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Growth Patterns During and After Treatment in Patients with Pediatric ALL: A Meta-Analysis

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Abstract

Background—Survivors of pediatric acute lymphoblastic leukemia (ALL) have a significantly higher body mass index (BMI) than their peers. Understanding the critical time periods in which patients with pediatric ALL are vulnerable to unhealthy weight gain will lay the groundwork for developing effectively timed interventions.

Procedure—We determined the growth patterns of patients with pediatric ALL during and after treatment through the conduct of a systematic review and meta-analysis. A search of MEDLINE, Scopus, and Web of Science was performed from its inception through May 2014. Studies met the inclusion criteria if they included at least 10 patients of pediatric ALL, and longitudinally assessed BMI at diagnosis and at least one time point after diagnosis

Results—Twenty-one studies met the inclusion criteria for the systematic review and 16 were included in meta-analysis. The mean increase in BMI z-score during treatment in 1,514 patients with pediatric ALL was 0.81 (95% CI: 0.25–1.38). Specifically, patients experienced substantial weight gain in early treatment ($=0.41$, 95% CI: -0.34 , 1.17) and again during maintenance $($ =0.34, 95% CI: -0.22 , 0.90). The mean increase in BMI z-score ranged between 0.52 and 0.89 beyond treatment completion. Subgroup analyses found unhealthy weight gain occurred regardless of patients' receipt of cranial radiation therapy, sex, and weight status at diagnosis.

The authors have no conflicts of interests to disclose.

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Conclusions—Patients with pediatric ALL experience unhealthy weight gain early in treatment, and increases in weight are maintained beyond treatment completion. Preventing early onset of obesity is a priority for improving the care and outcomes for patients with pediatric ALL.

Keywords

Obesity; Growth; Acute Lymphoblastic Leukemia; Pediatric; Survivors

INTRODUCTION

Advances in treatment have resulted in significant improvements in survival rates of childhood cancer. Approximately 90% of children diagnosed with acute lymphoblastic leukemia (ALL), the most common childhood cancer, can expect to be cured (1). This success translates into a growing population of long-term survivors (2). At the same time, cancer treatment at a young age also produces late-effects, including obesity (3, 4). Our meta-analysis in 1,742 survivors of pediatric ALL, identified a mean body mass index (BMI) z-score of 0.83, corresponding to the 80th BMI percentile (5). This indicates that pediatric ALL survivors have a significantly higher BMI than their peers.

Obesity adds additional risks to the increased chronic health conditions already experienced by pediatric ALL survivors. Lifestyle interventions promoting healthy eating and physical activity can potentially prevent unhealthy weight gain in pediatric ALL survivors, but a critical question that needs to be addressed is when to intervene. In a retrospective cohort of 83 patients with pediatric ALL diagnosed between 1985–2010, we found a rapid weight gain during induction and the first 6 months of maintenance therapy (6). There is now accumulating evidence suggesting that pediatric ALL survivors experience excessive weight gain early in treatment $(7, 8)$, and increases in weight tend to be maintained throughout the treatment and beyond.

This systematic review aims to identify growth patterns of patients with pediatric ALL by summarizing existing longitudinal evidence on changes in BMI z-score during treatment and years post treatment completion. Understanding the critical time periods in which patients with pediatric ALL are vulnerable to unhealthy weight gain will lay the groundwork for developing effective interventions to reduce premature mortality and serious morbidity in this population.

METHODS

We followed the Preferred Reporting Items for Systematic Review and Meta-Analyses (PRISMA) statement for reporting our results (9) (Supplemental Table I). A protocol was developed prior to the conduct of the systematic review and submitted to PROSPERO (#: CRD42014010315), an international register of systematic review protocols (10).

