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The Abilities of Body Mass Index and Skinfold Thicknesses to Identify Children with Low or Elevated Levels of Dual-Energy X-Ray Absorptiometry–Determined Body Fatness

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Abstract

Objective—To examine the accuracies of body mass index (BMI) and skinfold thicknesses in classifying the body fatness of 7365 8- to 19-year-old subjects in a national sample.

Study design—We used percent body fat determined by dual-energy x-ray absorptiometry (PBF_{DXA}) between 1999 and 2004. Categories of PBF_{DXA} and the skinfold sum (triceps plus subscapular) were constructed so that that numbers of children in each category were similar to the number in each of 5 BMI categories based on the Centers for Disease Control and Prevention growth charts.

Results—Approximately 75% of the children and adolescents who had a BMI-for-age 95th percentile (considered obese) had elevated body fatness, but PBF_{DXA} levels were more variable at lower BMIs. For example, only 41% of the boys who had a BMI < 25th percentile, had a similarly low PBF_{DXA}. The use of the skinfold sum, rather than BMI, slightly improved the identification of elevated levels of body fatness among boys ($P = .03$), but not among girls ($P > .10$). A low sum of the triceps and subscapular skinfold thicknesses was a better indicator of low PBF_{DXA} than was a low BMI, but differences were smaller among children with greater levels of body fatness. Among girls who had a PBF_{DXA} above the median, for example, BMI and the skinfold sum were correlated similarly ($r = 0.77-0.79$) with body fatness.

Conclusions—Both BMI and skinfold thicknesses are fairly accurate in identifying children who have excess body fatness. In contrast, if the goal is to identify children who have low body fatness, skinfold thicknesses would be preferred.

Body mass index (BMI, kg/m²) is used as a screening tool for overweight and obesity in various settings, and a high BMI among children is associated with adverse levels of various cardiovascular risk factors, the initial stages of atherosclerosis, and adult obesity.¹⁻³

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However, because BMI is based on weight and height, both of which change greatly during growth, a high BMI can reflect a high level of either fat mass or fat-free mass.⁴ A child with a high BMI is likely to have elevated body fatness,⁵ but lower levels of BMI are a poor indicator of body fatness among children.⁶ In addition, several investigators^{7,8} have reported that the correlation between BMI and more accurate measures of body fatness among children and adolescents is only moderate ($r < 0.70$).

The thickness of various skinfolds is thought to give a more direct indication of body fatness than does BMI, and despite their large measurement errors,⁹ skinfold thicknesses are widely used.¹⁰⁻¹² Most, but not all,¹³ studies have found that skinfold thicknesses are more strongly associated with the body fatness of children than is BMI.^{6,14-17} A stronger correlation, however, does not necessarily mean that skinfold thicknesses can more accurately identify children who have high levels of body fatness than can BMI. The greater correlation may reflect the more accurate prediction of low levels of body fatness by skinfold thicknesses.

Our objective was to determine whether the sum of 2 (subscapular and triceps) skinfold thicknesses is more strongly related to dual-energy X-ray absorptiometry (DXA)-calculated body fatness than is BMI in a nationally, representative sample of 8- to 19-year-old subjects ($n = 7365$). We also assessed whether differences are caused by the identification of children who have relatively low or relatively high levels of body fatness.

Methods

The 1999-2004 National Health and Nutrition Examination Survey (NHANES) is a representative, cross-sectional sample of the US civilian, noninstitutionalized population.¹⁸ NHANES 1999-2004 underwent institutional review board approval, and parental permission was obtained for minors younger than the age of 18 years. Children 7-17 years of age also were asked to provide documented assent. Consent was obtained for all subjects 18 years and older. Age was calculated as age in months at the time of examination. Race and ethnicity were self-reported, and in the current study, we classify subjects as non-Hispanic white, non-Hispanic black, Mexican American, and other. The overall examination rate for 6- to 19-year-old subjects in NHANES 1999-2004 was 85%.

DXA scans were acquired in NHANES 1999-2004 for boys and nonpregnant girls who were at least 8 years of age with the use of a Hologic QDR 4500A fan-beam densitometer (Hologic Inc, Bedford, Massachusetts).^{19,20} The scan for each survey participant was analyzed using Hologic Discovery software (version 12.1) by the Department of Radiology in the University of California, San Francisco. PBF_{DXA} (ie, percent body fat determined by DXA) was calculated as $100 \times (DXA - \text{estimated total fat mass} / DXA - \text{estimated total mass})$. To protect patient confidentiality, the 1999-2000 DXA data for 8- to 17-year-old girls are available only in the Research Data Center. The current analyses do not include girls who were examined in the 1999-2000 cycle.

