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# Association between diabetes and amyotrophic lateral sclerosis in Sweden

Daniela Mariosa, MS<sup>1,\*</sup>, Freya Kamel, PhD, MPH<sup>2</sup>, Rino Bellocco, PhD<sup>1,3</sup>, Weimin Ye, MD, PhD<sup>1</sup>, and Fang Fang, MD, PhD<sup>1</sup>

<sup>1</sup>Department of Medical Epidemiology and Biostatistics, Karolinska Institutet, Stockholm, Sweden

<sup>2</sup>Epidemiology Branch, National Institute of Environmental Health Sciences, Research Triangle Park, NC, USA

<sup>3</sup>Department of Statistics and Quantitative Methods, University of Milano-Bicocca, Milan, Italy

# Abstract

**Background**—Energy metabolism is altered in patients with amyotrophic lateral sclerosis (ALS) but the role of diabetes is largely unknown.

**Methods**—We conducted a population-based case-control study of 5,108 ALS cases and 25,540 individually matched population controls during 1991–2010. Information on ALS and preexisting diabetes was retrieved from the Swedish Patient Register to explore the association of ALS with diabetes overall and with insulin-dependent or non-insulin dependent diabetes specifically. We also studied variation of the association by diabetes duration and age.

**Results**—In total, 224 ALS cases (4.39%) and 1,437 controls (5.63%) had diabetes before the index date, leading to an overall inverse association between diabetes and ALS risk (OR=0.79, 95%CI=0.68–0.91). The association was strong for non-insulin-dependent diabetes (OR=0.66, 95%CI=0.53–0.81) but not for insulin-dependent diabetes (OR=0.83, 95%CI=0.60–1.15) and varied as a function of diabetes duration, with the strongest association observed around six years after first ascertainment of diabetes. The association was age-specific; the inverse association was noted only among individuals aged 70 or older. In contrast, for younger individuals (<50 years), preexisting insulin-dependent diabetes was associated with a higher ALS risk (OR=5.38, 95%CI=1.87–15.51).

**Conclusions**—Our study suggests that there is an association between diabetes and ALS, and highlights the importance of taking into account age, insulin dependence and diabetes duration. Future studies should explore whether the association is independent of BMI.

# INTRODUCTION

Amyotrophic lateral sclerosis (ALS) is a fatal motor neuron disease. In Sweden the agestandardized incidence rate has increased from 2.32 per 100,000 person-years in 1991–1993

Conflicts of interest: None.

<sup>&</sup>lt;sup>\*</sup>Correspondence to: Daniela Mariosa, Department of Medical Epidemiology and Biostatistics, Karolinska Institutet, Box 281, 17177 Stockholm, Sweden, Phone: +46-8-52486183, Fax: +46-8-314975, Daniela.mariosa@ki.se.

to 2.98 per 100,000 person-years in 2003–2005 [1], but the etiology of ALS remains largely unknown. Probable risk factors include exposure to heavy metals, pesticides, and smoking [2]. Known genetic variants are found in only a minor proportion of ALS cases without clear family history (90–95% of the total) and in two-thirds of familial cases [3].

Other diseases may share risk factors or genetic predisposition with ALS and, additionally, the pathological changes associated with specific diseases may alter future risk of ALS. Interestingly, a beneficial cardiovascular risk profile is associated with both ALS and frontotemporal dementia [4, 5]. Furthermore, a connection between metabolic disorders and ALS is supported by the fact that ALS patients are hypermetabolic and have impaired glucose tolerance [6, 7]. One study found that premorbid type 2 diabetes was associated with a 4-year delayed onset of ALS. However, two recent studies reported null associations with premorbid diabetes, except for a higher ALS risk associated with insulin-dependent diabetes before age 30 [8–10]. A literature review concluded that the role of diabetes in ALS remains unclear [11].

To address this issue, we investigated whether preexisting diabetes affects ALS risk in the Swedish population.

#### METHODS

#### Study design

We performed a nested case-control study using the Swedish Population and Housing Census conducted by Statistics Sweden in 1990. All residents who were born in Sweden and had never emigrated abroad nor been diagnosed with ALS before 1991 were selected for the study. Using the Swedish Patient, Causes of Death, and Migration Registers, we followed the study population from January 1<sup>st</sup>, 1991 to ALS diagnosis, death, emigration or December 31<sup>st</sup>, 2010, whichever came first. Cross-linkages were performed using the Swedish National Registration Numbers [12].

The Regional Ethical Review Vetting Board in Stockholm, Sweden approved the study.

