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# **Obsessive-compulsive disorder is associated with broad impairments in executive function: A meta-analysis**

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# **Abstract**

Obsessive-compulsive disorder (OCD) is a serious and often chronically disabling condition. The current dominant model of OCD focuses on abnormalities in prefrontal-striatal circuits that support executive function (EF). While there is growing evidence for EF impairments associated with OCD, results have been inconsistent, making the nature and magnitude of these impairments controversial. The current meta-analysis uses random-effects models to synthesize 110 previous studies that compared participants with OCD to healthy control participants on at least one neuropsychological measure of EF. The results indicate that individuals with OCD are impaired on tasks measuring most aspects of EF, consistent with broad impairment in EF. EF deficits were not explained by general motor slowness or depression. Effect sizes were largely stable across variation in demographic and clinical characteristics of samples, although medication use, age, and gender moderated some effects.

# **Keywords**

obsessive-compulsive disorder; executive function; meta-analysis

Obsessive-compulsive disorder (OCD) is a serious and often chronically debilitating condition, which affects 2–3% of the population (e.g., Kessler et al., 2005). OCD is characterized by obsessions (intrusive, distressing, and persistent thoughts and images) that are often accompanied by compulsions (ritualized, repetitive behaviors or mental acts) performed in an attempt to avoid or neutralize the distress resulting from obsessions, or according to rules that must be applied rigidly (American Psychiatric Association, 2000). Neuroimaging research has emphasized neurobiological abnormalities that may underlie the

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clinical and neuropsychological symptoms of OCD. Indeed, the current dominant model of OCD focuses on abnormalities in prefrontal-striatal circuits (see Menzies, Chamberlain, et al., 2008a for review) that support executive function (EF). Executive functions are a set of general-purpose cognitive control abilities, mainly supported by the prefrontal cortex (PFC), which allow individuals to regulate their thoughts and behaviors (e.g., Miyake & Friedman, 2012). EF regulates lower level cognitive processes (e.g., perception, motor responses) and thereby enables self-directed behavior towards a goal (e.g., Banich, 2009), allowing individuals to break out of habits, make decisions and evaluate risks, plan for the future, prioritize and sequence actions, and cope with novel situations. EF deficits thus have important consequences for daily life functioning, and may be major contributors to the lack of cognitive flexibility and perseverative, repetitive behaviors that are cardinal symptoms of OCD.

EF is best characterized as a set of separable but related cognitive processes that have both unique and shared individual differences, genetic influences, and neural substrates (e.g., Collette et al., 2005; Friedman et al., 2008; Miyake et al., 2000). One influential model of EF is the *unity/diversity model* (Friedman et al., 2008; Miyake et al., 2000; Miyake & Friedman, 2012), in which three fundamental aspects of EF are identified: (1) updating working memory, (2) shifting (e.g. between tasks), and (3) inhibition, as well as a common EF ability (which is related to both updating and shifting, and may subsume inhibition, Friedman et al., 2008; Miyake & Friedman, 2012). Updating is defined as monitoring and coding incoming information for task-relevance, and replacing no longer relevant information with newer, more relevant information. Shifting is defined as switching between task sets or response rules. Inhibition is defined as suppressing or resisting a prepotent (automatic) response in order to make a less automatic but task-relevant response. Common EF is posited to be the ability to monitor for and maintain goal and context information (Miyake et al., 2000; Miyake & Friedman, 2012). This hypothesis regarding the nature of common EF is compatible with the view that the central role of the frontal lobes is active maintenance of goals, plans and other task-relevant information, which may be essential for all aspects of EF (e.g., Miller & Cohen, 2001). Critically, the unity/diversity model of EF may be a useful vantage for the investigation of cognitive deficits and biases in psychopathology, as disorders such as OCD may be characterized by general (e.g., difficulty maintaining goals) or specific (e.g., difficulty shifting to a new set of behaviors) deficits in EF.

While updating, shifting and inhibition are important aspects of EF, this model in no way posits that these are the only components of EF. For example, working memory (WM) is often considered a component of EF. WM is defined as maintaining or manipulating information across a short delay when that information is not available in the environment. WM maintenance tasks (e.g., simple forward span tasks) require only keeping information in mind temporarily (i.e., 'holding on line'), and involves subsystems for active rehearsal and storage, whereas WM manipulation tasks (e.g., complex and backward span tasks) additionally require the reorganization of the information being maintained (e.g., Fletcher  $\&$ Henson,  $2001$ <sup>1</sup>. WM manipulation is strongly linked to other aspects of EF, while WM maintenance (sometimes called short term memory) is less closely linked to other aspects of

EF (e.g., Engle, Tuholski, Laughlin, & Conway, 1999). WM can also be divided into verbal and visuospatial components (e.g., Baddeley, 1992; 1996; Repovs & Baddeley, 2006). Given evidence for impaired visuospatial ability (e.g., block design and design copying tasks; Abramovitch et al., 2013) in individuals with OCD, which might affect visuospatial WM, it is thus important to evaluate visuospatial and verbal WM separately. In sum, when applying the unity/diversity model to a clinical population, it is important to keep in mind additional domains of EF that may be impacted in a particular disorder.

One challenge for the investigation of EF in clinical (or, indeed, any) populations stems from the fact that many complex tasks may tap multiple aspects of EF. For example, verbal fluency tasks (generating words starting with a certain letter or from a category) likely tap several cognitive processes (Rende, Ramsberger, & Miyake, 2002). However, they have been shown to form a distinct component separable from other EF components (Fisk & Sharp, 2004) and to depend on prefrontal function (e.g., Alvarez & Emory, 2006). Planning tasks are also complex, involving multiple cognitive demands (e.g., Goel & Grafman, 1995), and so may not represent a single EF ability. Notably, verbal fluency and planning tasks are frequently used in clinical studies, including studies of individuals with OCD. Such tasks may be commonly implemented in clinical research because they are viewed as more ecologically sensitive: the complexity of verbal fluency and planning tasks may make them more relatable to real-world tasks that require similar skills. Thus, there are both disadvantages (in terms of interpretability) and advantages (in terms of ecological validity) in the use of such complex EF tasks.

While there is growing evidence for EF impairments associated with OCD, results have been inconsistent, causing controversy about the nature and magnitude of these impairments (for review see Chamberlain, Blackwell, Fineberg, Robbins, & Sahakian, 2005; Kuelz, Hohagen, & Voderholzer, 2004b; Menzies, Chamberlain, et al., 2008a; Olley et al., 2007). Two recent meta-analyses targeting cognitive function more broadly in OCD found some evidence of impaired EF, but inconsistent effect sizes, likely due to the differences in the way they operationalized EF. Specifically, Abromovitch, Aromowitz, & Mittelman (2013) grouped tasks into composite measures of planning  $(d=0.44)$ , response inhibition  $(d=0.49)$ , set shifting/cognitive flexibility (which included verbal and design fluency and WAIS similarities in addition to traditional measures of shifting,  $d=0.52$ ), and verbal ( $d=0.34$ ) and spatial  $(d = 0.37)$  WM (which included measures of updating, in addition to WM maintenance and manipulation). Individual tasks and measures were not analyzed separately. Shin, Lee, Kim, & Kwon (2013) included some individual EF tasks, and in comparison to Abromovitch et al. (2013) found a much larger effect for planning  $(d=0.73)$ , smaller effects on shifting tasks  $(d=0.31-0.51)$  and verbal working memory  $(d=0.11)$ , and somewhat comparable effects on inhibition  $(d=0.55)$  and spatial WM  $(d=0.45)$  tasks. In addition, both meta-analyses included a relatively small number of studies in most EF

<sup>1</sup>Although updating and working memory manipulation are clearly related, and some may consider updating to be a subtype of WM manipulation processes, updating specifically requires adding and removing information from WM, while WM manipulation often involves reorganizing (e.g., reordering) information already held in WM. Thus, there are unique aspects of updating not shared with other types of WM manipulation. For example, the basal ganglia are thought to play a key role in gating information into and out of working memory during updating (e.g., Chatham et al., 2011; Frank, Loughry, & O'Reilly, 2001), which is of strong interest in regard to OCD given evidence of basal ganglia dysfunction.

