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Genetic Susceptibility to Dental Caries Differs between the Sexes: A Family-based Study

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Abstract

Many of the factors affecting susceptibility to dental caries are likely influenced by genetics. In fact, genetics accounts for up to 65% of inter-individual variation in dental caries experience. Sex differences in dental caries experience has been widely reported, with females usually exhibiting higher prevalence and severity of disease across all ages. The cause for this sex bias is currently uncertain, although may be partly explained by the differential effects of genetic factors between the sexes: gene-by-sex interactions. In this family-based study (N=2,663; 740 families; ages 1–93 years), we assessed dental caries via intra-oral examination and generated six indices of caries experience (DMFS, dfs, and indices of both pit-and-fissure surface caries and smooth surface

Declaration of Interests

All authors declare no conflicts of interest.

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caries in both primary and permanent dentitions). We used likelihood-based methods to model the variance in caries experience conditional on the expected genetic sharing among relatives in our sample. This modeling framework allowed us to test two lines of evidence for gene-by-sex interactions: (1) whether the magnitude of the cumulative effect of genes differs between the sexes, and (2) whether different genes are involved. We observed significant evidence of gene-by-sex interactions for caries experience in both the primary and permanent dentitions. In the primary dentition, the magnitude of the effect of genes was greater in males than females. In the permanent dentition, different genes may play important roles in each of the sexes. Overall, this study provides the first direct evidence that sex differences in dental caries experiences may be explained, in part, by gene-by-sex interactions.

Keywords

gene-by-environment interaction (GxE); tooth decay; sex differences; oral health

Introduction

Dental caries is the process of enamel or dentin demineralization caused by acid produced by cariogenic oral bacteria. This process is opposed by the natural function of saliva to remineralize dental tissue by supplying calcium and phosphate ions that incorporate into the crystalline structure of tooth enamel [Lukacs and Largaespada, 2006; ten Cate et al., 2008]. Caries progression occurs as a result of an imbalance in the processes of demineralization and re- mineralization, eventually leading to cavitations [Featherstone, 2008]. Many factors can affect the processes of demineralization and remineralization including bacterial flora, dietary and oral hygiene behaviors, saliva composition, flow rate, and pH buffering capacity, positional and morphological features of the teeth, fluoride exposures, and socioeconomic factors including access to oral health care [Martinez-Mier and Zandona, 2013]. Host genetics may influence many of these factors leading to inter-individual variation in susceptibility to caries. Indeed, previous studies have shown that dental caries is highly heritable, with 20-65% of variation attributable to genetics [Boraas et al., 1988; Bretz et al., 2005a; Bretz et al., 2006; Bretz et al., 2005b; Shaffer et al., 2012a; Shaffer et al., 2012b; Shaffer et al., 2013b; Shaffer et al., 2012c; Shuler, 2001; Wang et al., 2010]. The current consensus is that the genetics of dental caries may be truly complex, affected not only by many genetic variants, but also by important interactions between genetic and non-genetic factors which may change over the life course.

In conjunction with the environmental and genetic risk factors listed above, sex also affects susceptibility to caries, with epidemiological surveys usually showing females at higher risk and having greater numbers of affected tooth surfaces compared to males [Lukacs and Largaespada, 2006; Martinez-Mier and Zandona, 2013]. The causes of sex differences in dental caries experience are not fully understood, although possible explanations include earlier tooth eruption (and thus longer exposure to cariogenic processes) in females, as well as sex differences in dietary and oral hygiene behaviors, utilization of oral health care, hormones/physiology, and characteristics of saliva [Lukacs and Largaespada, 2006; Martinez-Mier and Zandona, 2013]. The differential actions of genes in men and women

also have been proposed [Ferraro and Vieira, 2010; Vieira et al., 2008]. Genetic variants related to tooth eruption, dietary preferences, physiology, saliva, or other unknown caries risk factors may have different effects in men versus women; such genetic effects are called gene-by-sex interactions.

