

# Genetic and environmental exposures constrain epigenetic drift over the human life course

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Epigenetic mechanisms such as DNA methylation (DNAm) are essential for regulation of gene expression. DNAm is dynamic, influenced by both environmental and genetic factors. Epigenetic drift is the divergence of the epigenome as a function of age due to stochastic changes in methylation. Here we show that epigenetic drift may be constrained at many CpGs across the human genome by DNA sequence variation and by lifetime environmental exposures. We estimate repeatability of DNAm at 234,811 autosomal CpGs in whole blood using longitudinal data (2–3 repeated measurements) on 478 older people from two Scottish birth cohorts—the Lothian Birth Cohorts of 1921 and 1936. Median age was 79 yr and 70 yr, and the follow-up period was ~10 yr and ~6 yr, respectively. We compare this to methylation heritability estimated in the Brisbane Systems Genomics Study, a cross-sectional study of 117 families (offspring median age 13 yr; parent median age 46 yr). CpG repeatability in older people was highly correlated (0.68) with heritability estimated in younger people. Highly heritable sites had strong underlying *cis*-genetic effects. Thirty-seven and 1687 autosomal CpGs were associated with smoking and sex, respectively. Both sets were strongly enriched for high repeatability. Sex-associated CpGs were also strongly enriched for high heritability. Our results show that a large number of CpGs across the genome, as a result of environmental and/or genetic constraints, have stable DNAm variation over the human lifetime. Moreover, at a number of CpGs, most variation in the population is due to genetic factors, despite some sites being highly modifiable by the environment.

[Supplemental material is available for this article.]

Epigenetics refers to the chemical modifications of DNA that regulate gene expression without altering the underlying DNA sequence. DNA methylation (DNAm), the covalent addition of a methyl group to a cytosine nucleotide primarily in the context of a CpG dinucleotide, is among the best-studied epigenetic mechanisms and plays a crucial role in mammalian development (Li et al. 1992; Tate et al. 1996), X-chromosome inactivation (Panning and Jaenisch 1998), imprinting (Li et al. 1993), cell-lineage specification (Bröske et al. 2009; Trowbridge et al. 2009; Challen et al. 2012), and maintaining genome stability (Lengauer et al. 1997; Bellacosa et al. 1999). In mammalian development, shortly after fertilization, DNAm marks are largely erased. Genome-wide and tissue-specific *de novo* methylation takes place following blastocyst formation. Primordial germ cells also undergo demethylation, followed by *de novo* methylation during gametogenesis and establishment of sex-dependent imprints. In contrast with sequence

variants, DNAm is dynamic and is influenced by many factors, including lifestyle, environment, and genetics (Jirtle and Skinner 2007; Mathers et al. 2010; Alegría-Torres et al. 2011).

Two forms of inheritance of epigenetic marks have been demonstrated: genetic inheritance and epigenetic inheritance. The latter is independent of DNA sequence and may occur through incomplete erasure of epigenetic marks during epigenetic reprogramming at the gamete and zygote stages. The genetic inheritance of an epigenetic state occurs when an individual's DNA sequence affects the epigenetic state. A large number of common single nucleotide polymorphisms (SNPs) have been shown to affect methylation at CpG sites that are typically in close proximity to the SNP, i.e., located in *cis* (Kerkel et al. 2008; Gibbs et al. 2010; Bell et al. 2011; Grundberg et al. 2012), with some SNPs explaining up to 90% of the variation in methylation at the associated CpG site (Gibbs et al. 2010). These genetic variants are referred to as DNA

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methylation quantitative trait loci (mQTL), and methylation at the associated CpG sites is highly heritable (McRae et al. 2014). Recent estimates of mean genome-wide DNAm heritability (proportion of variance explained by additive genetic factors) are 12%–18% in blood (Bell et al. 2012; Grundberg et al. 2013; McRae et al. 2014), but locus-specific levels have shown to be variable. Mean heritability across tissues vary, with estimates of 5% in placenta and 7% in human umbilical vein endothelial cells (HUVECs) (Gordon et al. 2012).

Age-related changes in DNAm are well documented, and a large number of age-associated methylation sites have been identified across the genome (Bocklandt et al. 2011; Hannum et al. 2013; Horvath 2013). The epigenome diverges with age, even within monozygotic twins, with older twins having less similar epigenetic profiles than younger twins (Fraga et al. 2005). The differences in the epigenome of genetically identical individuals have been attributed to differences in external factors, such as changes in lifestyle and environment, but also to internal stochastic factors, such as small defects in transmitting epigenetic information through successive cell divisions or maintaining it in differentiated cells, a process termed “epigenetic drift” (Fraga et al. 2005). Epigenetic differences in genetically identical twins may be key to understanding disease discordance in monozygotic twins.

Understanding what factors contribute to DNAm variation is important not only for understanding the role of DNAm in complex disease traits, but also for informing future epigenetic studies that take a cross-sectional approach. There have been relatively few longitudinal studies on DNAm, and these have focused either on a handful of selected loci (Talens et al. 2010; Wong et al. 2010) or only on loci whose DNAm changes with age (Madrigano et al. 2012; Florath et al. 2013), and others have used only global measures of methylation (Bjornsson et al. 2008). Therefore, it is still not fully known how stable DNAm levels across the genome are over time and to what extent these are influenced by heritable versus environmental factors.