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analyses (e.g., 12 and 6 studies respectively for planning, compared to 28 in the current meta-analysis), potentially accounting for the variability in effect sizes. Such inconsistencies suggest the need for a larger-scale meta-analysis than has previously been performed, to improve the reliability of EF estimates. The present analysis also uses the well-articulated model of Miyake and Friedman, in conjunction with other perspectives on EF, to provide a more specific rationale for decomposing EF tasks.

Specifically, in addition to variable effect sizes, previous meta-analyses have reported considerable variability in the specific pattern of impairment across different components of EF. Such differences may derive from discrepancies in how EF was operationalized, e.g., how domains of EF were defined. In one case, measures were combined into composites which do not conform to established models of EF such as the unity/diversity model (e.g., fluency tasks, which tap multiple aspects of EF, were grouped with shifting tasks, and updating tasks were grouped with WM tasks), and in the other case only a handful of individual EF tasks were included, and composite measures were not analyzed. In sum, while these previous meta-analyses are valuable in providing a survey of cognitive function in OCD more broadly, they do not permit testing specific hypotheses about EF impairment in OCD.

The current meta-analysis thus addresses these limitations in the extant literature by taking a theoretically-driven approach, applying well-established models of EF to comprehensive analyses, to test competing hypotheses about the nature of EF impairments associated with OCD. At least four hypotheses regarding executive dysfunction in OCD have been proposed, positing that individuals with OCD have (1) a broad impairment in EF, (2) specific impairments in the shifting and/or inhibition components of EF, (3) general slowing of motor responses that accounts for apparent EF deficits, or (4) co-occurring depression that accounts for EF deficits.

# **Hypothesis 1: Broad impairment in EF**

Evidence exists that individuals with OCD have abnormalities in a prefrontal-striatal network that is critical for EF, suggesting that EF may be broadly impaired in individuals with OCD. A meta-analysis of fMRI studies reporting case-control comparisons during a variety of cognitive tasks found evidence for activation abnormalities in a wide PFC network including anterior cingulate cortex (ACC), lateral PFC, and orbitofrontal cortex, as well as in the striatum (caudate and putamen; Menzies, Chamberlain, et al., 2008a). For many different EF tasks, there is joint recruitment of these regions (e.g., Duncan & Owen, 2000). Meta-analyses of neuroimaging studies have found reliable activation of dorsal and ventral lateral PFC and ACC for inhibition (Nee, Wager, & Jonides, 2007), shifting (Wager, Jonides, & Reading, 2004), WM (Wager & Smith, 2003), and verbal fluency (Costafreda, David, & Brammer, 2009), and a narrative review concluded that these regions were also active for planning (Collette, Hogge, Salmon, & Van der Linden, 2006). Orbitofrontal cortex has been implicated in evaluating the reward probabilities associated with different response options (e.g., see Krain et al., 2006 for a meta-analysis), and thus could affect performance across EF tasks, especially when response feedback or reward is involved.

Thus, the frontal-striatal model of OCD would predict broad impairment across multiple aspects of EF that are all supported by the prefrontal areas which are altered in individuals with OCD. Supporting this hypothesis, meta-analyses have reported deficits on a wide variety of EF tasks in individuals with OCD (Abramovitch et al., 2013; Shin et al., 2013). However, as noted above, the magnitude of these effects is inconsistent across previous meta-analyses, and composite measure analyses conforming to established models of EF were not conducted, making it impossible to effectively compare the magnitude of deficits across different aspects of EF. Thus, the breadth and magnitude of EF impairments associated with OCD has not been clearly established, and others have argued that such impairments are an artifact of co-occurring depressing or general motor slowing (see Hypotheses 2–4).

#### **Hypothesis 2: Specific impairment in shifting and/or inhibition**

Since highly repetitive behaviors and thoughts are the hallmarks of OCD, it has been proposed that individuals with OCD have particular difficulty shifting attention between different cognitive representations and behaviors, and/or inhibiting inappropriate responses (e.g., Bannon, Gonsalvez, Croft, & Boyce, 2002; Chamberlain et al., 2005; Olley et al., 2007). Hypothesis 2 is distinct from Hypothesis 1 in that it does not predict equivalent impairment on other aspects of EF, which would show deficits only to the extent that tasks designed to assess other aspects of EF also tap inhibition or shifting (e.g., planning tasks may require inhibiting incorrect moves and verbal fluency tasks may require shifting between subcategories). Thus, Hypothesis 2 is inconsistent with the general EF impairment posited by Hypothesis 1. Although impairments on shifting and inhibition tasks have been reported (Abramovitch et al., 2013; Shin et al., 2013), it is unclear whether the magnitude of these deficits is larger than deficits detected in other aspects of EF, a pattern would more clearly suggest the presence of impairments specific to shifting and inhibition.

# **Hypothesis 3: Apparent EF deficits are actually due to general motor response slowing**

Individuals with OCD are often significantly slower than healthy individuals in completing everyday tasks such as eating and dressing (e.g., Hymas, Lees, Bolton, & Head, 1991), and may perform more poorly on timed than untimed tasks (e.g., Alarcón, Libb, & Boll, 1994). It has thus been proposed that individuals with OCD show a general slowing of motor responses, potentially because of abnormalities in the neuromotor system (e.g., Hymas et al., 1991). However, others have argued that response slowing in individuals with OCD is limited to EF tasks, and is not secondary to OCD symptoms such as checking (e.g., Bucci et al., 2007). In addition, there has been no systematic evaluation of whether individuals with OCD are significantly impaired on untimed, accuracy-based measures of EF, which would suggest that EF deficits cannot be attributed to general motor slowing. Critically, previous meta-analyses did not conduct separate analyses of accuracy measures (Abramovitch et al., 2013) or included only a few accuracy-based measures, which were not compared to RTbased measures (Shin et al., 2013).

A broader impairment in *general processing speed* has also been proposed to account for impaired task performance associated with psychopathology i.e., that the rate of processing limits performance on higher level operations because if processing steps are carried out too slowly, the products of earlier operations may be lost or no longer relevant by the time later operations occur (Nebes et al., 2000). However, in its current form this hypothesis is not empirically falsifiable. Since cognitive slowing is posited to affect even un- timed and unspeeded tasks, impairments on self-paced accuracy measures of EF would not be considered evidence against this hypothesis. Nor would greater impairment on EF tasks than processing speed tasks, since it is always possible to argue that more complex tasks may require more processing steps and are thus more affected by cognitive slowing. Thus, although the motor speed hypothesis can be empirically evaluated (including by metaanalyses), evaluation of the general processing speed hypothesis must await more complete specification of the theory in a way that makes it empirically falsifiable.

# **Hypothesis 4: Apparent EF deficits are actually due to co-occurring**

### **depression**

There is a high rate of comorbidity between OCD and depression, with more than 70% of individuals with a primary diagnosis of OCD also experiencing a mood disorder during their lifetime (61% MDD; Brown, Campbell, Lehman, Grisham, & Mancill, 2001). Depression is also associated with broad impairments in EF (see Snyder, 2012 for a meta-analysis). Thus, some have argued that EF deficits in individuals with OCD are actually due to co-occurring depression rather than OCD per se (e.g., Basso, Bornstein, Carona, & Morton, 2001). Indeed, several studies have found that the effects of OCD on EF were no longer significant after controlling for co-occurring depressive symptoms (Aycicegi, Dinn, Harris, & Erkmen, 2003; Basso et al., 2001; Moritz, Kloss, Jahn, Schick, & Hand, 2003). However, other studies have found that co-occurring depressive symptoms do not account for EF deficits in individuals with OCD (Abramovitch, Dar, Schweiger, & Hermesh, 2011; Nedeljkovic et al., 2009), and most studies have not investigated the effects of co-occurring depression.

# **Current Meta-Analysis**

The current meta-analysis synthesizes previous research findings and applies wellestablished models of EF to test the four hypotheses outlined above. Additionally, we examined the potential moderating effects of demographic (age and gender) and clinical (OCD symptom severity, psychotropic medication use, and co-occurring depression) variables on EF effect sizes. Findings are discussed in light of the barriers that may limit interpretation of the prior literature, and suggestions for potential solutions and future directions are presented.