While it is currently unknown whether gene-by-sex interactions are important for dental caries, recent findings from genetic studies have suggested that this may be the case. Sex chromosomes have been historically regarded as important for sexual dimorphism [Mank, 2009], and in a family-based study, Vieira et al. identified a locus on the X-chromosome showing suggestive linkage to dental caries (p=5E-5) [Vieira et al., 2008]. Similarly, a recent genome-wide association scan (GWAS) of dental caries by Zeng et al. implicated two highly-homologous genes on opposite arms of the X-chromosome, BCOR (p=4E-7) and BCORL1 (p=5E-6) [Zeng et al., 2013]. The same genetic variant in BCORL1 was also reported in a GWAS of novel caries phenotypes in the same sample (p=3E-6) [Shaffer et al., 2013a]. Mutations in BCOR cause oculofaciocardiodental (OFCD) syndrome, a disorder presenting craniofacial and dental anomalies including cleft palate, radiculomegaly, delayed dentition, oligodontia, persistent primary teeth, and defective tooth enamel [Gorlin et al., 1996; Ng et al., 2004]. BCORL1 shows high sequence similarity to BCOR, although its function is unknown. Additionally, notable candidate genes reside on the X-chromosome, such as AMELX, which codes amelogenin, the major protein component of the enamel matrix. Other genes on the X-chromosome, such as MST4 and FGF13, also may influence susceptibility to dental caries [Kuchler et al., 2014]. While X-inactivation in females is traditionally thought to balance the gene dosage between diploid females and haploid males for genes on the X-chromosome, recent studies have shown that 15% of genes on the Xchromosome escape inactivation to some degree, and another 10% show varying patterns of inactivation [Carrel and Willard, 2005]. Thus, X-linked genes may be involved in gene-bysex interactions via dosage effects.

While some genes related to sex differences, such as dental caries experience, may reside on the X-chromosome, the majority of gene-by-sex interactions likely involve autosomal loci [Wijchers et al., 2010]. Differential patterns of autosomal gene expression for males and females may occur for a number of reasons, including the response to estrogen or other sex hormones. Though others have speculated a possible role for gene-by-sex interactions on dental caries [Ferraro and Vieira, 2010; Vieira et al., 2008], no direct evidence for gene-by-sex interactions has yet been reported. In the present study we have used a family-based approach to explore this question. Specifically, we have extended our previously-published models of dental caries heritability [Shaffer et al., 2012c; Wang et al., 2010] to include the contribution of gene-by-sex interactions in explaining the observed correlation in dental caries experience among biological relatives.

Methods

The present study is a direct extension of the variance components models reported by Wang et al. [Wang et al., 2010] and Shaffer et al. [Shaffer et al., 2012c]. Whereas these previous studies reported the heritability of dental caries phenotypes (for the primary and permanent dentitions [Wang et al., 2010], and for pit and fissure vs. smooth tooth surfaces

[Shaffer et al., 2012c]), in the current study we have extended the statistical models to include additional parameters quantifying the heritability in males and females separately, and allowing us to directly test for evidence of gene-by-sex interactions. Below we briefly describe the methods of recruitment, data collection, and statistical analysis used in this study. The details of our extended statistical model are presented in the Appendix.

Recruitment

Household-based recruitment for the Center for Oral Health Research in Appalachia, cohort 1 (COHRA1) initiative was carried out in Allegheny, Washington, and McKean counties in Pennsylvania, and Webster and Nichols counties in West Virginia. Eligible households contained at least one biological parent-child pair, and all members of an eligible household were offered admission to the study regardless of legal or biological relationships or oral health status. Written informed consent was obtained for all adult participants; assent with written parental or guardian consent was obtained for all underage participants. All aspects of recruitment and data collection for the COHRA1 study were approved by the Institutional Review Boards of the University of Pittsburgh and West Virginia University. In total, 2,663 participants from 740 biological families of 1 to 20 members (with mean of 4.7 members) were enrolled. Table 1 shows the number of relative pairs (e.g., parent-offspring, siblings) available for analysis. All biological relationships were validated using genome-wide genetic marker data [Cornelis et al., 2010; Laurie et al., 2010] and standard relationshiptesting methods [O'Connell and Weeks, 1998]. Additional details regarding the design of the COHRA1 study have been previously published [Polk et al., 2008; Shaffer et al., 2012c; Wang et al., 2010].

