Influenza-associated mortality in Thailand, 2006–2011

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Background Influenza-associated mortality in subtropical or tropical regions, particularly in developing countries, remains poorly quantified and often underestimated.

We analyzed data in Thailand, a middle-income tropical country with good vital statistics and influenza surveillance data.

Methods We obtained weekly mortality data for all-cause and three underlying causes of death (circulatory and respiratory diseases, and pneumonia and influenza), and weekly influenza virus data, from 2006 to 2011. A negative binomial regression model was used to estimate deaths attributable to influenza in two age groups (<65 and ≥65 years) by incorporating influenza viral data as covariates in the model.

Results From 2006 to 2011, the average annual influenza-associated mortality per 100 000 persons was $4\cdot0$ (95% CI: -18 to 26). Eighty-three percent of influenza-associated deaths occurred among persons aged >65 years. The average annual rate of influenza-associated deaths was $0\cdot7$ (95% CI: $-8\cdot2$ to 10) per 100 000 population for person aged <65 years and 42 (95% CI: -137 to 216) for person aged ≥ 65 years.

Discussion In Thailand, estimated excess mortality associated with influenza was considerable even during non-pandemic years. These data provide support for Thailand's seasonal influenza vaccination campaign. Continued monitoring of mortality data is important to assess impact.

Keywords Influenza, mortality, Thailand, tropical.

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Introduction

The global burden of influenza is unknown but thought to be considerable. In 2008, the global estimate of influenza-associated severe acute lower respiratory illness in children <5 years was 1 million cases, and the World Health Organization (WHO) suggests that 3–5 million severe cases occur in persons of all ages each. In temperate regions, influenza usually causes annual outbreaks during the winter season (November–March in the Northern Hemisphere and June–September in the Southern Hemisphere). In temperate regions, influenza-associated mortality in subtropical or tropical regions, particularly in developing countries, remains poorly quantified and often underestimated. The pattern of influenza epidemics in tropical regions is less distinct and more diffuse. Seasonal influenza epidemics can sometimes occur twice or more throughout a year.

The exact burden of influenza morality is difficult to estimate, and the challenges in counting influenza-associated deaths include the following: testing of hospital patients for influenza (particularly in elderly persons) is uncommon, influenza is rarely specifically recorded on death certificates, and many deaths that may be causally related to influenza occur after virus can be detected. Therefore, estimation of influenza-associated deaths or hospitalizations often relies on statistical modeling rather than on direct measurement. Various methodological approaches have been used to estimate the excess deaths associated with the circulation of influenza virus in temperate region.⁶⁻¹⁰ However, the best approach for modeling influenza in tropical regions is unclear. Further, although there are results for wealthy, tropical countries (e.g., Singapore), 11 the findings may be different from Thailand, where the severity of infections may be compounded by a higher prevalence of underlying illness or other factors.. In this study, we used a regression model to examine the impact of influenza, by virus type and subtype, on deaths in Thailand, while adjusting for potential confounding effects by other co-circulating influenza virus subtypes.

Estimating the burden of influenza mortality is important to help guide vaccination programs, evaluate the use of diagnostic tests and antiviral drugs, and plan for seasonal epidemics and future pandemics. In this analysis, we applied a negative binomial model to the weekly counts of deaths and viral data to explore the seasonal effect of influenza on mortality and provided estimates of excess mortality associated in Thailand by death category, age group, and influenza subtypes for the years 2006 through 2011.

Methods

Mortality data

Weekly electronic mortality data for years 2006 through 2011 were obtained from the Bureau of Policy and Strategy, Ministry of Public Health (MoPH), Thailand. Deaths were categorized into three groups based on codes from the International Classification of Diseases, Tenth Revision [ICD-10]: circulatory diseases (*ICD-10* codes I00-99), respiratory diseases (ICD-10 codes J00-99), and pneumonia and influenza (P&I) (ICD-10 codes J10-18). Other variables included age at death and location of death (hospital or community). We defined non-pandemic years as every year except 2009.

Viral surveillance data

In 2004, the National Institute of Health, MoPH launched the national influenza sentinel surveillance system to monitor virus circulation in patients with influenza-like illness presenting at outpatient clinics in 11 sites throughout Thailand. 12 Between 2004 and 2009, each site was first instructed to enroll a convenience sample of up to five patients per week with ILI for a total of 20 patients per month. In September 2009, the sample size was increased to 10 patients per week (five from children aged <15 years and five from persons aged \geq 15 years). These data, collected systematically throughout the year, represent an unbiased sample of the timing of influenza activity and are appropriate for analyzing seasonal trends. All specimens were tested by reserve-transcription polymerase chain reaction for influenza viruses. We obtained the weekly numbers of influenza viruses (by type and subtype) and the total number of specimens tested, from the National Institute of Health, and calculated weekly positivity using weekly number of specimens as the denominator.

