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Which alcohol use disorder criteria contribute to the association of *ADH1B* with alcohol dependence?

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Abstract

Although alcohol dependence (AD) is approximately 50% heritable, little is known about how specific genetic loci affect AD risk. In a genome-wide association study (GWAS), we identified highly significant associations between two population-specific functional variants in the alcohol dehydrogenase 1B gene (*ADH1B*) and AD in African-Americans (AAs; rs2066702) and European-Americans (EAs; rs1229984). In the current study, we determined which specific diagnostic criteria contributed to the observed associations of *ADH1B* SNPs with AD. Our analysis included both the DSM-IV and DSM-5 diagnostic systems. We also investigated the relationship of *ADH1B* variants to the maximum number of drinks consumed in a 24-hour period (MaxDrinks), a presumed intermediate phenotype of AD. We found that, although all criteria made strong individual contributions to the associations, the largest contributions came from those reflecting neuroadaptation: tolerance (rs2066702) and withdrawal (rs1229984). Overall, evidence for association with DSM-5 criteria was slightly stronger than for DSM-IV criteria. For rs2066702, results were similar for DSM-IV and DSM-5 criteria. However, the most significant DSM-5 criterion associated with rs1229984 was alcohol-related social/interpersonal problems. Both *ADH1B* variants were associated with MaxDrinks, a measure of innate tolerance, and MaxDrinks mediated the associations between *ADH1B* and alcohol outcomes. We replicated the findings for rs2066702 and tolerance in an independent sample of AAs. Taken together, these results suggest that variation in *ADH1B* affects the adaptation to heavy drinking, highlighting

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population-specific differences in genetic risk for AUD. They also suggest that the revisions reflected in DSM-5 AUD may enhance the utility of that diagnosis for gene finding.

Keywords

Alcohol dehydrogenase 1B; DSM; tolerance

Introduction

Alcohol dependence (AD) is a complex psychiatric disorder with an estimated heritability of about 50% (Goldman *et al*, 2005). Understanding the genetic influences on AD could inform diagnosis, treatment, and prevention efforts. However, relatively little is known regarding the actions of specific genetic loci on AD risk. A major focus of the AD candidate gene literature has been on genes that encode alcohol-metabolizing enzymes, such as alcohol dehydrogenase 1B (*ADH1B*), which catalyzes the oxidation of alcohol to acetaldehyde. Several *ADH1B* variants alter the enzyme's kinetics, including the missense single nucleotide polymorphisms (SNPs) rs2066702 (C→T; Arg370Cys) and rs1229984 (G→A; Arg48His). Both of these substitutions result in enhanced *ADH1B* enzymatic activity that, following alcohol consumption, increases the production of acetaldehyde, producing such aversive effects as facial flushing, tachycardia, and nausea (Crabb *et al*, 2004; Edenberg, 2007). The rs2066702*T and rs1229984*A alleles are relatively uncommon in European-American (EA) and African-American (AA) populations. The rs1229984*A allele is prevalent in Asian populations, where it was first shown to decrease risk for AD (Chen *et al*, 1999; Thomasson *et al*, 1991; Li *et al*, 2011; Thomasson *et al*, 1991). Subsequent studies demonstrated that these alleles are also protective from AD in AA and EA populations (Bierut *et al*, 2012; Li *et al*, 2011; Luo *et al*, 2006; Meyers *et al*, 2013).

We recently published a genome-wide association study (GWAS) of AD in groups of AA and EA subjects (Gelernter *et al*, 2014). Two approaches were employed: a case-control model that used a dichotomous AD diagnosis from the fourth edition of the Diagnostic and Statistical Manual of Mental Disorders (DSM-IV; American Psychiatric Association, 1994) as the phenotype, and an ordinal trait model where the phenotype was the number of DSM-IV criteria endorsed, adjusted for the number of criteria endorsed for other major substance dependence diagnoses (cocaine, opioid, and nicotine). Rs2066702 and rs1229984 were among the top genome-wide significant findings for both models, and consistent with the literature, we found that the rs2066702*T (in AAs) and rs1229984*A (in EAs) alleles protected against AD. Thus, although collectively, results from candidate gene and genome-wide studies demonstrate that *ADH1B* variants reduce AD risk, the question remains as to whether these variants influence specific aspects of AD and if so, which ones. The primary goal of the present study was to assess the contributions of specific diagnostic criteria to the observed associations of *ADH1B* SNPs with AD.

A second phenotype related to alcohol metabolic activity is the maximum number of drinks consumed in a 24-hour period ("MaxDrinks"; Bierut *et al*, 2012; Kapoor *et al*, 2013; Saccone *et al*, 2000). Because as an intermediate phenotype, MaxDrinks may be more

closely related to the specific genetic mechanisms and biological pathways underlying AD risk, its use may enhance gene identification (Goldman and Ducci, 2007). Although other phenotypes have been used to examine the genetic influences on drinking behavior, such as the frequency of heavy drinking (Heath *et al*, 2011), MaxDrinks has been the most widely studied and has the advantage of being relatively easy to measure.

As expected of an intermediate phenotype, MaxDrinks is correlated with AD (Dick and Bierut, 2006; Dick *et al*, 2011; Grant *et al*, 2009; Kendler *et al*, 2010) and also has a heritability of about 50% (Saccone *et al*, 2000). In a study of adolescents, MaxDrinks was significantly correlated with the score on the Self-Rating of the Effects of Alcohol (SRE) First 5, a measure of initial sensitivity to alcohol, described as reflecting “innate tolerance” (Schuckit *et al*, 2005). In addition to its association with AD, *ADH1B* has been associated with MaxDrinks (rs1229984; Bierut *et al*, 2012; Macgregor *et al*, 2009; Meyers *et al*, 2013), as was a null mutation in another well-studied gene encoding a metabolic enzyme, *ALDH2*, in a Chinese population (Quillen *et al*, 2014). Thus, investigation of the relationship between *ADH1B* variants and MaxDrinks could help to elucidate the mechanisms underlying the observed association with AD.

Changes in alcohol-related diagnoses could also influence associations with variation in alcohol metabolizing genes. The DSM-IV (American Psychiatric Association, 1994) differentiated alcohol use disorders (AUDs) into AD (based on the endorsement of three of seven criteria) and alcohol abuse (based on the endorsement of any one of four criteria). In the recent revision of the DSM (DSM-5; American Psychiatric Association, 2013), alcohol abuse and dependence criteria were combined to yield a single diagnosis of alcohol use disorder (AUD). This decision was based on findings from item response theory analyses that consistently demonstrated that 10 of the 11 combined criteria map onto a unidimensional continuum of severity (Borges *et al*, 2010; Proudfoot *et al*, 2006; Saha *et al*, 2006). The utility of the DSM-5 diagnosis for gene identification has not yet been tested.

