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Perioperative blood transfusion in gynecologic oncology surgery: Analysis of the National Surgical Quality Improvement Program Database

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Abstract

Objective—To use a large-scale multi-institutional dataset to quantify the prevalence of packed red blood cell transfusions and examine the associations between transfusion and perioperative outcomes in gynecologic cancer surgery.

Methods—The American College of Surgeons National Surgical Quality Improvement Program (NSQIP) participant use file was queried for all gynecologic cancer cases between 2010 and 2012. Demographic, preoperative and intraoperative variables were compared between transfusion and non-transfusion groups using chi-squared, Fisher's exact and Wilcoxon rank-sum test. The primary endpoint was 30-day composite morbidity. Secondary endpoints included composite surgical site infections, mortality and length of stay.

Results—A total of 8,519 patients were analyzed, and 13.8% received a packed red blood cell transfusion. In the multivariate analysis, after adjusting for key clinical and perioperative factors, including preoperative anemia and case magnitude, transfusion was associated with higher composite morbidity (OR = 1.85, 95% CI 1.5 – 2.24), surgical site infections (OR 1.80, 95% CI

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1.39 – 2.35), mortality (OR 3.38, 95% CI 1.80 – 6.36) and length of hospital stay (3.02 days v. 7.17 days, $p < 0.001$).

Conclusions—Blood transfusions are associated with increased surgical wound infections, composite morbidity and mortality. Based on our analysis of the NSQIP database, transfusion practices in gynecologic cancer should be scrutinized. Examination of institutional practices and creation of transfusion guidelines for gynecologic malignancies could potentially result in better utilization of blood bank resources and clinical outcomes among patients.

INTRODUCTION

Blood is a precious, costly resource that is often over utilized and transfused with great variation in clinical practice. According to the US Department of Health and Human Services there were more than 13.5 million units of blood transfused in 2011 at an average cost of \$225.42/unit [1]. Perioperative surgical transfusion rates in gynecologic oncology patients fluctuate greatly with some studies reporting rates as low as 3% [2] and others as high as 77% [3–5]. This wide variation in practice patterns may be attributed to vague clinical practice guidelines combined with conflicting data in cancer patients.

Several large randomized controlled trials have suggested that a more restrictive transfusion protocol in surgical and critically ill patients is associated with improved clinical outcomes [6–10]. Although there have been no randomized controlled trials in oncology patients, there is ample evidence in the colorectal cancer surgery literature to suggest that blood transfusions themselves are immunosuppressive and associated with increased rates of infection, perioperative morbidity, disease progression and mortality [11, 12].

There is compelling evidence that questions the liberal use of blood transfusion in colorectal surgery and critically ill patients; however, uncertainties remain about the application of these data to gynecologic cancer patients. There are limited data examining the effects of blood transfusions on perioperative outcomes after gynecologic cancer surgery. Furthermore, to date, most of the studies in gynecologic cancer have been single-institution studies evaluating outcomes in a single disease site such as cervix or ovary. Awareness of national blood transfusion practices in gynecologic oncology could potentially result in better utilization of blood bank resources and both short- and long-term clinical outcomes among patients. We hypothesized that blood transfusions are associated with increased morbidity in gynecologic oncology surgical patients. We used a large-scale multi-institutional dataset, the National Surgical Quality Improvement Program database, to quantify the prevalence of perioperative blood transfusion and examine the effect of transfusion on perioperative outcomes.

MATERIALS AND METHODS

The American College of Surgeons National Surgical Quality Improvement Program (ASC-NSQIP) is a multi-institutional comprehensive database containing perioperative information on surgical patients. Trained risk-assessment nurses in participating hospitals collect preoperative patient characteristics, intraoperative data and 30-day morbidity and mortality. The specific methodology has been reported previously [13]. De-identified patient

information is available to all participating institutions through the ASC-NSQIP participant use file (PUF).

The ASC-NSQIP PUF was queried for all gynecologic cases between 2010 and 2012 and limited to cases with ICD-9 codes associated with malignant gynecologic neoplasms (vulva, vagina, cervix, uterus, and ovary). CPT codes for which the transfusion rate was zero were excluded based on the findings by Bernard *et al.*[14] Extreme outliers were excluded from the analysis which included patients with preoperative transfusion greater than 4 units, those undergoing emergency procedures, pelvic exenteration, or procedures with operative time less than 30 minutes.

