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Evaluating the impact of spatio-temporal smoothness constraints on the BOLD hemodynamic response function estimation: an analysis based on Tikhonov regularization

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Abstract

Recently we have proposed the use of Tikhonov regularization with temporal smoothness constraints to estimate the BOLD fMRI hemodynamic response function (HRF). The temporal smoothness constraint was imposed on the estimates by using second derivative information while the regularization parameter was selected based on the generalized cross-validation function (GCV). Using one-dimensional simulations, we previously found this method to produce reliable estimates of the HRF time course, especially its time to peak (TTP), being at the same time fast and robust to over-sampling in the HRF estimation. Here, we extend the method to include simultaneous temporal and spatial smoothness constraints. This method does not need Gaussian smoothing as a pre-processing step as usually done in fMRI data analysis. We carried out twodimensional simulations to compare the two methods: Tikhonov regularization with temporal (Tik-GCV-T) and spatio-temporal (Tik-GCV-ST) smoothness constraints on the estimated HRF. We focus our attention on quantifying the influence of the Gaussian data smoothing and the presence of edges on the performance of these techniques. Our results suggest that the spatial smoothing introduced by regularization is less severe than that produced by Gaussian smoothing. This allows more accurate estimates of the response amplitudes while producing similar estimates of the TTP. We illustrate these ideas using real data.

Keywords

ER-fMRI; HRF estimation; Tikhonov regularization; GCV

Introduction

Blood oxygen level dependent (BOLD) signal in fMRI (Ogawa *et al* 1990) has become one of the main imaging tools to study the brain *in vivo*. Common assumptions made in fMRI data analysis are the linearity and time invariance of the observed time series (Friston *et al*

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1994). While in general these assumptions are not true (Boynton *et al* 1996, Friston *et al* 2000), they produce reasonable approximations in many practical situations. A linear and time-invariant system can be characterized by an impulse response function that when convolved with the system's input, defines its output. For BOLD fMRI signals, the input of the system is a given stimulation pattern while the output is the observed time series that is compounded by noise due to the scanning process and biological processes. The so-called hemodynamic response function (HRF) plays the role of the impulse response function at each voxel of the brain.

The detection of functional activity in fMRI is often carried out by assuming some fixed shape of the HRF and solving a voxel-wise regression model (Friston et al 1995). Though this approach has rendered great services to the neuroimaging community, it has the intrinsic problem of biasing detection toward the assumed shape. This opens the possibility of a loss of statistical power in situations where the elicited HRF greatly differs from the assumed shape, as previously reported in epileptic (Lu et al 2006) and stroke (Bonakdarpour et al 2007) patients as well as in normal subjects (Fuhrmann Alpert et al 2007, Marrelec et al 2003). It has been proposed that the estimated HRF time course can be used to extract valuable information about the timing of neuronal events in different regions of the brain (Menon et al 1998, Miezin et al 2000). While caution should be exercised in this type of analysis due to differences in vasculature across brain regions (Miezin et al 2000), this is potentially another useful application of fMRI. Thus, estimating HRF without assuming any shape is very desirable even though it is a complex problem confounded by different factors including low signal-to-noise ratios and nonlinearities of the BOLD fMRI signals. In addition, the huge number of voxels that characterizes the fMRI images makes this problem computationally challenging. Different approaches have been proposed to estimate the HRF time course. These include window averaging (Dale and Buckner 1997), FIR (Dale 1999, Lange et al 1999), Bayesian regularization (Goutte et al 2000, Marrelec et al 2003, Woolrich et al 2004) and different versions of Tikhonov regularization (Vakorin et al 2007, Zhang et al 2007). Recently, we proposed (Casanova et al 2008) the use of Tikhonov regularization with boundary and temporal smoothness constraints (Goutte et al 2000, Marrelec et al 2003) for HRF estimation. We explored its properties in different contexts using one-dimensional simulations, and we found it to be fast and reliable for estimating HRF. The regularization parameter was selected using generalized cross-validation (Carew et al 2003, Wahba 1990). Here, we extend the algorithm by adding a spatial smoothness constraint. The basic idea is to use information from neighboring voxels when estimating the HRF at a given voxel. The spatial smoothness constraint is imposed by using second derivative information. Similar ideas have been implemented before in the context of the MEG-EEG inverse problem by Pascual-Marqui (1994). Penny et al (2005) used spatial priors based on second derivative information similar to those used by LORETA in their Bayesian approach for fMRI data analysis. We relied on Tikhonov regularization with GCV to choose the regularization parameter. Our method imposes not only spatial but also temporal smoothness on the estimated HRF and does not require prior Gaussian smoothing of the images which is often performed in fMRI data analysis as a preprocessing step. Though useful in many respects, the smoothing step introduces distortions in the data. These distortions are particularly evident in voxels at the edges of the brain since voxels belonging