Literature Search

We searched MEDLINE, Scopus, and Web of Sciences from inception through May 27, 2014 to identify studies that investigated obesity in patients or survivors of pediatric ALL. We used medical subject heading (MeSH) and text words related to obesity (*obesity, weight,*

body composition, growth, etc.) in combination with survivors (*survivors, remission, disease free survival, etc.*) and acute lymphoblastic leukemia (*acute lymphoblastic leukemia, leukemia, precursor cell lymphoblastic leukemia-lymphoma, etc*.) (10). We consulted a research librarian in specifying the search and searched reference lists of eligible studies and relevant narrative reviews to identify additional studies that met inclusion criteria. While we did not set any language restrictions for the search, we did not screen studies without abstracts available in English.

Eligibility Criteria and Study Selection

Two authors (FFZ and SL) independently screened the titles and abstracts to identify potentially eligible studies using the software Abstrackr [\(http://www.cebm.brown.edu/](http://www.cebm.brown.edu/software) [software](http://www.cebm.brown.edu/software)). Discrepancies were jointly reviewed to reach consensus. These two authors then independently screened the full texts of the identified studies to determine which were eligible for review. Studies were eligible if they met the following criteria: (1) were research articles published in peer-reviewed journals; (2) included at least 10 patients who were diagnosed with ALL prior to age 21 years; (3) assessed BMI z-score or percentile according to a reference population (e.g., the 2000 Centers for Disease Control and Prevention (CDC) growth reference (11)); and (4) followed BMI z-score or percentile longitudinally at diagnosis and at least one time point after diagnosis. We excluded review articles, case reports and studies that assessed only weight or only height as the outcome. We also excluded studies that assessed survivors who received hematopoietic stem-cell transplantation (HSCT) based on the rationale that HSCT survivors might experience a different pattern of growth compared to other ALL survivors.

Data collection and extraction

Two authors (FFZ and SL) extracted data and each verified the other's extracted information. Discrepancies were resolved by consensus. For each eligible study, we extracted the following information (1) author, year and country of publication; (2) characteristics of the study population (sample size, years at diagnosis, mean age at diagnosis, treatment protocol, percentage of receiving cranial radiation therapy (CRT), percentage of males, and mean BMI z-score at diagnosis); (3) number of time points assessed; (4) length of follow-up from diagnosis to each of the time points; (5) primary outcomes (BMI z-score or percentile) reported at each time point; and (6) findings by receipt of CRT, sex, and weight status at diagnosis when available. BMI z-score or percentile calculated based on age- and sex-specific BMI cut-offs of a reference population was used to evaluate growth patterns because for individuals < 20 years of age BMI normally changes with age and varies by sex (12). The BMI z-score indicates the number of standard deviations the measurement is away from the age- and sex-specific mean value in the reference population. A BMI z-score >0 or BMI $>50th$ percentile indicates a higher-thanaverage BMI.

Assessment of Study Validity/Quality Assessment

There is no standard scale to assess quality in observational studies (13, 14). We modified the Newcastle-Ottawa Scale (15) to create a checklist including the following quality items:

whether the study adequately described the patients characteristics including age at diagnosis, sex distribution, and treatment protocols and/or years at diagnosis, whether the study clearly described the length of follow-up and the sample size at each follow-up, and whether the study clearly descried the methods of assessing the outcome, and provided the standard deviation, standard error, confidence interval or p-value for the outcome assessed at each follow-up (Supplemental Table II).

Statistical Analysis

In light of clinical heterogeneity in study characteristics, we performed random-effect metaanalysis to combine studies that reported the mean and standard deviation/error of BMI zscore at diagnosis and at least one time point after diagnosis. Studies were grouped and evaluated at seven time periods including (1) diagnosis; (2) end of induction (approximately 1 month post diagnosis); (3) start of maintenance (approximately 6–9 months post diagnosis); (4) end of treatment (approximately 2–3 years post diagnosis); (5) short-term follow-up ($\langle 2 \rangle$ years post-treatment); (6) median-term follow-up ($2 - 4.9$ years posttreatment); and (7) long-term follow-up (≥5 years post-treatment). If multiple BMI z-scores were reported within one time period, we chose the one with the most complete data (i.e. the largest sample size).