A total of 8,519 cases were included for analysis. The demographic data assessed included: age, body mass index (BMI), ethnicity (Hispanic and non-Hispanic) and race (white, black, other). Medical comorbidities and risk factors analyzed included: American Society of Anesthesiologists (ASA) class, presence of disseminated cancer, presence of ascites, receipt of neoadjuvant chemotherapy within 30 days of surgery, smoking, steroid use, hypertension requiring medication management, dyspnea, COPD, disease site (uterus, ovary, vagina/vulva, or cervix), preoperative bleeding disorders and more than 10% body weight loss in last six months. Perioperative factors evaluated included: preoperative labs (including hematocrit, INR, platelets and albumin), operating time, anesthesia time, procedure complexity, wound classification and procedure type. Procedure complexity was assessed by using total work relative value scales (WRVU), which has been previously shown to be an appropriate surrogate marker for surgical complexity [13]. Procedure type was defined as minimally invasive surgery (MIS) or open. Perioperative variables with less than 1% incidence were excluded.

Patients were divided into two groups: those who received a blood transfusion and those who did not receive a blood transfusion. The variable for transfusion includes those patients who received a transfusion in the operating room until up to 72 hours post-operatively. The primary endpoint for the study was 30-day composite morbidity (based on the occurrence of 1 or more of the 20 adverse events defined by NSQIP, excluding transfusion, which are listed in Figure 1). Secondary endpoints examined were: 30-day composite infectious morbidity (superficial, deep or organ/space surgical site infections), the 20 adverse events defined by NSQIP, mortality, and length of stay.

Summary statistics were used to describe demographic, preoperative and intraoperative variables. Chi-squared and Fisher's exact tests were used to test for differences between those who received a blood transfusion and those who did not receive a blood transfusion for categorical variables, and the Wilcoxon rank-sum test was used to compare medians between groups for continuous variables.

Univariate logistic regression was used to model the logit of the probability of composite morbidity as a function of whether or not a patient received a transfusion and several other potential prognostic factors. A saturated model including all factors with a $P < 0.20$ was built and backward elimination was used in a multivariate analysis to construct a parsimonious model, removing factors one at a time until all remaining factors remained

statistically significant. Preoperative hematocrit was retained as a continuous variable in all models. Adjusted odds ratios and corresponding 95% confidence intervals for each factor remaining in the model are reported. $P < 0.05$ was considered statistically significant. This modeling strategy was repeated for composite surgical site infections (SSI). However, since there were only 53 events for mortality, a forward selection strategy was used to build a multivariate model. The model began with transfusion (yes/no), and then the factor with the smallest P was added and the model was refit. All factors with $P < 0.05$ were retained and this process was repeated until no remaining factors could enter the model. This strategy avoided overfitting the model. All analyses were performed using STATA™ 13.0 for Macintosh (StatCorp LP, College Station, Texas). The study was approved by The University of Texas MD Anderson Cancer Center Institutional Review Board.

RESULTS

We identified 8,906 patients with the diagnosis of gynecologic malignancy in the NSQIP database. Three hundred eighty-seven patients were excluded for the following reasons: emergency case ($n=61$), exenterative procedure ($n=88$), preoperative transfusion more than 4 units ($n=79$), operative time less than 30 min ($n=74$) and CPT codes with transfusion rate of zero ($n=85$). Of the 8,519 patients who met our inclusion criteria, 1178 or 13.8% (95% CI 13.1% – 14.6%) received a blood transfusion within 72 hours of surgery.

Procedures were grouped according to primary CPT code and organ system (Table 1). Laparoscopy was the most common procedure performed ($n=3,916$), followed by open abdominal hysterectomy ($n=2,483$) then laparotomy for tumor reductive surgery ($n=1,773$). Laparotomy associated with a tumor reductive surgery had the highest propensity for blood transfusion with 35.08% of patients receiving at least one transfusion, followed by vaginectomy (23.53%) and laparotomy for adnexal surgery (18.18%). Laparoscopy was associated with the lowest likelihood of having a transfusion (2.32%).

