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### **Gene-environment interaction involving recently identified colorectal cancer susceptibility loci**

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#### **Abstract**

**BACKGROUND—**Genome-wide association studies have identified several single nucleotide polymorphisms (SNPs) that are associated with risk of colorectal cancer (CRC). Prior research has evaluated the presence of gene-environment interaction involving the first 10 identified susceptibility loci, but little work has been conducted on interaction involving SNPs at recently identified susceptibility loci, including: rs10911251, rs6691170, rs6687758, rs11903757, rs10936599, rs647161, rs1321311, rs719725, rs1665650, rs3824999, rs7136702, rs11169552, rs59336, rs3217810, rs4925386, and rs2423279.

**METHODS—**Data on 9160 cases and 9280 controls from the Genetics and Epidemiology of Colorectal Cancer Consortium (GECCO) and Colon Cancer Family Registry (CCFR) were used to evaluate the presence of interaction involving the above-listed SNPs and sex, body mass index (BMI), alcohol consumption, smoking, aspirin use, post-menopausal hormone (PMH) use, as well as intake of dietary calcium, dietary fiber, dietary folate, red meat, processed meat, fruit, and vegetables. Interaction was evaluated using a fixed-effects meta-analysis of an efficient Empirical Bayes estimator, and permutation was used to account for multiple comparisons.

**RESULTS—**None of the permutation-adjusted p-values reached statistical significance.

**CONCLUSIONS—**The associations between recently identified genetic susceptibility loci and CRC are not strongly modified by sex, BMI, alcohol, smoking, aspirin, PMH use, and various dietary factors.

**IMPACT—**Results suggest no evidence of strong gene-environment interactions involving the recently identified 16 susceptibility loci for CRC taken one at a time.

#### **Keywords**

Colorectal Cancer; Gene-Environment Interaction; Polymorphism; Single Nucleotide; Genetic Predisposition to Disease; Diet

#### **INTRODUCTION**

Colorectal cancer (CRC) is the third most common cancer among men and women in the United States [1]. To date, genome-wide association studies (GWAS) have identified a number of single nucleotide polymorphisms (SNPs) that are associated with risk of this cancer [2–14]. There is much interest in identifying whether demographic and lifestyle factors modify the association between genetic variants and CRC, as finding evidence of gene-environment (GxE) interaction may help guide future prevention strategies. Furthermore, understanding GxE interaction may shed light on the mechanisms by which genetic polymorphisms affect risk of CRC, as well as the underlying biology of this disease. The SNPs identified to be associated with CRC thus far only account for a small fraction of the estimated heritability of CRC [15,16], and it has been suggested that one factor contributing to this 'missing heritability' is gene-environment (GxE) interaction [17,18].

We previously reported on gene-environment interaction for the first 10 identified susceptibility loci [19]. Since the time of that publication, 16 additional SNPs have been associated with CRC, including: rs10911251 (1q25.3), rs6691170 (1q41), rs6687758 (1q41), rs11903757 (2q32.3), rs10936599 (3q26.2), rs647161 (5q31.1), rs1321311 (6p21), rs719725 (9p24), rs1665650 (10q26.12), rs3824999 (11q13.4), rs7136702 (12q13.13), rs11169552 (12q13.13), rs59336 (12q24.21), rs3217810 (12p13.32), rs4925386 (20q13.33), rs2423279 (20p12.3) [3,4,7,8,10,14]. Few studies have evaluated the presence of interaction involving these recently identified susceptibility loci [8,20–24]. Although it has been suggested that sex may interact with rs4925386 [22], no interaction has been observed between sex and rs719725 [8, 21,24], rs6691170 [22], rs10936599 [22], or rs11169552 [22]. Of the newly identified susceptibility loci, only rs719725 [8, 21,23] and SNPs highly correlated with rs719725 [20] have been evaluated for interaction with environmental factors such as body mass index (BMI), alcohol consumption, smoking, medication use, and diet. No statistically significant GxE interactions were observed in these studies; however statistical power to detect interaction may have been limited due to insufficient sample sizes. We have therefore evaluated whether environmental risk factors for CRC modify the associations between these genetic polymorphisms and CRC risk using data on 9160 cases and 9280 controls in the Genetics and Epidemiology of Colorectal Cancer Consortium (GECCO) and the Colon Cancer Family Registry (CCFR). The following environmental and demographic factors were included in our study: sex, BMI, alcohol use, smoking, aspirin use, post-menopausal hormone (PMH) use, dietary intake of calcium, fiber, folate, red meat, processed meat, fruit, and vegetables. These 'environmental factors' have been loosely defined so as to include lifestyle factors and personal characteristics associated with CRC risk [25–35].