Here, we used our GWAS sample of EA and AA subjects that was carefully assessed for substance use disorder diagnoses to gain insight into the mechanism by which *ADH1B* variants influence AD risk. We hypothesized that specific criteria contributed to the observed associations with AD, which we sought to identify. We also investigated the impact of the change from DSM-IV AD to DSM-5 AUD on the strength of association of *ADH1B* variants. A secondary aim of the study was to explore the relationship between *ADH1B* variants and MaxDrinks, a proposed intermediate phenotype for AD.

Materials and Methods

Subjects and phenotyping procedures

Subjects were recruited from five sites for studies of the genetics of drug and alcohol dependence: Yale University School of Medicine (APT Foundation; New Haven, CT), the University of Connecticut Health Center (Farmington, CT), McLean Hospital (Harvard Medical School; Belmont, MA), the Medical University of South Carolina (Charleston, SC), and the University of Pennsylvania (Philadelphia, PA). Subjects gave written informed consent as approved by the institutional review board at each site, and certificates of

confidentiality were obtained from NIDA and NIAAA. The current study was restricted to AA and EA subjects (with Hispanic subjects assigned genetically to one of those groups). Based on the findings of the published GWAS, analyses of rs2066702 were conducted in the AA part of the sample ($P=2.18\times 10^{-9}$ in the GWAS, not significant in EAs (probably because of minimal information)) and analyses of rs1229984 were conducted in the EA part of the sample ($P=6.75\times 10^{-14}$ in the GWAS, not significant in AAs (probably because of minimal information)). Although the sample consisted largely of unrelated individuals, there was a subset of subjects from small nuclear families in both populations. Subjects were interviewed with the computer-assisted Semi-Structured Assessment for Drug Dependence and Alcoholism (SSADDA; Pierucci-Lagha *et al*, 2005; 2007), a psychiatric interview that assesses physical, psychological, social, and psychiatric manifestations of substance use disorders. This part of the sample constituted most of the “discovery” sample for our published AD GWAS.

The following DSM diagnostic criteria were evaluated with the SSADDA: giving up or reducing important social, occupational, or recreational activities because of alcohol use (“activities given up”); a great deal of time spent in activities necessary to obtain, use, or recover from alcohol use (“much time spent using”); continuing alcohol use despite persistent or recurrent physical or psychological problems (“physical/psychological problems”); persistent desire or unsuccessful efforts to cut down or control alcohol use (“repeated attempts to quit”); a need for markedly increased amounts of alcohol to achieve intoxication or desired effect or markedly diminished effect with continued use of the same amount of alcohol (“tolerance”); using alcohol in larger amounts or over a longer period than was intended (“used larger amounts/longer”); the characteristic withdrawal syndrome for alcohol or alcohol is used to relieve or avoid withdrawal symptoms (“withdrawal”); a strong desire or urge to use alcohol (“craving”); recurrent alcohol use in situations in which it is physically hazardous (“hazardous use”); recurrent alcohol use resulting in a failure to fulfill major role obligations at work, school, or home (“neglected major roles”); and continued alcohol use despite having persistent or recurrent social or interpersonal problems caused or exacerbated by the effects of the alcohol (“social/interpersonal problems”). During the interview, subjects were also asked “In your lifetime, what is the largest number of drinks you have ever had in a 24-hour period?” (“MaxDrinks”).

Genotyping and quality control

Samples were genotyped on the Illumina HumanOmni1-Quad v1.0 microarray at the Center for Inherited Disease Research and the Yale Center for Genome Analysis. Following quality control, 5,697 individuals and 889,659 SNPs were available for imputation. Detailed genotyping and quality control methods are given in Gelernter *et al* (2014). Both rs2066702 and rs1229984 were well imputed and genotype dosages were rounded to the nearest integer for this analysis (i.e., 0, 1, 2).

We verified subjects’ self-reported ancestry with principal component analysis (PCA) implemented by the SmartPCA component of EIGENSOFT (Patterson *et al*, 2006; Price *et al*, 2006). The first principal component separated AA subjects from EA subjects (with Hispanic subjects clustering with one of these groups). We then conducted PCA analyses in

each sample separately and used the first three principal components as covariates in the statistical analyses to correct for residual population structure.

Statistical analyses

Chi-squared tests were used to evaluate the differences in DSM diagnosis and criterion prevalence in each population group. Tetrachoric and biserial correlations among DSM criteria and among DSM criteria and MaxDrinks, respectively, were computed with the psych package in R (Revelle, 2014). To account for within-family correlations in responses (MaxDrinks, rs2066702, and rs1229984), all of our analyses used robust sandwich standard errors, implemented through the geepack package (Halekoh *et al*, 2006) in R (R Development Core Team, 2014), with an exchangeable correlation structure specified. For each model, age, gender, and the first three ancestry PCs were included as covariates. In both the AA and EA samples, the MaxDrinks phenotype was Winsorized to address overestimation of self-reported values using a cutoff of 50 drinks. For both SNPs, due to the small number of minor-allele homozygotes, we combined the heterozygote and homozygote minor-allele groups.

Using a logistic regression model, we tested for association between rs2066702 in AAs and rs1229984 in EAs and the number of DSM-IV AD criteria and DSM-5 AUD criteria. We did not test for association with DSM-IV AD or DSM-5 AUD diagnoses, but used the criterion count regardless of clustering. SNP genotype group was the dichotomous response variable and the 7 (DSM-IV) or 11 (DSM-5) criteria were the explanatory variables. We used the fmsb package in R (Nakazawa, 2012) to calculate the percentage of variance explained by these SNPs (Nagelkerke's R^2). We note that we could not account for family structure in the calculations of Nagelkerke's R^2 . We tested for association between rs2066702 and rs1229984 and MaxDrinks using a linear regression model, where MaxDrinks was the response variable and SNP genotype group was the dichotomous explanatory variable.

We also ran sequential logistic regression analyses, with the dichotomous rs2066702 and rs1229984 genotypes as responses, to determine the relative contribution of each DSM criterion (for either the set of 7 DSM-IV dependence criteria or 11 DSM-5 AUD criteria). In Step 1, each criterion was entered into the model separately without adjustment for other criteria. In Step 2, models were adjusted for the criterion that yielded the most significant Wald z-statistic in Step 1. This process was repeated until no other criteria contributed significantly to the model (defined as $P < 0.05$). Model fit was assessed with the Quasi-likelihood under Independence Model Criterion (QIC) goodness of fit statistic (Pan, 2001) calculated with the MESS package in R. To ensure that our results were not confounded by the severity of the disorder, we performed a secondary analysis comparing the fit of the model including all criteria versus the fit of the models where one criterion was dropped out (i.e., 7 models for DSM-IV and 11 models for DSM-5). In all analyses, the major allele group (T/T for rs2066702 and G/G for rs1229984) served as the reference; the major allele for both SNPs was also the risk allele. Thus, in Tables 2 through 5, an odds ratio greater than one for a criterion indicates that a subject endorsing that criterion was more likely to have the risk allele than a subject not endorsing that criterion (all else being equal).