Comparison of the demographics and preoperative characteristics of patients who received a blood transfusion and those who did not are displayed in Table 2. Compared to those patients who did not receive a transfusion, patients who were transfused were more likely to be older, thinner, non-white, have a higher ASA class, have disseminated cancer, dyspnea, ovarian cancer, a bleeding disorder. Comparison of preoperative laboratory variables between patients who received a blood transfusion and those who did not are displayed in Table 3. Compared to those patients who did not receive a transfusion, patients who were transfused were more likely to have a lower preoperative hematocrit, and a preoperative albumin less than 3 ($P < 0.001$ for all). While our primary interest was to evaluate the association of morbidity with transfusion use, we considered the above factors in our multivariate analysis of morbidity in an effort to account for potential bias in differences between those patients who did and did not receive transfusions. Importantly, we accounted for presence of disseminated cancer and preoperative anemia in our multivariate analysis. Table 4 displays the comparison of intraoperative variables in patients who were transfused and those who were not. Compared to those patients who did not receive a transfusion, patients who were transfused were more likely to have longer operating time, longer

anesthesia time, increased surgical complexity, contaminated or dirty wounds and undergone laparotomy ($P < 0.001$ for all).

In the univariate analysis, transfusion was associated with higher composite morbidity (9.06% v. 25.13%, OR 3.37; 95% CI 2.89 – 3.93), increased composite surgical site infections (4.05% v. 10.95%, OR 2.92; 95% CI 2.35 – 3.62), and increased mortality (0.37% v. 2.21%, OR 6.11; 95% CI 3.56 – 10.50). Transfusion was also associated with increased rates of wound disruption, pneumonia, unplanned intubations, pulmonary embolism, ventilator use for longer than 48 hours, progressive and acute renal failure, urinary tract infections, deep vein thrombosis or thrombophlebitis, sepsis and septic shock ($P < 0.01$ for all). A forest plot of risk of 30-day postoperative outcomes by transfusion status is presented in Figure 1. The ORs and 95% CIs for the NSQIP 30-day adverse events are displayed in this figure. Hospital length of stay was increased by 4.15 days for those receiving transfusions (3.02 days v. 7.17 days, $P < 0.001$).

All perioperative characteristics with a $P < 0.20$ were taken into consideration in the multivariate analysis and included: BMI, race, ASA classification, presence of disseminated cancer, ascites, tobacco use, steroid use, hypertension, dyspnea, cancer type, bleeding disorder, weight loss, procedure complexity, wound class, preoperative hematocrit, preoperative platelet count and procedure type. Preoperative hematocrit was modeled as both a categorical and continuous variable in the multivariate analysis for morbidity and surgical site infection without a significant difference in magnitude of odds ratios; therefore, preoperative hematocrit was considered as a continuous variable in all models. Forest plots displaying multivariate logistic regression analysis for composite morbidity, composite surgical site infection and mortality are presented in Figure 2. After adjusting for the key clinical and perioperative factors which included: BMI, ASA class, tobacco use, hypertension, disease type, bleeding disorder, procedure complexity, OR time, wound classification, preoperative hematocrit and procedure type, as shown in a forest plot in Figure 2, transfusion was associated with 1.85 times increased odds of composite morbidity (OR 1.85; 95% CI 1.53 – 2.24). In the risk-adjusted analysis we also found that blood transfusions increased the odds of composite SSI (OR 1.80; 95% CI 1.38 – 2.34) and mortality (OR 3.38; 95% CI 1.80 – 6.35).

We performed several subgroup analyses to evaluate the effect of blood transfusion on different cancer types and procedure classifications. The directionality and statistical significance of the association of transfusion with increased morbidity were the same regardless of the disease type or procedure classification. In the subset of patients with ovarian cancer, transfusion was associated with a 2.68 times increased odds of composite morbidity (OR 2.68; 95% CI 2.10 – 3.42) after adjusting for BMI, presence of ascites, hypertension, bleeding disorder, procedure complexity, OR time, preoperative hematocrit and procedure type. In those patients who had undergone a tumor reductive surgery, transfusion was associated with a 1.77 times increased odds of composite morbidity (OR 1.77; 95% CI 1.33 – 2.35) after adjusting for the presence of ascites, hypertension, procedure complexity, operating time and preoperative hematocrit.

