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Contralateral Prophylactic Mastectomy after Unilateral Breast Cancer: A Systematic Review & Meta-Analysis

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INTRODUCTION

Despite an overall trend towards less invasive oncologic care in the United States (US), rates of contralateral prophylactic mastectomy (CPM) in women diagnosed with unilateral breast cancer (UBC) have more than doubled over the past 15 years.¹ The increased prevalence of CPM is thought to reflect pervasive overestimation of metachronous contralateral breast cancer (MCBC) risk by breast-cancer patients,^{2–4} increased dissemination of personalized genetic and immunohistochemical information to patients,⁵ improved post-mastectomy reconstruction options,^{6–8} and exposure to internet-based information that is often contradictory. It is unclear whether CPM is associated with improved survival or decreased recurrence in UBC patients, all of whom are at increased risk for MCBC,^{9,10} i.e., contralateral breast cancer (CBC) diagnosed subsequent to an index cancer. Definitions of MCBC vary throughout the literature. Depending on a given researcher's decision as to what period of time is sufficiently long to distinguish a synchronous contralateral breast cancer (SCBC) from a metachronous one, MCBC has been defined as a new CBC diagnosed anywhere from one month to two years after an index tumor.¹¹ But the magnitude of MCBC risk is not uniformly distributed among patients with UBC: among women without a BRCA mutation, less than 10% would be expected to eventually develop MCBC,^{2,12} but among women with a family history of breast cancer and/or an identified genetic mutation in BRCA1 or BRCA2, incidence of MCBC has been estimated to be anywhere from 12% to 47%.^{13–15} CPM has historically been prescribed for these higher risk patients as a means through which to decrease MCBC and, concomitantly, mortality associated with MCBC. But even among this subset of breast-cancer patients, the efficacy of CPM in improving long-term clinical outcomes is questionable.

Mirroring the difficulties of establishing a uniform definition of MCBC, survival – overall, breast-cancer-specific, and disease-free – in women with UBC has been defined in variable ways throughout the literature, and reports of the potential survival benefit CPM might confer on recipients have been similarly inconsistent. Among recent studies examining the relationship between CPM and overall survival (OS), neither Chung and colleagues' 2012 study⁶ nor the 2000 study by Peralta et al.¹⁶ demonstrated a CPM-associated benefit with regards to OS. Peralta and colleagues did, however, report prolonged disease-free survival (DFS), defined as time to any breast-cancer event (namely, a recurrent or second primary breast cancer including newly diagnosed CBCs) among CPM recipients. In contrast, Bedrosian et al.'s 2010 study based on Surveillance, Epidemiology, and End Results (SEER) data, Boughey et al.'s 2010 study from the Mayo Clinic, and Herrinton et al.'s 2005 Cancer Research Network study all reported a OS advantage potentially conferred by CPM; however, there are important subtleties in their findings.^{17–19} In the SEER data study by Bedrosian and colleagues, the observed CPM-associated survival benefit demonstrated in the full analysis was found in subgroup analysis to stem largely from the strong survival benefit (4.8%) conferred on young (i.e., under the age of 50) CPM recipients with early-stage (I and II), estrogen-receptor (ER)-negative disease who – by virtue of having more years to live and more aggressive tumor biology at baseline – had a higher absolute lifetime risk of MCBC compared to their older and ER-positive counterparts.¹⁷ In their cohort, Boughey et al. found CPM to be associated with improved OS but not with breast-cancer-specific survival (BCSS) and this discrepancy could be ascribed to CPM recipients' being healthier at baseline, a conjecture supported by the fact that the 9% survival difference between recipients and non-recipients was greater than the absolute rate of CBCs in non-recipients (8.1%).¹⁸ Finally, in Herrinton et al.'s study, the 3.6% difference in breast-cancer-specific mortality (BCM) between CPM recipients and non-recipients (8.1% vs. 11.7%) is greater than the absolute reduction in CBC (0.5% vs. 2.7%), making it difficult to attribute the difference in disease-specific mortality to the effects of CPM and suggesting there may be some other contributing factor.¹⁹ Thus, it is unclear to what extent the observed survival benefit reported in these studies is secondary to decreased (though, notably, not eliminated) risk of MCBC following removal of contralateral breast tissue;⁹ selection bias, specifically confounding patient characteristics, such as younger age,^{9,17,20–22} that are both independently associated with better baseline health and a greater likelihood of undergoing CPM; or to receipt of treatments – such as tamoxifen and bilateral oophorectomy – that decrease the risk of BCM and/or all-cause mortality.^{23,24}

Here, we present the results of a systematic review and meta-analysis of CPM in female patients with a personal history of UBC. Although a Cochrane review on prophylactic mastectomy (both CPM in UBC patients as well as bilateral prophylactic mastectomy for prevention of a first breast cancer) was published in 2004 and updated in 2010,²⁵ our review is the first to include meta-analyses of clinical outcomes, focuses solely on CPM as a method of risk reduction in patients with breast-cancer diagnoses, and includes several large-scale studies published after the Cochrane review's 2010 update. Our intention is to provide a quantitative summation of current evidence that can serve as a succinct guide for clinical practice and can inform the development of future research examining the efficacy of CPM in both average- and high-risk breast-cancer patients.

METHODS

Data collection

The aims of this project were to examine whether CPM for women with UBC is associated with significant improvements in OS (primary outcome) as well as the following secondary outcomes: BCM, incidence of CBC, and rates of distant/metastatic recurrence (DMR).

A medical librarian developed search strategies (Appendix 1) for Medline/PubMed, EMBASE, Scopus, ClinicalTrials.gov, the Cochrane Database of Systematic Reviews, the Cochrane Central Register of Controlled Trials, and the Health Economic Evaluations Database using a combination of standardized index terms and plain language to cover the concepts of CPM, UBC, randomized controlled trials (RCTs), and observational studies as comprehensively as possible. Searches were limited to English-language studies published through March 2012 using standard limitations provided by the databases. We also contacted the authors of 4 studies and asked whether they would share their study data with us in a format that would be more amenable to inclusion in our meta-analyses; the authors for two of these studies agreed to do so.^{17,26}

Study selection

Our review was limited to published RCTs and observational studies that included and compared patients who had and had not received CPM. Case series and convenience samples were only included if they reported incidence of SCBC in CPM recipients. Conference abstracts were excluded. We defined CPM as any simple (total), subcutaneous, skin-sparing, modified radical, or radical mastectomy performed on a breast with no clinical or radiographic evidence of malignancy for the purpose of preventing CBC in a patient with a history of UBC. Included studies had to report survival, cause-specific mortality, CBC incidence, and/or recurrence for women who underwent CPM at any time after being diagnosed with primary UBC.

