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Effect of hemoglobin adjustment on the precision of mercury concentrations in maternal and cord blood

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Abstract

The cord-blood mercury concentration is usually considered the best biomarker in regard to developmental methylmercury neurotoxicity. However, the mercury concentration may be affected by the binding of methylmercury to hemoglobin and perhaps also selenium. As cordblood mercury analyses appear to be less precise than suggested by laboratory quality data, we studied the interrelationships of mercury concentrations with hemoglobin in paired maternal and cord blood samples from a Faroese birth cohort (N = 514) and the Mothers and Children's Environmental Health study in Korea (n=797). Linear regression and structural equation model (SEM) analyses were used to ascertain interrelationships between the exposure biomarkers and the possible impact of hemoglobin as well as selenium. Both methods showed a significant dependence of the cord-blood concentration on hemoglobin, also after adjustment for other exposure biomarkers. In the SEM, the cord blood measurement was a less imprecise indicator of the latent methylmercury exposure variable than other exposure biomarkers available, and the maternal hair concentration had the largest imprecision. Adjustment of mercury concentrations both in maternal and cord blood for hemoglobin improved their precision, while no significant effect of the selenium concentration in maternal blood was found. Adjustment of blood-mercury concentrations for hemoglobin is therefore recommended.

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Keywords

blood analysis; hemoglobin; methylmercury; prenatal exposure; selenium

1. Introduction

The cord-blood mercury (Hg) concentration has been suggested as the best risk indicator in regard to methylmercury-associated developmental neurotoxicity (Grandjean et al., 2007). However, as a biomarker of prenatal methylmercury exposure, the total Hg concentration in cord blood is associated with imprecision that exceeds the level suggested by laboratory quality assurance data (Grandjean et al., 2005; Grandjean et al., 2007). Some of this imprecision may be due to variable binding of methylmercury (MeHg) to erythrocytes, in which mercury binds to hemoglobin (Sakamoto et al., 2004). Previous studies have documented that Hg concentrations are higher in cord blood than in the corresponding maternal blood, likely due to the easy transfer of MeHg through the placenta (Kajiwara et al., 1996; Morrissette et al., 2007; Iyengar et al., 2001), and the higher hematocrit in newborns compared to their mothers (Stern et al., 2003). For this reason, standardization of the blood-Hg concentration to the one in erythrocytes has been recommended (Sakamoto et al., 2004). Adjustment for the hemoglobin concentration would likely be even better, although the impact on the imprecision has not been determined so far.

Another factor of possible relevance is that selenium (Se) is thought to bind to MeHg (Harris et al., 2003), thus possibly affecting the toxicokinetics of the latter. Hence, Se status could conceivably interfere with the transplacental transfer of MeHg and thus the partition between mother and fetus. However, previous studies of Se-MeHg interactions have mainly focused on impacts on MeHg toxicity under particular exposure regimens that may not reflect human exposures. The earliest experimental studies showed that Se reduced the acute toxicity of MeHg injected into rats, thus suggesting the notion that Se may form complexes with MeHg in the blood, thereby decreasing the bioavailability of both elements (Ganther et al., 1972). More recent research in rodents supports that antioxidant nutrients, including Se, in the diet may alter the reproductive and developmental toxicity associated with MeHg exposure (Beyrouty et al., 2006). As Se is known to co-exist with MeHg in fish and sea mammals (Burger et al., 2007a; Burger et al., 2007b; Cabanero et al., 2005; Kaneko et al., 2007), a potential toxicokinetic interaction may occur in regard to transplacental transfer of MeHg from maternal seafood diets. Although human evidence on this possibility is not available, we considered Se as a covariate.

Imprecision of the exposure parameter is a crucial concern, because the exposure parameter in routine statistical calculations is usually treated as an independent variable without error (Grandjean and Budtz-Jørgensen, 2007). However, all biomarkers are subject to imprecision, and non-differential errors tend to bias the dose-response relationship toward the null (Fuller, 1987). To take into account the imprecision, a useful approach is to employ a structural equation model, where confounders and effect variables are included (Budtz-Jørgensen et al., 2002; Grandjean et al., 2007). In a Faroese birth cohort, the average total imprecision (expressed as the coefficient of variation) for the cord-blood Hg concentration

was found to be about 25% (Grandjean et al., 2005; Grandjean et al., 2007), a magnitude large enough to bias apparent dose-response relationships. The imprecision for hair Hg measurements is much greater. As only a very small part of such imprecision can be ascribed to laboratory variability, identification of other error sources is important.

