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## Whole Grain Intake and Mortality: Two Large Prospective Studies in U.S. Men and Women

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### Abstract

**Importance**—Higher intake of whole grains has been associated with a lower risk of major chronic diseases, such as type 2 diabetes and cardiovascular disease (CVD), although limited prospective evidence exists regarding whole grains' association with mortality.

**Objective**—To examine the association between dietary whole grain consumption and risk of mortality.

**Design, Setting and Participants**—Two large prospective cohort studies. The study population comprised 74,341 women in the Nurses' Health Study (1984–2010) and 43,744 men in the Health Professionals Follow-Up Study (1986–2010), who were free of CVD and cancer at baseline.

**Main Outcomes and Measures**—Hazard ratio for total mortality and mortality due to CVD and cancer according to quintiles of whole grain consumption, which was updated every two or four years by using validated food frequency questionnaires.

**Results**—We documented 26,920 deaths during 2,727,006 person-years of follow-up. After multivariate adjustment for potential confounders, including age, smoking, body mass index, physical activity, and a modified alternative healthy eating index, higher whole grain intake was

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associated with a lower total and CVD mortality but not cancer mortality: the pooled hazard ratios (HRs) and 95% confidence intervals (CIs) across quintiles of whole grain intake were 1.0, 0.99 (0.95–1.02), 0.98 (0.95–1.02), 0.97 (0.93–1.01), and 0.91 (0.88–0.95) for total mortality ( $P_{\text{trend}} < 0.001$ ); 1.0, 0.94 (0.88–1.01), 0.94 (0.87–1.01), 0.87 (0.80–0.94) and 0.85 (0.78–0.92) for CVD mortality ( $P_{\text{trend}} < 0.001$ ); or 1.0, 1.02 (0.96–1.08), 1.05 (0.99–1.12), 1.04 (0.98–1.11), and 0.97 (0.91–1.04) for cancer mortality ( $P_{\text{trend}} = 0.43$ ). The pooled multivariable-adjusted HRs (95% CIs) for each serving/day increment of whole grain consumption were 0.95 (0.93–0.98), 0.91 (0.87–0.96), and 0.98 (0.94–1.02) for total, CVD and cancer mortality, respectively. Similar inverse associations were observed between bran intake and CVD mortality, with a pooled HR of 0.80 (95% CI 0.73–0.87,  $P_{\text{trend}} < 0.0001$ ), whereas germ intake was not associated with CVD mortality after adjustment for bran intake.

**Conclusions and Relevance**—These data indicate that higher whole grain consumption is associated with lower total and CVD mortality in U.S. men and women, independent of other dietary and lifestyle factors. These results are in line with recommendations that promote increased whole grain consumption to facilitate disease prevention.

### Keywords

whole grains; mortality; prospective cohort study

## INTRODUCTION

Whole grains have been widely recommended in numerous dietary guidelines as a healthful food.<sup>1,2</sup> In comparison to refined carbohydrates, whole grains contain many beneficial nutrients and phytochemicals that primarily reside in the outer layers of grains that are removed during milling processes to produce refined grain products. In laboratory research and human feeding trials, whole grains, as well as constituents of whole grains, such as insoluble fiber, magnesium, and phytochemicals, consistently have beneficial effects on glucose metabolism,<sup>3–5</sup> blood lipids,<sup>6</sup> endothelial function,<sup>7</sup> antioxidant activity,<sup>8</sup> and inflammation.<sup>9,10</sup> In addition, epidemiological studies have consistently found inverse associations between whole grain intake and lower risk of type 2 diabetes (T2D)<sup>11–13</sup> and cardiovascular diseases (CVD).<sup>12–14</sup>

While these lines of evidence suggest beneficial effects of whole grain intake on the overall health, data regarding whole grain intake and mortality was not entirely consistent. For instance, whole grain intake was significantly associated with lower total mortality risk in the Iowa Women's Health Study,<sup>15,16</sup> the Atherosclerosis Risk in Communities (ARIC) study,<sup>17</sup> and the Norwegian County Study.<sup>18</sup> In contrast, null associations were found among a healthy elderly population<sup>19</sup> or diabetes patients.<sup>20</sup> Regarding cause-specific mortality, an inverse association between whole grain intake and CVD mortality was consistently observed in previous studies,<sup>15,16,18–21</sup> although only three studies evaluated cancer mortality and reported a non-significant reduced cancer mortality risk among whole grain eaters.<sup>15,16,18</sup> The inconsistency of results may be explained in part by the heterogeneity in dietary assessments, baseline exclusion criteria, and population demographic and lifestyle characteristics. In particular, most previous studies only assessed the eating frequency of whole grain foods that may contain various proportions of actual

whole grain contents. For example, some studies defined whole grain foods using the proportion of whole grain or bran  $\geq 25\%$  as the criterion and examined total servings per day of whole grain foods in relation to disease risk.<sup>17,19</sup> This approach might bring in residual measurement error because the absolute amount of whole grain varies among the foods. Moreover, only one study, which was conducted among type 2 diabetic women, explicitly examined the intake of added bran and germ in relation to total and cause-specific mortality.<sup>20</sup>

Therefore, in the current study we investigated the association between whole grain intake and total and cause-specific mortality in two large cohort studies with repeated assessments of diet and extended length of follow-up: the Nurses' Health Study (NHS) and the Health Professionals Follow-up Study (HPFS). We also evaluated the association between bran and germ consumption and total and cause-specific mortality.

## METHODS

### Study Populations

The NHS is a prospective cohort study of 121,700 female registered nurses from 11 U.S. states aged 30–55 years initiated in 1976. The HPFS cohort was established in 1986 with an enrollment of 51,529 U.S. male health professionals from all 50 states aged 32–87 years. Through 2010, a response rate exceeding 90% has been achieved in both cohorts.

For the current investigation, we excluded participants with cancer, stroke, or coronary heart disease (CHD) at baseline (1984 for NHS, 1986 for HPFS); those who had incomplete information for dietary data, or who reported implausible total energy intake ( $<500$  or  $>3500$  kcal/day for NHS, and  $<800$  or  $>4200$  kcal/day for HPFS). After exclusions, a total of 74,341 women and 43,744 men remained in the analysis.

The study protocol has been approved by the Human Research Committee of Brigham and Women's Hospital in Boston and the Harvard School of Public Health. The completion of the self-administered questionnaire was considered to imply informed consent.