Using path analyses implemented in Mplus version 7.3 (Muthén and Muthén, 1998-2012), we tested for indirect effects of *ADH1B* on alcohol outcomes with MaxDrinks as a mediator variable. Path analysis allows for the parsing of indirect and direct effects of a predictor on an outcome. Individuals with missing values for MaxDrinks were excluded from the analysis (n=101 AAs, 50 EAs). A maximum likelihood estimator was used for continuous outcomes (DSM-IV and DSM-5 criterion counts) and a weighted least squares means and variance adjusted parameter estimator was used for binary outcomes (individual criteria). Age, sex, and the first three ancestry principal components were used as covariates in the analyses. The significance of the indirect effect of *ADH1B* on alcohol phenotypes via MaxDrinks was estimated by multiplication of the regression coefficients of *ADH1B* on MaxDrinks and of MaxDrinks on the alcohol phenotype. To address non-normality of indirect effects, bias-corrected bootstrapped 95% confidence intervals were estimated from 1,000 draws.

Replication analyses

We sought to replicate our findings in publicly available data from the Study of Addiction: Genetics and Environment (SAGE). Data were obtained from the database of Genotypes and Phenotypes (dbGaP, accession number phs000092.v1.p1). The sample consists of 1,103 unrelated AA subjects and 2,733 unrelated EA subjects. We used statistical methods largely identical to those described above, with some exceptions. Because not all required criteria for DSM-5 were available, we tested for replication using only DSM-IV AD criteria. Additionally, we did not include ancestry PCs as covariates for this analysis, as they were not available. We did not embed models in GEE because, in contrast to the primary study sample, there were no related subjects in the SAGE sample. In the replication mediation analyses, we excluded subjects with missing data (n=24 AAs, 21 EAs).

Results

Sample characteristics

A summary of sample demographics and characteristics is presented in Table 1. The GWAS sample consisted of 3,318 African-American (AA) and 2,379 European-American (EA) subjects. Due to missing data, 17 AAs and 11 EAs were excluded from the analyses presented here. Both samples were, on average, approximately 40 years old, largely male, never married, with a high school education, and mostly unemployed at the time of the study. The EA sample had a higher prevalence of DSM-IV AD, DSM-IV alcohol abuse, and DSM-5 severe AUD ($P<0.001$, Table S1). Of the 11 DSM-5 AUD criteria, the most commonly endorsed criterion in both populations was “used larger amounts/longer” (64% in AAs, 70.5% in EAs, Table S1). With the exception of the “repeated attempts to quit” criterion, all criteria showed significantly higher endorsement in EAs ($P<0.001$, Table S1).

In the phenotype data, we observed substantial collinearity among the 11 DSM-5 criteria (which include all 7 of the DSM-IV dependence criteria), with tetrachoric correlations ranging from 0.61-0.89 in AAs and 0.60-0.86 in EAs (all correlations $P<2\times 10^{-16}$, Figure 1). MaxDrinks was moderately correlated with each criterion (biserial, 0.47-0.63 in AAs, 0.43-0.60 in EAs; Figure 1) and with the overall number of criteria endorsed (DSM-IV: $r=0.55$, $P<2\times 10^{-16}$ and DSM-5: $r=0.57$, $P<2\times 10^{-16}$).

Association of ADH1B variants with MaxDrinks and number of DSM criteria

Consistent with our prior findings (Gelernter *et al*, 2014), we observed statistically significant associations between the number of DSM-IV criteria and rs2066702 in AAs ($\beta=0.09$, $SE=0.014$, Wald test $P=1.9\times 10^{-9}$, $R^2 = 0.028$, Figure 2) and rs1229984 in EAs ($\beta=0.18$, $SE=0.024$, Wald $P= 1.4\times 10^{-13}$, $R^2 = 0.029$, Figure 2). Subjects with the rs2066702*C/C genotype or the rs1229984*G/G genotype (i.e., the major allele homozygote groups for both SNPs) had higher criterion counts than the combined heterozygote and minor allele homozygote group. We also observed significant associations of both *ADH1B* variants and the number of DSM-5 criteria endorsed (rs2066702 $\beta=0.06$, $SE=0.009$, Wald test $P=1.4\times 10^{-9}$, $R^2=0.067$, Figure 2; rs1229984 $\beta=0.12$, $SE=0.009$, Wald test $P= 5.3\times 10^{-14}$, $R^2=0.067$, Figure 2). Interestingly, we also observed that EAs endorsed more criteria than AAs (DSM-IV $P=8.7\times 10^{-11}$; DSM-5 $P<2\times 10^{-16}$). Neither of these analyses was reported by Gelernter *et al* (2014).

Our hypothesis that these SNPs would be associated with the maximum number of drinks consumed in a single 24-hour period (MaxDrinks) was also supported. Individuals in the major allele homozygote groups (rs2066702 T/T and rs1229984 G/G) for both *ADH1B* SNPs reported greater MaxDrinks (rs2066702 $\beta=2.67$, $SE=0.49$, Wald $P=6.4\times 10^{-8}$, Figure 3; rs1229984 $\beta=5.17$, $SE=0.75$, Wald $P=5.2\times 10^{-12}$, Figure 3). Similar to what we found for the number of DSM-IV and DSM-5 criteria, EAs had higher MaxDrinks than AAs ($P<2.0\times 10^{-16}$).

Sequential logistic regression analyses: rs2066702 in AAs

We hypothesized that, despite collinearity among the DSM criteria, a limited number of them were primarily responsible for the overall association between the *ADH1B* SNPs and the number of DSM-IV or DSM-5 criteria endorsed. To test this hypothesis, we performed sequential logistic regression analyses, evaluating the contributions of each criterion to the model, adjusted for age, sex, and the first three ancestry PCs (base model, $QIC=4230.84$). In Table 2, we present the results for DSM-IV criteria and rs2066702 in AAs. In Step 1, when the criteria were entered into the base model separately (without adjustment for other criteria), all were significantly associated with rs2066702 genotype, with tolerance being the strongest predictor (Wald z-ratio=28.58, $P=9.0\times 10^{-8}$). In Step 2, when we repeated the same analysis after adjusting for tolerance, much time spent using was the strongest predictor (Wald z-ratio=9.17, $P=0.002$ after adjustment for tolerance). In Step 3, after adjusting for both tolerance and much time spent using, no other criterion significantly predicted genotype. Thus, differences in tolerance and much time spent using largely accounted for the association between rs2066702 and AD. This was underscored by the finding that the model fit was better after adjustment for tolerance and much time spent using than after adjustment for all DSM-IV criteria ($QIC=4198.70$ vs. 4204.80).

We obtained very similar results for AAs using the DSM-5 criteria (Table 3). Again, tolerance was the strongest predictor in Step 1 (Wald z-ratio=28.58, $P=9.0\times 10^{-8}$) and much time spent using was the strongest predictor in Step 2 (Wald z-ratio=9.17, $P=0.002$), with no significant contributors in Step 3. In addition, the fit for the model adjusted for tolerance and

much time spent using was better than the fit for all DSM-5 criteria (QIC=4198.70 vs. 4211.60).