When longitudinal data are available, one measure of stability over time is the repeatability ( $R_m$ ) of methylation levels.  $R_m$  is the correlation between repeated measures of the same individual (Falconer and Mackay 1996) and is calculated as the ratio of the between-individual variance over the total variance. Variance between individuals can be partitioned into variance due to genetic factors and variance due to permanent differences in the environment between individuals. The genetic variance can be further partitioned into additive and nonadditive (dominant and epistatic effects) genetic variance, and the proportion of phenotypic variance due to additive genetic factors is the (narrow sense) heritability ( $h^2$ ). Since heritability is one of the determining factors of repeatability,  $h^2$  estimates should not exceed  $R_m$  estimates in the same sample, and  $R_m$  therefore sets an upper bound for  $h^2$  (see Methods). By comparing the  $R_m$  and  $h^2$ , we can assess the extent to which variation between individuals is due to additive genetic factors.

To assess the stability of DNAm across the genome over time, we estimated methylation  $R_m$  at 234,811 autosomal CpG probes in whole blood using longitudinal data on 478 older, initially healthy, Scottish individuals from two birth cohorts—the Lothian Birth Cohorts of 1921 (LBC1921) and 1936 (LBC1936) (Deary et al. 2004, 2007, 2012). Participants had two or three repeated measures of DNAm over a period of ~10 yr in LBC1921 and ~6 yr in LBC1936. Median ages at baseline (wave 1) were 79 yr and 70 yr, in LBC1921 and LBC1936, respectively. To assess the extent to which additive genetic factors contribute to DNAm variation, we compared the  $R_m$  for each CpG site to their  $h^2$  estimated in the Brisbane Systems Genomics Study (BSGS). The BSGS is a cross-sectional

study of 117 families comprising adolescent monozygotic (MZ) and dizygotic (DZ) twins, their siblings and their parents, with an age range of 9–75 yr (offspring median age 13 yr; parent median age 46 yr) (Powell et al. 2012).

## Results

### Data

In LBC, DNAm was measured in 2460 whole blood samples (from 1520 individuals) at 485,512 CpG sites across the genome using the Infinium HumanMethylation450 array. After quality control (QC) analysis and probe filtering (see Methods), 2195 samples (from 1366 individuals) and 249,741 methylation probes remained. Only 478 of the 1366 individuals in the QC'ed data had repeated measures of DNAm (186 individuals with two measures and 292 individuals with three measures), and these were used for estimating repeatability. At the three waves of sample collection, LBC1921 participants had median age 79 ( $n = 446$ ), 87 ( $n = 175$ ), and 90 ( $n = 82$ ) yr, and LBC1936 participants had median age 70 ( $n = 920$ ), 72 ( $n = 299$ ), and 76 ( $n = 273$ ) yr (Table 1). In BSGS, DNAm was measured in 614 blood lymphocyte samples using the Infinium HumanMethylation450 array. All 614 samples passed QC with DNAm data available on 261,450 autosomal CpG sites.

### Beta value heteroscedasticity

Methylation levels are summarized as beta values ranging between 0 and 1, where a value close to zero indicates that all copies of the CpG site in the sample were completely unmethylated (no methylated molecules were measured), and a value close to one indicates that every copy of the site was methylated. Six percent of CpG sites (14,930 of 249,741) were either unmethylated (beta  $\leq 0.1$ ) across all samples or completely methylated (beta values  $\geq 0.9$ ) across all samples. The beta value has significant heteroscedasticity, and the standard deviation of the beta value in these very low and very high beta ranges is greatly compressed (Supplemental Fig. S1; Du et al. 2010). Any variation in methylation signal at these high and low ranges is likely to be measurement noise. Since  $R_m$  is a function of both within-individual and between-individual variance, individual consistency will only be evident in the repeatability if there is between-individual variation in the sample (Supplemental Fig. S2). Therefore, despite these CpGs having extremely stable methylation in the entire sample, their estimates of repeatability were very low (Supplemental Fig. S3A). This was also the case for heritability estimates (Supplemental Fig. S3B). Based on assessments of the relationship between mean beta values, beta variance, and  $R_m$ , we excluded these 14,930 probes.  $R_m$  estimates for 234,811 probes were therefore considered in further analysis.

**Table 1. Participant characteristics**

	LBC1921	LBC1936
Total sample size	703	1492
Wave 1: median age	79 yr	70 yr
Wave 2: median age	87 yr	72 yr
Wave 3: median age	90 yr	76 yr
Wave 1: percentage males	39%	51%
Wave 1: number of current smokers	31	103
Wave 1: number of ex-smokers	213	385
Wave 1: number of never smokers	201	432