## **Method**

#### **Inclusion and Exclusion Criteria**

To be eligible for inclusion, studies were required to include a group of individuals with a diagnosis of OCD, and a healthy control group with no diagnosed psychopathology. Studies were included if they tested participants on at least one EF task and reported sufficient

information to calculate effect sizes. EF tasks were defined as detailed in the Coding Procedures section. Studies were excluded if they investigated OCD in samples of participants with organic brain damage (e.g. following head-injury).

#### **Search Strategies**

Searches were conducted in PubMed and ISI Web of Science for papers published through October 2013 using the keywords *obsessive compulsive* paired with *executive function, working memory, response inhibition, inhibitory control, shifting, task switching, planning, verbal fluency, cognitive,* or *neuropsychological*. In addition, a search for unpublished studies was conducted by emailing the corresponding authors of papers included in the meta-analysis, and searching ProQuest for unpublished dissertations and masters theses. The first author, who is experienced in EF research, conducted the search and screening procedures. An initial screen for study eligibility was conducted by examining titles to eliminate studies that clearly did not meet the inclusion criteria. Next, the abstracts of all remaining articles were examined, and if an article appeared likely to meet the inclusion criteria the full text was obtained and checked for inclusion criteria. In addition, the reference lists of included articles, and articles citing included articles, were screened for any studies missed in the database search process. Publication bias was assessed using the trim and fill method (Duval & Tweedie, 2000).

#### **Coding Procedures**

**Tasks—**The types of tasks used in the included studies determined the specific aspects of EF covered in the present meta-analysis. Tasks were coded as tapping one of the following EF components, as detailed below: shifting, inhibition, updating, verbal and visuospatial WM, planning, and verbal fluency. This list is not meant to be exhaustive of all EF abilities, but rather includes all the EF tasks commonly included in the OCD literature. The first author, who is highly experienced with EF research, coded all studies. In addition, for 25% of studies, the coauthors, who are also highly experienced with EF research, coded the EF component measured by each EF task, blind to the first author's coding. Intercoder agreement for the included EF tasks was high  $(99\%)^2$ ; thus, the first author's coding was used in all analyses. Descriptions of the included tasks tapping each EF construct, their dependent measures, and the number of studies reporting each, are provided in Table 1. For each construct, all tasks were included in a composite measure analysis. All tasks and measures reported by at least three studies were also analyzed individually in separate analyses.

In addition, two types of motor speed measures reported by studies in the meta-analysis were included, to determine if there is general motor slowing associated with OCD, as proposed by Hypothesis 3. The Trail Making Test Part A (TMT-A, *k=*32) shares the motor speed and sequencing demands of the Trial Making Test Part B (TMT-B), but does not require shifting like the TMT-B. In addition, 10 studies reported one or more general motor

<sup>&</sup>lt;sup>2</sup>In addition, the following tasks were nominated for inclusion by one author each, but were excluded because there was not agreement that they clearly assessed a single, specific aspect of EF: (1) verbal fluency switching (excluded because it mixes verbal fluency and shifting demands), (2) Rey Complex Figures immediate recall (excluded because it mixes working memory and episodic/ incidental memory demands), and (3) delayed response task (excluded because working memory demands appear to be minimal).

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speed measures, including simple reaction time  $(k=4)$ , choice reaction time  $(k=2)$ , finger tapping (tap fingers as quickly as possible;  $k = 1$ ), and grooved pegboard (put pegs into holes in a board as quickly as possible;  $k = 4$ ). These tasks were included in a general motor speed composite score.

**Moderator Variables—**Information was coded on current OCD symptom severity, age, gender, psychotropic medication use, and co-occurring depression.

**Symptom severity—Total Yale-Brown Obsessive Compulsive Scale (Y-BOCS;** Goodman, Price, Rasmussen, Mazure, Fleischmann, et al., 1989b) scores were reported by 81% of studies. The Y-BOCS is the most frequently used questionnaire to assess OCD symptom severity, and has good reliability and validity (Goodman, Price, Rasmussen, Mazure, Delgado, et al., 1989a; Goodman, Price, Rasmussen, Mazure, Fleischmann, et al., 1989b). It is a clinician-rated 10-items scale, with each item rated from 0 (no symptoms) to 4 (severe symptoms), providing a total range of 0–40.

**Age—**The mean age of the OCD group was included as a continuous variable in metaregression analyses. Age was reported by all studies.

**Gender—**The percentage of females in the OCD group was included as a continuous variable in meta-regression analyses. Gender was reported by 96% of studies.

**Medication—**The percentage of the OCD group currently taking psychotropic medications was coded for each sample. Medication usage was reported by 81% of studies. Many studies only reported the total number of participants using medication; thus, a more detailed analysis of the types or duration of medication could not be conducted. However, when type was reported, medications were generally antidepressants.

**Co-occurring depression—**Because of the diversity of depression measures reported and the lack of detailed depression reporting in many studies, continuous measures of cooccurring depressive symptoms could not analyzed. Instead, the presence of co-occurring depression or depressive symptoms in the sample was coded as a categorical variable. The sample was coded as containing individuals with co-occurring depression or depressive symptoms if (a) any OCD participants were reported to have a comorbid diagnosis of a depressive disorder and/or (b) mean depressive symptoms on a standard depression questionnaire were reported in the clinical range. The sample was coded as containing participants without co-occurring depression or depressive symptoms only if (a) *and* (b) were not met (since the absence of diagnosed depression does not preclude clinically significant levels of depressive symptoms). The clinical range was defined as follows, using published cut-point norms: Hamilton Depression Rating Scale >7 (Kearns et al., 1982), Montgomery–Asberg Depression Rating Scale >7 (Kearns et al., 1982), Beck Depression Inventory  $> 9$  (Beck, 1978), Beck Depression Inventory – II  $> 13$  (Beck, Steer, & Brown, 1996), Children's Depression Inventory > 12 (Kovacs, 1983). Applying both criteria, 55% of samples were coded as having co-occurring depression/depressive symptoms, 18% as having no co-occurring depression/depressive symptoms, and 27% did not report enough information to code. Co-occurring depression was included as a categorical variable in meta-

analyses of variance whenever there were at least three studies in the smaller category. In addition, samples with no co-occurring depression/depressive symptoms were analyzed separately to provide a conservative test of the hypothesis that EF deficits in OCD are driven by co-occurring depression.

#### **Statistical Methods**

For each study, effect sizes comparing the performance of the OCD and control groups on each EF measure were calculated as Cohen's *d*. The sign of *d* was set such that a positive value always indicated poorer performance for the OCD group relative to the control group (e.g. lower accuracy, higher error rates, or longer RTs). Before analyses were conducted, effect sizes were adjusted, weighted, and screened for outliers. First, since it has been demonstrated that *d* is slightly over-estimated when sample sizes are small, Hedges' (1980) small sample bias correction was applied (*dadj*=*d*(1−(3/4*N*)−9, where *N*= number of participants in both samples combined). Second, since sampling error is also higher for smaller sample sizes, effect sizes were weighted by sample size using inverse variance weights  $(w = (2(n_1 + n_2)n_1n_2)/(2(n_1 + n_2)^2 + n_1n_2d^2)$ , where  $n_1$  and  $n_2$  are the number of participants in the OCD and control groups respectively). Finally, analyses were screened for outliers with effect sizes +/− 3 SD from the mean effect size in each analysis; only two such outliers were detected: one was excluded from the shifting composite analysis and one from the digit span forward and verbal WM composite analyses (see Table 2 footnotes a–c).

Only one effect size from each sample comparison was included in each analysis to avoid statistical dependence. When there were three or more studies reporting a particular task, individual tasks were analyzed separately, as described in Coding Procedures. In addition, composite effect sizes were calculated by averaging effect sizes within a construct (e.g. all inhibition measures). In addition, when individuals were tested more than once (e.g., at different points in treatment), only task performance at the first testing timing was analyzed, as practice effects may diminish the EF demands of tasks.

