Z = or subgroup differen	tental), 9 (Control) 57; Chi ² = 2.18, df = 1 (P = 1.04 (P = 0.30) rces: Not applicable	o = 0.14); ² =54%			

Review: Iron for restless legs syndrome

mparison: 3 Quali tcome: 2 Quality	ty of Life of Life (continuo	us measure)			2	Std.	
subgroup	Experimental		Control		Differ	ence	Weight
	Z	Mean(SD)	z	Mean(SD)	IV,Random	1,95% CI	
2009	24	56.5 (49.1)	61	19.5 (51.7)		•	68.2 %
2000	8	13.45 (13.52)	13	1.69 (16.85)	-		31.8 %
95% CI)	32		32		•	♦	100.0 %
:neity: Tau ² = (verall effect: Z	0.0; Chi ² = 0.00, = 2.75 (P = 0.00	df = 1 (P = 0.99); 360)	² =0.0%				
ubgroup diffen	ences: Not applic	able					

Std. Mean Difference

IV,Random,95% CI

0.72 [-0.19, 1.63] 0.72 [0.10, 1.35]

0.72 [0.21, 1.24]

Favours experimental

Favours control -7

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Comparison 4 Patient satisfaction with treatment, Outcome I Withdrew to use other RLS meds Analysis 4.1

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	Review: Iron for restless legs syndrome

Outcome: I Withdrew to use other RLS meds

	0.1 0.2 0.5 1 2 5 10			
	-			
			(P = 0.19)	Test for overall effect: $Z = 1.30$
		. =53%	$n^2 = 6.41$, df = 3 (P = 0.09); l^2	Heterogeneity: Tau ² = 0.27; C ¹
			, 34 (Control)	Total events: 16 (Experimental)
$0.62 \ [\ 0.30, \ 1.28 \]$	¢	72	70	Total (95% CI)
0.0 [0.0, 0.0]		0/7	11/0	Wang 2009
0.42 [0.02, 9.34]		1/14	11/0	Sloand 2004
0.28 [0.12, 0.65]	ł	19/31	5/29	Grote 2009
1.27 [0.46, 3.50]	•	3/7	6/11	Earley 2009
0.74 [0.41, 1.33]	•	11/13	5/8	Davis 2000
H,Random,95% CI	H,Random,95% CI	N/n	N/n	
Risk Ratio M	Risk Ratio M	Control	Experimental	Study or subgroup

Analysis 5.1

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Comparison 5 Change in PLMS, Outcome 1 Change in PLMS

Review: Iron for restless legs syndrome

Comparison: 5 Change in PLMS

Outcome: I Change in PLMS

Std. Mean Difference	IV,Random,95% CI	-0.20 [-0.81, 0.41]	-0.17 [-1.12, 0.78]	-0.19 [-0.70, 0.32]						
Weight		70.8 %	29.2 %	100.0 %					2	trol
Std. Mean Difference	IV,Random,95% CI		•	¢				-	- 0	perimental Favours con
	Mean(SD)	-4.6 (17.1)	-11.7 (32.9)					_	-2	Favours es
Control	Z	61	7	26	:0%					
	Mean(SD)	-7.6 (13)	-24.8 (89.4)		$(P = 0.96); I^2 = C$					
Favours experimental	Z	23	Ξ	34	= 0.0; $Chi^2 = 0.00$, $df = 1$	Z = 0.72 (P = 0.47)	erences: Not applicable			
Study or subgroup		Allen 2009	Earley 2009	Total (95% CI)	Heterogeneity: Tau ² =	Test for overall effect:	Test for subgroup diffe			

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Analysis 6.1

Outcome: I Change in Subjective Sleep quality

Review: Iron for restless legs syndrome Comparison: 6 Subjective sleep quality

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Std. Std. Std. Mean Weight Difference	1,95% CI IV,Random,95% CI	5 8.4 % 0.52 [-0.10, 1.13]		100.0 % 0.19 [-0.58, 0.95]				-	- 2
Differ	IV,Random		•	ł				-	-2 -1 0
	Mean(SD)	35.1 (75.2)	-3.68 (16.9)						
Control	Z	61	13	32); I ² =52%				
	Mean(SD)	75.8 (79)	-8.23 (13.49)		df = 1 (P = 0.15		able		
Experimental	Z	24	æ	32	0.16; Chi ² = 2.09,	z = 0.48 (P = 0.63)	ences: Not applic		
Study or subgroup		Allen 2009	Davis 2000	Total (95% CI)	Heterogeneity: Tau ² = 1	Test for overall effect: Z	Test for subgroup differ		