Consistent with our findings in AAs that tolerance contributed most to the association in the sequential logistic regression analyses, when we included all criteria in the model (either DSM-IV or DSM-5), tolerance was the only criterion that was significant (DSM-IV Wald z-ratio=6.22, $P=0.013$; DSM-5 Wald z-ratio=5.01, $P=0.025$; Table S2).

Sequential logistic regression analyses: rs1229984 in EAs

We present the results for DSM-IV criteria and rs1229984 for EAs in Table 4. In Step 1, withdrawal was the strongest predictor of rs1229984 genotype (Wald z-ratio=35.98, $P=2.0\times 10^{-8}$). In Step 2, after adjusting for withdrawal, used larger/amounts longer was the next best predictor (Wald z-ratio=10.95, $P=9.3\times 10^{-4}$). In Step 3, only tolerance contributed significantly to the model that was adjusted for withdrawal and used larger amounts/longer (Wald z-ratio=4.52, $P=0.033$). In Step 4, after adjustment for withdrawal, used larger amounts/longer, and tolerance, no other criteria contributed significantly. Confirming that withdrawal, used larger amounts/longer, and tolerance largely explained the association with DSM-IV AD in EAs, the fit for the model that adjusted for these predictors was better than that for the model incorporating all DSM-IV criteria (QIC=1821.38 vs. 1826.00).

Table 5 displays the results for DSM-5 criteria in EAs. In contrast to what we observed for the DSM-IV criteria, the social/interpersonal problems criterion was the strongest predictor of rs1229984 genotype (Wald z-ratio=37.66, $P=8.4\times 10^{-10}$). In Step 2, after adjusting for social/interpersonal problems, withdrawal was the next best predictor (Wald z-ratio=14.70, $P=1.3\times 10^{-4}$). In Step 3, tolerance was the best predictor, similar to what we observed for the DSM-IV model (Wald z-ratio=4.50, $P=0.034$). After adjustment for social/interpersonal problems, withdrawal, and tolerance, no other criteria contributed significantly. The fit for these three criteria was better than that for the model incorporating all DSM-5 criteria (QIC=1818.96 vs. 1830.30)

When we included all DSM-IV criteria in the model, withdrawal was the only criterion that remained significant (Wald z-ratio=4.91, $P=0.027$; Table S3). However, in the full DSM-5 model, no individual criterion was significant.

To ensure that our results were not confounded by the severity of AD/AUD, we performed an alternative analysis in which we compared the fit of the model that included all criteria to the fit of the models in which one criterion was dropped out (i.e., 7 models for DSM-IV and 11 models for DSM-5). For both SNPs and both diagnostic systems, we found that the major criteria identified by adding in criteria to the model were the same that impacted the model fit most when they were dropped from the full model (tolerance, much time spent using for rs2066702; withdrawal, used larger amounts/longer for rs1229984; Table S4).

Mediation by MaxDrinks of the associations between ADH1B SNPs and alcohol phenotypes

Using path analysis, we tested the hypothesis that MaxDrinks mediated the relationship between *ADH1B* and the DSM criterion count and individual criteria. The results for the

mediation analyses are shown in Table S7 (rs2066702 in AAs) and Table S8 (rs1229984 in EAs). In AAs, we found significant indirect effects of rs2066702 via MaxDrinks on the DSM-IV criterion count, the DSM-5 criterion count, and individual criteria (Table S7), suggesting that MaxDrinks partially mediates the associations between rs2066702 and alcohol phenotypes. For some criteria, the direct effect of rs2066702 was no longer significant when MaxDrinks was included in the model (used larger amounts/longer, neglected major roles, physical/psychological problems, and repeated attempts to quit), suggesting that in these cases mediation by MaxDrinks is substantial. Similarly, in EAs the indirect effect of *ADH1B* via MaxDrinks was significant for all alcohol phenotypes (Table S8). However, in this population, the inclusion of MaxDrinks did not render any direct associations non-significant, suggesting that the mediation by MaxDrinks may be weaker in EAs.

Replication results

Consistent with the findings in our sample, we observed statistically significant associations between both *ADH1B* SNPs and both the number of DSM-IV criteria endorsed (rs2066702: $\beta=0.08$, $SE=0.026$, $P=0.002$; rs1229984: $\beta=0.006$, $SE=0.002$, $P=0.003$; Figure S1) and MaxDrinks (rs2066702: $\beta=2.21$, $SE=0.89$, $P=0.013$; rs1229984: $\beta=2.88$, $SE=0.90$, $P=0.001$; Figure S1).

We present the results for the sequential logistic regression replication analyses in Tables S5 and S6. Consistent with the findings in our sample, we observed that tolerance was able to explain the association between rs2066702 and number of DSM-IV criteria in the SAGE AA replication sample. After adjusting for tolerance in Step 2, no other criteria were significant (Table S5). In contrast, we did not replicate our findings for rs1229984, i.e., withdrawal did not contribute to the association (Table S6).

Results for the mediation analyses in the SAGE sample are found in Tables S9 and S10. Similar to the findings in our sample, there was a significant indirect effect of *ADH1B* SNPs on the DSM-IV criterion count and individual criteria via MaxDrinks in both populations. For rs2066702 in AAs, the direct effect of *ADH1B* was no longer significant when MaxDrinks was present in the model for nearly all phenotypes (excluding tolerance). For rs1229984 in EAs, we also observed that the direct effects of *ADH1B* were no longer significant for nearly all phenotypes (excluding repeated attempts to quit).

Discussion

The results of this study demonstrate that a limited number of diagnostic criteria contributed to the associations observed between *ADH1B* variants and AD in our recent GWAS (Gelernter *et al*, 2014). In each model, once these criteria were accounted for, the addition of other criteria had little impact. Specifically, we found that tolerance and much time spent using largely explained the association between rs2066702 and the number of DSM-IV AD criteria endorsed by AAs. In contrast, withdrawal, used larger amounts/longer, and tolerance explained the association between rs1229984 and the number of DSM-IV AD criteria endorsed by EAs. Extension of these results to the DSM-5 criteria for AUD yielded findings that for rs2066702 were very similar to those for DSM-IV AD, but for rs1229984 social/

interpersonal problems was the most significant DSM-5 criterion. We also found that both of the *ADH1B* SNPs (rs2066702 in AAs and rs1229984 in EAs) were associated with the maximum number of drinks consumed in a 24-hour period (also reported in Sartor *et al*, 2014 and observed by K. Xu, personal communication, June 15, 2014). We found that MaxDrinks was moderately correlated with DSM criterion counts and that it partially mediated the associations between *ADH1B* SNPs and the DSM-IV criterion count, the DSM-5 criterion count, and individual DSM criteria, consistent with it being an intermediate phenotype of the DSM diagnosis of AD/AUD.