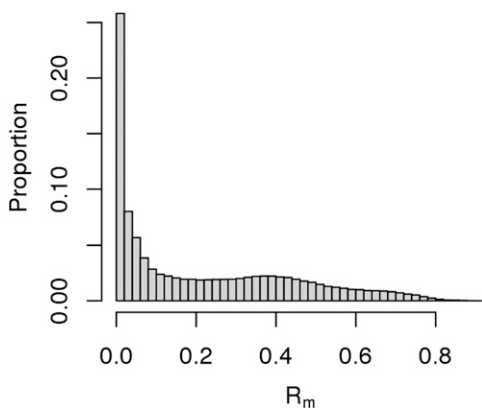
### Genome-wide repeatability

The distribution of  $R_m$  for the 234,811 autosomal CpG sites is shown in Figure 1. We found that 13% (31,273 probes) had  $R_m \geq 0.5$ . To show that the CpG sites with higher  $R_m$  were not observed simply by chance, we used permutation to derive the null distribution. The empirical distribution of  $R_m$  values was not consistent with the null hypothesis of chance associations (Supplemental Fig. S4). For comparison, we also looked at the distribution of  $R_m$  for the 234,811 CpG sites in seven pairs of technical replicates available in the BSGS data (no replicate data was available in the LBCs). The  $R_m$  distribution was shifted to the right, with the majority of probes having high  $R_m$  values (60% with  $R_m \geq 0.5$ ) (Supplemental Fig. S5). However, around 25,000 probes had repeatability of zero even in technical replicates. These may represent poor-performing probes. A much larger number of replicate samples would be needed to accurately estimate the true proportion of repeatable probes and identify those that warrant removal from further analyses. With a larger number of replicates, the standard error in the  $R_m$  estimation will decrease, and it is likely the  $R_m$  distribution would shift further to the right, with most probes either having repeatability close to one or close to zero, providing greater power to distinguish the poor-performing probes.

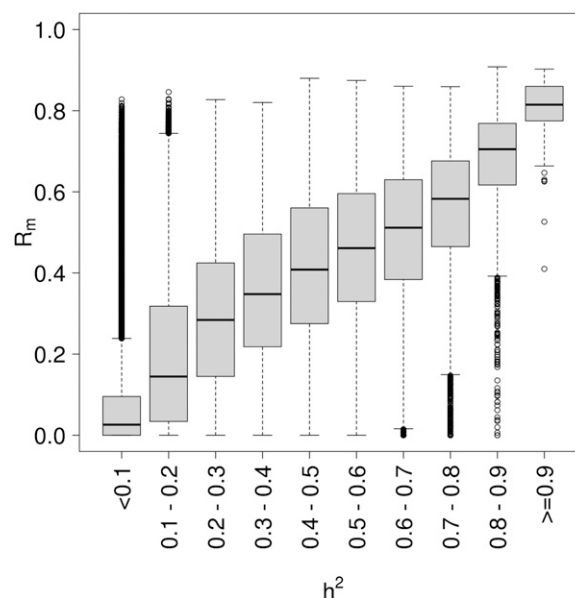
### Correlation between repeatability and heritability

We then correlated probe  $R_m$  to probe  $h^2$ , which had been estimated in the younger BSGS family cohort (McRae et al. 2014). All 234,811 probes were present in the QC'ed BSGS data set. We observed a high correlation of 0.68 ( $P$ -value  $< 1.0 \times 10^{-16}$ ) between  $R_m$  and  $h^2$  (Fig. 2). Given the relationship of both  $R_m$  and  $h^2$  with mean beta values (Supplemental Fig. S3) in the two separate data sets, one might expect the correlation between probe  $R_m$  and  $h^2$  to be confounded by the mean beta values of the CpG probes. However, we found a strong positive correlation between  $R_m$  and  $h^2$  in different bins of mean beta values (Supplemental Fig. S7), suggesting that the observed correlation between  $R_m$  and  $h^2$  is not confounded by mean beta values.

Both genetic and/or environmental factors may be responsible for high DNAm repeatability. Probes with high  $h^2$  generally had high  $R_m$  consistent with a genetic constraint. To demonstrate this, we used the sum of ranks to select 10 CpG sites with the highest  $R_m$  and  $h^2$  and looked for associations with SNPs in the BSGS data. For all 10 CpG sites, we identified significant associations between methylation levels and SNPs (association  $P$ -value  $< 8.3 \times 10^{-9}$ , representing significance at the 5% level after Bonferroni correction