#### **MATERIALS AND METHODS**

#### **Study participants**

Study participants were drawn from either case-control studies (Ontario Familial Colorectal Cancer Registry [OFCCR], Darmkrebs: Chancen der Verhuetung durch Screening [DACHS], Diet, Activity and Lifestyle Survey [DALS], Colon Cancer Family Registry [CCFR], Colorectal Cancer Studies 2&3 [Colo2&3], and the Postmenopausal Hormone study within the Colon Cancer Family Registry [PMH-CCFR]) or from case-control studies

nested within prospective cohorts: Health Professionals Follow-up Study [HPFS], Nurses' Health Study [NHS], Physicians' Health Study [PHS], Prostate, Lung, Colorectal, and Ovarian Cancer Screening Trial [PLCO], Women's Health Initiative [WHI], Multiethnic Cohort Study [MEC], and the VITamins And Lifestyle [VITAL] study. More detailed information on these studies can be found in Table 1 and in the Supplemental Methods. All participants gave informed consent and studies were approved by their respective Institutional Review Boards.

#### **Outcome**

Colorectal cancer (CRC) cases included in this study were defined as invasive colorectal adenocarcinoma (ICD codes 153–154). Cases were confirmed by medical record, pathology report, or death certificate. Controls in these case-control studies and nested case-control studies were selected based on study-specific eligibility and matching criteria, as detailed in the Supplemental Methods.

#### **Genotype Data**

Gene-environmental interaction was evaluated for 16 SNPs located at recently identified CRC susceptibility loci, including: rs10911251 (1q25.3), rs6691170 (1q41), rs6687758 (1q41), rs11903757 (2q32.3), rs10936599 (3q26.2), rs647161 (5q31.1), rs1321311 (6p21), rs719725 (9p24), rs1665650 (10q26.12), rs3824999 (11q13.4), rs7136702 (12q13.13), rs11169552 (12q13.13), rs59336 (12q24.21), rs3217810 (12p13.32), rs4925386 (20q13.33), rs2423279 (20p12.3) [3,4,7,8,10,14].

DNA for genotyping was largely obtained from blood samples, though DNA was also obtained from buccal swabs for VITAL participants and for a subset of participants from DACHS, MEC, and PLCO. Genotyping was conducted on several different platforms and several of the studies were genotyped in sets, Therefore, in describing the genotyping platform and in presenting data on genotyping quality in Supplemental Table 1, results are presented by study set. However, we have presented results in tables and figures by overall study population. The Illumina HumanHap BeadChip Array System was used to genotype SNPs for the following studies: Colo2&3, DACHS1, DALS2, MEC, PLCO2, PMH-CCFR, VITAL, WHI2 (300k); DALS1, WHI1 (550k); WHI1 (550kduo); DALS1, WHI1 (610k); DACHS2, HPFS1, HPFS2, NHS1, NHS2, PHS1+2 (730k), as described previously [9]; OFCCR samples were genotyped using Affymetrix platforms [14]. All genotyping underwent quality-control checks, including concordance checks for blinded and unblinded duplicates, as well as examination of sample call rates, SNP call rates, and, in controls, Hardy-Weinberg Equilibrium (HWE). Samples with gender discrepancies were excluded, as were persons who reported a racial/ethnic group other than "white;" European ancestry was confirmed in GWAS samples using principal components analysis.