We sought to replicate our findings in the SAGE sample, which is the only publicly available GWAS dataset for AD other than our own. For both SNPs, we replicated the associations that we observed in our sample with the number of AD criteria endorsed and MaxDrinks (Figure S1) and found that MaxDrinks partially mediated the associations between *ADH1B* SNPs and the DSM-IV criterion count and individual criteria (Table S7 and S8). We also replicated our finding that tolerance explained the association between rs2066702 and DSM-IV AD in AAs (Table S5). However, we did not replicate the finding of an association between withdrawal and rs1229984 in EAs (Table S6). One possible explanation for the lack of association of rs1229984 with alcohol withdrawal was that the SAGE EA sample was much less severely affected than our EA sample. This was evidenced by the difference in mean number of AD criteria endorsed in the risk genotype group (4 criteria in our sample, which is moderately severe vs. 2 criteria in SAGE, which is subthreshold for DSM-IV AD). The replication effort was also limited by the fact that we were unable to test the DSM-5 criteria, as these were not available in the SAGE dataset. Thus, this finding must be validated in other samples.

A recently published article examined the relationship between rs1229984 and individual DSM-IV AUD criteria in an Israeli population-based sample (Kilcoyne *et al*, 2014). The authors found that tolerance, repeated attempts to quit, used larger amounts longer, physical/psychological problems, hazardous use, and social problems were significantly associated with *ADH1B*. Although the methods differed somewhat from those used in our study, it is notable that this study did not identify an association between withdrawal and rs1229984. However, similar to the SAGE sample, the Israeli sample was less severely affected than our sample and the overall prevalence of withdrawal was much higher in our sample (40.8% in our sample vs. 14.7% in the Israeli sample), suggesting that the contrasting findings may be a result of differences in AUD severity.

Results from the sequential logistic regression analyses, the analyses of the effect of genotype on MaxDrinks, and the MaxDrinks mediation analyses converge on a role for *ADH1B* in influencing heavy alcohol consumption. Tolerance and withdrawal were the top criteria contributing to the associations with rs2066702 and rs1229984, respectively (though the finding for withdrawal did not replicate in the SAGE sample). These criteria are thought to be manifestations of adaptation (metabolic and/or neuronal) to sustained alcohol intake, but they appear to have different underlying mechanisms (Littleton, 1998). These criteria showed moderate correlation in both samples (AA: $r=0.68$, EA: $r=0.65$, Figure 1). Tolerance was the only criterion that emerged as a significant predictor of genotype in all models tested, and for rs2066702, it was the only criterion significant in the full model for both

DSM-IV and DSM-5 criteria (Table S2). The significant contribution of tolerance to the models presented here is also noteworthy in light of the association with MaxDrinks, which is shown in Figure 3. MaxDrinks is correlated with innate tolerance, or the initial lack of sensitivity to alcohol's stimulant and sedative effects (Chung and Martin, 2009; Schuckit *et al*, 2005). While one model postulates that increased tolerance to alcohol's effects predicts future risk for AD (Schuckit, 1980), recent longitudinal studies have shown that greater sensitivity to the subjective effects of alcohol may predict AD risk (King *et al*, 2014). Irrespective of the direction of the effect, it is clear that tolerance is an important aspect of AUD and based on our findings, further investigation of the effects of *ADH1B* on tolerance is warranted.

The influence of *ADH1B* on alcohol consumption and subsequent alcohol outcomes is further underscored by the results of our mediation analysis. We found that for all outcomes (the DSM-IV and DSM-5 criterion counts, individual criteria), MaxDrinks significantly mediated the association with *ADH1B* polymorphisms, suggesting that *ADH1B* may exert its effect on risk for AUD through alcohol consumption. In their recent study Kilcoyne *et al* also examined the role of MaxDrinks in mediating the effects of rs1229984 on alcohol outcomes in their Israeli sample. They found that, for criteria significantly associated with *ADH1B*, the effect of MaxDrinks explained 23-74% of the associations (Kilcoyne *et al*, 2014). Although the methods from our study differed from Kilcoyne *et al* (path analysis vs. multiple regression), both studies support the notion that alcohol consumption is an important mediator of the associations between *ADH1B* and AUD.

Our analysis of the association of rs1229984 and DSM-5 criteria in EAs showed that the social/interpersonal problems criterion was the most important predictor of genotype, suggesting that *ADH1B* genotype may predict social consequences of alcohol use. A study in Asian-Americans found that *ALDH2* genotype predicted social consequences of drinking, however this was not observed for *ADH1B* in that sample (Hendershot *et al*, 2009). The grouping of alcohol dependence and abuse criteria together has been criticized for the addition of three additional psychosocial criteria to the AD diagnosis (Meyer, 2011), which are thought to be less informative for diagnosis than the biological criteria that form the basis for the dependence syndrome. However, the reliability of the social/interpersonal problems is among the highest of all criteria for AUD (test-retest $\kappa=0.77$, inter-rater $\kappa=0.60$; C. Denis, personal communication, May 2, 2014). Additionally, in a staging analysis, heavy drinking with associated social problems was found to be one of the earliest symptoms to occur in the development of AUDs (Martin *et al*, 1995). Our results suggest that psychosocial consequences are an important aspect of AUDs that are potentially influenced by variation in *ADH1B*.

Because both of the variants examined in this study were in *ADH1B*, the differential association of the variant with diagnostic criteria provides evidence for either population-specific or variant-specific effects. Because each variant is sufficiently informative for analysis in only one population, the specificity of effect attributable to the variants themselves cannot be evaluated. This extends the findings from the original GWAS, in which rs2066702 was not significantly associated with DSM-IV AD in EAs and rs1229984 was not significantly associated with DSM-IV AD in AAs. However, it is important to note that the

minor allele frequencies of rs2066702 and rs1229984 were low in EAs and AAs, respectively (1% and 2%; Gelernter *et al*, 2014). Although tolerance was a recurring theme for both variants, it showed a stronger influence on the association of rs2066702 in AAs than for rs1229984 in EAs. The findings reported here suggest that the structure of the syndromes (AD in DSM-IV and AUD in DSM-5) differs substantially by population. Furthermore, we found that AA subjects had lower DSM-IV and DSM-5 criterion counts and a lower number of MaxDrinks than EAs. These results corroborate findings from the 2001-2002 National Epidemiologic Survey on Alcohol and Related Conditions (NESARC), which reported a lower incidence of alcohol dependence among AAs than EAs (3.29% vs 5.10%; Grant *et al*, 2004). But unlike the NESARC study, which was representative of the U.S. population, we cannot exclude ascertainment bias as an influence on this observation in our sample.