Therefore, we assessed exposure biomarker imprecision and the impact of adjustment for hemoglobin and Se. We utilized data from birth cohort studies in the Faroe Islands and in Korea (Mothers and Children's Environmental Health, MOCEH).

2. Materials and methods

2.1 Subjects

A cohort of 514 singleton births was assembled at the National Hospital in the Faroe Islands during a 20-month period in 2007–2009. This North Atlantic population is of mainly Scandinavian origin, relatively uniform, and is covered by a modern health care service. The Faroese are of particular interest in environmental epidemiology, as pilot whale is among the traditional food items eaten as part of their marine diet. Pods of this small whale species are occasionally caught and the meat and blubber are shared locally. Because of the high MeHg concentration in the meat (Julshamn et al., 1987), the Faroese have a high average exposure to this contaminant and a wide range of exposure levels that depend on whale availability (Budtz-Jørgensen et al., 2004; Grandjean et al., 1992). At parturition, we obtained whole blood from the cord immediately after clamping. Maternal blood and hair was obtained approximately two weeks after parturition, when the mother brought the infant in for a scheduled health check-up. Cord blood and maternal blood were analyzed for total Hg and hemoglobin, while the maternal blood was also analyzed for Se. Complete samples sets were available from 514 subjects.

The MOCEH study was carried out in Korea to determine the effects of maternal environmental exposure on fetal and postnatal growth and development. All pregnant women living in the targeted study site (i.e., Seoul, Cheonan, and Ulsan), who were in their first trimester of pregnancy at the time of screening, were eligible. The maternal blood samples were obtained during late pregnancy (28–42 gestational weeks). The recruitment period in the MOCEH study was from 2006 to 2008 for the hospitals and clinics. A total of 921 women were eligible, and 124 women were not included due to exclusions or missing values in major variables. The study subjects were restricted to those for whom the maternal and cord blood Hg levels were assessed. Overall, 797 participants were eligible for this study.

2.2 Subjects and biomarker assessment

The Faroese blood samples were analyzed for total Hg by cold-vapor atomic absorption spectrometry on a Direct Mercury Analyzer DMA-80 (Milestone, Italy). Maternal hair-Hg was determined as previously described (Dalgard et al., 1994) by flow injection cold-vapor technique (FIMS-400 and AS-90; PerkinElmer, Wellesley, MA, USA). Quality assurance data showed imprecision levels below 5% in agreement with previous studies (Grandjean et al., 2005; Grandjean et al., 2007). Selenium in blood was measured by Zeeman atomic furnace absorption spectrometry using an AA800 Zeeman Atomic Absorption

Spectrophotometer fitted with THGA endcapped L'vov type graphite tubes, an EDL Selenium lamp current 280 mA, and autosampler AS90 (all from Perkin Elmer). In regard to imprecision, the relative standard deviation was 3.5% for within-run repeatability and 5.2% between runs. Hemoglobin was determined by spectrometry on a Cobas Mira analyzer. Total hemoglobin was determined in 50 μ L whole blood by Drabkin's cyanomethemoglobin method on a Cobas Mira S analyser (Roche, Basel, Switzerland) using a 9.3 mmol/L hemoglobin standard from Cypress Diagnostics (Langdorp, Belgium) for calibration.

Mercury exposure biomarkers in Korea have been previously described (Kim et al., 2011). The analytical method for total Hg was similar (DMA-80, Milestone, Italy) with standard quality control procedures. The coefficient of variation (CV) of the Hg concentration between runs and within each run was 3.0% and 2.5%, respectively. Hemoglobin concentrations in MOCEH samples were analyzed using sodium lauryl sulfate (SLS)-hemoglobin (Sysmex XE 2100D, CA, USA).

2.3 Statistical analysis

Following descriptive analyses, logarithmic transformations were used for the mercury biomarkers due to skewed distributions, and geometric means were calculated. Interrelationships between the transformed exposure biomarkers were determined by Pearson's correlation coefficients.