to the background are used in the smoothing operation introducing a bias toward zero of the signal. Here we present the results of different two-dimensional simulations comparing the relative performance of these two different versions (temporal versus spatio-temporal) of Tikhonov regularization. We focus these simulations on examining the impact of the smoothing introduced by the Gaussian filter or by the spatio-temporal Tikhonov regularization when estimating the HRF. Special attention is directed to the edges of the brain and activations. We also use real data to illustrate these concepts.

Materials and methods

Linear model

If we assume that the fMRI response is linear and time invariant, then the BOLD signal *y* at a given voxel can be represented as

y = Xh + n, (1)

where y is a $N \times 1$ vector representing the fMRI signal from a voxel, N is the number of time samples, $X = [X_1X_2 \dots X_{N_e}], h = [h_1, h_2 \dots h_{N_e}]^T$, where X_i and h_i are the stimulus convolution matrix and the vector representation of the HRF corresponding to the event *i*, respectively and N_e is the number of events. The stimulus convolution matrix X_i is generated based on the stimulus sequence. The dimension N_s of the vector h_i determines a fixed timespan of the HRF and its time resolution. Finally, *n* is additive noise with covariance matrix V.

In this work, we model the drift usually observed in the fMRI times series using orthogonal polynomials, the number of events was set to $N_e = 1$, and the noise was considered to be independent and identical distributed (i.i.d.) Gaussian (V = I). In order to account for the autocorrelation present in the fMRI time series, the temporal covariance structure should be estimated based on the acquired data (Bullmore *et al* 2001, Worsley *et al* 2002) and used to pre-whiten the fMRI time series before applying the model. So, our linear model will be defined as

$$y = Xh + Pl + n$$
, (2)

where *P* is a $N \times M$ matrix containing a basis of *M* orthonormal polynomial functions that takes a potential drift into account. The highest order of the polynomials is M - 1 and *l* is the $M \times 1$ vector of the drift coefficients. We took M = 3 in our implementation to model first-and second-order drift commonly observed in fMRI times series.

HRF estimation methods based on regularization

Tikhonov regularization with temporal smoothness constraint (Tik-GCV-T)-

The Tikhonov solution of (2) with temporal smoothness constraining (Casanova *et al* 2008) is obtained by solving the following optimization problem:

$$\min_{h,l} \{ \|Xh + Pl - y\|_2^2 + \lambda^2 \|L_1h\|_2^2 \}, \quad (3)$$

where λ is the regularization parameter. The solution of (3) is given by

$$h_{\lambda} = (X_{\perp}^{T} X_{\perp} + \lambda^{2} Q)^{-1} X_{\perp}^{T} y$$

$$Q = L_{1}^{T} L_{1}, \quad L_{1} = \begin{pmatrix} -2 & 1 & 0 & \dots & 0 \\ 1 & -2 & 1 & \dots & 0 \\ 0 & 1 & -2 & \dots & 0 \\ \vdots & \vdots & \vdots & 0 & 0 \\ 0 & \dots & \dots & 1 & -2 \end{pmatrix} \quad (4)$$

$$X_{\perp} = (I - P (P^{T} P)^{-1} P^{T}) X = (I - P P^{T}) X.$$

The last equality results from orthogonality of *P* in our case. The resulting X_{\perp} amounts to removing the drift from the stimulus convolution matrix (Liu *et al* 2001). L_1 is a $(N_s - 2)(N_s - 2)$ matrix embodying the discrete second derivative (Marrelec *et al* 2003) and we assumed the boundary conditions $h_1 = h_{N_s} = 0$ (Goutte *et al* 2000) that implies removing two columns from matrix *X*. For the particular case when $\lambda = 0$ this method reduces to the common least squares (LS).