The associations of *ADH1B* with DSM-5 criteria tended to be slightly stronger than those with DSM-IV criteria (Figure 2). These results are interesting because they argue against the contention that excess sensitivity and inadequate specificity in DSM-5 could lead to over diagnosis and less reliable, noisier phenotypes (Meyer, 2011). Although in one study (Mewton *et al*, 2011), there was a 61.7% increase in the prevalence of AUD from DSM-IV to DSM-5, others have found only modest increases (11.3%: Agrawal *et al*, 2011; 5.1%: Edwards *et al*, 2013; 0.8%: Peer *et al*, 2013). In a large sample of twin pairs, Edwards *et al*. (2013) noted similar heritability estimates and genetic correlations for DSM-IV and DSM-5 diagnoses. Furthermore, a recent study found higher test-retest and inter-rater reliabilities for DSM-5 substance use disorder diagnoses than DSM-IV diagnoses in a sample interviewed using the SSADDA, the diagnostic instrument used in the current study (C. Denis, personal communication, May 2, 2014). Taken together, the findings suggest that the revision of the diagnostic categories and criteria in DSM-5 may increase the utility of the AUD diagnosis for gene finding.

There are limitations to our study that should be considered. First, as demonstrated in the correlation analyses presented in Figure 1, there was substantial collinearity among the 11 DSM-5 diagnostic criteria. Although this is consistent with the results from the IRT analyses showing that the criteria lie along a single continuum (Borges *et al*, 2010; Proudfoot *et al*, 2006; Saha *et al*, 2006), it may have limited our ability to discern which criteria contribute to the observed associations. In addition, MaxDrinks was based entirely on self-report, which may not reflect the true value in all cases. Furthermore, estimates of MaxDrinks on one occasion may be less reliable than estimates of alcohol-related criteria, which have minimum frequency and/or duration thresholds. The operationalization of certain criteria in the SSADDA may have affected our results. Tolerance is operationalized as a dichotomous variable based on a cutoff (“after drinking for some years, needed 50% more alcohol to get an effect”), which may be subject to interpretation. Similarly, while equivalent to the DSM-5 criterion, the operationalization of the craving criterion in the SSADDA (“a strong desire or urge to use alcohol”) as a single item may not be optimal for use in diagnoses and genetic analyses (Agrawal *et al.*, 2014). Nevertheless, the reliabilities of these items in the SSADDA were fair to excellent (tolerance: test-retest $\kappa=0.88$, inter-rater $\kappa=0.54$; craving: test-retest $\kappa=0.69$, inter-rater $\kappa=0.69$; C. Denis, personal communication, May 2, 2014). Lastly, the affected sample that we analyzed had a variety of co-occurring substance dependence diagnoses (i.e., those involving cocaine, nicotine, marijuana, and opioids). In our analyses,

to avoid over-fitting of the models, we included the individual alcohol criteria as predictors without controlling for the other major substance use disorder criteria that were endorsed by subjects.

Our study had several strengths. We analyzed a large dataset that allowed us to detect significant effects in a GWAS (Gelernter *et al*, 2014). Furthermore, the phenotype data analyzed here were obtained using a diagnostic instrument with good diagnostic and criterion-level reliability (Pierucci-Lagha *et al*, 2005; 2007). As a result of our previous GWAS, we had very strong prior associations to consider here, providing a firm foundation for an examination of the specific criteria contributing to the observed associations. Additionally, our study was strengthened by the analysis of specific variants in two different populations, which allowed us to examine population-specific effects. Lastly, the comparison of DSM-IV and DSM-5 results was a novel aspect of the study that addresses the utility of the changes in the diagnostic system for research on AUD.

This study represents an effort to further elucidate the effects of *ADH1B* on AUD, with the aim of better understanding the specific effects of polymorphisms associated with AUD, and the ultimate goal of generating biomarkers for risk. It should be noted that the observed effects of *ADH1B* variants are small, which is consistent with findings from the majority of GWAS for complex traits. Further studies will be necessary to evaluate the clinical utility of these findings. Presumably other variants show specific association with other DSM criteria, and it would be of great interest to identify them. Additionally, further research that examines the effects of *ADH1B* variation on adaptations resulting from heavy alcohol consumption is warranted.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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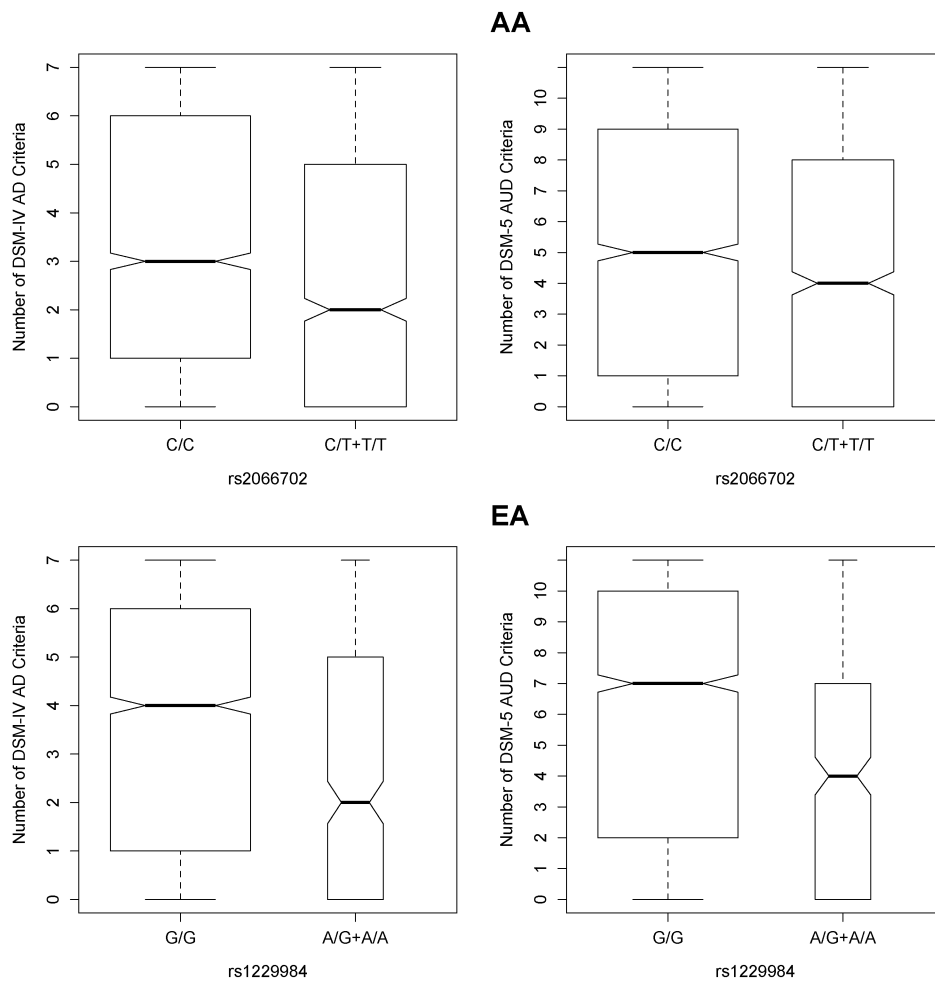


Figure 2. Association of *ADH1B* single nucleotide polymorphisms (SNPs) with the number of DSM-IV alcohol dependence (AD) and DSM-5 alcohol use disorder (AUD) criteria endorsed in African-Americans (AA) and European-Americans (EA). Top panel: AA individuals homozygous for the major allele of rs2066702 (C/C) showed higher number of DSM-IV (left) and DSM-5 (right) criteria endorsed ($P=1.9\times 10^{-9}$ and $P=1.4\times 10^{-9}$). Bottom panel: EA individuals homozygous for the major allele of rs1229984 (G/G) showed higher number of DSM-IV (left) and DSM-5 (right) criteria endorsed ($P=1.4\times 10^{-13}$ and $P=5.3\times 10^{-14}$).

