# Comparative coronary risks of apixaban, rivaroxaban and dabigatran: a meta-analysis and adjusted indirect comparison

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#### AIMS

There are concerns regarding increased risk of acute coronary syndrome with dabigatran. We aimed to assess whether alternative treatment options such as rivaroxaban or apixaban carry a similar risk as compared with dabigatran.

#### **METHODS**

We searched MEDLINE and EMBASE for randomized controlled trials of apixaban, dabigatran or rivaroxaban against control (placebo, heparin or vitamin K antagonist). We pooled odds ratios (OR) for adverse coronary events (acute coronary syndrome or myocardial infarction) using fixed effect meta-analysis and assessed heterogeneity with *I*<sup>2</sup>. We conducted adjusted indirect comparisons to compare risk of adverse coronary events with apixaban or rivaroxaban vs. dabigatran.

#### RESULTS

Twenty-seven randomized controlled trials met the inclusion criteria. Dabigatran was associated with a significantly increased risk of adverse coronary events in pooled analysis of nine trials (OR 1.45, 95% Cl 1.14, 1.86). There was no signal for coronary risk with apixaban from nine trials (pooled OR 0.89, 95% Cl 0.78, 1.03) or rivaroxaban from nine trials (pooled OR 0.81, 95% Cl 0.72, 0.93). Overall, adjusted indirect comparison suggested that both apixaban (OR 0.61, 95% Cl 0.44, 0.85) and rivaroxaban (OR 0.54; 95% Cl 0.39, 0.76) were associated with lower coronary risk than dabigatran.

Restricting the indirect comparison to a vitamin K antagonist as a common control, yielded similar findings, OR 0.57 (95% CI 0.39, 0.85) for apixaban vs. dabigatran and 0.53 (95% CI 0.37, 0.77) for rivaroxaban vs. dabigatran.

#### CONCLUSIONS

There are significant differences in the comparative safety of apixaban, rivaroxaban and dabigatran with regards to acute coronary adverse events.

### Introduction

A number of new oral anticoagulants have been developed that could be used in place of low molecular weight heparins or oral vitamin K antagonists to prevent arterial and venous thrombotic events. These new agents include direct thrombin inhibitors (i.e. dabigatran) and factor Xa inhibitors (i.e. rivaroxaban and apixaban). However, the absence of direct head to head trials makes it difficult to quantify the comparative safety of these three agents,

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which we believe is an important consideration for patients and clinicians who are trying to decide on the preferred option. A number of researchers have used the technique of adjusted indirect comparison (AIC) to compare these agents. In an earlier meta-analysis, we used AIC to demonstrate that rivaroxaban was associated with relatively higher efficacy and slightly increased risk of haemorrhage compared with dabigatran in thromboprophylaxis in hip and knee surgery [1]. Since then, other systematic reviews of the new anticoagulants (including apixaban) in atrial fibrillation or thromboprophylaxis have not identified any major differences in efficacy [2, 3].

The potential cardiac harm of dabigatran was raised by Uchino & Hernandez's systematic review demonstrating a significant association between dabigatran and risk of myocardial infarction (MI) or acute coronary syndrome (ACS) [4]. A more recent review (limited only to trials that recruited >1000 participants) also indicated that dabigatran use was associated with increased risk of coronary events, but no similar signal was detected with rivaroxaban, apixaban and ximelagatran [5]. Heterogeneity in trial populations and control interventions were important limitations of both meta-analyses. One crucial point in this debate is whether warfarin (used as the control intervention) provides a cardioprotective role, instead of dabigatran itself being harmful [6]. If that hypothesis were true, then we might also expect to see similar cardiovascular benefits with warfarin in trials against other new oral anticoagulants.

Given the availability of alternative drugs of comparable efficacy (such as apixaban and rivaroxaban), clinicians and patients who are concerned about the threat of acute coronary events with dabigatran would be greatly aided by information on the presence or absence of any similar signal with apixaban or rivaroxaban. Despite the absence of head to head trials, we are able to use adjusted indirect comparisons based on common control agents (such as warfarin) to assess quantitatively the relative likelihood of acute coronary events with apixaban or rivaroxban as compared with dabigatran. Hence, we conducted a metaanalysis of randomized trials and multiple treatments comparison of acute coronary events with the new oral anticoagulants.

