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Efficacy of Antibiotic Prophylaxis in Children with Vesicoureteral Reflux: Systematic Review and Meta-Analysis

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Abstract

Purpose—Controversy exists regarding the use of continuous antibiotic prophylaxis vs observation in the management of children with vesicoureteral reflux. The reported effectiveness of continuous antibiotic prophylaxis in children with reflux varies widely. We determined whether the aggregated evidence supports use of continuous antibiotic prophylaxis in children with vesicoureteral reflux.

Materials and Methods—We searched the Cochrane Controlled Trials Register, clinicaltrials.gov, MEDLINE[®], EMBASE[®], Google Scholar and recently presented meeting abstracts for reports in any language. Bibliographies of included studies were then hand searched for any missed articles. The study protocol was prospectively registered at PROSPERO (No. CRD42014009639). Reports were assessed and data abstracted in duplicate, with differences resolved by consensus. Risk of bias was assessed using standardized instruments.

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[†]Financial interest and/or other relationship with Eli Lilly & Co.

Results—We identified 1,547 studies, of which 8 are included in the metaanalysis. Pooled results demonstrated that continuous antibiotic prophylaxis significantly reduced the risk of recurrent febrile or symptomatic urinary tract infection (pooled OR 0.63, 95% CI 0.42–0.96) but, if urinary tract infection occurred, increased the risk of antibiotic resistant organism (pooled OR 8.75, 95% CI 3.52–21.73). A decrease in new renal scarring was not associated with continuous antibiotic prophylaxis use. Adverse events were similar between the 2 groups. Significant heterogeneity existed between studies (I² 50%, p = 0.03), specifically between those trials with significant risk of bias (eg unclear protocol descriptions and/or lack of blinding).

Conclusions—Compared to no treatment, continuous antibiotic prophylaxis significantly reduced the risk of febrile and symptomatic urinary tract infections in children with vesicoureteral reflux, although it increased the risk of infection due to antibiotic resistant bacteria. Continuous antibiotic prophylaxis did not significantly impact the occurrence of new renal scarring or reported adverse events.

Keywords

antibiotic prophylaxis; meta-analysis; pediatrics; review; vesico-ureteral reflux

Primary vesicoureteral reflux is a common condition that is present in 1% to 10% of all children in the United States.¹ In the setting of a febrile urinary tract infection reflux is reportedly present in a third of children.² In affected children reflux is associated with an increased risk of recurrent pyelonephritis and, hence, an increased risk of renal scarring.^{2,3} Typical interventions for children with reflux include antireflux surgery (endoscopic, laparoscopic or open) and continuous antibiotic prophylaxis. The purpose of continuous antibiotic prophylaxis is to keep the urine "sterile" so that the risk of retrograde renal infection will be decreased. Since a significant proportion of reflux cases will spontaneously resolve with time,^{4,5} many authors recommend a conservative approach, ie continuous antibiotic prophylaxis, as the initial management option in children, reserving surgical intervention for those in whom continuous antibiotic prophylaxis is ineffective at preventing urinary tract infection.

Significant controversy and treatment related variability still exist regarding VUR management.⁶ In particular the effectiveness of CAP at decreasing infections in children has recently been called into question.⁷ Recent RCTs investigating the effect of CAP on prevention of urinary tract infection in children with reflux have shown conflicting results.^{8–15} Further clouding this picture is the fact that, as noted in a recent Cochrane Review,¹⁶ many of these RCTs have significant design or reporting flaws that limit their impact. However, with the recent publication of the National Institutes of Health sponsored RIVUR trial,¹² it is unclear whether the accumulated data on CAP have shifted enough to affect treatment recommendations. We evaluated the accumulated literature on the effectiveness of CAP for children with VUR and determined the extent to which reported success rates for CAP have been influenced by underlying patient or study level factors.

