Original Article Prognostic value of p53 alterations in human osteosarcoma: a meta analysis

Dong Yao^{1,2*}, Guo-Hong Cai^{2*}, Jing Chen^{2*}, Rui Ling³, Sheng-Xi Wu², Yong-Ping Li³

¹Department of Orthopedic Surgery, Shanxi Provincial Corps Hospital of The Chinese People's Armed Police Force, 36 Shifan Street, Taiyuan 030006, Shanxi, China; ²Department of Anatomy, Histology and Embryology & K.K. Leung Brain Research Centre, Preclinical School of Medicine, Fourth Military Medical University, Xi'an 710032, PR China; ³Department of Vascular and Endocrine Surgery, Xijing Hospital, Fourth Military Medical University, Xi'an 710032, PR China. ^{*}Equal contributors.

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Abstract: Tumor suppressor gene *p53* functions as the guardian of the human genome and mutations in *p53* contribute to cancer development. However, studies that investigated the potential of *p53* as a prognostic marker in osteosarcoma patients have yielded inconclusive results. Based on recommendation of the Cochrane Collaboration, this meta-analysis was conducted using data from the 17 published studies to evaluate the association of *p53* alterations with clinical outcome of osteosarcoma patients. Different databases, including MEDLINE, PsycINFO, Scopus, EMBASE, and the Cochrane Central Register of Controlled Trials (CENTRAL) were searched. Prognostic value of *p53* alterations was determined by risk ratio (RR). The data showed that *p53*-positive immunostaining tended to associate with decreased 2-year survival rates (RR, 1.94; 95% Cl, 1.43 to 2.64; p < 0.0001, l² = 10%). However, the prediction value of RR was smaller with *p53* expression than with *p53* mutations. Moreover, patients who received neoadjuvant chemotherapy and surgery tended to have a stronger association between *p53*-positive staining and 2-year mortality compared to the patients treated with surgery only. However, *p53*-positive staining was not associated with 3-year (RR, 1.64; 95% Cl, 0.84 to 3.20; *P* = 0.15; l² = 56%) and 5-year survival (RR, 1.25; 95% Cl, 0.78 to 2.01; *P* = 0.36; l² = 70%). The data from the current study suggest that *p53*-positive osteosarcoma only predicted a decreased short-term survival rate, but not 3- or 5-year survival.

Keywords: p53, osteosarcoma, survival, meta-analysis

Introduction

Osteosarcoma (OS) is an aggressive malignant bone tumor occurring frequently in children and adolescents. To date, the incidence of osteosarcoma remains high, accounting for the eighth most common childhood cancer and the sixth leading cancer in children under age 15, although recent advances in treatment of osteosarcoma have led to significant improvements in patient outcome [1]. Tumor metastasis frequently occurs in approximately 40% of such patients, indicating tumor resistance to cytostatic chemotherapy [2]. Thus, it is urgently needed to develop and identify biomarkers to predict prognosis and treatment outcome for these patients.

Tumor suppressor gene p53 functions as the guardian of the human genome and mutations in p53 contribute to human carcinogenesis [3].

p53 is localized at chromosome 17 band p13.1 where loss of heterozygosity, deletion, and mutation frequently occur [3]. p53 functions to maintain the stability of the genome [4] and acts as "the guardian of DNA", especially when cells are under stress (such as DNA damage, aberrant proliferative signals, heat shock, or hypoxia) [5]. The wild-type p53 protein regulates genes that are involved in DNA repair, cell cycle checkpoint, and apoptosis [6-8]. In early studies, p53 was found to be frequently mutated in osteosarcoma [9] and subsequent studies investigated the clinical significance of p53 mutations or overexpression of p53 protein in osteosarcoma [10-18]. For example, previous studies showed that p53 expression was associated with a poor response to chemotherapy and worsened survival of patients [11, 16], whereas in other studies the data were inconclusive [15, 18]. In 2004, Pakos et al. conducted a meta-analysis, which suggested that p53



Figure 1. The flow chart of included studies.

alterations might be associated with a poor survival of osteosarcoma patients [19]. However, this controversy continued with the emergence of more recent studies [20-27]. We therefore conducted an updated meta-analysis of all available studies for association of p53 expression or p53 mutations with clinical outcome of osteosarcoma patients.