Data extraction

Data were independently abstracted and verified by two coders. Effect measures were deconstructed into component values (i.e., measures of incidence, length of follow-up) for both CPM and non-CPM cohorts. When this was not possible, study authors were contacted for primary data components of effect measures. In addition to outcome data, information on potentially confounding study-level characteristics (BRCA carrier status, family history) and cohort-level characteristics (age; length of follow-up; race; stage; lymph-node involvement; family history; receipt of chemotherapy, radiotherapy, and/or endocrine therapy; ER status; index tumor histology; multifocality/multicentricity; BRCA carrier status; and receipt of prophylactic oophorectomy) were extracted with the intention of conducting stratified subgroup analyses and meta-regression, respectively.

Statistical Analysis

Included studies were independently assessed by two coders using a modified version of Downs and Black's methodological quality checklist in which item 27 on sample-size calculation and power was converted into a yes/no question.²⁷ In addition, publication bias

was assessed by visual appraisal of a funnel plot, which plots the log of each study's effect size against its standard error. If asymmetry suggesting publication bias was apparent, a Harbord test was performed, with two-tailed $p < 0.05$ considered significant.

Relative risk (RR) was selected as the primary meta-analytic measure of association because not having individual patient follow-up time precluded calculation of hazard ratios. Risk difference (RD), i.e., absolute risk reduction, was also calculated. Statistical heterogeneity across trials was tested by Cochran's Q statistic. An alpha value of 0.5 was taken to indicate between-trial heterogeneity, which is represented by I^2 values. Fixed-effects models were used for meta-analyses in which there was no evidence of between-study heterogeneity, while random-effects models were used when there was significant heterogeneity. Meta-analyses were conducted to calculate pooled RRs and RDs for OS, BCM, incidence of MCBC, and incidence of DMR. Random-effects meta-analysis of proportions was used to calculate pooled incidence of SCBC.

Stratified subgroup and bivariate meta-regression analyses were conducted for all potential confounders for which there were at least two (for stratification) or three (for meta-regression) studies with sufficient data for analysis. We report pooled RRs and RDs for OS, BCM, MCBC, and DMR and pooled proportions for SCBC with 95% confidence intervals (CI) at two-tailed $p < 0.05$ significance. Statistical analyses were conducted in STATA 12 (Stata, College Station, Texas).

RESULTS

Study selection

The initial database search yielded 79 reports, reduced to 61 after de-duplication (Figure 1). Abstracts for all 61 reports were screened, and 8 studies were initially selected as potentially appropriate for meta-analysis inclusion.^{6,16,17,19,21,24,28,29} An additional 14 studies were selected for bibliographic review,^{1,8,25,30–40} for a total of 22 articles that were initially read in full in addition to having their bibliographies reviewed. Bibliographic review yielded 32 additional references, whose abstracts were also screened. Thus, a total of 93 abstracts (61 from the database search and 32 from bibliographic review) were screened. Of these, 53 were excluded, leaving 40 reports for full-text review. Twenty-three reports were excluded for not meeting eligibility criteria following full-text review, leaving 17 studies for qualitative synthesis. For three of these studies, we were unable to obtain primary data amenable to meta-analysis,^{6,22,41} leaving 14 studies for quantitative synthesis (Table 1).^{13,16–19,21,24,26,42–47} On the modified Downs and Black study quality and bias assessment checklist (maximum score 28), the median score for included studies was 18 (range 14–22). Inter-rater reliability (i.e., Kappa score) was substantial at 0.8. Only one study had an *a priori* sampling strategy¹⁹ designed to optimize statistical power, and none were RCTs. A funnel plot representing potential publication bias among the studies included in the OS meta-analysis was fairly symmetric (Figure 2), and the Harbord test indicated no significant publication bias ($p = 0.627$).

Results of meta-analyses (Table 2)

Overall survival (OS)—In random-effects meta-analysis (6 studies; CPM n = 10,666, no CPM n = 145,490; Figure 3), OS was 9% more likely for CPM recipients than for those who did not undergo CPM (RR = 1.09 [95% CI 1.06, 1.11, p<0.001]), and seven more UBC patients out of every 100 might be expected to survive if they received CPM (RD = 7.4% [95% CI 5.6%, 9.3%, p<0.001]).^{17–19,21,24,26} Of the six studies included, only one – Kiely et al.²¹ – had an individual effect size that was not statistically significant, and it was one of only two studies in this meta-analysis – the other being Van Sprundel et al.²⁴ – conducted outside the US.

Breast-cancer specific mortality (BCM)—In random-effects meta-analysis (4 studies; CPM n = 10,120, no CPM n = 142,105), CPM recipients had a BCM rate that was 31% lower than that of patients who did not undergo CPM (RR = 0.69 [95% CI 0.56, 0.85, p=0.001]).^{17–19,24} Out of 100 women with UBC, an additional 3 to 4 might not die of breast cancer if they received CPM (RD = –3.5% [95% CI –4.0, –3.0%, p<0.001]). Two (Bedrosian et al. and Herrinton et al.) of the four included studies had individual BCM rates that were significantly lower among CPM recipients, and both of these studies attempted to control for confounders – including receptor status, non-surgical treatments received, and family history – that might affect interpretation of post-CPM outcomes.^{17,19}

Synchronous contralateral breast cancer (SCBC)—The pooled incidence of SCBC among CPM recipients was 4.8% (9 studies; CPM = 3438; 95% CI 3.4, 6.2%)^{13,16,19,21,24,43–45,47} and is comparable to recent estimates of SCBC as reported by King et al. (5.9%)⁴⁵ and Chung et al. (6.6%).⁶

Metachronous contralateral breast cancer (MCBC)—In random-effects meta-analysis (8 studies; CPM n = 2325, no CPM n = 4840), CPM was associated with a 96% reduction in MCBC (RR = 0.04 [95% CI 0.02, 0.08, p<0.001]).^{16,18,19,21,23,24,26,42} Notably, CPM was not associated with an absolute reduction in the risk of MCBC incidence (RD = –18.0% [95% CI –42.0%, 5.9%, p=0.118]).