**Figure 1.** Distribution of  $R_m$  for 234,811 autosomal CpG sites in the Lthian Birth Cohorts (LBC).

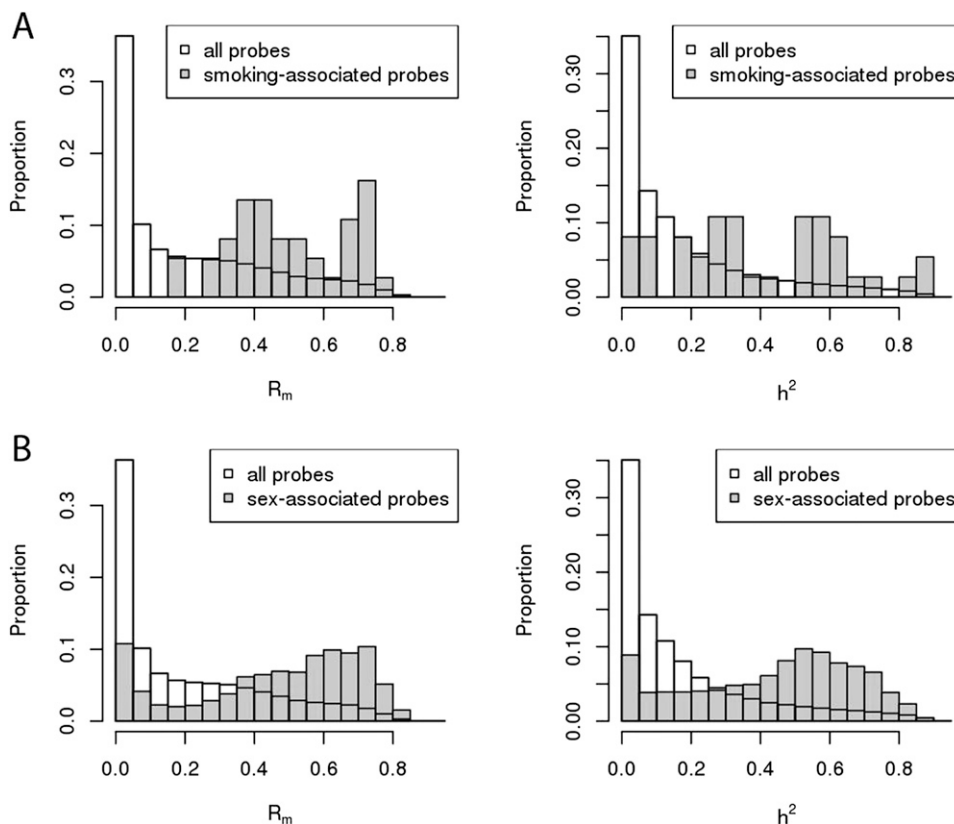


**Figure 2.** Correlation between  $R_m$  and  $h^2$ . The distribution of  $R_m$  estimated in the LBC cohorts by bins of  $h^2$  estimated in the BSGS cohort. Correlation between the two measures was 0.68.

for 6,005,138 SNP association tests performed on the BSGS genotype data). For nine of 10 CpG sites, the associated SNP was located  $< 150$  kb from the CpG site; and for seven of these, the SNP was located within 1 kb. The effect sizes of these single *cis*-acting SNPs were very large, explaining 50%–85% of the variation in methylation at the CpG site (Supplemental Table S1). For one of these associations, the SNP was located on a different chromosome to the CpG site but still explained 30% of the variation in methylation at this site. Using the sum of ranks, we also selected 10 probes that had highest  $R_m$  and lowest  $h^2$  (ranking probes by decreasing  $R_m$  and increasing  $h^2$ ), and none had significant associations with any SNPs, suggesting that the high repeatability may be due to environmental effects.

### Repeatability and heritability of smoking-associated methylation sites

We hypothesized that CpG sites with high  $R_m$  but low  $h^2$  may be under strong influence of environmental factors. To test this, we assessed the distribution of  $R_m$  and  $h^2$  of CpG sites whose methylation levels were associated with an environmental factor—smoking. Using LBC1936 as a discovery sample (731 never smokers versus 139 current smokers), we identified 305 significantly associated autosomal CpG sites (at  $P$ -value  $< 2.13 \times 10^{-7}$ , representing significance at the 5% level after Bonferroni correction for 234,811 tests). Using LBC1921 as a replication sample (333 never-smokers versus 39 current smokers), 37 sites in or near 27 genes were validated ( $P$ -value  $< 0.00016$ ; Bonferroni correction for 305 tests). Of the identified genes, 16 had previously been identified to contain CpGs that were differentially methylated in response to smoking status (Supplemental Table S2; Joubert et al. 2012; Shenker et al. 2013; Zeilinger et al. 2013). Looking at the distribution of  $R_m$  and  $h^2$ , we observed a strong enrichment for probes with high  $R_m$  (Fig. 3). Although some of these CpG sites also had high  $h^2$ , the enrichment for high heritability probes was not as substantial (Fig. 3). The discordance in the distribution of  $R_m$  and  $h^2$  for smoking-associated CpG sites was more obvious when



**Figure 3.** Distribution of  $R_m$  and  $h^2$  for smoking- and sex-associated CpG sites. (A) 37 smoking-associated autosomal probes. (B) 1687 sex-associated autosomal probes.

looking at the larger subset of the 305 CpG sites that were identified in the LBC1936 discovery sample. For this set of 305 probes, a much larger proportion had high  $R_m$  compared to those with high  $h^2$  (Supplemental Fig. S8). The correlation between  $R_m$  and  $h^2$  for the 37 smoking-associated probes was 0.68 ( $P$ -value =  $3.0 \times 10^{-6}$ ), suggesting that even for sites where methylation is influenced by environment, additive genetic factors have a strong influence. We found significant SNP associations for 12 of the 37 smoking probes (32%), 10 of which were with *cis*-acting SNPs (located within 500 kb of CpG site), one was with a SNP on the same chromosome but >500 kb away, and one with a SNP on a different chromosome. A lookup of two independent smoking-associated probes with the highest  $R_m$  and  $h^2$  identified strong

underlying *cis*-SNP effects, each explaining ~10% of the methylation variation at these CpG sites, similar to that explained by smoking status itself (Table 2). The most significant association with smoking status in our data was with a previously reported CpG probe (cg05575921) in the *AHRR* gene. This probe had high  $R_m$  of 0.74 but relatively low  $h^2$  of 0.24, with no significant SNP associations found in the BSGS data.