The dependence of the cord blood Hg concentration on hemoglobin and Se was first assessed in two regression models. In both models the cord blood Hg concentration was the dependent variable while hemoglobin and Se were independent variables. As an additional covariate, Model A included the maternal Hg blood concentration while Model B instead used the maternal hair concentration for comparison purposes. The adjustment for another MeHg exposure biomarker is critical as it will remove a possible association between selenium and level of prenatal mercury exposure. The analysis will then reveal whether the cord blood mercury concentration depends on selenium in children with the same exposure level. Regression coefficients are expressed as change of the Hg concentration in cord blood associated with a doubling of other exposure biomarkers. We report two-sided *p*-values.

As a supplement, a structural equation model was developed for the exposure biomarkers (Budtz-Jørgensen et al., 2002). In these models, observed variables are assumed to depend on common latent variables which may in turn depend on covariates. Thus, each of the exposure markers (*M*-*Hg*) was assumed to be manifestations of the true (unobserved) exposure (*Hg*) and a random error ($\varepsilon_{\rm m}$):

$$\log(M-Hg) = \alpha_{\rm m} + \lambda_{\rm m} \log(Hg) + \varepsilon_{\rm m}$$

When a least three independent exposure markers are available, the model can be estimated. Here we chose to express the latent exposure on the scale of the cord blood concentration. This was done by fixing the factor loading (λ_m) at 1 for the cord-blood concentrations. With this restriction, a one-unit increase in log-Hg will on average lead to a one-unit increase in log cord blood Hg. In these models, a scale for the latent variables is required, but this

choice has no implications for the correlations between the markers and the latent variable or the size of the error standard deviation (Bollen, 1989; Sanchez et al., 2005). The imprecision of the exposure marker is then modelled in terms of the random error term, ε_m , which is assumed to follow a normal distribution with mean of zero and a standard deviation (σ_m). When a natural log transformation is used, standard deviation of the error term is a mathematical approximation to the error coefficient of variation (CV) of the untransformed concentrations (Budtz-Jørgensen et al., 2007).

Model 1 included only the three Hg concentrations (Fig. 2) while model 2 included the available information about hemoglobin (Hb) and Se as covariates (Fig 3). Thus, the latent mercury exposure (Hg) was allowed to depend linearly on all three variables. In addition, the cord blood Hg concentration was allowed to depend directly on the Hb in the cord blood and the Se level. Similar effects were included for the maternal blood concentration. These direct effects were incorporated by modifying the model equations to include selenium and Hb:

 $\log(M - Hg) = \alpha_{\rm m} + \lambda_{\rm m} \log(Hg) + \beta_{\rm mhb} \log({\rm Hb}) + \beta_{\rm mse} \log {\rm Se} + \varepsilon_{\rm m},$

This equation is very similar to the regression models described above, but the SEM has the advantage of using all available information to estimate a true (latent) exposure level and it takes into consideration the imprecision of the variables. Parameters were estimated using the maximum likelihood method and the fit of the more complex Model 2 was assessed by the chi-square test comparing the expected the observed covariance matrix and by Root Mean Square Error of Approximation (RMSEA) (Kline, 2012). A model is considered to have a good fit if the chi-square test has a p-value above 5% and RMSEA below 5%.

3. Results

The geometric means of the Hg concentrations in cord blood averaged 50% higher than those of post-parturition maternal blood samples in the Faroes and closer to 70% higher than the Korean maternal blood collected during pregnancy. The interquartile range (IQR) and total range for the Hg concentrations reflected substantial variation, whereas the range of Hb and Se was much narrower (Table 1). Mercury in maternal blood correlated well with concentrations in maternal hair and in cord blood, but much less closely with Se in maternal blood (Fig 1).

Multiple regression models showed a positive effect of Hb on the cord blood concentration (Table 2). This effect was statistically significant after adjustment for maternal blood concentration (Model A) and the maternal hair concentration (Model B). The possible effect of Se was more uncertain. Thus, no effect of Se was apparent in Model A, but a significant positive effect was seen in model 2 when adjustment included only mercury in maternal hair. However, blood-Se was closely associated with maternal blood-Hg, and the regression coefficient for maternal blood-Hg was greater than that for maternal hair-Hg as predictors of cord blood Hg.