Random effects meta-analytic models were used for all analyses. Although many metaanalyses have used fixed-effects models in the past, random effects models are now considered more appropriate, as there are likely to be many sources of variability between study samples beyond sampling error, violating the assumptions of fixed-effects models (Raudenbush, 2009). Importantly, random effects models allow inferences to be drawn about a broader population of studies, rather than just about the samples tested.

Analyses were conducted using the SPSS meta-analysis macro developed by David B. Wilson (Wilson, 2006). For each analysis, weighted mean effect sizes with 95% confidence intervals were calculated. The null hypothesis that the mean effect size is zero was tested with the *z* statistic at the alpha = .05 significance level. Heterogeneity in effect sizes was tested with the  $Q_t$  statistic (Hedges & Olkin, 1985).  $Q_t$  quantifies the degree to which the studies contributing to each weighted mean effect size can be considered homogeneous. If  $Q_t$  is significant, it suggests that there are substantive differences between the studies in that analysis. Publication bias was assessed with the trim-and-fill method (Duval & Tweedie, 2000). In addition, separate analyses were conducted including only published studies to determine whether the inclusion of unpublished studies affected the results. To ensure that

effects were not driven by failure to match OCD and control samples on IQ, separate analyses were also conducted including only those studies that reported that IQ did not significantly differ between groups.

Moderator analyses were conducted via mixed-effects models with method of moments estimation. Current symptom severity (Y-BOCS), OCD group age, and medication status (% receiving psychotropic medications) were included as continuous variables in separate and combined meta-regression analyses. Co-occurring depression (individuals with clinicallysignificant levels of depressive symptoms and/or depressive disorder diagnosis) was included as a categorical variable in meta-analyses of variance whenever there were at least three studies in the smaller category. Moderator analyses were conducted only for measures with 20 or more effect sizes, as analyses with fewer studies have inadequate power and may produce unstable estimates (Marín-Martínez & Sánchez-Meca, 1998; Sánchez-Meca & Marín-Martínez, 1998).

# **Results**

The search process identified 110 studies for inclusion, 104 of which were published in peer-reviewed journals and 6 of which were unpublished. An additional 26 papers were screened but were excluded because they did not report sufficient information to calculate effect sizes (*k*=11), did not include a healthy control group (*k*=7), included the same data as another paper in the meta-analysis (*k*=4), pre-matched/pre-trained subjects on task performance (*k*=2), or used tasks not clearly classifiable into an EF component (*k*=2; see Supplementary Information, Table S1). Included studies are marked with an \* in the reference list. The full data sets used in the analyses are available from the authors upon request.

In total, the included studies comprised 6,315 participants (3,162 individuals with OCD and 3,153 healthy control participants). On average, individuals with OCD in the included studies were 31.36 years old (SD=7.16, range  $12-49<sup>3</sup>$ ), and there were equal numbers of males and females with OCD (50.14% female, SD=20.40, range 0–100%). Across studies, the average Y-BOCS (Goodman et al., 1989) score was 23.07 (SD=4.32, range 3–30), corresponding to a moderate level of symptom severity, and 38.90% of individuals with OCD were currently taking psychotropic medications  $(SD=37.35$ , range  $0-100\%$ ).

**Weighted mean effect sizes—**Individuals with OCD performed more poorly than healthy control participants on most EF tasks, but the magnitude of these impairments varied somewhat depending on the aspect of EF and the task. Table 2 reports the effect size (*d*) for each measure, along with their 95% confidence intervals and significance test; these effect sizes, with 95% confidence intervals, are also plotted in Figure 1. Table 2 also provides the test for homogeneity of effect sizes across samples (*Q)*, and tests for sensitivity of the results

<sup>3</sup>Six studies had adolescent participants (ages 12–14). Since adult and non-adult studies could potentially differ, supplementary analyses were conducted with studies of adult participants (ages 18+) only. Effect sizes were virtually identical with the adolescent samples excluded as when they were included. Thus, all studies are included in the analyses reported here.

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to publication bias, inclusion of unpublished studies, and IQ matching, as detailed in following sections.

**Inhibition—**There were significant small to medium effects of OCD groups compared to healthy groups for the inhibition composite measure  $(d = .37)$ , Stroop incongruent condition time ( $d = .55$ ), interference ( $d = .36$ ) and accuracy ( $d = .39$ ), and stop signal task reaction times (SSRT;  $d = .62$ ), but only a small and non-significant effect for accuracy on the go/nogo task  $(d = .24)$ .

**Shifting—**There was a moderate and significant effect of group for shifting composite scores  $(d = .50)$ . Examining the individual shifting tasks, the largest effect size was for TMT-B (*d* = .54). However, the effect size was as large for the TMT-A, which does not require shifting  $(d = .57)$ , suggesting that this effect may be primarily driven by slowed general motor speed or sequencing, and not shifting per se. Thus, it may be more informative to focus on shifting tasks that are not confounded by general motor speed demands because they are self-paced and have accuracy, rather than RT, outcome measures (see Table 1). These all have somewhat smaller, but significant, effects (Intradimensional/ Extradimensional (ID/ED) Shift, *d* = .50; Object Alternation/Delayed Object Alternation Test (OAT/DAT), *d* = .32; Wisconsin Card Sorting Task (WCST; *d* = .44). The only shifting task which did not show a significant effect was cued task switching  $(d = .35)$ , which was marginal; however, this should be interpreted with caution as there were few studies using this task  $(k = 3)$ .

**Updating—**Only four studies tested updating, all with an n-back test. Nonetheless, there was a large effect size  $(d = .71)$ , although confidence intervals are wide due to the small number of studies.

**Verbal WM—**There was a small but significant effect of group on overall verbal WM composite scores  $(d = .22)$ . Separating measures into those requiring manipulation of information in WM versus simple WM maintenance, there was a small but significant effect of group on WM manipulation composite scores  $(d = .31)$ , with a marginal effect for digit span backward  $(d = .21)$ . There was no evidence for meaningful verbal WM maintenance impairments in individuals with OCD. There were very small and non-significant effects for verbal WM maintenance composite scores  $(d = .07)$ , and digit span forward  $(d = .08)$ .

**Visuospatial WM—**There were significant effects of OCD groups compared to healthy control groups for visuospatial WM composite scores (*d* = .47), block span (*d* = .43), and self-ordered pointing  $(d = .62)$ . The effect size for delayed-match-to-sample (DMTS) was comparable in magnitude to those for block span and composite scores (*d* = .49), but did not reach significance due to the small number of studies using this task  $(k = 3)$ .

**Verbal Fluency—**There was a small but significant effect of group for verbal fluency composite scores ( $d = .36$ ), with comparable and significant effects for phonemic ( $d = .39$ ) and semantic  $(d = .34)$  verbal fluency.

**Planning—**There was a significant effect for planning composite scores  $(d = .44)$ . Since some studies used latency measures (i.e., time until first move), which may be influenced by overall slowing, latency and accuracy measures were also analyzed separately. There were similar and significant effects for accuracy measures ( $d = .44$ ) and latency measures ( $d = .$ 42).

**General Motor Speed—**There was a significant effect of OCD groups compared to healthy groups for general motor speed measures (*d* = .34), although a great deal of variability between studies led to a wide confidence interval  $(d = .05-.63)$ . As noted above, there was also a significant effect for the TMT-A, the comparison task for the TMT-B, which requires both general motor speed and sequencing  $(d = .57)$ .

#### **Heterogeneity Analyses**

There was significant heterogeneity among effect sizes for all measures except Stroop accuracy, stop signal (marginal), ID/ED shift, cued task switching, verbal WM maintenance, digit span forward, block span, DMTS (marginal), and phonemic verbal fluency. There may be multiple sources of this variability. First, some variability is likely due to differences in methodology across studies. In composite score analyses, tasks are likely to vary in sensitivity (e.g., standard neuropsychological tests are less sensitive to subtle impairments than those designed to assess individual differences in the normal range). Even in analyses of single tasks and measures, task versions may vary in sensitivity (e.g., the standard neuropsychological version of the Stroop task, with separate blocks of neutral and incongruent stimuli, is easier than versions in which trial types are intermixed). Given the myriad variations in tasks, it is not possible to account for this variation. Second, some variability is due to differences in the demographic characteristics of the patient groups included in each study (see Moderator Analyses). Finally, there are likely additional unmeasured moderators, given the known heterogeneity of clinical profiles, genetics, and neurobiology in all diagnostic categories, including OCD.