### Ascertainment of Diet

Intakes of whole grains and other foods were assessed using validated food frequency questionnaires (FFQs) every 2 to 4 years.<sup>22,23</sup> The FFQs inquired about average consumption of foods (with a pre-specified portion size) during the previous year using nine categories of intake frequency ranging from " $<1$ /month" to " $6+$ /day". Open-ended questions were included for breakfast cereal brand names and foods that were not listed on the FFQ.

Intakes of whole grain (g/d) were estimated from all grain-containing foods (rice, bread, pasta, and breakfast cereals) according to the dry weight of the whole grain ingredients in each food.<sup>24–26</sup> Whole grain consumption from breakfast cereal was derived from more than 250 brand name cereals based on information provided by product labels and breakfast cereal manufacturers. In our study, whole grains included both intact and pulverized forms containing the expected proportion of bran, germ, and endosperm for the specific grain types. By definition, the following foods/ingredients were considered as whole grains: whole

wheat and whole wheat flour, whole oats and whole oat flour, whole cornmeal and whole corn flour, whole rye and whole rye flour, whole barley, bulgur, buckwheat, brown rice and brown rice flour, popcorn, amaranth, and psyllium. In the FFQ, we further asked the frequency of consuming added bran (oat bran and other bran) and added wheat germ. Intakes of bran and germ were derived directly from whole grain foods and those added to foods. Total bran and total germ are the sum of intakes from both sources.

### Ascertainment of Mortality

Deaths were reported by the next of kin or the postal authorities or by searching the National Death Index.<sup>27</sup> More than 97% of deaths can be identified in these cohorts.<sup>27</sup> For all deaths, we sought death certificates and, when appropriate, requested permission from the next of kin to review medical records. The underlying cause of death was assigned according to the International Classification of Diseases, 8th Revision (ICD-8). In this analysis we also specifically considered deaths due to CVD (ICD-8 390.0–458.9 or 795.0–795.9) or cancer (ICD-8 140.0–207.9).

### Assessment of Covariates

In both cohorts, information on body weight, medical history, lifestyle characteristics (e.g., cigarette smoking and physical activity), disease diagnoses (diabetes, hypertension, and hypercholesterolemia), and other characteristics were collected at baseline and in biennial validated questionnaires. Alcohol and other dietary information was assessed and updated by validated FFQs. Detailed descriptions on the validity and reproducibility of self-reported body weight, physical activity, and alcohol consumption have been published elsewhere.<sup>28–30</sup> A modified alternative healthy eating index (aHEI) was calculated based on intakes of ten foods and nutrients predictive of chronic disease risk, including fruits, vegetables, nuts and legumes, red/processed meat, sugar sweetened beverages, alcohol, sodium, *trans* fat, long-chain n-3 fats and other polyunsaturated fats.<sup>31</sup>

### Statistical Analysis

We calculated person-years of follow-up from the return date of the first FFQ to the date of death or January 31, 2010 for HPFS or June 30, 2010 for NHS, whichever came first. We calculated and used the energy-adjusted residuals of whole grain intakes to control for total energy intake. To better represent long-term or habitual intake, and to minimize within-person variation, we created and used the cumulative average of energy-adjusted whole grain intake from all available dietary questionnaires from baseline through the end of follow-up.<sup>32</sup> We replaced missing values in each FFQ with cumulative averages based on prior assessments. To minimize the possibility of reverse causation bias, we stopped updating diet after participants reported a diagnosis of diabetes, stroke, or CHD, because diagnosis of these conditions led to changes in whole grain intake (Table S1). We then carried forward the cumulative averages of dietary variables prior to the development of these diseases to represent diet for later follow-up.<sup>33</sup> Due to differences in sex, follow-up time, and the questionnaires in the two cohorts, we conducted analyses separately for each cohort to facilitate better control of confounding. Cox proportional hazard models stratified on age (months) and calendar time (2-year intervals) were used to investigate the association

of whole grain, bran and germ consumption with total and cause-specific mortality. We used hazard ratios (HRs) to measure relative risks in higher intake quintiles in comparison with participants in the lowest quintiles, and 95% confidence intervals (CIs) were calculated for HRs. In multivariate analysis, we controlled for age and ethnicity, as well as time-varying covariates, including BMI, smoking status, alcohol intake, physical activity, multivitamin use, aspirin use, a family history of heart disease, cancer, or diabetes, and history of hypertension, high cholesterol, or diabetes at baseline, total energy (kcal; in quintiles), and modified aHEI (whole grain excluded). In NHS, we also adjusted for menopausal status and postmenopausal hormone use. Proportional hazards assumption was tested by evaluating the significance of the interaction term between quintiles of whole grain consumption and time period of follow-up, and no violation of the proportional-hazards assumption was found ( $P>0.05$  for all tests). Tests for trend were conducted by assigning the median value to each category and modeled this value as a continuous variable. In addition, we used restricted cubic spline regressions with 4 knots to examine dose-response relationships between whole grain intake and risk of mortality. Tests for non-linearity used the likelihood ratio test, comparing the model with only the linear term to the model with the linear and the cubic spline terms. We estimated the association of substituting whole grains for refined grains, red meat, and potato. The calculation of substitution effects was based on the differences of  $\beta$  coefficients between whole grains and a specific food, and their variances and covariance matrix were used to derive the 95% CI for the point estimate.<sup>34</sup> The estimates of association across the two studies were pooled using a fixed-effects model.

All analyses were performed using SAS, version 9.3 (SAS Institute Inc, Cary, NC). All  $p$  values presented are two-tailed, and  $p$  values below 0.05 were considered statistically significant.

## RESULTS

Table 1 presents baseline characteristics by quintiles of whole grain intake. Men and women with higher intake of whole grains were more likely to be physically active and have a history of high cholesterol, were less likely to be current smokers, and had lower alcohol intake. In addition, a higher whole grain intake was associated with a better diet quality as reflected by the higher aHEI score.

