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Hemoglobin A1c Assay Variations and Implications for Diabetes Screening in Obese Youth

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Abstract

Background—Standardization of the hemoglobin A1c (A1c) assay has led to its increasing utilization as a screening tool for the diagnosis of prediabetes and type 2 diabetes in youth. However, significant A1c assay variability remains and has implications for clinical management.

Objective—To describe our center's experiences with A1c results in youth and to evaluate intermethod differences and their clinical implications.

Subjects—75 youth (ages 10–18 years old), BMI 85th%ile participated.

Methods—72 participants had two A1c values performed on the same sample, one via immunoassay (DCA Vantage Analyzer, $A1c₁$) and the other via high performance liquid chromatography (Bio-Rad Variant II, $A1c_2$). 19 had A1c run on two immunoassay devices (A1c₁) and Dimensions Vista, $A1c_3$).

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Results—Mean age of participants was 13.9 years, BMI% 97.89, 33% male, 16% white, 21% black, and 61% Hispanic. Mean A1c₁ was 5.68% \pm 0.38 vs. a mean A1c₂ of 5.73% \pm 0.39, p=0.049. Concordance in diabetes status between methods was achieved in 79% of subjects. 19 subjects with A1c₃ results had testing performed an average of 22 \pm 9 days prior to A1c₁. Mean A1c₃ was 6.24% \pm 0.4, compared to a mean A1c₁ of 5.74% \pm 0.31, (p<0.0001). A1c₁ was on average systematically −0.5±0.28 lower compared to A1c₃. There was poor agreement in diabetes classification between $A1c_1$ and $A1c_3$, with a concordance in classification between methods of only 36.8%.

Conclusions—Clinically significant inter-method A1c variability exists that impacts patient classification and treatment recommendations. In the screening of obese youth for diabetes, A1c results should be interpreted with caution.

Keywords

Hemoglobin A1c; prediabetes; type 2 diabetes; obesity

Introduction

Standardization of hemoglobin A1c (A1c) methodologies by the National Glycohemoglobin Standardization Program (NGSP) to the Diabetes Control and Complications Trial (DCCT), which demonstrated direct relationships between A1c and diabetes outcomes, has promoted widespread use of A1c testing. In response, the American Diabetes Association (ADA) incorporated A1c into the diagnostic criteria for diabetes in 2010 (<5.7% normal, 5.7–6.4% prediabetes, 6.5% diabetes). (1) Despite lack of validated studies in pediatrics, these cut points have been extrapolated to youth, leading to increased A1c screening for diabetes by pediatricians (2, 3) and increased subspecialty referrals for abnormal A1c values. Our clinical experience suggested that abnormal A1c values obtained in outside hospitals were often normal when repeated at our institution. Our objective was to formally analyze differences between A1c results measured by multiple methodologies in a sample of overweight or obese adolescents.

Methods

Between March 2011 and December 2012, 75 overweight or obese participants were recruited from general pediatric clinics and referrals to the endocrine clinic at Children's Hospital Colorado for a larger ongoing trial at this center. Inclusion criteria were ages 10–18 years, BMI 85th%ile, and not on medications affecting glucose metabolism. A1c was obtained via immunoassay on a Siemens DCA Vantage Analyzer™ (Tarrytown, NY), A1c1, for all 75 participants. 72 (96%) participants also had an A1c performed on the same sample by high performance liquid chromatography (HPLC; Bio-Rad Variant II, Hercules, CA), A1c2. In addition, 19 (25%) participants also had A1c results obtained from the same outside hospital central lab operating a Siemens Dimension Vista® (Tarrytown, NY), A1 c_3 . All three A1c devices are NGSP certified and have documented traceability to the DCCT reference method. The laboratory reference range for the A1 c_1 DCA Vantage AnalyzerTM is 4.2–6.3%, with no distinction between normal and prediabetes. The reported reference ranges for the Bio-Rad Variant II, $A1c_2$, and the Siemens Vista, $A1c_3$, are identical to ADA

cutpoints for defining normal glycemia, prediabetes, and diabetes. $Alc₁$ and $Alc₃$ are immunoassay devices that may be utilized as point-of-care (POC) analyzers but, in this report, are operated by central laboratories at large tertiary care hospitals with rigorous quality control.