Methods

Identification of eligible and relevant studies

Based on the recommendations of the Cochrane Collaboration, we performed this meta-analysis. To do so, we considered all studies for association of p53 expression and/or p53 alterations with osteosarcoma outcomes. We searched different electronic databases, including MEDLINE (January 1980 to December 2013), PsycINFO (January 1980 to December 2013), Scopus (January 1980 to December 2013), EMBASE (January 1980 to December 2013), and the Cochrane Library (Issue 11 of 12, Dec 2013). The search was limited to human studies in all languages and types of publications. The search terms used were: osteosarcoma, p53, TP53, p53 protein, p53 mutation, and 17p13 gene and the full search strategy were illustrated in Figure 1 for numbers of studies reviewed and analyzed. Such strategy was developed for MEDLINE and was adapted for the other electronic databases. References of retrieved studies were screened and we then contacted the investigators to request additional data when key information relevant to the meta-analysis was missing. All studies on the relationship between TP53 status and clinical outcome (death) were eligible for this meta-analysis, regardless of the method of detection [immunohistochemistry (IHC) for measuring protein levels and reverse transcription-PCR (RT-PCR) techniques for identifying mutations or other gene changes].

Definitions and standardizations

For consistency, "p53" stands for the gene, while p53 is for protein, and "p53 status" is to cover both the gene and protein as a marker. Nuclear accumulation of mutant p53 protein, which are induced by p53 alterations, can be detected by immunohistochemistry (IHC) [10]. However, accumulation of p53 protein detected by IHC does not necessarily correspond to p53 mutations measured by RT-PCR [28]. Thus, an overall analysis was considered for all data, regardless of whether protein expression or mutation was being evaluated. For example, for studies using IHC only, we used prespecified rules to standardize the p53 status as much as possible to define a positive p53 status based on different cut-off thresholds. We defined positive p53 protein expression as nuclear staining in at least 10% of tumor cells, a standard used by most studies [27]. When different definitions were used, we accepted the cutoff point closest to the 10% level [19]. The clinical outcome used was mortality of the patient. Clinical outcomes were standardized to include 24, 36, or 60 months follow-up in all studies.

Inclusion criteria

Original studies were considered for inclusion in this meta-analysis if they met with the following criteria: i) The patients were diagnosed pathologically as osteosarcoma; ii) treatments of patients included radiotherapy, chemotherapy, surgery, or a combination of both; iii). The 2-year, 3-year or 5-year survival rates were reported; and iv). The comparison between patients with low or undetectable p53 and patients with upregulated p53 was performed in terms of the survival rate.

Data extraction

Two investigators (D. Y. and J. C.) independently screened the titles and abstracts of all potentially eligible studies. The full text articles were then assessed independently by two other investigators (Y-P. L. and G-H. C.) to determine whether the articles met the inclusion criteria. After that, three other investigators (S-X. W., R.L. and G-H. C.) independently extracted data (study characteristics and results) using data extraction forms and then the collected data were entered into Rev-Man 5.1 using the double-entry system. Point estimates for selected variables were extracted and checked by the other two reviewers. In case of disagreement between these two reviewers, a consensus was achieved through discussion among all of the reviewers. A record of reasons for excluding studies was kept.

Data collection and analysis

We collected the following data from each study: i). General study information, such as title, authors, publication source and publication year; ii). Characteristics of study population (e.g. sample size, patient age, and osteosarcoma classification); iii). Treatment data, such as neoadjuvant chemotherapy and surgery; iv). p53 status, such as expression or mutation; v). Detection of p53 status methods (e.g. IHC, antibody used, IHC cut-off point, and PCR amplification of the exons).

After that, the meta-analyses were performed using Rev-Man analyses software (Rev-Man 5.1) according to Cochrane Handbook for Systematic Reviews of Interventions [29]. Data on the predictive ability of p53 overexpression or p53 alterations for outcome were combined from all 17 studies using RR for 2-year, 3-year and 5-year mortality. Measurements of the graphs published in the articles were used if we could not get the raw data from the authors. When only the standard error was reported, it was converted into standard deviation [29]. I² statistics were used to measure heterogeneity of the studies. If the I² value was less than 50%, a fixed-effects meta-analysis was applied, whereas if the I² value was 50% or more, the random-effects meta-analysis was performed [29]. Sensitivity analyses were performed and aisual assessment of the funnel plot calculated by RevMan Analyses software was used to investigate the potential publication bias.

