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Higher Expression of Topoiscinerase II Alpha Is an Independent Marker of Vicroased Risk of Cancer-specific Death in Patients with Clear Cell Renal Cell Carcinoma

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Abstract

Background—Turnor-based biomarkers of or the for predients with clear cell renal cell carcinoma (ccRCC) remain limited, especially for those with low-risk disease. Type IIa

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topoisomerase (TOPOIIa) is a well known biomarker of DNA replication and a target for antineoglastic agents, but it has not been evaluated a biomarker of ccRCC outcome.

Object ve— To evaluate the association of TOPOIIa expression in ccRCC and risk of cancerspecific ceath following surgery.

Design, satting, s. d participar.cs—Two independent cohort studies were studied in tertiary referral urology practices in the United States. We identified cohorts of 1378 (analytic) and 279 (validation) patients who underweit the phrectomy for clinically localized ccRCC and had paraffin tumor assume available. TOPOIIa expression was assessed using immunohistochemistry and scored as the number of positive cells per square inillimeter.

Outcome measurements and statistical anziysis—Our primary end point was cancerspecific survival (CSS). We evaluated TCT Offa expression as a continuous variable and diche tomized as low versus high. For associations with CSS, we use I Kaplan-Meier curves and Cox regression n odels

Results and limitations—In both cohorts, patients who had high TOPOIIa expression were approximately threatings more likely to experience ccPCC death than those with low expression (hazard ratio [VR]: 2.75; 55% confidence interval [C1], 2.12–3.56; p = 1.79E-14 and HR: 3.45; 95% CI, 1.34–9.68; p = 0.2104, respectively). Multivariable adjustment for pathologic features of aggressiveness did not explain these associations, and stratified analysis suggests that the association is more pronounced among patients with low-risk disease is defined by the Mayo Clinic stage, size, grade, and necroils score

Conclusions—Highe. TOPOLL expression is independently accounted with increased risk of cancer death among patients undergoing surgery for ccRCC, and the prognomic value is pronounced among patients with low stak diserbe. Evaluation of TOPOLL in ccRCC provides the opportunity to help guide postsurgical curveillance for ccRCC patients as well as inform the design of more targeted clinical trials and novel treatment strategies.

Keywords

Kidney neoplasms; Carcinoma; renal cel', Tum or biomarkers; Biologic; Survi 1

1. Introduction

Mortality rates for renal cell carcino in (RCC) have been ruling steadily for >3 decades [1]. During the same time period, there has been little change in 5 yr survival to provide the same time period, there has been little change in 5 yr survival to provide the same time period. (approximately 65%) [2]. Moreover, the small observes increases in survival can be attributed in part to a lead-time bias associated with a rise in the incidental detection of small, clinically dubid as tumors [3,4]. Taken together, these trends underscore the need to continue efforts to improve our understanding of the factors that predict KCC aggressiveness, particularly among the growing number of individuals diagoos d with low-risk RCC.

DNA topoisomerases are enzyme: that manage the 'sopologic state of DNA in the optiony introducing temporary single- or double-su and breaks in the DNA [5]. Through these strand breaks, the topoisomerase enzymes allow for a wide voliety of viscential LNLA metabolic

reactions including replication, manscription, recombination, and chromatin remodeling [6,7]. Several in restignative teams is ave reported that higher intratumor expression levels of top oiso nerase enzymes are on indicetor of poor prognosis in a variety of human cancers [8–10]. Of interest, drugs targeting topoisomerase enzymes have been developed and represent some of the most successful arugs used to area human malignancies [11]. Despite this wellinown action with cancer cogressiveness, the potential role of topoisomerases in the path ogenesis and prognosis of RCC remains unknown.

Me avated by this gap in understanding, we used two large independent cohort studies to analyze and van late the hypothesis that bigher tumor protein expression levels of the type Ha topoiccinetase (TOPOHE) are associated with increased risk of cancer-specific death following surgery for localized clear cell RCC (cck CC). Moreover, we explore the specific hypothesis that this essociation is more pronounced among patients with low-risk ccRCC.