Statistical Analysis

Simple linear regression and Deming regression, which assumes measurement error in both X and Y, were used to explore the relationship between $Alc₁$ vs. $Alc₂$ and $Alc₁$ vs. A1c₃. Regression coefficients were reported as intercept \pm SE and $\beta \pm$ SE and the regression equation for the lines of best fit were also reported. Multiple linear regression was used to adjust for time differences between $A1c_1$ and $A1c_3$. Bland-Altman plots, in which the difference in paired values is plotted against the mean of the paired values, explored the bias between A1c₁ vs. A1c₂ and A1c₁ vs. A1c₃. p<0.05 was considered significant. Paired t-tests were used to compare A1c types. Cohen's kappa (k), a measure of inter-rater reliability used to compare two categorical methods of classification, was used to measure agreement in diabetes status. Fasting plasma glucose (FPG) and 2hour plasma glucose (2hr PG) after 75 g OGTT were available for these participants and concordance in diabetes classification between the different A1c assays with both FPG and 2hr PG were also calculated. Analyses were performed in SAS 9.3 (Cary, NC) and GraphPad Prism 5.0. The protocol was approved by the Colorado Multiple Institutional Review Board and appropriate consent and assent were obtained.

Results

Participants had a mean age of 13.9 years, mean BMI% 97.89, and were 33% male, 16% Non-Hispanic White (NHW), 21% African American (AA), and 61% Hispanic (H). Table 1 presents study population characteristics by gender and race/ethnicity. Of the $N = 75$ who had A1c₁, N = 72 had A1c₂ and N = 19 had A1c₃ results. Thus these patients are subsets of the $Alc₁$ data, however, there was no difference in distribution of sex or race/ethnicity for A1c₂ and A1c₃ compared to A1c₁.

Agreement between A1c1 and A1c²

Mean A1c₁ was $5.68\% \pm 0.38$ compared to a mean A1c₂ of $5.73\% \pm 0.39$, p=0.049. As expected, there was a significant relationship between A1c₁ and A1c₂ ($\mathbb{R}^2 = 0.67$, p < 0.0001). Linear regression as depicted in Figure 1a demonstrates that $A1c₂$ estimates were generally higher than $A1c_1$. In the Bland Altman plot in Figure 1b, $A1c_1$ was systematically 0.05 ± 0.23 lower, compared to A1c₂, p=0.049. There was reasonable agreement between A1c₁ and A1c₂ in diabetes status classification, ($k = 0.60$ (0.42–0.79)), with concordance in diabetes status achieved in 79% of subjects (Table 2a).

Agreement between A1c1 and A1c³

The 19 subjects with A1c₃ results had testing performed an average of 22 ± 9 days prior to A1c₁. Mean A1c₃ was 6.24% \pm 0.4, compared to a mean A1c₁, in the same 19 subjects, of 5.74% \pm 0.31, (p<0.0001). There was a significant relationship between A1c₁ and A1c₃, R² $= 0.50$ (p=0.0007) (Figure 2a) but the Bland Altman plot demonstrates that A1c₁ was

systematically lower than $A1c_3$ by $0.50 \pm 0.28\%$, (p<0.0001) (Figure 2b). There was little agreement in diabetes classification between $A1c₁$ and $A1c₃$, with a concordance of 36.8%, $(k = -0.44 (-0.69 - (-0.19))$ (Table 2b).

As African Americans (AA) have been reported to have a higher A1c for a given mean blood glucose, a subanalysis was performed removing AA subjects from the groups. When AA subjects were removed from the analysis, the A1c differences between methodologies were enhanced. With AA subjects removed from the comparison between $A1c_1$ and $A1c_2$, N $= 56$, kappa = 0.61 (0.41–0.82), and the A1c difference = -0.07 +/- 0.29, p = 0.03. For A1c₁ versus A1c₃, without AA subjects, N = 14 and the A1c difference = $0.61 +/-0.19$ %, p < 0.0001. Kappa = -0.64 , however the sample size was small.

To further explore the discrepancy in diabetes categorization between the subjects with both A1c₁ and A1c₂ and the subjects with both A1c₁ and A1c₃, we compared results to FPG and 2hr PG diabetes categories. 71 subjects with both $A1c₁$ and $A1c₂$ had FPG and 2hr PG data. For this group, the concordance between A1c1 and FPG = 48% and between A1c1 and 2hr $PG = 52\%$. For A1c2, the concordance rate with FPG = 44% and with 2hr PG = 58%. For the 19 subjects with both A1c1 and A1c3, A1c1 concordance rates with FPG = 42% and with 2hr PG 53%. Concordance in diabetes classification between A1c3 and FPG = 5% and between A1c3 and $2^hFG = 16%$. (Tables 3a-h)

Discussion

In this report, we demonstrate clinically important intra-individual A1c variability. When subjects had an A1c repeated on the same sample using two different NGSP certified methodologies, there was strong correlation between the two absolute values, but substantial discordance when the results were interpreted as categorical variables to determine diabetes classification. Even more importantly, testing performed on the same patient by the same assay methodology, but in different institutions, was systematically different and concordance in classification of diabetes status was poor. While lifestyle modifications between the two tests could explain the observed change in part, this explanation is unlikely to account for the consistency of the 0.5% lowering of A1c nor the degree of change noted within an average of only 3 weeks.

