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The Temporal Course of the Influence of Anxiety on Fairness Considerations

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Abstract

This study investigated the potential causes of anxious people's social avoidance. The classic ultimatum game (UG) was utilized in concert with electroencephalogram (EEG) recording. Participants were divided into two groups according to levels of trait anxiety as identified by a self-report scale. The behavioral results indicate that high-anxious participants were more prone to reject human-proposed than computer-proposed unequal offers compared to their low-anxious counterparts. The event-related potential (ERP) results indicate that the high-anxious group showed a larger feedback-related negativity (FRN) when receiving unequal monetary offers than equal ones, and a larger P3 when receiving human-proposed offers than computer-proposed ones, but these effects were absent in the low-anxious group. We suggest anxious people's social avoidance results from hypersensitivity to unequal distributions during interpersonal interactions.

Keywords

Anxiety; social decision-making; Ultimatum Game; fairness evaluation; equality; event-related potential (ERP); feedback-related negativity (FRN); P3

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Author Contributions

Conceived and designed the experiments: YL, TW, CF, RG. Performed the experiments: YL, TW, CF. Analyzed the data: YL, TW, DZ. Wrote the manuscript: YL, TW, LSB, RG. Contributed materials and analysis tools: YJL. Provided lab equipment for running the study: YJL.

Conflict of Interest

The authors declare no conflict of interest.

Introduction

Linking Anxiety with Social Decision-making

Anxiety, an unpleasant emotional state characterized by automatic cognitive bias to threat-related information, plays a significant role in our everyday lives (D. M. Clark, 1999). In its exaggerated form, anxiety interrupts daily function such as social skills (J. V. Clark & Arkowitz, 1975). Patients suffering from anxiety disorders deliver poor performance in interpersonal interactions (Hirsch, Meynen, & Clark, 2004). For example, anxiety disorders are associated with spousal conflict (Whisman, Sheldon, & Goering, 2000). This social impairment may result from decision-making biases that lead anxious individuals to choose avoidance strategies in interpersonal situations (Grecucci et al., 2012). Indeed, levels of anxiety are negatively correlated with the willingness to communicate with strangers (Dumont, Nishida, & Nakayama, 2005; Samochowiec & Florack, 2010). When deciding whether to engage in a risky but potentially rewarding social activity (e.g., dating an unknown person), high-anxious people's expectations are more pessimistic than those of non-anxious people (Wray & Stone, 2005). Investigating the relationship between anxiety and social decision-making may have important implications for the treatment of social dysfunction and thus help improving anxious people's quality of life.

Social Decision-making Paradigm: Ultimatum Game

The Ultimatum Game (UG) is a classic paradigm for social decision-making research (Güth, Huck, & Müller, 2001; Güth, Schmittberger, & Schwarze, 1982). In this game, one of two players ("proposer") splits an amount of money between the two. The other player ("responder") may accept the offer, but also has the right to reject it, in which case neither players receive anything. Numerous studies have reported that the acceptance rates significantly decrease when responders receive disadvantageous unequal offers, often interpreted to indicate that people have a taste for fairness (Camerer & Thaler, 1995; see Feng et al., 2013 for the discussion on the relationship between equality and fairness). In addition, acceptance rates of offers from putative human proposers are lower than those from the computer, indicating a motivation to punish nonreciprocators during social interaction (Rilling & Sanfey, 2011; Sanfey, Rilling, Aronson, Nystrom, & Cohen, 2003). In other words, UG responders consider both the equality factor and the social-interaction factor when making decisions.

The processing of both factors might be affected by the feeling of anxiety. Anxious people's stronger emotional reactivity might lead to oversensitive reactions to inequitable distribution (Grecucci et al., 2012). However, anxious people might be less likely to reject unequal offers because they tend to avoid social confrontation with others (Mennin, Heimberg, Turk, & Fresco, 2002). The behavioral results of Grecucci et al. (2012) support the second hypothesis, which highlights the importance of the social-interaction factor. In contrast, our recent electrodermal study suggests that the skin conductance responses to unequal offers were positively correlated with levels of anxiety in the high-anxious group, indicating that the equality factor affected high-anxious people's emotional responses to offers (T. Wu, Luo, Broster, Gu, & Luo, 2013). Despite these discrepancies, both studies reveal that the impact

of anxiety on UG performance is significant. It was hypothesized that anxious people's UG behavior manifests a tradeoff between the equality and social-interaction factors, and these two studies may have captured different aspects of this tradeoff. The present study aimed to test this theory.