Our analyses used the NHANES DXA Multiple Imputation Data Files.²⁰ Approximately 9.5% of the children and adolescents in the current study were missing at least one DXA measurement, and because missingness was associated with BMI, body fatness, and other

characteristics, an analysis of only nonmissing data could be biased.²¹ Multiple imputation, performed by National Center for Health Statistics (NCHS) using sequential regression²⁰ with 5 imputations, was used to estimate missing DXA values from nonmissing DXA measurements and other characteristics such as race/ethnicity, age, BMI, and skinfold thicknesses. Although it has been suggested that a larger number of imputations may be necessary to arrive at accurate estimates of CIs and *P* values,^{22,23} these imputations were performed by NCHS before these reports were available.

Weight and height were measured using standardized techniques and equipment. BMI was calculated as weight (kg)/height (m)², and BMI-for-age z-scores (that account for sex and age) were calculated relative to children who participated in national studies between 1963-1965 and 1988-1994.²⁴ Overweight is defined as a BMI, relative to the 2000 Centers for Disease Control and Prevention (CDC) growth charts,²⁴ between the 85th and 94th percentiles for a child's sex and age; obesity is defined as a BMI \geq 95th percentile of this reference population. On the basis of these cut points, 16% of children in the current study were considered to be overweight, and 17% were obese. Because we were interested in comparing the accuracies of BMI and skinfold thicknesses over the entire distribution of body fatness, we also used the 25th and 50th percentiles of BMI in the CDC reference population as cut-points in the analyses. Approximately 16% of subjects in the current study had a BMI $<$ 25th percentile, 18% had a BMI that was between the 25th and 49th percentiles, and 33% had a BMI between the 50th and 84th percentiles.

The thickness of the triceps and subscapular skinfolds were measured to the nearest 0.1 mm by the use of Holtain skinfold calipers. These values were missing for about 7% (subscapular) and 4% (triceps) either because of measurement difficulties or because the skinfold exceeded the capacity of the caliper (45.0 mm).²⁵ Because the probability of a missing data for the skinfold thicknesses was not random, with more than 70% of these children having a BMI \geq 95th percentile, we used the "Amelia" package in R (<http://cran.r-project.org/web/packages/Amelia/index.html>)^{26,27} to impute values for the missing skinfolds based on levels of sex, race, age, BMI, PBF_{DXA}, and other characteristics. The skinfold thicknesses were log-transformed to avoid imputing improbably large values and to linearize the associations with PBF_{DXA}. We imputed one skinfold thickness value for each of the 5 sets of imputed DXA values; each set contained (if originally missing) an imputed value for the PBF_{DXA} and the 2 skinfold thicknesses.

We focus on the ability of 5 categories ($<$ 25th percentile, 25th-49th percentile, 50th-84th percentiles, 85th-94th percentiles, and \geq 95th percentile) of BMI-for-age and the sum of the triceps and subscapular skinfold thicknesses (SF sum) to correctly classify the body fatness of children. Because there is little agreement on the classification of "excess body fatness" among children or adults,²⁸ we constructed 5 categories (low, low-normal, moderate, slightly elevated, and elevated) for both SF sum and PBF_{DXA}, with each category having a similar number of children as the corresponding category of BMI-for-age.

This type of classification, with similar numbers of children in each BMI, SF sum, and percent body fat category, eliminates a possible bias in comparing screening performance. For example, if one compared extremely high values of the SF sum with less extreme values

of BMI in identifying high levels of body fatness, the positive predictive value of the SF sum would likely be greater. This could be true even if BMI was a more accurate indicator of body fatness, and results from the differing prevalences of high levels of BMI and the SF sum.

We used quantile regression²⁹ to calculate sex- and age-specific cut-points for 5 categories of the SF sum and percent body fat. For example, because 18% of the boys in the current study had a BMI-for-age 95th percentile, we considered boys with a PBF_{DXA} at or above the 82nd percentile to have an elevated level of PBF_{DXA} , and boys with a SF sum at or above the 82nd percentile to have an elevated SF sum. Age was modeled by the use of restricted cubic splines to account for nonlinearity, and a similar process was used to determine cut-points for other categories of the SF sum and body fatness. For example, because 14% of the girls had a BMI-for-age below the CDC 25th percentile, low levels of the SF sum and PBF_{DXA} were defined as values <14th (age-specific) percentile. Levels of PBF_{DXA} corresponding to each of the 5 categories of BMI-for-age, along with the terminology used throughout the article (Figure; available at www.jpeds.com).