Among the 16 included studies, eight (6, 16–22) explicitly reported the mean BMI z-score for the overall cohort, one (23) reported the mean BMI percentile, which was converted to BMI z-score based on the standard normal curve, and we obtained the mean BMI z-score directly from authors in one study (24). For the remaining six studies, the mean BMI z-score was calculated as a weighted average based on subgroup values (using sample size as weights) in five studies (7, 25–28), or was estimated based on median BMI z-score in one study (29). For standard deviation (SD) of the BMI z-score, nine studies (6, 16–23) explicitly reported the standard deviation/error or confidence interval of the BMI z-score or percentile, and we obtained SD directly from authors in one study (24). For the remaining six studies, the SD was either calculated as a pooled standard deviation based on subgroup values assuming equal variance in five studies (7, 25–28), or estimated from median and range in one study (29) using methods described by Hozo et al (30). Among all 16 included studies, the mean and SD of the BMI z-score were estimated based on graphs using DigitizeIt [\(http://www.digitizeit.de/](http://www.digitizeit.de/)) in three studies (22, 25, 29).

Changes in BMI z-score during and after treatment were calculated for the following time periods using methods described in the Cochrane Handbook (31) from diagnosis to: (1) end of induction; (2) start of maintenance; (3) end of treatment; (4) $\lt 2$ years post treatment; (5) 2 -4.9 years post treatment; and (6) 5 years post treatment. Briefly, the changes in BMI zscore were calculated as the difference between the two time points. Only two studies explicitly reported changes in BMI z-score and SDs of the changes (21, 25). We estimated the SDs of the mean changes (SD_{change}) using the SDs for each periods (SD_{follow-up}=SD of mean BMI z-score at the longer follow-up time period; $SD_C=SD$ of mean BMI z-score at diagnosis or the shorter follow-up time period) and an imputed correlation coefficient (Corr) for the pre-post measurements using this formula $SD_{change} = (SD_{follow-up}^2 + SD_C^2 2 \times \text{Corr} \times \text{SD}_{\text{follow-up}} \times \text{SD}_{\text{C}}$). We used the correlation coefficients between time periods in

our previous study as Corr (6) (Supplemental Table III) to impute SDs of the changes. To evaluate whether the growth pattern differed by key study characteristics, we performed subgroup meta-analysis in patients of pediatric ALL treated with CRT and those treated without CRT separately, in female and male patients separately, and in patients with BMI zscore θ and $\leq \theta$ at diagnosis separately.

We performed sensitivity analyses to assess the robustness of our findings. First, we performed the analysis after excluding six studies for which the mean BMI z-score or its standard deviation was not explicitly reported. We also repeated the analysis after excluding three studies for which the mean and SD of BMI z-score was estimated based on graphs using DigitizeIt. In addition, we performed a "leave-one-out meta-analysis" to evaluate the impact of individual studies on the summary estimates (32).

We assessed between-study heterogeneity using Cochran's Q statistic (33) and the I^2 index. (34) The Cochran's Q was considered statistically significant at P_Q < 0.1. The I^2 index, ranging between 0 and 100%, quantified the extent of heterogeneity beyond chance, with higher values indicating greater inconsistency across studies. All analyses were conducted using Stata version IC/12.1 (Stata Corp., College Station, TX, 2012). Statistical significance was defined as a two-sided p-value <0.05 for all tests except those for heterogeneity.

RESULTS

Included Studies

Our initial search identified 1,265 studies from MEDLINE, 522 studies of Web of Science, and 292 studies from Scopus, for a total of 2,079 studies. After screening titles and abstracts, 40 non-overlapping studies were considered potentially eligible and were retrieved for full text review. Of these, 19 studies were excluded, and 21 were included in this systematic review that reported results on 1,791 pediatric ALL survivors (Figure 1). Tables 1 summarizes the characteristics of the 21 studies.