Electrophysiological Indices

Event-related potentials (ERPs) are an important neuroscience technique to relate behavioral performance and brain electrical activity during cognitive tasks (Rugg & Coles, 1996). By utilizing the ERP, the present study investigated anxious individuals' fairness considerations when evaluating UG offers. Two ERP components, namely the feedback-related negativity (FRN) and P3, were chosen for the analyses.

The FRN represents a rapid evaluation of the subjective value of current events along a "good-no good" dimension, with "no good" events elicit a larger FRN than "good" events (Holroyd, Hajcak, & Larsen, 2006; Holroyd, Larsen, & Cohen, 2004). During the UG, the FRN amplitude is more pronounced in response to unequal offers than to equal offers (Alexopoulos, Pfabigan, Lamm, Bauer, & Fischmeister, 2012; Boksem & De Cremer, 2010; Campanha, Minati, Fregni, & Boggio, 2011; Van der Veen & Sahibdin, 2011).

The P3 component also plays an important role in the evaluative process; its amplitude responds to all the major properties of an event. For instance, many studies suggest that during monetary gambling, the FRN responds specifically to outcome valence while the P3 encodes both valence and magnitude (Gu et al., 2011; Philiastides, Biele, Vavatzanidis, Kazzer, & Heekeren, 2010; Y. Wu & Zhou, 2009). In the same vein, social decision-making literature suggests that P3 amplitude is sensitive both to equality considerations and to social context, such that unequal offers elicit a larger P3 response than equal offers (Hewig et al., 2011), and the same is true when contrasting other-relevant and self-relevant properties (Y. Wu, Hu, van Dijk, Leliveld, & Zhou, 2012). Consequently, the FRN and P3 were chosen to detect the brain processing of the equality and social-interaction factors.

Previous research has linked FRN and P3 amplitudes to anxiety. High-anxious participants show a smaller FRN elicited by negative feedback during monetary gambling (Gu, Ge, Jiang, & Luo, 2010; Gu, Huang, & Luo, 2010). In response to novel stimuli, the P3 became larger as anxiety levels increased (Grillon & Ameli, 1994), which might be linked with an altered mechanism of attention allocation (see Polich, 2007 for the relationship between the P3 and attentional resources). The neurotransmitter activities underlying the FRN and P3 components, including dopamine (Holroyd & Coles, 2002) and norepinephrine (Polich & Criado, 2006), are related to anxiety symptoms (Goddard et al., 2010; Yoshioka, Matsumoto, Togashi, & Saito, 1996). These findings support our idea that both the FRN and P3 would be sensitive to anxiety during social decision-making.

Experimental Hypotheses

In the current study, the classic UG paradigm was utilized with electroencephalogram (EEG) recording. Due to the characteristics of the FRN and P3 described above, we predicted that both components would evidence the influence of anxiety on the equality factor, such that the main effect of equality would be more prominent in the high-anxious group than in the

low-anxious group. Further, we predicted that the influence of anxiety on the social-interaction factor would be evidenced on the P3, such that the main effect of social interaction would be more prominent in the high-anxious group than in their low-anxious counterparts.

Method

Participants

Two hundred and fifty-nine undergraduate students (all Chinese) were recruited from the campus of Beijing Normal University. They completed the Chinese version of the Trait form of Spielberger's State-Trait Anxiety Inventory (STAI-T), which has demonstrated good internal consistency reliability, convergent validity, and discriminate validity (Shek, 1993; Spielberger, Gorsuch, Lushene, Vagg, & Jacobs, 1983). In this sample, individuals with STAI-T scores in the lower 25% of the distribution were assigned to the low-trait anxiety (LTA) group, whereas individuals in the upper 25% of the distribution were assigned to the high-trait anxiety (HTA) group (Gu, Ge et al., 2010; T. Wu et al., 2013). From those who were suitable for these criteria, we randomly chose 40 students and asked them to participate in the formal experiment, with both the LTA group and the HTA group consisting of 20 participants (LTA group: 8 females, mean age: 22.7 ± 2.0 years; HTA group: 12 females, mean age: 20.4 ± 2.2 years). An independent-samples *t*-test confirmed significantly lower levels of trait anxiety among LTA participants (31.2 ± 3.9 ; [23, 36]) compared to HTA participants (57.0 ± 6.6 ; [48, 69]) ($t(38) = 15.13, p < .001$). Participants had normal or corrected-to-normal vision and had no history of psychiatric, medical, or neurological illness. All were right-handed. All participants provided written informed consent prior to the experiment.