Review: Iron for restless legs syndrome

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udy or subgroup Experimental Control Risk n/N n/N H,Rando Allen 2009 10,24 9,19		
HR. 2009 H.Randon Alen 2009 10.0.4 9/19	 Ratio Weight M- 	t Risk Ratio M-
Allen 2009 10/24 9/19	.m.95% Cl	H,Random,95% CI
		5 0.88 [0.45, 1.72]
Grote 2009 18/29 14/31	35.4 %	5 [0.85, 2.22]
Sloand 2004 10/11 8/14	34.9 %	5 [0.97, 2.60]
Davis 2000 9/14 0/14 –	30%	5 [1.21, 297.89]
otal (95% CI) 78 78	100.0 %	0.85, 2.27
tal events: 47 (Experimental), 31 (Control) terogeneity: Tau ² = 0.12; Chi ² = 6.16, df = 3 (P = 0.10); l ² =5.1%		
st for overall effect: $Z = 1.31$ (P = 0.19)		
st for subgroup differences: Not applicable		

Analysis 8.1 Comparison 8 Gastrointestinal side effects, Outcome 1 Gastrointestinal side effects

Review: Iron for restless legs syndrome

Comparison: 8 Gastroir	itestinal side effects					
Outcome: I Gastrointe:	stinal side effects					
Study or subgroup	Experimental	Control	Risk H,Randc	c Ratio M- m,95%	Weight	Risk Ratio M. H,Random,95%
Davis 2000	5/14	0/14			6.6 %	11.00 [0.67, 181.83]
Earley 2009	5/11	0/7		Ì	6.8 %	7.33 [0.47, 115.10]
Grote 2009	2/29	2/31			14.4 %	1.07 [0.16, 7.10]
Sloand 2004	2/11	3/14			% 6:61	0.85 [0.17, 4.23]
Allen 2009	11/24	4/19			52.3 %	2.18 [0.82, 5.76]
Total (95% CI)	89	85		•	100.0 %	1.97 [0.96, 4.06]
Total events: 25 (Experime Heterogeneity: Tau ² = 0.0 Test for overall effect: $Z =$	intal), 9 (Control) 1; Chi ² = 4.07 , df = 4 (f 1.84 (P = 0.066)	⁵ = 0.40); l ² =2%				
iest ior subgroup difference	es: Ivot applicable		-	-		
			0.1 0.2 0.5 1	2 5 10		
		Few	er in experimental	ewer in control		

Analysis 9.1 Comparison 9 Drop out due to adverse event, Outcome 1 Drop out due to adverse event

1			
-			
-			
-			
4			
-			

Review: Iron for restless legs syndrome

Comparison: 9 Drop ou	t due to adverse event				
Outcome: I Drop out d	ue to adverse event				
Study or subgroup	Experimental	Control	Risk Ratio M- H Random 95%	Weight	Risk Ratio M- H Random 95%
	N/N	N/n			
Davis 2000	3/14	0/14	•	28.1 %	7.00 [0.39, 124.14]
Grote 2009	3/29	1/31	•	47.8 %	3.21 [0.35, 29.11]
Sloand 2004	11/0	1/14	•	24.1 %	0.42 [0.02, 9.34]
Total (95% CI)	54	59	\	100.0 %	2.44 [0.53, 11.24]
Total events: 6 (Experimen Heterogeneity: $Tau^2 = 0.0$;	tal), 2 (Control) Chi ² = 1.82, df = 2 (P :	= 0.40); ² =0.0%			
Test for overall effect: Z =	1.15 (P = 0.25)				
Test for subgroup differenc	es: Not applicable				
		0	01 0.1 1 10 100		
		Favours	experimental Favours contr	0	

Appendix 1. Search Strategy

Our MEDLINE search used the following strategy:

- 1. RANDOMIZED CONTROLLED TRIAL [pt]
- 2. CONTROLLED CLINICAL TRIAL [pt]
- 3. RANDOMIZED [tiab]
- 4. PLACEBO [tiab]
- 5. DRUG THERAPY [sh]
- 6. RANDOMLY [tiab]
- 7. TRIAL [tiab]
- 8. GROUPS [tiab]
- **9.** #1 or #2 or #3 or #4 or #5 or #6 or #7 or #8
- 10. animals [mh] not (humans [mh] and animals [mh])
- **11.** #9 not #10
- 12. IRON [mh]
- **13.** IRON COMPOUNDS
- 14. IRON-DEXTRAN COMPLEX
- 15. FERROUS COMPOUNDS
- 16. IRON, DIETARY
- **17.** FERROUS SULFATE
- **18.** FERROUS GLUCONATE
- **19.** FERROUS FUMARATE
- 20. IRON SUCROSE
- **21.** FERRIC GLUCONATE
- 22. IRON DEXTRAN
- **23.** #12 or #13 or #14 or #15 or #16 or #17 or #18 or #19 or #20 or #21 or #22
- 24. RESTLESS LEGS SYNDROME [mh]
- 25. NOCTURNAL MYOCLONUS SYNDROME [mh]
- 26. PERIODIC LIMB MOVEMENT*
- **27.** PERIODIC LEG MOVEMENT*
- 28. EKBOM* and SYNDROME
- 29. PARASITOSIS

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- **30.** #28 not #29
- **31.** RESTLESS LEG*
- **32.** NOCTURNAL MYOCLONUS
- **33.** PSYCHOMOTOR AGITATION
- **34.** AKATH*
- **35.** #24 or #25 or #26 or #27 or #30 or #31 or #32 or #33 or #34
- **36.** #11 and #23 and #35

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Plain Language Summary

Iron for restless legs syndrome

Restless legs syndrome (RLS) is a common medical condition that results in uncomfortable urges to move the legs, especially in the evening and at night, and often interferes with sleep. Low blood levels of iron are frequently seen in people who have RLS and the lack of iron may be part of the cause of RLS. Iron can be supplemented either in pill form or through injections into the bloodstream. This review was performed to see if iron supplements are effective in reducing the symptoms of RLS. Six studies of iron were included, which together involved only 192 subjects. Results from the studies were conflicting, with some studies showing that iron was not effective but others showing some help for patients' feelings of restlessness or discomfort. Because of the different ways in which the studies were done, we could not combine results from all of the studies to come up with an overall judgement of whether or not iron is effective. Two of the studies were limited to specific sub-groups of RLS patients, who might be expected to respond to iron differently than would the RLS group as a whole. The study of RLS patients with severe kidney disease showed a benefit of iron therapy. The study of RLS patients with low blood levels of iron did not consistently show a benefit of iron therapy at all time points. Iron did not cause any more side effects than the placebo medication. More studies are needed before we will be able to determine whether iron therapy should be used for patients with RLS.

	Random sequence generation (selection bias)	Allocation concealment (selection bias)	Blinding (performance bias and detection bias)	Incomplete outcome data (attrition bias)	Selective reporting (reporting bias)	Other bias	
Allen 2009	?	÷	?	•	?	•	
Davis 2000	?	÷	•	•	?	•	
Earley 2009	?	•	?	•	?	•	
Grote 2009	•	•	•	•	?	•	
Sloand 2004	?	÷	•	?	?	?	
Wang 2009	•	+	•	•	?	•	

Figure 1. Risk of bias summary: review authors' judgements about each risk of bias item for each included study

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	Expe	erimen	tal	C	ontrol		Mean Difference		Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% Cl	IV, Random, 95% Cl
Allen 2009	-8.9	8.52	24	-4	6.11	19	30.7%	-4.90 [-9.28, -0.52]	
Earley 2009	-10.1	5.1	11	-12	11.5	7	13.5%	1.90 [-7.14, 10.94]	
Grote 2009	-8.7	7.9	29	-7.3	7.7	31	33.1%	-1.40 [-5.35, 2.55]	
Wang 2009	-10.3	7.4	11	-1.14	5.64	7	22.7%	-9.16 [-15.21, -3.11]	<
Total (95% CI)			75			64	100.0%	-3.79 [-7.68, 0.10]	-
Heterogeneity: Tau² = 7.83; Chi² = 6.20, df = 3 (P = 0.10); l² = 52% -10 -10 -5 0 5 10 Test for overall effect: Z = 1.91 (P = 0.06) Favours experimental Favours control									

Figure 2. Change in IRLS severity scale score