Patients and Methods

Search Strategy

We searched MEDLINE, EMBASE, Cochrane Controlled Trials Register, www.clinicaltrials.gov and Google Scholar electronic databases for studies published between January 2010 and May 2014 in any language based on PRISMA (Preferred Reporting Items for Systematic Reviews and Meta-Analyses) guidelines.¹⁷ We additionally evaluated all studies previously included in systematic reviews of this topic.¹⁶ This date range was chosen to provide a contemporary selection of series. We used the search terms "vesicoureteral reflux," "vesicoureteric reflux," "vesico-ureteral reflux" and "vesico-ureteric reflux" (see Appendix).

Reference lists of included studies were manually screened for any additional series. We also manually searched for unpublished abstracts presented at relevant scientific meetings, including meetings of the American Urological Association, Society for Pediatric Urology, American Academy of Pediatrics Section on Urology, Pediatric Academic Societies, World Congress of Endourology, Société Internationale d'Urologie, International Pediatric Nephrology Association and the European Association of Urology. Before the formal literature search the protocol was prospectively registered at PROSPERO (No. CRD42014009639).

Selection Criteria

Inclusion criteria consisted of age 18 years or younger and history of VUR treated with CAP. Study patients were compared to individuals 18 years or younger with VUR undergoing no treatment or treatment with placebo (controls).

We chose to include only RCTs that described the number of patients treated as well as the fraction in whom treatment was successful. No study was excluded based on method of analysis, definition of success, language of publication, perceived quality or susceptibility to bias. In cases of ambiguity or where study reporting made evaluation difficult we attempted to err on the side of inclusiveness.

Data Abstraction

Two reviewers (H-HSW, JCR) independently examined all study abstracts in duplicate, with disagreements resolved by consensus. Full text articles appearing to meet selection criteria were reviewed, and study data were abstracted in the same manner. Reports published in languages other than English were translated by study authors fluent in that language and/or by institutional translation staff. Abstracted data included patient level factors (age, VUR grade, gender, UTI rates, new renal scarring on nuclear scan, adverse events) and study level factors (country of origin, year of publication, conflict of interest disclosure, study funding). COI was identified by publication of disclosure.¹⁸

Missing Data and Author Contact

In cases of missing or unreported data we attempted to contact corresponding study authors by email. Data updates/confirmations were received from 5 of 8 included studies.^{9,10,12–14}

In cases where authors were unable to provide the missing information we excluded the study from only those analyses requiring the missing data.

Outcome Assessment

Our primary outcome was the odds ratio of having febrile or symptomatic UTI. Secondary outcomes included new renal scarring, antibiotic resistance and any adverse effects related to CAP.

Risk of Bias Assessment

Bias assessment was performed by 2 study authors (H-HSW, JCR) using the Cochrane Collaboration checklist.¹⁹ Differences were resolved by consensus discussion. Funnel plots were visually assessed for evidence of publication bias. A stratified subgroup analysis was performed between studies deemed to have a high risk and those with a low risk of bias.

Statistical Methods

Descriptive statistical analyses were performed using OR and 95% confidence intervals, as appropriate. For univariate pooling DerSimonian-Laird random effects models were constructed.²⁰ Study heterogeneity was assessed using the Higgins-Thompson I² method.²¹ Given the anticipated small number of eligible studies, metaregression was a priori planned as a series of bivariate models to evaluate whether study or patient level factors were associated with study outcomes. All statistical analyses were performed using Stata®/SE, version 11.0 and Review Manager, version 5.2.9 (Nordic Cochrane Centre, Cochrane Collaboration, Copenhagen, Denmark).

Results

Search Results

A total of 1,547 publications were identified using our search criteria. Of these studies 16 (67 publications) were selected for full review. Eight studies met all criteria for inclusion, and were included in the pooled qualitative review and the quantitative meta-analysis (supplementary fig. 1, http://jurology.com/).