As not all of the SNPs of interest were genotyped on each platform, we imputed SNPs to the CEPH collection (CEU) population in HapMap II. Imputation was used only if a minor allele frequency (MAF) of >1% could be assumed and satisfactory overall imputation accuracy ( $\mathbb{R}^2$  >0.3) was achieved. Imputation quality was high for all SNPs of interest (average R<sup>2</sup>>0.85), except rs3217810 (average R<sup>2</sup>=0.49) and rs11903757 (average R<sup>2</sup>=0.69)

(Table 2). For each SNP included in our analyses, the number of studies in which that SNP was imputed or genotyped is provided in Table 2. All SNPs are presented in terms of number of risk alleles, with 0 corresponding to no risk alleles, and 2 corresponding to 2 risk alleles. Directly genotyped SNPs are coded as 0, 1, or 2 risk alleles, and imputed SNPs are instead coded in terms of the expected number of risk alleles ("dosage" between 0 and 2) [36]. The risk allele designation for each SNP was determined by the discovery studies, as presented in Table 2. The SNP details by study, including the risk allele frequency (RAF), imputation  $\mathbb{R}^2$ , and HWE among controls are provided in Supplemental Table 1.

#### **Environmental data and harmonization procedure**

Environmental and demographic exposures evaluated for GxE interaction include: sex, BMI, alcohol consumption, smoking, aspirin use, PMH use, dietary intake of calcium, fiber, folate, red meat, processed meat, fruit, and vegetables [25–35].

Data on environmental exposures were self-reported at either in-person interview or in structured self-administered questionnaires. As data collection instruments differed across studies, a multi-step, iterative data harmonization procedure was used. After the common data elements (CDEs) were identified, the questionnaires and data dictionaries of each study were examined to identify specific elements that could be mapped to these CDEs. These data elements were then written to a common data platform and then transformed via an SQL programming script, allowing these variables to be combined into a single dataset with common definitions, standardized coding, and standardized permissible values. This mapping procedure and resulting values were reviewed for quality assurance, with range and logic checks performed to assess data distributions within and between studies. After examining the data, outlying samples were truncated to the minimum or maximum value of the established range for each variable.

The harmonized alcohol variable was categorized as follows:  $\langle 1 \text{ g/day}, 1 - \langle 28 \text{ grams/day}, 0 \rangle$ 28+ grams/day. BMI was modeled as a scaled variable (BMI [kg/m<sup>2</sup>]/10), with underweight persons (BMI<18.5) excluded in analyses of BMI to avoid concern that underweight persons may have had occult disease at the time of exposure assessment.

Smoking was defined in two ways, a binary never/ever variable and a 5-level pack-year variable (never smoking, 4 study-specific quartiles of pack-years smoked). Aspirin use was defined as a binary variable, with yes indicating regular use of aspirin at the time of reference (with study-specific definitions varying across studies); similarly, PMH use was defined as a binary variable, with yes indicating any current use of PMH at the time of reference, and analyses of PMH use were limited to women.

All dietary variables (dietary calcium intake, dietary fiber intake, dietary folate intake, red meat consumption, processed meat consumption, vegetable consumption, fruit consumption) were categorized into quartiles. Calcium, fiber, and folate were limited to dietary intake. These quartiles were sex- and study-specific, with the coding of the quartiles corresponding to the median value of the quartile within each sex and study. After combining data across studies, we then scaled these variables to a unit reflective of the distribution of each dietary variable; the scaled units are as follows: calcium (500 mg/day), fiber (10 g/day), folate (500

mcg/day), processed meat (servings/day), red meat (servings/day), vegetable (5 servings/ day), and fruit (5 servings/day). As some of the studies included in our meta-analysis collected information in categories that did not allow for conversion to these quartiles, we have also examined consumption of processed meat, red meat, vegetable, and fruit as lessrich (but more inclusive) binary variables, with the threshold between low and high consumption defined by sex-and study-specific medians. HPFS and NHS were excluded from analyses of fiber and the 4-level processed meat variable, as comparable data for these variables were not available at the time of study initiation. DACHS was excluded from analyses pertaining to the 4-level fruit and vegetable variables due to substantial differences in how these variables were assessed and defined. For all environmental exposures, the referent group corresponds to the lowest level of exposure.