Distant/metastatic recurrence (DMR)—In fixed-effects meta-analysis (5 studies; CPM n = 953, no CPM n = 3323), DMR was found to be 36% less likely in CPM patients (RR = 0.64 [95% CI 0.51, 0.81, p<0.001]), and approximately 5 more women out of 100 might be expected to avoid DMR by undergoing CPM (RD = –4.9% [95% CI –7.2%, –2.6%, p<0.001]).^{16,18,24,26,42} Only two^{18,26} of the five studies in this meta-analysis had a statistically significant association between CPM and DMR, though all five studies had the same directionality of effect. King and colleagues found that the decrease in DMR they observed was attenuated and rendered insignificant when they adjusted for age and treatments received.²⁶

Stratified subgroup analysis: familial/genetic risk (FGR) – BRCA carrier status and family history—We conducted subgroup analyses in which we separated both the two studies – Metcalfe et al.²³ and Van Sprundel et al.²⁴ – in which all patients were BRCA mutation carriers as well as the two studies – Boughey et al. (2010)¹⁸ and Kiely et al.²¹ – in

which all patients had a family history positive for breast cancer and analyzed these four studies together to account for the higher familial/genetic risk (FGR) these patients have for CBC (synchronous and metachronous)^{13,48,49} and the greater benefit these patients might concomitantly derive from CPM. In the study by Boughey and colleagues, all study participants had a parent, sibling, or second-degree relative with breast cancer¹⁸ while all the participants in Kiely et al.'s study were members of families enrolled in the Kathleen Cunningham Foundation Consortium for Research into Familial Breast Cancer (kConFab), a collaborative research registry of Australian and New Zealander families with multiple cases of breast cancer.²¹

In stratified meta-analysis, CPM was not associated with improved OS when looking at the studies in which all patients had elevated FGR (RR = 1.09 [95% CI 0.97, 1.24, p=0.157]; RD = 6.6% [95% CI -1.2%, 14.3%, p=0.096]; 3 studies; CPM n = 618, no CPM n = 1318)^{18,21,24} but continued to be associated with significantly improved survival in the studies including patients with varying levels of FGR (RR = 1.10 [95% CI 1.09, 1.11, p<0.001]; RD = 8.4% [95% CI 7.8%, 8.9%, p<0.001]; 3 studies; CPM n = 10,048, no CPM n = 144,172).^{17,19,26} Likewise, BCM was not significantly decreased among studies in which all patients had elevated FGR (RR = 0.66 [95% CI 0.27, 1.64, p=0.283]; RD = -4.2% [95% CI -9.5%, 1.1%, p=0.123]; 2 studies; CPM n = 464, no CPM n = 454),^{18,24} but for the studies with patients of all FGR levels, CPM continued to be associated with decreased BCM (RR = 0.63 [95% CI 0.56, 0.70, p<0.001]; RD = -3.5% [95% CI -4.0%, -3.0%, p<0.001]; 2 studies; CPM n = 9656, no CPM n = 141,651).^{17,19}

In contrast to the full analysis, both the relative and absolute risks of MCBC incidence (Figures 4 and 5) were significantly decreased among patients with elevated FGR (RR = 0.04 [95% CI 0.02, 0.09, p<0.001]; RD = -24.0% [95% CI -35.6%, -12.4%, p=0.013]; 4 studies; CPM n=764, no CPM n=1654),^{18,21,23,24} but only the RR of MCBC was significantly decreased among patients of all risk levels (RR = 0.08 [95% CI 0.01, 0.46, p=0.005]; RD = -11.1% [95% CI -5.9%, 37%, p=0.240]; 4 studies; CPM n=1561, no CPM n=3186).^{16,19,26,42} The effect size and directionality of the pooled measures for SCBC and DMR did not appreciably differ in this subgroup analysis from those of the full meta-analysis.

Meta-regression

Bivariate meta-regression analyses demonstrated no significant covariate effect on outcomes (Appendix 2).

DISCUSSION

Summary and contextualization of main results

In our full meta-analysis, CPM was associated with significant relative and absolute increases in the OS of recipients when compared to non-recipients. CPM was also associated with lower relative and absolute risks of BCM and DMR. But while CPM was associated with a decreased relative risk of developing MCBC, it was not associated with an absolute reduction in MCBC risk. In stratified meta-analysis with subgroup divisions based on

whether or not participants' had elevated FGR, results differed from those of the full analysis. Among patients with varying levels of FGR, CPM continued to be associated with improved OS, BCM, and DMR, and was associated with a relative but not an absolute reduction in MCBC risk. Among patients with known elevated FGR, however, there was no relative or absolute improvement in OS or BCM, but there were significant reductions in the relative and absolute risks of both MCBC and DMR among CPM recipients relative to non-recipients.

Although our primary outcome was OS, we focus our discussion on MCBC because we suspect it is less likely to be confounded than any of the other comparative measures in our analysis. There are many more known and unknown covariates – including but not limited to aspects of a patient's health that are unrelated to her having breast cancer – associated with a breast-cancer patient's likelihood of survival than there are with a patient's likelihood of developing another breast cancer. Furthermore, a number of the demographic characteristics known to confer a pro-survival selection bias on patients undergoing CPM (e.g., noninvasive histology)²⁶ may not or only to a lesser extent impact the odds of a patient's going on to experience MCBC. Finally, through stratification, we controlled for family history and BRCA carrier status, two covariates that are well-documented risk factors for both index and metachronous breast cancers,⁵⁰ and this subgroup analysis enabled a more insightful understanding of the impact of CPM on MCBC.

In our full meta-analysis, there was no significant decrease in the absolute risk of MCBC incidence when comparing CPM recipients and non-recipients with varying levels of FGR for MCBC. This finding is critically important: it confirms that the risk of MCBC in UBC patients, regardless of whether they undergo CPM or not, is very low at baseline, and it strongly suggests that the decreased rates of mortality observed when comparing CPM recipients to non-recipients in the general population are not attributable to a treatment-derived decrease in MCBC incidence, but rather to *other* covariates unrelated to the decrease in MCBC risk that CPM would directly provide. This finding is further supported by the results of the subgroup analysis in which we stratified studies according to whether or not they exclusively included patients with known elevated FGR. It is well-established that UBC patients with carcinogenic genetic mutations and/or family histories of breast cancer are significantly more likely to develop MCBC,^{13,14} and it was only among these high-risk patients that we saw a significant decrease in both the relative and absolute risks of MCBC following CPM, thus implying that, based on our analysis, CPM should be limited to this subset of UBC patients, if performed at all.