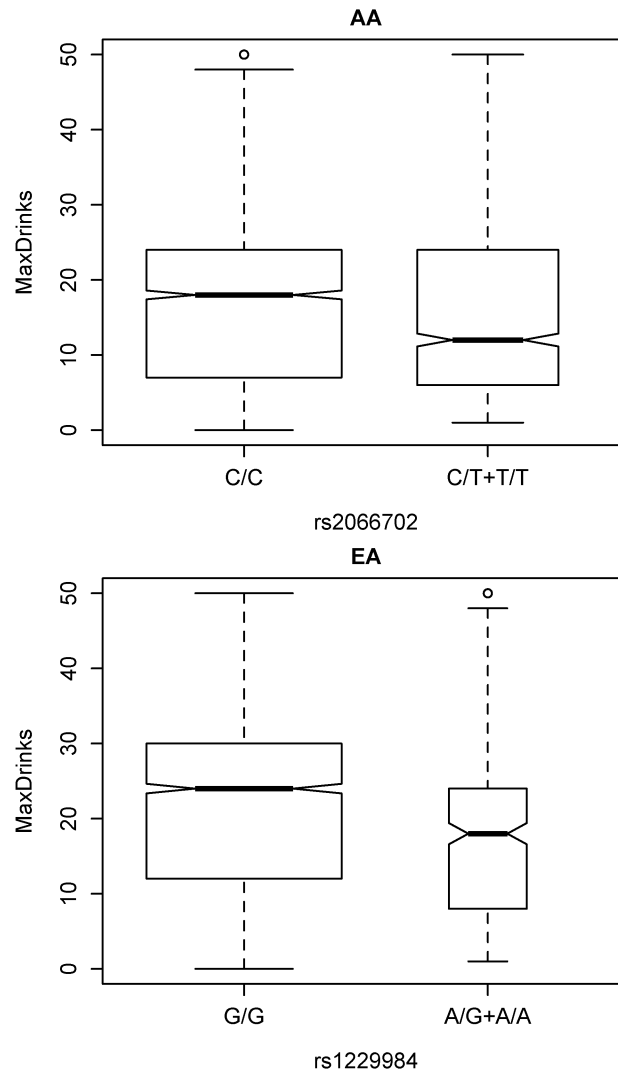


Figure 3. Association of *ADH1B* variants with the maximum number of drinks consumed in a 24-hour period (MaxDrinks) in African-Americans (AA) and European-Americans (EA). Top panel: AA individuals homozygous for the major allele of rs2066702 (C/C) reported greater MaxDrinks ($P=6.4 \times 10^{-8}$). Bottom panel: EA individuals homozygous for the major allele of rs12289984 (G/G) reported greater MaxDrinks ($P=5.2 \times 10^{-12}$).

Table 1

Sample demographics and characteristics.

	AA (n=3,301)	EA(n=2,368)
Age (yrs) Mean (SD)	41.5 (9.1)	38.1 (10.5)
Sex (% male)	52.0	58.8
Marital Status (% married) (% divorced/separated) (% widowed) (% never married)	13.9 23.1 2.3 60.7	13.6 28.8 2.0 55.6
Maximum education level (yrs) Mean (SD)	12.0 (1.99)	11.8 (2.32)
Employment (% currently employed (% full time))	36.5 (59.7)	32.5 (61.4)
Maximum drinks consumed in a 24-hr period Mean (SD)	18.2 (14.0)	23.6 (14.0)
DSM-IV/DSM-5 criterion count Mean (SD)	3.1 (2.5)/4.8(4.0)	3.6 (2.6)/5.8 (3.9)
rs2066702 genotype counts n=C/C, n=C/T+T/T	2162, 1139	-
rs1229984 genotype counts n=G/G, n=A/G+A/A	-	2044, 324

Abbreviations: AA, African-American; EA, European-American; AD, alcohol dependence; AUD, alcohol use disorder

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Table 2

Results from the sequential logistic regression analysis of rs2066702 genotype on DSM-IV AD criteria in African-Americans.

Criterion	Step 1 (unadjusted)			Step 2 (adjusted for tolerance)			Step 3 (adjusted for tolerance and much time spent using)					
	OR (95% CI)	Wald test (z-ratio)	P	QIC	OR (95% CI)	Wald test (z-ratio)	P	QIC	OR (95% CI)	Wald test (z-ratio)	P	QIC
Tolerance	1.48 (1.28-1.70)	28.58	9.0×10 ⁻⁸	4207.36	-	-	-	-	-	-	-	-
Much time spent using	1.48 (1.27-1.72)	26.12	3.2×10 ⁻⁷	4205.41	1.30 (1.10-1.53)	9.17	0.002	4198.70	-	-	-	-
Activities given up to use	1.43 (1.23-1.65)	23.10	1.5×10 ⁻⁶	4210.42	1.25 (1.07-1.47)	7.66	0.006	4201.57	1.15 (0.95-1.38)	2.12	0.145	4198.91
Withdrawal	1.47 (1.25-1.73)	21.75	3.1×10 ⁻⁶	4211.31	1.28 (1.07-1.52)	7.28	0.007	4201.90	1.17 (0.96-1.42)	2.57	0.109	4198.45
Used larger amounts/ longer	1.38 (1.19-1.60)	18.23	2.0×10 ⁻⁵	4213.45	1.17 (0.98-1.39)	3.14	0.076	4205.12	1.08 (0.89-1.30)	0.68	0.409	4199.70
Physical/ psychological problems	1.37 (1.17-1.59)	16.19	5.7×10 ⁻⁵	4217.09	1.18 (0.99-1.39)	3.55	0.059	4205.81	1.07 (0.89-1.29)	0.54	0.463	4200.37
Repeated attempts to quit	1.26 (1.10-1.45)	10.44	0.001	4221.03	1.07 (0.92-1.26)	0.78	0.377	4207.81	0.99 (0.84-1.17)	0.01	0.909	4200.65

Abbreviations: AD, alcohol dependence; OR, odds ratio; CI, confidence interval; P, p-value; QIC, quasi-likelihood under the independence model criterion

Table 3

Results from the sequential logistic regression analysis of rs2066702 genotype on DSM-5 AUD criteria in African-Americans.