Tikhonov regularization with spatio-temporal smoothness constraints (Tik-GCV-ST)—We have implemented a version of Tikhonov regularization with both spatial and temporal smoothness regularization. The basic idea is to estimate the HRF at a given voxel and its neighbors simultaneously using an additional second derivative spatial smoothness constraint. The same temporal smoothness and boundary constraints used in equation (3) are imposed on all HRFs being simultaneously estimated. The Tikhonov functional for a given voxel is then defined by

$$\min_{\tilde{h},\tilde{l}} \left\{ \left\| (I_{N_{n+1}} \otimes X)\tilde{h} + (I_{N_{n+1}} \otimes P)\tilde{l} - \tilde{y} \right\|_{2}^{2} + \lambda^{2} \left\| L_{s} L_{t}\tilde{h} \right\|_{2}^{2} \right\} \\
L_{s} = \begin{pmatrix} L_{2} \\ L_{3}^{1} \\ L_{3}^{2} \\ L_{3}^{3} \\ L_{3}^{4} \end{pmatrix}, \quad \tilde{h} = \begin{pmatrix} h_{0} \\ h_{1} \\ h_{2} \\ h_{3} \\ h_{4} \end{pmatrix}, \quad \tilde{l} = \begin{pmatrix} l_{0} \\ l_{1} \\ l_{2} \\ l_{3} \\ l_{4} \end{pmatrix}, \quad \tilde{y} = \begin{pmatrix} y_{0} \\ y_{1} \\ y_{2} \\ y_{3} \\ y_{4} \end{pmatrix}, \quad (5)$$

where N_n is the number of cardinal neighboring voxels (see figure 1) based on a four neighborhood system, h is the $(N_s - 2) * (N_n + 1) \times 1$ vector that contains the stacked HRFs of the voxel of interest and its N_n neighbors. Similarly \tilde{y} and l are the $N(N_n + 1) \times 1$ and $M(N_n + 1) \times 1$ vectors containing the stacked data and drift parameters. The symbol \otimes is the matrix kronecker product. L_1 is the same as in equation (3) and L_2 (2D Laplacian) and L_3^i are $(N_s - 2) \times ((N_s - 2)(N_n + 1))$ matrices that embody spatial second derivative constraints on the voxel of interest and its neighbors. The two-dimensional discrete Laplacian L_2 is computed by assigning to the diagonal elements the number of neighbors of the analyzed

voxel. It has to be noted that this is not used by LORETA (Pascual-Marqui *et al* 1994) nor the one used in Penny *et al* (2005) since the weight we assigned to the voxel of interest is

 N_n . Our Laplacian by itself is singular but our choice for $L_3^{(i)}$ renders the matrix L_s nonsingular. Specifically, we chose it as the second derivative with weight fixed to -2 for a given neighboring voxel and 1 for the voxel of interest. The advantage of our choice of the second derivative constraint is that it avoids the bias toward zero of parameter estimates in the voxels at the boundary of the brain produced by Gaussian smoothing of the data or other choices of regularization operators. The L_s , h, \tilde{l} , \tilde{y} displayed after equation (5) correspond to the situation when the voxel has all of its four neighbors present (see figure 1) or it does not lie in the brain boundary. The computation of L_s , L_t requires for each voxel a customized treatment according to the number of neighboring voxels N_n .

The \hat{h} that minimizes the problem defined in equation (5) is

$$\hat{h} = \left(X_B^T X_B + \lambda^2 L^T L\right)^{-1} X_B^T \tilde{y} \quad (6)$$

where

$$X_{\scriptscriptstyle B} {=} (I_{\scriptscriptstyle N_n} \otimes X_{\perp}), \quad L {=} L_s L_t, \quad \hat{h} {=} \begin{pmatrix} \hat{h}_0 \\ \hat{h}_1 \\ \hat{h}_2 \\ \hat{h}_3 \\ \hat{h}_4 \end{pmatrix}.$$

After simultaneously estimating the HRFs for the voxel of interest and its neighbors, we retain \hat{h}_0 as the final estimator of the HRF and discard the others.

The selection of the regularization parameter in both algorithms is based on generalized cross-validation GCV (Golub *et al* 1979, Wahba 1990). In our implementation, we used a MATLAB package for regularization available publicly at http://www2.imm.dtu.dk/~pch/Regutools/regutools.html) (Hansen 1994).