Description of Studies

Of the 8 included studies all were randomized controlled trials.^{8–15} Thus far, only 1 study has been published as an abstract.⁹ A total of 1,594 patients were included in the final metaanalysis, nearly all of whom were diagnosed with VUR after a UTI (see table). Mean patient age ranged from 8.6 to 21.3 months (median 12 to 24). Female patients made up 37% to 93% of the study cohorts. VUR grades also varied between studies, with some including low grade only, some high grade only and some grade I to IV disease.^{10–15,22} Followup ranged from 1 to 3 years. Included study populations were from the United States, Italy, France, Sweden, Norway, Australia, Chile and Spain. No study author had a significant COI, although COI data were not reported for all studies.^{9,11,14}

Risk of Bias in Included Studies

The risk of bias graph is illustrated in supplementary figure 2 (http://jurology.com/), and the risk of bias summary is illustrated in figure 1. The major bias of included studies was lack of adequate blinding. Only 2 studies provided a placebo that was similar to the active antibiotic,^{10,12} while the remainder provided no treatment in the control arm. Therefore, those 6 studies were deemed to have high potential for performance and detection biases. Three studies failed to provide adequate detail regarding patient randomization or allocation and/or outcome reporting.^{9,11,15} These series were deemed as having an unclear risk in those categories. Lastly an additional source of bias was from the source of urine samples. Two of the studies included bagged urine samples,^{13,15,22} which have an increased risk of contamination and falsepositive culture compared to catheterized specimens.²³ There was little visual evidence of publication bias noted on funnel plots (fig. 2 and supplementary fig. 3, http://jurology.com/).

Effects of Interventions

CAP significantly reduced the risk of febrile or symptomatic UTI in children with reflux (pooled OR 0.63, 95% CI 0.42–0.96, p = 0.03, fig. 3). When we stratified by the susceptibility of each study to bias, those studies at lower risk of bias revealed an even more significant protective effect from CAP (pooled OR 0.51, 95% CI 0.35–0.73, p = 0.0003). There was no statistically significant impact of CAP on febrile or symptomatic UTI in studies with a higher risk of bias (p = 0.34).

Regarding the development of antibiotic resistance, CAP was associated with an increased risk of resistant bacteria (pooled OR 8.75, 95% CI 3.52–21.73, p <0.0001). This risk remained significant after stratifying between studies based on risk of bias (supplementary fig. 4, http://jurology.com/). CAP failed to demonstrate any significant impact on new renal scarring or antibiotic related adverse events. Bivariate metaregression models showed no significant association between the likelihood of febrile or symptomatic UTI and duration of followup (p = 0.5), gender (p = 0.1), placebo (vs no treatment, p = 0.5) or study year (p = 0.1).

Discussion

In this study pooled RCT results reveal that CAP was associated with a 37% decrease in the odds of febrile or symptomatic urinary tract infection in children with reflux (pooled OR 0.63). This effect estimate differs from the most recent Cochrane Review,¹⁶ largely due to the completion of the recent RIVUR trial. The fact that a subset analysis of the only 2 RCTs noted to have a low risk of bias^{10,12} demonstrated an even stronger effect of CAP (49% decrease, pooled OR 0.51) would seem to lend further support to the argument that CAP does, indeed, have a protective effect against the development of UTI in children with reflux.

Despite this positive effect, rates of new renal scarring were not significantly associated with CAP. There are many potential reasons for this finding, most of which ultimately depend on the impressively low rate of renal scarring in all included studies. First, the rather short-term

followup of all included studies may preclude detection of renal scarring formation. In addition, the "healthy volunteer" phenomenon may have a role, ie patients who are enrolled in a RCT may be less likely to have renal scars develop due to parental/provider vigilance and care at the first signs of UTI. Furthermore, most of these studies, including the PRIVENT (Prevention of Recurrent Urinary Tract Infection in Children with Vesicoureteric Reflux and Normal Renal Tracts) and RIVUR trials, contained patients recruited after the first or second UTI. Given that the risk of renal scarring after only 1 to 2 UTIs is quite low,²⁴ it is not surprising that renal scarring events were uncommon. As the risk of renal scarring has been found to increase exponentially with an increasing number of UTIs, it stands to reason that if CAP prevents UTI, it should also prevent renal scarring. However, this supposition cannot be proved based on the data in this metaanalysis. It is noteworthy that recent studies have shown higher rates of renal scarring (approximately 15%),^{25,26} suggesting that these RCT participants may have had lower renal scar rates than the general population.