We then performed SEM analyses to further explore the interrelationships between the exposure biomarkers and a latent exposure variable based on the analyses available from the Faroes (Tables 3 and 4). The extended SEM (Fig. 3) showed a good fit to the data (p = 0.32 for difference between observed and predicted covariances, RMSEA = 0.016). In agreement with the multiple regression analyses, this model showed a strong positive effect (p<0.001) of Hb on the cord blood-Hg concentration for fixed level of the latent exposure (Table 3). A similar effect was also seen for the maternal blood-Hg concentration. Selenium did not seem to affect the observed Hg concentrations, although it was strongly related to the latent mercury concentration (p<0.001) as an indication that higher MeHg exposures were associated with increased Se intakes.

In the simple structural equation model without information about Se and Hb (Fig. 2), the cord blood measurement was the most precise exposure marker, and the maternal blood measure was only slightly inferior, as reflected by the error CVs (Table 4, Model 1). As expected, the maternal hair concentration had the largest error component (Table 4). The extended SEM (Fig 3) showed the same ordering of the markers, but after adjustment for Hb, the error CV for the cord blood concentration decreased from about 11% (Table 4, Model 1) to about 3% (Table 4, Model 2), which is similar to the analytical imprecision.

4. Discussion

An imprecise exposure assessment may complicate confounder adjustment and lead to an underestimation of dose-effect relationships (Budtz-Jørgensen et al., 2003; Carroll, 1998). Therefore, this problem may be exaggerated by potential confounders that are correlated with the exposure. In regression analyses, inclusion of such variables – when measured with better precision – may further add to the average bias toward the null hypothesis (Budtz-Jørgensen et al., 2003), even in cases where the potential confounder has no independent effect on the outcome. Thus, assessment of total exposure biomarker imprecision is an important issue to environmental epidemiology studies. However, the imprecision of exposure data in epidemiological studies is usually unknown (Grandjean et al., 2007) and is often ignored.

The present study examined the precision of the cord-blood Hg concentration while considering the impact of adjustment for blood covariates. Using standard regression techniques we have demonstrated a strong positive dependence of the cord blood-Hg concentration on the Hb level, also after adjustment for other mercury exposure biomarkers. Because three exposure parameters were available, a structural equation analysis could also be conducted. This analysis confirmed the regression results and showed in addition that the Hg concentration in cord blood is less imprecise than the concentrations in maternal hair and maternal blood, and that adjustment for Hb further improves the precision of this exposure biomarker. A higher Hb concentration (or hematocrit) is known to result in greater MeHg binding in whole blood and thus a higher Hg concentration in erythrocytes (Sakamoto et al., 2004), although this procedure may be more cumbersome than the analysis of whole blood for both mercury and hemoglobin.

Hemoglobin concentrations in maternal blood are also of importance. In the present study, a lower hematocrit during pregnancy than two weeks after parturition is the likely explanation of the differences in blood Hg concentrations between the two study populations. Due to this difference, cord blood has a Hg concentration about 50% greater than the one in maternal blood two weeks after parturition and 70% higher than first-trimester maternal blood.

While the total Hg concentrations includes both MeHg and inorganic Hg species in routine analyses, the Hg concentration in cord blood reflects mainly the methylated form, for which the placenta does not constitute a barrier (Kelman et al., 1982). Thus, speciation of Hg in cord blood is not needed. In contrast, Hg in maternal blood may not be a precise measure of fetal MeHg exposure due to the presence of inorganic mercury. Still, our results suggest that the imprecision of maternal blood Hg remains far better than the one for hair Hg. It is also noteworthy that the loading factor for maternal blood Hg was higher in the Faroes study than in the Korean study. This result is probably due to the maternal blood in the Faroes being sampled much closer to parturition where the cord blood was obtained.

Available evidence suggests that the cord blood Hg concentration is the best risk indicator for MeHg neurotoxicity (Grandjean et al., 2002; National Research Council, 2000). Given the larger imprecision of the maternal hair Hg parameter and its known variation with hair type and hair color (Grandjean et al., 2002), this highly feasible and widely used exposure biomarker is less appropriate for prenatal MeHg exposure assessment. The present study shows that adjustment of blood-based biomarkers by the Hb concentration will further improve the precision of the Hg concentration in both maternal and cord blood. This adjustment may also provide better precision than in previous studies that rely on cord blood concentrations (Grandjean et al., 2002; Grandjean and Budtz-Jørgensen, 2007).