#### ALS cases

We identified individuals with a first ALS diagnosis during follow-up through the Swedish Patient Register. Since 1964/1965, the Swedish National Board of Health and Welfare has compiled data on individual hospital discharges. This Inpatient Register achieved national coverage in 1987. The Outpatient Register contains information on visits to hospital-based specialist care from 2001 onward. Each record in the Swedish Patient Register (for Inpatient and Outpatient Registers collectively) includes discharge diagnoses coded by the Swedish Revisions of the International Classification of Diseases (ICD) codes (ICD-7 before 1969, ICD-8 1969–1986, ICD-9 1987–1996, and ICD-10 since 1997). ALS patients were defined as individuals with a main diagnosis of ALS (ICD-9 335.2 or ICD-10 G12.2). Hospital discharge data on ALS diagnosis are believed to be of high sensitivity and specificity [13].

#### Controls

Using incidence density sampling, we randomly selected five controls per case individually matched on year of birth, sex and area of residence (Northern, Central, or Southern Sweden). Individuals who were missing information on area of residence were excluded from the study. We defined the index date for cases as their date of diagnosis, and for controls as the date of diagnosis of their matched cases. To be eligible, the controls had to be alive and free of ALS on the index date.

#### Exposure

We assessed history of diabetes from 1987 to the index date through the Swedish Patient Register. Since 1987 this Register has had nationwide coverage for all hospital discharges and since 2001 it has also covered >80% hospital-based specialist care visits. We identified diagnoses of diabetes using the codes ICD-9 250 and ICD-10 E10-E11. In a sensitivity analysis, we also used information on use of antidiabetic medications extracted from the Swedish Prescribed Drug Register in order to include additional diabetes patients. This was done for the subsample with index dates between 2006 and 2010 since the Prescribed Drug Register contains information on all dispensed prescriptions of the Swedish population from July 1<sup>st</sup>, 2005 onward. All drugs are coded according to the Anatomical Therapeutic Chemical Classification System (ATC codes). Antidiabetic drugs are coded A10A or A10B.

We classified diabetes onset as early if the age at first ascertainment was 30 years or as late otherwise. Diabetes duration was defined as the time interval between the first available diabetes record in the Patient Register and the index date. In the main analyses, these two variables were calculated using information available after 1987. Since restricting ascertainment to 1987 onward might underestimate the duration and overestimate the age at diabetes onset, in a sensitivity analysis we further identified diabetes patients from records during the years 1964–1986 when the Inpatient Register was available but not yet nationwide (ICD-7 260, ICD-8 250).

We categorized patients as insulin-dependent if their ICD-10 code was E10 and as noninsulin-dependent if it was E11. In a sensitivity analysis, we also assigned insulin-dependent status to patients who had ever been prescribed insulin (ATC code A10A), and non-insulindependent status to those who were prescribed other drugs (ATC code A10B). The analyses on insulin dependence excluded individuals with unknown insulin-dependence status (i.e., absence of ICD-10 codes or with both ICD-10 E10 and E11).

Because diabetes without complications was recorded with specific ICD codes (ICD-9 250.0, ICD-10 E10.9 and E11.9), we defined cases of diabetes with complications as those with at least one diagnosis of diabetes with complications before the index date.

#### Covariates

The Swedish Population and Housing Census in 1990 provided information on socioeconomic status (Blue collar, White collar, Self-employed/Farmers, Other). The Swedish Education Register was established in 1985 by Statistics Sweden and annually collects information on the highest level of formal education attained for each Swedish

resident. Education achieved before the index date was categorized as "9 years", "9–12 years" and "University or doctoral studies".

#### Statistical analysis

Analyses were performed using Stata ver. 12.1 (StataCorp, Texas, USA). We fitted conditional logistic regression models to estimate the odds ratios (ORs) for the association of ALS with diabetes, with 95% confidence intervals (CIs). We adjusted estimates for matching variables only or additionally for education and socioeconomic status. We repeated control sampling 100 times and plotted the observed distribution of the ORs. We performed analyses stratified by age at index date (<50, 50–59, 60–69, 70–79, 80+ years), sex, and calendar period of index date (1991–2000 versus 2001–2010). We tested multiplicative interaction between diabetes and age at index date by adding a continuous interaction term to the models. Departure from additivity was assessed by computing the Relative Excess Risk due to Interaction (RERI). Given that pre-clinical ALS symptoms might lead to intensified diagnostic evaluation, we repeated the analysis disregarding diabetes diagnosed during the 3-year time window before index date.