Statistical Analyses

All analyses were performed with the survey and mitools packages in R,^{27,30} and accounted for the sample weights, sample design, and multiple imputations. We focused on the relation of the skinfold thickness and BMI to PBF_{DXA} , but additional analyses were performed that examined associations with total fat mass (kg).

We incorporated the uncertainty of the multiple imputations (for PBF_{DXA} and log skinfold thicknesses) into all standard errors³¹ by analyzing each of the 5 imputation sets separately. Estimates were then averaged over the 5 sets, and the total variance was calculated as the within-imputation variance plus $(1 + 1/5)$ times the between-imputation variance.^{20,32} The accuracy of the imputations was assessed by overimputing; a process in which each observed value was treated as if it had been missing. The correlation between the overimputed and actual values was strong ($r \sim 0.93$), and almost all of the observed values were within the 90% CIs of the overimputed values.

Although the skinfold thicknesses were log transformed before the imputation process, BMI levels, to a lesser extent, were positively skewed. However, as we did not impute any BMI values, this variable was not transformed. It has been shown³³ that even if the data are highly skewed, the results of regression models are accurate if the sample is relatively large (eg, >500).

The 15th, 50th, and 85th percentiles or proportions and SEs were used to summarize various characteristics of the sample. Sex-specific regression models were used to predict levels of PBF_{DXA} from race, age, BMI, and the skinfold thicknesses; age and BMI were modeled with restricted cubic splines. We focused on the sex-specific (weighted) correlations of PBF_{DXA} with BMI and SF sum. To control for the influence of age, these correlations were based on the residuals of regression models in which each body size measure was regressed on age. To determine the statistical significance of the observed differences (eg, the correlation between PBF_{DXA} and BMI minus the correlation between PBF_{DXA} and SF sum),

we calculated standard errors with the “withReplicates” function,³⁰ which uses jackknife replicate weights. The estimated variances were then combined across the 5 imputations, and *P* values were calculated using *t* tests with 44 degrees of freedom.

The abilities of elevated levels of BMI and the SF sum to identify children with elevated levels of PBF_{DXA} were examined using the: (1) positive predictive value (eg, the proportion of children with an elevated BMI who had an elevated PBF_{DXA}); (2) the positive likelihood ratio (prevalence of an elevated BMI among children who had an elevated PBF_{DXA} divided by its prevalence among children without an elevated PBF_{DXA}); and (3) the kappa statistic,³⁴ a measure of chance-corrected agreement. The variance of the differences between kappa statistics also were calculated with the “withReplicates” function.³⁰ A similar process was used to compare the accuracies of BMI and the SF sum in the identification of children with low PBF_{DXA} levels.

Results

Various characteristics of the sample are shown among boys and girls in Table I (available at www.jpeds.com). The distributions of several variables were positively skewed, and we focus on the 15th, 50th, and 85th percentiles; in a normal distribution, the 15th and 85th percentiles would be about 1 SD from the mean. Overall, 14% of the children were non-Hispanic black and 11% were Mexican-American. The median BMI-for-age level of these children was 0.5 standard deviations greater than the median in the 1963-1994 CDC reference population, but BMI levels did not differ between boys and girls. In contrast, median levels of the SF sum and PBF_{DXA} were about 40% to 50% greater among girls than boys.

As assessed by the multiple R^2 values of various sex-specific models predicting PBF_{DXA} from race, age, and either BMI or the SF sum, the SF sum accounted for more of the variability in PBF_{DXA} than did BMI (Table II). Multiple R^2 values for models that included BMI-for-age, in addition to race and age, resulted in R^2 values of approximately 0.75 (second row), but the use of skinfold thicknesses yielded R^2 values of 0.86 (boys) and 0.81 (girls). As shown in Table II, even after we accounted for the information conveyed by BMI-for-age, the SF sum provided additional information on PBF_{DXA} among both boys (R^2 increases from 0.75 to 0.87), and girls (0.76 to 0.84).

Table III shows sex-specific correlations between BMI and the SF sum with PBF_{DXA}. (We accounted for age in these analyses by using the residuals of regression models that predicted levels of PBF_{DXA}, BMI and the SF sum from sex and age.) BMI was less strongly correlated with PBF_{DXA} than was the SF sum among both boys ($r = 0.83$ vs 0.91 , $P < .001$ for difference) and girls ($r = 0.84$ vs 0.89 , $P < .001$). These differences varied somewhat by age, with a smaller difference seen among 18- to 19-year-olds (boys: $r = 0.87$ vs 0.91 ; girls: $r = 0.89$ vs 0.88). However, the magnitudes of the associations with PBF_{DXA}, particularly for BMI, were substantially weaker among children whose body fatness was below the median for their sex and age. Among these thinner children, correlations ranged from $r = 0.49$ - 0.59 for BMI and from $r = 0.68$ - 0.75 for the SF sum. In contrast, among children with greater levels of body fatness, correlations ranged from 0.76 to 0.84. Furthermore, there was

little difference in the magnitudes of the associations with BMI and SF sum among girls who had levels of PBF_{DXA} above the median ($r = 0.77$ and $r = 0.79$).