In NHS, during 26 years of follow-up (1,798,063 person-years), we documented 15,106 deaths, of which 2,989 were CVD deaths and 5,964 were cancer deaths (of which 499 due to colorectal cancer, 1417 lung cancer, and 895 breast cancer). In HPFS, with up to 24 years of follow-up (928,943 person-years), we documented 11,814 deaths, of which 3,621 were CVD deaths and 3,921 were cancer deaths (of which 423 due to colorectal cancer, 739 lung cancer, and 564 prostate cancer). Table 2 shows associations between whole grain consumption and total and cause-specific mortality. In age-adjusted analyses, higher intake of whole grain was significantly associated with lower total and cause-specific mortality. Further adjustment for other potential confounders, especially physical activity, smoking, and BMI, attenuated these associations, although the statistical significance remained for associations with total and CVD mortality. The pooled HRs (CIs) comparing extreme whole grain intake levels were 0.91 (0.88–0.95;  $P_{\text{trend}}<0.001$ ) for total mortality, and 0.85 (0.78–

0.92;  $P_{\text{trend}} < 0.001$ ) for CVD mortality. In contrast, the same multivariate adjustment abolished the associations for cancer mortality; the pooled HR (95% CI) was 0.97 (0.91–1.04;  $P_{\text{trend}} = 0.43$ ) comparing extreme intake levels. Further analyses showed that whole grain intake was not significantly associated with mortality due to major types of cancer, including colorectal cancer, lung cancer, breast cancer or prostate cancer (Table S2). The association for non-CVD and non-cancer mortality was showed in Table S3.

We further estimated that every serving (28 g)/day of whole grain consumption was associated with 5% (95% CI 2%–7%) lower total mortality or 9% (95% CI 4%, 13%) lower CVD mortality. Consistently, we observed a monotonic relationships between whole grain consumption and total mortality ( $P$  for linearity = 0.047), and CVD mortality ( $P$  for linearity = 0.003), and no evidence of non-linearity was observed (Figure 1).

By using the same model adjustment, we further observed that each serving (28 g)/day consumption of refined grains was associated with a small reduction in total mortality (pooled HR = 0.98, 95% CI 0.97, 0.99), by not with CVD mortality (pooled HR = 0.99, 95% CI 0.97, 1.01) or cancer mortality (pooled HR = 0.98, 95% CI 0.97, 1.00). In the substitution analyses, replacing 1 serving of refined grains or total red meat with 1 serving of whole grains daily was associated with lower CVD mortality: 8% (pooled HR, 0.92; 95% CI, 0.88, 0.97) for replacing refined grains, and 20% (pooled HR, 0.80; 95% CI, 0.75, 0.86) for replacing red meat (Figure 2). The corresponding substitution estimates were 4% and 10% for total mortality, respectively. Replacement of potato was not significantly associated with total or CVD mortality. Meanwhile, no significant associations were found for cancer mortality in substitution analyses (Figure S1). The associations between whole grain intake and CVD mortality largely persisted among participants with various risk profiles defined by age, BMI, physical activity, smoking status, and aHEI score (Table 3, all  $P_{\text{interaction}} > 0.21$ ). In addition, whole grain without added bran or germ was also associated with lower total and CVD mortality to a similar extent (Table S4).

In several sensitivity analyses we updated participants' diet throughout follow-up regardless of disease occurrence and applied a 4-year lagged between exposure and the occurrence of deaths; further stopped updating diet after intermediate (hypertension and hypercholesterolemia) diagnosis; or adjusted for incidence of intermediate outcomes; or used baseline whole grain intakes which did not account for different trends of whole grain intakes over time; or used a multiple imputation procedure with 5 rounds of imputation to impute missing data of exposures and covariates and repeated the analyses. The association between whole grain intake and mortality did not change materially in these analyses (Table S5).

The age-adjusted correlation coefficient for association between intake of whole grain and total bran was 0.87 for women and 0.85 for men; and these figures were both 0.79 for association of whole grain and total germ in men and women. Total bran consumption was significantly associated with lower total and CVD mortality. The pooled HRs (95% CIs) comparing extreme quintiles of total bran intake were 0.80 (0.73–0.87,  $P_{\text{trend}} < 0.001$ ) for CVD mortality (Table 4), and 0.94 (0.90–0.99,  $P_{\text{trend}} = 0.02$ ) for total mortality (Table S6) in multivariate adjustment models. Added bran had similar benefits as naturally-occurring bran

(Table S7). We did not find a significant association between total germ intake and risk of mortality after further adjustment for total bran intake.

## DISCUSSION

In these two cohorts of U.S. men and women, we found that a higher whole grain intake, with or without added bran or germ, was associated with a reduced mortality, especially deaths due to CVD. These associations were independent of demographic and lifestyle predictors of mortality, as well as the overall dietary quality, and largely persisted among participants with various risk profiles.

The associations between whole grain foods and total and cause-specific mortality have been examined in several prior investigations. In the first investigation among the Iowa Women's Health Study, Jacobs et al. found that baseline whole grain foods (dark bread, whole grain breakfast cereal, popcorn, and other foods that contain various whole grain contents) were significantly associated with a lower CHD mortality, which was primarily ascribed to dark bread.<sup>16</sup> In updated analyses in the same cohort with extended follow-up, total servings of whole grain foods were significantly associated with a lower all-cause and CVD mortality, but not cancer mortality.<sup>15</sup> In the ARIC study, total whole grain food consumption was significantly associated with a reduced all-cause mortality, but associations with cause-specific mortality were not examined.<sup>17</sup> In a Norwegian cohort, dark bread was associated with lower total and various cause-specific mortality with similar strength.<sup>18</sup> In contrast, in studies conducted among a healthy elderly population<sup>19</sup> or diabetes patients,<sup>20</sup> whole grain intakes were significantly associated with a lower CVD mortality only. Overall, results regarding lower CVD mortality are concordant with numerous previous studies showing inverse associations between whole grain intake and risk of diabetes,<sup>11–13</sup> hypertension,<sup>17,35,36</sup> and CVD.<sup>12–14</sup> Several clinic trials also demonstrated beneficial effects of whole grain intake on CVD risk factors, such as lipid profiles,<sup>37,38</sup> blood pressure,<sup>39,40</sup> insulin sensitivity and glucose metabolism.<sup>5,21,39</sup>

In contrast, associations between whole grain intakes and cancer mortality remain inconclusive. Consistent with our results, several previous investigations have reported an inverse association between whole grain intake and cancer mortality that was much attenuated after further adjustment for other healthful lifestyle and dietary factors correlated with whole grain intake.<sup>15,16,18</sup> Associations between whole grain intake and cancer incidence may depend on the population characteristics and vary by specific types of cancer. For example, a recent meta-analysis of 25 prospective studies demonstrated that whole grain intake was associated with a reduced risk of colorectal cancer,<sup>41</sup> whereas other studies did not find significant associations with endometrial cancer,<sup>42,43</sup> ovarian cancer,<sup>44</sup> breast cancer,<sup>45,46</sup> or prostate cancer.<sup>47</sup> Consistent with the evidence of incident cancer, we observed that whole grain consumption was associated with lower mortality due to colorectal cancer in men, but no significant associations with mortality due to lung cancer, prostate cancer, or breast cancer were found. Future studies with larger sample size of various cancer caused death are needed to replicate our observation.