2. Patients and methods

2.1. Patient selection

After institutional review) oard approval, we identified 663 patients treated with radical nuphreatomy or population-sparing surgery (NSS) for unitateral, sporadic, noncystic, organcol fined (ie, N0 or Nx, M0) ccRCC between 1990 and 2006 from the Mayo Clinic Rochester Nephrectomy Registry. Of these, 1464 gratients (%3%) had paraffin-embedded tissue blocks available for immunohistoch mical (PCC) staining and available outcome data, and this group represents our analytic cohort. For our validation, cohort, we identified 415 patients from the Mayo Clinic Florida Nephrectomy Registry for ated with radical nephrectoing of NSS for unilateral, poradic, noncystic ccRCC be ween 2000 and 2011. Of these, 337 (81%) had tissue blocks and outcome data available, and this group represents our validation cohort. We discuss further loss of cases in both collocity resulting from failure of IHC staining in the Results section. Of note, the underlying patient calchment areas for Mayo Rochester and vitayo Florida are separated by >1000 miles (1000 kilometers) and as such represent geographical up and culturally unique populations within the United States.

2.2. Data collection

For both cohorts, we abstracted follow-up data from the registry efforts at each institution. Briefly, these data are continely updated and maintained through a combination of active (mail-out questionnaires) and passive (medicol record, linkage to national databases) surveillance by experienced clinical coordinaters [12]. Loss to follow-up is <5% for orth registry efforts. In addition, we abstracted data on relevant clinic pathelogic covariates including age at surgery, gender, symptoms at presentation, Eastern Coordinater (ABCC) primary tumor classification, regional 'graph node involvement, distant metastases, the 2010 Affect TNM stage groupings, tumor size, naclear grade, and presence of congulative timer necrosis. To obtain the pathologic features in a standar lize I fashion, one applying and organized the output of the pathologic features in a standar lize I fashion, one applying and easing should be determined and the pathologic features in a standar lize I fashion, one applying and easing should be determined and the pathologic features in a standar lize I fashion of the action of the pathologic features in a standar lize I fashion of the action of the pathologic features in a standar lize I fashion of the action of the pathologic features in a standar lize I fashion of the action of the pathologic features in a standar lize I fashion of the metastory in the pathologic features in a standar lize I fashion of the metastory in the pathologic features in a standar lize I fashion of the metastory in the pathologic features in a standar lize I fashion of the metastory in the pathologic features in a standar lize I fashion of the metastory in the pathologic features in the pathologic features and the pathologic features and the pathologic features in the pathologic features and the pathologic features in the pathologic features at the pathologic features and the pathologic features at the pathologic fea

2.3. Type IIa topoisomerase expression

We identified a parafine encoder older with representative tumor tissue for each patient in both covorts and obtained a C pm-thick slide for IHC. Technicians in our core facility performed IHC staining for TOP OIIa using a monoclonal antibody and the respective protocol from Leica Arieror ystems (Buffato Grove, IL, USA). One of our study pathologists (J.C.C.) trained a certified current of Buffato Grove, IL, USA). One of our study pathologists (J.C.C.) trained a certified current he staining pattern was recorded as the average of the number of positive tumor cells in each or five representative high-powered fields using a Leica DMR microscope (Leica Microsystems, Wetzlar, Germany). With a 10/25 eyepiece and a ×40 objective, the Leica DMR has an objectime 1 diameter of 0.625 mm², resulting in a high-powered field of 0.307 mm². As each, TOPOIIa expression was quantified as the number of positive tumor cells per square millimete. For the purposes of evaluating intrarater reliability, we selected a random sample of 50 cross from the analytic cohort for review by the same cytotechnologist. Sin ilouty, to assess interrater agreement, we ran tom is sameled 100 cases from the analytic cohort for independent review by a urologic pathologist. (K.J.W.).