#### **Statistical analysis**

Analyses of main effects of SNPs and environmental factors and GxE interaction were adjusted for age, sex, and study center. Analyses involving genetic data were further adjusted for population substructure (first 3 principal components using EIGENSTRAT [37]); analyses corresponding to the following dietary variables were further adjusted for energy intake if available: calcium, fiber, folate, fruit consumption, and vegetable consumption. Analyses of the Physicians' Health Study were further adjusted for smoking, as participants were matched on smoking status.

To assess the best model fit for each SNP, we compared an unrestricted model to logadditive, dominant, and recessive models using a likelihood ratio test [19]. All SNPs were best modeled using a log-additive model, except for rs59336; this SNP was modeled dominantly, given that the unrestricted model outperformed both the additive and recessive models.

The model form of environmental variables was also assessed. The best model form for the alcohol variable and 4-level dietary variables was assessed using a likelihood ratio test to compare a model with unrestricted categorical variables to a reduced model with a single linear variable. The likelihood ratio test indicated that modeling alcohol categorically significantly outperformed the linear alcohol variable; therefore, alcohol was modeled using unrestricted categorical variables. However, all of the 4-level dietary variables (fruit consumption, vegetable consumption, red meat consumption, processed meat consumption, fiber intake, folate intake, and calcium intake) were modeled as single linear variables, given that the unrestricted categorical variable did not outperform the linear variable. To assess the best model form for BMI ( $\left[\frac{kg}{m^2}\right]$ /10) and pack-years smoked (5-level variable), we used a likelihood ratio test to compare a model with and without a quadratic term; the addition of the quadratic term did not improve the model fit for either of these variables, and therefore both BMI ( $\left[\frac{\text{kg}}{m^2}\right]$ /10) and smoking (5-level variable) were modeled linearly.

To test for interaction, an efficient Empirical Bayes (EB) shrinkage method was used, which is a weighted sum of the case-only test and the traditional case-control method [38]. In the event that the assumption of gene-environment independence appears to hold, more weight is given to the more powerful case-only method; if this assumption is violated, more weight is given to the case-control estimate, which does not assume gene-environment

independence. This approach affords the greater power of the case-only analysis, while protecting against bias in the event of gene-environment dependence. All results for metaanalyses were obtained using a fixed-effects model, and for each meta-analysis performed, we examined the corresponding p-value for heterogeneity across studies (Supplemental Table 2).

Given that 288 tests were performed (16 SNPs\*18 environmental factors) and some of the environmental variables were correlated with one another, permutation was used to account for multiple testing and correlations among variables. Each analysis was performed 2000 times using a permuted case-control status in each run, after which the Westfall and Young stepdown procedure was applied to derive an adjusted p-value for each interaction [39]. These adjusted p-values were then used to assess the presence of interaction at the alpha=0.05 level. All other p-values are termed nominal p-values.

Data harmonization was performed in SAS and T-SQL, while all other analyses were performed in R.

#### **RESULTS**

Our study population included a total of 18,440 persons, including 9160 cases and 9280 controls. Of the 18,440 persons included, 8110 (44.0%) were male and 10,330 (56.0%) were female.

The marginal associations of the SNPs with CRC risk are presented in Table 2. In this consortium of studies, of the 16 SNPs studied, 12 showed evidence of association with CRC risk as initially discovered, with p-values <0.05. Though not statistically significant, three of the remaining SNPs (rs6687758, rs6691170, and rs10936599) showed evidence of association in the expected direction [4]. However, for one SNP, rs1665650, the significant risk allele in our study (C) did not match the risk allele as it was discovered (T)[7]. One SNP, rs59336, showed evidence of heterogeneity across studies in its marginal association with CRC (p-het:  $1.5\times10^{-3}$ ).

The marginal associations between the environmental factors and CRC are presented in Table 3. Increasing folate intake, NSAID use, PMH use, low alcohol intake, and increasing consumption of calcium, vegetable, fruit, and fiber were associated with reduced risk of CRC, whereas high alcohol consumption, increasing red and processed meat consumption, smoking, and high BMI were associated with increased CRC risk. The main effect of sex is not presented due to matching on this variable. As can be seen in Supplemental Table 3, the main effects of the environmental variables tend to be stronger in case-control studies than in cohort studies.