An interesting finding of our study is that intakes of bran but not germ were significantly associated with a reduced CVD mortality. Consistently, results from previous analyses in the NHS and HPFS also suggested that the bran component, but not germ, was significantly associated with reduced risk of diabetes,<sup>11</sup> hypertension,<sup>36</sup> CHD,<sup>25</sup> or CVD mortality among those with diabetes,<sup>20</sup> after mutual adjustments of bran and germ. These lines of evidence suggest that the association for germ may not be independent of that of bran. Another possible explanation for null associations for germ may be that absolute intake level for germ in our study is rather low. The observed significant associations for bran are in line with proposed mechanisms that attribute the benefits of whole grains primarily to nutrients and phytochemicals existing in the bran portion.<sup>25</sup> Bran is a rich source of fiber, B-group vitamins, vitamin E, magnesium and phytochemicals, which may potentially explain whole grains' favorable effects.<sup>48</sup> For instance, fiber, primarily found in the bran, may reduce the risk of certain chronic diseases, in particular CVD, metabolic syndrome, diabetes, and certain cancers.<sup>49,50</sup> Antioxidant phytochemicals found in wheat bran fractions, such as phenolic acids and alkylresorcinols may modulate cellular oxidative status and prevent biologically important molecules such as DNA, proteins, and membrane lipids from oxidative damage.<sup>51</sup> In addition, magnesium has potentially favorable effects on insulin sensitivity and diabetes risk,<sup>3</sup> blood pressure,<sup>52</sup> and cardiovascular health.<sup>53</sup>

The strengths of our study include a large sample size, a high follow-up rate, long duration of follow-up, repeated assessments of diet, multivariate adjustment, and assessments of whole grain contents from various food sources. In addition, all participants were health professionals, which may help minimize potential confounding by educational attainment or socioeconomic status. There are also several limitations of our study. First, although we carefully adjusted for multiple dietary and lifestyle factors, residual or unmeasured confounding might still exist and may thus hinder causal inference based on these observations. Second, measurement errors in whole grain intake and other dietary factors are inevitable, although the FFQs used in our cohorts have been validated against diet records with reasonable reproducibility and validity.<sup>22,23</sup> Because of the prospective study design, misclassification of whole grain intake was unlikely to be correlated with study outcome ascertainment and therefore more likely to attenuate true associations toward the null. Moreover, we calculated cumulative averages for dietary intakes to reduce random measurement errors and represent long-term dietary habits.<sup>32</sup> On the other hand, despite that we calculated whole grain intake from all relevant foods in comparison to previous studies that focused on whole grain foods only, the inverse association with CVD mortality was consistently observed in the current and previous investigations, suggesting that the associations are largely robust to various degrees of measurement errors. Finally, the participants included in our study were predominantly middle-aged and older health professions with European ancestry, and it is unknown whether our findings could be generalized to other demographic or ethnic groups.

In summary, our data from two large prospective cohort studies consistently showed significant inverse associations of whole grain intake and mortality, especially CVD mortality. In addition, bran portion of the whole grain foods, as well as bran added to foods, was significantly associated with a lower CVD mortality. These findings further support current dietary guidelines that recommend increasing whole grain consumption to facilitate



primary and secondary prevention of chronic diseases, and also provide promising evidence suggesting diet enriched with whole grains may confer benefits toward extended life expectancy.

## Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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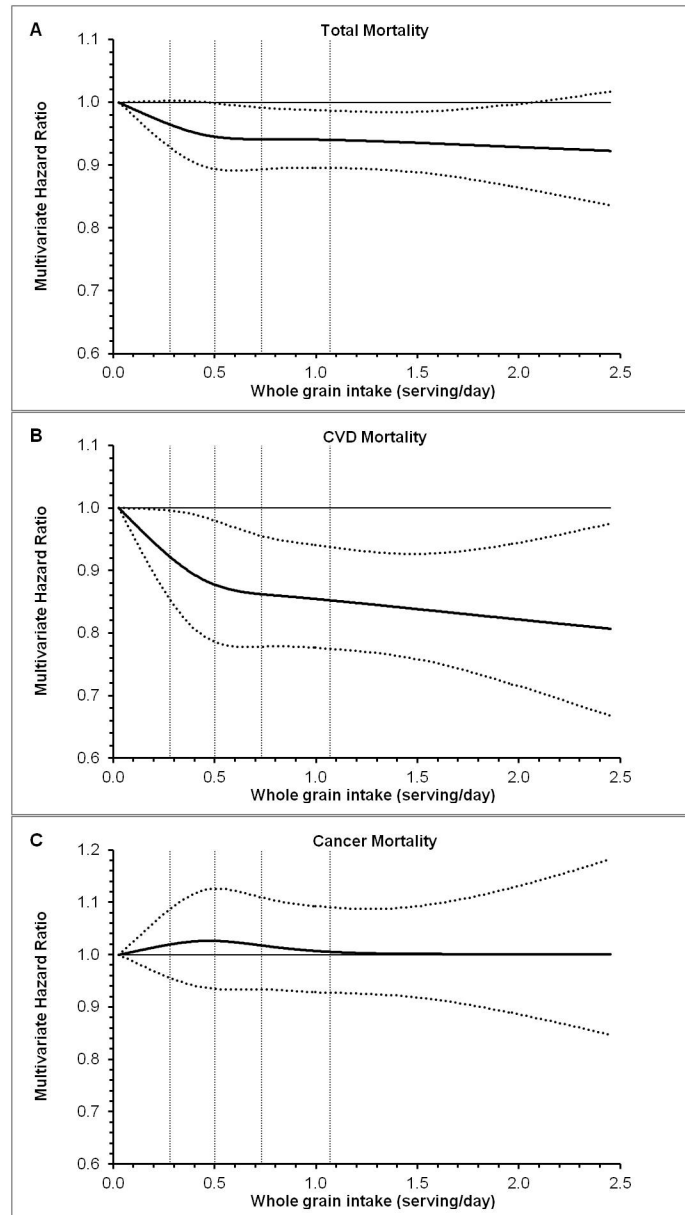
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**Figure 1. Dose-response Relationships between Whole Grain Consumption and Total and Cause-specific Mortality**