## Methods

### Eligibility criteria

We considered parallel group randomized controlled trials of apixaban, dabigatran and rivaroxaban for prophylaxis/ treatment of venous thromboembolism, ACS or atrial fibrillation. Comparator arms could consist of placebo, other anticoagulant or antiplatelet agents (examples include low molecular weight heparin, vitamin K antagonist, aspirin or clopidogrel). We excluded trials where the duration of intervention was less than 1 week, as well as those where the planned duration of intervention differed between the two arms.

## Search strategy

We searched MEDLINE and EMBASE in April 2012 using the search strategy shown in Appendix S1. We regularly updated the search through PubMed on a monthly basis (most recent December 2013) based on automated electronic notifications for any new articles on PubMed. We also checked the bibliographies of included trials for any relevant studies.

## Study screening and data abstraction

Pairs of reviewers (CSK and SP) or (CSK and JKY) independently and in duplicate assessed the eligibility and extracted numerical outcomes data from the included studies. The data extracted were then checked by a senior reviewer (YKL). The team obtained full consensus on inclusion of the studies and data extraction after resolving any discrepancies though discussion.

The primary outcome of interest was ACS adverse events (encompassing non-ST segment elevation MI as well as ST segment MI and unstable angina). Where ACS outcomes were not explicitly reported, we extracted data on MI as reported by the trial authors.

## Study characteristics and quality assessment

We recorded the dose and duration of interventions and comparators, follow-up, as well as the indications for therapy.

YKL and CSK independently assessed study validity. This included consideration of randomization sequence generation, allocation concealment, blinding, outcome ascertainment of acute coronary events (including whether pre-specified and independently adjudicated) and the possibility of selective reporting or missing data [7].

## Quantitative data synthesis

We used RevMan 5.2 (Nordic Cochrane Center, Kobenhavn) to conduct a fixed effect meta-analysis for dichotomous outcomes, and generated pooled odds ratios (OR), with 95% confidence intervals (CI). The main analysis was on an intention to treat basis. We used the inverse variance method to pool odds ratios obtained from different subgroups of trials.

We assessed statistical heterogeneity using the  $l^2$  statistic, with  $l^2$  values of 30–60% representing a moderate level of heterogeneity [8]. If there was substantial heterogeneity ( $l^2 > 60\%$ ), we planned to use random effects metaanalysis and to explore potential sources of heterogeneity.

We performed AIC with Bucher's method [9] using ITC software (Canadian Agency for Drugs and Technologies in Health, Ottawa, Canada) [10]. Here, pooled ORs from the separate dabigatran, rivaroxaban and apixaban meta-

analyses were indirectly compared using common controls. As there may have been different indications for therapy, we also performed the indirect comparisons based on groups of patients with similar disease conditions.

We estimated the number needed to treat (NNT) using Visual Rx 3.0 (http://www.nntonline.net/visualrx/).

## Results

#### Study selection, design and methodology

The process of study selection is shown in Figure 1. Twenty-seven trials met the inclusion criteria (nine dabigatran, nine rivaroxaban and nine apixaban) [11–36]. The control groups varied with a variety of agents such as enoxaparin (nine trials), warfarin (five trials), low molecular weight heparin and then warfarin (five trials), placebo (five trials), aspirin (two trials) and low molecular weight heparin initially before placebo (one trial). There were 18 trials in prevention or treatment of venous thromboenbolism in different settings, five trials in ACS and four trials in atrial fibrillation. The follow-up ranged from 40 days to 36 months. These results along with the dosages and regimes of treatment groups, time of follow up and follow-up are shown in Appendix S2.

#### Quality assessment

The quality of studies is shown in Appendix S3. Most trials had some degree of lost to follow-up which ranged from 1 to 323 patients who were excluded from safety analysis, withdrew consent or discontinued treatment. All the studies had adequate randomization sequence generation, allocation concealment and reporting of outcomes. Aside from two studies which were unclear, the cardiac outcomes reported for all studies were pre-specified [16, 31] and the majority (24/27 trials) conducted independent adjudication.