Although long-term CAP was associated with a higher rate of UTIs due to antibiotic resistant bacteria, there was no increase in associated adverse events reported in the CAP group compared to controls. Given its overall protective effect against UTI, coupled with the low rate of adverse events, CAP seems to be a reasonable management option to prevent recurrent febrile or symptomatic UTI in children with reflux. These benefits should be weighed against the obvious trade-off in terms of increased antibiotic resistance. Some authors have reported successful discontinuation of antibiotic prophylaxis in children with persistent reflux without subsequent UTI,^{27,28} implying that there may be certain populations at lower risk for UTI in whom the benefit of long-term CAP may not be adequate to justify the risks. Unfortunately, to our knowledge it is not yet feasible to reliably identify such a population.

One potential impact of our results involves the use of radiological imaging, particularly voiding cystourethrography, in children after a first febrile UTI. Given that the recent AAP guidelines were formulated based on data suggesting a lack of protective effect of CAP against UTIs,²³ our findings may prompt reconsideration of those guidelines. We would tend to agree with the RIVUR authors, who noted, "Decision analysis and cost-effectiveness analysis may help to clarify the clinical and financial trade-offs, which may help clinicians and families reach more informed decisions about the advisability of imaging in young children."¹²

This series should be interpreted in light of its limitations. As all of these studies included patients diagnosed with VUR after an initial UTI, these results may not apply to children diagnosed with VUR for other reasons (eg prenatal hydronephrosis and sibling VUR). As with any systematic review, our analyses were limited by the available data from the included studies. As seen in supplementary figure 2, many published studies of CAP were judged to have a risk or potential risk of significant bias. Selection bias could not be ruled out in studies lacking a clear description of recruitment, exclusion, randomization and/or allocation methodology. Many of the included studies were not blinded, which may introduce performance and detection bias. Thus, as with any meta-analysis, we must rely on

a potentially imperfect data set. Nevertheless, to our knowledge this data set also represents the best data currently available.

Additionally it is noteworthy that we conducted our risk analyses at the study level instead of at the patient level. This is a form of ecological bias, where the aggregate results across studies are different from the raw results within studies. Similarly our analysis of author factors such as COIs are highly dependent on the veracity of author disclosures. While we have no reason to suspect that groups were not wholly forthright in their disclosures, it is noteworthy that the omission or false report of COI in even a small number of groups could have significantly impacted our findings.

Finally, we noted fairly significant heterogeneity in our pooled analysis, especially in the outcome group of febrile or symptomatic UTI (I^2 55%, p = 0.04, fig. 3). This heterogeneity was principally driven by RCTs judged to have a higher risk of bias. Not surprisingly, in addition to the various biases due to study design and execution, many of the potentially important cohort characteristics such as age, VUR grade, CAP regimen and definition of outcomes also varied greatly among these studies. However, the 2 RCTs with a low risk of bias revealed consistently minimal heterogeneity (I^2 0%) across all outcome measures.

Conclusions

Compared to no treatment or placebo, CAP significantly reduced the risk of febrile and symptomatic urinary tract infections in children with VUR, although it increased the risk of infection due to antibiotic resistant bacteria. The protective effect of CAP was more prominent in studies deemed to have a low risk of bias. CAP did not significantly impact the rate of new renal scarring or reported treatment related adverse events.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

Appendix

Search Terms and Strategy

Medline

("vesico-ureteral reflux" OR "vesicoureteral reflux" OR "vesico-ureteric reflux" OR "vesicoureteric reflux" OR "kidney reflux") AND " "2010/01/01"[PDat] : "2014/05/09" [PDat]"