The results for the 288 gene-environment interactions tested are presented in Supplemental Table 2. In analyses adjusted for age, sex, study center, and population substructure (principal components), six interactions had a nominal p-value <0.01: rs6691170\*PMH use (no/yes), rs3217810\*dietary fiber intake (per 10 g/day), rs3217810\*dietary folate intake (per 500 mcg/day), rs7137602\*vegetable consumption (per 5 servings/day), rs10936599\*sex, and rs719725\*fruit consumption (high vs low) (Table 4). The strongest interaction was

between rs6691170 and PMH, with an interaction odds ratio (OR) of 1.22 (95% CI: 1.08– 1.39), and a nominal p-value of  $1.74 \times 10^{-3}$  (p-value heterogeneity=0.18; results presented in Table 4). After accounting for multiple comparisons, the adjusted p-value for the PMHrs6691170 interaction did not reach statistical significance (adjusted p-value=0.30) (Table 4). No other interactions were statistically significant after accounting for multiple comparisons.

#### **DISCUSSION**

In our meta-analysis of 9160 CRC cases and 9280 controls, after adjustment for multiple comparisons, we found no statistical evidence to support that the associations between recently identified susceptibility loci and CRC are modified by environmental factors, including sex, BMI, smoking, alcohol, aspirin use, PMH use, and various dietary factors.

We confirmed expected associations between CRC and environmental factors studied, as well as between CRC and 12 of the recently identified SNPs. Four variants did not replicate in this study population, including SNPs located at 1q41 (rs6687758, rs6691170), 3q26.2 (rs10936599), and 10q26.2 (rs1665650); nonetheless, the direction of association for three of these SNPs (rs6687758, rs6691170, and rs10936599) was the same in our study as prior studies [4,22,40]. However, the risk allele for rs1665650 in our study did not match the one reported [7]. This may be due to differences in the underlying linkage patterns given the ethnic differences in populations studied (the discovery study by Jia et al. was conducted among Asian populations, whereas our study included only persons of European descent). However, it remains unclear why rs6687758, rs6691170, and rs10936599 did not replicate in GECCO. It may be that the distribution of environmental factors in our population differs from that of the populations in which these genetic variants were discovered, though, as noted, none of the environmental factors studied here interacted with these genetic variants.

None of the interactions studied was statistically significant after adjustment for multiple comparisons. This may be because there is truly no interaction between these genetic and environmental factors or it may be that power is still limited to detect modest or weak interactions despite our large sample size. In our analyses of 9160 cases and 9280 controls, we are adequately powered to detect interactions with an interaction OR in the range of 1.21 to 1.29 for MAF in the observed range (0.16–0.49), assuming a main effect of 1.08 for logadditive SNPs, a main effect of 1.22 for binary environmental risk factors, and an alpha of  $1.74 \times 10^{-4}$  (Bonferonni p-value of  $1.74 \times 10^{-4}$ = [0.05/288]). However, as analyses of PMH use were limited to women (4284 cases, 4695 controls), we were underpowered to detect an OR in this range and therefore a larger sample size may be needed to more thoroughly evaluate interactions between PMH use and recently identified susceptibility loci. This evidence builds upon a prior analysis conducted within GECCO in which we examined the presence of gene-environment interaction for the first 10 identified susceptibility loci [19]. In that paper, we observed only one statistically significant gene-environment interaction, between rs16892766 (8q23.3) and vegetable consumption. Taken together, there is very little evidence for gene-environment interaction involving known susceptibility loci within GECCO, though a larger sample size may be needed to evaluate interaction. Our data suggests that GxE with known susceptibility loci may not account for the missing

investigating the presence of genome-wide GxE interaction with a variety of environmental factors. It may also be informative to evaluate GxE by anatomic subsite or by molecular characteristics, such as microsatellite instability; however, an even larger sample size would be needed for such analyses.