Reapportioning mortality data

Of all-cause deaths among person aged <65 and ≥65 years, ill-defined deaths (i.e., no cause listed) accounted for 25% (58% of these occurred in the community) and 54% (87% occurred in the community), respectively. To account for the high proportion of ill-defined cause of death, we reapportioned weekly ill-defined deaths by age (<65 and ≥65 years), and location (hospital and community) to the various categories (P&I, respiratory, circulatory). We first excluded the ill-defined deaths and calculated the proportion of deaths

due to P&I, respiratory and circulatory among deaths recorded in-hospital by age (<65 and ≥65 years) under the assumption that the coding for an in-hospital death was more accurate than for a community death. This assumption is based on the knowledge that deaths occurring outside the hospital are recorded by non-medical, civil registrars and based on lay reports from relatives. ¹³ We then reapportioned all ill-defined deaths (including both those that occurred inhospital and those in the community) to the various categories (P&I, respiratory, circulatory) so that the final proportion in each category were equivalent to those found in the in-hospital deaths.

We created two separated databases, one without reapportioning the deaths and the other with reapportioning.

Climate data

Meteorological parameters including weekly mean temperature and absolute humidity were obtained from the Thai Meteorological Department.

Negative binomial regression model

A negative binomial regression model was used to estimate deaths attributable to influenza in two age groups (<65 and ≥65 years) by incorporating influenza viral data as covariates in the model. To assume an additive relationship between the exposure to influenza and resulting mortality, we used an identity link function in the models. The models included independent variables that comprised weekly percentage of specimens testing of confirmed infection with seasonal influenza A, seasonal influenza B, and influenza A (H1N1) pdm09 during a given week. We tested various lags (no lag, 1 week, and 2 weeks) to account for delays between infection and death. To account for the baseline, the models include either harmonic terms for annual [sine (θ) and cosine (θ) , where $\theta = 2 \pi \text{week}/52.25$ and semiannual [sine (2 θ) and cosine (2θ)] periods or climate factors [weekly mean temperature and absolute humidity, using linear terms or nonlinear terms (natural cubic spline function)]. Selection of the most statistically meaningful proxy for influenza activity was based on fitting 13 models to ≥65 years respiratory deaths and comparing Akaike information criterion (AIC) values. We obtained the smallest AIC values for a model with a 2-week lag and smoothing spline of temperature and humidity and report those results. A detailed description of the model fitting procedure is provided in Table S1.

In this model,

$$Y_{t} = \alpha \exp{\{\beta_{0} + \beta_{1} * \text{Time} + \beta_{2} * \text{Time}^{2} + \beta_{3} * \text{NCS (Temperature}) + \beta_{4} * \text{NCS (Humidity)} + \beta_{5} * [A(H1N1)] + \beta_{6} * [A(H3N2)] + \beta_{7} * [B] + \beta_{8} * [Pdm09H1N1]^{4},$$

where Y_t is the number of deaths in week t and α is the offset term and is equal to the log of the population size. β_0 represented the intercept, β_1 and β_2 represented coefficients associated with secular trends, β_3 represented coefficients associated with a natural cubic spline (NCS) of weekly mean temperature, β_4 represented coefficients associated with a NCS of absolute humidity, β_5 - β_9 represented coefficients associated with the percentages of specimens testing positive for each influenza virus type and subtype during a given week. All viral terms had a 2-week lag. The models were run separately by death category (P&I, respiratory, and circulatory), age group (<65 and ≥65 years), and location of death (hospital or community, only for the data without reapportioned deaths). The analysis was conducted both ways, with and without reapportioning these deaths with the latter included as supplementary tables and figures.

To estimate excess mortality associated with influenza, we used the following procedure:

- 1. Calculate the expected mortality with the full model (E0).
- 2. Calculate expected mortality with the same fitted model, but the influenza terms were set as zero (E1).

E0-E1 was the estimate of mortality due to influenza.

Confidence intervals were calculated based on the standard errors of model coefficients. We note that these confidence intervals provide only a minimum estimate of uncertainty.

The statistical analyses were carried out using SAS version 9.3 (SAS Institute, Cary, NC, USA).

Sensitivity analysis

As a sensitivity analysis, we modeled all-cause deaths. We compared the excess deaths obtained using the all-cause deaths as outcome in the final models with estimates using respiratory and circulatory deaths as outcomes in the same models.