### Pooled analysis of ACS adverse events

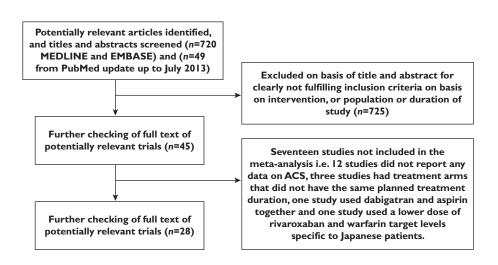
The results for individual outcomes for each trial are shown in Appendix S4.

Pooled estimates for each oral anticoagulant are presented according to control intervention (Figures 2, 3 and 4).

There was no signal of increased risk of ACS with apixaban from nine trials involving 47 559 participants (OR 0.89, 95% Cl 0.78, 1.03) or rivaroxaban from nine trials involving 52 979 participants (OR 0.81, 95% Cl 0.72, 0.93). We did not detect any evidence for statistical heterogeneity in either meta-analysis ( $l^2 = 0\%$ ). In contrast, dabigatran appeared to be associated with ACS in data from nine trials involving 36 966 patients (OR 1.45, 95% Cl 1.14, 1.86,  $l^2 = 0\%$ ). Here we emphasize that statistically significant estimates with dabigatran were demonstrated, not only in patients with atrial fibrillation [21], but also in another trial focused on extended prophylaxis for venous thromboembolism [28].

Statistical testing between subgroups of trials according to anticoagulant showed that the pooled risk of ACS was significantly different between apixaban and dabigatran (P = 0.0007) or between rivaroxaban and dabigatran (P = 0.0001). No difference was observed between the subgroups of trials involving apixaban and rivaroxaban (P = 0.33).

Overall, the adjusted indirect comparison yielded an OR of 0.61 (95% CI 0.44, 0.85) for apixaban vs. dabigatran and an OR of 0.54 (95% CI 0.39, 0.76) for rivaroxaban vs. dabigatran with regard to acute coronary adverse events.



#### Figure 1

Flow chart describing study selection of randomized controlled trials of dabigatran, rivaroxaban or apixaban with myocardial infarction events

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	Apix	aban	Cont	trol		Odds Ratio		Odds Ratio	
Study or Subgroup [reference]	Events	Total	Events	Total	Weight	M-H, Fixed, 95% (		1-H, Fixed, 95% Cl	
I.I.I VTE prophylaxis									-
ADVANCE-I [19]	I.	1599	4	1596	<b>0.9</b> %	0.25 [0.03, 3.82]	ן ←		
ADVANCE-2 [18]	1	1555	1	1550	0.2%	1.00 [0.06, 15.95	j •		
ADVANCE-3 [17]	5	2708	3	2699	0.7%	1.66 [0.40, 6.96	]		
APROPOS [16]	3	623	1	305	0.3%	1.47 [0.15, 14.20]	]		
Subtotal (95% CI)		6485		6150	2.2%	0.96 [0.38, 2.40]	]		
Total events	10		9						
Heterogeneity: Chi <sup>2</sup> =2.16, df=3 (P Test for overall effect: Z=0.09 (P=0		=0%							
1.1.2 ACS									
APPRAISE-I [12]	13	635	20	611	4.7%	0.62 [0.30, 1.25]	_		
APPRAISE-2 [13]	267	3705	284	3687	61.7%	0.93 [0.78,  .  ]		<b></b>	
Subtotal (95% CI)		4340		4298	<b>66.4</b> %	0.91 [0.77, 1.08]		•	
Total events	280		304						
Heterogeneity: Chi <sup>2</sup> = 1.22, df = 1 (P	=0.27); I <sup>2</sup>	=   8%							
Test for overall effect: Z=1.12 (P=0	).26)								
I.I.3 AF									
ARISTOTLE [15]	90	9120	102	9081	23.7%	0.88 [0.66, 1.17]		- <b></b> -	
AVERROES [14]	24	2808	28	2791	6.5%	0.85 [0.49, 1.47]		<b>-</b> _	
Subtotal (95% CI)		11928		11872	30.2%	0.87 [0.68, 1.12]		•	
Total events	114		130						
Heterogeneity: Chi <sup>2</sup> =0.01, df=1 (P	=0.92); I <sup>2</sup>	=0%							
Test for overall effect: Z = 1.07 (P=0									
1.1.4 Extended VTE prophylaxis									
AMPLIFY-EXT [11]	5	1657	4	829	1.2%	0.62 [0.17, 2.33]			
Subtotal (95% CI)		1657		829	1.2%	0.62 [0.17, 2.33]			
Total events	5		9						
Heterogeneity: Not applicable									
Test for overall effect: Z=0.70 (P=0	).48)								
Total (95% CI)		24410		23149	100.0%	0.89 [0.78, 1.03]			
Total events	409		447					•	
Heterogeneity: Chi <sup>2</sup> =3.81, df=8 (P		=0%					HH	-+	
Test for overall effect: $Z = 1.58$ ( $P = 0$							0.1 0.2	0.5 I 2 5 I(	)
Test for subgroup differences: Chi <sup>2</sup>		=3 (P=	0.94). / <sup>2</sup> =	•0%			Favours a	pixaban Favours control	
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#### Figure 2