One of the major strengths of this study is the large sample size. This is especially important, as this is the largest study examining GxE involving these SNPs and prior studies have cited the need for a larger sample size when evaluating gene-environment interaction [20,23]. Furthermore, we used an Empirical Bayes approach so as to derive additional power from the use of case-only analyses [38]. Another advantage is that we used a standardized harmonization procedure to combine environmental data across studies.

Nonetheless, a limitation of this study is measurement error. As measurement error can bias estimates of interaction in GxE analyses [41,42], we evaluated the best model form for environmental and genetic factors to minimize the measurement error present in our variables. Regardless, harmonizing data across studies necessarily yields simpler variables, potentially leading to some loss of information in our environmental data and attenuation of effect estimates. For example, our PMH use variable is limited to a binary variable and does not incorporate information on other potentially important characteristics of use.

Furthermore, our consortium includes both retrospective and prospective studies, and these types of studies have different sources of error. The main exposure effects varied somewhat by study design (Supplemental Table 3), likely due to differential measurement error and/or selection bias in case-controls studies or the variable time period between baseline questionnaire and cancer diagnosis in the prospective studies (the average time between baseline and cancer diagnosis ranged from approximately 3–11 years across prospective studies). However, gene-exposure interactions are not subject to selection bias under the assumption that genotype does not influence participation (conditional on exposure and disease status) [43]. Despite these concerns, the associations between all environmental variables and CRC were in the expected directions. Indeed, it is notable that the environmental variables show relationships almost entirely consistent with the large body of earlier epidemiologic work [25–35]. Even so, this loss of richness of environmental data is a limitation common to consortia-based studies; however, it is this harmonization of environmental data which allows for the sample size needed to evaluate GxE.

Finally, we examined GWAS-identified SNPs and therefore our analyses do not include all genetic polymorphisms associated with CRC risk. These GWAS-identified SNPs are unlikely to be the underlying functional (i.e., disease-causing) variant; instead, they tag correlated variants that may have functional importance in CRC development. If these causal SNPs are not well tagged, a study that directly genotypes these causal SNPs would yield stronger associations [44,45] and improve power to detect GxE interactions.

In conclusion, our study suggests that the associations between recently identified CRC susceptibility loci and CRC are not strongly modified by known environmental factors. Our findings, along with those of our prior GxE paper [19] suggest that there may be limited gene-environment interaction involving the first 26 identified susceptibility loci and common CRC risk factors. However, large studies incorporating richer harmonized environmental data and causal SNPs may be needed to uncover the presence of weak to moderate gene-environment interaction. Further work is needed to evaluate the presence of genome-wide GxE interaction involving rare variants and multifactorial interaction.

#### **Supplementary Material**

Refer to Web version on PubMed Central for supplementary material.

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A subset of control samples were genotyped as part of the Cancer Genetic Markers of Susceptibility (CGEMS) Prostate Cancer GWAS [46], Colon CGEMS pancreatic cancer scan (PanScan) [47, 48], and the Lung Cancer and Smoking study. The prostate and PanScan study datasets were accessed with appropriate approval through the dbGaP online resource ([http://cgems.cancer.gov/data/\)](http://cgems.cancer.gov/data/) accession numbers phs000207v.1p1 and phs000206.v3.p2, respectively, and the lung datasets were accessed from the dbGaP website [\(http://www.ncbi.nlm.nih.gov/gap\)](http://www.ncbi.nlm.nih.gov/gap) through accession number phs000093 v2.p2. For the lung study, the GENEVA Coordinating Center provided assistance with genotype cleaning and general study coordination, and the Johns Hopkins University Center for Inherited Disease Research conducted genotyping.