20 years ago when the DCCT and UKPDS first demonstrated a direct relationship between glycemic control and risk for diabetes complications, lack of comparability among A1c methods prevented implementation of A1c-based guidelines for diabetes screening and management. As a result of the efforts of the NGSP, standardization has greatly reduced variability among A1c methodologies leading to incorporation of A1c as an approved method for diabetes screening. However, it is important for providers utilizing A1c as a diagnostic tool to be aware of the clinically significant variability that remains, despite national standardization. Even prior to incorporation of A1c into the diagnostic guidelines, variability among DCCT aligned A1c assays was described (4–6). In 2012, allowed variability was +/−7% meaning that for a reference A1c of 6.5%, the allowed A1c for NGSP/DCCT certification ranged between 6.0% and 7.0%. These limits of acceptability are continuing to improve and in 2013 decreased to +/− 6%. However, while some degree of

A1c measurement variability is acceptable for monitoring of glycemic control on a single device in a patient with known diabetes, even the current standards of allowed variability become an important problem for screening and diagnosis when strict categorical ADA cutpoints are applied and the allowable certified range around a patient's value on differing devices runs from non-diabetic to diabetic.

The problem of inconsistent categorization of diabetes classification is not unique to A1c. Poor reproducibility of fasting plasma glucose (FPG) and 2 hour plasma glucose (2hrPG) have previously been demonstrated (7–11). In a study by Libman et al, only 30% of children with impaired glucose tolerance (IGT) on an initial oral glucose tolerance test (OGTT) had IGT during a second OGTT performed within 2 weeks. The percent positive agreement between the two OGTTs in this study for impaired fasting glucose (IFG) and IGT was low at 22.2 and 27.3% respectively. Another study on the prevalence of dysglycemia found that in a population of 2501 adolescents averaging 14.3 yrs of age, 175 had IFG on an initial screen, but only 11 (6%) of these participants had IFG upon repeat testing (11).

The validity of A1c as a screening or diagnostic tool in youth has also been questioned. Several studies have found poor sensitivity of A1c for detecting diabetes when compared to FPG and 2hrPG (12–14). In the data presented here, concordance rates in diabetes categorization between the different A1c assays with FPG and 2hrPG were generated. A1c³ demonstrated the lowest concordance with FPG and 2hrPG, although overall concordance rates of A1c assays to FPG and 2hrPG were poor. However, although OGTT is considered the gold-standard for diagnosing diabetes, we must bear in mind that OGTT, FPG, and A1c cut-points have all been extrapolated from adult populations to identify the A1c at which microvascular complications (largely retinopathy) are evident. None of these criteria have been validated in youth and the implications of these cut-points for diagnosing prediabetes and diabetes in youth are unclear. Given these unknowns, the argument has been made that the A1c may be no less valid than the OGTT in children (15, 16). Regardless of the controversy over the validity of A1c as a diagnostic tool in youth, A1c has become increasingly utilized by general practitioners as a screening tool over the OGTT due to the theoretical advantages of the A1c test as an index of chronic glycemia, lack of dependence on fasting, and greater convenience of sampling (2, 3), and further studies to understand the implications of A1c and OGTT in predicting glucose abnormalities in children are necessary.

It should be noted that the ADA has approved only traditional central lab methodologies, and not POC devices, for the diagnosis of diabetes. However, the list of NGSP certified A1c devices (17) includes numerous POC assays that have been shown to demonstrate acceptable precision in relation to DCCT standards when operated and maintained by a reliable lab. Concerns arise over the precision of POC devices because they are not required to participate in ongoing proficiency testing. Additionally, POC devices with a coefficient of variation (CV) 3% may have clinically significant analytic errors that are not reliably detected by clinical laboratories despite routine calibration (4). With an analytical CV of 3%, for example, a measured A1c of 6% may have an uncertainty of measurement ranging from 5.6–6.4% (18) and this may lead devices to increasingly underestimate or overestimate A1c relative to NGSP references, leading to large inter-method differences over time (4, 5,

19). On the other hand, POC methods have advantages over traditional central lab techniques including a faster turnaround time, ease of operation, and ability to handle highvolume testing. With improved standardization, many central labs have abandoned traditional high-performance liquid chromatography techniques in favor of these more rapid POC assays. Thus, given the increasing adoption of POC assays by large laboratories, the methodology can no longer simply be determined based on location of test performance (back-office versus central lab) and is usually not obvious to the general practitioner ordering the A1c screening test.