Apixaban and risk of acute coronary syndrome

## Adjusted indirect comparison based on common controls

As the indirect comparison may have been affected by differences in choice of pharmacological agent in the control arms, we performed an analysis based solely on trials that used vitamin K antagonists as the comparator against the new oral anticoagulant. This demonstrated an OR of 0.57 (95% CI 0.42, 0.84) for apixaban *vs.* dabigatran and an OR of 0.51 (95% CI 0.37, 0.72) for rivaroxaban, using vitamin K antagonists as the common control. (Table 1 and Figure 5).

# Adjusted indirect comparisons stratified according to clinical indication

As the indirect comparison may have been affected by differences in the disease state between participants in the trials, we compared the effects of the anticoagulants in a subgroup of trials covering similar indications. (Figure 6) There was evidence across a number of clinical settings demonstrating a lower risk of acute coronary events with apixaban when compared with dabigatran. Adjusted indirect comparison with rivaroxaban against dabigatran also yielded similar findings of lower risk of acute coronary

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	Riva	roxabaı	n Cor	ntrol		Odds Ratio	Odds Ratio
Study or Subgroup [reference]	Events	Total	Events	Total	Weight	M-H, Fixed, 95%	CI M-H, Fixed, 95% CI
I.2.I VTE prophylaxis							
RECORDI [31]	3	2266	6	2275	I.2%	0.50 [0.13, 2.01]	·
RECORD3 [32]	1	1254	2	1277	0.4%	0.51 [0.05, 5.62]	
RECORD4 [36]	i i	1584	3	1564	0.6%	0.33 [0.03, 3.16]	
Subtotal (95% CI)	•	5104		5116	2.2%	0.46 [0.16, 1.31]	
Total events	5						
Heterogeneity: Chi <sup>2</sup> =0.11, df=2		$I^2 = 0\%$					
Test for overall effect: $Z = 1.46$ (P		,					
	,						
1.2.2 ACS							
ATLAS ACS TIMI 46 [33]	67	2331	44	1160	11.2%	0.75 [0.51, 1.11]	ı —∎∔
ATLAS ACS TIMI 51 [34]	384	10350	229	5176	<b>57.6</b> %	0.83 [0.70, 0.98]	
Subtotal (95% CI)		12681		6336	<b>68.8</b> %	0.82 [0.70, 0.95]	$\bullet$
Total events	45 I		273			-	
Heterogeneity: Chi <sup>2</sup> =0.23, df=1	(P=0.63)	$I^2 = 0\%$					
Test for overall effect: Z = 2.55 (P		,					
	,						
1.2.3 AF							
ROCKET-AF [35]	101	7131	126	7133	24.3%	0.80 [0.61, 1.04]	
Subtotal (95% CI)		7131		7133	24.3%	0.80 [0.61, 1.04]	
Total events	101		126				
Heterogeneity: Not applicable							
Test for overall effect: Z = 1.67 (P	9=0.10)						
1.2.4 VTE treatment							
EINSTEIN-DVT [29]	6	1731	2	1718	0.4%	2.98 [0.60, 14.81	1
EINSTEIN-PE [30]	15	2419	21	2413	4.1%	0.71 [0.37, 1.38	-
Subtotal (95% CI)	15	4150	21	4131	4.5%	0.91 [0.50 1.65	-
Total events	21		23		110/0		,
Heterogeneity: Chi <sup>2</sup> = 2.64, df = 1		· 12 - 47					
Test for overall effect: Z=0.31 (P	• •	,,, – 02	/0				
lest for overall effect: 2=0.31 (r	-0.75)						
I.2.5 Extended VTE prophylaxis							
EINSTEIN-EXT [29]	3	602	Т	595	0.2%	2.97 [0.31, 28.68	]
Subtotal (95% CI)		602		595	0.2%	2.97 [0.31, 28.68	
Total events	3		Т				
Heterogeneity: Not applicable							
Test for overall effect: Z=0.94 (P	9=0.35)						
-	-						
Total (95% CI)		29668		23311	100.0%	0.81 [0.72, 0.93]	Ⅰ ◆
Total events	58 I	•	434				
Heterogeneity: Chi <sup>2</sup> =5.43, df=8	• •	); / <sup>2</sup> = 0%					0.1 0.2 0.5 1 2 5 10
Test for overall effect: Z=3.14 (P				•			Favours apixaban Favours control
Test for subgroup differences: Cl	hi <sup>2</sup> = 2.58,	df=4 (/	P=0.63);	; <b>I</b> <sup>2</sup> = 0%			