Meta-Analysis of Changes in BMI Z-Score During and After Treatment

Sixteen studies provided (6, 7, 16–21, 23–29, 35) data for mean and SD of BMI z-score or percentile for at least two time points during and after treatment and were included in the meta-analysis. Baseline BMI z-scores ranged from −0.64 to 0.64 across studies. In 14 studies (6, 7, 16, 17, 19–28) that examined changes in BMI z-score from diagnosis to end of treatment, a significant increase in pooled BMI z-score was observed in 1,514 patients with pediatric ALL (Figure 2) ($=0.81, 95\%$ CI: 0.25, 1.38). When different treatment phases were considered, a rapid weight gain occurred during the early treatment (i.e. from diagnosis to start of maintenance) ($=0.41, 95\%$ CI: $-0.34, 1.17$) in 990 patients with pediatric ALL. Specifically, there was a substantial increase in pooled BMI z-score during induction $(=0.67, 95\% \text{ CI: } -0.72, 2.06)$ (i.e. from diagnosis to end of induction), followed by a decrease in pooled BMI z-score during consolidation ($=$ −0.50, 95% CI: −1.78, 0.77) (i.e. from diagnosis to start of maintenance). Weight gain occurred again from start of maintenance to end of treatment ($=0.34, 95\%$ CI: $-0.22, 0.90$) in 966 patients with pediatric ALL (6, 7, 16, 17, 21, 22, 24, 25) (Supplemental Figure 1), although the increase in

Weight gain that occurred during treatment persisted beyond completion of treatment. Eleven studies evaluated growth patterns after treatment completion. The increase in pooled BMI z-score was 0.89 (95 CI: −0.34, 2.11) from diagnosis to <2 years post-treatment in 208 patients with pediatric ALL, was 0.79 (95 CI: -0.54 , 2.13) from diagnosis to 2 – 4.9 years post-treatment in 370 patients with pediatric ALL, and was 0.52 (95 CI: −0.90, 1.94) from diagnosis to 5 years post-treatment in 378 patients with pediatric ALL (Figure 2, Supplemental Figure 2). Little between-study statistical heterogeneity was observed for changes in BMI z-score during and after treatment with I^2 all approaching 0 (Supplemental Figures 1 and 2).

Sensitivity Analysis

Sensitivity analysis demonstrated similar growth patterns during and after treatment in patients with pediatric ALL after excluding studies that did not explicitly report the mean or SD of the BMI z-score/percentile or after excluding studies that obtained the mean and standard deviation of the BMI z-score based on graphs (Supplemental Figure 3). Leave-oneout meta-analysis did not find any specific study significantly impacted the meta-analysis results.

Assessment of Quality and Reporting

All 21 studies clearly defined the outcome and described the methods used to calculate BMI z-score/percentile. Most studies clearly stated patients' age at study evaluation (95%), sex distribution (95%), and treatment protocols or years at diagnosis (71%). Fifteen studies (71%) described the length of follow-up, and 19 (90%) provided the sample size at the follow-up. Thirteen studies (62%) provided SD, standard error, confidence interval, or presented sufficient data for the estimation of SD (Supplemental Table II). Because all studies met at least five of the seven quality assessment criteria, we did not perform subgroup meta-analysis by study quality.

Growth Patterns in Patients with pediatric ALL by Key Study Characteristics

CRT—Four studies (6, 27, 28, 36) reported changes in BMI z-score separately for patients treated with CRT and those treated without CRT; one study (26) exclusively included patients treated with CRT; and five studies (16, 17, 19, 22, 24) exclusively included patients treated without CRT. Among the 10 studies, two (27, 36) did not provide SD for BMI zscore at diagnosis and were excluded in meta-analysis. Pooled changes in BMI z-score from diagnosis to end of treatment was 0.95 (95% CI: −0.06, 1.96) in the 127 survivors treated with CRT based on four studies, and was 0.97 (95% CI: 0.41, 1.54) in the 668 survivors treated without CRT based on eight studies. Growth patterns were also similar in patients treated with and without CRT beyond treatment completion (Figure 3A).