### Repeatability and heritability of sex-associated methylation sites

We also identified and assessed the distribution of  $R_m$  and  $h^2$  of CpG sites that were differentially methylated between males and

**Table 2.** *cis*-SNP associations with sex- and smoking-associated probes with high  $R_m$  and  $h^2$

Probe	Exposure	CpG location	Closest gene to CpG	$R_m$	$h^2$	<i>cis</i> -SNP	SNP location	mQTL association <i>P</i> -value in BSGS	Proportion of methylation variance explained by SNP in BSGS data	Proportion of variance explained in methylation by exposure	
										LBC1936	LBC1921
cg21117965	Sex	2:220325369	<i>SPEG</i>	0.79	0.89	rs7202333	2:220315970	$2.7 \times 10^{-71}$	0.61	0.03	0.05
cg07187855	Sex	6:30854161	<i>DDR1</i>	0.80	0.84	rs3130779	6:30872203	$2.6 \times 10^{-34}$	0.29	0.04	0.07
cg12803068	Smoking	7:45002919	<i>MYO1G</i>	0.78	0.85	rs6976664	7:45014304	$2.76 \times 10^{-12}$	0.10	0.08	0.05
cg18316974	Smoking	1:92947035	<i>GFI1</i>	0.70	0.83	rs2774947	1:93348423	$1.92 \times 10^{-13}$	0.12	0.07	0.06

(BSGS) Brisbane Systems Genomics Study; (mQTL) methylation quantitative trait loci; (*cis*-SNP) located  $\leq 500$  kb from CpG site.

females. Sex can be considered as a genetic factor, but also as an unmodifiable lifelong “exposure,” and differences between males and females would be confounded by any gender-related differences in environment or lifestyle. Using LBC1936 as a discovery sample (732 females and 760 males) and modeling an individual as a random effect, 3810 autosomal CpGs were significantly associated with sex ( $P$ -value  $< 2.13 \times 10^{-7}$ ; Bonferroni correction for 234,811 tests). Of these, 1687 were replicated ( $P$ -value  $< 1.30 \times 10^{-5}$ ; Bonferroni correction for 3810 tests) in the LBC1921 data (409 females and 294 males). These were located in or near 1157 genes (Supplemental Table S3). Sex-associated CpG probes were significantly enriched for probes with high  $R_m$  and  $h^2$  (Fig. 3; Supplemental Fig. S8). It is important to note that  $R_m$  and  $h^2$  were calculated after adjusting for sex; therefore, the high  $R_m$  and  $h^2$  are not being driven simply by mean differences between the two groups. The correlation between  $R_m$  and  $h^2$  for these CpG sites was 0.76 ( $P$ -value  $< 1.0 \times 10^{-16}$ ). This suggests that methylation at these sites is strongly influenced by genetic factors. Of the 1687 probes, 561 (33%) had significant underlying SNP associations, of which 519 (93%) were located within 500 kb from the CpG site. A lookup of two sex-associated probes with high  $R_m$  and  $h^2$  identified strong underlying *cis*-SNP effects explaining 29% and 61% of the variation in methylation (Table 2), which was much higher than the proportion of variance explained by sex (3%–7%). One of these CpG sites (cg07187855) is located within a regulatory region upstream of the discoidin domain receptor tyrosine kinase 1 (*DDR1*) gene, which based on mice studies, plays an essential role in mammary gland development (Vogel et al. 2001).

## Discussion

Although observations of changes in global (or average) methylation over time have been previously reported (Fraga et al. 2005; Bjornsson et al. 2008) using population-based, longitudinal methylation analyses, we have identified a large number of autosomal CpG sites across the genome whose methylation levels are stable over time. It has been suggested by studies on global methylation that epigenetic drift with age is a genome-wide phenomenon that may occur due to the accumulation of small defects in the transmission of epigenetic marks through successive cell divisions (Fraga et al. 2005). We identify genetic and environmental factors that appear to be responsible for DNAm stability over time at a number of CpG sites. Such factors may therefore be constraining drift at these CpG sites.

The strong correlation between repeatability estimated in older individuals to heritability estimated in younger individuals suggests, first, that additive genetic factors are the main contributors to methylation variation at a large proportion of the measured CpG sites; and second, that CpG sites under strong genetic influence appear to be stable almost over the human life course. Heritability estimates for any given CpG site can change, and can be age-, tissue-, and population-specific. Methylation at the same CpG site measured over an individual's lifetime can have different genetic and environmental effects influencing it so that the variances become a function of age. However, despite the differences between the LBC and BSGS participant characteristics, especially the age of the participants, we find that CpG sites under strong genetic influence have stable methylation even later on in life. Heritability can be thought of as the repeatability within related individuals. For example, if stability is solely due to genetic factors, then the repeatability within a person over time is the same as the correlation between monozygotic twins (and this correlation is the

heritability). Since  $h^2$  is partly derived from the correlation between parents and offspring, high  $h^2$  in BSGS data indicates that methylation levels are stable within the age range of the BSGS samples (9–75 yr). If the same probes have high  $R_m$  in LBC, this indicates that methylation is also stable from 70–90 yr. Therefore, for probes with both high  $h^2$  and  $R_m$ , methylation repeatability is high throughout the age ranges in BSGS and LBC (9–90 yr).