Additional analyses of total fat mass (data not shown), rather than percent body fat, indicated that fat mass was more strongly correlated with BMI ($r = 0.94$ - 0.96) than with the SF sum ($r = 0.85$ - 0.87) among boys and girls. Furthermore, the SF sum did not convey additional information on fat mass if the BMI level was already known.

Table IV shows a cross-classification of the 5 PBF_{DXA} categories with those of BMI (top) and SF sum (bottom). As assessed by the proportion of children along the diagonal, the SF sum was a more accurate indicator of PBF_{DXA} than was BMI among both boys (60% vs 48% along the diagonal) and girls (59% vs 51%). (Similar differences were seen for the intraclass correlations across the 5 categories, with correlations of 0.65-0.66 for BMI among boys and girls vs 0.76-0.79 for the SF sum.) These differences, however, were largely due to the more accurate classification of thinner children by the SF sum. For example, whereas only 41% of boys in the lowest BMI category were also in the lowest PBF_{DXA} category, the comparable proportion (row percent) for the SF sum was 61%. In contrast, differences in the abilities of BMI and the SF sum to correctly identify children with elevated levels of PBF_{DXA} were much smaller among both boys (75%, BMI vs 79%, SF sum) and girls (76% for both BMI and SF sum). The classification of slightly elevated levels of PBF_{DXA} (category #4) was relatively poor for both BMI and SF sum.

Table V focuses on the screening performance of BMI and the SF sum in the identification of elevated (left columns) or low (right) levels of PBF_{DXA} based on the results of several 2×2 tables. In general, elevated levels of both BMI and SF sum were good indicators of elevated PBF_{DXA} levels (Table V) with positive likelihood ratios ranging from 14 to 18 among both boys and girls. (For example, boys with an elevated PBF_{DXA} were 14.3 times more likely to have an elevated BMI than were those with a lower PBF_{DXA} level.) As assessed by kappa statistics (H_0 : no difference in the abilities of BMI and SF sum to classify body fatness), an elevated SF sum was a better indicator of an elevated PBF_{DXA} level among boys (kappas of 0.70 vs 0.75, $P = .03$) but not girls (kappas of 0.71 for both measures). Kappa statistics varied somewhat across race-ethnicity groups, but the only statistically significant difference between BMI and the SF sum was among non-Hispanic black boys (0.68, BMI vs 0.77, SF sum). The positive predictive values, however, were lower among non-Hispanic black girls, with values of 0.61 (BMI) and 1.64 (SF sum) versus values of 0.79 or more among other girls.

In general, a low level of the SF sum was a much better indicator of a low PBF_{DXA} level than was a low BMI (Table V) among both boys (kappas of 0.28 vs 0.53) and girls (0.40 vs 0.55); $P < .001$ for both differences. Furthermore, BMI appeared to perform worse among non-Hispanic black and Mexican-American girls (kappas of 0.34 and 0.29) than among non-Hispanic white girls (kappa = 0.41). Only 28% of Mexican-American boys and girls with a low BMI had a low PBF_{DXA} .

Discussion

It is frequently assumed that thickness of skinfolds at various sites, typically expressed as a sum or as percent body fat based on published equations,⁶ is a better indicator of body fatness than is BMI.^{10,35} Our results indicate that although the SF sum (subscapular plus triceps) is more strongly associated with PBF_{DXA} than is BMI, with correlations of about 0.90 (SF sum) versus 0.84 (BMI), the importance of this difference may depend upon the objective of the study. We found that the largest difference between BMI and SF sum was in their abilities to correctly identify children with a low level of PBF_{DXA}: a low BMI-for-age could identify only 40% (boys) to 50% (girls) of these thin children, and the SF sum could identify 75%. In contrast, differences between BMI and SF sum were greatly reduced among children with elevated body fatness. Among boys, an elevated SF sum was slightly more accurate than was an elevated BMI (kappas of 0.75 and 0.70) in identifying those with a high PBF_{DXA} level. Among girls, however, BMI and the SF sum performed equally well (kappas of 0.71 for both measures).