In appreciation of the strong relationship between anxiety and depression (L. A. Clark & Watson, 1991; Stavrakaki & Vargo, 1986), the Chinese version of Zung's self-rating depression scale (SDS) was used to assess self-reported symptoms of depression in order to control the potential effect of depression during statistical analysis. Both clinical and non-clinical studies have established excellent reliability and validity of the SDS (Shu, 1993; Zung, Richards, & Short, 1965). An independent-samples *t*-test revealed that the two groups differed significantly in SDS scores (LTA: 29.0 ± 6.0 , HTA: 40.0 ± 6.0 ; $t(38) = 5.71, p < .001$).

See Figure 1 for the distributions of the STAI-T and SDS scores. The correlation between STAI-T and SDS was not significant in either the LTA group ($r = -.109, p = .649$) or the HTA group ($r = .247, p = .293$).

Procedure

The task procedure replicated that of Wu et al. (2013). To reinforce the social nature of the task, participants were told that they would play the UG together with three other anonymous college students, but in fact no other people were playing the game. Participants received no further information about the identities of the other supposed players. Before the task, each participant was instructed about the rules in the UG and was informed that his/her

payment depended on his/her choice in the task; the higher the score a participant earned, the higher payment s/he would receive at the end of the experiment.

Figure 2 illustrates the experimental procedure of an exemplar trial. Each trial began with the presentation of a central fixation cross for 1.5 - 2 s (randomized across trials). Afterwards, an offer proposed by a pseudo-player or the computer was presented for 2 s. Participants played as responders and decided whether or not to accept the offer by pressing the “F” or “J” buttons on the keyboard with their left or right index fingers (the button assignments were counterbalanced across participants). After participants finished their decisions, they waited for 0.8-1.2 s and received the feedback which informed about the results of the current trial.

There were six blocks in the task. In three of the blocks (human-proposer blocks), participants were told that the offers were proposed by the other three players, whereas in the other three blocks (computer-proposer blocks), the offers were proposed randomly by the computer. Each block consisted of 50 trials (300 trials in total), with each block separated by a short interval. Before the start of the task, participants were told that the offers in human-proposer blocks were randomly selected from the other three players’ proposals, that they would be unable to identify which proposer suggested the offer in a given trial, and that the “proposers” would not be able to know whether the participants accepted or rejected their offers (Boksem & De Cremer, 2010). Unbeknownst to the participants, all offers were actually proposed by the computer in predetermined pseudorandom sequences. Each block included 20 equitable trials (10 offers of 50:50 and 10 offers of 40:60), 20 inequitable trials (10 offers of 10:90 and 10 offers of 20:80), and 10 moderate inequitable offers (30:70).

Electrophysiological Recording and Measures

The electroencephalogram (EEG) was recorded from 64 scalp sites using tin electrodes mounted in an elastic cap (NeuroScan Inc.) with an online reference to the left mastoid and off-line algebraic re-reference to the average of the left and right mastoids. The electrooculogram (EOG) was recorded for the purpose of artifact correction. Horizontal EOG was recorded from electrodes placed at the outer canthi of both eyes. Vertical EOG was recorded from electrodes placed above and below the left eye. All inter-electrode impedance was maintained at $< 5 \text{ k}\Omega$. EEG and EOG signals were amplified with a 0.05 – 100 Hz online band-pass filter and continuously sampled at 500 Hz/channel.

During the offline analysis, ocular artifacts were removed from the EEG signal using a regression procedure implemented with Neuroscan software (Semlitsch, Anderer, Schuster, & Presslich, 1986). After 30 Hz low-pass digital filtering through a zero phase shift, the EEG data were epoched time-locked following the onset of each offer. Separate EEG epochs of 1000 ms were baseline-corrected by subtracting from each sample the average activity of that channel during the $-200 - 0 \text{ ms}$ baseline period. Any trial in which EEG voltages exceeded a threshold of $\pm 100 \mu\text{V}$ during the recording epoch was excluded from further analysis. As a result, $15.3 \pm 14.0\%$ of the epochs were rejected across participants.