2.4. Statistica methods

For our exarysis in both cohorts, we explored the magnitude of the association of continuous TOI OIL expression and RCC-specific death by using Cox proportional hazards regression mode's and summarized the results with he card ration (nPs) and 95% confidence intervals (CIs). Smoothing spines were used to explore the functional + rm of the continuous TOPOIIa, v nich was quantified as the number of possive cella ser square millimeter, and it was determine a that the square-root transform, tion of TOPOL'a h, d a linear relationship with cancer-specific surrival (CS^{\circ}). Thus, for both the an^{\circ}/₁ uc ard validation cohorts, the continuous TOPOIIa variable was quantified as the number of positive cells per millimeter in the Cox regression models. In the Cox models, Vic first estimated the age-adjusted association of Torona expression with time to RCC-structfic destin. Then, to assess the association of TCTCITA expression with KCC-specific down after compelling for other known predictors of cci CC catcorie, we constructed Co's moder that adjusted for individual pathologic feetures chaggressively sa well as a convolute scoing system (Mayo Clinic stage, size, side, and necro is [SIGN] score). We also stated TOPOIIa expression as a dicho, mized variable (ie, high vs low). To estimate a cut point for dichotomizing TOPOIIa expression into high versus-low expression, we used the analytic cohort and chose the cut point that maximized the concordance index. As a result tanors with TOPOIIa expression ~16.6 positive cells per square millinget categorized as "Jow"; those ≥16.6 were categorized as "high" in the validation cohort, 've dichet united TOPOIIa expression using the same cut count as for the analytic cohort. We analyzed occure dance index values to compare the predictive ability of various models with and without the addition of the TOPOIIa expression variab¹. All concord, nee indices w/re internally validated using a bootstrap metho lolor, y proposed by Harrell et al [13] and therefore represent optimism-corrected estimates of prognostic accuracy. To further place the potential prognostic value of TOPOIIa correction, we evaluated Kaplan-Meica curves and HR estimates from Cox models stratified by May J Clinic SolCN score (12, lov. - 0-3, intermediate = 4-7, and high = 8-11) [14-16].

Finally, we used Pointon correlation coefficients to evaluate intra- and interrater agreement for our method of quantifying TOP Alla staining. Our statistical analyses were performed using the R programming language, v.2.15. All tests were two sided, and p values <0.05 were considered statistically significant.

3. Results

3.1. '.ssociation of type IIa topoixomerase with pathology and renal cell carcinomaspecific death (analytic cohort)

For the analytic cohort, 1378 of 1464 patients (94%) had successful staining of TOPOIIa, and the mannievel of TOPC IIa expression was 13.7 pusitive tumor cells per square millimeter (median: 7.0; min = 0.0 : ax = 2777). Of note, we observed no statistically significant differences in demographic or clinical features between the 1378 patients in our final cohort and the 235 who were excluded for 'ack o' tissue, follow-up, or successful IHC staining. In our dianotomization of TOPOIIa, 332 patients (24.1%) had tumors classified as higi (≥16.6 ^T.)POIIa-positive tumor cells ramminater). In Table 1, we provide a comparison of standard c inicopathologic features by dishotomized TOPOIIa status (low vs righ). Those tamors classified as TOFOII , high had more aggressive pathologic features including larger size (p < 0.0001), high er gic $\frac{1}{2}$ (p < 0.0001), later stage (p < 0.0001), presence of necrosis (p < 0.0001), sarconic toid features (z = 0.0006), and higher Mayo Clin c SS IGN score (p < 0.0001). Estimates of +1.c age-ad; ted and multivariable associations of TOPOIP: expression variables with cancer specific death are summarized in Table 2... When modeled as a continuous variable, we noted evidence of a linear increase in the risk of lancer specific death with increasing TCFOIIa expression (HR: 1.266; 95% CI, 1.210–1.326; $\mu < 2.0\text{E-16}$) after adjusting for a get "GrOIIa wis nodeled as the number of positive tu nor colle for millimeter. When we dichotomized TOP IIa expression, patients with high TC $^{\text{o}}$ OIIa eventsion were nearly three times more likely to experience cancerspecific death than patients who had low TOPOIIa expression (AR: 2.750; 95% CI, 2.123-3.561; p = 1.795-14) after adjusting for a e. Multivational adjustment for a variety of known predictors of cckCC outcome resulted in slight attenu tion of the as sociation of TOPOIIa expression with risk of ancer-sperific death (Table 2a). To quantify the prognostic ability of TOPOIIa, we provide e aunates of the optimism-correct d conco. datue incides for models with and without adjustment for TOPOIIa in Table 2b. Of ante, we observed similar HRs and 95% CIs when we modeled recurrence of dilease as the end point of interest indead of death from RCC (data not shown).