The stability of DNA methylation over time is an important factor to consider when drawing conclusions about causality in epidemiological studies, especially in prospective cohorts, where only a single biological sample has been collected but phenotypic information on participants is collected at subsequent phases. For CpG sites where environment/lifestyle plays the major role, the length or continuity of exposure may be key. For example, for cg05575921, the most strongly associated probe with smoking status, we observed that the mean methylation level at baseline in ex-smokers (unadjusted beta mean = 0.83; SD = 0.087) was significantly lower than those who had never smoked (unadjusted beta mean = 0.88; SD = 0.05) ( $P < 1 \times 10^{-16}$ ), but also significantly higher than current smokers (unadjusted beta mean = 0.61; SD = 0.11) ( $P < 1 \times 10^{-16}$ ). Information on when ex-smokers gave up smoking would allow us to determine the longevity of the effects of smoking on methylation, which would be important in determining the long-term exposure risk in diseases where smoking is a known risk factor. CpG sites influenced by smoking were relatively stable over time. However, although some CpGs had high heritability, overall the genetic contribution was not as strong as that observed for the sex-associated CpG sites. This was evident in the higher correlation between  $R_m$  and  $h^2$  for the sex-associated CpGs compared to that for the smoking-associated CpGs. CpG sites with differential methylation between males and females will be of interest in the context of diseases that are more prevalent in one sex compared to the other. It remains to be seen whether the stability and heritability of DNAm translates to gene expression levels. A comparison between methylation QTLs and expression QTLs would be needed to determine if common genetic variation affects gene expression via changes in DNAm.

There are a number of factors to consider when interpreting the results from these analyses: First, repeatability is a well-established measure for assessing reproducibility (Sokal and Rohlf 1995), but it is only evident if there is between-individual variation in the population. It is therefore not a suitable method for estimating reproducibility in extremely stable measures where there is little between-individual variation, and is likely to underestimate the true number of repeatable CpGs. Our analyses on replicate samples did identify a number of probes with repeatability of zero. These may be probes that are truly stable over time or probes that due to technical reasons do not perform well. Under the null distribution of the test statistic for repeatability, the expected proportion of probes with zero repeatability is 50%. Using this, we can estimate a lower limit of the true proportion of probes that have a nonzero repeatability as  $2*(0.5 - P)$ , where  $P$  is the proportion of observed probes with  $R_m = 0$ . In the LBC data,  $P = 0.16$ , thus we conclude that at least 68% of probes have nonzero repeatability over the time course studied. Similarly, from our analysis of seven replicates ( $P = 0.1$ ), the upper-bound of the proportion of probes that have zero repeatability through either showing no variation in DNA methylation levels or through technical issues is 20%. This indicates that although our estimated proportion of repeatable results in LBC is downwardly biased by the presence of probes with technical issues or having no variation in DNA methylation, it is not markedly biased.

Second, CpGs with mean beta values from 0.4–0.6 showed the highest  $R_m$  and  $h^2$  values (Supplemental Fig. S3). This may partly reflect the heteroscedasticity of the beta value and the reliance of  $R_m$  estimation on the presence of between-individual variation. However, CpGs with strong underlying SNP effects, especially if these are common SNPs, have high between-individual variation with mean methylation around 0.5 (individuals in the three genotype groups tend to have beta values around 0, 0.5, and 1). Therefore, high  $R_m$  and  $h^2$  for probes with mean beta values in the medium methylation range (0.4–0.6) is likely to reflect underlying mQTLs.

Third, although the Infinium HumanMethylation450 array comprehensively evaluates promoter regions and CpG islands as well as other potentially relevant intergenic regions, such as regulatory regions (Sliker et al. 2013), it only interrogates a small subset of the ~28 million CpG sites in the human genome. Therefore, our results may not be representative of all CpG sites across the entire human genome. Finally, whole blood, from which the DNA samples were extracted, consists of a heterogeneous cell population, and it has been shown that these functionally distinct populations have unique DNA methylation profiles (Reinius et al. 2012) which could confound analyses. However, we found a very high correlation (0.98) between  $R_m$  estimated from data adjusted and unadjusted for cell count.

In summary, we show that genetic and environmental factors may constrain epigenetic drift. We also show that methylation at CpG sites measured in blood has a number of the same characteristics as other “complex traits.” First, variation in methylation at any given CpG site can be influenced by both genetic and environmental factors; second, heritable variation can be mapped by associations with SNPs; and third, robust and replicable associations with environmental factors can be found. Therefore, methylation levels at a single CpG site can be treated as a complex trait.

## Methods

### Subjects

The Lothian Birth Cohort 1921 (LBC1921): This cohort is the basis of a longitudinal study of aging (Deary et al. 2004, 2012). All participants were born in 1921, and most completed a cognitive ability test (Moray House Test No. 12) at about age 11 yr in the Scottish Mental Survey 1932 (SMS1932) (Scottish Council for Research in Education 1933). The LBC1921 study attempted to follow up individuals who might have completed the SMS1932 and resided at about age 79 yr in the Lothian region (Edinburgh and its surrounding areas) of Scotland; 550 people ( $n = 234$ ; 43% men) were successfully traced and participated in the study from age 79 yr. To date, there have been four additional follow-up waves at average ages of 83, 87, 90, and 92 yr. The cohort has been deeply phenotyped during the later-life waves, including blood biomarkers, white blood cell counts, cognitive testing, and psychosocial, lifestyle, and health measures. DNA methylation was measured in subjects at an average age of 79 yr ( $n = 515$ ), 87 yr ( $n = 181$ ), and 90 yr ( $n = 87$ ), and data from these waves were used for analyses.