Among the limitations of this study, we did not analyze serum-Hg concentrations of maternal blood to obtain an indication of the amount of inorganic Hg species. Still, our results show that the imprecision based on whole-blood analysis only was only slightly greater than for cord blood. Also, we did not analyze Se in the MOCEH study. However, no effect of Se could be determined in the Faroes data, so this omission is probably inconsequential. The molar ratio between Hg and Se is relevant for possible interactions between the elements (Newland et al., 2006). The Faroese data show an average Se molar excess of about a 50-fold above Hg concentrations. This excess suggests that all Faroese children are Se sufficient and that MeHg toxicity is independent of any relative degree of Se sufficiency. Se in maternal blood showed a positive correlation with the maternal Hg concentration, which is consistent with other studies (Barany et al., 2003; Bates et al., 2006; Svensson et al., 1992), possibly because both substances originated, at least in part, from the same marine food sources (Karita et al., 2002; Svensson et al., 1995; Svensson et al., 1992). However, a positive correlation between Hg and Se in maternal blood does not necessarily imply that Se can moderate the toxicity of MeHg in humans. While Hg in cord blood was weakly associated with Se in maternal blood, the SEM revealed no statistically significant association between maternal Se and the cord blood Hg. These findings imply that any protective effects of Se against MeHg toxicity are not manifested via impaired transport through the placenta. The present study therefore shows that any toxicokinetic interference

by Se would have no important effect on the partition of MeHg between maternal and fetal circulations.

Recently, a Japanese study also found no significant positive correlation between Hg and Se concentrations in cord blood (Sakamoto et al., 2010). However, experimental studies examining Hg and Se exposure in relation to neurodevelopment have produced inconclusive evidence, especially at levels comparable to human exposures (Reed et al., 2006). A Faroese birth study failed to identify any protective effect of Se against MeHg neurotoxicity (Choi et al., 2008). Selenium therefore seems to be of marginal, if any, interest in regard to MeHg toxicity in populations that are Se sufficient.

The two cohort studies demonstrated that Hg concentrations in cord blood strongly depend on the Hg concentrations in maternal blood, but the regression coefficients (and their confidence intervals) for maternal blood differed between the Faroese birth cohort study and the MOCEH study. This difference may be due to the different times of maternal blood sampling in the two studies. In the Faroes cohort, the maternal blood was obtained two weeks after parturition, i.e., at a time when the blood volume had approached non-pregnant levels. In the Korean cohort, maternal blood was sampled during late pregnancy when the blood volume was expanded, thus leading to lower concentrations of both Hb and Hg. While hemoglobin concentrations were taken into account, physiological differences in MeHg distribution in the body may have played a role as well. Also, the greater interval between paired sample collections in the Korea study may have increased the variability of relative Hg concentrations in the blood samples. In addition, the results may be affected by distinct racial/ethnic/cultural differences, as biases of the chemical analyses would seem unlikely. Also, Caucasian hair grows faster than Oriental hair, and the hair growth rate generally decreases with age (Lam, 2005). Such ethnically-related differences are likely independent of seafood consumption (Caldwell et al., 2009) but may affect the incorporation of Hg in hair.

A previous calculation of the total imprecision (Budtz-Jørgensen et al., 2003) suggested that the cord-blood Hg concentration had a total imprecision CV of about 25%, while maternal hair had one almost twice as high. In this calculation, the frequency of whale meat consumption was used as the third parameter needed to calculate the latent exposure variable. Because the three variables are independent and reflect different aspects of dietary MeHg exposure during pregnancy, the CV of 25% is probably a good reflection of the total variability. The present study included the maternal blood Hg, which constitutes the source of the MeHg that leads to a Hg concentration in the cord blood. The same applies to the hemoglobin in both blood samples, both of which likely depend on the maternal iron status. Thus, these measurements may not be fully independent, and the error variance of the latent variable may be underestimated for this reason. However, the error SD decreased from 0.11 to 0.03 after adjustment for hemoglobin in both maternal blood and cord blood in Faroese study. Selenium alone had only a negligible effect. Thus, although the variables may not be fully independent, the improvement in error CV achieved by hemoglobin adjustment would strongly support this way of expressing the blood Hg concentrations. Further, the imprecision of the hair-Hg remained much higher than for the blood-Hg concentrations.