Results

Study selection

In this study, we first searched MEDLINE, PsycINFO, Scopus, EMBASE, and the CENTRAL

databases and reviewed a total of 840 published studies (**Figure 1**). Initially, we excluded 808 publications, 175 of which contained animal experiments, 465 of which were not in osteosarcoma, 50 of which were not for p53, and 118 of which were excluded because they were either comments, editorials, reviews, case reports, or duplicated publications. We obtained 32 publications that met our inclusion criteria, but additional 15 publications were eventually excluded because of lack of full text [30] or detailed data [31-44]. Finally, we obtained the remaining 17 studies for this meta-analysis [10-18, 20-27].

Description of included studies

The detailed characteristics of the included studies were shown in Table 1. Overall, 595 patients were included in this analysis. The median or mean age of patients was 24.6 years old, ranging between 15 years [13, 16, 20] and 67 years old [12]. Seven studies [11-13, 15, 17, 18, 20] were conducted on osteosarcoma in high histological grades, while three studies [10, 21, 27] were conducted on osteosarcoma in low or intermediate histological grades and four studies [14, 22, 24, 25] were on osteosarcoma in all histological grades. However, there was no grade data in three studies [16, 23, 26]. Patients in 5 of these 17 studies received surgery treatment only [12, 18, 20, 24, 26], whereas patients in 12 studies were treated with neoadjuvant chemotherapy and surgery. p53 status was shown as p53 gene/protein expression in 13 studies, while remaining 4 studies only showed p53 mutation [13, 16, 18, 26].

p53 gene/protein expression was analyzed by using immunohistochemistry, while *p*53 mutation was assessed using PCR. Eight studies used 10% as the cutoff value for p53 protein positivity, whereas different thresholds (0-25%) were used in the remaining reports (**Table 1**). In most studies, clone D0-7 antibody was used immunohistochemically to detect expression of p53 protein. Two-year survival rates differed significantly (P < 0.001) across the thirteen eligible studies (ranged between 6% and 52%), which may be due to differences in patient populations (e.g., tumor grade and stage) and/or treatment options.

The meta-analyses

Based on Cohen categories for evaluating the magnitude of effect sizes, p53-positive status

Author (yrs.)	N	Age (mean yrs.)	HG (Grade, N)	Treatment	TP53	Method	IHC antibody	IHC cutoff	PCR Exons	Death in 2 yrs. n (%)
Papai (1997) [10]	21	20	17 (Grade IIb) NC + surgery expression IHC DO-7 5% 4 (Grade IIa)		5%		8 (38)			
Goto (1998) [11]	32	16	23 (Grade III) 9 (Grade IV)	NC + surgery	expression	IHC/PCR	DO-7/Rsp53	> 0%	MS	14 (44)
Jensen (1998) [12]	25	67	9 (Grade III) 16 (Grade IV)	surgery	expression	IHC	D0-7	10%		11 (44)
Yokoyama (1998) [13]	17	15	8 (Grade III) 7 (Grade IV)	NC + surgery	mutation	PCR*			4-8	1 (6)
Gorlick (1999) [14]	53	17	11 (Grade I) 24 (Grade II) 10 (Grade III) 8 (Grade IV)	NC + surgery	expression	IHC	1801/D0-7	> 0%		16 (30)
Tsuchiya (2000) [16]	27	15	NA	NC + surgery	mutation	PCR*			5-9	11 (41)
Uozaki (2000) [17]	70	16	43 (Grade III) 27 (Grade IV)	NC + surgery	expression	IHC	D0-7	10%		17 (24)
Oda (2000) [15]	25	17	6 (Grade III) 19 (Grade IV)	NC + surgery	expression	IHC	NR	10%		4 (16)
Kawaguchi (2002) [18]	23	55	8 (Grade III) 15 (Grade IV)	surgery	mutation	IHC/PCR	1801	10%	5-9	12 (52)
Tsai (2004) [21]	22	16	3 (Grade II) 11 (Grade IIa) 8 (Grade III)	NC + surgery	expression	IHC	D0-7	5%		9 (41)
Ferrari (2004) [20]	19	15	NA	surgery	expression	IHC	D0-7	25%		3 (16)
Kaseta (2008) [22]	35	30	4 (Low grade) 31 (High grade)	NC + surgery	expression	IHC	D0-7	25%		NA
Ozger (2009) [23]	45	20	NA	NC + surgery	expression	IHC	D0-7	10%		6 (13)
Boulytcheva (2010) [24]	40	NA	4 (Grade I) 22 (Grade II) 2 (Grade III) 12 (Grade IV)	surgery	expression	IHC	D0-7	10%		NA
Hu (2010) [25]	44	25	5 (Grade I) 16 (Grade II) 16 (Grade III) 7 (Grade IV)	NC + surgery	expression	IHC	NR	10%		NA
Seidinger (2011) [26]	41	NA	NA	surgery	mutation	IHC/PCR	D0-7	NR	10	12 (29)
Wu (2012) [27]	56	Range (13-37)	56 (Grade IIb)	NC + surgery	expression	IHC	D0-7	10%		NA