Sex—Two studies (6, 26) reported changes in BMI z-score separately in male and female survivors. Pooled changes in BMI z-score from diagnosis to end of treatment were 1.19 (95% CI: −0.26, 2.64) in the 64 female survivors of pediatric ALL and was 0.95 (95% CI:

−0.52, 2.42) in the 59 male survivors. Increase in BMI z-score was observed beyond treatment completion in both male and female survivors although it was slightly higher in female survivors (Figure 3B). Both male and female survivors experienced weight gain post treatment completion.

Weight status at diagnosis—Eight studies (20–22, 25–29) assessed changes in BMI zscore in patients with a mean BMI z-score <0 (i.e. below-average BMI) at diagnosis, and eight $(6, 7, 16–19, 23, 24)$ in patients with a mean BMI z-score θ (i.e. average or aboveaverage BMI) at diagnosis. During treatment, weight gain was observed regardless of weight status at diagnosis, although meta-analysis combining cohorts of 667 patients with belowaverage BMI showed a greater increase in pooled BMI z-score ($=1.55$, 95% CI: -0.30 , 3.39) than that combining cohorts of 862 patients with average/above-average BMI ($=0.57$, 95% CI: −0.27, 1.42). The same pattern persisted beyond treatment completion (Figure 3C).

DISCUSSION

This systematic review synthesized the longitudinal evidence for the growth patterns in patients with pediatric ALL during and after treatment. The meta-analysis demonstrated a significant increase in mean BMI z-score during treatment in 1,514 patients with pediatric ALL, and increases in weight persist beyond treatment completion. The meta-analysis also found unhealthy weight gain is prevalent in patients with pediatric ALL regardless of receipt of CRT, sex, and patients' weights status.

Survivors of pediatric ALL have significantly high risk for obesity (5) but we did not know at which time period they experienced excessive weight gain. This meta-analysis found that patients with pediatric ALL gained substantial weight during early treatment (i.e. from diagnosis to start of maintenance), and again during the maintenance phase of the treatment. Overall, the mean increase in BMI z-score during treatment was 0.8 in patients with pediatric ALL. For a 5-year old male with an average BMI (i.e. BMI z-score $= 0$ or BMI $=$ 50th percentile) at diagnosis who becomes 7-year old after treatment completion, this increase corresponds to 4 lbs. extra weight gain compared to his peers assuming normal growth in height. These findings support the notion that cancer treatment period is a sensitive window for unhealthy weight gain in patients with pediatric ALL.

Importantly, the meta-analysis suggested that patients with pediatric ALL did not return to their pre-treatment weight and were consistently more overweight or obese than their peers beyond treatment completion. From diagnosis to several years post treatment completion, the mean increase in BMI z-score remained high, ranging from 0.5 to 0.9. These findings, taken together, suggest that excessive weight gain occurred early in treatment is unlikely to be reversed after children complete cancer treatment.

Cancer and its treatment can affect energy intake and expenditure through complex pathways. CRT has been associated with obesity in pediatric ALL survivors but CRT has been gradually replaced by intrathecal and systemic chemotherapy. The meta-analysis supports the notion that patients treated under modern protocols that do not involve CRT still experience substantial weight gain during and after treatment. Corticosteroids are

known to be critically involved in regulating energy intake, storage, and mobilization. Prolonged use of corticosteroids has demonstrated effects on increasing energy intake and percentage of body fat in patients with pediatric ALL (37, 38). Strategies to reduce the use of steroids during maintenance may ameliorate some of the unhealthy weight gain seen during this treatment period. Some of the international pediatric oncology groups have omitted steroids completely during maintenance (39) while others are investigating a reduction in steroids during maintenance (40). Other chemotherapeutic agents such as anthracyclines and vincristine may also contribute to weight gain through impairing cardiovascular fitness and muscle strength and subsequently reducing levels of physical activity (41–43). Although these changes were originally thought to be acute responses to cancer treatment, recent evidence indicates that children had difficulty reversing unhealthy eating habits and sedentary behaviors after treatment completion (41, 44, 45). Cancer treatment may have long-lasting impacts on survivors' dietary intake and physical activity, contributing to unhealthy weight gain in this population.