Data collection and phenotype definitions

Dental caries was assessed via intra-oral examination by dentists or research dental hygienists calibrated at least annually across all sites. Interclass Correlation Coefficient (ICC) analysis was applied to quantify the consistency of caries assessments among and within the examiners. High correlation rates were observed for both inter (ICC > 0.99) and intra-examiner reliability (ICC 0.86 to 0.99) scores [Polk et al., 2008; Wendell et al., 2010]. Tooth surfaces were scored by visual inspection with a dental explorer using methods consistent with the National Center for Health Statistics Dental Examination Procedures Manual [2001] and recommended by the PhenX Toolkit [Hamilton et al., 2011] for comparability across genetic and epidemiological studies. Each tooth surface was classified as sound, precavitated decay, cavitated decay, filled/restored, missing due to decay, or missing due to reasons other than decay. Third molars were excluded from data collection. Edentulous participants were excluded from analysis.

DMFS index (i.e., the number of decayed, missing due to decay, or filled *permanent* tooth surfaces) was calculated for all participants with one or more *permanent* tooth present; dfs index (i.e., the number of decayed or filled *primary* tooth surfaces) was calculated for all participants with one or more *primary* teeth present. These caries indices were generated regardless of the participant's age. However, in general, younger children provided only the dfs index, older children provided both dfs and DMFS indices, and adolescents and adults provided DMFS index. Precavitated decay (i.e., "white spot" lesions) was included in caries

indices based on previous work showing that precavitated decay contributes to (rather than detracts from) the heritability of caries phenotypes [Wang et al., 2010]. The cause of each missing permanent tooth was provided by the participant as one of six possible reasons: trauma, orthodontia, decay, periodontitis, having never erupted, and other. Only missing teeth due to decay were included in DMFS indices.

Because caries risk factors may not uniformly impact surfaces across the dentition, partial DMFS and dfs indices also were calculated for two categories of tooth surfaces: pit and fissure (PF) surfaces, which included occlusal surfaces of molars and premolars, buccal surfaces of mandibular molars, and lingual surfaces of maxillary molars; and smooth (SM) surfaces which included all other tooth surfaces. In total, six dental caries phenotypes were considered in this study: (1) DMFS, (2) PF DMFS, (3) SM DMFS, (4) dfs, (5) PF dfs, (6) SM dfs. Data manipulations, descriptive statistics, and non-parametric (Wilcoxon) tests for sex differences were performed in the R statistical suite (R Foundation for Statistical Computing, Vienna, AU).

Statistical approach

The variance components method as implemented in SOLAR [Almasy and Blangero, 1998] was used to partition the phenotypic variance into environmental, heritable, and residual error components by conditioning on the biological relationships among the participants of the study. Likelihood methods were used to estimate model parameters, and statistical significance was determined by comparing full and constrained models using the likelihood ratio test. We extended the traditional heritability models reported by Wang et al. [Wang et al., 2010] and Shaffer et al. [Shaffer et al., 2012c] by further partitioning the genetic variance, σ_{G}^2 , into separate sex-specific genetic variances, σ_{GM}^2 and σ_{GF}^2 , and by including a new parameter, the male-female genetic correlation, ρ_G , to describe the covariance between opposite-sex relative pairs. Models were adjusted for sex and age. See Appendix for model details.

The extended model allowed us to calculate sex-specific heritabilities (i.e., the proportion of phenotypic variance due to genetics in males and females) simultaneously. Moreover, the extended model provided a framework for testing two lines of evidence for gene-by-sex interactions: (1) whether the magnitude of the genetic variance differs between males and females (i.e., test $\sigma_{GM}^2 \neq \sigma_{GF}^2$), and (2) whether the male-female genetic correlation differs from 100% (i.e., test $\rho_G = 1.0$). A significant difference in the magnitude of the sex-specific genetic variances indicates that genes cumulatively have a larger role in caries experience in one sex compared to the other. A value for genetic correlation that is significantly less than 1.0 roughly indicates that different genes may be important for males than for females, and vice versa. Both of these statistical tests can provide evidence for the role of gene-by-sex interactions.