#### **Sensitivity Analyses**

**Publication Bias—**Effect sizes for Duval & Tweedie's (2000) trim-and-fill analyses are reported in Table 2. Overall, there was little evidence of publication bias. Across analyses, effects were very robust to the trim-and-fill procedure: on average, weighted mean *d* was only 0.05 lower than for untrimmed analyses, and no significant measures became nonsignificant.

**IQ Matching—**Effect sizes for IQ matched samples only are reported in Table 2. Across analyses, effects were very robust to IQ matching: on average, for studies that matched groups on IQ, weighted mean *d* actually was .04 *higher* than those for all samples combined. Nearly all analyses that were significant for all samples remained significant when restricted to IQ-matched samples. Two significant effects, for the verbal WM overall and manipulation composite scores, had slightly reduced effect sizes and became non-significant due to low power, as there were few IQ-matched samples  $(k=11, k=8)$ . In addition, the effect size for Stroop accuracy was higher for IQ-matched samples, but the effect became non-significant due to low power  $(k=4)$ . IQ-matched samples could not be analyzed for

cued task switching, n-back, and DMTS because there were fewer than three IQ-matched samples.

#### **Moderator Analyses**

Effect sizes were largely stable across variation in demographic and clinical characteristics of the samples, although age and gender did moderate some effects. Specifically, was some evidence for increasing effect sizes with increasing age and increasing percentage of female participants, while symptom severity did not moderate effect sizes. Importantly, moderator analyses indicate that deficits in EF associated with OCD are not driven by co-occurring depressing or medication use.

**Depression—**Comparisons of samples with and without co-occurring diagnosis of a depressive disorder or elevated depressive symptoms are reported in Table 3. Table 3 gives effect sizes for samples without co-occurring depression (*absent*) and those with possible co-occurring depression (*possible*), along with the confidence interval and significance test for each group, and the test for the significance of the difference in effect sizes between the group (*Q Between*). Across analyses, effects were very robust to depression. Depression was only a significant moderator for one task measure, Stroop interference, with larger effect sizes for samples *without* co-occurring depression diagnosis or elevated depressive symptoms. However, this finding should be interpreted with caution since there were few Stroop interference studies in participants without depression (*k*=5). In addition, for the TMT-A comparison measure, effects were marginally larger for samples with possible cooccurring depression.

On average, across the 20 EF analyses for which there were enough studies for meta-ANOVA analysis, effect sizes were very similar for those with and without co-occurring diagnosis of a depressive disorder or elevated depressive symptoms (Δ*d* = 0.04). Nearly all EF analyses that were significant with all samples included remained significant when restricted to samples without co-occurring depression diagnosis or elevated depressive symptoms, with the exception of stop signal SSRT, verbal WM overall and manipulation composite scores, and semantic verbal fluency, all of which had low power due to having few samples without co-occurring depression (*k*=3–5).

**Continuous moderators—**Meta-regression analyses for continuous moderators are reported in Table 4. For each measure with sufficient studies to conduct meta-regression analyses, Table 4 provides the regression coefficients for each moderator, with their associated 95% confidence intervals and significance test. Age significantly moderated visuospatial WM composite scores, and marginally moderated shifting composite scores, verbal fluency composite scores, phonemic verbal fluency, and planning composite scores, such that effect sizes were larger for older samples. These effects remained significant or marginal when controlling for gender (visuospatial WM  $z = 2.15$ ,  $p = .032$ ; shifting composite scores  $z = 2.53$ ,  $p = .011$ ; verbal fluency composite  $z = 2.62$ ,  $p = .009$ ; phonemic verbal fluency  $z = 1.97$ ,  $p = .049$ ; planning composite scores  $z = 1.91$ ,  $p = .056$ ). Controlling for medication use, the effect of age remained significant or marginal for visuospatial WM composite scores ( $z = 1.82$ ,  $p = .069$ ), verbal fluency composite scores ( $z = 2.00$ ,  $p = .046$ ),

and planning composite scores  $(z = 1.71, p = .087)$ , while the effect of age became nonsignificant on shifting composite scores  $(z=1.31, p=.189)$  and phonemic verbal fluency  $(z=1.15, p=.250)$ . Controlling for symptom severity, the effect of age remained significant or marginal on visuospatial WM ( $z = 2.09$ ,  $p = .037$ ), shifting composite scores ( $z = 1.70$ ,  $p = .$ 089), and planning composite scores ( $z = 1.97$ ,  $p = .049$ ), while the effects on verbal fluency composite scores ( $z = 1.09$ ,  $p = 0.275$ ) and phonemic verbal fluency ( $z = 1.40$ ,  $p = 0.162$ ) became non-significant.

There was a significant effect of gender on shifting composite scores, and marginal effects of gender on TMT-B and verbal WM composite scores, such that samples with more female participants had worse performance. Controlling for age, the effect of gender remained significant on shifting composite scores ( $z = 2.48$ ,  $p = .013$ ) and verbal WM composite scores ( $z = 2.02$ ,  $p = .044$ ), but became non-significant on TMT-B ( $z = 1.63$ ,  $p = .104$ ). Controlling for medication, the effect of gender remained significant for shifting composite scores ( $z = 2.53$ ,  $p = .011$ ) but became non-significant on TMT-B ( $z = 1.54$ ,  $p = .123$ ) and verbal WM composite scores  $(z = 1.26, p = .207)$ . Controlling for symptom severity, the effect of gender remained significant for shifting composite scores ( $z = 2.25$ ,  $p = .025$ ), but became non-significant for TMT-B ( $z = 1.21$ ,  $p = .228$ ) and verbal WM composite scores ( $z = 1.21$  $= 1.30, p = .192$ .

The percentage of individuals with OCD taking psychotropic medication marginally moderated inhibition composite scores, such that samples with a higher percentage of medicated participants exhibited worse performance. This effect remained marginal after controlling for age ( $z = 1.76$ ,  $p = .078$ ) and gender ( $z = 1.68$ ,  $p = .092$ ), but not symptom severity  $(z = 1.30, p = .194)$ . OCD symptom severity, as assessed by the Y-BOCs, did not significantly moderate any analyses.

# **Discussion**

#### **Evaluating Hypotheses: EF is Broadly Impaired in OCD**

In sum, the current meta-analysis found that in comparison to their healthy peers, individuals with OCD exhibited significantly impaired performance on tasks measuring most aspects of EF, with most effect sizes in the  $d = 0.3-0.5$  range<sup>4</sup>. These effects were not due to failure to match groups on IQ, or to publication bias. The results are consistent with Hypothesis 1, which posits a broad impairment across multiple aspects of EF that may be driven by dysfunction in prefrontal-striatal circuits (e.g., Kuelz et al., 2004b; Menzies, Chamberlain, et

<sup>4</sup>Comparing the results of the current meta-analysis to those of the previous meta-analyses that included some EF measures, effect sizes for individual tasks reported by Shin et al. (2013) were fairly consistent with the current results for analyses in which Shin et al. (2013) included ten or more studies, while analyses with only 5–6 studies both over and under-estimated effects compared to the current study, likely reflecting imprecision in these estimates due to the small number of studies. Comparing the current results to those of Abromovitch et al. (2013) is not straightforward, because they did not report analyses of individual measures, and grouped measures into composites that do not always align with the theoretically-motivated grouping of measures used in the current metaanalysis. For the most comparable composite measures across meta-analyses, planning effect sizes were identical (*d*=0.44), and inhibition effect sizes were somewhat smaller in the current study (*d*=0.37 vs. 0.49). For the less comparable composites, the shifting composite measure effect size in the current study was similar to that of the shifting/fluency/WAIS similarities measure in Abromovitch et al. (2013; *d*=0.50 vs. 0.55), the verbal WM/updating composite (*d*=0.34) was between those for verbal WM (*d*=0.22), and updating  $(d=0.71)$  in the current study, and the spatial WM/updating composite  $(d=0.37)$  was lower than both the visuospatial WM ( $d=0.47$ ) and updating effect sizes in the current study (potentially due to the relatively small number of studies included in the Abromovitch et al. (2013) analysis, and differences in the included tasks).