According to previous literature, there are essentially two ways to calculate the FRN amplitude, that is, either using the grand-averaged waveforms or creating difference wave between “error” and “correct” trials (Holroyd, Pakzad-Vaezi, & Krigolson, 2008). The latter approach is based on the assumption that the FRN appears only on “error” trials, but this has been criticized because the differential ERPs could be caused by neural activity associated with either condition (Foti, Weinberg, Dien, & Hajcak, 2011; Holroyd et al., 2008). In addition, we are interested in the group effect on the FRN in both the unequal condition and the equal condition. Thus in the current study, we directly measured the grand-averaged waveforms rather than the difference wave so as to observe the ERPs in two conditions separately.

FRN amplitudes were measured as the mean values within the 300 – 350 ms time window following offer presentation. According to visual inspection of grand-averaged waveforms and their scalp distributions, the FRN was maximal in the frontal area. Hence, six frontal electrodes in this area (F1, Fz, F2, FC1, FCz, FC2) were chosen for further analysis of the FRN. P3 amplitudes were measured as the mean values within the 400 – 600 ms time window. The P3 was maximal in the parietal-occipital area. Hence, four parietal-occipital midline electrodes (CPz, Pz, POz, Oz) were chosen for analysis of the P3.

Data Analysis

Acceptance rates were analyzed using three-way Equality (equal vs. unequal) \times Proposer (human vs. computer) \times Anxiety (LTA vs. HTA) ANOVA test, while the Electrode was added into the analyses on the FRN and P3 amplitudes as the fourth factor. As mentioned above, the Electrode factor included six levels for the FRN and four levels for the P3. The acceptance rates for equal (i.e., 40:60 and 50:50) and unequal (i.e., 10:90 and 20:80) offers were calculated respectively (Harle & Sanfey, 2007; Hewig et al., 2011). Moderate unequal offers (30:70) were excluded from analysis, since previous studies suggest that UG players hold diverse opinions about whether such offers should be regarded as fair (Halko, Hlushchuk, Hari, & Schurmann, 2009; Hewig et al., 2011; van't Wout, Chang, & Sanfey, 2010).

For all the analyses, the significance level was set at .05. Greenhouse-Geisser correction for ANOVA tests was used when appropriate. Post-hoc comparisons were evaluated using the Newman–Keuls method. Partial eta-squared (η_p^2) values were provided to demonstrate effect size where appropriate, such that 0.05 represents a small effect, 0.10 represents a medium effect, and 0.20 represents a large effect (Cohen, 1973).

Results

Behavioral Results

See Figure 3 for the illustration of behavioral data. A significant main effect of the Equality factor was confirmed ($F(1, 38) = 182.75, p < .001, \eta_p^2 = .83$), with the acceptance rate being lower for unequal offers ($34.9 \pm 4.6\%$) than equal offers ($97.7 \pm 1.1\%$). The main effect of the Proposer factor was also significant ($F(1, 38) = 39.35, p < .001, \eta_p^2 = .51$), with the acceptance rate being lower when participants received offers in the human-proposer

condition ($58.0 \pm 2.3\%$) than those in the computer-proposer condition ($74.5 \pm 3.1\%$). The Proposer \times Equality interaction was significant ($F(1, 38) = 36.55, p < .001, \eta_p^2 = .49$). The Proposer \times Anxiety interaction was significant ($F(1, 38) = 6.86, p = .013, \eta_p^2 = .15$). The Proposer \times Equality \times Anxiety interaction was significant ($F(1, 38) = 5.57, p = .022, \eta_p^2 = .13$); while both the LTA and the HTA group accepted fewer unequal offers proposed by human than by computer, this effect appeared to be weaker in the LTA group ($F(1, 38) = 6.71, p = .013, \eta_p^2 = .15$) than in the HTA group ($F(1, 38) = 35.51, p < .001, \eta_p^2 = .48$). Finally, when adding depression as a covariate in the ANOVA test, the Proposer \times Equality \times Anxiety interaction ($F(1, 37) = 5.58, p = .024, \eta_p^2 = .13$) remained significant, indicating that the influence of anxiety on UG performance was not attributable to the effect of depression.