DMR was still found to be lower among CPM recipients in the subgroup with elevated FGR, implying that the CPM-associated decrease in MCBC may significantly prevent metastasis from these subsequent cancers but does not improve OS. This apparent discrepancy may reflect selection bias, with CPM recipients' having less aggressive tumor biology at baseline. But the evidence regarding CPM receipt and index tumor biology in this cohort of women is contradictory. There is evidence that larger tumor size is associated with choosing to undergo CPM,^{6,9,51} perhaps because larger tumors predispose patients to choose mastectomy over breast conservation therapy (BCT) and women are more likely to undergo CPM if they are already getting a therapeutic mastectomy.⁶ Furthermore, there is some

evidence that women with elevated FGR are actually more likely to have biologically aggressive primary breast cancers that are more likely to recur, metastasize, and result in death and that they are also at increased risk for other extramammary malignancies.⁴⁸ But other studies have reported an association between small tumor size and receipt of CPM.²⁶ Notably, women with elevated FGR are more likely to undergo radiographic surveillance – including magnetic resonance imaging (MRI) of the breast – that leads to diagnosis of index tumors at earlier stages.⁵² Accordingly, these women may actually have *less* advanced disease at diagnosis than CPM recipients without known elevated FGR,⁵² their index tumors may be less likely to spread, and these women would appear to have lower DMR without a significant difference in OS relative to non-recipients. However, we suspect that the observed discrepancy between CPM's association with DMR and OS among patients with elevated FGR is due to the increasingly non-lethal nature of breast cancer, even after regional or (to a lesser extent) distant disease extension. Over time, stage for stage, fewer women are dying of breast cancer as a result of improved treatments.⁵³ Hence, diagnosis with advanced disease is less likely than it once was to be associated with death such that we might see CPM have an advantageous effect on MCBC and concomitantly on DFS but not on OS.

Quality of evidence

A meta-analysis is only as good as its constitutive studies, which ideally would be RCTs. None of the studies in this meta-analysis were RCTs because none have been conducted on CPM. In our assessment of overall study quality, the 14 studies included in our meta-analysis were of moderate methodological rigor for the purposes of quantitative synthesis. A number of the studies in our meta-analysis made significant efforts to reduce selection bias due to confounders, but we recognize that selection bias is the greatest hindrance to interpreting our results, and we anticipated this issue by using meta-regression and stratification. Bivariate meta-regression failed to demonstrate any significant confounding, though we recognize that these analyses may have been underpowered due to each one's having a small number of observations. However, our stratified subgroup analysis was very illuminating, as it revealed the extent to which CPM-associated differences in outcome could or could not be ascribed to changes in MCBC incidence.

We would have wished to be able to analyze other patient-specific and disease-specific characteristics that might place patients at significantly lower or increased risk for survival and/or CBC incidence, but by virtue of the data we were able to extract, our stratified analysis was limited to patients with either known BRCA carrier status or a documented family history of breast cancer, who we grouped together as having elevated FGR for both primary and secondary breast cancer incidence. Clinical prediction models such as the Gail⁵⁴ and Tyrer-Cuzick (i.e., International Breast Cancer Intervention Study [IBIS])⁵⁵ models provide fuller assessments of breast-cancer risk than simply the binary characteristic of having a family history of breast cancer or not, but these models were not consistently utilized in the studies we ultimately included in this review.

Some might argue that individual patient data (IPD) would be required to enable confounder adjustment through multivariate regression, propensity scores, or instrumental variables.

However, we suspect that the benefits of a pooled analysis using IPD from currently available CPM studies would be limited for a number of reasons. First of all, it is not always the case that IPD analysis leads to different conclusions from study-level meta-analysis, especially for models looking at overall estimates of exposure/treatment (in this case, CPM) effects.⁵⁶ Thus, it is entirely possible that the benefits of IPD – adjustment for confounders, better time-to-event analysis – might not outweigh the significant costs in personnel and time. Furthermore, our meta-regression analyses suggest that using IPD from existing studies might not solve the problem of confounding because many of the covariates (e.g., socioeconomic status³⁹ and surgeon gender⁹) impacting who among breast-cancer patients receives and benefits from CPM will only rarely and inconsistently have been recorded. Finally, several of the included studies used data from premiere medical centers dedicated to the treatment of cancer, though in reality most patients in the US and throughout the world receive their oncologic surgery through generalists. Thus, the benefits observed in our meta-analysis must be considered with the understanding that realization of said benefits depends not only on who receives CPM but also by whom and in what medical context it is performed.

Implications for practice

In short, CPM decreases MCBC incidence in breast-cancer patients with BRCA mutations and/or family histories of breast cancer, both of which place women at increased risk for MCBC, but CPM does not appear to confer a survival benefit even within this subset of patients. Among breast-cancer patients not otherwise at high risk for MCBC, the improvement observed in OS and BCM is likely secondary to selection bias, as CPM recipients may be more likely to have other characteristics – adequate health insurance,^{39,57} early-stage tumors²⁶ – that strongly correlate with improved survival; furthermore, the risk of MCBC in most UBC patients is already low and likely does not warrant the morbidity of CPM. We recommend that UBC patients without known elevated FGR be advised against CPM, while patients with elevated FGR should be advised that while CPM would significantly decrease their risk of MCBC, it is unlikely to prolong their lives.

We do not claim that UBC patients with elevated FGR constitute the only type of high-risk group that might benefit from CPM, but it is possible that the conclusions reached regarding this group of women could also be extended after further study to women who, for other reasons, are at high risk for MCBC. Accordingly, a critical next step in evidence-based clinical practice is refinement of our criteria as to who is at high risk for MCBC. Some of the studies in our meta-analysis were limited to early-stage patients^{18,23,44} and amongst all of the studies, the racial and/or ethnic diversity of included studies' participants – if reported at all – was very low. Given that CBC is more common in women with late-stage tumors^{58,59} and in women of African ancestry,⁶⁰ the benefits of CPM in these populations could be significant but are, as of yet, underexplored. CPM is also more commonly chosen by women with health insurance,^{39,57} so it will be important to assess the extent to which implementation of the Affordable Care Act in the US impacts the demographics of who gets CPM and the extent to which they realize its potential benefits, particularly if access to CPM alternatives (e.g., MRI surveillance) is significantly affected by insurance type.

Improvements in breast-cancer treatment and diagnosis should also mitigate use of CPM. For example, UBC patients treated with tamoxifen, trastuzumab, or other receptor-targeted therapy for their index breast cancers not only decrease their chances of BCM but also simultaneously treat any occult malignancy or atypia that might exist in contralateral breast tissue.^{61–65} Furthermore, index tumors, SCBCs, and MCBCs are all increasingly likely to be diagnosed at earlier stages as a result of higher rates of screening participation,⁶⁶ improvements in mammography, and increased utilization of breast MRI,^{67–73} and these early-stage tumors are less likely to be associated with recurrence and death.^{74–77} Currently, the Society of Surgical Oncology (US) only recommends CPM be considered for breast-cancer patients in whom contralateral surveillance would be difficult, post-reconstructive breast symmetry would be improved, and risk reduction would be significant secondary to strong family history, known predisposing genetic mutations, and/or biopsy-proven high-risk pathology (atypical ductal hyperplasia, atypical lobular hyperplasia, and lobular carcinoma in situ).⁴ We concur that CPM should be limited to patients at high risk for MCBC, but we suspect our current assessment of risk will need to be broadened to incorporate other clinical characteristics – including tumor grade and molecular subtype – at the same time as we collectively refine physicians’ counseling of patients.