It has previously been emphasized that although BMI is a useful surrogate for body adiposity among fatter children, it is “almost useless” in assessing the body fatness of normal-weight children.^{6,36} However, even when the goal is to identify children with low levels of body fatness, BMI may be of some use. We found, for example, that boys with a low (<16th percentile) PBF_{DXA} were 3.3 times more likely to have a low BMI than were other boys. This, however, should be contrasted with our findings that boys with an elevated PBF_{DXA} were 14 times more likely to have an elevated BMI than were other boys. It is likely that these differing associations between BMI and body fatness may account for the inter-study differences that have been observed between BMI and PBF_{DXA}.^{7,8,36} For example, Kerruish et al⁸ reported a correlation of only $r = 0.46$ between BMI and PBF_{DXA}, but they focused on girls with anorexia nervosa (mean PBF_{DXA}, 14%). In contrast, we and others³⁷ have found multiple R^2 values of 0.75 or greater for the prediction of PBF_{DXA} from BMI (or 1/BMI), race, sex, and age in more representative samples.

It should also be realized that although approximately 25% of the children with an elevated BMI did not have an elevated PBF_{DXA} level, most (~80%) of these misclassified children had a PBF_{DXA} level considered to be slightly elevated, corresponding to BMIs in the overweight category (between the CDC 85th and 94th percentiles). Of the children with a BMI-for-age 95th percentile, only 5% had a PBF_{DXA} corresponding to levels in the moderate or normal range. However, there was a wide range of levels of body fatness among children who had slightly lower, but still high, BMIs (85th-94th percentiles). Among these overweight children, approximately 40%-50% had a body fatness that was in the expected range, but about 20% had a greater-than-expected PBF_{DXA} level and another 30% had a PBF_{DXA} corresponding to levels between the CDC 50th and 84th percentiles.

We also found that the screening ability of BMI for elevated body fatness varied across race-ethnic groups, with BMI having a lower positive predictive value among non-Hispanic black children. Whereas about 80% of non-Hispanic white and Mexican-American children with a BMI CDC 95th percentile had an elevated PBF_{DXA}, these predictive values were 61% (girls) and 65% (boys) among non-Hispanic black children. This difference is comparable

with that observed by Flegal et al³⁸ despite differences in the cut-points used for body fatness. The authors of previous studies^{37,39} have reported that at similar levels of age and BMI, the mean PBF_{DXA} of non-Hispanic black children is approximately 2%-3% lower than that of non-Hispanic white children. If the 95th percentile of BMI is used to identify children who have excess body fatness, these differences could result in a large number of false positives among non-Hispanic black children.

Although we found that the use of skinfold thickness significantly improved the identification of boys (but not girls) who had a high PBF_{DXA}, it is not certain if the higher positive predictive value (75%, BMI vs 79%, SF sum) would have a substantial impact on screening and interventions. The errors associated with skinfold thicknesses measurements can be large,^{9,40} particularly among inexperienced observers, and these measurements, some of which require disrobing, are generally more intrusive than are those for weight and height. It is also known that the accuracy of skinfolds in predicting body fatness varies according to the selected sites and the equation used. For example, Bray et al⁶ found that most skinfold thickness equations were better predictors of body fatness (determined from a 4-compartment model) than was BMI, but that one resulted in very poor prediction (multiple R² of 0.51). Furthermore, 29% of obese children in the current study had at least one skinfold that could not be measured. We imputed these missing values, but in practice, it may be difficult to use skinfold thicknesses to track the progress of extremely obese persons over time.

There are several limitations of the current analyses that should be considered. DXA estimates of body fatness are known to vary across machines, manufacturers, and software versions.^{6,41-46} Because the Hologic QDR-4500-A (Hologic Inc, Bedford, Massachusetts) in the current study has been found⁴⁷⁻⁴⁹ to overestimate lean mass among adults, the NCHS decreased the recorded DXA lean mass values by 5% and added an equivalent weight to each subject's fat mass.⁵⁰ This adjustment may be the reason for the greater PBF_{DXA} values in this sample than in other studies of children who have similar BMI levels. However, if this proportional adjustment was valid across the range of body fatness in the current study, it is unlikely to have substantially influenced our findings concerning the screening performances of BMI and skinfold thicknesses. Relying on a fixed cutpoint (eg, 30%) for high body fatness would be influenced by a bias, but we used cut-points that resulted in equivalent numbers of children with elevated levels BMI, SF sum, and PBF_{DXA}. A somewhat-similar approach has been taken by other investigators^{51,52} in the construction of body fat reference curves. It should also be realized that we were primarily interested in PBF_{DXA}, rather than fat mass, as a measure of body fatness. Additional analyses of total fat mass, however, indicated that this characteristic was: (1) more strongly correlated with BMI than with skinfold thicknesses; and (2) that skinfolds provided no additional information beyond that conveyed by BMI.