To study the impact of insulin-dependence on the association between diabetes and ALS, the subjects were classified as "No diabetes", "Insulin-dependent diabetes" or "Non-insulin-dependent diabetes". To estimate the effect of diabetes duration on ALS risk, we employed dummy variables (0–2, 2–4, 4–6, 6–8, 8–10, 10+ years) and cubic regression splines with 5 knots (percentiles: 5, 27.5, 50, 72.5, 95).

### RESULTS

Among the 7,673,847 Swedish residents followed from 1991 to 2010, we identified 5,108 newly diagnosed cases of ALS and selected 25,540 matched controls.

#### Main analyses

Descriptive statistics (Table 1) showed that ALS cases were more educated (Chi squared test p=0.02) and more likely to be self-employed or white collar (p=0.02) than controls of the same age, sex and area of residence.

In total, 5.63% of controls and 4.39% of cases had diabetes before the index date. Diabetes was inversely associated with ALS; the result was similar after further adjustment for education and socioeconomic status (Table 2). After repeating control sampling 100 times, the median multivariable OR was 0.80 (Figure 1). The association between diabetes and ALS varied with age at the index date (test for linear trend p<0.01, RERI=0.67); compared to the controls, ALS cases diagnosed before the age of 50 years were more likely and cases diagnosed after age 70 were less likely to have preexisting diabetes (Table 2). After stratifying by sex or calendar period of ALS diagnosis the associations did not vary greatly (Table 2).

Only three cases and five controls had diabetes before age 30. Nevertheless, we observed suggestive evidence for a positive association with early onset diabetes (OR=12.05, 95%CI=2.34–62.21). Analysis by diabetes duration revealed a non-linear trend (Figure 2A).

Among the study participants 302 had insulin-dependent diabetes and 889 had non-insulindependent diabetes. Diabetes patients with unknown insulin dependence (77 cases and 393 controls) were on average older and less educated than those whose status was known (data not shown). The association with ALS was stronger for non-insulin-dependent diabetes (OR=0.65, 95%CI=0.52–0.79) than for insulin-dependent diabetes (OR=0.81, 95%CI=0.58– 1.12). Results were similar after further adjustment for education and socioeconomic status (Table 3). Both insulin-dependent and non-insulin dependent diabetes were associated with decreased risk at older ages, but the former was associated with increased risk at younger ages (Table 3).

In total, 1,076 subjects had diabetes with complication. We observed similar inverse associations with ALS for diabetes without any complication (OR=0.79, 95%CI=0.62–1.01) and for diabetes with complications (OR=0.78, 95%CI=0.65–0.94).

#### Sensitivity analyses

Using the Prescribed Drug Register and restricting the analyses to participants with index dates during the period 2006–2010 (n=9,678), 10.38% of controls and 7.56% of cases had diabetes. The inverse association between diabetes and ALS persisted (OR=0.70, 95%CI=0.57–0.86); again the association was observed for non-insulin-dependent diabetes (OR=0.72, 95%CI=0.55–0.94) but not for insulin-dependent diabetes (OR=1.36, 95%CI=0.78–2.37).

Considering diabetes ascertained after 1964, we noted an increase in diabetes duration compared to duration ascertained after 1987 (median increase: 5 years). A U-shaped relationship was still seen for diabetes duration and ALS risk (Figure 2B). In total, 35 subjects had diabetes ascertained at age 30 years (13 cases and 22 controls); the OR for early onset diabetes was 3.25 (95%CI=1.61–6.53), and for late onset diabetes 0.74 (95%CI=0.63–0.85). Duration of diabetes was longer for early onset compared to late onset diabetes (t-test of the equality of means p<0.01). When considering only late onset diabetes, ALS risk decreased with increasing diabetes duration, with a plateau beyond six years after diabetes ascertainment (Figure 2C).

Finally, we confirmed our findings when considering only diabetes diagnosed 3 or more years before the index date (OR for overall diabetes 0.66, 95%CI=0.54–0.80).

#### DISCUSSION

We found an inverse association between premorbid diabetes and ALS that was mainly noted for non-insulin-dependent diabetes and ALS at age 70 and above. Conversely, for insulin-dependent diabetes, we found a positive association before the age of 50 years. The group of insulin-dependent diabetics in our study likely included both type 1 diabetes and type 2 diabetes treated with insulin. Since the life expectancy for type 1 diabetes is estimated as 70 years of age [14], we speculate that the inverse association between insulin-dependent

diabetes and ALS after age 70 may be largely due to insulin-dependent type 2 diabetes. Furthermore, diabetes with onset at age 30 or earlier, mainly insulin-dependent diabetes, appeared to be a risk factor for ALS, both in our main analysis and in our sensitivity analysis. Worth noting is that the latter provided a more meaningful estimate for the effect of early-onset diabetes because it relies on a better ascertainment of age at diabetes onset. Our results were consistent with an English study that found a positive association of ALS with early onset diabetes (<30 years) (relative risk 3.94) [9]. However, given the small numbers, our findings must be interpreted with care.