#### Figure 3

Rivaroxaban and risk of acute coronary syndrome

events with rivaroxaban across all clinical settings. The Forest plot illustrates that the direction of effect in the indirect comparisons were generally consistent, with no significant statistical heterogeneity. Even if we excluded studies of new oral anticoagulants in patients presenting with ACS itself (Acute Coronary Trials in Figure 6), the AIC OR was 0.61 (95% CI 0.43, 0.87) for apixaban vs. dabigatran and 0.56 (95% CI 0.40, 0.79)

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	Dabig	atran	Con	trol		Odds Ratio		Odds	Ratio		
Study or Subgroup [reference]	Events	Total	Events	Total	Weight	M-H, Fixed, 95% (		M-H, Fixe	ed, 95% Cl		
I.3.I VTE prophylaxis											
RE-MODEL [23]	10	1402	4	699	<b>4.6</b> %	1.25 [0.39, 3.99 <sup>°</sup>	1				
RE-NOVATE [24]	13	233 I	9	1162	10.4%	0.72 [0.31, 1.69]					
RE-NOVATE II [22]	1	1036	I	1019	0.9%	0.98 [0.06, 15.75	] ←				→
Subtotal (95% CI)		4769		2880	15.8%	0.89 [0.46, 1.72]	]				
Total events	24		14								
Heterogeneity: Chi <sup>2</sup> =0.57, df=2 (	P=0.75);	l <sup>2</sup> = 0%									
Test for overall effect: Z=0.35 (P:	=0.72)										
1.3.2 ACS											
RE-DEEM [26]	32	1505	4	373	5.4%	2.00 [0.70, 5.70]					
Subtotal (95% CI)		1505		373	5.4%	2.00 [0.70, 5.70]					
Total events	32		4			- / -					
Heterogeneity: Not applicable											
Test for overall effect: Z = 1.30 (P	=0.19)										
1.3.3 AF											
RE-LY [21]	175	12091	63	6022	71.8%	1.39 [1.04, 1.86]					
Subtotal (95% CI)		12091		6022	71.8%	1.39 [1.04, 1.86]			-		
Total events	175	,	63								
Heterogeneity: Not applicable	175		05								
Test for overall effect: Z=2.22 (P	= 0.03)										
1.3.4 VTE treatment											
RE-COVER [27]	5	1273	3	1266	2.6%	I.66 [0.40, 6.69	1				
RE-COVER II [20]	5	1275	i	1288	0.9%	5.05 [0.59, 43.26]			_		→
Subtotal (95% CI)	5	2553	•	2554	3.5%	2.50 [0.78, 7.99]	-				-
Total events	10		4								
Heterogeneity: $Chi^2 = 0.72$ , df = 1 (		l <sup>2</sup> = 0%	•								
Test for overall effect: Z = 1.55 (P		/-									
1.3.5 Extened VTE prophylaxis											
RE-MEDY 2011 [28]	13	1430	3	1426	2.6%	4.35 [1.24, 15.30]	-			•	<b>→</b>
RE-SONATE [28]	1	681	I.	682	0.9%	1.00 [0.06, 16.04]				-	→
Subtotal (95% CI)		2111		2108	3.4%	3.51 [1.15, 10.68]	]				
Total events	14	•	4								
Heterogeneity: Chi <sup>2</sup> =0.90, df=1 (		l <sup>2</sup> =0%									
Test for overall effect: Z=2.21 (P:	=0.03)										
Total (95% CI)		23029		13937	100.0%	1.45 [1.14, 1.86]			•		
Total events	255		89								
Heterogeneity: Chi <sup>2</sup> =7.54, df=8 (	P=0.48);	l <sup>2</sup> = 0%					0.1 0.	2 0.5	   2	5	
Test for overall effect: Z=3.02 (P:			-							-	
Test for subgroup differences: Ch	i <sup>2</sup> = 5.83, d	f=4 (P	= 0.21); <i>1</i> <sup>2</sup>	=31.4%			ravours	s dabigatran	Favou	rs contro	51