The results of the current study indicate that sum of the thicknesses of triceps and subscapular skinfolds is more strongly associated with DXA-calculated body fatness of children than is BMI. This stronger association, however, does not necessarily indicate that skinfolds rather than BMI should always be used in the classification of body fatness of children. If the objective is to identify children with low body fatness, which may be

associated with slower bone development,⁵³ skinfold thicknesses are superior. However, if the goal is to identify girls with elevated levels of PBF_{DXA} , who are at increased risk for obesity-related complications, skinfold thicknesses and BMI perform equally well. If the goal is to identify boys with elevated levels of percent body fat, skinfold thickness provides some additional information concerning health risks and correctly identifies an additional 4 of 100 boys who have an elevated PBF_{DXA} , but the importance of this additional information is uncertain. The relatively small improvement obtained with the measurement of skinfold thicknesses should be balanced with the additional training needed to standardize these measurements and the difficulties in obtaining these measurements among obese children.

Glossary

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| BMI | Body mass index |
| CDC | Centers for Disease Control and Prevention |
| DXA | Dual-energy X-ray absorptiometry |
| NCHS | National Center for Health Statistics |
| NHANES | National Health and Nutrition Examination Survey |
| PBF_{DXA} | Percent body fat determined by dual-energy x-ray absorptiometry |
| SF sum | Sum of the triceps and subscapular skinfold thicknesses |

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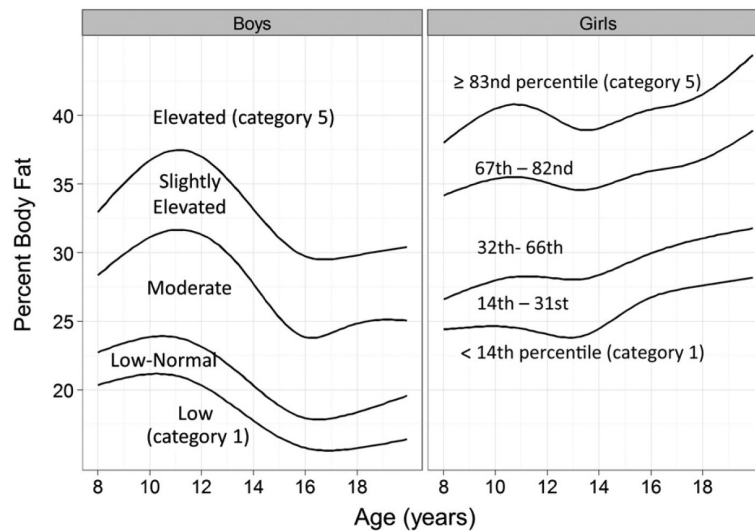


Figure.

Cut-points for the 5 categories of PBF_{DXA} by sex (boys, left panel; girls, right panel) and age. Within each sex, the 4 curves were estimated with the use of quantile regression (see Methods)²⁹ and represent the boundaries of the 5 body fatness categories. These cut-points were chosen so that the number of children in each PBF_{DXA} category would be similar to the number of children in the corresponding BMI-for-age category. The names of these categories are shown for boys, and the percentile cut points are shown for girls. The cut-points for the 5 categories of body fatness among boys were the 17th, 35th, 66th, and 82nd percentiles.

Table I

Descriptive characteristics of 8- to 19-year-old subject by sex: NHANES 1999-2004

| Characteristic | Boys | Girls |
|---|-------------------------------|-------------------|
| n (unweighted) | 4493 | 2872* |
| Age, years | 13.9 (9.8, 18.2) [†] | 13.8 (9.9, 17.9) |
| BMI, kg/m ² | 20.8 (16.8, 26.9) | 21.1 (17.0, 27.3) |
| BMI-for-age [‡] | 0.5 (-0.8, 1.8) | 0.5 (-0.6, 1.7) |
| % obese [§] | 18 ± 1% | 17 ± 1% |
| % overweight or obese [§] | 34 ± 1% | 34 ± 1% |
| % non-Hispanic white | 61% | 62% |
| % non-Hispanic black | 15% | 14% |
| % Mexican-American | 11% | 11% |
| Skinfold thicknesses, mm | | |
| Triceps | 11.2 (7.0, 22.1) | 17.2 (10.4, 27.3) |
| Subscapular | 9.1 (5.7, 19.4) | 12.4 (7.2, 23.7) |
| SF sum | 20.2 (13.4, 41.5) | 30.2 (18, 50.1) |
| DXA-calculated percent body fat, PBF _{DXA} | 23.6 (17.2, 34.6) | 32.3 (25.6, 41.1) |

* Data for girls are from 2001-2004 rather than 1999-2004.