Results

Mortality data

From January 2006 to December 2011, an annual mean of 399 853 deaths (range, 389 696 in 2006 to 413 209 in 2011)

occurred in Thailand (42% occurred in hospital). An average of 13 554 (3.4%) underlying P&I deaths, 27 268 (6.8%) underlying respiratory deaths, and 37 112 (9.3%) underlying circulatory deaths occurred each year. Table 1 shows the annual mean number and proportion of deaths by cause of death. Forty-one percent of total deaths were coded as cause ill-defined and redistributed to the various categories (P&I, respiratory, circulatory) by location and age group.

Viral surveillance data

During the 6-year period, 21 560 specimens were tested for influenza viruses. There were 2961 (14%) positive results for influenza viruses. The annual mean number of tests positive was 14% (range 10–18%) for influenza A viruses, and 6.5% (range $4\cdot1-10\%$) for influenza B viruses (Table S2).

Estimates of influenza-associated deaths

For all ages combined, there was year-to-year variability in the annual number of influenza-associated deaths; the mean for all years was 2511 (4.0 per 100 000) (Table 2). The average influenza-associated mortality rate for non-pandemic years was 4.3 per 100 000, and for the pandemic year, 2009, it was 2.4 per 100 000. The rates in pandemic year for both age groups were lower compared with the other years (<65 years: 0.6 versus 0.7 per 100 000 and ≥65 years: 24 versus 26 per 100 000). The annual death rate attributable to influenza was the highest in the year 2008 (5.6). Eighty-three percent (2094/2511) of the average annual estimated influenza-associated deaths occurred among person age ≥65 years. The average annual rate of influenza-associated deaths was 0.7 per 100 000 population for person aged <65 years and 42 for person aged ≥65 years. Among adults aged <65 years and ≥65 years, influenza A (H3N2) and B were associated with the most deaths. The average annual number of influenzaassociated deaths was 1401 (2.2 per 100 000) for underlying P&I causes, 2751 (4.3 per 100 000) for underlying respiratory causes, and 529 (0.8 per 100 000) for underlying circulatory causes (Table 3). Deaths among persons aged ≥65 years

Table 1. Average annual number and proportion of deaths by cause of death, location of death, and age group in Thailand, 2006–2011

| | | In-hospital d | In-hospital deaths (42%) | | Community deaths (58%) | | | |
|--|--------------------|--------------------|--------------------------|--------------|------------------------|-----------------|--------------|--|
| Cause of death | Total <i>n</i> (%) | <65 years (55%) | ≥65 years (45%) | All ages | <65 years (40%) | ≥65 years (60%) | All ages | |
| Annual average number of deaths Proportion of deaths | 399 853 | 92 178 | 75 962 | 168 139 | 92 519 | 139 195 | 231 714 | |
| Respiratory disease | 27 268 (6.8) | 7481 (8.1) | 12 361 (16) | 19 842 (12) | 3575 (3.9) | 3851 (2.8) | 7426 (3.2) | |
| Pneumonia and influenza | 13 554 (3.4) | 4583 (5.0) | 7279 (9.6) | 11 861 (7.1) | 935 (1.0) | 757 (0.5) | 1693 (0.7) | |
| Circulatory disease | 37 113 (9.3) | 11 118 (12) | 14 482 (19) | 25 600 (15) | 5634 (6.1) | 5879 (4-2) | 11 513 (5.0) | |
| Ill-defined | 163 474 (41) | 19 513 (21) | 14 615 (19) | 34 128 (20) | 27 194 (29) | 102 152 (73) | 129 345 (56) | |

Table 2. Estimated annual influenza-associated deaths in Thailand, 2006-2011 (with the apportioned ill-defined deaths)