3.2. Association of type IIa topoiscinerase with pathelogy and rener cell carcinomaspecific death (validation cohort)

For our validation cohort, 27 i of 337 patients (22%) had successible claiming of TOPOIIa, and the mean level of TOPOIIa expression was 9.9 positive tumor cells per square millimeter (median: 3.5; minimum = 0.0, maximum = 347.6). In our dic totomization of TOPOIIa, 35 patients (12.5%) had tumors classified as TOPOIIa high (\geq 16.5) TOPOIIa-positive tumor cells per square millimeter). Again two provide a comparison of stemard clinicopathologic features by dicholomized 10POIIa status (low vs high) in fable 1. As with the analytic cohort, we noted evidence that FOPOIIa-high tumors have more aggressive

pathologic features including higher grady (p = 0.0013). For comparison with our analytic cohor, we provide estimates of the age-adjusted and multivariable associations of TOPOIIa expression with cancer-specific death in Table 2a. Similar to the analytic cohort, we noted evidence of a linear increase in risk of cancer-specific death with increasing TOPOIIa expression (HR: 1.23°, 95% CI, 1.114–1.278; p = 7.68E-5) after adjusting for age; TOPOIIa was moduled as the number of positive turnor cells per millimeter. Moreover, when using the same cut point we used in the analytic cohort to dichotomize TOPOIIa expression, we once again noted that patients with high TOPOIIa empression were more than three times more likely to experience cancer-specific death than potents with low TOPOIIa expression (HR: 2.443; 95% CI, 4.337-8.877; p = 0.0104) after adiments with low TOPOIIa expression (HR: 2.443; 95% CI, 4.337-8.877; p = 0.0104) after adiments with low TOPOIIa expression (HR: 2.443; 95% CI, 4.337-8.877; p = 0.0104) after adiments with low TOPOIIa expression (HR: 2.443; 95% CI, 4.337-8.877; p = 0.0104) after adiments with low TOPOIIa expression (HR: 2.443; 95% CI, 4.337-8.877; p = 0.0104) after adiments of a ge. Interestingly, multivariable adjustment for known predictors of ccRCC outcome did not result in attenuation of the association of TOPOIIa expression, with risk chan cancer-specific death (Table 2a). We again closer real summar HR and 95% CIs when vie n odeled recurrence of disease as the end point of interest and of death from RCC (data not shown).

3.3. Stratified analysis (analytic cohort only)

Figure 1 displays the over all disparity in CSs for patients with low- and high-TOPOIIa expression (log-rank p = 1.55E-15). To ill istrate the potential prognostic value of TOPOIIa, we evaluated use ability of dichotomized TOPOIIa to further stratify patients following initial classification by Mayo Clinic SSIGN score. We note that this was primarily evident among lo v-risk patients (Fig. 20; log-rank p = 8.62E-5) and, to a slightly lesser extent, intermedia e-risk patients (Fig. 20; log-rank p = 0.0036). In contrast, TOPOIIa expression had a limited ability to stratify patients already predicted to be a high risk of RCC-specific death based on SSIGN score (Fig. 20; log-rank p = 0.33). In Table 3, we provide the ageadjusted F R and 95% CI must correspond with our stratified Kaplan-Meier curves. Given that in our validation cohort only six patients who were classified as low risk using the SSIGN score experienced a ccRCC-specific death, we and not attempt stratified analyses in this cohort.