[†] With the exception of the sample size, values are medians for continuous variables or percents for categorical variables. The distribution of each characteristic is indicated by the 15th and 85th percentile (for continuous variables) or by the SE (for categorical variables). The 15th and 85th percentiles are about 1 SD from the mean for a normally distributed variable.

[‡] Z-score (SD score) of subjects in the current study relative to those who were included in the 2000 CDC growth charts.²⁴

[§] Obesity is defined as a BMI-for-age ≥ CDC 95th percentile, and overweight is a BMI-for-age between the 85th and 94th percentiles.

Table IIMultiple R² values for various regression models predicting PBF_{DXA}

| Predictors* | Boys | Girls |
|----------------------------------|------|-------|
| Race, age (baseline) | 0.11 | 0.04 |
| + BMI | 0.75 | 0.76 |
| + triceps skinfold thickness | 0.85 | 0.77 |
| + subscapular skinfold thickness | 0.78 | 0.73 |
| + SF sum | 0.86 | 0.81 |
| Race, age, and BMI (baseline) | 0.75 | 0.76 |
| + triceps skinfold thickness | 0.86 | 0.83 |
| + subscapular skinfold thickness | 0.81 | 0.80 |
| + SF sum | 0.87 | 0.84 |

* All models include race and age (top rows) or race, age, and BMI (bottom rows). Age, BMI, and the skinfold thicknesses were modeled with the use of restricted cubic splines.

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Table IIIStratified correlations between PBF_{DXA}, BMI, and SF sum

| | Boys | | Girls | |
|---------------------------|-------------------|-------------------|-------------------|-------------------|
| | BMI | SF sum | BMI | SF sum |
| Overall | 0.83* | 0.91* | 0.84* | 0.89* |
| Age group, years | | | | |
| 8-11.9 | 0.86* | 0.92* | 0.84* | 0.90* |
| 12-14.9 | 0.78* | 0.91* | 0.86 [†] | 0.90 [†] |
| 15-17.9 | 0.84* | 0.90* | 0.83 [†] | 0.87 [†] |
| 18-19.9 | 0.87 [†] | 0.91 [†] | 0.89 | 0.88 |
| Race | | | | |
| Non-Hispanic whites | 0.83* | 0.91* | 0.85* | 0.89* |
| Non-Hispanic blacks | 0.84* | 0.92* | 0.85* | 0.91* |
| Mexican-Americans | 0.86* | 0.93* | 0.82* | 0.90* |
| Body fatness [‡] | | | | |
| <Median | 0.49* | 0.68* | 0.59* | 0.75* |
| Median | 0.76* | 0.84* | 0.77 | 0.79 |

P values assess whether the correlation between PBF_{DXA} and BMI is equal to the correlation between PBF_{DXA} and the SF sum.

* $P < .001$;

[†] $P < .05$. In a random sample of 3500, a correlation of 0.055 would be statistically significant ($H_0: r = 0$) at the 0.001 level.

[‡] Body fatness was categorized into 2 groups on the basis of the median PBF_{DXA} for a child's sex, age (year), and imputation number.

Table IV

Cross-classification of categories* of PBF_{DXA} with either BMI-for-age or SF sum

| BMI-for-age category | Boys: PBF _{DXA} category | | | | | Girls: PBF _{DXA} category | | | | |
|-------------------------|-----------------------------------|-----|-----|-----|--------------|------------------------------------|-----|-----|-----|--------------|
| | 1 (low) | 2 | 3 | 4 | 5 (elevated) | 1 (low) | 2 | 3 | 4 | 5 (elevated) |
| <25th percentile | 41% [†] | 35% | 23% | 2% | 0 | 48% | 32% | 19% | 1% | 0 |
| 25th-49th percentile | 32% | 28% | 37% | 3% | 0 | 29% | 30% | 39% | 1% | 0 |
| 50th-84th percentile | 14% | 20% | 50% | 14% | 2% | 7% | 21% | 54% | 16% | 2% |
| 85th-94th percentile | 0 | 4% | 28% | 46% | 22% | 1% | 1% | 35% | 43% | 20% |
| 95th percentile (obese) | 0 | 0 | 4% | 21% | 75% | 0 | 0 | 4% | 20% | 76% |
| SF sum category | | | | | | | | | | |
| 1 (low) | 61% | 28% | 1% | 0 | 0 | 62% | 32% | 6% | 0 | 0 |
| 2 (low-normal) | 28% | 42% | 30% | 0 | 0 | 27% | 40% | 33% | 0 | 0 |
| 3 (moderate) | 5% | 17% | 63% | 13% | 1% | 2% | 16% | 65% | 16% | 2% |
| 4 (slightly elevated) | 0 | 1% | 25% | 53% | 21% | 0 | 1% | 31% | 46% | 22% |
| 5 (elevated) | 0 | 0 | 1% | 20% | 79% | 0 | 0 | 4% | 20% | 76% |