| Column C | | A(H1N1) | _ | | A(H3N2) | | | Pdm09 H1N1 | H1N1 | | 8 | | | Total | | | Rate | Rate per 100 000 | 000 |
|--|-----------------------|---------|-------|------|---------|-------|------|------------|-------|------|-------|-------|-------|-------|---------|--------|------|------------------|------|
| 625 -3222 4151 228 -3467 3922 0 0 0 647 -3060 4228 1500 -10 351 12 993 2.4 -17 622 -3206 4143 1563 -1500 545 0 0 0 928 -2565 4470 292 -8808 14 169 47 -13 642 -3206 4143 1563 -1215 5240 0 0 0 1499 -2205 5470 592 -8808 14 169 47 -12 642 -3209 4376 841 -3611 427 1078 -3219 512 529 -14 6800 19275 3.5 -23 -23 -446 4246 2273 -1965 6250 133 -4186 299 429 985 -3281 4895 1524 -11 522 17 280 2.4 -22 -21 521 17 580 2.4 -22 -21 521 17 580 2.4 -22 -21 521 17 580 2.4 -22 -21 521 17 580 2.4 -22 -21 521 17 580 2.4 -22 -21 521 17 580 2.4 -22 -21 521 17 580 2.4 -22 -21 521 17 580 2.4 -22 -21 521 17 580 2.4 -22 -21 521 17 580 2.4 -40 -10 6.6 -10 6.0 6.0 -17 7 18 -12 6.0 -17 7 18 -12 6.0 -12 7 18 -12 7 | Year | Death | 95%CI | | Death | 12%56 | | Death | 12%26 | | Death | 12%26 | | Death | 12%56 | | Rate | 95%CI | |
| 625 - 3222 | All age | | | | | | | | | | | | | | | | | | |
| 215 - 3400 3710 1866 - 1690 5426 0 0 0 0 928 - 2265 4470 292 - 8801 14 169 47 - 113 261 - 3759 4137 633 - 2115 540 0 0 0 1499 - 2205 6 150 3534 - 810 15 107 56 - 113 261 - 3759 4137 633 - 3299 4576 841 - 3601 4278 356 - 3293 4289 1524 - 41252 17 20 262 - 3266 4271 776 - 3491 5045 955 - 3884 4647 1078 - 3319 5312 513 513 51 291 - 3865 4110 1223 - 2671 5122 643** - 3891 4408 916 - 3200 4 794 5511 292 - 3670 4104 1341 - 2546 5231 544 - 4037 4473 1027 - 2884 4895 708 - 11060 16 365 43 - 117 203 - 1257 1477 158 | 2006 | 625 | -3232 | 4151 | 228 | -3467 | 3922 | 0 | 0 | 0 | 647 | -3060 | 4328 | 1500 | -10351 | 12 993 | | -17 | 2 |
| 642 - 3206 4143 1563 - 2115 5240 0 0 0 1499 - 2206 5150 3534 - 8101 15 107 5 6 - 13 2 1 280 4 426 271 776 - 3491 5045 955 - 3884 4647 1078 231 2239 4789 512 2 229 17 800 24 - 22 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 | 2007 | 215 | -3400 | 3710 | 1866 | -1690 | 5426 | 0 | 0 | 0 | 928 | -2655 | 4470 | 2952 | -8308 | 14 169 | | -13 | 7 |
| 261 -3759 4137 633 -3299 4576 841 -3601 4278 356 -3593 4289 1524 -14 251 17 280 24 -22 3 | 2008 | 642 | -3206 | 4143 | 1563 | -2115 | 5240 | 0 | 0 | 0 | 1499 | -2206 | 5150 | 3534 | -8101 | 15 107 | | -13 | 7 |
| 3 | 2009 | 261 | -3759 | 4137 | 633 | -3299 | 4576 | 841 | -3601 | 4278 | 356 | -3593 | 4289 | 1524 | -14 252 | 17 280 | | -22 | 7 |
| 291 -3685 410 1223 -1965 6520 133 -4189 4299 985 -3281 5216 3315 -13 681 5.2 281 5.2 -2.1 297 -3670 4104 1341 -2546 5231 5.44 -4037 4473 1027 -2884 4895 2708 -110 60 16 365 4.3 -17 192 -1251 1471 25 -1337 1389 0 0 0 118 -1250 1474 335 -4430 4926 0.6 -7.7 14 200 -1222 1445 180 -1158 1514 0 0 0 0 172 -1143 1460 395 -4036 5132 0.8 -6.9 14 81 -1395 1487 72 -1361 1509 678 -1009 1860 66 -1377 1500 390 -5142 6936 0.7 -7.0 14 92 -1351 1515 259 -1250 1776 109 -1443 1581 179 -1341 1685 471 -4773 5798 0.7 -8.2 93 -1365 1460 140 -1275 1515 259 -1250 1776 199 -1443 1581 159 159 1460 395 -4036 5132 0.8 -9.4 94 -1365 1460 140 -1275 1515 259 -1250 1776 199 -1361 1509 281 150 281 281 150 281 281 150 281 281 281 281 281 281 281 281 281 281 | 2010 | m | -4266 | 4271 | 276 | -3491 | 5045 | 955 | -3884 | 4647 | 1078 | -3219 | 5312 | 2239 | -14860 | 19 275 | | -23 | M |