The biological mechanisms linking diabetes to ALS are as yet unclear. In an animal study, TDP-43, a protein encoded by the gene *TARDBP* which is mutated in ~4% of familial ALS cases, was shown to be a regulator of glucose and energy metabolism [15]. Specifically, overexpression of progranulin, an adipokine that mediates high-fat-induced insulin resistance, rescued mutant TDP-43 induced axonopathy in zebra fish ALS models [16,17]. Differences in associations of ALS with non-insulin-dependent type 2 and insulin-dependent type 1 diabetes may be related to differences in the pathophysiology of the two conditions. The former is characterized by insulin resistance, while the latter is an autoimmune disease destroying the beta-cells in the pancreas. Obesity, insulin resistance, or a reduction of insulin/insulin-like growth factor 1 (IGF1) signalling could all result in both type 2 diabetes and protection from neurodegeneration [18–20]. Furthermore, the two types of diabetes have both shared and independent genetic susceptibility [21].

Another potential explanation may be beneficial effects on ALS of treatments for noninsulin-dependent diabetes, such as metformin or pioglitazone. Metformin treatment in SOD1<sup>G93A</sup> transgenic mice produced small increases in motor neuron survival although it failed to significantly affect disease onset, progression or survival [22]. For pioglitazone, encouraging results have been reported in SOD1 mice, though it showed no beneficial effects on the survival of ALS patients as an add-on therapy to Riluzole [23].

The strength of the association between diabetes and ALS varied with time since first ascertainment of diabetes. Since the ascertainment date may have been years after diabetes onset and first diagnosis, the analysis by diabetes duration should be interpreted cautiously. The U-shaped association may have been a consequence of mixing insulin-dependent and non-insulin-dependent diabetes. In fact, when excluding subjects with diabetes onset at age 30 or earlier (mainly insulin-dependent diabetes), a consistently decreased ALS risk was noted. Such a pattern suggests a long induction period for diabetes to lower ALS risk, possibly because long-standing diabetes involves more profound pathophysiological changes; one possibility is greater dyslipidemia, a condition that is associated with longer survival in ALS patients [24]. This hypothesis may also explain why the decreased risk was observed only at age 70 and above.

The strengths of our study lie in the nationwide study design, the prospectively collected data, the complete follow-up, and the large sample size.

The main limitation of our study is the potential misclassification of diabetes status. The Inpatient Register data are known to have better positive predictive value than sensitivity for

diabetes [25]. Indeed, the prevalence of diabetes in our data increased substantially when antidiabetic medications were used as additional source for diabetes ascertainment. The cases of diabetes undetected by the Patient Register are likely to be milder forms; thus it would be speculative to generalize our study findings to all diabetes patients. Nevertheless, since the information was prospectively and independently collected using the Patient Register, the misclassification is expected to be non-differential.

Another limitation in our data is the lack of information on BMI. BMI may be associated with both diabetes, especially non-insulin-dependent diabetes, and ALS. Elevated BMI may also increase diabetes risk and partly explain the association between diabetes and ALS. In conclusion, our study provided evidence for an inverse association between diabetes and ALS. We emphasize the importance of taking into account age, insulin dependence and diabetes duration. Future studies should explore whether the association is independent of BMI.

#### Acknowledgments

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#### Figure 2.

Temporal relationship between diabetes duration and ALS. (A) Main analysis where diabetes was ascertained from 1987 onward. (B) Sensitivity analyses where diabetes was ascertained from 1964 onward. (C) Sensitivity analyses where diabetes was ascertained from 1964 onward and excluding the individuals with diabetes before age 30.