Simulated data

Signal generation—We generate sequences of two-dimensional 17×17 images, where activations have been added in different positions, to simulate one slice of the fMRI time series. For the voxels inside the activated regions, single event sequences generated with a geometric probabilistic distribution of the intertrial intervals (ITI) (Burock *et al* 1998) were convolved with the difference of gamma densities function (Worsley *et al* 2002)

$$h(t) = \mathrm{HR}\left(\left(\frac{t}{d_1}\right)^{a_1} \mathrm{e}^{-(\frac{t-d_1}{b_1})} - c\left(\frac{t}{d_2}\right)^{a_2} \mathrm{e}^{-(\frac{t-d_2}{b_2})}\right),$$

where $a_1 = 6$, $a_2 = 12$, $b_1 = b_2 = 0.9$, c = 0.35, HR = 0.3. In this case the time to peak (TTP) was 5.4 s, the hemodynamic response (amplitude) HR was the same for all activated voxels.

The fMRI signal was simulated by putting together the sequence of events at a finer temporal grid. The event onsets were approximated to the nearest point in the grid and then convolved with the test function h(t). The resulting signal was under-sampled to TR = 1 s and Gaussian white noise with a fixed standard deviation was added to the data together with a quadratic drift. The durations of the experiment and the HRF were set to 310 s and 20 s, respectively.

Measures of performance

The comparative results presented below are based on the relative errors for TTP, HR and the root mean square error (RMS) given in percentages. We estimate TTP and HR as

$$HR=\max(abs(h_i)), \quad i_{max}=\{i \in 1, 2, \dots, N_s | HR=h_i\}, \quad TTP=(i_{max}-1)*dt,$$

where dt is the time resolution of the HRF estimation. The measures of performance we will use in our simulations are

$$e(\text{TTP}) = \frac{1}{N_r N_{\text{av}}} \sum_{l=1}^{N_r} \sum_{k=1}^{N_{\text{av}}} \frac{\left|\text{TTP}_{l,k}^{\text{est}} - \text{TTP}^{\text{true}}\right|}{\text{TTP}^{\text{true}}} \times 100 \quad (7)$$

$$e(\mathrm{HR}) = \frac{1}{N_r N_{\mathrm{av}}} \sum_{l=1}^{N_r} \sum_{k=1}^{N_{\mathrm{av}}} \frac{\left|\mathrm{HR}_{l,k}^{\mathrm{est}} - \mathrm{HR}^{\mathrm{true}}\right|}{\mathrm{HR}_{l,k}^{\mathrm{true}}} \times 100 \quad (8)$$

$$e(\text{RMS}) = \frac{1}{N_r N_{\text{av}}} \sum_{l=1}^{N_r} \sum_{k=1}^{N_{\text{av}}} \frac{1}{\text{HR}_{l,k}^{\text{true}}} \sqrt{\frac{\sum_{i=1}^{N_s} (h_{l,k}^{i} \text{ est} - h_{l,k}^{i} \text{ true})^2}{N_s}} \times 100, \quad (9)$$

where N_r , N_{av} stand for the number of realizations of the noise and the number of activated voxels, respectively. $H_{l,k}^{true}$ and $H_{l,k}^{est}$ are the true and estimated amplitudes of the response for individual voxels inside the simulated activation. $\text{TTP}_{l,k}^{est}$ and TTP^{true} are the estimated and true time to peak for individual voxels inside the simulated activation. The latter is constant for all activated voxels (5.4 s). Finally, $h_{l,k}^{i^{est}}$ and $h_{l,k}^{i}$ true stand for the estimated and true HRF samples for individual voxels inside the simulated activation. These measures of the relative error were computed for the most efficient design (Dale 1999) generated using a random walk of 1000 iterations (Casanova *et al* 2008). The efficiency when optimizing the designs was computed according to Dale (1999) and taking into account the drift (Liu *et al* 2001).

Simulations

In all simulations we assumed that the HRF is smooth and that the fMRI signal was linear and time invariant, common assumptions in fMRI data analysis. The amplitude of the HRFs was the same for all activated voxels (HR = 0.3) in all simulations with the exception of the simulation 1 c) where a spatially varying profile of the HRF amplitudes was simulated. We compare the performance of Tikhonov regularization with temporal and spatio-temporal smoothness constraints, and we attempt to quantify the relative impact of these constraints on the HRF estimation in different situations. The least-square method is also included in some of these simulations as a reference since it has been used often in the past (Dale 1999, Miezin *et al* 2000, Serences 2004). In these simulations, TR = 1 s, the HRF estimation resolution was TR and the minimum inter trial interval (ITI_{min}) was 2 s. The time span of the HRF was fixed to 20 s. As a rule, when Tikhonov-GCV-ST is employed the data are never smoothed with a Gaussian filter.