| 991 - 3685 4110 1223 - 2671 5122 643** - 3891 4408 916 -3002 4794 2511 - 11 592 16 518 40 - 18 297 - 3670 4104 1341 - 2246 5231 544 - 4037 4473 1027 - 2884 4895 2708 - 11 060 16 365 43 - 17 298 - 3670 4104 1341 - 2246 5231 544 - 4037 4473 1027 - 2884 4895 2708 - 11 060 16 365 43 - 17 200 - 1222 1445 180 - 1138 1514 0 0 0 172 - 1143 1460 395 - 40450 4867 0.7 200 - 1222 1445 180 - 1158 1514 0 0 0 175 - 1143 1460 390 - 485 - 4030 5132 0.8 200 - 1222 1445 180 - 1158 1514 0 0 0 175 - 1143 1460 390 - 5142 6350 0.6 -88 200 - 1367 1515 259 - 1250 1776 1599 485 - 4030 5132 0.8 -6.9 200 - 1367 1518 299 - 1250 1776 199 - 1443 1581 179 1824 168 200 - 1367 1460 159 1460 159 1460 159 1460 159 1480 1480 1490 1490 1490 1490 1490 1490 1490 149 | 2011 | 0 | -4246 | 4246 | 2273 | -1965 | 6520 | 133 | -4189 | 4299 | 985 | -3281 | 5216 | 3315 | -13 681 | 20 281 | | -21 | Μ |
| years)* 192 | Average | 291 | -3685 | 4110 | 1223 | -2671 | 5122 | 643** | -3891 | 4408 | 916 | -3002 | 4794 | 2511 | -11 592 | 16 518 | | -18 | 7 |
| years)* 192 | Average | 297 | -3670 | 4104 | 1341 | -2546 | 5231 | 544 | -4037 | 4473 | 1027 | -2884 | 4895 | 2708 | -11 060 | 16 365 | 4.3 | -17 | 7 |
| 192 -1.251 1471 25 -1.337 1389 0 0 0 118 -1.250 1474 335 -4430 4926 0.6 -7.7 1.8 1.8 1.8 1.2 1.4 1.8 1.2 1.4 1.8 1.2 1.4 1.8 1.2 1.4 1.8 1.4 1 | (non-pandemic years)* | | | | | | | | | | | | | | | | | | |
| Fig. 1262 1471 1272 1472 1273 12 | ged ≥ ob years | 192 | _1251 | 1771 | 75 | _1337 | 1380 | C | C | C | 212 | _1250 | 1/7// | 335 | 7/130 | 7976 | 9.0 | 7.7 | α |
| Column C | 2002 | 92- | 1251 | 1337 | 217 | 2001 | 1512 | o c | o c | o c | 17.2 | 11/3 | 1760 | 207 | 4056 | 7867 | 0 0 | 7.0 | ο α |
| years)* 1 | 0008 | 200 | -1222 | 1445 | 180 | 1158 | 1514 | o C | 0 0 | 0 0 | 275 | -1076 | 1599 | 485 | -4030 | 5137 | 0 0 | 0.6 | 0 00 |
| 1 - 1547 1548 88 - 1458 1639 769 - 1065 2030 198 - 1362 1732 483 - 5432 6949 0.8 - 9-3 9 -1515 1515 259 - 1250 1776 109 - 1443 1581 179 - 1341 1685 471 - 5549 6557 0.8 - 9-4 90 -1515 1515 259 - 1256 1776 109 - 1443 1581 179 - 1341 1685 471 - 5549 6557 0.8 - 9-4 90 -1365 1466 140 - 1275 1557 519** - 1172 1824 168 - 1258 1575 417 - 4773 5798 0.7 - 8-2 92 -1359 1462 153 - 1258 1566 293 - 836 1204 188 - 1234 1590 434 - 4699 5686 0.7 - 8-0 433 -1981 2680 203 - 2130 2533 0 0 0 0 529 - 1810 2557 - 4252 9302 53 - 89 1 442 -1984 2698 1383 - 957 3726 0 0 0 756 - 1512 3010 2557 - 4252 9302 53 - 84 2 180 -2346 2650 561 - 1938 3067 163 - 2592 2418 290 - 1287 3580 1756 - 9428 12 326 35 - 188 20 -2719 2723 688 - 2033 3406 186 - 2278 2618 806 - 1940 3531 2844 - 8132 13 724 54 - 155 2 20 -2731 2632 2643 1084 - 1396 3565 124** - 2719 2584 748 - 1744 3219 2094 - 6819 10 679 46 - 138 2 20 -2311 2642 