Due to imprecisions of exposure parameters, calculated associations will tend to be biased toward the null, and inclusion of covariates may add to the bias (Budtz-Jørgensen et al., 2003). Most studies on MeHg biomarker associations have applied standard regression equations (Akagi, 1998; Bjornberg et al., 2003; Haxton et al., 1979; Phelps et al., 1980; Sherlock et al., 1982; Turner et al., 1980). An important limitation is that this approach includes measurement error only in the dependent variable (Fuller, 1987). A structural equation model allows for measurement error in both variables, and may therefore be a more appropriate choice for estimation of the relation between the two biomarkers.

The findings of this study support the use of cord blood as the best available exposure biomarker for MeHg, and its precision can be improved by adjustment for hemoglobin. Hg in maternal blood is also appropriate as a risk indicator after adjustment for hemoglobin, although the sampling time must be carefully considered. However, hair Hg concentrations are less precise and may result in bias toward the null. Although Hg and Se concentrations in maternal blood showed significant correlations, only a weak association was found between Se in maternal blood and Hg in cord blood, thus suggesting that any protective effects of Se against MeHg toxicity do not depend on toxicokinetic interference.

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Abbreviations

AAS	Atomic absorption spectrometry
AGFI	Adjusted Goodness of Fit Index
AIC	Akaike information criterion
CB-Hb	hemoglobin in cord blood
CB-Hg	Hg in cord blood
CFI	Comparative fit index
CV	Coefficient of variation
GFI	Goodness of fit index
GM	Geometric means
Hg	Mercury
IQR	Interquartile range

MB-Hb	Hemoglobin in maternal blood
MB-Hg	Hg in maternal blood
MB-Se	Selenium in maternal blood
MeHg	Methylmercury
MH-Hg	Hg in maternal hair
MOCEH	Mothers and Children's Environmental Health
Se	Selenium
SEM	Structural equation model
SLS	Sodium lauryl sulfate
TLI	Tucker-Lewis index

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Highlights

- The cord-blood mercury concentration reflects risk of developmental neurotoxicity
- This concentration may be imprecise due to binding to hemoglobin
- Structural equations were used to compare mercury in maternal and cord blood samples
- Adjustment of mercury concentrations for hemoglobin improved their precision

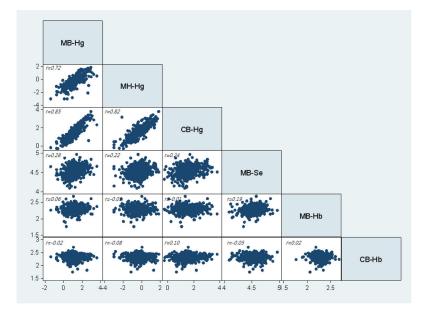


Figure 1.

Correlations between logarithmic transformations of biomarkers in the Faroese birth cohort. ^aMB-Hg, Hg in maternal blood; MH-Hg, Hg in maternal hair; Hg, latent variable; CB-Hg, MB-Se, Selenium in maternal blood; MB-Hb, Hemoglobin in maternal blood; Hg in cord blood; CB-Hb, hemoglobin in cord blood –

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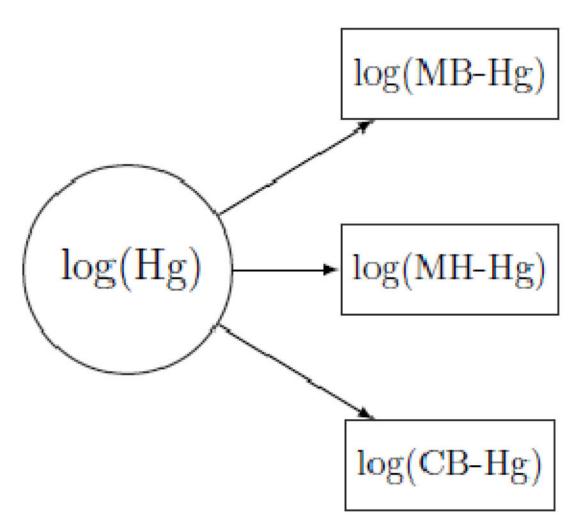


Fig. 2.

Structural equation model of latent methylmercury exposure (Hg), as indicated by mercury concentrations in maternal blood (MB-Hg), maternal hair (MH-Hg), and cord blood (CB-Hg), all of which are log transformed.