ERP Results

The FRN component—See Figure 4 for the FRN results in each group. The main effect of Electrode was significant ($F(5, 190) = 13.77, p < .001, \eta_p^2 = .27$), with the FRN being larger at Fz ($-1.43 \mu\text{V}$) and F2 ($-1.30 \mu\text{V}$) than other electrodes selected (p values $< .05$), but these two sites showed no significant difference with each other ($p = .132$). No significant main effect of the Equality factor ($F(1, 38) = 1.90, p = .176, \eta_p^2 = .05$), the Proposer factor ($F(1, 38) = 0.94, p = .338, \eta_p^2 = .02$), or the Anxiety factor ($F(1, 38) = 0.27, p = .604, \eta_p^2 = .01$) was detected. A significant interaction of Equality \times Anxiety was observed ($F(1, 38) = 4.74, p = .036, \eta_p^2 = .11$); no FRN difference reached significance in the LTA group ($F(1, 38) = 0.32, p = .575, \eta_p^2 = .01$), but the FRN was larger in response to unequal offers ($-1.60 \mu\text{V}$) compared to equal offers ($-0.89 \mu\text{V}$) in the HTA group ($F(1, 38) = 6.31, p = .016, \eta_p^2 = .14$). No other significant interaction was found (F values $< 2.94, p$ values $> .05$). When adding depression as a covariate in the ANOVA test, the Equality \times Anxiety interaction remained significant ($F(1, 37) = 4.73, p = .036, \eta_p^2 = .11$), indicating that the influence of anxiety on the FRN was not attributable to the effect of depression.

The P3 component—See Figure 5 for the P3 results in each group. The main effect of Electrode was significant ($F(3, 114) = 14.62, p < .001, \eta_p^2 = .28$), with the P3 being larger at Pz ($5.14 \mu\text{V}$) and POz ($5.06 \mu\text{V}$) than other electrodes selected (p values $< .05$), but these two sites showed no significant difference with each other ($p = .075$). The main effect of Equality was significant ($F(1, 38) = 12.85, p = .001, \eta_p^2 = .25$), indicating that the P3 was larger in response to unequal offers ($4.87 \mu\text{V}$) than to equal offers ($4.00 \mu\text{V}$). The main effect of Proposer was also significant ($F(1, 38) = 7.47, p = .009, \eta_p^2 = .16$); the P3 was larger when the proposer was believed to be human being ($4.86 \mu\text{V}$) and smaller when the proposer was known to be the computer ($4.00 \mu\text{V}$). The main effect of Anxiety was not significant ($F(1, 38) = 1.32, p = .257, \eta_p^2 = .03$). The interaction of Proposer \times Anxiety was significant ($F(1, 38) = 5.84, p = .021, \eta_p^2 = .13$); no P3 difference reached significance for the LTA group ($F(1, 38) = 0.05, p = .82, \eta_p^2 = 0.001$), but the P3 was larger in the human-proposer

condition (4.65 μV) than in the computer-proposer condition (3.22 μV) for the HTA group ($F(1, 38) = 13.25, p = .001, \eta_p^2 = .26$). Neither the Proposer \times Equality interaction ($F(1, 38) = 2.75, p = .105, \eta_p^2 = .07$) nor the Proposer \times Equality \times Anxiety interaction ($F(1, 38) = 0.57, p = 0.45, \eta_p^2 = .02$) was significant. No significant interaction between the experimental factors and Electrode was found (F values $< 2.25, p$ values $> .05$). When adding depression as a covariate in the ANOVA test, the Proposer \times Anxiety interaction remained significant ($F(1, 37) = 5.07, p = .03, \eta_p^2 = .12$).

Discussion

Our results replicated the well-established behavioral pattern that participants' acceptance rates decreased as levels of equality of UG offers decreased (Güth et al., 1982; Haselhuhn & Mellers, 2005). Also, the acceptance rates decreased in response to human-proposed offers compared to computer-proposed offers (see also Rilling & Sanfey, 2011; Sanfey et al., 2003). That is, both the equality factor and the social-interaction factor (i.e., the Proposer factor in the statistical analysis) manifested in the current study. Furthermore, the influence of anxiety on these two factors was evidenced on the behavioral performance and the ERPs, all of which remained significant after entering SDS scores as a covariate, thus ruling out the potential impact of depression on the current study.