DISCUSSION

Our results suggest that blood transfusions are associated with increased morbidity, surgical site infections and mortality in gynecologic oncology patients. Perioperative blood transfusions occurred in 13.8% of patients in the NSQIP database. This rate is on the lower end of reported transfusion rates in the gynecologic oncology literature and is likely attributed to the high percentage of minimally invasive cases in this database. Minimally invasive cases accounted for 47.51% of cases in our database, but only 2.37% of transfusions. Given that postoperative transfusion is increasingly being used as a quality metric, it is critical that risk-adjusted models account for surgical approach and other important risk factors that we identified including disseminated cancer, preoperative anemia and disease site.

Anemia in gynecologic cancer patients is common and its etiology multifactorial. Animal and human studies have shown that anemia is an indicator of poor prognosis [15–18]. In 1965, Evans was one of the first researchers to show the association between anemia and decreased survival in cervix cancer [15]. Since then, there have been numerous experimental studies demonstrating worse oncologic outcomes in anemic patients [15, 19]. Knowing that anemic patients have worse outcomes, there has traditionally been a trend towards liberal use of blood products to correct anemia in gynecologic oncology patients. In two recent studies, more than half of ovarian cancer patients received at least one perioperative transfusion [3, 20]. However, there is a dearth of data in the gynecologic cancer population supporting the concept that preemptive or perioperative transfusion actually improves clinical outcomes.

Several large randomized controlled trials have suggested that a more restrictive transfusion protocol in surgical and critically ill patients is associated with improvement in short- and long-term clinical outcomes. The Transfusion Requirements in Critical Care (TRICC) [7] study showed a lower mortality rate in ICU patients who were randomized to the restrictive transfusion group (transfused for Hb < 7 g/dL) compared to patients in the liberal group (transfused for Hb < 10 g/dL). Similarly, the Transfusion Trigger Trial for Functional Outcomes in Cardiovascular Patients Undergoing Surgical Hip Fracture Repair (FOCUS) [21] trial showed no difference in mortality in a cardiovascular high-risk cohort undergoing hip surgery randomized to restrictive (transfused for Hb < 8 g/dL) or liberal (transfused for Hb < 10 g/dL) approach. A third randomized controlled trial of patients with acute upper gastrointestinal bleeding revealed that a restrictive strategy (transfused for Hb < 7 g/dL) was associated with decreased mortality and adverse events compared to rates in the liberal group (transfused for Hb < 9 g/dL) [22]. Expert opinion supports a restrictive transfusion protocol utilizing hemoglobin triggers of 7 – 8g/dL[23–25].

Our data are consistent with previous studies in colorectal surgery using the NSQIP database. Bernard and colleagues demonstrated that one unit of packed red blood cells significantly increased the risk of morbidity, mortality, pneumonia and sepsis [14]. Halabi *et al.* confirmed these findings in a more recent analysis and demonstrated that blood transfusions were independently associated with increased morbidity, mortality, length of stay and SSI [11].

Although the data in colorectal surgery are quite compelling, there is currently minimal evidence available to guide practitioners regarding the effect of transfusion on perioperative outcomes in gynecologic oncology. In 2005, Abu-Rustum and colleagues at Memorial Sloan-Kettering Cancer Center demonstrated that transfusion was associated with increased venous thromboembolism risk [26]. Boone *et al.* reported that, after the adoption of a restrictive transfusion protocol at the University of Alabama, there was no significant difference in postoperative infections, thrombotic events or mortality between those patients who received an appropriate transfusion and those who were inappropriately transfused [27]. Bakkum-Gamez *et al.* demonstrated that receipt of an intraoperative PRBT was associated with an increased rate of surgical site infections [28].

While the pathogenesis of increased morbidity with transfusion cannot be answered by our study, there have been several theories that have proposed physiologic mechanisms for this effect. One proposed mechanism relates to the immunomodulatory effects of blood transfusions. Alloantigens present on donor blood cells elicit cytokine mediators and cellular responses that influence the inflammatory response system [29]. Blumberg *et al.* demonstrated that transfused patients had significant changes in immunologic laboratory tests such as decreased IL-2 production, decreased CD4 and natural killer cells as well as increased macrophage prostaglandin E2 production [30, 31]. Thus, it has been suggested that this alteration in immune function may predispose individuals who receive perioperative blood transfusion to postoperative complications and disease progression. Furthermore, evidence from the surgical oncology transfusion-related immunomodulation literature suggests a dose-dependent increase in perioperative adverse outcomes. So although the adverse impact of transfusion occurs with even one unit, increasing units amplifies this effect [14].