Embase

(vesico-ureteral reflux)/exp AND (2010:py OR 2011:py OR 2012:py OR 2013:py OR 2014:py) AND 'randomized controlled trial (topic)'/de

Page 7

Cochrane Controlled Trials Register

(("vesicoureteral reflux" (title, abstract, keywords) OR "vesicoureteric reflux" (title, abstract, keywords) OR "vesico-ureteral reflux" (title, abstract, keywords) OR "vesico-ureteric reflux" (title, abstract, keywords)) and limited to 2010-2014

Clinicaltrials.gov

"vesico-ureteral reflux" OR "vesicoureteral reflux" OR "vesico-ureteric reflux" OR "vesicoureteric reflux"

Google Scholar

("vesico-ureteral reflux" OR "vesico-ureteric reflux" OR "vesico-ureteric reflux" OR "vesicoureteric reflux") AND "randomized controlled trial"

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Abbreviations and Acronyms

- AAP American Academy of Pediatrics
- CAP continuous antibiotic prophylaxis
- COI conflict of interest
- **RCT** randomized controlled trial
- **RIVUR** Randomized Intervention for Children with Vesicoureteral Reflux
- UTI urinary tract infection
- **VUR** vesicoureteral reflux

	Random sequence generation (selection bias)	Allocation concealment (selection bias)	Blinding of participants and personnel (performance bias)	Blinding of outcome assessment (detection bias)	Incomplete outcome data (attrition bias)	Selective reporting (reporting bias)	Other bias
Craig 2002	?	?	•	•	•	•	?
Garin 2006	?	•	•	•	?	?	?
Montini 2008	•	•	•	•	•	•	•
Pennesi 2006	•	•	•	•	•	•	•
PRIVENT 2009	•	•	•	•	•	•	•
Reddy 1997	?	?	?	?	•	?	?
RIVUR 2014	•	•	•	•	•	•	•
Roussey-Kesler 2008	?	?	•	•	?	?	•
Swedish reflux trial	•	•	•	•	•	•	•

Figure 1.

Risk of bias detail. Green circles represent positive risk of bias. Red circles indicate negative risk of bias. Yellow circles signify unknown risk of bias.

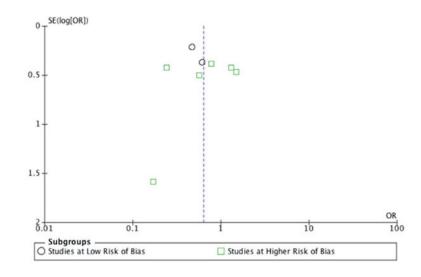


Figure 2.

Funnel plot of included studies for UTI. Circles indicate series with low risk of bias. Squares represent studies with higher risk of bias.

	CAL	Þ	Cont	rol		Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	IV, Random, 95% CI	Year	IV, Random, 95% CI
1.1.1 Studies at Low R	Risk of Bia	as						
PRIVENT 2009	14	122	21	121	14.8%	0.62 [0.30, 1.28]	2009	
RIVUR 2014	39	302	72	302	21.2%	0.47 [0.31, 0.73]	2014	
Subtotal (95% CI)		424		423	35.9%	0.51 [0.35, 0.73]		◆
Total events	53		93					
Heterogeneity: $Tau^2 = 0$	0.00; Chi ²	= 0.38	8, $df = 1$	(P = 0)	.54); I ² =	0%		
Test for overall effect: 2	Z = 3.61 (P = 0.0	0003)					
1.1.3 Studies at Highe	r Risk of	Bias						
Craig 2002	0	21	2	20	1.7%	0.17 [0.01, 3.82]	2002	•
Pennesi 2006	18	50	15	50	12.9%	1.31 [0.57, 3.03]	2006	
Garin 2006	13	55	10	58	11.6%	1.49 [0.59, 3.74]	2006	
Montini 2008	10	82	9	46	10.7%	0.57 [0.21, 1.53]	2008	
Roussey-Kesler 2008	13	103	19	122	14.2%	0.78 [0.37, 1.67]	2008	
Swedish reflux trial	10	69	28	68	13.1%	0.24 [0.11, 0.55]	2010	
Subtotal (95% CI)		380		364	64.1%	0.69 [0.37, 1.29]		◆
Total events	64		83					
Heterogeneity: $Tau^2 = 0$	0.33; Chi ²	= 12.	10, df =	5 (P =	0.03); I ²	= 59%		
Test for overall effect: 2	Z = 1.15 (P = 0.2	25)					
Total (95% CI)		804		787	100.0%	0.63 [0.42, 0.96]		•
Total events	117		176					
Heterogeneity: $Tau^2 = 0$	0.16; Chi ²	= 13.9	99, df =	7 (P =	0.05); I ²	= 50%		
Test for overall effect: 2								0.01 0.1 1 10 10 Favors CAP Favors Control
Test for subgroup diffe				1 (P =	0.40), I ²	= 0%		Favors CAP Favors Control