The Impact of Anxiety on Social Rejection

Regarding the behavioral data, the interaction between the equal factor and the social-interaction factor were sensitive to anxiety levels, such that the tendency of rejecting human-proposed unequal offers compared to computer-proposed unequal offers appeared to be stronger in the HTA group than in the LTA group. The ERP results may help with understanding this phenomenon, since the impact of anxiety on the equality and social-interaction factors was detected on the FRN and the P3, respectively.

Anxiety and Electrophysiological Responses to Proposals

The FRN amplitude was larger in response to unequal offers compared with equal offers in the HTA group, while the equality factor was not significant in the LTA group. Seeing that the classical FRN component differentiates between "favorable" and "unfavorable" outcomes (Hajcak, Moser, Holroyd, & Simons, 2006; Nieuwenhuis, Slagter, von Geusau, Heslenfeld, & Holroyd, 2005), our FRN results indicate that high-anxious participants were more prone to evaluate inequitable outcomes as unfavorable than were their low-anxious counterparts. Interestingly, the FRN following negative outcomes was smaller among high-anxious participants than among low-anxious ones in single-player gambling tasks (Gu, Ge et al., 2010; Gu, Huang et al., 2010), indicating that the influence of anxiety on outcome evaluation show context-dependent patterns. High-anxious participants' smaller FRN amplitudes in single-player scenarios were interpreted as evidence of their lower confidence of earning reward when outcome valence is probabilistic; in other words, negative outcomes match high-anxious people's pessimistic outcome expectations and thus elicit weaker neural responses (Gu, Huang et al., 2010). In contrast, in multiple-player scenarios, high-anxious individuals may be more likely to expect equitable distributions from their partners and

consequently, perceive stronger prediction errors than their low-anxious counterparts when the offers were inequitable (see also T. Wu et al., 2013).

The P3 component, which occurs immediately after the FRN temporally, showed an interaction between anxiety and the social-interaction factor. Specifically, the human-proposed offers elicited a larger P3 than the computer-proposed offers in the HTA group, but not in the LTA group. The P3 amplitude has been suggested to represent the motivational or affective significance of a stimulus (Gu et al., 2011; Martin & Potts, 2004; Yeung & Sanfey, 2004). Increased P3 amplitudes among anxious people has previously been reported in the studies using the oddball paradigm (Boudarene & Timsit-Berthier, 1997; Enoch, White, Harris, Rohrbaugh, & Goldman, 2001; Li, Hu, Liu, & Wu, 2011), indicating stronger motivations to detect novel or affectively salient information (Paulus, Feinstein, Simmons, & Stein, 2004). In the current study, an enhanced P3 among high-anxious participants indicates that these participants judged UG offers that reflect human players' benevolent or selfish intentions to be more motivationally salient than the offers provided randomly by the computer. In contrast, low-anxious participants might be less concerned about the social context of an offer, indicated by an unchanged P3 between the human- and computer-proposer conditions for the group.

Further Directions and Concluding Remarks

Contrary to our behavioral results, Grecucci et al. (2012) reported that anxiety-disordered patients accepted significantly more unequal offers than the controls. To explain this discrepancy, it is important to note that the current study and that of Grecucci et al. (2012) recruited participants from a normal sample and from patients suffering from generalized anxiety disorder (GAD), respectively. Normal levels of anxiety and pathological anxiety are qualitatively distinct (Belzung & Griebel, 2001). It might be that high-anxious individuals in nonclinical populations and GAD patients are driven by different motivations during social decision-making. While high-anxious individuals tend to express their negative feelings towards unequal distributions, patients might be more worried about the potential interpersonal consequence of rejections, because GAD is strongly associated with low self-confidence and negative self-image (Butler, Cullington, Hibbert, Klimes, & Gelder, 1987; Masi et al., 2004). Further research that clarifies this unsolved issue would have significance to both clinical and non-clinical studies.