Simulation 1—This simulation was designed to quantify the effects of smoothing the data with a Gaussian kernel and to assess the relative performance of the algorithm with spatio-temporal constraints versus Tikhonov-GCV-T and LS. The activation in all cases was simulated in the center of the image as shown in the left panel of figure 2. We analyze three situations: (a) the data are not smoothed and (b) the data are smoothed with a Gaussian isotropic kernel with full width half maximum (FWHM) of 2 pixels as a pre-processing step before using LS or Tik-GCV-T. These simulations are performed across different levels of Gaussian white noise (0.1–0.5 std). (c) The same as the previous one but we fixed the noise level (0.1 std) and we simulated a spatially varying activation profile. We used the Gaussian

function HR(x, y)=0.3exp $\left(-\frac{(x-x_0)^2+(y-y_0)^2}{\rho}\right)$, where x, y are the pixels coordinates and the pixel (x_0, y_0) lies at the center of the simulated circular activation in the left panel of figure 2. We experimented different widths of the Gaussian function by changing the parameter ρ given in the number of pixels.

Simulation 2—This simulation focuses on the performance of these methods exclusively at the edges. One activation is simulated on the edge of the region of interest (gray square) as shown in the right panel of figure 2. The boundary of the activation was determined. The HRFs and the measures of performance are computed only for the voxels lying on the boundary to quantify the impact of the edges on the HRF estimates by these methods. These simulations are performed across different levels of Gaussian white noise (0.1–0.5 std). The data were smoothed with a Gaussian isotropic kernel with full width half maximum (FWHM) of 2 pixels as a pre-processing step before using LS or Tik-GCV-T.

Simulation 3—These simulations were designed to compare the performance of the two regularization techniques in terms of the variability of the estimates. For four different ITI_{mean} values (3, 7, 10 and 15 s), the mean TTP and HR values, together with their 2.5% lower and upper tails, were computed for 200 realizations of the noise. Three different levels of white noise (0.1, 0.2 and 0.3 std) were explored. Other parameters in this simulation were TR = 1 s and temporal estimation resolution TR. The data were smoothed with a Gaussian

isotropic kernel with full width half maximum (FWHM) of 2 pixels as a previous step before using LS or Tik-GCV-T.

Real data

Stimulus presentation—A healthy volunteer (47 year old male) provided written consent for MRI scanning. The subject was instructed to maintain fixation on a gray cross in the center of a projection screen and to concentrate on the visual stimulus, a black and white checkerboard (250 ms duration) that encompassed \sim 3° of visual space displayed using MR-compatible goggles (Resonance Technology, www.mrivideo.com). The subject was instructed to respond as quickly and as accurately as possible to the occurrence of a single flash of the checkerboard using the right index finger on a keypad. The single event-related paradigm used a geometric design optimized in terms of efficiency (Dale 1999) among 10 000 events generated at random. It consisted of 78 events totaling 330.8 s. The ITI_{min} was 1.7 s, and there were null events built into the paradigm (p = 0.5) (Burock *et al* 1998).

Image acquisition—The experiment was conducted on a 1.5-T GE Twin-speed MRI scanner with EXCITE technology (GE Medical Systems, Milwaukee, WI). Functional imaging was performed in the axial plane using multi-slice gradient echo-planar imaging with a field of view of 24 cm (frequency) \times 15 cm (phase) and an acquisition matrix of 64 \times 40 (28 slices, 5 mm thickness, no gap, TR/TE = 2100 ms/40 ms, flip angle = 90°). A high-resolution structural image was obtained using a 3D spoiled gradient-echo sequence with matrix, 256 \times 256; field of view, 24 cm; section thickness, 3 mm with no gap between sections; the number of sections, 128; and in-plane resolution, 0.94 mm.