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**Table 1**

General characteristics of included studies General characteristics of included studies



**Table 2**

Associations between recently identified single nucleotide polymorphisms and colorectal cancer in the Genetics and Epidemiology of Colorectal Cancer Associations between recently identified single nucleotide polymorphisms and colorectal cancer in the Genetics and Epidemiology of Colorectal Cancer Consortium and Colon Cancer Family Registry Consortium and Colon Cancer Family Registry



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 ${}^{2}$  All SNPs modeled additively, with the exception of rs59336, which was modeled dominantly All SNPs modeled additively, with the exception of rs59336, which was modeled dominantly <sup>b</sup>SNPs identified to be associated with colorectal cancer risk in the following studies: rs10911251 (Peters et al. Gasteroenterology, 2013 [10]); rs6687758 (Houlston et al. Nat Genet, 2010 [4]); rs6691170<br>(Houlston et al. SNPs identified to be associated with colorectal cancer risk in the following studies: rs10911251 (Peters et al. Gasteroenterology, 2013 [10]); rs6687758 (Houlston et al. Nat Genet, 2010 [4]); rs6691170 (Houlston et al. Nat Genet, 2010 [4]); rs11903757 (Peters et al. Gasteroenterology, 2013 [10]); rs10936599 (Houlston et al. Nat Genet, 2010 [4]); rs647161 (Jia et al. Nat Genet, 2013 [7]); rs1321311

(Dunlop et al. Nat Genet, 2012 [3]); rs719725 (Zanke et a., Nat Genet, 2007 [14]; Kocarnik et al. Cancer Epidemiol Biomarkers Prev, 2010 [8]); rs1665650 (Jia et al. Nat Genet, 2013 [2013]); rs3824999 (Dunlop et al. Nat Genet, 2012 [3]); rs719725 (Zanke et a., Nat Genet, 2007 [14]; Kocamik et al. Cancer Epidemiol Biomarkers Prev, 2010 [8]); rs1665650 (Jia et al. Nat Genet, 2013 [2013]); rs3824999 (Dunlop et al. Nat Genet, 2012 [3]); rs3217810 (Peters et al. Gasteroenterology, 2013 [10]); rs7136702 (Houlston et al. Nat Genet, 2010 [4]); rs11169552 (Houlston et al. Nat Genet, 2010 [4]); rs5936 (Dunlop et al. Nat Genet, 2012 [3]); rs3217810 (Peters et al. Gasteroenterology, 2013 [10]); rs7136702 (Houlst, 2010 [4]); rs11169552 (Houlston et al. Nat Genet, 2010 [4]); rs59336 (Peters et al. Gasteroenterology, 2013 [10]); rs2423279 (Jia et al. Nat Genet, 2013 [7]); rs4925386 (Houlston et al. Nat Genet, 2010 [4]) (Peters et al. Gasteroenterology, 2013 [10]); rs2423279 (Jia et al. Nat Genet, 2013 [7]); rs4925386 (Houlston et al. Nat Genet, 2010 [4])

 $\mathbf{\hat{c}_{Risk}}$  base allele designation based on the literature Risk/base allele designation based on the literature

 $d$  dijusted for age, sex, study center, and population substructure (principal components  $1\hbox{--}3)$ Adjusted for age, sex, study center, and population substructure (principal components 1–3)

#### **Table 3**

Association between environmental factors and CRC in GECCO



ABBREVIATION: PMH (post-menopausal hormone); OR (odds ratio)

 $\alpha$ <sup>a</sup> Analyses adjusted for age, sex, and study center

b<br>Pack-year variable categorized into five groups: never smokers and study-specific quartiles of pack-years smoked

 $c<sub>A</sub>$  Analyses further adjusted for energy intake where available

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# **Table 4**

Gene-environment interactions with nominal interaction p-value <0.01 Gene-environment interactions with nominal interaction p-value <0.01



ABBREVIATIONS: PMH (post-menopausal hormone use); OR (odds ratio)

 $^4$ Interaction OR for SNP (log-additive for number of risk alleles) \* exposure (as categorized above) Interaction OR for SNP (log-additive for number of risk alleles) \* exposure (as categorized above)

 $b$  dijusted for age, sex, study center, and population substructure (principal components  $1\hbox{--}3)$ Adjusted for age, sex, study center, and population substructure (principal components 1–3)

 $\ensuremath{^\mathcal{C}}_\text{Analysis}$  further adjusted for energy in<br>take (where available) Analyses further adjusted for energy intake (where available)