Results

Demographic characteristics of the sample and descriptive statistics of the dental caries phenotypes are shown in Table 2. Among COHRA1 participants having at least one primary

tooth present, we did not observe sex differences in the proportion of primary tooth surfaces with caries (out of the total number of primary surfaces present; Wilcoxon p=0.33 for smooth surfaces; p=0.10 for pit and fissure surfaces). However, among COHRA1 participants having at least one permanent tooth present, the proportion of permanent tooth surfaces with caries (out of the total number of permanent surfaces present) was higher in females than in males (p=3.7E-5 for smooth surfaces; p=1.3E-6 for pit and fissure surfaces). More detailed comparisons of caries rates between males and females within specific age groups have been previously reported for the COHRA1 sample [Shaffer et al., 2012c; Wang et al., 2010].

Figure 1 shows heritability estimates for males and females separately, and combined. The combined heritability estimates came from the classic heritability models as reported in Wang et al. [Wang et al., 2010] and Shaffer et al. [Shaffer et al., 2012c] where genetic effects were assumed to be identical in males and females. The sex-specific heritabilities were simultaneously estimated from our extended model. For all six caries phenotypes, males showed greater heritability than females and greater than both sexes combined. For primary dentition phenotypes, the combined heritability estimates were partway between those of the sexes, whereas for the permanent dentition phenotypes, the combined heritability is defined as the proportion of phenotype variability attributable to genetic causes. In other words, these values are ratios, and therefore dependent not only on the magnitude of the genetic contribution, but also the magnitude of the total phenotype variance (which also includes the environmental contribution). Thus, strictly speaking, *heritability estimates* cannot be directly compared between the sexes to test for gene-by-sex interactions.

However, both sex-specific genetic variances and male-female genetic correlations can be used to test for gene-by-sex interactions. Table 3 shows estimates of these parameters. For caries phenotypes in the primary dentition, genetic variances were significantly greater in males than in females (p=0.002 to 0.028), but the genetic correlations were not different than 1.0 (p>0.05 for all). The greater genetic variance in males suggests that in the primary dentition the magnitude of the genetic effect (i.e., the cumulative role of genes on caries susceptibility) differs between the sexes. However, the fact that genetic correlations were not different than 1.0 suggests that the same set of genes may be involved in both sexes. This is consistent with the sex-specific and combined heritability estimates (Figure 1). For caries phenotypes in the permanent dentition, the genetic variances were not significantly different between males and females (p>0.05 for all). However, for DMFS and SM DMFS indices, the estimates of genetic correlation were both significantly less than 1.0 (p=0.004 and 0.038, respectively), and for the PF DMFS index, the estimate of genetic correlation showed a similar trend (i.e., $\rho_G = 0.50$), although was not quite statistically significant (p=0.09). These results suggest that the magnitudes of the genetic effects are similar between males and females, but that different genes or suites of genes may be involved in the different sexes. These results are also consistent with the observation that the sexspecific heritability estimates are both greater than the combined heritability estimate (Figure 1); the combined estimate reflects only the subset of genes contributing to caries

experience in both sexes, and is expected to be comparatively lower if different genes are involved.

Discussion

In this study we present the first direct evidence that gene-by-sex interactions may contribute to dental caries experience. We observed that in the primary dentition, the size of the role that genes play differed between males and females, whereas in the permanent dentition, different sets of genes may be involved. For both sexes, and across both dentitions, dental caries experience was moderately-to-highly heritable. Altogether, these results suggest that gene-by-sex interactions may be partly responsible to the observed sexual dimorphism in caries experience. Furthermore, these results suggest that future gene mapping studies seeking to identify the specific genes involved in dental caries may benefit by pursuing gene-by-sex interactions. Indeed, one of the dominant concerns borne by the growing GWAS literature is the problem of "missing heritability", i.e., the fact that specific genetic variants cumulatively account for only a fraction of the genetic variance observed in family studies. For dental caries, our results suggest that the "missing heritability" may be partly due to gene-by-sex interactions.