Multivariate hazard ratios are calculated in combined dataset of the Nurses' Health Study and the Health Professionals Follow-up Study, using restricted cubic spline regression with adjustment for gender, age (continuous), ethnicity (Caucasian, Asian, African American and Hispanic/others), BMI (<18.5, 18.5–22.9, 23.0–24.9, 25.0–29.9, 30.0–34.9, or 35.0 kg/m<sup>2</sup>), smoking status (never smoked, past smoker, currently smoke 1–14 cigarettes/d, 15–24 cigarettes/d, or ≥25 cigarettes/d), pack years smoked (0, 1–9, 10–24, 25–44, ≥45 pack-years), years since quitting for past smoker (not past smoker, years since smoking quitting <2, 3–5, 6–9 and 10 years), alcohol intake (0, 0.1–4.9, 5.0–9.9, 10.0–14.9, 15.0–29.9, and 30.0 g/d for men; 0, 0.1–4.9, 5.0–9.9, 10.0–14.9, and 15.0 g/d for women), physical

activity (quintiles), family history of diabetes, cancer and heart disease (yes, no), multivitamin use (yes, no), aspirin use at least once per week (yes, no), history of hypertension, high cholesterol, or diabetes at baseline, total energy (kcal/day, in quintiles), and modified alternative healthy eating index (in quintiles), which did not include whole grain. Solid lines represent point estimates and dashed lines are 95% confidence intervals. Vertical dashed lines represent cut-off points for quintiles of whole grain intake. A, total mortality; B, CVD mortality; C, cancer mortality.

Baseline Characteristics of Study Participants by Intake Levels of Whole Grains, the Nurses' Health Study (1984) and Health Professionals Follow-up Study (1986).\*

Table 1

Variable	Quintiles of whole grain intake in NHS				Quintiles of whole grain intake in HPFS			
	Q1	Q3	Q5		Q1	Q3	Q5	
N	14819	14630	14868		8724	8777	8776	
Whole grains, gram/day	4.3	14.7	35.6		5.8	21.8	52.6	
Bran, gram/day	0.4	2.2	8.7		0.5	3.8	11.7	
Germ, gram/day	0.1	0.6	1.8		0.2	0.9	2.7	
Age, years	50.2	50.2	50.3		53.3	53.2	53.2	
BMI, kg/m <sup>2</sup>	25.3	25.2	24.4		25.2	25.1	24.3	
Alcohol, g/d	9.4	6.8	4.7		14.5	11.6	8.1	
Physical activity, METs-hr/week	11.4	13.6	15.7		17.8	20.4	25.8	
Total energy, kcal/day	1711	1786	1653		1984	2060	1879	
Alternative healthy eating index <sup>†</sup>	42.8	45.6	49.5		46.0	49.9	54.6	
Smoking status, %								
Never smoker	36.3	44.6	50.2		41.9	50.6	56.1	
Past smoker	27.1	32.1	35.2		40.7	41.0	39.5	
Current smoker	36.6	23.3	14.6		17.5	8.4	4.3	
Family history of diabetes, %	29.9	29.5	29.0		23.7	24.1	24.0	
Family history of cancer, %	14.3	15.2	16.6		34.9	34.7	36.7	
Family history of MI, %	19.7	19.6	18.9		31.3	30.8	33.0	
Diabetes, %	2.7	2.7	3.5		2.2	2.4	3.3	
Hypertension, %	22.0	20.4	21.2		22.9	19.1	19.9	
High cholesterol, %	6.6	7.6	10.1		9.2	10.1	13.2	
Aspirin use, %	66.3	67.4	64.8		25.3	27.8	26.5	
Multivitamin use, %	30.2	37.0	44.9		34.7	41.0	50.3	

Abbreviations: BMI, body mass index; MI, myocardial infarction; MET, metabolic equivalent task, NHS, Nurses' Health Study; HPFS, Health Professionals Follow-up Study.

\* Values are mean for continuous variables or % for categorical variables. All variables were age-standardized, except age per se.

<sup>†</sup> Whole grains were excluded when calculating the alternative healthy eating index.

**Table 2**  
Hazard Ratio (95%CI) of Total Mortality and Cause-specific Mortality by Quintiles of Whole Grain Intake.

	Quintiles of Whole Grain Intake (g/d)					<i>P</i> <sub>trend</sub>
	Q1	Q2	Q3	Q4	Q5	
<b>Total mortality</b>						
<b>NHS</b>						
Median	4.2	9.7	14.7	21.1	33.0	
No. of cases	4122	2990	2744	2550	2700	
Person-years	356,895	359,688	359,945	361,086	360,449	
Model 1 <sup>*</sup>	1.0	0.72 (0.69–0.76)	0.63 (0.60–0.66)	0.54 (0.51–0.57)	0.49 (0.47–0.51)	< 0.001
Model 2 <sup>†</sup>	1.0	0.97 (0.93–1.02)	0.98 (0.93–1.03)	0.92 (0.87–0.97)	0.86 (0.81–0.90)	< 0.001
Model 3 <sup>‡</sup>	1.0	0.98 (0.93–1.03)	1.00 (0.95–1.05)	0.94 (0.89–0.99)	0.88 (0.84–0.93)	< 0.001
<b>HPFS</b>						
Median	5.9	14.4	22.1	31.3	47.8	
No. of cases	3056	2392	2117	2112	2137	
Person-years	183,911	185,779	186,319	186,660	186,274	
Model 1 <sup>*</sup>	1.0	0.80 (0.76–0.84)	0.68 (0.65–0.72)	0.66 (0.62–0.69)	0.62 (0.59–0.66)	< 0.001
Model 2 <sup>†</sup>	1.0	0.99 (0.94–1.05)	0.95 (0.90–1.01)	0.99 (0.93–1.05)	0.92 (0.86–0.97)	0.006
Model 3 <sup>‡</sup>	1.0	1.00 (0.94–1.05)	0.97 (0.91–1.02)	1.01 (0.95–1.07)	0.95 (0.89–1.00)	0.13
<b>Pooled<sup>§</sup></b>						
Model 1 <sup>*</sup>	1.0	0.75 (0.73–0.78)	0.65 (0.63–0.68)	0.59 (0.57–0.61)	0.54 (0.52–0.56)	< 0.001
Model 2 <sup>†</sup>	1.0	0.98 (0.95–1.02)	0.97 (0.93–1.01)	0.95 (0.91–0.99)	0.88 (0.85–0.92)	< 0.001
Model 3 <sup>‡</sup>	1.0	0.99 (0.95–1.02)	0.98 (0.95–1.02)	0.97 (0.93–1.01)	0.91 (0.88–0.95)	< 0.001
<b>CVD mortality</b>						
<b>NHS</b>						
No. of cases	860	597	535	450	547	
Person-years	359,859	361,853	361,937	362,951	362,354	
Model 1 <sup>*</sup>	1.0	0.69 (0.62–0.77)	0.59 (0.53–0.65)	0.45 (0.40–0.50)	0.46 (0.41–0.51)	< 0.001
Model 2 <sup>†</sup>	1.0	0.96 (0.87–1.07)	0.95 (0.85–1.06)	0.81 (0.72–0.91)	0.84 (0.75–0.93)	< 0.001