In sum, according to the ERP results, we suggest that high-anxious participants were more sensitive to the equality level of reward distribution than their low-anxious counterparts. Also, high-anxious participants were more sensitive to the underlying intention or motive of other people's actions. The interaction of these two factors on fairness consideration might explain high-anxious participants' altered socioeconomic behavior. That is, they were more likely to reject human-proposed unequal offers in the UG, despite this being regarded as a suboptimal strategy from the perspective of the "economic man" (Camerer & Fehr, 2006). In real-life, hypersensitivity to unfairness may contribute to high-anxious people's social avoidance, because unfair treatment is sometimes inevitable in social interactions. Inspired by the emotional regulation research, we suggest that training people who suffer from

anxiety symptoms to control their hyperactive responses would help improve their social performance (Hartley & Phelps, 2012).

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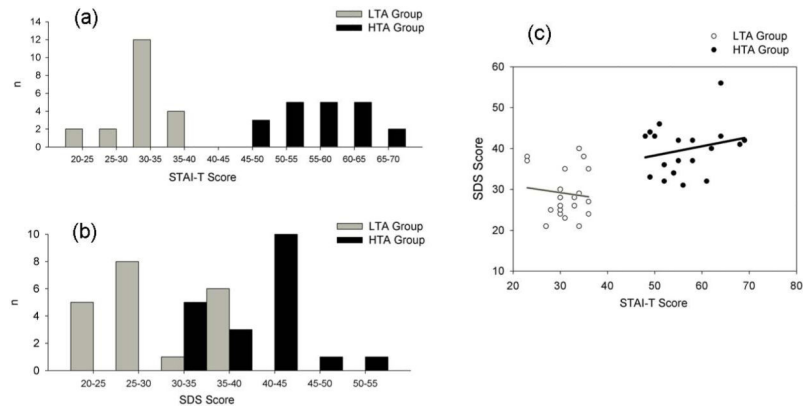


Figure 1. The distributions of the STAI-T (a) and SDS (b) scores, as well as their relationship (c), in the two groups. HTA: high-trait anxiety; LTA: low-trait anxiety.

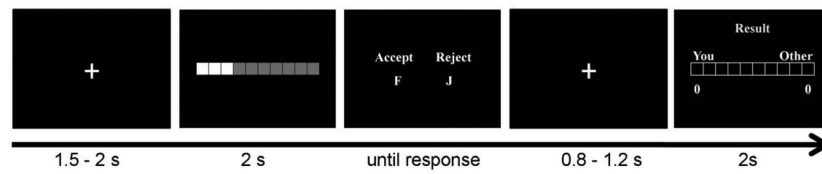


Figure 2.

Single-trial settings. A central fixation cross was presented for 1.5 - 2 s, which was followed by an offer (made by a pseudo-player or the computer) presented for a time window of 2 s. Immediately afterwards, participants had to decide whether to accept the proposal or not. After a short delay (0.8 - 1.2 s), a feedback screen informed participants about the results of the current trial.

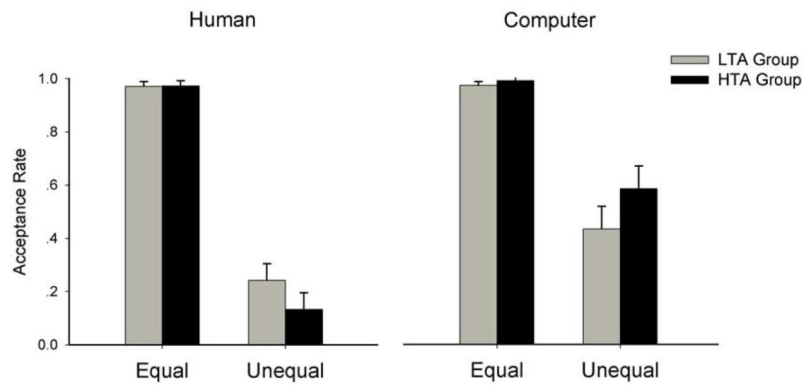


Figure 3. The illustration of the Proposer \times Equality \times Anxiety interaction on the acceptance rate. Error bars indicate 1 SE. HTA: high-trait anxiety; LTA: low-trait anxiety.