Studying the etiology of dental caries experience poses a unique challenge in that the manifestation of the disease is difficult to capture in a single numeric phenotype. DMFS and dfs indices were analyzed and presented in this study, although these phenotypes have some limitations. For example, DMFS and dfs indices do not account for the number of teeth present in the mouth and therefore the number of teeth at-risk of developing caries (which varies among participants due to patterns of eruption and exfoliation in children, and tooth loss due to reasons other than caries in adults). Thus, a young participant who has few permanent teeth erupted, for example, may be expected to have a lower DMFS score than a participant with the full permanent dentition present. For this reason, we also analyzed six analogous caries phenotypes representing the *proportion* of affected surfaces (e.g., the DMFS index divided by the number of permanent teeth present, and so on for the other caries indices). The results (not shown) are very similar to the indices presented herein. In addition, age was included in the models presented herein to account for such effects. Another limitation is that DMFS and dfs indices exhibit a skewed distribution, which violates the normality assumption of the variance components model. Therefore, we also analyzed analogous caries phenotypes generated by performing the probit (i.e., inverse normal) transformation [Bliss, 1934] on the percentiles of the residuals of the caries indices. Analysis of the transformed caries indices, which were exactly normally distributed, yielded similar results, except for the transformed PF DMFS, which did not show evidence of geneby-sex interaction (p=0.27; results not shown). Taken together, these observations suggest that our conclusions are largely robust to the choice of phenotypes chosen for analysis. Moreover, noise in the phenotype measurement (such as that due to uncertainty regarding cause of missing teeth or reason for restoration) would bias our analyses toward the null hypotheses of no genetic variance in either sex, and no genetic correlation between the sexes. Give that we observed significant test results, our analysis framework was robust to potential noise due to measurement errors in the caries phenotypes.

In interpreting these results, it is important to note that gene-by-sex interactions represent a specific class of gene-by-environment interactions. Moreover, the sex of a study participant may be correlated with numerous other environmental exposures that are not taken into account in our analysis. Hence, the evidence of gene-by-sex interactions reported herein includes the possible effects of other gene-by-environment interactions, where the environment is closely associated with sex. Therefore, this study does not offer insight into whether interactions are due to physiological differences, behavioral differences, differences in environmental exposures, or any other differences between the sexes. Nevertheless, these results mesh nicely with the position forwarded by Lukacs et al., that differences in physiology and saliva (which are likely influenced by genetics) may partly account for the observed sex differences [Lukacs and Largaespada, 2006].

This research was performed in an understudied, high-risk, rural population with poorer oral health than the US national average. Though caries rates are higher in this population, we speculate that the same risk factors that contribute to disease in our population also affect other populations, and our results are likely generalizable.

In conclusion, this study presents the first direct evidence that gene-by-sex interactions are involved in dental caries experience. Therefore, we advocate that genetics should be considered, along with other potential causes, in efforts to understand the multi-factorial nature of the sex bias in dental caries. Indeed, insight gleamed from GWAS reinforces our hypothesis that dental caries is very complex, and that the next frontier in understanding disease may come from multidisciplinary approaches exploring the interactions among host genetics, epigenetics, environmental exposures, and microbial flora. Toward this end, more work is needed to understand the sex differences in dental caries, and ultimately to fully understand caries etiology, which may lead to improved interventions to prevent or oppose the disease process.