There are limitations to the results presented here. The population with $A1c₃$ was a convenience sample of obese youth with available results for comparison, rather than a predefined group of subjects with A1c run by different devices on the same sample. We thus limited the analysis to individuals with $A1c₃$ values from the same hospital laboratory obtained within 1 month of study date. In addition, A1c represents average glucose levels in the preceding 3 months due to red cell lifespans of 60–120 days. However, A1c is weighted towards the most recent 3–4 weeks (20), which may have influenced the assay differences noted. Participants who were made aware of a recent abnormal A1c may also have instituted lifestyle changes that could have lowered A1c within this short time frame. Nevertheless, the A1c differences described here are consistent with results that would be expected based on the known variances of these devices.

Additionally, despite the small sample size, the statistically significant discrepancy reported here shows that further studies are needed to understand the implications of A1c.

In summary, there is persistent variability in current A1c methodologies that can lead to potential clinical pitfalls related to interpretation of a continuous variable in a categorical manner. Since a diagnosis of diabetes is dependent on A1c results that change category with a difference of 0.1%, the noted variability means that individuals with A1cs in the upper normal to mildly elevated range will be inconsistently categorized when testing is performed on different devices. Discrepant categorization leads to inconsistent diagnoses, which may confuse providers and families, lead to under- or over-treatment, and result in patient and parental anxiety and psychological stress. There are also implications of assay variability for the recruitment of patients into pediatric diabetes clinical trials, where A1c inclusion and exclusion criteria affect selection of potential participants to be screened. Finally, the intraassay variation identified here also raises caveats regarding the use of A1c, not only in screening, but for diabetes management, where caution should be exercised when comparing A1c 'trends' between different assays or devices.

We face an epidemic of obesity with rising rates of type 2 diabetes in youth. However, given the low prevalence of type 2 diabetes, even among high risk youth, reliable, efficient and cost-effective approaches to screening adolescents and detecting those at highest risk for progression to disease are needed. Measurement of A1c offers a number of potential advantages as a screening approach. However, variation in methodologies leads to inconsistent classification of and possible under- or over-treatment depending on the assay used or institution at which the assay is run. This variability in A1c results underscores the need for repeated measurements over time, consideration of a combination of screening

methods such as A1c and FPG to reduce the bias inherent in a single test, and the employment of clinical judgment to determine the appropriate course of management when screening each patient. Accordingly, the ADA recommendations require two repeated measures of A1c or plasma glucose meeting diabetes criteria or signs and symptoms of diabetes prior to making a diagnosis of diabetes. Until we have a clearer understanding of the A1c cut points and their implications in pediatrics, and until the precision of A1c assays is improved, A1c values obtained for screening purposes in obese youth should be interpreted with caution.

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Abbreviations

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 \mathbf{a}

Figure 1.

 a – Correlation between $A1c_1$ and $A1c_2$

b – Bland-Altman plot of the difference against the mean of A1c₁ and A1c₂ results

 a – Correlation bewteen $A1c_1$ and $A1c_3$

b – Bland-Altman plot of the difference against the mean of A1c₁ and A1c₃ results

Table 1

Characteristics of Study Population

Table 2

Table 3

There were no observations for FPG in the Diabetes category so Cohen's Kappa could not be calculated. The concordance rate was 48%

Cohen's Kappa = 0.12 (−0.08–0.31) which means there is no agreement. The Concordance rate was 52%.

There were no observations for FPG in the Diabetes category so Cohen's Kappa could not be calculated. The concordance rate was 44%

Cohen's Kappa = 0.20 (0.005–0.40) which means there is no agreement. The Concordance rate was 58%.

There were no observations for FPG in the Diabetes category so Cohen's Kappa could not be calculated. The Concordance rate was 42%.

There were no observations for 2hr PG in the Diabetes category so Cohen's Kappa could not be calculated. The Concordance rate was 53%

There were no observations for FPG in the Diabetes category so Cohen's Kappa could not be calculated. The Concordance rate was 5%

There were no observations for 2hr PG in the Diabetes category so Cohen's Kappa could not be calculated. The Concordance rate was 16%