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		ALS cases			Controls	
	Men	Women	Total	Men	Women	Total
Z	2,872	2,236	5,108	14,360	11,180	25,540
Mean age at the index date (SD), years	67.37 (12.18)	69.36 (12.18)	68.24 (12.22)	67.36 (12.18)	69.37 (12.18)	68.24 (12.22)
Diabetes mellitus at the index date, $n\left(\%\right)$	144 (5.0)	80 (3.6)	224 (4.4)	899 (6.3)	538 (4.8)	1,437 (5.6)
Area of residence, $n \ (\%)$						
Northern Sweden	707 (24.6)	519 (23.2)	1,226 (24.0)	3,535 (24.6)	2,595 (23.2)	6,130 (24.0)
Central Sweden	1,513 (52.7)	1,201 (53.7)	2,714 (53.1)	7,565 (52.7)	6,005 (53.7)	13,570 (53.1)
Southern Sweden	652 (22.7)	516 (23.1)	1,168 (22.9)	3,260 (22.7)	2,580 (23.1)	5,840 (22.9)
Highest achieved education, $n \ (\%)$						
9 years	1,243 (44.3)	1,050 (48.4)	2,293 (46.1)	6,395 (45.8)	5,539 (51.3)	11,934 (48.2)
>9 and 12 years	1,037 (37.0)	773 (35.6)	1,810 (36.4)	5,033 (36.0)	3,526 (32.7)	8,559 (34.5)
University or doctoral studies	523 (18.7)	346 (16.0)	869 (17.5)	2,550 (18.2)	1,730 (16.0)	4,280 (17.3)
Missing	69	67	136	382	385	767
Socioeconomic status, n (%)						
Blue collar	992 (34.5)	741 (33.1)	1,733 (33.9)	5,178 (36.1)	3,721 (33.3)	8,899 (34.8)
White collar	1,115 (38.8)	760 (34.0)	1,875 (36.7)	5,385 (37.5)	3,584 (32.1)	8,969 (35.1)
Self-employed/Farmer	337 (11.7)	312 (14.0)	649 (12.7)	1,598 (11.1)	1,475 (13.2)	3,073 (12.0)
Other	428 (14.9)	423 (18.9)	851 (16.7)	2,199 (15.3)	2,400 (21.5)	4,599 (18.0)

#### Table 2

Adjusted Odds Ratios (ORs) and their corresponding 95% confidence intervals (CIs) for the association of ALS with diabetes mellitus (DM).

Variables	No. of cases with DM	No. of controls with DM	OR <sup>a</sup> (95%CI)	OR <sup>b</sup> (95%CI)
Overall	224	1,437	0.77 (0.66–0.89)	0.79 (0.68–0.91)
Age at the index	a date (years)			
<50	10	18	2.90 (1.31-6.44)	3.15 (1.40-7.08)
50–59	21	88	1.20 (0.74–1.94)	1.20 (0.74–1.95)
60–69	58	303	0.96 (0.72–1.27)	0.95 (0.71–1.27)
70–79	93	641	0.71 (0.57–0.89)	0.71 (0.57–0.89)
80 and above	42	387	0.52 (0.37-0.72)	0.56 (0.40-0.78)
Sex				
Male	144	899	0.79 (0.66–0.95)	0.80 (0.67–0.96)
Female	80	538	0.73 (0.57–0.93)	0.77 (0.60-0.98)
Calendar period	1			
1991-2000	56	380	0.73 (0.55–0.97)	0.77 (0.58–1.04)
2001-2010	168	1,057	0.78 (0.66–0.92)	0.79 (0.67–0.93)

 $^{a}$ OR adjusted for age, sex and area of residence (matching factors).

 $^{b}$ OR adjusted for age, sex, area of residence, education and socioeconomic status; 136 cases and 767 controls with missing information on education and socioeconomic status were excluded in this analysis.

#### Table 3

Adjusted Odds Ratios (ORs) and their corresponding 95% confidence intervals (CIs) for the association of ALS with insulin-dependent and non-insulin-dependent diabetes mellitus.

	No diabetes	Insulin-dependent diabetes	Non-insulin-dependent diabetes
ALS cases, n (%)	4,863 (97.1)	43 (0.9)	104 (2.1)
Controls, n (%)	24,059 (95.8)	259 (1.0)	785 (3.1)
<b>OR</b> <sup><i>a</i></sup> (95%CI)	1.00 (Ref)	0.83 (0.60–1.15)	0.66 (0.53–0.81)
<b>OR</b> <sup><i>a</i></sup> by Age at the index date (years)			
<50	1.00 (Ref)	5.38 (1.87–15.51)	2.12 (0.37–12.10)
50-59	1.00 (Ref)	1.16 (0.54–2.51)	0.77 (0.32–1.83)
60–69	1.00 (Ref)	1.19 (0.65–2.19)	0.74 (0.49–1.12)
70–79	1.00 (Ref)	0.45 (0.25-0.82)	0.67 (0.49–0.91)
80 and above	1.00 (Ref)	0.46 (0.14–1.51)	0.51 (0.33-0.80)

<sup>*a*</sup>OR adjusted for sex, age, area of residence, education and socioeconomic status; 136 cases and 767 controls with missing information on education and socioeconomic status were excluded in this analysis.