#### Figure 4

Dabigatran and risk of acute coronary syndrome

for rivaroxaban vs. dabigatran with regards to acute coronary events.

#### Sensitivity analysis

Most of the weight in the dabigatran pooled analysis comes from an open label trial in atrial fibrillation [21]. Even if this trial were excluded, the elevated risk of acute coronary events with dabigatran remained significant, with a pooled OR of 1.62 (95% CI 1.03, 2.54).

More recently, the data on myocardial infarction in this trial were subjected to further *post hoc* analysis with a

revised number of MIs in both the dabigatran and warfarin arms [37]. Inclusion of this *post hoc* evaluation data in our meta-analysis did not lead to any major change in our pooled estimate of acute coronary events with dabigatran, OR of 1.38 (95% CI 1.10, 1.74).

#### Number needed to treat

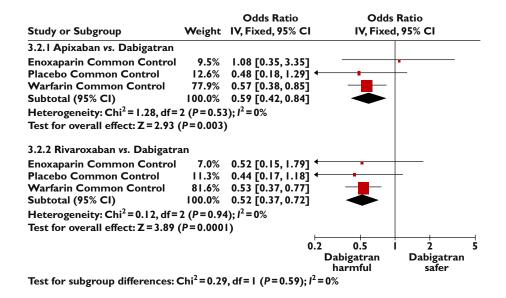
We used the acute coronary event rate of 1.31% (over a median of 2 years) from a large clinical trial (RELY-AF) [21], and applied the odds ratios from the AIC in estimating the absolute effects of using apixaban or rivaroxaban rather

Coronary risks with new oral anticoagulants **BICP** 

## Table 1

Indirect comparison of oral anticoagulants for risk of acute coronary syndrome, stratified according to common control intervention

Vitamin K antagonist as common control	Studies	Participants	Risk of ACS odds ratio (95% Cl)
Apixaban vs. vitamin K antagonist	1	18 201	0.88 (0.66, 1.17)
Rivaroxaban <i>vs.</i> vitamin K antagonist	3	22 545	0.82 (0.64, 1.04)
Dabigatran vs. vitamin K antagonist	4	26 076	1.54 (1.17, 2.02)
AIC via vitamin K antagonist	Apixaban <i>vs</i> . dabigat Rivaroxaban vs. dabig		0.57 (0.39, 0.85) 0.53 (0.37, 0.77)
Placebo as common control			
Apixaban vs. placebo	3	11 124	0.90 (0.76, 1.07)
Rivaroxaban vs. placebo	3	20 754	0.83 (0.71, 0.96)
Dabigatran vs. placebo	2	32 41	1.87 (0.71, 4.91)
AIC via placebo	Apixaban vs. dabigat Rivaroxaban vs. dabig		0.48 (0.18, 1.29) 0.44 (0.17, 1.18)
Enoxaparin as control			
Apixaban vs. enoxaparin	4	12 635	0.96 (0.38, 2.40)
Rivaroxaban vs. enoxaparin	3	10 220	0.46 (0.16, 1.31)
Dabigatran vs. enoxaparin	3	7649	0.89 (0.46, 1.72)
AIC via enoxaparin	Apixaban vs. dabigat	ran	1.08 (0.35, 3.35)
	Rivaroxaban vs. dabi	gatran	0.52 (0.15, 1.79)