3.4. Inter- and intrarate reliability for type lla copoisomerase stating lovel

We observed a high level of intrarater reliability for the quartation of FOPOIIa staining. The correlation coefficient among the 50 cases that were read trailed in a blinded fashion by our cytotechnologist (T.f.) was 0.77 ($p \le 0.001$). Similarly, we noted a high level of interrater reliability for the quantitation of TOPOIIa staining. The correlation coefficient among the 100 cases read by our cytotechnologist (T.H.) and an experimence are logic pathologist (K.J.W.) was 0.70 ($p \le 0.001$).

4. Discussion

Key advancements in the management of ccRCC patients continue to center of the need to more accurately pinpoint postoperative rick for cancel-related death to better infortin pritient surveillance and streamline the design of the next generation of clinical trues [17]. Related to this, there is a parallel need to i lent fy molecular features within ccRCC current rices that represent markers of disease as gressiveness, predictors of treatment response, and rational targets for the development of more the apeutics [18]. We present the original targets for the development of more the apeutics [18].

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supporting the finding that higher expression of TOPOIIa is associated with an increased risk of cancer-related death following surgery for clinically localized ccRCC. Aspects of our report that increase the ralue of our findings include (1) these associations remained after adjustment for known predictors of ccRCC aggressiveness, (2) we noted a specific association among the patients with low the disease (the largest growing subset of ccRCC patients), and (3) we ter ort high inter- and increase agreement with regard to quantifying TOP offa in ccRCC tiss ass.

Implications of our current results we trane ruther A scussion. Primary among these is the potential for 10 POIIa (or any single biomarker) to emerge as a meaningful clinical tool in the posterior inanagement of cirCC patients. In fact, although we have shown that TOPOIIa remains statistically associated with cancer-specific death after adjustment for SSIGN score and up? (p = 0.0099), the circle van improvement in concordance index deserver furthe discussion. The inco por tion of TOPOIIa into a prognostic model with $a_{\pm} e a_{\pm} SSIGN$ score will improve prediction for roughly 5 in 1000 patients. However, we have demonstrated in Table 3 and Figure 2 that this improvement in prediction will largely trice place for muchts who would other wise be classified as low risk using standard clini opathologic indices. We have proviously advorated for the sequential or stepwise use of tranor-based biomarkers in determining postcargica' ccRCC prognosis [19]. In other words, rather than seeking to immutably integrate a particular biomarker into an existing algo ithn, we support the use of biomarkers on an as-need, basis. The most cost-effective value for any tumor-basic piom, rker rests on its county to firs determine prognosis for a patient using reauly available routine pathology-based indices and algorithms. This effort can then be followed by further prognostic refinement by biomerker testing, where physician: and their patients doem it necessary. Along these lives, bur data suggest that any value of TOPO, Ta as a prognostic marker would most lilibry be in the specific subset of patients with 'ow-risk disease. Patients are often a nonfoliable with the notion they are at low risk for developing metastatic disease after surgery. For many, this fear is exacerbated by the absence of any guidelines as to how they can 'ower their lisk further; the lack of a screening marker for early lefection of recurrent dise. 2., and the reality that if metastatic disease develops, no the apies offer durable success. In contrast, patients who have intermediate- or high-risk ccPCC are often plyced on more rapid sur, sillar ce protocols (ie, imaging at 6 mo rather than 1 yr) or even encouraged to encoll in adjustant trials. Our data support that staining and analysis of TCI OIIa could be offered to patients who have lowrisk disease to provide additional information regarding the probability that they are among the 5–10% who will provide the metastatic discrete and die from their cancer. Ultimately the clinical value of TOPOIIa (or any biomarker) will most likely to realized when ad uvant therapies for ccRCC are approved and unsee biomarkers can be examined for their ability to predict response to therapy.