1188 - 1287 3665 105 - 2783 2668 839 - 1650 3305 2274 - 6361 10 679 46 - 1288 2 20 -2311 2642 1188 - 1287 3665 105 - 2783 2668 839 - 1650 3504 - 6819 10 679 46 - 1288 2 20 -2311 2642 1188 - 1287 3665 105 - 2783 2668 839 - 1650 3505 2274 - 6361 10 679 46 - 1288 2 20 -2311 2642 1188 - 1287 3665 105 - 2783 2668 839 - 1650 3505 2274 - 6361 10 679 46 - 1288 2 20 -2311 2642 1188 - 1287 3665 105 - 2783 2668 839 - 1650 3505 2274 - 6361 10 679 46 - 1288 2 20 -2311 2642 1188 - 1287 3665 105 - 2783 2668 839 - 1650 3505 2274 - 6361 10 679 46 - 1288 2 20 -2311 2642 1188 - 1287 3665 105 - 2783 2668 839 - 1650 3505 2274 - 6361 10 679 46 - 1288 2 20 -2311 2642 1188 - 1287 3665 105 - 2783 2668 839 - 1650 3505 2274 - 6361 10 679 46 - 1288 2 20 -2311 2642 1188 - 1287 3665 105 - 2783 2668 839 - 1650 3505 2274 - 6361 10 679 46 - 1288 2 20 -2311 2642 1188 - 1287 3665 105 - 2783 2668 839 - 1650 3505 2274 - 6361 10 679 46 - 1288 2 20 -2311 2642 1188 - 1287 3668 105 - 2783 2668 839 - 1650 3505 2274 - 6361 10 679 46 - 1288 2 20 -2311 2642 1188 - 1287 3668 10 | 600 | 8 | -1395 | 1487 | 72 | -1361 | 1509 | 678 | -1009 | 1860 | 99 | -1377 | 1500 | 330 | -5142 | 6356 | 9.0 | & & & | _ |
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| 92 -1359 1462 153 -1258 1566 293 -836 1204 188 -1234 1590 434 -4699 5686 0.7 -8.0 392 -1359 1462 153 -1258 1566 293 -836 1204 188 -1234 1590 434 -4699 5686 0.7 -8.0 433 -1981 2680 203 -2130 2533 | Average | 90 | -1365 | 1466 | 140 | -1275 | 1557 | 519** | -1172 | 1824 | 168 | -1258 | 1575 | 417 | -4773 | 5798 | 0.7 | -8.2 | _ |
| years)* 433 | Average | 92 | -1359 | 1462 | 153 | -1258 | 1566 | 293 | -836 | 1204 | 188 | -1234 | 1590 | 434 | -4699 | 2686 | 0.7 | -8.0 | _ |
| 433 -1981 2680 203 -2130 2533 0 0 0 529 -1810 2854 1165 -5921 8067 25 -125 | non-pandemic years)* | | | | | | | | | | | | | | | | | | |
| 433 -1981 2680 203 -2130 2533 0 0 0 529 -1810 2854 1165 -5921 8067 25 -125 149 -2138 2378 1652 -602 3914 0 0 0 756 -1512 3010 2557 -4252 9302 53 -89 442 -1984 2698 1383 -957 3726 0 0 0 1224 -1130 3551 3049 -4071 9975 63 -84 180 -2364 2650 561 -1938 3067 163 -2592 2418 290 -2216 2789 1194 -9110 10 924 24 -184 24 -2716 2789 1194 -9110 10 924 24 -184 24 -2716 2789 1194 -9110 10 924 24 -184 24 -2716 2789 1194 -9110 10 924 24 -185 9975 69 -84 185 9975 69 1975 69 | ged 65+ years | | | | | | | | | | | | | | | | | | |
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| 442 | 2007 | 149 | -2138 | 2378 | 1652 | -602 | 3914 | 0 | 0 | 0 | 756 | -1512 | 3010 | 2557 | -4252 | 9302 | 53 | -89 | 19 |
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| 205 —2311 2642 1188 —1287 3665 105 —2783 2668 839 —1650 3305 2274 —6361 10 679 46 —128 | Average | 201 | -2320 | 2643 | 1084 | -1396 | 3565 | 124** | -2719 | 2584 | 748 | -1744 | 3219 | 2094 | -6819 | 10 720 | 42 | -137 | 21 |
| non-pandemic years)* | Average | 205 | -2311 | 2642 | 1188 | -1287 | 3665 | 105 | -2783 | 2668 | 839 | -1650 | 3305 | 2274 | -6361 | 10 679 | 46 | -128 | 21 |
| | (non-pandemic years)* | | | | | | | | | | | | | | | | | | |