#### **Figure 5**

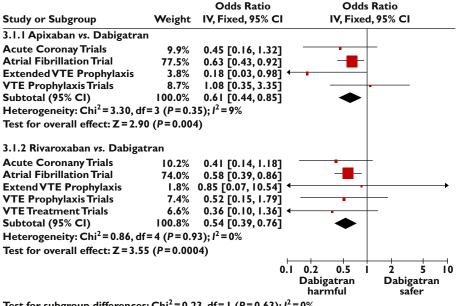
Adjusted indirect comparisons of oral anticoagulants, stratified according to common control intervention

than dabigatran. If apixaban were given to this group of patients instead of dabigatran, there would be five fewer acute coronary events per 1000 patients treated, and an NNT of 198 (95% Cl 143, 407) for this beneficial effect.

Similarly, if rivaroxaban were given to the group of patients instead of dabigatran, there would be six fewer acute coronary events per 1000 patients treated and a NNT of 175 (95% CI 133, 297) for this beneficial effect.

## Selective outcome reporting, dissemination bias and missing data

There were a number of trials with missing outcome data in the journal manuscript where we were unable to obtain the data from the authors or the clinical trials registry (Appendix S5). We also provide a list of studies where suitable data were available but the trial was excluded due to other reasons (Appendix S6).



Test for subgroup differences:  $Chi^2 = 0.23$ , df = 1 (P=0.63);  $l^2 = 0\%$ 

#### **Figure 6**

Adjusted indirect comparison of oral anticoagulants, stratified according to indication for treatment

## Discussion

Our meta-analysis of randomized controlled trials (involving more than 38 000 participants) clearly demonstrates a signal of increased coronary risk with dabigatran, whereas no such signal was seen in meta-analyses of trials that used apixaban (with >45 000 participants) or rivaroxaban (>50 000 participants) in patients with similar conditions. This signal was not completely eliminated even if we used post hoc re-adjudicated data from a large trial of dabigatran, or if we removed that trial altogether. In contrast, the relative lack of cardiac risk with apixaban or rivaroxaban was demonstrated through adjusted indirect comparison, stratified either according to common clinical indication or control therapy, against dabigatran.

We are conscious that dabigatran therapy can have beneficial effects on stroke prevention and we do not aim, in this meta-analysis, to make isolated judgments on whether the benefits of dabigatran outweigh any possible harm. Instead, our primary focus is on the comparative safety of dabigatran relative to other oral anticoagulants that are available as alternative agents for atrial fibrillation, or in patients with venous thromboembolism. Recent systematic reviews have concluded that there are no consistent differences in comparative efficacy of the three agents in atrial fibrillation [38], and that rivaroxaban has similar efficacy to dabigatran in patients with venous thromboembolism [39]. In situations where the available drug therapies are similarly efficacious, we strongly believe that patients and physicians involved in making treatment

choices should be fully informed on any potential differences in harm, particularly if there is a signal of coronary risk with one agent but not the alternative agents. Moreover, neither rivaroxaban nor apixaban appear to be associated with any significantly greater risk of bleeding than dabigatran [38, 39]. While the Canadian Cardiovascular Society have cautioned against dabigatran in patients with atrial fibrillation who are at high risk of coronary events, we are not aware of similar advice from other expert or requlatory bodies [40].