Criterion	Step 1 (unadjusted)				Step 2 (adjusted for tolerance)				Step 3 (adjusted for tolerance and much time spent using)			
	OR (95% CI)	Wald test (z-ratio)	P	QIC	OR (95% CI)	Wald test (z-ratio)	P	QIC	OR (95% CI)	Wald test (z-ratio)	P	QIC
Tolerance	1.47 (1.28-1.70)	28.58	9.0×10 ⁻⁸	4207.36	-	-	-	-	-	-	-	-
Much time spent using	1.48 (1.27-1.72)	26.12	3.2×10 ⁻⁷	4205.41	1.29 (1.09-1.53)	9.17	0.002	4198.70	-	-	-	-
Activities given up to use	1.43 (1.23-1.65)	23.10	1.5×10 ⁻⁶	4210.42	1.25 (1.07-1.47)	7.66	0.006	4201.57	1.15 (0.95-1.38)	2.12	0.145	4198.91
Withdrawal	1.47 (1.25-1.73)	21.75	3.1×10 ⁻⁶	4211.31	1.28 (1.07-1.52)	7.28	0.007	4201.90	1.17 (0.96-1.42)	2.57	0.109	4198.45
Used larger amounts/ longer	1.38 (1.19-1.60)	18.23	2.0×10 ⁻⁵	4213.45	1.17 (0.98-1.39)	3.14	0.076	4205.12	1.08 (0.90-1.30)	0.68	0.409	4199.70
Hazardous use	1.37 (1.18-1.59)	17.68	2.6×10 ⁻⁵	4212.99	1.19 (1.01-1.40)	4.34	0.037	4203.78	1.00 (0.99-1.01)	1.67	0.196	4198.54
Neglected major roles	1.46 (1.25-1.69)	17.14	3.5×10 ⁻⁵	4207.50	1.27 (1.07-1.51)	7.77	0.005	4200.91	1.00 (0.99-1.01)	1.93	0.165	4198.95
Social/ interpersonal problems	1.36 (1.17-1.57)	17.06	3.6×10 ⁻⁵	4212.69	1.16 (0.98-1.37)	2.90	0.088	4204.78	1.00 (0.99-1.01)	0.24	0.621	4200.10
Craving	1.47 (1.25-1.73)	16.60	4.6×10 ⁻⁵	4215.33	1.28 (1.07-1.52)	4.58	0.032	4204.02	1.00 (0.99-1.01)	1.06	0.303	4199.60
Physical/ psychological problems	1.36 (1.17-1.59)	16.19	5.7×10 ⁻⁵	4217.09	1.18 (0.99-1.39)	3.55	0.059	4205.81	1.07 (0.89-1.29)	0.54	0.463	4200.37
Repeated attempts to quit	1.26 (1.09-1.45)	10.89	9.7×10 ⁻⁴	4221.03	1.07 (0.92-1.25)	0.78	0.377	4207.81	0.99 (0.84-1.17)	0.01	0.909	4200.65

Abbreviations: AD, alcohol dependence; OR, odds ratio; CI, confidence interval; P, p-value; QIC, quasi-likelihood under the independence model criterion

Table 5

Results from the sequential logistic regression analysis of rs1229984 genotype on DSM-5 AUD criteria in European-Americans.

Criterion	Step 1 (unadjusted)				Step 2 (adjusted for social/interpersonal problems)				Step 3 (adjusted for social/interpersonal problems and withdrawal)				Step 3 (adjusted for social/interpersonal problems, withdrawal, and tolerance)			
	OR (95% CI)	Wald test (z-ratio)	P	QIC	OR (95% CI)	Wald test (z-ratio)	P	QIC	OR (95% CI)	Wald test (z-ratio)	P	QIC	OR (95% CI)	Wald test (z-ratio)	P	QIC
Social/interpersonal problems	2.13 (1.67-2.71)	37.66	8.4 × 10 ⁻¹⁰	1833.83	-	-	-	-	-	-	-	-	-	-	-	-
Withdrawal	2.35 (1.78-3.10)	35.98	2.0 × 10 ⁻⁹	1835.52	1.28 (1.07-1.52)	14.70	1.3 × 10 ⁻⁴	1821.52	-	-	-	-	-	-	-	-
Used larger amounts/longer	2.07 (1.62-2.65)	33.58	6.8 × 10 ⁻⁹	1843.72	1.52 (1.12-2.04)	7.20	0.007	1827.77	1.35 (0.99-1.84)	3.63	0.057	1819.54	1.27 (0.92-1.75)	2.04	0.153	1818.83
Tolerance	1.96 (1.55-2.47)	31.17	2.4 × 10 ⁻⁸	1845.21	1.51 (1.15-1.98)	8.89	0.003	1826.84	1.70 (1.24-2.33)	4.50	0.034	1818.96	-	-	-	-
Much time spent using	2.08 (1.60-2.69)	30.57	3.2 × 10 ⁻⁸	1844.16	1.57 (1.17-2.11)	9.00	0.003	1827.51	1.30 (0.94-1.79)	2.58	0.108	1821.12	1.22 (0.88-1.70)	1.44	0.230	1819.60
Repeated attempts to quit	1.95 (1.53-2.50)	28.19	1.1 × 10 ⁻⁷	1842.80	1.47 (1.10-1.95)	6.95	0.008	1827.10	1.25 (0.94-1.68)	2.29	0.130	1820.07	1.18 (0.87-1.60)	1.12	0.291	1819.10
Physical/psychological problems	1.99 (1.53-2.57)	27.13	1.9 × 10 ⁻⁷	1844.96	1.51 (1.12-2.04)	7.39	0.007	1827.47	1.19 (0.85-1.68)	1.05	0.306	1822.21	1.14 (0.81-1.62)	0.57	0.450	1820.30
Craving	2.14 (1.61-2.86)	27.07	2.0 × 10 ⁻⁷	1848.69	1.28 (1.07-1.52)	9.82	0.002	1826.99	1.00 (0.99-1.01)	2.75	0.097	1821.48	1.30 (0.90-1.88)	2.01	0.157	1819.67
Neglected major roles	1.88 (1.47-2.40)	25.04	5.6 × 10 ⁻⁷	1849.32	1.32 (0.96-1.80)	2.89	0.089	1832.94	0.99 (0.98-1.00)	0.24	0.627	1823.31	1.02 (0.73-1.43)	0.02	0.897	1820.99
Activities given up to use	1.78 (1.39-2.28)	21.05	4.5 × 10 ⁻⁶	1851.79	1.27 (0.93-1.74)	2.33	0.127	1832.39	1.04 (0.74-1.46)	0.05	0.823	1823.61	0.98 (0.69-1.39)	0.02	0.902	1821.45
Hazardous use	1.74 (1.36-2.22)	19.49	1.0 × 10 ⁻⁵	1856.26	1.21 (0.90-1.61)	1.59	0.208	1833.60	0.99 (0.98-1.01)	0.68	0.408	1822.38	1.06 (0.77-1.43)	0.09	0.759	1820.74

Abbreviations: AD, alcohol dependence; OR, odds ratio; CI, confidence interval; P, p-value; QIC, quasi-likelihood under the independence model criterion