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Appendix

The variance components modeling framework was used to estimate the additive genetic variance (and thereby estimate the "narrow sense" heritability) of a trait by comparing phenotype measurements across all pairs of relatives and unrelated participants while conditioning on the expected genetic sharing (quantified by twice the kinship coefficient, 2 φ) between the pairs. For example, parent-offspring pairs are expected to share 50% of their genome, sibling pairs share 50%, half-sibling pairs share 25%, uncle-nephew pairs share 25%, grandparent-grandchild pairs share 25%, cousin pairs share 12.5%, etc. The standard

$$\begin{split} p_{i} = \mu + \sum_{j}^{n} \beta_{j} X_{ji} + g_{i} + e_{i}, \\ \text{where } p_{i} \text{ is the phenotype for the i-th participant, } \mu \text{ is the overall trait mean, } X_{ij} \text{ and } \beta_{j} \text{ are} \\ \text{the values and regression coefficients, respectively, for the i-th participant and j-th covariate,} \\ g_{i} \text{ is the additive genetic effect for the i-th participant, and } e_{i} \text{ is the residual environmental} \\ \text{effect for the i-th participant. For this study we included sex and age as covariates: } p_{i} = \mu + \\ \beta_{sex} \text{sex}_{i} + \beta_{age} \text{age}_{i} + g_{i} + e_{i}. \text{ The corresponding variance model takes the form:} \\ \sigma_{p}^{2} = \sigma_{sex}^{2} + \sigma_{G}^{2} + \sigma_{E}^{2}, \\ \text{where } \sigma_{p}^{2} \text{ is the variance of the phenotype, } \sigma_{sex}^{2} \text{ and } \sigma_{age}^{2} \text{ are the} \\ \text{components of the phenotype variance attributable to the additive effects of all genes), and } \sigma_{E}^{2} \\ \text{is the residual environmental variance (attributable to un-modeled and unmeasured environment factors). The genetic covariance between two relatives, R_{1} and R_{2} (e.g., parents and offspring, etc.) is COV(R_{1}, R_{2}) = 2\phi\sigma_{G}^{2}. \\ \text{This model assumes identical genetic effects in both males and females, and was used to estimate the combined heritability defined as \\ \end{array}$$

follows: $h^2 = \frac{\sigma_G^2}{\sigma_p^2 - \sigma_{sex}^2 - \sigma_{age}^2}$. Maximum likelihood methods were used to estimate these model parameters simultaneously. The likelihood ratio test was used to determine statistical significance of parameters by comparing the likelihood of a model where the parameter is estimated to the constrained model where the parameter is set equal to its value under the null hypothesis). The test statistic follows the χ^2 distribution for covariate parameters and a 50:50 mixed distribution of a χ^2 and a point mass at zero for genetic parameters.

To test for the actions of gene-by-sex interactions we extended the standard variance components model, above, following the approach by Blangero and others [Blangero, 1993; Brown et al., 2004; Martin et al., 2002]. The extended variance model partitions the genetic and environmental variances into sex-specific components: σ_{GM}^2 and σ_{EM}^2 for males, and σ_{GF}^2 and σ_{EF}^2 for females. Under the extended model, the genetic covariance between same-sex relative pairs is $COV(M_1, M_2) = 2\phi\sigma_{GM}^2$ for males and $COV(F_1, F_2) = 2\phi\sigma_{GF}^2$ for females. In contrast, the genetic covariance between opposite-sex relative pairs is $COV(M, F) = 2\phi\sigma_{GM}\sigma_{GF}\rho_G$, where ρ_G is the male-female genetic correlation, which quantifies the degree to which the genetic influence on the trait is common to both sexes. In the absence of geneby-sex interactions, the sex-specific genetic variances will be equivalent (i.e., $\sigma_{GM}^2 = \sigma_{GF}^2$) and genetic correlation will equal 100% (i.e., $\rho_G = 1.0$). On the other hand, in presence of geneby-sex interactions, the sex-specific genetic variances will differ (i.e., $\sigma_{GM}^2 \neq \sigma_{GF}^2$, indicating the role of genes is larger in one sex compared to the other) and/or the genetic correlation will be less than 100% (i.e., $\rho_G < 1.0$, suggesting that different genes are involved in the two

sexes). Gene-by-sex interactions were tested via likelihood ratio tests, which compared models where all parameters were simultaneously estimated to constrained models where the sex-specific genetic variances were forced to be equivalent and where the genetic correlation was forced to equal 100%. Rejection of constrained models provides evidence of gene-by-sex interactions.