Table 1. Patien	t characteristics	in each study
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Note: HG, histological grades; N, number; NC, neoadjuvant chemotherapy; *PCR/single-strand conformational polymorphism.

tended to be associated with a poor 2-year survival rate and a higher risk of death within 2 years (**Figure 2A**, RR, 1.94; 95% Cl, 1.43 to 2.64; P < 0.0001). The test for heterogeneity showed that these studies were not heterogeneous ($I^2 = 10\%$, **Figure 2A**). RR was smaller in studies of p53 protein expression than in studies of p53 alterations (**Table 2**). These studies showed that patients received neoadjuvant chemotherapy and surgery tended to have a stronger association of p53-positive status with 2-year mortality when compared to patients treated with surgery only.

However, our further analysis showed that p53-positive status was not associated with 3-year survival (**Figure 2B**, RR, 1.64; 95% Cl, 0.84 to 3.20; P = 0.15) and 5-year survival (**Figure 2C**, RR, 1.25; 95% Cl, 0.78 to 2.01; P = 0.36). The test for heterogeneity showed that these studies were more heterogeneous ($I^2 = 56\%$, **Figure 2B**; $I^2 = 70\%$, **Figure 2C**). Nevertheless, there was no further layer analysis conducted when considering the small size effect and limited number of included studies. In this case, only the random-effects model was performed when the I^2 value was 50% or more.