Image Processing-Images were motion corrected within SPM99, normalized to Montreal Neurological Institute space using image header information (Maldjian et al 1997) combined with the SPM99 normalization, and resampled to $4 \times 4 \times 5$ mm using sinc interpolation. Statistical parametric maps were generated using SPM99 from the Wellcome Department of Cognitive Neurology, London, England, and implemented in Matlab (The Mathworks Inc., Sherborn, MA) with an IDL (Research Systems Inc., Boulder, CO) interface. The data sets were smoothed using an $8 \times 8 \times 10$ mm full-width-half-maximum Gaussian kernel. The data were high-pass filtered, detrended and globally normalized using the corresponding options in the SPM estimation module. Regional activity was detected using SPM by fitting a regression model based on the stimulus time series convolved with the canonical HRF and the first derivative. The SPMs were thresholded at uncorrected p < p0.000 01 and corrected for multiple comparisons using the random field theory functions for family wise error rate (cluster size test p < 0.05 corrected). The significantly activated region of the visual cortex containing 371 voxels was selected to carry out the HRF estimation. The SPM pre-processed data (without high pass filtering and global normalization) were fed to a set of MATLAB programs with our implementation of the different regularized HRF estimation techniques. The detrending of the time series was carried out as part of the estimation process as shown in equation (2) by including the polynomial term. No high-pass filtering or global normalization was applied. When Tik-GCV-T was applied the data were smoothed with the Gaussian kernel described above. The time span was fixed to 21 s. To obtain full brain coverage in this data set a relatively long TR (2.1 s) was used. We employ a

time resolution of TR/4 to estimate the HRF. We have shown before (Casanova *et al* 2008) that LS performs worse than Tik-GCV-T when a time resolution finer than TR is used for HRF estimation. Therefore we do not present results for real data with LS. The computer used in all analyzes was a dual-core AMD processor system running at 3 GHz with 8 GB of RAM and 64 bit Redhat Linux OS.

Results and discussion

Comparisons of simulations across different noise levels using non-smoothed data are shown in figure 3. Tikhonov-GCV-ST is more accurate than the other two methods for all measures except for HR when the signal-to-noise ratio is high. Figure 4 shows the results of the same simulations for the situation when the data are smoothed with a Gaussian kernel prior to using LS and Tik-GCV-T. In this case, the estimation accuracy of these three methods comes closer in terms of the TTP and mean square errors measures (equations (7, 9)) though Tik-GCV-ST produces more accurate estimates of the response amplitude. In fact, in this last case, the Gaussian smoothing decreases the quality of the amplitude estimates. These results suggest that the smoothing introduced by spatial regularization is less severe than that resulting from the isotropic Gaussian kernel. Figure 5 shows the results for a spatially varying activation profile introduced by using a Gaussian function for a fixed level of white noise (0.1 std). We observe a very similar qualitative result as before. Tik-GCV-ST produces better estimates of the amplitude and the HRF time course in general.

Figure 6 shows the results of simulations that quantify the difference in performance of the three methods when estimating the HRF at the edges for 100 realizations of the noise and across different signal-to-noise ratios (0.1–0.5 std). The spatio-temporal method produces similar estimates of TTP and is often more accurate when estimating the HRF amplitude. The results of the simulation 3 showed a very similar performance of the two regularization techniques in terms of bias and variability when estimating the TTP. However the estimates of the response amplitude were less biased and variable when the Tik-GCV-ST is used. We did not observe major differences of the time course metrics estimates across the different inter trial intervals (3, 7, 10 and 15 s) of the random designs.

Real data results

In figure 7, an activated region of the visual cortex is superimposed on the corresponding anatomical slice. A horizontal row crossing the activation is highlighted. HRFs across voxels lying in this row were estimated using both regularization methods. Figure 8 illustrates different three-dimensional views of the estimated HRFs using Tik-GCV-T (top row) and Tik-GCV-ST (bottom row) for the voxels along the horizontal row highlighted in figure 7. In both cases, 40 s of the HRF time course has been estimated. Though the spatiotemporal method produces a rougher image, a very good synchronization of the time to peaks is observed. Tik-GCV-ST also produces estimates with a higher amplitude than Tik-GCV-T, potentially increasing the sensitivity of this technique.

It should be mentioned that the spatio-temporal algorithm is much slower than its temporal counterpart, due to the increased scale of the regression problem to be solved. In a whole brain computational experiment using the same real data set that contains around 23 000

voxels, Tik-GCV-T can be performed in less than 3 min while Tik-GCV-ST did the same task in 41 min.