Physician recommendations have been repeatedly demonstrated to have a significant impact on patients’ decision to undergo CPM.^{33,47} As a profession, we are guilty of providing breast-cancer patients with ever increasing amounts of information but with insufficient contextualization regarding their options for treatment and postsurgical cosmesis. CPM is not costless. While many women report satisfaction with both CPM and their post-mastectomy reconstructions, others have reported post-CPM issues with self-esteem, body image, and mental and sexual health.^{41,47,78–82} Physicians in general, and surgeons in particular, must work to educate patients as to their individual risk of CBC and to better inform them of the costs and benefits of CPM as well as alternatives to CPM such as MRI surveillance and pharmaceutical prophylaxis.

Implications for research

We have hypothesized that the improvement in OS and other outcomes seen in our full meta-analysis is secondary to confounding, but we recognize that there may be other unexplored reasons. For example, while young women with breast cancer might have better overall health at baseline and are more likely to choose CPM than not, young age at diagnosis is actually a negative prognosticator with regards to breast cancer, with younger women having worse survival and clinicopathologic features compared with older women.⁸³ However, young CPM recipients are also more likely to have family histories of breast cancer⁹ and to have had more screening and diagnostic imaging prior to diagnosis,⁶ so they might actually have lower stage disease than other young breast-cancer patients. Accordingly, one cannot be sure to what extent all of these epidemiologic factors might interact in predicting the benefit or harm of CPM in this population. Given the low probability of a future RCT on CPM, prospective, population-based studies akin to that conducted by Kiely and colleagues²¹ would be helpful to interrogate post-CPM outcomes and the interaction between baseline FGR, age at diagnosis, the levels of radiographic surveillance and prophylactic therapy (e.g., tamoxifen, oophorectomy) patients with elevated

FGR receive, and the extent to which these levels vary regionally and internationally. We encourage the collection of CPM data in prospectively maintained databases such as kConFab's²¹ and hope that such databases will be increasingly inclusive of women from all racial, ethnic, and socioeconomic backgrounds.

In conclusion, CPM may hold benefits for UBC patients with elevated FGR for MCBC, but based on the findings of our meta-analysis, its use is not justified in breast-cancer patients not otherwise at elevated risk for developing MCBC. And even among the subset of high-risk patients with elevated FGR, CPM is associated with decreased MCBC incidence but not with improved survival. The temporal and financial challenges of conducting an RCT on CPM will likely preclude one from ever being conducted; post-enrollment periods of at least 10 or 20 years would be required to allow for substantive accumulation of breast-cancer events, and it would be difficult, if not impossible, to achieve that degree of longitudinal observation in retrospective analyses from single or even multiple institutions.⁵⁸ Accordingly, the contribution of our study, the only quantitative summation of the literature on CPM, is significant: we have demonstrated that, given the minimal decrease in MCBC risk conferred by CPM in the general population of UBC patients, CPM should not be offered to those whose FGR does not otherwise place them at high risk for MCBC.

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Appendix 1 (Online only). PubMed database search strategy for systematic review and meta-analysis of contralateral prophylactic mastectomy

(“Women”[Mesh] OR “Female”[Mesh] OR “women” OR “woman” OR “female” OR “females”) AND (“Breast Neoplasms”[Mesh] OR “Hereditary Breast and Ovarian Cancer Syndrome”[Mesh] OR “breast carcinomas” OR “breast carcinoma” OR “Breast Neoplasms” OR “Breast Neoplasm” OR “Breast Tumors” OR “Breast Tumor” OR “Mammary Carcinomas” OR “Mammary Carcinoma” OR “Mammary Neoplasm” OR “Mammary Neoplasms” OR “Breast Cancer” OR “breast cancers” OR “Cancer of the Breast” OR

"Cancer of Breast" OR "Mammary Ductal Carcinomas" OR "Mammary Ductal Carcinoma"
 OR "Breast Invasive Ductal Carcinoma" OR "Hereditary Breast and Ovarian Cancer
 Syndrome" OR "HBOC Syndrome" OR "HBOC Syndromes" OR "BRCA1" OR "BRCA2"
 OR "breast gland cancer" OR "breast gland neoplasm" OR "mamma cancer" OR
 "mammary cancer" OR "mammary gland cancer" OR "breast adenocarcinoma" OR
 "mammary adenocarcinoma" OR "breast carcinogenesis" OR "breast cancerogenesis" OR
 "mammary gland carcinogenesis" OR "breast carcinoma" OR "mamma carcinoma" OR
 "breast metastasis" OR "mammary gland metastasis" OR "breast tissue metastasis" OR
 "breast sarcoma" OR "mammary gland sarcoma" OR "mammary sarcoma" OR
 "cystosarcoma phylloides" OR "breast phyllod tumor" OR "cysto sarcoma phylloides" OR
 "cystosarcoma" OR "giant fibroadenoma" OR "phyllodes tumor" OR "phyllodes tumor")
 AND ("prevention and control" [Subheading] OR "prevention and control" OR "preventive
 therapy" OR "prophylaxis" OR "preventive measures" OR "prevention" OR "control" OR
 "asepsis" OR "disease eradication" OR "protection" OR "prophylactic" OR
 "Prophylactically") AND ("Mastectomy"[Mesh] OR "Mastectomy" OR "Mastectomies" OR
 "Mammectomy" OR "Mammectomies" OR "breast amputation" OR "breast resection" OR
 "Halsted operation") AND ("unilateral" AND "contralateral") AND ("Mortality"[Mesh] OR
 "mortality" [Subheading] OR "Mortality" OR "Mortalities" OR "Case Fatality Rate" OR
 "Case Fatality Rates" OR "Death Rate" OR "Death Rates" OR "survival" OR "disease free"
 OR "cancer free" OR "asepsis" OR "Recurrence"[Mesh] OR "Recurrence" OR
 "Recurrences" OR "Relapse" OR "Relapses" OR "Recrudescence" OR "Recrudescences"
 OR "cancer recidive" OR "cancer regeneration") NOT (("Animals"[Mesh] NOT ("Animals"
 [Mesh] AND "Humans"[Mesh]))

Limits: English

Appendix 2 (Online only). Results of bivariate meta-regression analyses

Meta-regression analyses for the outcome OS were conducted using covariates representing
 age (5 studies, $p=0.491$), length of follow-up (5 studies, $p=0.107$), receipt of chemotherapy
 (3 studies, $p=0.776$), and lymph node involvement (4 studies, $p=0.410$). For the outcome
 BCM, there was only sufficient data to conduct meta-regression analyses using age (3
 studies, $p=0.639$) and length of follow-up (3 studies, $p=0.380$). Meta-regression analyses
 were conducted for the outcome MCBC using the following covariates: age (6 studies,
 $p=0.651$); length of follow-up (7 studies; $p=0.404$); stage (4 studies; $p=0.238$); receipt of
 systemic (5 studies, $p=0.856$), endocrine (4 studies, $p=0.732$), and radiation therapy (3
 studies, $p=0.684$); family history (5 studies, $p=0.504$); and lymph-node status (3 studies,
 $p=0.597$). For the outcome DMR, meta-regression was conducted for six covariates: age (4
 studies, $p=0.683$); length of follow-up (4 studies, $p=0.544$); stage (3 studies, $p=0.683$);
 receipt of endocrine (3 studies, $p=0.664$) and systemic therapy (3 studies, $p=0.752$); and
 family history (3 studies, $p=0.576$). There were insufficient observations to enable meta-
 regression analysis of any outcomes examining race, ER status, index tumor histology,
 multifocality/multicentricity, BRCA carrier status, or receipt of prophylactic oophorectomy
 as covariates.