p53 alterations in human osteosarcoma

A	Experimental		Control			Risk Ratio	Risk Ratio		
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	M-H, Fixed, 95% CI		
Gorlick 1999	3	8	13	45	11.6%	1.30 [0.48, 3.55]			
Goto 1998	4	9	10	23	16.6%	1.02 [0.43, 2.43]	+		
Hu 2010	2	19	1	25	2.5%	2.63 [0.26, 26.92]			
Jensen 1998	2	6	9	19	12.7%	0.70 [0.21, 2.40]			
Kawaguchi 2002	4	5	8	18	10.2%	1.80 [0.91, 3.54]			
Oda 2000	2	7	2	18	3.3%	2.57 [0.44, 14.87]			
Ozger 2009	5	23	1	22	3.0%	4.78 [0.61, 37.75]			
Papai 1997	7	9	2	12	5.0%	4.67 [1.26, 17.34]			
Seidinger 2011	3	3	9	38	5.2%	3.59 [1.85, 6.99]			
Tsai 2004	3	6	6	16	9.6%	1.33 [0.48, 3.70]			
Tsuchiya 2000	9	15	3	12	9.8%	2.40 [0.83, 6.95]			
Uozaki 2000	5	9	8	44	8.0%	3.06 [1.30, 7.20]	_ _ _		
Yokoyama 1998	0	4	1	13	2.3%	0.93 [0.04, 19.38]			
Total (95% CI)		123		305	100.0%	1.94 [1.43, 2.64]	•		
Total events	49		73						
Heterogeneity: Chi ² =	13.26, df =	12 (P =	0.35); l ² =	= 10%					
Test for overall effect:	Z = 4.23 (F	> < 0.00	01)			-	0.02 0.1 1 10 50		
B			,			Far	vours experimental Favours control		
D	Experime	ental	Contro			Risk Ratio	Risk Ratio		
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% C	I M-H. Random. 95% CI		
Boulytcheva 2010	6	12	11	36	23.0%	1.64 [0.77, 3.46]	1-		
ferrari 2004	1	6	8	13	9.4%	0.27 [0.04, 1.71]			
Hu 2010	5	19	9	25	20.2%	0.73 [0.29, 1.83]			
Ozger 2009	7	23	1	22	8.3%	6.70 [0.90, 50.08]			
Tsuchiya 2000	10	15	3	12	18.1%	2.67 [0.94, 7.57]			
Wu 2012	30	40	4	16	21.0%	3.00 [1.26, 7.14]			
Total (95% CI)		115		124	100.0%	1 64 [0 84 3 20]	•		
Total overts	50	115	36	124	100.070	1.04 [0.04, 0.20]	•		
Hotorogonoity: Tau ² = (0.36. Chi2 -	11 26	df - 5 (D -	- 0.05)	12 - 56%		++		
Test for overall effect: 7	$7 = 1 \Lambda \Lambda (P)$	= 0.15)	ui – 5 (F -	- 0.03)	, 1 = 50 %		0.005 0.1 1 10 200		
	- 1.44 (P	- 0.13)				Fa	avours experimental Favours control		
С	Experime	ental	Contro	bl		Risk Ratio	Risk Ratio		
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% C	M-H, Random, 95% Cl		
Boulytcheva 2010	12	12	24	36	28.8%	1.45 [1.13, 1.87]	-		
Hu 2010	8	19	20	25	21.8%	0.53 [0.30, 0.92]			
Kaseta 2008	15	21	6	11	20.8%	1.31 [0.72, 2.39]			
Ozger 2009	10	23	6	22	16.1%	1.59 [0.70, 3.64]			
Tsuchiya 2000	10	15	3	12	12.5%	2.67 [0.94, 7.57]			
Total (95% CI)		90		106	100.0%	1.25 [0.78, 2.01]	₹		
Total events	55		59						
Heterogeneity: Tau ² =	0.19; Chi ² =	= 13.25,	df = 4 (P =	= 0.01)	; l ² = 70%				
Test for overall effect:	Z = 0.91 (P	= 0.36)				F	avours experimental Favours control		

Figure 2. The meta-analysis of p53 status association with patients' risk of death. A: Within two years. B: Within three years. C: Within five years.

Risk of bias in these included studies

Sensitivity analyses were performed to assess the effect of limitations on the evaluation of studies using the 10% IHC cut-off point for p53 expression. The survival difference was somewhat stronger and formally statistically significant in studies using the 10% IHC cut-off value (**Table 2**). Moreover, the funnel plots of analysis of p53-positive status for association with 2-year mortality confirmed a symmetric distribution and suggested that there was non-publication bias (**Figure 3**).

Discussion

A previously published meta-analysis showed a significant association of *p*53 alterations (*p*53 gene mutation or loss of heterozygosity) with 2-year survival [19], while a meta-analysis published in 2013 showed that high p53 expression associated with a poorer prognosis for patients with osteosarcoma [45]. Our current study is remarkably different from these previous studies and we assessed that i). In contrast to the previous studies, we have paid more attention to evaluating the association between

Studies	Cases/total cases	l² (%)	Risk ratio (95% CI)			
All	13/428)	13.26	1.94 [1.43, 2.64]			
IHC only	8/288	7	2.05 [1.33, 3.15]			
Studies on expression	9/320	14	1.81 [1.24, 2.66]			
Studies on mutation	4/108	0	2.28 [1.38, 3.77]			
Treatment: NC + surgery	10/339	0	2.06 [1.41, 3.02]			
Treatment: surgery only	3/89	72	1.84 [0.75, 4.52]			
Sensitivity analyses						
Specific 10% cutoff	7/256	12	2.22 [1.46, 3.40]			

 Table 2. Association of TP53 expression with patient 24 months mortality

Note: If the l^2 value was less than 50%, a fixed-effects meta-analysis was applied. If the l^2 value was 50% or more, the random-effects meta-analysis was used. Abbreviations: CI, confidence interval; NC, neoadjuvant chemotherapy.