Our study is limited by several factors. First, the study design is retrospective and perioperative differences exist between the two study groups. Second, we are limited to the data collected by NSQIP and therefore are inherently missing important clinical variables that have been shown to impact outcomes, such as time from diagnosis until surgery and cancer stage. While we do not have time from diagnosis until surgery, the presence of metastatic disease has been shown to as a surrogate for advanced cancer stage and more important for predicting 30-day morbidity and mortality in NSQIP studies [32, 33]. NSQIP also provides limited data on important clinical variables that may influence the receipt of blood transfusion such as estimated blood loss, hematocrit nadir and postoperative hematocrit. The database is incomplete as several fields are not mandatory; this results in a high percentage of missing values, especially for laboratory values. NSQIP has also changed its definitions for transfusion variables. Starting in 2010, NSQIP combined intraoperative and postoperative transfusion into one category and stopped collecting number of units transfused. Thus, we have only considered data from 2010–2012, but unfortunately have access to neither total units of blood transfused nor the breakdown of intraoperative and postoperative transfusion. Furthermore, NSQIP only collects data on blood transfusions up to 72 hours post-operatively; therefore, there may be individuals in the non-transfusion group who actually received a blood transfusion later in their hospital course.

To our knowledge, this study is the first evaluation of perioperative outcomes with respect to blood transfusions in gynecologic oncology patients using a national dataset. Our study found that, similar to the colorectal literature, transfusions are associated with increased surgical morbidity, infections and perioperative mortality in gynecologic cancer patients even after controlling for the presence of disseminated cancer and preoperative anemia. The American Medical Association's Physician Consortium for Performance Improvement and the Joint Commission have both identified blood transfusions as one of the five treatments that are overused. Our study contributes to the compelling evidence from other surgical specialties that transfusions are associated with increased length of hospital stay and perioperative complications. Based on our analysis of the NSQIP database, transfusion practices in gynecologic cancer should be scrutinized. Our data suggest that individuals and institutions who have not previously modified their practice to a restrictive blood transfusion program should do so to minimize unnecessary transfusions. In addition, the creation and adoption of restrictive transfusion guidelines for gynecologic cancer surgery will likely yield better utilization of blood bank resources and improve clinical outcomes among patients.

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Research Highlights

- 13.8% of gynecologic surgical patients received a perioperative blood transfusion.
- Transfusion is independently associated with increased perioperative morbidity.
- Transfusion increases risk of perioperative mortality and surgical site infections.