Limitations of our current study war, ant further discussion. Chief an ong these is that our sample size for the validation cohort was smaller than the analytic cohort. That said it is worth noting that for our primary analysis (estimating association with TOP DLa), our validation cohort was adequately powered to report the same associations we observed in the analytic cohort as statistically significent. Additional limitations include patient, populations from tertiary referral centers with limited racial contents. The total of the focus

on only one enzyme in the topoisonerase family; quantitation of TOPOIIa staining that did not in orporate a measure of staining intensity; exclusion of individuals because of lack of start on SSIGN score; to more the use or railed IHC staining; and follow-up that, although start lard zed, was observational and not performed as part of a clinical trial. Nevertheless, the streng hs of this in restignation include the two-stage cohort design, use of the same cut point in both cohorts, pentralized pathology review, adjustment for well-known predictors of ccRCC outcome, additional stratified analyses, use of a commercially available monoclonal antibuly, and demonstration of high inter- and inter rater agreement for quantification of TCPOIIa expression.

5. Conclusions

We provide the first evidence that higher expression of TOPOIIa in ccRCC tissues is associated with an increased rist, of cancer specific death independent of other known pathologic medictors of RCC outcome. Mor wer, this association is more pronounced among patients with low-risk disease.

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Take-ho ne message

The ab lity to identify which priments who undergo surgery for clear cell renal cell carcinonia (ccRCC) will progress and die from their disease remains challenging. Evaluation of type IIa isopoist merase expression can be used to augment standard athologie indices and but identify ccRCC matters who have aggressive disease.

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Fig. 1.

Estimated cancer-specific survival following surgery by dichotomized type IIa opois surerase expression for 1378 pottents with clear cell renal cell carcinoma (analytic cohort). RCC = renal cell continuous; TCrOIIa = type IIa topoisomerase.

Fig. 2.

Estimated cancer-specifie survival following engry by dichotomized type IIa opois therase expression among patients with (a) low, (b) intermediate, and (c) high Mayo Clinic tage size, grade, and hecrosis stores. RCC = renal cell carcinoma; SSIGN = stage, size, grade, and necrosis; TOPCHa = type IIa topoisomerase.

Table 1

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	Analyti	c cohort $(n = 1378)$		Validatio	on cohort (<i>n</i> = 279	
	TOPOIIa low	TOPOIIa High	<i>p</i> value	TOPOIIa low	TOPOIIa High	<i>p</i> value
nder no. (%)			0.5515			0.4427
Female	358 (34.2)	120 (36.1)		81 (33.2)	9 (25.7)	
Male	688 (65.8)	212 (63.9)		163 (66.8)	25 (74.5	
ge at surgery, yr			0.7822			, 2001
Mean	63.0	63.4		02.8	65.1	
Median	63.9	64.6		63.6	67.2	
Range	19.8–90.2	27.8-88 2		27.2. 92.1	44.8-85.6	
mor size, cm			<0., 001			t 4366
Mean	5.8	6.8		4.9	5.1	
Median	5.0	0،ر		4.0	4.5	
Range	0.: -29.0	0.8-2 1.0		.3-14.5	2.0-12.0	
vM stage, no. (%)			<0.0001			0.9031
W'ssing	3	1		-	0	
	68. (65)	163 (49.2)		175 (72.0)	26 (/4.3)	
	: J1 (14.5)	37 (11.2)		3^ (13.2)	0 0.0)	
п	20° (19.9)	125 (37.8)		35 (14.4)	9 (2.5.7)	
	3 (0.3)	6 (1, ⁸)		1 (0.4)	r (0,r)	
icit ^{ar} grade, no. (° o)			<0., 001			u.0013
Missing	د	J		1	0	
1	5.1 (8.5	10 (,.0)		19 / 1 °,	0(0.0)	
2	518 49.5,	, n2 (31 c)		169 (69.5)	19 (54.3)	
3	2 98 (5 3.0)	175 (52.1)		47 (19.3)	13 (37.1)	
4	37(3.2)	42 (12.7)		8 (3.3)	3 (8.6)	
n zulative tumor neerosi , no. %)			<0.0001			0.8161
Mi sin _c	0	0		7	0	
No	876 (83 7)	201 (60.5)		198 (81.8)	28 (80.0)	
Y's	170 (16.3)	131 (39.5)		44 (18.2)	7 (20.0)	