al., 2008a). The exception was verbal WM maintenance, where task performance for individuals with OCD was comparable to controls. However, this finding is not incompatible with Hypothesis 1, since simple maintenance of information in WM (as opposed to manipulation) is not strongly linked to other aspects of EF (e.g., Engle et al., 1999). Hypothesis 2, which posits a specific impairment in shifting and/or inhibition, was not supported, as effect sizes in other EF domains were equivalent to those for shifting and inhibition. The results are thus consistent with the theory that individuals with OCD have impairments in the unitary component of EF (i.e., common EF), posited to be the ability to actively maintain task goals and use this information to effectively bias lower-level processes (Friedman et al., 2008; Miyake & Friedman, 2012). Although other explanations are also possible (e.g., multiple specific aspects of EF could be independently impaired in OCD), impairment in common EF is the most parsimonious interpretation. It is also possible that individuals with OCD have deficits in common EF as well as processing-specific impairments in shifting and/or updating (recall that there is no inhibition-specific component, e.g., Friedman et al., 2008). Indeed, the largest effect size in the meta-analysis was for updating working memory (n-back), which is believed to depend critically on striatal gating of information into prefrontal cortex (e.g., Chatham et al., 2011; Frank, Loughry, & O'Reilly, 2001; Hazy, Frank, & O'Reilly, 2007). This suggests that individuals with OCD might have specific updating impairments in addition to common EF impairments, since both striatal and prefrontal dysfunction may contribute to deficits on updating tasks. Future research using a latent variable approach is needed to address these possibilities, as discussed in Future Directions.

Hypothesis 3, which posits that apparent EF deficits are due to general motor response slowing, was not supported. The current analysis revealed that individuals with OCD do exhibit significant general motor slowing, and are especially slowed on the TMT-A, which requires both motor speed and sequencing. However, significant impairments were also detected on accuracy measures from self-paced EF tasks, with effect sizes as large or larger than many of the response time tasks. Thus, while individuals with OCD do have slowed responses even on simple general motor speed tasks, these deficits cannot fully account for deficits on EF tasks. (However, as discussed in the introduction, it is impossible to rule out a deficit in general processing speed – as opposed to general motor speed—that could potentially reduce accuracy).

Finally, co-occurring depression does not account for EF deficits in OCD as posited by Hypothesis 4. OCD samples with no depression diagnoses and low levels of depressive symptoms were significantly and equivalently impaired on almost all measures of EF. This raises the question of how EF deficits associated with OCD and MDD (e.g., Snyder, 2012) are related to one another. One possibility is that prefrontal abnormalities that lead to impairments in EF may be transdiagnostic risk factors for psychopathology, including OCD and depressive disorders including MDD (e.g., Nolen-Hoeksema & Watkins, 2011; see below, Relating deficits across cognitive domains and disorders). It is also possible that OCD and MDD have independent effects on EF that might be detected with more sensitive continuous analyses of depressive symptoms, which were not possible here because of the wide variety of depression measures reported in the primary literature.

Effect sizes were largely stable across variation in demographic characteristics of the samples, although there was some evidence for larger deficits for older OCD groups (for shifting, visuospatial WM, verbal fluency and planning). This finding warrants further research, as empirical studies have not investigated age effects. In addition, medication use was associated with larger impairments on inhibition composite scores, and a higher percentage of female participants was associated with larger impairments in shifting and verbal WM. Although these effects were not found for any other measures, they may warrant further empirical study, as some medications may have cognitive side-effects, and one previous study found larger EF impairments for women with OCD on some measures (Mataix-Cols et al., 2006). The fact that symptom severity did not moderate effect sizes suggests that EF impairment may be a stable trait associated with OCD rather than fluctuating with current symptoms. However, this finding must be interpreted with caution given the relatively narrow range of severity levels in the included studies.

#### **Limitations**

There are several limitations in the conclusions that can be drawn from the current metaanalysis, due to limitations in the primary literature. First, co-occurring depression was coded as a categorical variable. This was necessary because the primary literature reports a wide variety of depression measures, which cannot easily be converted into a single continuous measure. The categorical depression measure (no co-occurring depression, versus any amount of co-occurring depression) provides a conservative test demonstrating that EF deficits are present in non-depressed individuals with OCD. However, this categorical measure limits the ability to detect the extent to which co-occurring depression might contribute to larger EF deficits in individuals with OCD. Future research could address this issue in several ways–individual studies could examine correlations between depression and EF performance in samples with OCD, and increased reporting of a common set of depression measures across studies, or psychometric studies to allow conversion of different measures to a common scale, would allow future meta-analyses to use continuous depression measures.

Second, the current meta-analysis is limited in its ability to determine to what extent EF deficits are related to specific OCD symptoms, versus anxiety more broadly. To address this issue, there is a need for increased reporting and analyzing of more detailed information about co-occurring anxiety disorders and anxiety symptoms, as well as more specific sets of OCD symptoms (e.g., compulsions and obsessions separately). Moving towards this more dimensional approach holds promise for uncovering mechanisms of psychopathology that may be obscured by heterogeneous diagnostic categories (e.g., Insel et al., 2010). Finally, as discussed in Future Directions, the current meta-analysis is limited by the types of EF tasks included in the primary literature. Specifically, many of these tasks are too broad to answer fine-grained questions about specific aspects of EF.

#### **Implications for the Frontal-Striatal Model**

Consistent with the EF deficits reviewed here, individuals with OCD have been found to have structural and functional abnormalities in PFC (for reviews see Menzies, Chamberlain, et al., 2008a; Nitschke & Heller, 2005). Earlier versions of the frontal-striatal model posited

a specific deficit in orbitofrontal function (e.g., Graybiel & Rauch, 2000). However, both the results of the current meta-analysis (which found deficits in EF tasks known to depend primarily on other areas of PFC; e.g., Nee et al., 2007; Wager et al., 2004), and more recent versions of the frontal-striatal model based on neuroimaging evidence (Menzies, Chamberlain, et al., 2008a), suggest that function is disrupted in a wider PFC network not limited to orbitofrontal cortex.

Functional neuroimaging during EF tasks has revealed activation differences between individuals with OCD and healthy controls across a wide PFC network, including anterior cingulate (e.g., Fitzgerald et al., 2005; Maltby, Tolin, Worhunsky, O'Keefe, & Kiehl, 2005; Yucel et al., 2007) and dorsolateral and ventrolateral PFC (Maltby et al., 2005; Roth et al., 2007; van den Heuvel, Veltman, Groenewegen, Cath, et al., 2005a), in addition to orbitofrontal cortex (Maltby et al., 2005; Roth et al., 2007). However, both hyperactivation and hypoactivation have been found across studies. Thus, while there is strong evidence for differences in PFC function in individuals with OCD compared to controls, the direction of these differences is unclear, and may depend on task or individual characteristics yet to be differentiated.

## **Future Directions**

Given the compelling evidence that individuals with OCD are impaired on most EF tasks, we would argue that there is no longer a need for further case-control studies of performance on standard neuropsychological measures of EF. Rather, there is now the opportunity to build on the foundation of such previous studies to better understand the specific mechanisms and causal processes contributing to EF deficits in OCD, and to move towards translational applications. To do so, we advocate for (1) better assessment of EF using multiple, specific, measures of different EF components, based on well-established EF models, (2) probing deficits at multiple levels of analysis, (3) investigating how EF deficits are related across disorders, and how EF deficits are related to deficits in other cognitive domains, and (4) using longitudinal, mediational, and behavior genetic approaches to probe possible causal links between EF deficits and OCD.

**More precise assessment of EF deficits—**Future research would benefit from the use of more sensitive and specific measures of each aspect of EF, and by the use of multiple measures. We argue that investigating how specific aspects of OCD are related to specific EF components is critical for elucidating the cognitive, neural, and genetic mechanisms involved. Many previous studies, including many in the current meta-analysis, have used EF measures that are too broad to answer these fine-grained questions. For example, verbal fluency tasks have been a perennial favorite for assessing EF. However, they and other complex neuropsychological tests tap a wide variety of cognitive processes, including not only aspects of EF but also non-executive abilities (Miyake, et al., 2000). As a result, impairments on such measures are difficult to interpret. This concern can be addressed by using tasks designed to specifically place demands on one aspect of EF, while keeping other demands minimal (e.g., Aron, 2008; Miyake, et al., 2000).