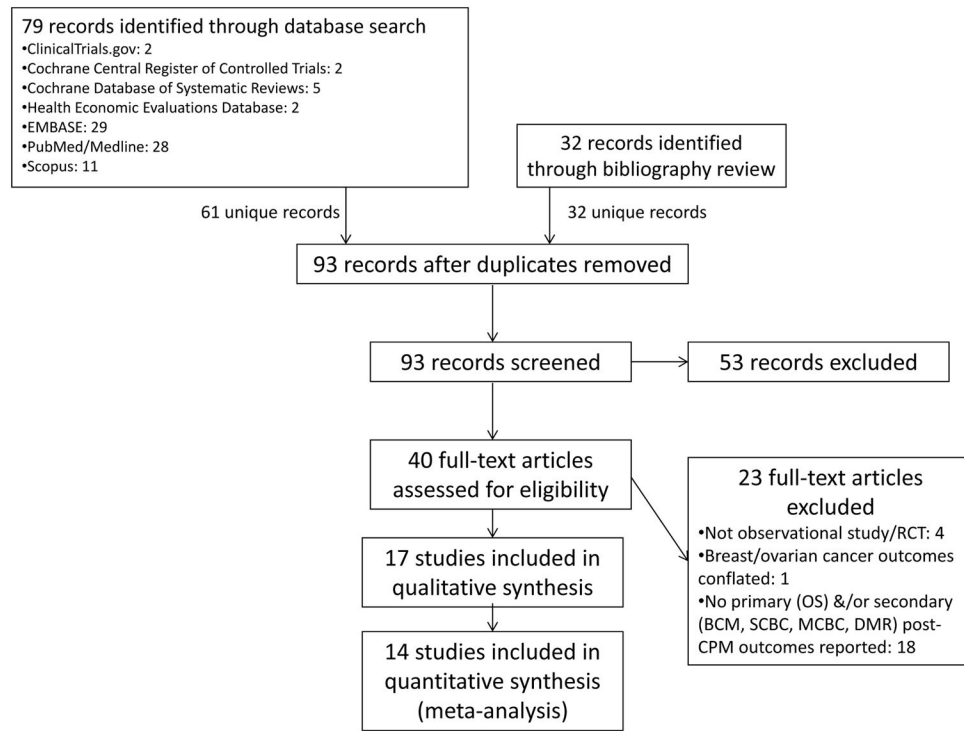


Figure 1.
Flow diagram of articles screened and selected for meta-analysis

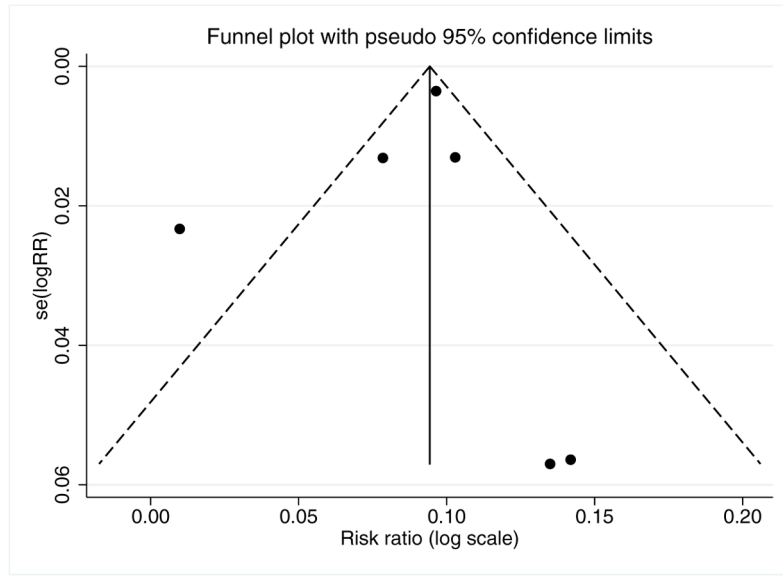


Figure 2. Funnel plot assessing publication bias in overall survival meta-analysis