	Analyti	ic cohort (<i>n</i> = 1378		Validati	ion cohort (<i>n</i> = 279	
	TOPOIIa low	TOPOIIa High	<i>p</i> value	TOPOIIa low	TOPOT.a His h	<i>p</i> value
Sarcomatoid differentiation, no. (%)			0.0006			0.4911
No	1038 (99.2)	320 (96.4)		. 40 (98.4)	34 (97.1)	
Yes	8 (0.8)	12 (3.6)		4 (1.6)	1 (2 °)	
SSIGN category, no. (%)			<0.0001			0 7772
Missing	161	40		٢,	0	
0-3	009 (75. j)	1.57 (53.8)		1 11 (72 9)	24 (6 ⁹ .5)	1
4-7	1.5 20.9	92 (2 (.5)		4.5(19.0)	8 '22.9)	
8+	31 (2.5)	4,7 (14.)		(2.1)	3 (8.6)	
T JPOI a = t pe II . topoiso. teras : SSI	GN = stage, 'ize, '	rade, and h. crosh.				

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(a) Age-adjusted and multivariable associations of type IIa topoisomerase expression with clear cell renal cell carcinoma-specific death in the nalytic cohort and validation cohorts; (b) optimism-corrected concordance indices in the analytic and validation cohorts*

	Analytic cohort		Validation cohort	
	HR (95% CI)	<i>p</i> value	HR (95% CI)	<i>p</i> value
Continuous TOPOIIa (square root)				
Adjusted for age	1.266 (1.210–1.326)	<2.0E-16	1.239 (1.114–1.378)	7.68E-05
Adjusted for age + gender	1.275 (1.217–1.336)	<2.0E-16	1.226 ($1.101 - 1.36c$)	J.000208
Adjusted for age + sarcomatoid	1.244 (1.188–1.303)	<2.0E-16	(792, 1-221, 198, 11.201, 197)	1.84E-05
Adjusted for age + TNM stage	1.171 (1.117–1.229)	8.88 3-11	1. 81 (1 071–1.30.)	9.000844
Adjusted for age + tumor size	1.171 (1. 20-1. 25)	3.11 3-12	1.2. 7 (1 34–1.393)	1 23E-05
Adjusted for age + nuclear grade	ر 142 (1.08، -1.19) ر	5.93E-08	1.211 (1.08 3-1.348)	0.6 10443
Adjusted for age $+ ne$, sis	1.1 '2 (1.061–1. 56)	9.02E-06	1.161 (.052 1.281)	0.0, 310
Adjusted for age + S5 'GN'. 30re	1.07 ' (1.018–1.139,	0.10 90	1.099 (0. ¹ 95–1 215)	5.u639
r ichotomous TC DOIIa high vs low)				
Adj' sted for a re	2. 50 (2 123–3.561)	1.7. ⁵ E-14	3.445 (1.337–8.877)	0.0104
tdjusted for $a_{c} e + g$ and r	2.76 1 (2.1 2 5.58 1)	1.33E-14	3.192 (1.233–8.264)	5. 168
djusted for age + sa. som toid	2.025 (2.022-2. tul)	4.08E-13	3.721 (1.431–°. 77)	1.00% 75
At justed for age + TN M stare	1 % 7.2 (1.438-2.438)	3.19E-06	3 = 2, (1.335 -9.41 5)	110.1
Adjus. rd for ree + tr nor size	2.397 (1.848–3.110)	4.652-1	4. 106 (.493 0.74 ')	0.6 1585
Au, usted for age nuclear grade	1.805 (1.385-2.53)	1.24 E-05	2.6.5 (1. 14-7. 60)	0 ,467
Adjusted for age + necrosis	1.726 (*.322-2 252)	5.39E 05	3.705 (1.40 7-9.7)5)	رد 300.0
Adjusted for age + SSIGN score	1.5, 2 (1. 07-2 037)	و 3800.1	. 232 (1 114–9.37),	0.0309
<u>h.</u>				
	naluti	in the start		Validation

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Including TOPOIIa

Without TOPOIIa

Including TOPOIIa

in ut l PPOIIa

0.665 (0.0232) 0.657 (0.0504)

0.605 (0.0063) 0.594 (0.0208)

0.663 (0.000270) 0.662 (0.00471)

0.580 (0.00214) 0.575 (0.00287)

v'ontin 'tou. TOP On a (squ'ire rou).