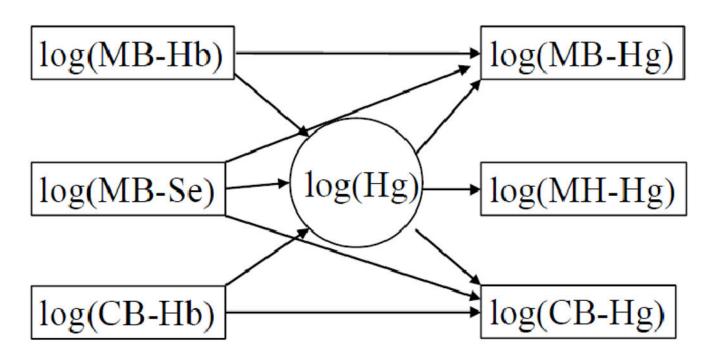


Fig. 3.

Extended structural equation model that allows the hemoglobin concentrations in maternal and cord blood to affect the mercury concentrations in the same media while also allowing the maternal selenium concentration to affect the blood-mercury concentrations.

Distribution summaries of mercury, selenium and hemoglobin.^a

Biomarkers	Ν	Geometric mean ^a	Interquartile range	Total range
Faroese birth cohort study				
Maternal blood Hg (µg/L)	499	15.30	10.4 - 21.15	1.25 – 188.1
Cord blood Hg (µg/L)	490	22.95	15.1 - 33.40	3.85 - 222.4
Maternal hair Hg (µg/g)	504	0.71	0.44 - 1.10	0.05 - 6.27
Selenium (µg/L)	496	1,132	1,025 - 1,235	718 – 1,715
Maternal hemoglobin (mmol/L)	498	9.73	9.14 - 10.40	5.84 - 14.57
Infant hemoglobin (mmol/L)	488	10.12	9.51 - 10.83	5.74 - 14.95
MOCEH study				
Maternal blood Hg (µg/L)	797	15.45	2.24 - 4.12	0.40 - 92.30
Cord blood Hg (µg/L)	796	25.90	3.95 - 7.00	0.23 - 24.10
Maternal hemoglobin (mmol/L)	558	11.92	11.30 - 12.80	6.10 - 17.70

 $^{\it a}$ To convert into nmol/L, multiply results in $\mu g/L$ by 5 (mercury) and 12 (selenium).

Linear regression coefficients for increase (in percent) in the cord blood mercury concentration associated with a doubling of different exposure biomarkers in the Faroese birth cohort.^a

	Cord blood mercury (µg/L)	
	Percent increase	95% CI
Model A		
Maternal blood mercury (µg/L)	82.7	76.6 to 89.0
Maternal blood selenium (µg/L)	1.26	-14.1 to 19,4
Cord blood hemoglobin (mmol/L)	53.3	27.6 to 84.1
Model B		
Maternal hair mercury (µg/L)	63.2	58.4 to 68.3
Maternal blood selenium (µg/L)	25.4	5.5 to 49.1
Cord blood hemoglobin (mmol/L)	79.0	47.0 to 118.0

Structural equation coefficients for increase (in percent) in mercury marker concentration in the Faroese birth cohort associated with a doubling in selenium and hemoglobin concentrations after adjustment for latent true mercury concentration.

Predictor	Percent increase	95% CI
	Effect on cord blood-Hg	
Cord blood-hemoglobin	74.71	49.44 to 104.3
Maternal blood-selenium	-5.66	-23.4 to 16.14
	Effect on maternal blood-Hg	
Maternal blood-hemoglobin	26.49	3.87 to 54.03
Maternal blood-selenium	15.75	-8.12 to 45.82

Estimated factor loading (λ) and error CV calculated for biomarkers of prenatal mercury exposure in a Faroese birth cohort by a structural equation model.

Biomarker	Factor loading	Error CV
Model 1		
Cord blood-Hg	1	0.105
Maternal blood-Hg	0.858	0.323
Maternal hair-Hg	0.981	0.420
Model 2*		
Cord blood-Hg	1	0.032
Maternal blood-Hg	0.833	0.324
Maternal hair-Hg	0.973	0.416

* Also adjusted for hemoglobin and selenium.