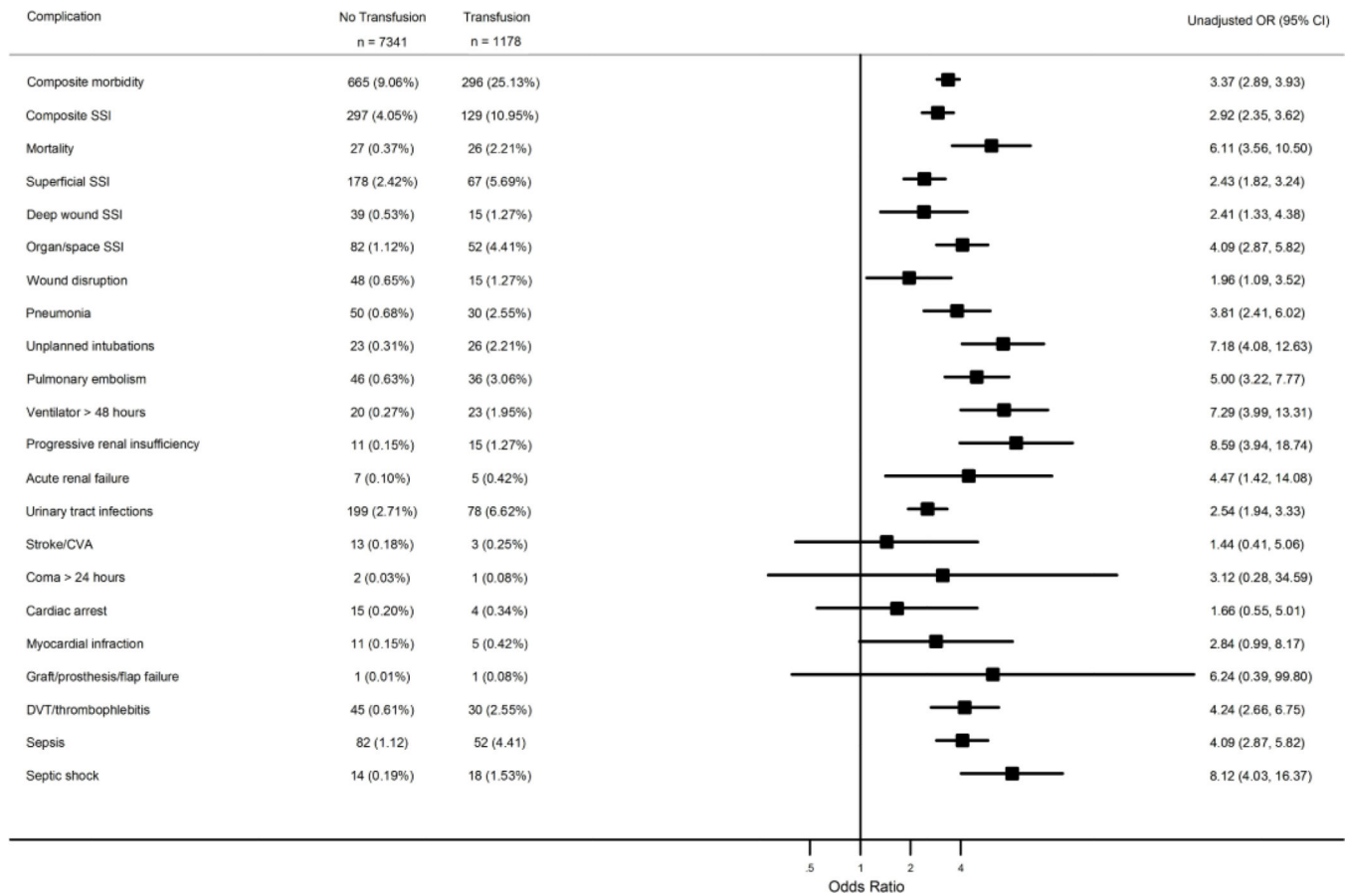


Figure 1. Forest plot of risk of 30-day postoperative outcomes by transfusion status. Odds ratios and 95% confidence intervals from univariate analysis of 30-day postoperative outcomes as a function of risk group (transfusion status), with number (%) of events by risk group