Figure 3.

Forest plot of febrile or symptomatic UTI. Squares indicate odds ratios. Diamonds represent summary measures (center of diamond) and associated confidence intervals (lateral tips of diamond).

	RIVUR Trial ¹²	Swedish Reflux Trial ⁸	PRIVENT Trial ¹⁰	Montini et al ¹³	Roussey-Kesler et al ¹⁵	Garin et al ¹¹	Pennesi et al ¹⁴	Craig et al ⁹
No. pts	607	137	243	128	225	113	100	41
Followup (yrs)	2	2	1	1	1.5	1	2	3
COI	Grants (National Institutes of Health)	Grants (Sweden)	Grants (Australia)	Grants (Italy)	Grants (France)	Not reported	Not reported	Not reported
Age at randomization	Median 12 mos (range 2 – 72)	Mean 21.3 mos (range 12 -24)	0-18 Yrs	2 –84 Mos	Mean 11.2 mos (range 1 – 36)	Median 24 mos (range 3 mose12 yrs)	Mean 8.6 mos (range 0 – 30)	0 –3 Mos
% Female	92	93	62	Not reported	69	81	52	37
VUR grades included	I –IV	VI- III	IV	III-II	III- I	IV -V	II –IV	Not reported
Intervention group:								
CAP	Trimethoprim- sulfamethoxazole	Trimethoprim, nitrofurantoin, cefadroxil	Trimethoprim- sulfamethoxazole	Trimethoprim- sulfamethoxazole, amoxicillin/clavulanic acid	Trimethoprim- sulfamethoxazole	Trimethoprim- sulfamethoxazole, nitrofurantoin	Trimethoprim- sulfamethoxazole	Trimethoprim- sulfamethoxazole
No. febrile or symptomatic UTI/total No.	39/302	10/68	14/122	10/82	13/103	13/55	18/50	0/21
No. new renal scarring/total No.	18/220	0/68	0/109	2/187	Not reported	5/55	0/50	0/21
No. antibiotic resistant (any pathogen)/total No.	26/38	8/10	8/13	Not reported	10/13	Not reported	18/18	Not reported
No. adverse events/total No.	153/302	Not reported	2/122	Not reported	Not reported	0/55	Not reported	Not reported
Control group:								
Strategy	Placebo	No treatment	Placebo	No treatment	No treatment	No treatment	No treatment	Placebo
No. febrile or symptomatic UTI/total No.	72/305	28/69	21/121	9/46	19/122	10/58	15/50	2/20
No. new renal scarring/total No.	19/227	9/68	1/101	2/108	Not reported	2/58	0/50	0/20
No. antibiotic resistant (any pathogen)/total No.	17/69	9/25	3/19	Not reported	7/19	Not reported	0/15	Not reported
No. adverse events/total No.	165/305	Not reported	5/121	Not reported	Not reported	0/58	Not reported	Not reported

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Wang et al.

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Patient and study characteristics