Lothian Birth Cohort 1936 (LBC1936): All participants were born in 1936 (Deary et al. 2007, 2012). Most had taken part in the Scottish Mental Survey 1947 at mean age 11 yr as part of national testing of almost all children born in 1936 who attended Scottish schools on June 4, 1947 (Scottish Council for Research in Educa-

tion 1949). The cognitive test administered was the same Moray House Test No. 12 used in the SMS1932. A total of 1091 participants ( $n = 548$ ; 50% men) who were mostly living in the Lothian area of Scotland were recontacted in later life. Extensive phenotyping, including white blood cell counts, has also been carried out in this study, with data collection waves at three time points. DNA methylation was measured at mean age 70 yr ( $n = 1005$ ), 73 yr ( $n = 336$ ), and 76 yr ( $n = 332$ ), and data from these waves were used for analyses.

Brisbane Systems Genetics Study (BSGS): Participants were originally recruited into an ongoing study of the genetic and environmental factors influencing cognition and pigmented nevi. DNA methylation was measured on 614 individuals from 117 families of European descent. Families consist of adolescent monozygotic (MZ;  $n = 67$  pairs) and dizygotic (DZ;  $n = 111$  pairs) twins, their siblings ( $n = 119$ ), and their parents ( $n = 139$ ). Children have a mean age of 14 yr (range 9–23) and parents 47 yr (range 33–75) (Powell et al. 2012).

### Ethics

Following informed consent, venesected whole blood was collected for DNA extraction in both LBC1921 and LBC1936. Ethics permission for the LBC1921 was obtained from the Lothian Research Ethics Committee (Wave 1: LREC/1998/4/183). Ethics permission for the LBC1936 was obtained from the Multi-Center Research Ethics Committee for Scotland (Wave 1: MREC/01/0/56), the Lothian Research Ethics Committee (Wave 1: LREC/2003/2/29). Written informed consent was obtained from all subjects. The BSGS study was approved by the Queensland Institute for Medical Research Human Research Ethics Committee. All participants gave informed written consent.

### DNA methylation

#### LBC1921 and LBC1936

DNA was extracted from whole blood samples in LBC1921 at MRC Technology, Western General Hospital, Edinburgh (LBC1921), and the Wellcome Trust Clinical Research Facility (WTCRF), Western General Hospital, Edinburgh (LBC1936), using standard methods. Methylation typing of 485,512 probes was performed at the WTCRF. Bisulphite converted DNA samples were hybridized to the Infinium HumanMethylation450 array using the Infinium HD Methylation protocol and Tecan robotics (Illumina). Raw intensity data were background-corrected and normalized using internal controls, and methylation beta values were generated using the R minfi package (Aryee et al. 2014). Quality control (QC) analysis was performed to remove probes with a low (<95%) detection rate at  $P < 0.01$ . Manual inspection of the array control probe signals was used to identify and remove low quality samples (e.g., samples with inadequate hybridization, bisulfite conversion, nucleotide extension, or staining signal). The Illumina recommended threshold was used to eliminate samples with a low call rate (samples with <450,000 probes detected at  $P < 0.01$ ). Samples whose predicted sex, based on XY probes, did not match reported sex were excluded. Probes on the X and Y chromosomes were also removed. After these QC steps, 450,726 autosomal probes remained. Beta values were corrected for effects of sample plate, BeadChip, position on BeadChip, hybridization date, white blood cell counts (basophils, eosinophils, monocytes, lymphocytes, neutrophils), and sex using a generalized linear model with a logistic link function. Residuals from this model were used in further analyses.

## BSGS

DNA was extracted from peripheral blood lymphocytes by the salt precipitation method (Miller et al. 1988) from samples that were time matched to a sample collection of PAXgene tubes for gene expression studies in the Brisbane Systems Genetics Study. Bisulphite converted DNA samples were hybridized to the 12 sample Infinium HumanMethylation450 array using the Infinium HD Methylation protocol and Tecan robotics (Illumina). The proportion of probes with detection  $P$ -value  $< 0.01$  was examined to confirm strong binding of the sample to the array. The probability of a probe within a sample either being called as missing or with a detection  $P$ -value  $< 0.001$  was estimated from the average rate across all probes and samples. A threshold for probes showing significant deviation from random missingness (or excess poor binding) was determined by testing against a binomial distribution for the number of samples at the 0.05 significance level with a Bonferroni correction for the number of probes. Any probe with more than 11 individuals with missing data or more than five individuals with detection  $P$ -values  $> 0.001$  were removed. Probes on the X and Y chromosomes were also removed. After QC, 614 samples and 458,836 autosomal probes remained. Raw intensity values were background corrected and normalized to internal controls using the Genome Studio software and beta values generated. Percentage of monocytes, B cells, NK cells, CD4+ T cells, CD8+ T cells, and granulocytes were predicted using a published prediction algorithm (Houseman et al. 2012). Beta values were corrected for array, sex, age, age<sup>2</sup>, sex  $\times$  age, sex  $\times$  age<sup>2</sup>, and predicted white blood cell counts using a generalized linear model with a logistic link function, and residuals used in subsequent analyses. BSGS methylation data are available from the NCBI Gene Expression Omnibus (GEO; <http://www.ncbi.nlm.nih.gov/geo/>) under accession number GSE56105.