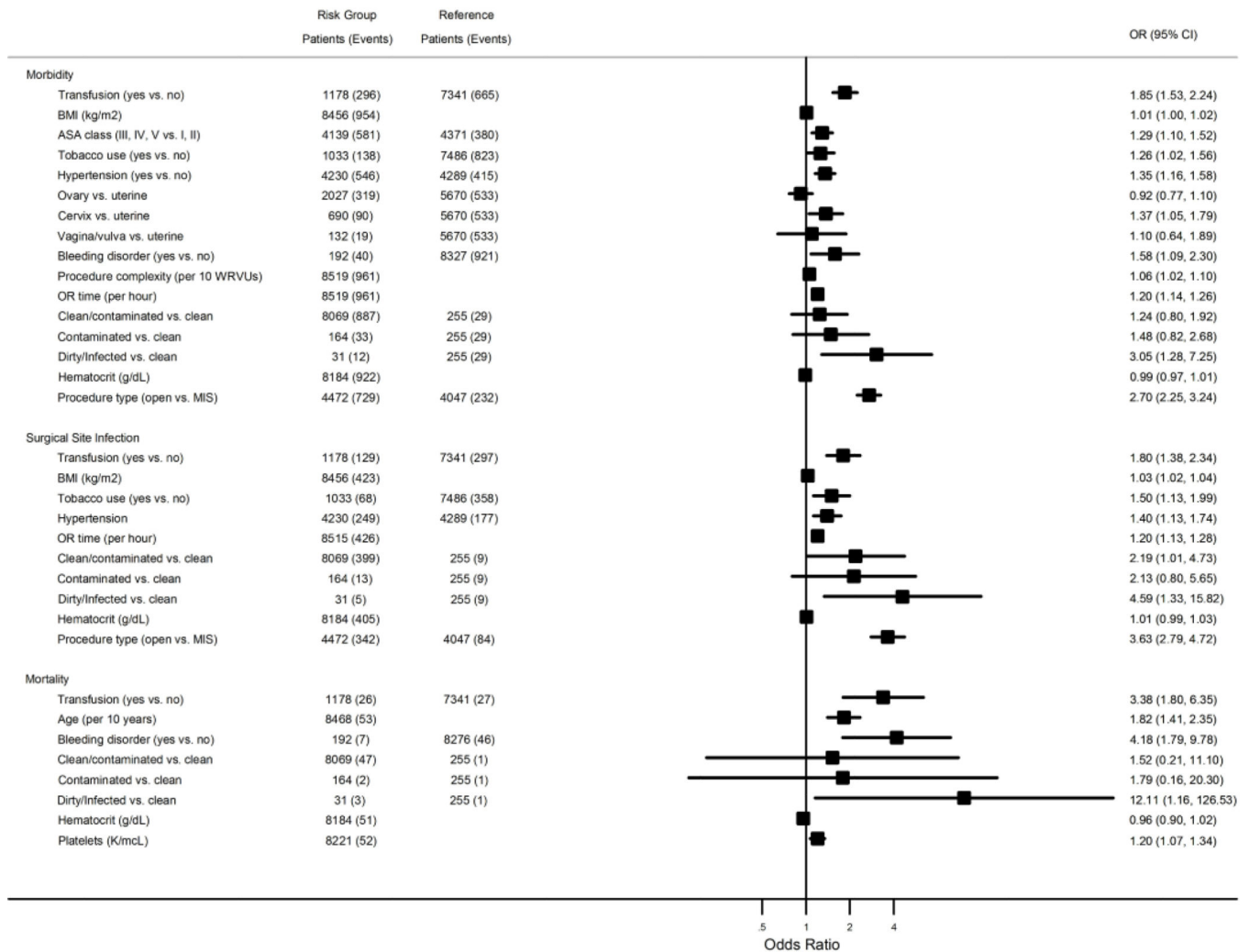


Figure 2. Forest plots displaying multivariate logistic regression analysis for composite morbidity, composite surgical site infection and mortality. Adjusted odds ratios and 95% confidence intervals from multivariate analysis of morbidity, composite surgical site infection, and mortality as a function of various risks, with number of patients and events in each risk group

Table 1

Procedure groups included in analysis by CPT code

Organ/system	CPT	N	Patients Transfused, N (%)
Vulvectomy	56630 – 56637	100	8 (8.00%)
Vaginectomy	57106 – 57111	17	4 (23.53%)
Trachelectomy	57530 – 57531	15	2 (13.33%)
Laparotomy, Hysterectomy	58150 – 58210	2483	426 (17.16%)
Vaginal Surgery, Hysterectomy	58260 – 58285	131	5 (3.82%)
Laparoscopy	58542 – 58573	3916	91 (2.32%)
Laparotomy, adnexal surgery	58720	33	6 (18.18%)
Laparotomy, Tumor Reductive Surgery	58950 – 58960	1773	622 (35.08%)

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Table 2

Demographic and preoperative characteristics in patients who were transfused and those who were not transfused