Of note, statistical power to detect heritability (i.e., $h^2 > 0.0$) was high. However, power to detect gene-by-sex interactions was comparatively lower, due to the burden of estimating a greater number of model parameters, and the comparatively smaller number of informative relative pairs for estimating each parameter. Therefore, non-significant results should be cautiously interpreted as lack of evidence for gene-by-sex interactions, as opposed to evidence that interactions are nonexistent.

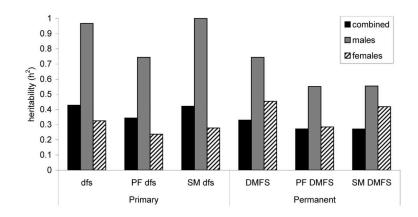


Figure 1.

Heritability estimates in males and females, combined and separately, for dental caries indices in the primary (dfs) and permanent (DMFS) dentitions, and in pit and fissure (PF) and smooth (SM) tooth surfaces.

Table 1

Biological relationships with the COHRA1 sample

Relative pairs	Ν
Parent-offspring	1,736
Siblings	676
Grandparent-grandchild	60
Half-siblings	322
Avuncular (i.e. uncle-nephew)	124
First cousins	98
Other relatives	35
Total related pairs	3,051
Within kinship unrelated pairs (i.e. spouses, etc.)	739
Total pairs	3,790

Table 2

Demographic characteristics and summary statistics of dental caries phenotypes

Summary statistics	
Sample size	2,663
Number of kinships	740
Size of kinships, mean (range)	4.72 (1–20)
Self-reported whites, %	89.64
Females, %	55.61
Age in years, mean (SD; range)	19.83 (15.29; 1–93)
Primary dentition (N=1058)	
dfs, mean (SD; range)	3.93 (7.32; 0–53)
PF dfs, mean (SD; range)	1.85 (3.09; 0–16)
SM dfs, mean (SD; range)	2.24 (4.82; 0–37)
Permanent dentition (N=1937)	
DMFS, mean (SD; range)	14.34 (18.76; 0–122)
PF DMFS, mean (SD; range)	6.33 (6.24; 0–24)
SM DMFS, mean (SD; range)	8.18 (13.82; 0–98)

Extended genetic model parameters and tests for gene-by-sex interactions

				p-values	les
trait	$\sigma^2_{_{ m GM}}$	$\sigma^2_{ m GF}$	PG	$\sigma^2_{ m _{GM}}{=}\sigma^2_{ m _{GF}}$	$\rho_{G}=1.0$
primary dentition (N=1052)	ion (N=10)52)			
dfs	59.3	15.0	0.40	0.0015 a	$0.07 \ b$
PF dfs	8.1	2.0	0.56	0.0282	0.20
SM dfs	25.4	5.8	0.34	0.0018	0.14
permanent dentition (N=1933)	tition (N=	=1933)			
DMFS	165.6	165.6 124.5	0.25	0.45	0.0038
PF DMFS	12.5	7.8	0.50	0.35	0.09
SM DMFS	74.0	67.6	0.15	0.87	0.0356
$\sigma_{\rm GM}^2$ = genetic variance in males	variance i	n males			
$\sigma^2_{ m GF}$ = genetic variance in females	ariance ir	ו females			
$\rho \mathbf{G} = \mathbf{g}\mathbf{e}\mathbf{n}\mathbf{e}\mathbf{t}\mathbf{i}\mathbf{c}$ correlation between males and females	rrelation b	oetween I	males ar	d females	
bold = p-values less than 0.05	less than (0.05			

b rejection of the null hypothesis implies that different genes contribute to the cumulative genetic effect in the two sexes

a rejection of the null hypothesis implies that the magnitude of the genetic effect differs between sexes