Conclusions

In this work, we have proposed a new method based on Tikhonov regularization which estimates the BOLD fMRI hemodynamic response function with spatio-temporal smoothness constraints. This method does not require smoothing the data with a Gaussian filter; instead the smoothness is introduced by the spatial regularization. We have studied its relative performance in relation to its temporal version through two-dimensional simulations. The effect of data smoothing was analyzed, and we found that when the data are not smoothed the spatio-temporal method often produces more accurate estimates of the HRF according to all measures. Although smoothing the data with a Gaussian kernel causes the performance of the two algorithms to be more similar in terms of estimated TTP, it produces less accurate estimates of the HRF amplitude when used as a previous step to Tik-GCV-T. Simulations performed only on the edges of the simulated brain and activations show that these differences, as expected, are even greater in those areas. We observed in both simulations and real data analysis that both algorithms produced robust estimates of the TTP across the activated regions, suggesting the possibility of generating robust TTP maps. In terms of variability of the estimates, both methods show very similar performance when estimating TTP but the spatio-temporal algorithm shows less variability and bias when estimating the response amplitude.

The spatio-temporal method was much more time consuming, although in the simulated situations (single event related designs) it was feasible. However, the performance of the latter could be improved by using more sophisticated matrix computation algorithms that take advantage of the structure of these matrices (sparsity, blocks, etc), coding in C or FORTRAN or distributed computing, which will be pursued in a future study.

In this work, we assumed the noise to be Gaussian and i.i.d. or that an efficient prewhitening procedure is available to be applied as a previous step. A thorough simulation study will be performed in the future to evaluate the impact of the fMRI data spatiotemporal correlation structure on the Tik-GCV-ST HRF estimates. The spatial smoothness constraint used here is two dimensional, similar to that used by Penny et al (2005) although fMRI is intrinsically a 3D process. The method can be expanded to three dimensions by using a 6-neighborhood system (Pascual-Marqui et al 1994) or more complex neighborhood systems. However, the performance of the method in its current 2D form for some specific situations is not known, such as when an activation is located in a plane perpendicular to that of the 2D estimation or when the inherent fMRI activation is not smooth temporally. These can arise for example, in some fMRI paradigms involving speech (Birn et al 1999) or when one is interested in examining BOLD activation in the sub-thalamic nucleus (a small structure) whose spatial extent may not involve more than a few voxels. Finally, a shortcoming of our Tikhonov-regularization-based methods is that in contrast to their Bayesian counterparts (Ciuciu et al 2003, Marrelec et al 2003, Woolrich et al 2004), they lack a signal detection scheme. So far we have relied on them to study HRF estimation properties in simulations or, for real data, we have used SPM software to detect activations

before estimating the HRF. The development of detection methods will be the focus of future research.

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Appendix. Description of matrices L2 and L3i

These matrices are illustrated for the case of voxels with four neighbors and $N_s = 12$. In all cases the dimension is $(N_s - 2) \times ((N_s - 2) (N_n + 1))$.

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Figure 1.

All possible situations for the 4-neighborhood system are shown. The first row corresponds to the case when the voxel is not in the boundary of the brain ($N_n = 4$). The second, third, forth and fifth correspond to the cases when 1, 2, 3 or 4 voxels are missing or $N_n = 3, 2, 1, 0$, respectively. The missing ones are highlighted in red bold.



Figure 2.

In the left and right panels the simulated activations used in simulations 1 and 2 are depicted. They contain 49 and 60 voxels, respectively.



Figure 3.

Results for unsmoothed data. In general Tikhonov-GCV-ST produces more accurate estimates of the HRF across the activated region.



Figure 4.

The spatio-temporal estimates of the HRF time course metrics are compared with the estimates produced by LS and Tik-GCV-T preceded by the application of a Gaussian filter. We recall the reader that no filtering of the data is applied when Tik-GCV-ST is used. The TTP estimates produced by Tikhonov-GCV-T are very similar to those generated by its spatio-temporal counterpart. However, the estimates of the response amplitude are more accurate with Tik-GCV-ST.



Figure 5.

The signal-to-noise ratio was fixed at 0.1 std and a spatially varying activation profile was simulated using a Gaussian function. Different widths of the Gaussian were tested by changing the parameter ρ . Again Tik-GCV-ST produces better estimates of the HRF response amplitude and the time course.



Figure 6.

Measures of performance computed at the edge of the activation. As expected, the difference in performance when estimating the response amplitude is even greater than in the absence of an edge (see figure 4).



Figure 7.

An activated region in the visual cortex is overlaid on the corresponding anatomical slice. A row crossing the activation horizontally is highlighted.



Figure 8.

Different views of the estimated HRFs by Tik-GCV-T (top row) and Tik-GCV-ST (bottom row), respectively, across the horizontal line shown in figure 8. Both methods produce well-synchronized time to peaks. However, the spatio-temporal method estimates higher response amplitudes.