		Quintiles of Whole Grain Intake (g/d)					<i>P</i> <sub>trend</sub>
		Q1	Q2	Q3	Q4	Q5	
Model 3 <sup>‡</sup>		1.0	0.97 (0.87–1.08)	0.96 (0.86–1.08)	0.82 (0.73–0.92)	0.86 (0.76–0.96)	0.001
<b>HPFS</b>							
No. of cases		973	721	664	630	633	
Person-years		185,791	187,312	187,630	187,998	187,645	
Model 1 <sup>*</sup>		1.0	0.76 (0.69–0.84)	0.68 (0.61–0.74)	0.62 (0.56–0.68)	0.58 (0.52–0.64)	< 0.001
Model 2 <sup>‡</sup>		1.0	0.92 (0.83–1.02)	0.92 (0.83–1.02)	0.90 (0.81–1.00)	0.82 (0.74–0.91)	< 0.001
Model 3 <sup>‡</sup>		1.0	0.92 (0.84–1.02)	0.92 (0.83–1.02)	0.91 (0.82–1.01)	0.84 (0.75–0.93)	0.002
<b>Pooled<sup>§</sup></b>							
Model 1 <sup>*</sup>		1.0	0.73 (0.68–0.78)	0.63 (0.59–0.68)	0.54 (0.50–0.58)	0.52 (0.48–0.56)	< 0.001
Model 2 <sup>‡</sup>		1.0	0.94 (0.87–1.01)	0.93 (0.87–1.01)	0.86 (0.80–0.93)	0.83 (0.77–0.89)	< 0.001
Model 3 <sup>‡</sup>		1.0	0.94 (0.88–1.01)	0.94 (0.87–1.01)	0.87 (0.80–0.94)	0.85 (0.78–0.92)	< 0.001
<b>Cancer mortality</b>							
<b>NHS</b>							
No. of cases		1449	1170	1162	1094	1089	
Person-years		359,291	361,309	361,345	362,318	361,852	
Model 1 <sup>*</sup>		1.0	0.80 (0.74–0.86)	0.77 (0.71–0.83)	0.68 (0.63–0.73)	0.60 (0.56–0.65)	< 0.001
Model 2 <sup>‡</sup>		1.0	1.02 (0.94–1.10)	1.10 (1.02–1.19)	1.06 (0.98–1.15)	0.98 (0.90–1.07)	0.73
Model 3 <sup>‡</sup>		1.0	1.02 (0.94–1.10)	1.10 (1.02–1.19)	1.06 (0.98–1.15)	0.99 (0.91–1.07)	0.82
<b>HPFS</b>							
No. of cases		976	790	708	725	722	
Person-years		185,746	187,228	187,599	187,911	187,560	
Model 1 <sup>*</sup>		1.0	0.84 (0.76–0.92)	0.72 (0.66–0.80)	0.71 (0.64–0.78)	0.67 (0.61–0.74)	< 0.001
Model 2 <sup>‡</sup>		1.0	1.01 (0.92–1.11)	0.97 (0.88–1.07)	1.01 (0.91–1.11)	0.94 (0.85–1.04)	0.29
Model 3 <sup>‡</sup>		1.0	1.01 (0.92–1.11)	0.98 (0.88–1.08)	1.01 (0.91–1.12)	0.95 (0.86–1.05)	0.40
<b>Pooled<sup>§</sup></b>							
Model 1 <sup>*</sup>		1.0	0.81 (0.77–0.86)	0.75 (0.70–0.80)	0.69 (0.65–0.73)	0.63 (0.59–0.67)	< 0.001
Model 2 <sup>‡</sup>		1.0	1.02 (0.96–1.08)	1.05 (0.99–1.12)	1.04 (0.97–1.10)	0.96 (0.90–1.03)	0.30

Quintiles of Whole Grain Intake (g/d)						
	Q1	Q2	Q3	Q4	Q5	<i>P</i> <sub>trend</sub>
Model 3 <sup>‡</sup>	1.0	1.02 (0.96–1.08)	1.05 (0.99–1.12)	1.04 (0.98–1.11)	0.97 (0.91–1.04)	0.43

Abbreviations: CVD, cardiovascular diseases; CI, confidence interval; NHS, Nurses' Health Study; HPFS, Health Professionals Follow-up Study.

\* Age-adjusted.

<sup>‡</sup> Further adjusted for ethnicity (Caucasian, Asian, African American and Hispanic/others), BMI (<18.5, 18.5–22.9, 23.0–24.9, 25.0–29.9, 30.0–34.9, or ≥35.0 kg/m<sup>2</sup>), smoking status (never smoked, past smoker, currently smoke 1–14 cigarettes/d, 15–24 cigarettes/d, or ≥25 cigarettes/d), pack years smoked (0, 1–9, 10–24, 25–44, ≥45 pack-years), years since quitting for past smoker (not past smoker, years since smoking quitting <2, 3–5, 6–9 and 10 years) alcohol intake (0, 0.1–4.9, 5.0–9.9, 10.0–14.9, 15.0–29.9, and ≥30.0 g/d for men; 0, 0.1–4.9, 5.0–9.9, 10.0–14.9, and ≥15.0 g/d for women), physical activity (quintiles), family history of diabetes, cancer and heart disease (yes, no), multivitamin use (yes, no), aspirin use at least once per week (yes, no), history of hypertension, high cholesterol, or diabetes at baseline, total energy (kcal/day, in quintiles). For women, postmenopausal status and postmenopausal hormone use were further adjusted for.

<sup>‡</sup> Further adjusted for modified alternative healthy eating index (in quintiles), which did not include whole grains.

<sup>§</sup> Pooled hazard ratios were calculated using a fixed-effects model.