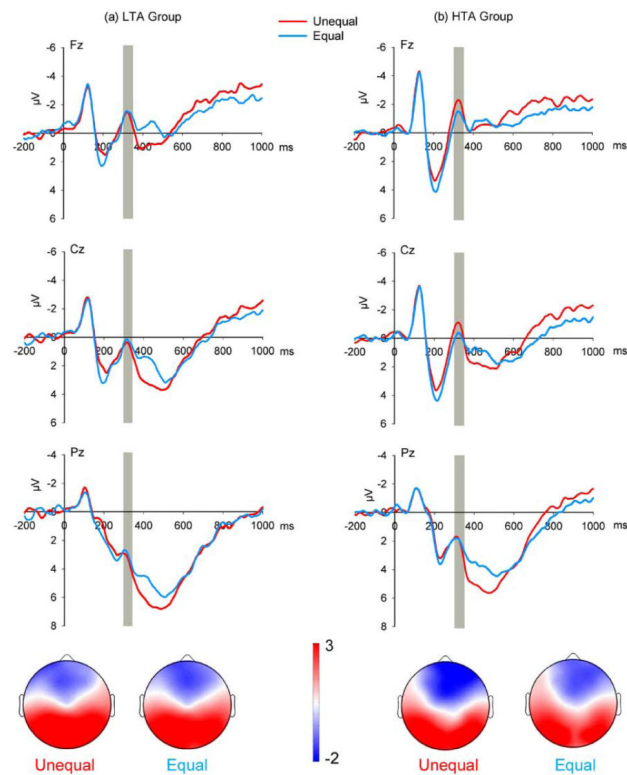


Figure 4.

The effect of the Equality factor on grand-average ERPs evoked by offers in the low-trait anxiety (LTA) group (a) and the high-trait anxiety (HTA) group (b) at mid-line recording sites Fz, Cz, and Pz. The time point “0” indicates offer presentation onset. The gray shaded areas indicate the 300–350 ms time window for the calculation of the mean value. The scalp topographies of each condition are presented beneath. Note: The waveforms in this figure are presented to provide a comparison of the midline scalp topographies for the FRN results. The electrodes used in the ANOVA were F1, Fz, F2, FC1, FCz, and FC2.

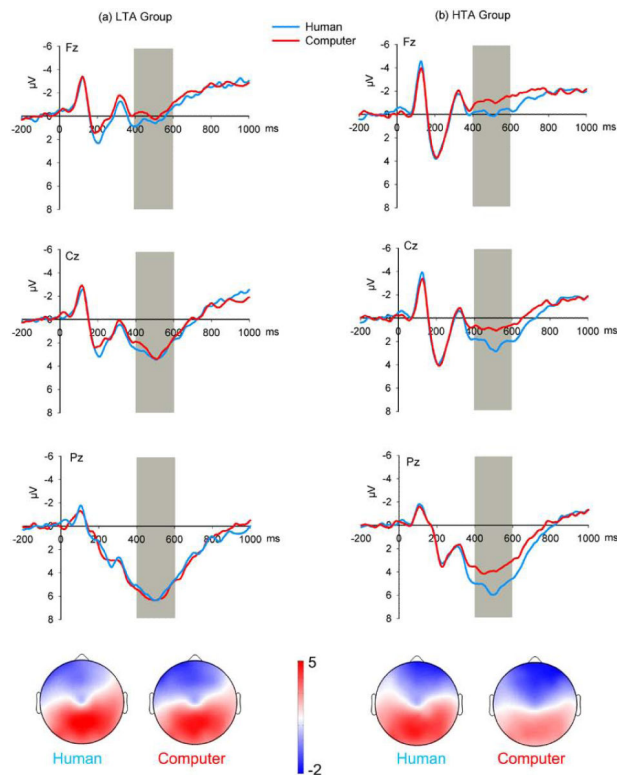


Figure 5.

The effect of the Proposer factor on grand-average ERPs evoked by offers in the low-trait anxiety (LTA) group (a) and the high-trait anxiety (HTA) group (b) at mid-line recording sites Fz, Cz, and POz. The time point “0” indicates offer presentation onset. The gray shaded areas indicate the 400–600 ms time window for the calculation of the mean value. The scalp topographies of each condition are presented beneath. Note: The waveforms are presented to provide the midline scalp topographies for the P3 results. The electrode used in the ANOVA were CPz, Pz, POz, and Oz.