## Genotyping and methylation quantitative trait loci (mQTL) analysis

All BSGS individuals were genotyped on the Illumina Human610-Quad BeadChip. Full details of genotyping procedures is given elsewhere (Medland et al. 2009). Standard QC filters were applied, leaving 528,509 SNPs. The remaining genotyped SNPs were phased using HAPI-UR (Williams et al. 2012) and imputed using 1000 Genomes Phase I Version 3 (The 1000 Genomes Project Consortium 2010) with IMPUTE2 (Howe et al. 2011, 2012). Raw imputed SNPs were filtered to remove any SNPs with low imputation quality as defined by an  $R^2 < 0.8$ . Subsequent quality control removed SNPs with minor allele frequency (MAF)  $< 0.05$ , those with Hardy-Weinberg equilibrium (HWE)  $P < 1 \times 10^{-6}$ , and a missing rate  $> 10\%$ . After filtering, 6,005,138 SNPs remained for further analysis. Association of methylation data (residuals after adjusting for covariates) with imputed genotype probabilities was performed using Merlin (Abecasis et al. 2002).

## Repeatability

The repeatability measure ( $R_m$ ) for each methylation probe was calculated as the ratio of the between-individual variance over the total variance ( $V_{total}$ ). These quantities can be extracted from a random effects linear regression with individual as random effect

$$R_m = \frac{\sigma_\alpha^2}{\sigma_\alpha^2 + \sigma_\epsilon^2},$$

where  $\sigma_\alpha^2$  and  $\sigma_\epsilon^2$  are the between-subject and residual (within-subject) variances, respectively; and  $V_{total} = \sigma_\alpha^2 + \sigma_\epsilon^2$ . The *lmer()* function in the lme4 R CRAN package (Bates et al. 2014) was used to perform the random effects linear regression.

Since the variance between individuals can be partitioned into variance due to genetic factors ( $V_G$ ) and variance due to permanent differences in environment between individuals ( $V_{Ep}$ ), we can define  $R_m$  as

$$R_m = \frac{V_G + V_{Ep}}{V_{total}}.$$

The genetic variance can be further partitioned into additive and nonadditive (dominant and epistatic effects) genetic variance, so that  $R_m$  can be defined as

$$R_m = \frac{V_{G_{additive}}}{V_{total}} + \frac{V_{G_{non-additive}}}{V_{total}} + \frac{V_{Ep}}{V_{total}}.$$

The proportion of phenotypic variance due to additive genetic factors ( $V_{G_{additive}}/V_{total}$ ) is the (narrow sense) heritability ( $h^2$ ). Based on this equation, we can see that  $h^2$  estimates should not exceed  $R_m$  estimates in the same sample, and  $R_m$  therefore sets an upper bound for  $h^2$ . By comparing the  $R_m$  and  $h^2$ , we can assess the extent to which variation between individuals is due to additive genetic factors.

Probes that contained SNPs or had been identified previously as potentially cross-hybridizing (Price et al. 2013) showed some inflation in  $R_m$  and  $h^2$  (Supplemental Fig. S9). Therefore, to avoid potential biases due to effects of SNPs on array binding, we removed all probes that encompass SNPs annotated in the 1000 Genomes Project data (The 1000 Genomes Project Consortium 2010) and high-quality SNPs in dbSNP (Price et al. 2013). Probes where all samples had beta  $\leq 0.1$  or all samples had beta  $\geq 0.9$  were excluded from analysis, as the lack of between-individual variance means  $R_m$  or  $h^2$  cannot be accurately estimated for these probes. The null distribution for  $R_m$  was derived using data permutation.

## Heritability

The heritability for each probe was estimated in BSGS by partitioning its variance into additive genetic ( $V_{G_{additive}}$ ) and environmental ( $V_{Ep}$ ) components by fitting a linear mixed model of the form

$$y = \mu + Za + e,$$

where  $y$  is the vector of the covariate-adjusted methylation level;  $\mu$  is the mean value;  $a$  is the additive genetic effects; and  $e$  is the unique environmental effects (residuals). The model was fitted using QTD (Abecasis et al. 2000), which uses maximum likelihood to estimate the variance components. Probes that contained SNPs or had been identified previously as potentially cross-hybridizing (Price et al. 2013) were excluded from analysis.

## Association analysis

To identify probes associated with smoking status and sex, data from all waves was analyzed using a linear mixed model, where individual was modeled as a random effect. LBC1936 samples were used for discovery and LBC1921 as a replication sample. Association with smoking (never smokers versus current smokers) was performed on residuals corrected for sample plate, BeadChip, position on BeadChip, hybridization date, sex, cell count, and wave. For sex-association analysis, residuals uncorrected for sex were used, and smoking status (current/ex/never) was added as a covariate in the model. Smoking information was available at all waves for 693 of the 1366 (51%) individuals; and for the vast majority of

these individuals, smoking status did not change across the different waves. Therefore, baseline smoking status was used for all samples in the association analysis. In a mixed model analysis, the proportion of variance explained by fixed effects (marginal  $R^2$  value) for each probe can be extracted using the `r.squaredGLMM()` command in the `MuMIn` R package (Barton 2014). For the sex-association analysis, since both sex and smoking were fitted as fixed effects, to get the proportion of variance explained by sex alone, the marginal  $R^2$  had to be extracted from a linear mixed model that included only sex but not smoking. Probes significant at the 5% level after Bonferroni correction in the LBC1936 discovery sample ( $P < 2.13 \times 10^{-7}$  for 234,811 tests) were validated in the LBC1921 replication sample.

## Data access

LBC data have been submitted to the European Genome-phenome Archive (EGA; <https://www.ebi.ac.uk/ega/home>) under accession number EGAS00001000910.

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