^{*}Average for non-pandemic years (excluded 2009). **Three-year average.

Table 3. Annual estimated number and rate of influenza-associated deaths for underlying P&I, respiratory, and circulatory deaths in Thailand, 2006–2011 (with the apportioned ill-defined deaths)

| | Influenza | l | | | | | | | |
|---------------------|------------------------|--------|--------|-----------------------|--------|-----|--|--|--|
| Cause of death | Number of deaths | 95% CI | | Rate per 1 000 000 | 95% CI | | | | |
| Respiratory of | disease | | | | | | | | |
| <65 years | 657 | -1788 | 3101 | 1.1 | -0.3 | 5.3 | | | |
| ≥65 years | 2094 | -4611 | 8801 | 42 | -93 | 178 | | | |
| All ages | 2751 | -6400 | 11 902 | 4.3 | -10.1 | 19 | | | |
| Pneumonia a | and influen: | za | | | | | | | |
| <65 years | 370 | -1256 | 2007 | 0.6 | -2.2 | 3.4 | | | |
| ≥65 years | 1031 | -3201 | 5259 | 21 | -64.7 | 106 | | | |
| All ages | 1401 | -4457 | 7266 | 2.2 | -7.0 | 11 | | | |
| Circulatory disease | | | | | | | | | |
| <65 years | 0 | -3363 | 3075 | 0.0 | -5.8 | 5.3 | | | |
| ≥65 years | 529 | -5934 | 6797 | 11 | -120 | 138 | | | |
| All ages | 529 | -9297 | 9871 | 0.8 | -14⋅7 | 16 | | | |

accounted for 74% (1031/1401), 76% (2094/2751), and 100% (529/529) of the overall estimated average annual influenza-associated deaths with underlying P&I, respiratory, and circulatory causes, respectively.

The average annual rate of influenza-associated deaths for adults aged ≥65 years was 21 per 100 000 persons for P&I deaths, 42 per 100 000 for respiratory deaths, and 11 per 100 000 for circulatory deaths (Table 3). Among adults aged <65 years, the average annual rate of influenza-associated deaths was 0·6 deaths per 100 000 persons for P&I causes and 1·1 per 100 000 for respiratory causes. The weekly number of deaths from 2006 to 2011, and excess deaths where mortality exceeded the baseline, is shown in Figure 1; there are no discrete peaks indicating influenza epidemics. In the sensitivity analysis of all-cause deaths, the average annual rate of influenza-associated deaths was 2·0 (95%CI: −24 to 27) per 100 000 population for person aged <65 years and 99 (95% CI: −322 to 516) for person aged ≥65 years.

Estimates of influenza-associated deaths without reapportion ill-defined death

Estimated mean annual influenza-associated deaths without reapportioning ill-defined death are shown in Table S3. The average annual number of influenza-associated deaths from underlying P&I, respiratory, and circulatory diseases was 632 (1·0 per 100 000), 1308 (2·1 per 100 000), and 295 (0·5 per 100 000, Table S4), respectively. The estimated number of influenza deaths from the model with reapportioned ill-defined deaths and from the non-reapportion model for persons aged <65 was similar, while for those age ≥ 65 years,

the estimates from reapportioned models were approximately 3 times higher than from the non-reapportion model. Compared with Figure 1, similar patterns were observed with no obvious peak for the excess circulatory and respiratory deaths and overall lower magnitude (Figure S1).

Discussion

Influenza was associated with excess death in Thailand, and the estimated excess mortality was considerable (average 4.0 per 100 000 persons), with 83% of the influenza-associated deaths among persons ≥65 years. The estimated annual excess morality in persons ≥65 years was 42 per 100 000 persons. These finding are similar, but slightly lower than those from a recent study in Thailand that used a Bayesian model to estimate excess mortality due to seasonal influenza; this study found an estimated 6.1 annual excess deaths per 100 000 population and 68 per 100 000 among person ≥60 years. 14 Reasons for the differences could be due to the use of different outcomes for the models; Copper and colleagues use all-cause mortality, whereas we focused on respiratory and circulatory deaths. In addition, it could also be the results of the slightly different years of data analyzed, the use of three instead of two age groups in the model, or the use of the Bayesian model.

Our estimates of annual influenza-associated underlying P&I, underlying respiratory and underlying circulatory deaths in Thailand for the all-ages group were two times lower than the estimates of excess influenza deaths observed in tropical Singapore,¹¹ subtropical Hong Kong,¹⁵ and temperate United States.¹⁶ When looking by age, we observe that the decreased rates in elderly persons account for these country-specific differences. Some of these differences may be explained by the different age distribution in these countries; Hong Kong has a higher percentage of the population ≥65 years (14%) compared with Thailand (9.8%) or Singapore (8.1%). ¹⁷ In Thailand, the proportion of influenza-associated deaths was higher among the elderly (persons ≥65 years), which is consistent with data from Hong Kong and the United States where >80% of influenzaassociated deaths occurred among the elderly. 15,16 However, our estimates for influenza-associated deaths in persons age ≥65 years were consistently lower than those in Singapore (46.9 per 100 000 for P&I, and 155.4 per 100 000 for circulatory and respiratory) and Hong Kong (39-3 per 100 000 for P&I and 102 per 100 000 for circulatory and respiratory). 11,15 We suspect that our rates in elderly may be an underestimation resulting from inaccurate coding of death; the high proportion of ill-defined causes of death in our data set support this theory. 13 Other factors that might have contributed to the differences in mortality impact between countries include geographical differences, age distribution, previous influenza exposure history of the