Eikelboom et al. have made a number of observations regarding the associated coronary risk with dabigatran [6]. One possibility is that dabigatran causes acute coronary events while the other is that warfarin carries greater efficacy in preventing such events. However, our analysis did not find any inherent superiority of warfarin in reducing acute coronary events when compared with rivaroxaban or apixaban. Conversely, when we restricted our analysis to trials with warfarin as the common control intervention, dabigatran was found to increase significantly coronary events in a pooled analysis of four trials, and to have greater associated cardiac risk in the adjusted indirect comparison against rivaroxaban or apixaban. While Eikelboom et al. suggest that placebo controlled trials are the optimal method of examining risk of coronary events with dabigatran, we identified only two large placebo controlled trials available, with one demonstrating a nonsignificantly increased rate of coronary events in the dabigatran group [26]. The question surrounding placebo controlled trials is scientifically interesting but perhaps

somewhat moot, given that clinicians and patients have the choice of apixaban and rivaroxaban. Here, we believe the most clinically relevant question is whether dabigatran might be any more harmful than other available agents that have been tested in similar settings against common control interventions.

Our findings are consistent with other metaanalyses that have hypothesized increased ACS risk with dabigatran [4, 5]. Uchino & Hernandez's meta-analysis was the first review to demonstrate the association between dabigatran and increased risk of MI or ACS across different controls and indications for anticoagulant use [4], while Mak's meta-analysis assessed the risk associated with other oral anticoagulants including apixaban, rivaroxaban and ximelagatran and suggested the possibility of differential risk [5]. A recent meta-analysis of five RCTs also reported a statistically significant 48% relative increase in MI with dabigatran [41]. The main weaknesses of these existing reviews is the pooling of data from patients with different conditions, as well as variation in control interventions, and we have aimed to overcome these limitations by conducting an analysis stratified by indication of treatment as well as the type of control. Moreover, none of the previous reviews has evaluated comparative cardiac safety through the use of adjusted indirect comparison.

There is as yet no clear pharmacological mechanism to account for any potential cardiovascular risk with dabigatran, although long term use of a related agent (ximelagatran) has been linked with pro-inflammatory effects in coronary ischaemia [42]. It has been suggested that dabigatran may also give rise to these unfavourable inflammatory effects which may increase atherosclerotic thrombotic events [4]. We have previously postulated that the disparity in efficacy between dabigatran (a direct thrombin inhibitor) and other new oral anticoagulants (direct factor Xa inhibitors) may be related to site of action on the clotting cascade [1].

Our review has several strengths. All the studies included were mainly high quality randomized controlled trials. We excluded trials which could have biased the results such as those with unequal treatment duration for oral anticoagulant and control arms. For our analysis, we were able to stratify by clinical indication as well as control arms and we were also able to use adjusted indirect comparison to directly compare different oral anticoagulants using a common control group. Our results have several limitations. There were 12 trials which we identified which did not report ACS/MI results. We attempted to contact the authors for additional data but we did not receive any response. Differences in trial methods, patient characteristics and outcome measures can affect the validity of adjusted indirect comparisons and we have attempted to tackle this by matching the trials more closely through a number of sensitivity analyses stratifying by common clinical indication or common control.

Our quantitative evaluation has identified significant differences in the comparative coronary risks of apixaban, dabigatran and rivaroxaban. Although we appreciate that head to head RCTs would be preferable, the absence of such RCTs means that our adjusted indirect comparison represents the next best option in helping inform the treatment decision for patients and physicians who are concerned about coronary adverse events with oral anticoagulants.

## **Competing Interests**

C.S. Kwok, S. Pradhan, J.K. Yeong and Y.K. Loke have no competing interests to declare. All authors have completed the Unified Competing Interest form at http:// www.icmje.org/coi\_disclosure.pdf (available on request from the corresponding author) and declare no support from any organization for the submitted work.

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## **Supporting Information**

Additional Supporting Information may be found in the online version of this article at the publisher's web-site:

#### **Appendix S1**

Search strategy

#### Appendix S2

Study design, indications for therapy and intervention arms in the randomized control trials

#### **Appendix S3**

Study quality assessment for included randomized controlled trials

#### **Appendix S4**

Table of main outcomes from individual trials

#### **Appendix S5**

Potentially eligible studies that could not be included in meta-analysis because of missing or unreported coronary events data

#### **Appendix S6**

Acute coronary syndrome data were reported but the trials were excluded for other reasons

#### File S1

PRISMA statement