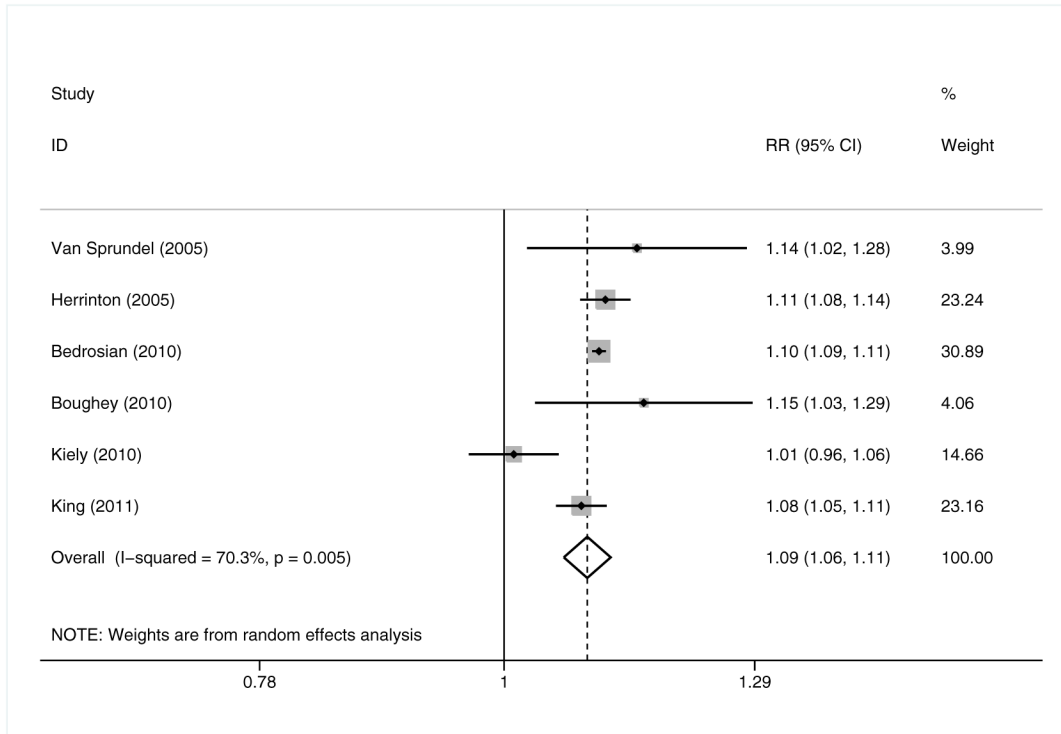


Figure 3.
Overall survival – Relative Risk

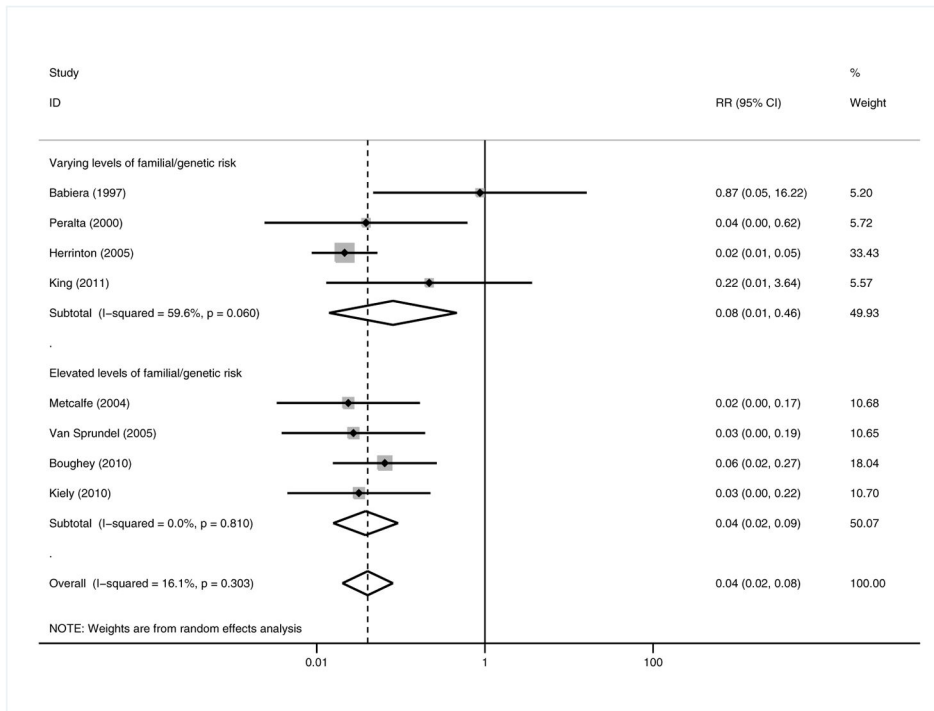


Figure 4. Metachronous contralateral breast cancer: elevated familial/genetic risk subgroup analysis – Relative Risk

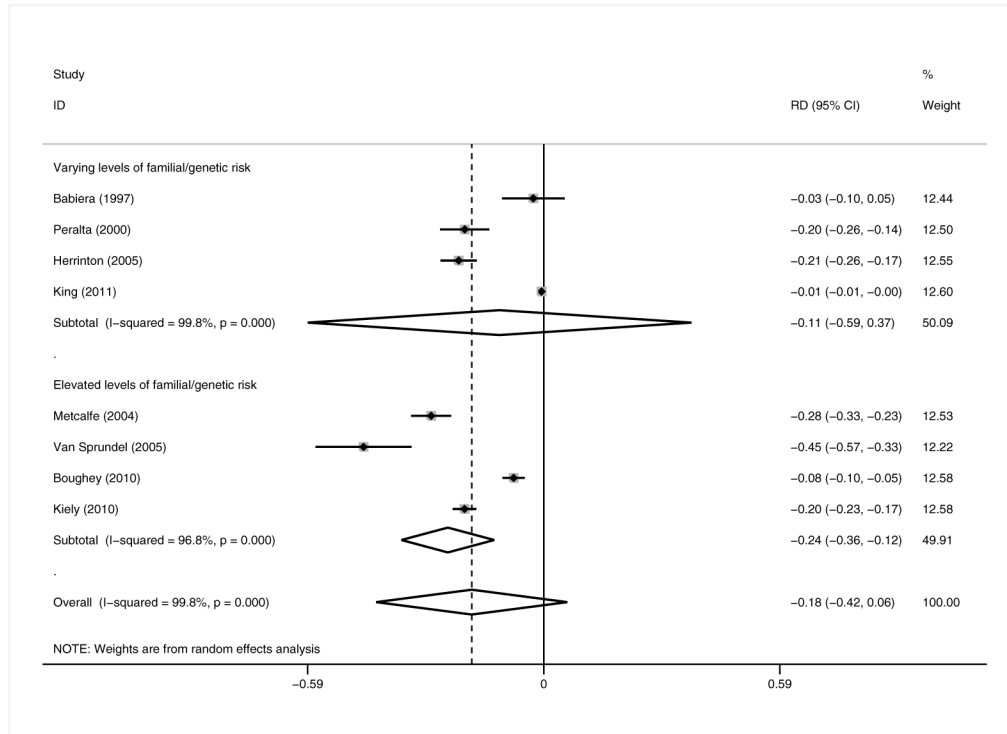


Figure 5. Metachronous contralateral breast cancer: elevated familial/genetic risk subgroup analysis – Risk Difference