	No Transfusion group (n = 7341)	Transfusion group (n = 1178)	p-value*
Age, years			<0.001
Mean ± sd	60.77 ± 12.24	62.50 ± 12.64	
Median (range)	61 (18 – 89)	63 (19 – 89)	
Missing	43	8	
BMI, kg/m ²			<0.001
Mean ± sd	32.80 ± 9.53	29.72 ± 8.56	
Median (range)	31 (10.97–102.50)	27.63 (15.79–76.89)	
Missing	52	11	
Race			
White	5,749 (88.38%)	845 (82.36%)	<0.001
Black	510 (7.84%)	115 (11.21%)	<0.001
Other	246 (3.78%)	66 (6.43%)	<0.001
Missing	836	333	
Ethnicity			
Hispanic	490 (6.67%)	72 (6.11%)	0.470
ASA class			<0.001
I	261 (3.56%)	25 (2.12%)	
II	3631 (49.46%)	454 (38.54%)	
III	3248 (44.24%)	645 (54.75%)	
IV	192 (2.62%)	52 (4.41%)	
V	1 (0.01%)	1 (0.08%)	
None assigned	8 (0.11%)	1 (0.08%)	
Disseminated cancer	894 (11.38%)	284 (42.84%)	<0.001
Presence of ascites	179 (2.44)	38 (3.23%)	0.111
Chemotherapy < 30 days	69 (0.94%)	60 (5.09%)	0.421
Tobacco Use	902 (12.29)	131 (11.12%)	0.255
Diabetes	1201(16.37)	207 (17.57%)	0.304
Steroid use	124 (1.69)	28 (2.38%)	0.098
Hypertension	3662 (49.88%)	568 (48.22%)	0.288
Dyspnea	511 (6.96%)	130 (11.04%)	<0.001
COPD	179 (2.44%)	38 (3.23%)	0.111
Cancer			<0.001
Uterus	5239 (71.37%)	431 (36.59%)	
Ovary	1369 (18.65%)	658 (55.86%)	
Vagina/Vulva	118 (1.61%)	14 (1.19%)	
Cervix	615 (8.38%)	75 (6.37%)	
Bleeding disorder	140 (1.91%)	52 (4.41%)	<0.001
>10% weight loss in last 6 months	87 (1.19%)	64 (5.43%)	<0.001

BMI = body mass index, ASA = American Society of Anesthesiologists class; COPD = chronic obstructive pulmonary disease

* Chi-squared and Fisher's-exact test for categorical variables and the Wilcoxon rank-sum test to compare medians between groups for continuous variables

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Table 3

Comparison of preoperative laboratory variables in patients who were transfused and those who were not transfused

	No Transfusion group (n = 7341)	Transfusion group (n = 1178)	p-value*
Hematocrit, g/dL			<0.001
36	5698 (80.87%)	484 (42.53%)	
30 < 36	1179 (16.73%)	463 (40.69%)	
> 21 – < 30	157 (2.23%)	186 (16.34%)	
21	12 (0.17%)	5 (0.44%)	
Missing	295	40	
INR			0.040
1.5	3277 (98.70%)	669 (97.66%)	
>1.5	43 (1.30%)	16 (2.34%)	
Missing	4021	493	
Albumin			<0.001
3	129 (3.01%)	118 (14.55%)	
> 3	4160 (96.99%)	693 (85.45%)	
Missing	3052	367	

BMI = body mass index, ASA = American Society of Anesthesiologists class; COPD = chronic obstructive pulmonary disease

* Chi-squared and Fisher's-exact test for categorical variables and the Wilcoxon rank-sum test to compare medians between groups for continuous variables

Table 4

Intraoperative variables in patients who were transfused and those who were not transfused

	No Transfusion group (n = 7341)	Transfusion group (n = 1178)	p-value*
Operation time, minutes			<0.001
Mean ± sd	172.94 ± 78.81	220.04 ± 102.96	
Median (range)	160 (30 – 1028)	203 (31 – 919)	
Missing	4	0	
Anesthesia Time, minutes			<0.001
Mean ± sd	228.76 ± 93.72	286.99 ± 111.34	
Median (range)	212 (52 – 858)	273.50 (71 – 910)	
Missing	4426	686	
Procedure complexity			<0.001
Mean ± sd	32.31 ± 14.14	47.97 ± 25.28	
Median (range)	31.63 (5.27 – 199.55)	40.05 (12.16 – 187.34)	
Wound Classification			
Clean	200 (2.72%)	55 (4.67%)	0.001
Clean/Contaminated	7010 (95.49%)	1059 (89.90%)	<0.001
Contaminated	113 (1.54%)	51 (4.33%)	<0.001
Dirty/Infected	18 (0.25%)	13 (1.10%)	<0.001
Procedure type			<0.001
MIS	3951 (53.82%)	96 (8.15%)	
Open	3390 (46.18%)	1082 (91.85%)	

* Chi-squared and Fisher's-exact test for categorical variables and the Wilcoxon rank-sum test to compare medians between groups for continuous variables