* Categories of PBF_{DXA} and the SF sum were constructed so the that number of children in each category would be the same as the number of children in each BMI-for-age category.

[†] Values are row percents, representing the proportion of children in the specified BMI or SF sum category who are in specified category of body fatness.

Table V
Screening characteristics of BMI and SF sum for elevated and low levels of PBF_{DXA}

| Sex | Race-ethnicity | Measure | Children with elevated PBF _{DXA} | | | | Thin children (Low PBF _{DXA}) | | | |
|-------|---------------------|---------|---|---------------------------|---------------------------|-------------------|---|---------------------------|---------------------------|--------------------|
| | | | Prevalence | Positive predictive value | Positive likelihood ratio | Kappa statistic | Prevalence | Positive predictive value | Positive likelihood ratio | Kappa statistic |
| Boys | Overall | BMI | 0.18 [†] | 0.75 | 14.3 | 0.70 [†] | 0.17 | 0.41 | 3.3 | 0.28 [§] |
| | | SF sum | 0.18 | 0.79 | 18.0 | 0.75 [†] | 0.17 | 0.61 | 7.7 | 0.53 [§] |
| Girls | Overall | BMI | 0.17 | 0.76 | 15.3 | 0.71 | 0.14 | 0.48 | 5.7 | 0.40 [§] |
| | | SF sum | 0.17 | 0.77 | 15.7 | 0.71 | 0.14 | 0.62 | 10.0 | 0.55 [§] |
| Boys | Non-Hispanic Whites | BMI | 0.16 | 0.77 | 16.8 | 0.71 | 0.16 | 0.37 | 3.6 | 0.29 [§] |
| | | SF sum | 0.16 | 0.80 | 12.0 | 0.76 | 0.17 | 0.54 | 7.1 | 0.51 [§] |
| | Non-Hispanic Blacks | BMI | 0.19 | 0.65 | 11.0 | 0.68 [‡] | 0.17 | 0.71 | 4.4 | 0.29 [§] |
| | | SF sum | 0.17 | 0.75 | 17.4 | 0.77 [‡] | 0.24 | 0.86 | 11.2 | 0.56 [§] |
| | Mexican-Americans | BMI | 0.24 | 0.82 | 12.7 | 0.72 | 0.16 | 0.28 | 3.7 | 0.26 [§] |
| | | SF sum | 0.23 | 0.85 | 16.5 | 0.73 | 0.13 | 0.50 | 9.5 | 0.52 ^{§§} |
| Girls | Non-Hispanic Whites | BMI | 0.16 | 0.80 | 19.9 | 0.73 | 0.15 | 0.47 | 5.6 | 0.41 [‡] |
| | | SF sum | 0.16 | 0.79 | 18.2 | 0.73 | 0.15 | 0.59 | 9.0 | 0.55 [‡] |
| | Non-Hispanic Blacks | BMI | 0.26 | 0.61 | 7.5 | 0.66 | 0.10 | 0.67 | 7.6 | 0.34 [§] |
| | | SF sum | 0.24 | 0.64 | 8.9 | 0.69 | 0.13 | 0.81 | 16.0 | 0.56 [§] |
| | Mexican-Americans | BMI | 0.17 | 0.81 | 17.8 | 0.70 | 0.13 | 0.28 | 4.8 | 0.29 [§] |
| | | SF sum | 0.16 | 0.87 | 27.3 | 0.73 | 0.12 | 0.45 | 9.8 | 0.51 [§] |

*Values represent various screening characteristics of BMI or the SF sum in the identification of children with elevated levels of PBF_{DXA} (left side) or low levels of PBF_{DXA} (right side). For example, a cross-classification of elevated levels of BMI and PBF_{DXA} among boys (first row, top left) indicated that the positive predictive value of an elevated BMI was 0.75 with a kappa statistic of 0.70.

P-values are based on the differences between kappa statistics for BMI and the SF sum, and were calculated for each imputation using jackknife replicate weights:

[†] P < .05;

[‡] P < .01;

$P < .001$ [§]

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