Table 1

Studies included in Contralateral Prophylactic Mastectomy (CPM) Meta-Analysis

Study	Year	Country	Study Design	Data Source	CPM N= (F/U)	No CPM N= (F/U)	Year Range	Meta-Analyses
Babiera et al. ⁴²	1997	United States	Retrospective cohort	M.D. Anderson Cancer Center, Houston, TX	18 (52 mos) ^a	115 (70 mos) ^a	1978–1993 ^d	MCBC, DMR
Bedrosian et al. ¹⁷	2010	United States	Retrospective cohort	SEER	8748 (47 mos) ^a	95,283 (47 mos) ^a	1998–2003 ^d	OS, BCM
Boughey et al. ⁴³	2006	United States	Retrospective cohort	M.D. Anderson Cancer Center, Houston, TX	382 (NR)	-	2001–2005 ^e	SCBC
Boughey et al. ¹⁸	2010	United States	Case-control	Mayo Clinic, Rochester, MN	385 (18 yrs) ^a	385 (16.4 yrs) ^a	1971–1993 ^e	OS, BCM, MCBC, DMR
Goldflam et al. ⁴⁴	2004	United States	Case series	M.D. Anderson Cancer Center, Houston, TX	239 (7.8 yrs) ^b	-	1987–1997 ^e	SCBC
Herrinton et al. ¹⁹	2005	United States	Retrospective cohort	Cancer Research Network	908 ^f ; 1072 ^g (5.7 yrs) ^a	46,368 ^f ; 317 ^g (4.8 yrs) ^a	1979–1999 ^d	OS, BCM, SCBC, MCBC
Kiely et al. ²¹	2010	Australia/New Zealand	Prospective cohort	kConFab	154 (8 yrs) ^b	864 (11.7 yrs) ^b	Up to 2008 ^e	OS, SCBC, MCBC
King et al. – JCO ²⁶	2011	United States	Retrospective cohort	Memorial Sloan-Kettering Cancer Center, New York, NY	407 (4.4 yrs) ^{a,c}	2572 (6.8 yrs) ^{a,c}	1997–2005 ^e	OS, MCBC, DMR
King et al. – A Surg ⁴⁵	2011	United States	Case series		407 (4.4 yrs) ^a	-		SCBC
McDonnell et al. ¹³	2001	United States	Case series	Mayo Clinic, Rochester, MN	745 (10 yrs) ^a	-	1960–1993 ^d	SCBC
Metcalfe et al. ²³	2004	Canada/United States	Retrospective cohort	The Hereditary Breast Cancer Clinical Study Group	146 (9.2 yrs) ^b	336 (9.2 yrs) ^b	1975–2000 ^d	MCBC
Montgomery et al. ⁴⁷	1999	United States	Convenience sample	National Prophylactic Mastectomy Registry	296 (4.9 yrs) ^a	-	1954–1998 ^e	SCBC
Peralia et al. ¹⁶	2000	United States	Retrospective cohort	City of Hope National Medical Center, Duarte, CA	64 (6.8 yrs) ^b	182 (6.8 yrs) ^b	1973–1998 ^e	SCBC, MCBC, DMR
Van Sprundel et al. ²⁴	2005	The Netherlands	Retrospective cohort	Leiden University Medical Center; The Netherlands Cancer Institute/Antoni van Leeuwenhoek Hospital, Amsterdam	79 (7.4 yrs) ^b	69 (10.5 yrs) ^b	Up to 1993 ^d	OS, BCM, SCBC, MCBC, DMR

^a Median;^b Mean;

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^c For overall survival meta-analysis, minimum follow-up time data = 2 years, median=6.1 years; CPM n=392, no CPM n=2521;

^d dates of diagnosis;

^e dates of procedure;

^f OS, BCM;

^g SCBC, MCBC

A Surg, *Annals of Surgery*; BCM, breast-cancer-specific mortality; CA, California; DMR, distant/metastatic recurrence; F/U, mean/median follow-up time; ILR, ipsilateral locoregional recurrence; JCO, *Journal of Clinical Oncology*; kConFab, Kathleen Cunningham Foundation consortium for Research into Familial Breast Cancer; MCBC, metachronous contralateral prophylactic mastectomy; MD, Maryland; MN, Minnesota; mos, months; NR, not reported; NY, New York; OS, overall survival; SCBC, synchronous contralateral breast cancer; SEER, Surveillance, Epidemiology, and End Results database; TX, Texas; yrs, years

Table 2

Summary of Meta-Analysis Results

Analysis	Outcome	Measure	Point Estimate	95% CI	p value	No. of Studies	CPM (n=)	No CPM (n=)
All studies ^{17-19,21,24,26}		RR	1.09	1.06, 1.11	<0.001	6	10,666	145,490
		RD	7.4%	5.6%, 9.3%	<0.001			
FGR – Elevated ^{18,21,24}	OS	RR	1.09	0.97, 1.24	0.157	3	618	1318
		RD	6.6%	-1.2%, 14.3%	0.096			
FGR – Varying ^{17,19,26}		RR	1.10	1.09, 1.11	<0.001	3	10,048	144,172
		RD	8.4%	7.8%, 8.9%	<0.001			
All studies ^{17-19,24}		RR	0.69	0.56, 0.85	0.001	4	10,120	142,105
		RD	-3.5%	-4.0%, -3.0%	<0.001			
FGR – Elevated ^{18,24}	BCM	RR	0.66	0.27, 1.64	0.283	2	464	454
		RD	-4.2%	-9.5%, 1.1%	0.123			
FGR – Varying ^{17,19}		RR	0.63	0.56, 0.70	<0.001	2	9656	141,651
		RD	-3.5%	-4.0%, -3.0%	<0.001			
All studies ^{13,16,19,21,24,26,43,44,47}		RR	4.8%	3.4%, 6.2%	-	9	3438	-
		Rate	5.7%	1.8%, 9.6%	-			
FGR – Varying ^{13,16,19,43-45,47}		RR	4.8%	3.2%, 6.3%	-	7	3205	-
		Rate	4.8%	3.2%, 6.3%	-			
All studies ^{16,18,19,21,23,24,26,42}		RR	0.04	0.02, 0.08	<0.001	8	2325	4840
		RD	-18.0%	-42.0%, 5.9%	0.118			
FGR – Elevated ^{18,21,23,24}	MCBC	RR	0.04	0.02, 0.09	<0.001	4	764	1654
		RD	-24.0%	-35.6%, -12.4%	0.013			
FGR – Varying ^{16,19,26,42}		RR	0.08	0.01, 0.46	0.005	4	1561	3186
		RD	-11.1%	-5.9%, 37%	0.240			
All studies ^{16,18,24,26,42}	DMR	RR	0.64	0.51, 0.81	<0.001	5	953	3323
		RD	-4.9%	-7.2%, -2.6%	<0.001			

Analysis	Outcome	Measure	Point Estimate	95% CI	p value	No. of Studies	CPM (n=)	No CPM (n=)
FGR – Elevated ^{18,24}		RR	0.71	0.53, 0.94	0.018	2	464	454
		RD	-5.9%	-10.7%, -1.0%	0.017			
FGR – Varying ^{16,26,42}		RR	0.58	0.40, 0.83	0.003	3	489	2869
		RD	-4.4%	-6.7%, -2.0%	<0.001			

BCM, breast-cancer-specific mortality; CI, confidence interval; CPM, contralateral prophylactic mastectomy; DMR, distant/metastatic recurrence; FGR, familial/genetic risk; OS, overall survival; MCBC, metachronous contralateral breast cancer; SCBC, synchronous contralateral breast cancer; RD, risk difference; RR, relative risk