**Table 3** Pooled Hazard Ratio (95% CI) of CVD Mortality according to Quintiles of Whole Grain Intake by Various Characteristics of Participants.\*

	Quintiles of Whole Grain Intake					<i>P</i> <sub>trend</sub>	<i>P</i> <sub>interaction</sub> <sup>†</sup>
	Q1	Q2	Q3	Q4	Q5		
<b>Stratified by age</b>							
<i>Age &lt; 65 years</i>							
No. of cases	320	216	201	146	150		
Person-years	343,367	348,369	340,124	325,628	299,984		
Multivariate-adjusted hazard ratio (95% CI)	1.0	0.94 (0.78–1.12)	1.01 (0.84–1.21)	0.83 (0.68–1.02)	0.88 (0.71–1.09)	0.13	
<i>Age ≥ 65 years</i>							
No. of cases	1513	1102	998	934	1030		
Person-years	202,282	200,795	209,443	225,321	250,015		
Multivariate-adjusted hazard ratio (95% CI)	1.0	0.96 (0.88–1.04)	0.93 (0.86–1.01)	0.88 (0.81–0.96)	0.84 (0.78–0.92)	< 0.001	0.48
<b>Stratified by BMI</b>							
<i>BMI &lt; 25 kg/m<sup>2</sup></i>							
No. of cases	1121	804	704	686	764		
Person-years	263,214	256,847	261,339	278,390	316,187		
Multivariate-adjusted hazard ratio (95% CI)	1.0	0.97 (0.88–1.06)	0.97 (0.87–1.08)	0.92 (0.83–1.02)	0.87 (0.79–0.96)	< 0.001	
<i>25 ≤ BMI &lt; 30 kg/m<sup>2</sup></i>							
No. of cases	408	317	323	250	288		
Person-years	180,616	192,918	197,649	192,915	171,614		
Multivariate-adjusted hazard ratio (95% CI)	1.0	0.93 (0.80–1.08)	0.96 (0.82–1.12)	0.82 (0.72–0.95)	0.81 (0.69–0.96)	0.06	
<i>BMI ≥ 30 kg/m<sup>2</sup></i>							
No. of cases	304	197	172	144	128		
Person-years	101,820	99,400	90,578	79,643	62,198		
Multivariate-adjusted hazard ratio (95% CI)	1.0	0.85 (0.70–1.04)	0.79 (0.67–0.92)	0.99 (0.82–1.19)	0.85 (0.68–1.06)	0.11	
<b>Stratified by aHEI</b>							
<i>aHEI median level</i>							
No. of cases	624	536	562	553	742		
Person-years	209,786	265,049	302,545	336,987	390,075		
Multivariate-adjusted hazard ratio (95% CI)	1.0	0.97 (0.88–1.06)	0.95 (0.86–1.05)	0.91 (0.81–1.01)	0.86 (0.77–0.97)	0.007	0.73

	Quintiles of Whole Grain Intake					<i>P</i> <sub>Trend</sub>	<i>P</i> <sub>Interaction</sub> <sup>†</sup>
	Q1	Q2	Q3	Q4	Q5		
<i>aHEI</i> < median level							
No. of cases	1209	781	638	528	437		
Person-years	335,560	283,811	246,682	213,584	159,723		
Multivariate-adjusted hazard ratio (95% CI)	1.0	0.94 (0.84–1.06)	0.94 (0.83–1.06)	0.85 (0.75–0.95)	0.85 (0.76–0.95)	0.002	0.21
<b>Stratified by physical activity</b>							
<i>Physical activity &lt; median level</i>							
No. of cases	652	542	543	482	559		
Person-years	290,024	328,705	351,504	368,725	383,318		
Multivariate-adjusted hazard ratio (95% CI)	1.0	0.93 (0.85–1.02)	0.92 (0.84–1.02)	0.90 (0.82–1.00)	0.85 (0.77–0.95)	0.006	
<i>Physical activity &gt; median level</i>							
No. of cases	1181	775	657	599	620		
Person-years	255,324	220,155	197,721	181,847	166,480		
Multivariate-adjusted hazard ratio (95% CI)	1.0	0.96 (0.85–1.08)	0.95 (0.84–1.06)	0.83 (0.73–0.93)	0.84 (0.74–0.95)	0.002	0.35
<b>Stratified by smoking status</b>							
<i>Never</i>							
No. of cases	587	492	494	449	553		
Person-years	217,996	242,354	260,296	276,409	290,455		
Multivariate-adjusted hazard ratio (95% CI)	1.0	1.03 (0.91–1.16)	1.03 (0.91–1.17)	0.90 (0.79–1.02)	0.90 (0.80–1.02)	0.03	
<i>Ever</i>							
No. of cases	1246	825	706	632	626		
Person-years	327,351	306,506	288,930	274,163	259,344		
Multivariate-adjusted hazard ratio (95% CI)	1.0	0.90 (0.82–0.98)	0.89 (0.81–0.98)	0.86 (0.78–0.96)	0.81 (0.74–0.90)	< 0.001	

Abbreviations: CI, confidence interval; BMI, body mass index; aHEI, alternative health eating index

\* Models were adjusted for age (years), ethnicity (Caucasian, Asian, African American and Hispanic/others), BMI (<18.5, 18.5–22.9, 23.0–24.9, 25.0–29.9, 30.0–34.9, or ≥35.0 kg/m<sup>2</sup>), smoking status (never smoked, past smoker, currently smoke 1–14 cigarettes/d, 15–24 cigarettes/d, or ≥25 cigarettes/d), pack years smoked (0, 1–10, 10–24, 25–44, 45 pack-years), years since quitting for past smoker (not past smoker, years since smoking quitting <2, 3–5, 6–9 and 10 years), alcohol intake (0, 0.1–4.9, 5.0–9.9, 10.0–14.9, 15.0–29.9, and ≥30.0 g/d for men; 0, 0.1–4.9, 5.0–9.9, 10.0–14.9, and ≥15.0 g/d for women), physical activity (quintiles), family history of diabetes, cancer and heart disease (yes, no), multivitamin use (yes, no), aspirin use at least once per week (yes, no), history of hypertension, high cholesterol, or diabetes at baseline, total energy (kcal/day, in quintiles) and the modified aHEI which did not include whole grain in (quintiles). BMI (kg/m<sup>2</sup>), physical activity (METs-hr/week), past smoking history/number of cigarettes currently smoked per day, and aHEI score were further adjusted for in analyses that were stratified by these variables. For women, postmenopausal status and postmenopausal hormone use were further adjusted for.

<sup>†</sup> *P*<sub>Interaction</sub> was calculated using likelihood-ratio test.

**Table 4**

Hazard Ratio (95%CI) of CVD Mortality by Quintiles of Total Bran and Germ Intakes.