/ e - gender

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	Analyti	c cohort	Validati	on cohort
	Without TOPOIIa	Including TOPOIIa	Without TOPOII •	Includin 7 TOPOIIa
Age + sarcomatoid	0.606 (0.000114)	0.676 (0.00455)	0.649 (0.0314)	0. 585 (0266)
Age + TNM stage	$0.810\ (0.00108)$	0.826 (0.000984)	1.785 (0.0105)	0.8. 2 (0. 130)
Age + tumor size	0.787 (1.34E-05)	0.812 (0.1 00605)	0.75 1 (0.00498)	0.832 (0.00 31)
Age + nuclear grade	0.789 (0.000524)	0.8(9 (0. 0149)	, 848 (t 00871)	P.320 (t. 0085.5)
Age + necrosis	0.705 (0.,00935)	P.S. 7 (0.0184)	1.82. (0.00 115)	0.833 (n)187,
Age + SSIGN score	0.873 (1.00, 73)	0.87° (0.0 0757)	(. 877 (J. 90725)	U.885 (0.00775)
Dichotomous TOPOIIa (high vv low)-				
Age	0.58 (0.0 259)	0.6~2 [3.2' E-05]	(605 (0.0069)	0.676 (0.0118)
A _{ξ} a + ξ and ar	0.576 9.00.84)	7.662 (0.0 326)	5.594 (0.0201)	0.667 (0.03°5)
$A_{\xi} $ $\Rightarrow + s_{i}$ rcont stoid	0.606 (0. 1007. 9)	0.6.'5 (0.0 110)	0.651 (0.0192)	0.687 (0.02 5)
Ag :+ TN M sta te	0.811 (0.0, 0265	0.828 (0.000557)	0.785 (0.0120)	7.808 (P. 5.135)
Ag + tume - size	0.785 (0.00 - 80)	0.806 (0.00260)	0.787 (0. 1118,	0.804 (0.01+2)
Age + nuclea grace	0.788 (0.000245)	0.807 (0.00130)	9.844 (0.0. 29)	(1021) 834
Age + necrosis	0.797 (1.55E-05)	0.815 (0. 1007. 3)	0. ?25 (0.005, 1)	0.8-, 5 (0.0.9)
Age + 3SIGN score	0.874 (0 000822)	0.879 (0.0, 0226	0.8.5 (0.00164)	0.883 (0. 90874)
a = مرينا فطعت المعالم ا	io; S ^r IGN = stage, s ₁ e	s, grade, and nec osis; TO	$\overline{OP}(\sqrt{IIa} = f_{y}pe IIa to, vol$	soi terase.

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Age-adjusted association (hazard ratio and 95% confidence interval) of high type IIa topoisomerase expression with cancer-specific death strot fied by

Mayo Clinic stage, size, grade, and necrosis score (analytic cohort only)

p value

HR (95% CI)

p value

HR (95% CI)

p value

HR (95% CI)

High risk (SSIGN = 8–11)

Intermediate risk (SSIGN = 4–7)

Low risk (SSIGN = 0-3)

Table 3

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SSIGN = stage, size, grade, and necrosis; HR = hazard ratio; CI = conF, conF, $inte val; ^{T} JPOIIe = typ$. Ila topoisomeras

0.50

1.48 (0 ° 1-2.71)

-0.035

1.52 (1.03-2.24)

0.0002

3.55 (1.80-6.98)

1.0 (reference)

I

1.0 (reference)

TOPOIIa Low High 3

1.0 (reference)