Patients with below-average BMI at diagnosis seemed to experience a larger degree of weight gain during and after treatment compared to those with average or above-average BMI at diagnosis, which may reflect a greater impact of cancer treatment on "catch-up growth" for those who were leaner at diagnosis. Nevertheless, we did not find that the overall trajectory of changes in BMI z-score differed significantly by receipt of CRT, sex, and weight status during and after treatment, although our meta-analysis was limited by the number of longitudinal studies with available data for longer periods of follow-up to be included.

It is important to note that cancer treatment only attributes partially, less than 50%, to the elevated obesity risk in childhood cancer survivors (3). It is critical to further improve our understanding of modifiable behaviors such as nutritional intake and physical activity in association with obesity risk in childhood cancer survivors, and in particular, to elucidate the immediate and long-term effect of cancer treatment on hypothalamic-pituitary function and subsequent energy balance profiles by different patient and treatment characteristics.

Limitations should be considered when interpreting our findings. Our systematic review comprised heterogeneous studies that included patients from different countries, treated under different protocols, and with different weight status at diagnosis. Studies were also conducted over a period of time over which the treatment protocols have changed. We explored this clinical heterogeneity with subgroup analyses and accounted for unexplained variability through random effects models. Our meta-analysis did not find substantial differences in changes in BMI z-scores when comparing subgroups based on patient- and treatment- related characteristics. However, it is possible that true findings between these subgroups exist, but that our subgroup analyses were under-powered to detect them. Although BMI is widely used to measure obesity, it may not correspond well to the degree of body fatness in children.(46) A recent study suggested that the use of BMI led to underestimation of body fatness in adult survivors of childhood cancer. (47)

In conclusion, our meta-analysis suggests that patients with pediatric ALL are at risk becoming overweight or obese early in treatment, and increases in weight are likely to be

maintained throughout treatment and beyond. Excessive weight gain occurs in all patients with pediatric ALL across treatment received, sex, and weight status at diagnosis. Given that approximately 90% of children and adolescents treated for ALL will be cured, our finding have important implications for pediatric oncologists, general pediatricians, and internal medicine/family medicine physicians, all of whom will provide long-term care to the growing cohort of childhood cancer survivors. Our results strongly suggest the needs for incorporating weight management early in treatment to prevent the early onset of obesityrelated morbidities in patients with pediatric ALL.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

Acknowledgments

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Figure 2. Changes in BMI z-score during and after treatment in patients with pediatric ALL

Δ (95% confidence interval) represents pooled random-effects meta-analysis results for the mean change in BMI z-score since diagnosis at each time period during and after treatment. $I²$ quantifies the extent of heterogeneity across studies. n represents the number of studies included in each meta-analysis.

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Figure 3. Changes in BMI z-score during and after treatment in patients with pediatric ALL by patient and treatment characteristics

A. Changes in BMI z-score by receipt of cranial radiation therapy (CRT). **B.** Changes in BMI z-score by sex. **C**. Changes in BMI z-score by weight status at diagnosis.

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Characteristics of studies that assessed changes in BMI z-score during and after treatment in patients with pediatric ALL Characteristics of studies that assessed changes in BMI z-score during and after treatment in patients with pediatric ALL

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Table II

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 b Mean and SD of the BMI z-score were obtained based on graphs using Digitizelt *b*Mean and SD of the BMI z-score were obtained based on graphs using DigitizeIt

 \int Median of the BMI z-score was presented because mean was not reported. *c*Median of the BMI z-score was presented because mean was not reported.