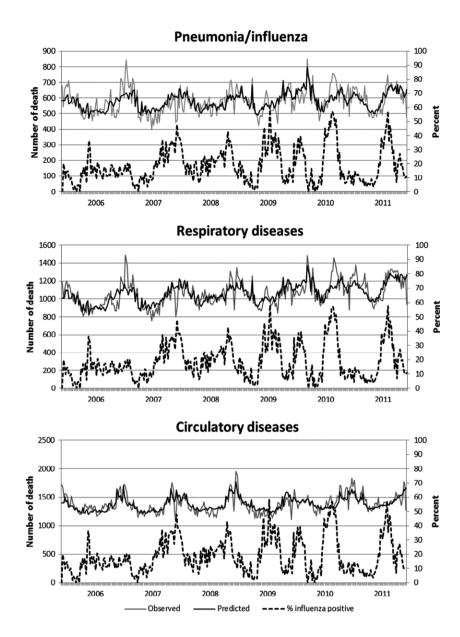


Figure 1. Weekly observed and predicted influenza-associated deaths from underlying P&I and respiratory and circulatory diseases (with the reapportioned ill-defined deaths).

population, influenza vaccination coverage, use of antiviral drugs, access to health care, and use of public health mitigation strategies. To address the issue of the high proportion of ill-defined causes of deaths in our data set, we performed a sensitivity analysis. We modeled all-cause deaths as outcome and our estimate (99 per 100 000, 95%CI: −322 to 516) for influenza-associated deaths in persons age ≥65 years was more comparable with the estimate in Singapore (167·8 (95% CI: 107·0−229·5) for all-cause) and the estimate in Hong Kong (136·1 (95% CI: 83·7−188·4) for all-cause).

Although there were peaks in influenza-associated mortality in Thailand during 2006 through 2011, they were not consistent in their timing. This is in contrast to trends

observed in temperate climates and likely reflects the year-round circulation of influenza viruses which may attenuate any seasonal mortality pattern. The majority of deaths from seasonal influenza occurred among people aged 65 year or older, but in the pandemic, the proportion of deaths among younger persons increased. This may suggest that there was some immunologic protection in persons who were exposed to A(H1N1) viruses before the 1957 pandemic, as has been demonstrated in other countries. 18,19

Our estimates of influenza-associated deaths have several methodological limitations. First, the quality of mortality statistics in Thailand was considered poor because a large proportion had a poorly defined cause of death. ^{13,20} Further, the validity of the cause of death statistics was questionable as

many out-of hospital deaths were coded by persons not medically qualified to determine the cause of death.²¹ Thus, misclassification of cause of death may have resulted in an underestimate of our specific mortality outcomes of interest, particularly in our excess mortality results that did not reapportion the ill-defined cause of death. Second, the relative distribution of influenza A and B viruses is known to vary by age, but the viral surveillance data were not robust enough to stratify by age and provide meaningful estimates by age group. Third, with only one cause of death listed, we were unable to assess contributing causes of death.²¹

These data are important to help guide the introduction of prevention strategies, such as seasonal influenza vaccination. Despite the widespread adoption of seasonal influenza vaccine recommendations in middle-income countries that have sufficient economic and public health resources to support vaccination programs, in Thailand seasonal influenza vaccine has been administered to <1% of the population annually and does not meet the need of the identified target groups.²² As vaccine use continues to increase, mortality should be monitored to assess the impact of the vaccination campaign.

Disclaimer

The findings and conclusions in this report are those of the authors and do not necessarily represent the views of the Centers for Disease Control and Prevention or the Thailand Ministry of Public Health.

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Supporting Information

Additional Supporting Information may be found in the online version of this article:

Figure S1. Weekly observed and predicted influenza-associated deaths from underlying P&I, respiratory and circulatory diseases (without the reapportioned ill-defined deaths).

Table S1. Model fitting procedure.

Table S2. Number of specimen tested and percent positive for influenza viruses by year, 2006–2011.

Table S3. Estimated average number of annual influenza-associated deaths with underlying P&I, respiratory, and circulatory causes by age group and location (without reapportion of ill-defined deaths).

Table S4. Estimated annual influenza-associated death rates in Thailand (without the reapportioned ill-defined deaths).