	Quintiles of Total Bran or Germ Intake (g/d)					<i>P</i> <sub>trend</sub>
	Q1	Q2	Q3	Q4	Q5	
<i>Total bran</i>						
<b>NHS</b>						
Median	0.7	2.0	3.5	5.7	10.4	
No. of cases	895	557	535	466	536	
Person-years	360,229	360,479	362,651	362,983	362,612	
Model 1 <sup>*</sup>	1.0	0.65 (0.58–0.73)	0.61 (0.54–0.68)	0.46 (0.41–0.53)	0.44 (0.39–0.50)	< 0.001
Model 2 <sup>†</sup>	1.0	0.88 (0.79–0.99)	0.94 (0.83–1.06)	0.80 (0.70–0.91)	0.79 (0.69–0.90)	< 0.001
Model 3 <sup>‡</sup>	1.0	0.89 (0.79–0.99)	0.94 (0.83–1.07)	0.80 (0.70–0.92)	0.80 (0.70–0.91)	0.001
<b>HPFS</b>						
Median	0.7	2.6	5.0	8.2	15.0	
No. of cases	930	691	653	660	687	
Person-years	185,617	186,559	188,474	187,981	187,744	
Model 1 <sup>*</sup>	1.0	0.80 (0.71–0.88)	0.73 (0.65–0.82)	0.64 (0.57–0.72)	0.56 (0.50–0.63)	< 0.001
Model 2 <sup>†</sup>	1.0	0.93 (0.84–1.04)	0.93 (0.83–1.05)	0.86 (0.77–0.97)	0.79 (0.70–0.89)	< 0.001
Model 3 <sup>‡</sup>	1.0	0.93 (0.84–1.04)	0.93 (0.83–1.05)	0.86 (0.77–0.97)	0.80 (0.71–0.90)	< 0.001
<b>Pooled<sup>§</sup></b>						
Model 1 <sup>*</sup>	1.0	0.72 (0.67–0.78)	0.67 (0.62–0.73)	0.56 (0.51–0.61)	0.50 (0.46–0.55)	< 0.001
Model 2 <sup>†</sup>	1.0	0.91 (0.84–0.98)	0.93 (0.86–1.02)	0.83 (0.76–0.91)	0.79 (0.72–0.87)	< 0.001
Model 3 <sup>‡</sup>	1.0	0.91 (0.84–0.98)	0.94 (0.86–1.02)	0.84 (0.77–0.91)	0.80 (0.73–0.87)	< 0.001
<i>Total germ</i>						
<b>NHS</b>						
Median	0.2	0.4	0.6	0.9	1.6	
No. of cases	1419	1176	1150	1090	1129	
Person-years	359,732	359,899	362,047	362,371	362,066	
Model 1 <sup>*</sup>	1.0	0.94 (0.84–1.05)	0.86 (0.76–0.98)	0.86 (0.76–0.99)	0.96 (0.84–1.09)	0.02

	Quintiles of Total Bran or Germ Intake (g/d)					<i>P</i> <sub>trend</sub>
	Q1	Q2	Q3	Q4	Q5	
Model 2 <sup>‡</sup>	1.0	1.06 (0.95–1.19)	1.00 (0.88–1.14)	1.03 (0.91–1.18)	1.09 (0.95–1.24)	0.49
Model 3 <sup>‡</sup>	1.0	1.07 (0.95–1.20)	1.01 (0.89–1.15)	1.04 (0.91–1.19)	1.11 (0.97–1.27)	0.30
<b>HPFS</b>						
Median	0.2	0.6	0.9	1.3	2.3	
No. of cases	908	763	662	795	793	
Person-years	185,631	186,437	188,459	187,880	187,637	
Model 1 <sup>*</sup>	1.0	0.91 (0.82–1.01)	0.81 (0.72–0.90)	0.89 (0.79–1.00)	0.93 (0.82–1.04)	0.08
Model 2 <sup>‡</sup>	1.0	1.03 (0.93–1.14)	0.93 (0.83–1.05)	1.02 (0.91–1.15)	1.02 (0.90–1.14)	0.93
Model 3 <sup>‡</sup>	1.0	1.03 (0.93–1.14)	0.94 (0.84–1.05)	1.03 (0.92–1.16)	1.03 (0.92–1.16)	0.71
<b>Pooled<sup>§</sup></b>						
Model 1 <sup>*</sup>	1.0	0.93 (0.86–1.00)	0.83 (0.76–0.90)	0.88 (0.80–0.96)	0.94 (0.86–1.02)	0.008
Model 2 <sup>‡</sup>	1.0	1.04 (0.97–1.13)	0.96 (0.89–1.05)	1.03 (0.94–1.12)	1.05 (0.96–1.14)	0.66
Model 3 <sup>‡</sup>	1.0	1.05 (0.97–1.13)	0.97 (0.89–1.06)	1.04 (0.95–1.13)	1.07 (0.98–1.17)	0.39

Abbreviations: CI, confidence interval; NHS, Nurses' Health Study; HPFS, Health Professionals Follow-up Study.

Total bran and total germ were mutually adjusted in all models.

\* Age-adjusted.

<sup>‡</sup> Further adjusted for ethnicity (Caucasian, Asian, African American and Hispanic/others), BMI (<18.5, 18.5–22.9, 23.0–24.9, 25.0–29.9, 30.0–34.9, or ≥35.0 kg/m<sup>2</sup>), smoking status (never smoked, past smoker, currently smoke 1–14 cigarettes/d, 15–24 cigarettes/d, or ≥25 cigarettes/d), pack years smoked (0, 1–9, 10–24, 25–44, ≥45 pack-years), years since quitting for past smoker (not past smoker, years since smoking quitting <2, 3–5, 6–9 and 10years), alcohol intake (0, 0.1–4.9, 5.0–9.9, 10.0–14.9, 15.0–29.9, and ≥30.0 g/d for men; 0, 0.1–4.9, 5.0–9.9, 10.0–14.9, and 15.0 g/d for women), physical activity (quintiles), family history of diabetes, cancer and heart disease (yes, no), multivitamin use (yes, no), aspirin use at least once per week (yes, no), history of hypertension, high cholesterol, or diabetes at baseline, total energy (kcal/day, in quintiles). For women, postmenopausal status and postmenopausal hormone use were further adjusted for.

<sup>‡</sup> Further adjusted for modified alternative healthy eating index (in quintiles), which did not include whole grains.

<sup>§</sup> Pooled hazard ratios were calculated using a fixed-effects model.