In addition, including multiple measures of each aspect of EF would allow construction of latent or composite measures, which greatly increase construct validity and power (e.g., Miyake, et al., 2000). While the current meta-analysis found broad impairment in EF, which is most parsimoniously explained by a deficit in common EF, latent variable approaches are also needed to answer the key question of whether impairments across multiple aspects of EF are due to impairment of the common EF component, or multiple separate impairments, or both. For example, such a design could address the question of whether deficits on EF tasks are fully accounted for by deficits in common EF (i.e., once common EF is accounted for, there is no longer any evidence of process-specific impairments that are applicable only to individual aspects of EF), or whether there are processing-specific deficits not fully accounted for by common EF impairments (e.g., an updating-specific impairment related to striatal dysfunction). Research designs needed to test these models pose logistical challenges for research in clinical populations, as they require longer testing sessions to collect multiple measures and large (200+) sample sizes. However, overcoming these challenges– for example by administering tasks in several shorter sessions to reduce fatigue and collaborating across sites to increase sample size– may pay large dividends for determining the specific cognitive and neural mechanisms affected in individuals with OCD.

**Understanding deficits at multiple levels of analysis—**While the current metaanalysis demonstrates impairments in EF at the level of behavioral task performance, future research is needed to investigate the specific neural mechanisms contributing to EF deficits in OCD. For example, human genetic studies and animal models have suggested a role for abnormalities in the dopamine, serotonin and glutamate systems in OCD (e.g., Albelda & Joel, 2012; Pauls, 2008; Rolls, Loh, & Deco, 2008). A promising area of research is the integration of what is known about the role of these neurotransmitter systems in PFC networks with behavioral and neuroimaging evidence for EF impairments in individuals with OCD, building a more detailed model of the neurobiology of OCD that spans multiple levels of analysis.

**Relating deficits across cognitive domains and disorders—**While the current meta-analysis focused on EF, cognitive deficits associated with OCD are not restricted to EF, with meta-analytic evidence for deficits of a similar magnitude in processing speed, episodic memory, and attention (Abromovitch et al., 2013; Shin et al., 2013; Woods, Vevea, Chambless, & Bayen, 2002). An important question is thus how these deficits are related to one another. Some may be independent, for example, general motor slowing may be related to dysfunction in premotor-striatal loops, which are adjacent to, but separate from, the PFCstriatal loops involved in EF (e.g., Haber & Calzavara, 2009). In other cases, a common deficit may lead to impairments across domains. For example, some have argued that poor performance on memory tasks by individuals with OCD is largely attributable to EF deficits, which impact the ability to generate and implement organizational strategies (e.g., grouping words semantically) during encoding and retrieval (e.g., Olley et al., 2007). To test these possibilities, future empirical research is needed to test relations among performance in different domains (e.g., whether memory impairments are fully mediated by EF impairments).

In addition, OCD is far from the only psychiatric disorder associated with EF deficits, with similar effect sizes to those in OCD for MDD (*d* = 0.3–0.7; e.g., Rock, Roiser, Riedel, & Blackwell, 2013; Snyder, 2013) and attention deficit hyperactivity disorder (*d* = 0.3–0.7; e.g., Frazier, Demaree, & Youngstrom, 2005; Walshaw, Alloy, & Sabb, 2010) and somewhat larger deficits in schizophrenia (*d* = 0.7–1.3; e.g., Forbes, Carrick, McIntosh, & Lawrie, 2009; Meshalam-Gately, Giuliano, Goff, Faraone, & Seidman, 2009) and bipolar disorders (*d* = 0.4–0.8; e.g., Kurtz & Gerraty, 2009; Mann-Wrobel, Carreno, & Dickinson, 2011). Hence a second important question is the extent to which EF deficits are shared, or differ, across disorders.

Importantly, while effect sizes differ somewhat across disorders, the same broad pattern of impairment across multiple aspects of EF is found in each, suggesting that PFC abnormalities that lead to impairments in EF may be transdiagnostic risk factors for psychopathology (e.g., Buckholtz & Meyer-Lindenberg, 2012; Nolen-Hoeksema & Watkins, 2011). This general vulnerability may combine with unique genetic, neurobiological, and environmental factors to produce divergent trajectories, leading to different disorders. However, when more detailed measures at multiple levels of analysis are considered in some cases these shared behavioral deficits may arise from distinct neural mechanisms (e.g., perturbations in different neurotransmitter systems; e.g., Gigante et al., 2012; Luykx et al., 2012). Future studies that systematically investigate relations between EF impairments, risk factors, and psychopathology are needed to refine understanding of how such broad EF deficits arise across disorders, and which aspects of these deficits are transdiagnostic versus disorder specific.

**Possible Causal Links between OCD and EF—Previously research has left the** question of how OCD and EF impairments are causally related largely unaddressed. For example, the vast majority of studies, including in the current meta-analysis, have used cross-sectional case-control designs, that do not provide any information about temporal precedence that could contribute to understanding causal links. There are (at least) three non-mutually-exclusive possibilities for these causal relationships. First, individual differences in neurobiology could both confer risk for OCD and lead to EF deficits. Specifically, alterations in PFC function may cause impaired EF as well as contribute to OCD symptoms, such as perseverative behaviors and the inability to inhibit compulsions (e.g., Menzies, Chamberlain, et al., 2008a). Supporting this model, differences in PFC and impairments in EF are present in non-affected relatives of individuals with OCD (e.g., Cavedini, Zorzi, Piccinni, Cavallini, & Bellodi, 2010; Menzies, Williams, et al., 2008b; Rajender et al., 2011), suggesting that they may represent endophenotypes for OCD. This possibility can be further tested with future behavior genetic research testing for shared genetic influence on EF and OCD, as well as by longitudinal studies tracking children at risk for OCD (e.g., due to family history) before onset.

The hypothesis that PFC and basal ganglia dysfunction may play a causal role in OCD is also supported by evidence that acquired brain damage to these areas can cause obsessions and compulsions along with EF impairments (e.g., see Coetzer, 2004 for review). Experimentally induced lesions to PFC and basal ganglia structures in animal models thus present an obvious target for testing causal models of OCD. However, such animal models

have been difficult to develop and validate, potentially due to the large differences in prefrontal structure and function between rodents and humans (e.g., see Albelda & Joel, 2012 for review). As an alternative approach, further research that carefully evaluates the location of lesions associated with the onset of OCD symptoms (e.g., through voxel based lesion symptom mapping), and associated EF deficits, may help to shed light on possible causal associations between dysfunction in particular brain areas and OCD.

Second, OCD could directly cause EF deficits. For example, OCD-related intrusive thoughts may consume cognitive resources and interfere with the ability to maintain task goals. This possibility could be tested experimentally (e.g., by triggering intrusive thoughts prior to tasks). However, this direct effect is unlikely to be the only causal path, as there is some evidence that EF deficits remain stable after OCD symptoms have remitted (Bannon, Gonsalvez, Croft, & Boyce, 2006; Roh et al., 2005). Further research with individuals in remission is needed to confirm this possibility. If true, this finding suggests that even when treatment is considered successful from the perspective of eliminating affective symptoms and behaviors, persistent EF deficits may continue to undermine daily functioning and quality of life. It may therefore be useful to make deficits in EF as a focus for intervention a topic for future translational research.

Third, poor EF could contribute to the development or maintenance of OCD. Current models of anxiety suggest individuals with poor emotion regulation abilities engage in frequent and excessive worrying (Borkovec & Roemer, 1995; Mennin, Heimberg, & Turk, 2002). Meanwhile, better EF is linked to more effective emotion regulation strategies (e.g., Ochsner & Gross, 2005). Thus, it may be that poor EF leads individuals to develop alternative, but less adaptive or efficient, strategies for coping with emotional challenges that ultimately increase levels of OCD symptoms. For example, the anxiety reduction hypothesis posits that compulsions are a maladaptive strategy for reducing anxiety associated with intrusive thoughts, and are thus reinforced through avoidance learning (e.g., see Clark, 2004 for review). Thus, reduced ability to use EF to engage more adaptive anxiety-regulation strategies (e.g., shifting attention away from the anxiety-provoking thought) could contribute to the development or maintenance of compulsive behaviors. Because this path is speculative, longitudinal research is necessary to explore EF, emotion regulation, and the onset of OCD or escalation of OCD symptoms.