Figure 3. Analysis of the publication bias.

TP53 status and long-term survival of osteosarcoma patients and ii). We included more studies; in this meta-analysis, we found that p53 alterations (either p53 protein overexpression detected by IHC or *p53* mutation detected by RT-PCR) associated with poor 2-year survival of osteosarcoma patients, particularly in studies that evaluated *p53* mutations. However, our current data showed p53 alterations didn't have any associations with 3-year survival or 5-year survival of the patients. Thus, our current data demonstrated that p53 alterations could only predict short clinical outcome, but not the longer-term survival of osteosarcoma patients.

It is true that the outcome of osteosarcoma patients has significantly improved throughout the last two decades and 5-year overall survival rate has reached between 50% and 70% [46-48]. In this regards, research in the field would pay more attention to predict 3 or 5-year survival of the patients [20, 22-25, 27]. In terms of

biomarker study using p53 status, the prognostic value of p53 status for long-term survival of osteosarcoma patients seemed to be more controversial, whereas positive results were shown for 2-year survival rate [20, 22-25, 27]. In our current study, the prognostic value of p53 for long-term survival of osteosarcoma patients was limited. Further studies using a larger sample size are needed to confirm it.

p53 alterations contribute to tumorigenesis as an early event and the detection of p53 alterations could help determine the

features of osteosarcoma (such as tumor grade, type, aggressiveness, and metastatic potential) [19, 49]. However, our current results showed that patients who received neoadjuvant chemotherapy and surgery seemed to have a stronger association between TP53positive status and 2-year mortality when compared to patients who were treated with surgery only. This piece of data was inconsistent with others [19, 49]. The reason for this controversial data may be due to the limitation of patients who were treated with surgery only.

Furthermore, in the current meta-analysis, we found that association of poor 2-year survival of patients with presence of p53 alterations was observed stronger than that with the IHC data. It is true that IHC can only detect protein in the case of p53 alterations for p53 point mutation, but IHC can't detect protein in the case of p53 deletion, frame-shift mutation, or early stop codon mutations [50]. Thus, there is no straightforward correlation between IHC and RT-PCR [51]. Other investigators have suggested that the combination of IHC and RT-PCR data may provide complementary prognostic information [52]. However, this has not been accomplished over time; most association studies continued using IHC or PCR individually rather than in combination in the recent 10 years. Thus, studies on p53 mutation could be more reliable.

However, our current meta-analysis does have its limitations. First, there was significant heterogeneity in the results for association between long-term survival (3-year/5-year survival) and p53 status. Considering the small

size effect and limited number of included studies, we did not apply a further layer of analysis and only performed the random-effect model analysis. Secondly, although we had made our best effort to get the full text of all published studies, there were still some studies that failed to be included in our meta-analysis due to the lack of detailed data. Thirdly, some statistical methods used in our current study may be limited, such as using l² to assess the amount of heterogeneity in random-effects meta-analysis [53] and visual assessment of the funnel plot for excluding a publication bias. Fourthly, we didn't assess the association between p53 alterations and some osteosarcoma features (such as tumor type, aggressiveness, and metastatic potential), which may be related to osteosarcoma outcomes.

But, our current meta-analysis did obtain the following data: i). p53 alterations positive status associated with poor short-term survival of patients with osteosarcoma, particularly in osteosarcoma with p53 mutations; ii). p53 alterations didn't associate with the long-term survival of the patients; and iii). Patients received neoadjuvant chemotherapy and surgery had a stronger association of p53 alterations with a 2-year mortality when compared to those treated with surgery only. However, further studies with a larger sample size will confirm the prognostic value of p53 for long-term survival of patients with osteosarcoma and detection of p53 mutations could be the better choice for future study of p53 alterations.

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Disclosure of conflict of interest

None.

Address correspondence to: Dr. Sheng-Xi Wu, Department of Anatomy, Histology and Embryology & K.K. Leung Brain Research Centre, Preclinical School of Medicine, Fourth Military Medical University, Xi'an 710032, PR China. E-mail: wushengxifmmu@163.com; Dr. Yong-Ping Li, Department of Vascular and Endocrine Surgery, Xijing Hospital, Fourth Military Medical University, Xi'an 710032, PR China. E-mail: liyongpingdt@sina.com

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