Understanding causal links between EF deficits and OCD will be critical for developing strategies for prevention and remediation. For example, if EF deficits precede and contribute to the development of OCD, those at risk (e.g., due to family history) may benefit from early intervention to train EF or teach compensatory strategies. In addition, regardless of the initial direction of the causal arrow, poor EF may reduce the effectiveness of therapeutic interventions. For example, deficits in EF could interfere with Cognitive Behavioral Therapy (e.g., Mohlman & Gorman, 2005), particularly with techniques such as Exposure and Response Prevention, which requires top-down processes such as shifting (e.g., from elaboration of automatic, catastrophic thoughts to metacognitive awareness of thoughts), inhibitory control (e.g., resisting engaging in compulsive behaviors), and updating (e.g., replacing maladaptive beliefs and behaviors with adaptive ones). EF deficits may therefore present a major barrier to successful completion of treatment.

#### **Conclusions**

In sum, OCD is associated with broad EF impairment, and these deficits cannot be accounted for by co-occurring depression or general motor slowing. These results are consistent with theories that posit that prefrontal dysfunction is a contributing factor in OCD, but additional research is needed to determine the causal links between EF impairments and OCD and build a more detailed model of the neurobiology of these impairments. A better understanding of when and how EF impairments arise for individuals with OCD may have important implications for treatment, such as pharmacological interventions targeting specific aspects of prefrontal function, or training programs to improve EF or teach compensatory strategies to mitigate the effects of EF impairments. Given the centrality of EF to our ability to successfully navigate daily life, such research has the potential to improve outcomes for many individuals affected by OCD.

# **Supplementary Material**

Refer to Web version on PubMed Central for supplementary material.

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## **Figure 1.**

Weighted mean effect sizes for all analyses. Error bars are 95% confidence intervals. Measures for which the lower error bar does not pass the vertical line are significantly greater than 0. Compared to healthy control participants, individuals with OCD are significantly impaired on most EF tasks. Executive function (EF) composite measures are indicated with diamond symbols, and individual measures within each EF component are indicated by circle symbols in the same color. Black circles indicate non-EF comparison measures. Comp. = Composite score; WCST = Wisconsin Card Sorting Test; OAT/DAT = Object alternation task/delayed alternation task; TMT–B = Trail Making Test Part B; TMT– A =Trail Making Test Part A; ID/ED = Intradimensional/Extradimensional; WM = working memory; DMTS = delayed-match-to-sample; VF = verbal fluency; TOL/TOH = Tower of London, Tower of Hanoi.



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 $\cdot$  : EF Tasks Included in the Meta-Analysis J  $\mathbf{M}$ ś, نة **ال**موادية. Ė Ŕ



**Construct Task Description Outcome Measure(s) # Of Studies**

Description

 $\rm{Task}$ 

 $$ 

# Of Studies

Outcome Measure(s)



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Preferred measure if reported.

 $^\wedge$  Analyzed individually as well as part of composite measure. Analyzed individually as well as part of composite measure.



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**Table 2**

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alternation task; TMT–B = Trail Making Test Part B; TMT–A =Trail Making Test Part A; ID/ED = Intradimensional/Extradimensional; WM = working memory; DMTS = delayed-match-to-sample; VF = Q. werbal fluency; TOL/TOH = Tower of London, Tower of Hanoi. verbal fluency; TOL/TOH = Tower of London, Tower of Hanoi.

 $a_{\text{One}}$  outlier excluded (d=2.66). With this outlier included the weighted mean effect size is d=0.53. *a*One outlier excluded (*d*=2.66). With this outlier included the weighted mean effect size is *d*=0.53.

 $b_4$  One outlier excluded (d=2.66). With this outlier included the weighted mean effect size is d=0.47. *b*a One outlier excluded (*d*=2.66). With this outlier included the weighted mean effect size is *d*=0.47.

 $^{6}$ One outlier excluded (d=1.96). With this outlier included the weighted mean effect size is d=0.14.  $c^{\prime}$ One outlier excluded (*d*=1.96). With this outlier included the weighted mean effect size is *d*=0.14.

 $d_{\text{One outlier excluded (}d = -0.58)}$ . With this outlier included the weighted mean effect size is  $d=0.37$ *d*One outlier excluded (*d* =−0.58). With this outlier included the weighted mean effect size is *d*=.0.37

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*\**

significant  $(p < .05)$ 

 Author Manuscript**Author Manuscript**  **Table 3**

Depression Moderator ANOVA Analyses Depression Moderator ANOVA Analyses





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Note. d= weighted mean Cohen's d effect size; CI= confidence interval; LL= lower limit; UL = upper limit; v = random-effects variance component; Q within= within-group (residual) heterogeneity; Q *Q* within= within-group (residual) heterogeneity; *Q*  Making Test Part B; TMT-A =Trail Making Test Part A; ID/ED = Intradimensional/Extradimensional; WM = working memory; DMTS = delayed-match-to-sample; VF = verbal fluency; TOL/TOH = Making Test Part B; TMT–A =Trail Making Test Part A; ID/ED = Intradimensional/Extradimensional; WM = working memory; DMTS = delayed-match-to-sample; VF = verbal fluency; TOL/TOH = between = between-group (moderator) heterogeneity; Comp. = composite score; WCST = Wisconsin Card Sorting Test; OAT/DAT = Object alternation task/delayed alternation task; TMT-B = Trail between = between-group (moderator) heterogeneity; Comp. = composite score; WCST = Wisconsin Card Sorting Test; OAT/DAT = Object alternation task/delayed alternation task; TMT–B = Trail *Note.* d= weighted mean Cohen's *d* effect size; CI= confidence interval; LL*=* lower limit; UL = upper limit; *v* = random-effects variance component; Tower of London, Tower of Hanoi. Tower of London, Tower of Hanoi.

Depression possible = average depressive symptom questionnaire scores in clinical range and/or individuals with diagnosis of any depressive disorder. Depression absent = average depressive symptom *1*Depression possible = average depressive symptom questionnaire scores in clinical range and/or individuals with diagnosis of any depressive disorder. Depression absent = average depressive symptom questionnaire scores below clinical range and no patients with a diagnosed depressive disorder. questionnaire scores below clinical range *and* no patients with a diagnosed depressive disorder.

*\** Significant (*p*<.05). *#* Marginal (*p*<.10).

Moderator regression analyses Moderator regression analyses



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*Clin Psychol Sci*. Author manuscript; available in PMC 2016 March 01.

*Note.*  $d$ = weighted mean Cohen's d effect size; CI= confidence interval; LL= lower limit; UL = upper limit; v = random-effects variance component; Q within = within-group (residual) heterogeneity; Q model = moderator he *Note. d*= weighted mean Cohen's d effect size; CI= confidence interval; LL= lower limit; UL = upper limit; v = random-effects variance component; Q within = within-group (residual) heterogeneity; Q model = moderator heterogeneity; Comp. = composite score; WCST = Wisconsin Card Sorting Test; TMT–B = Trail Making Test Part B; TMT–A =Trail Making Test Part A; VF = verbal fluency.

Gender 1992 −0.002 −0.0002 0.000 0.000 0.000 0.000 0.0000 0.0000 0.0000 0.001 0.001 0.001 0.001 0.001 0.001 0.0 Medication −0.156 −0.000 −0.001 0.000 0.000 −0.80 .426 26 0.04 0.63 (1) .426 24.55 (24) .430

.430

24.55 (24)

426

 $0.63(1)$ 

 $0.04$ 

 $26$ 

 $.426$ 

 $-0.80$ 

 $0.000$ 

 $0.000$ 

 $-0.001$ 

 $-0.000$ 

 $-0.156$ 

Medication